International Capital Flows and House Prices: Theory and Evidence*

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Abstract

The last fifteen years have been marked by a dramatic boom-bust cycle in real estate prices and accompanied by economically large fluctuations in international capital flows. We argue that changes in international capital flows played, at most, a small role in driving house price movements in this episode and that, instead, the key causal factor was a *financial market liberalization* and its subsequent reversal. Using observations on credit standards, capital flows, and interest rates, we find that a bank survey measure of credit supply, by itself, explains 53 percent of the quarterly variation in house price growth in the U.S. over the period 1992-2010, while it explains 66 percent over the period since 2000. By contrast, once we control for credit supply, various measures of capital flows, real interest rates, and aggregate activity–collectively–add less than 5% to the fraction of variation explained for these same movements in home values. Credit supply retains its strong marginal explanatory power controlling for expectations of future economic events and becomes even more important, relative to expectations, in predicting future house price movements. **Keywords**: housing boom and bust, global savings glut, foreign capital flows. **JEL**:F20,F32,G12,G21

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1 Introduction

The last fifteen years have been marked by a dramatic boom-bust cycle in real estate prices, a pattern unprecedented both in amplitude and in scope that affected many countries around the globe and all regions within the United States (Figure 1). Over the same period, there were economically large fluctuations in international capital flows. Countries that exhibited the largest house price increases also often exhibited large and increasing net inflows of foreign capital that bankrolled sharply higher trade deficits. Economists have debated the role of international capital flows in explaining these movements in house prices and asset market volatility more generally. A common hypothesis is that house price increases are positively related to a rise in the country's net foreign inflows, either because they directly cause house price increases (perhaps by lowering real interest rates), or because other factors simultaneously drive up both house prices and capital inflows. In this article, we study both theory and evidence that bears on this hypothesis, focusing on the unprecedented boom-bust cycle in housing markets that took place over the last 15 years.

We argue that changes in international capital flows played, at most, a small role driving house price movements in this episode and that, instead, the key causal factor was a *financial market liberalization* and its subsequent reversal that took place in many countries largely independently of international capital flows. Financial market liberalization (FML hereafter) refers to a set of regulatory and market changes and subsequent decisions by financial intermediaries that made it easier and less costly for households to obtain mortgages, borrow against home equity, and adjust their housing consumption.

By contrast, net capital flows into the United States over both the boom and the bust period in housing have followed a separate path, driven to great extent by foreign governments' regulatory, reserve currency, and economic policy motives. Consider the value of foreign holdings of U.S. assets minus U.S. holdings of foreign assets, referred to hereafter as *net foreign asset holdings* in the U.S., or alternatively, as the U.S. *net liability position*. A positive change in net foreign asset holdings indicates a capital inflow, or more borrowing from abroad.[1] As we show below, from 1994 to 2010, only the change in net foreign

[1] What we have defined as net foreign asset holdings, or the U.S. net liability position, is equal to the negative of the U.S. *net international investment position* in the U.S. Bureau of Economic Analysis balance of payments system. A country's resource constraint limits its expenditures on (government and private) consumption and investment goods, fees, and services, to its domestic output *plus* the change in the market value of its net liabilities (*minus* the change in the net international investment position). Thus a country's ability to spend in excess of domestic income in a given period depends positively on the change in its net

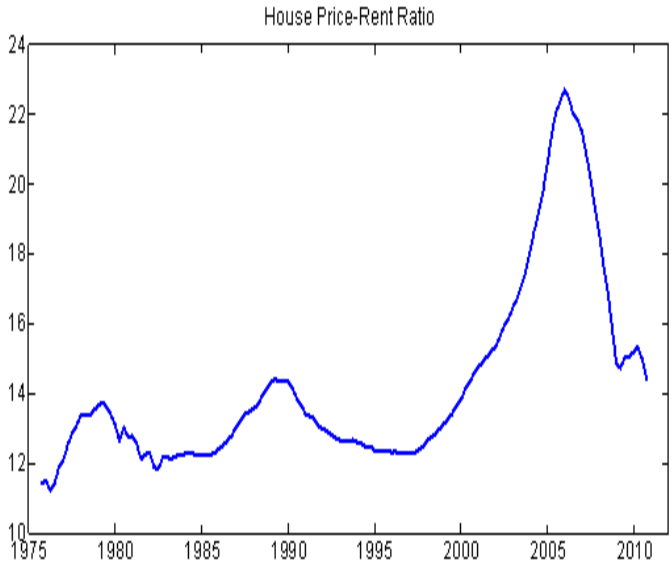


Figure 1: Price-Rent Ratio in the United States. The figure plots an aggregate price-rent ratio index for the United states from 1975:Q4-2010:Q4. Rent is rent for primary residence, constructed from the Shelter component of the Consumer Price Index for all urban consumers, SA, last month of each quarter. Data available from the Bureau of Economic Analysis of the U.S. Department of Commerce. Price is the CoreLogic National House Price Index (SA, Jan.2000=100). The price-rent ratio has been normalized to equal the level, in 1975:Q4 of the quarterly Price-Rent ratio constructed from the flow of funds housing wealth and National Income and Products data on housing consumption.

holdings of U.S. *securities* (equities, corporate, U.S. Agency and Treasury bills and bonds) show any discernable upward trend. Moreover, among securities, the upward trend has been driven almost entirely by an increase in net foreign holdings of U.S. assets considered to be safe stores-of-value, specifically U.S. Treasury and Agency debt. Yet inflows into these securities, rather than declining during the housing bust, have on average continued to increase. Importantly, foreign demand for U.S. "safe" assets is dominated by Foreign Official Institutions, namely government entities that have specific regulatory and reserve currency motives for holding U.S. Treasuries and other U.S.–backed assets, and that face both legal and political restrictions on the type of assets that can be held (Kohn (2002)). Such entities take extremely inelastic positions, implying that when these holders receive funds to invest, they buy U.S. Treasuries regardless of price (Krishnamurthy and Vissing-Jorgensen (2010)).

foreign liabilities.

We interpret these recent events through the lens of the theoretical models in Favilukis, Ludvigson, and Van Nieuwerburgh (2009), henceforth FLVNa, and Favilukis, Ludvigson, and Van Nieuwerburgh (2011), henceforth FLVNb. These papers study the economic consequences of both the U.S. FML (and its reversal) and, at the same time, empirically calibrated fluctuations in net U.S. capital inflows into the riskless bond market. The model environment is a two-sector general equilibrium framework with housing and non-housing production where heterogeneous households face uninsurable idiosyncratic and aggregate risks. Given the assets available in the model economy and collateralized financing restrictions, individuals can only imperfectly insure against both types of risk. We argue that these frameworks can account for the observed boom-bust pattern in house prices simultaneously with the continuing trend towards greater net capital inflows into U.S. securities over both the boom and the bust.[2] Fluctuations in the model's price-rent ratio are driven by changing risk premia, which vary endogenously in response to cyclical shocks, the FML, and its subsequent reversal. In FLVNa, house prices rise in the boom period because of a relaxation of credit constraints and decline in housing-related transactions costs, both of which reduce risk premia. Conversely, the reversal of the FML raises housing risk premia and causes the housing bust.

In contrast to the FML, an inflow of foreign money into domestic bond markets plays a small role in driving home prices in the models of FLVNa and FLVNb, despite its large depressing influence on interest rates. The reason is that a capital inflow into the safe bond market–by itself–*raises* risk premia on housing and equity, as domestic savers are forced out of the safe bond market and into risky securities. (We emphasize the words 'by itself' here because this increase in risk premia is more than offset by the simultaneous *decline* in risk premia caused by the FML during the boom.) At the same time, the capital inflow stimulates residential investment and an expected increase the housing stock. So while low interest rates in isolation tend to raise home prices, these factors tend to reduce them, thereby limiting the scope for a capital inflow to increase home prices. It follows that the sharp rise in price-rent ratios during the boom period must be attributed to an overall decline in risk premia and *not* to a fall in interest rates. Many alternative theories that can account for the positive correlation between house prices and capital inflows in the boom period are not able to explain the bust period, in which house prices collapsed but inflows into countries

[2]There was considerable volatility in the changes of net foreign asset holdings in the U.S. during the financial crisis in 2008 and 2009. Nevertheless, we show below that the changes in holdings were still higher at the end of the sample in 2010 than they were at the peak of the housing boom in 2006.

like the U.S. continued.

By FML we mean an outward shift in the availability of credit, at any given initial level of credit demand and borrower quality. This includes, as in the U.S. housing boom, an increase in maximal loan-to-value (LTV) ratios (e.g., the fraction of loans with combined–first and second–mortgage LTV ratios above 80% or 90%), an increase in the availability of new mortgage contracts (option-ARMs, interest only and negative amortization loans, loans to households with low FICO scores), a reduction in documentation requirements (asset and income verification), a rapid increase in the use of private-label securitization, and a reduction in fees (as well as in time and effort) associated with refinancing a mortgage or obtaining a home equity line of credit. The widespread relaxation of credit standards is well documented (see discussion below). Consistent with this evidence, microeconomic evidence in Mian and Sufi (2009) shows that mortgage credit expansion and house price growth in the boom were concentrated in areas with a large fraction of subprime mortgages and securitization of these mortgages, and *not* in areas with improved/ing economic prospects. Thus, this component of credit availability to households–accompanied by government deregulation of financial institutions and widespread changes in the way housing assets were financed and traded–appears to have fluctuated, to great extent, independently of current and future economic conditions.

But credit availability can also change endogenously in response to fluctuations in the aggregate economy and to revisions in expectations about future economic conditions, including house price growth. This information is reflected immediately in collateral values that constrain borrowing capacity. As in classic financial accelerator models (e.g., Bernanke and Gertler (1989), Kiyotaki and Moore (1997)), *endogenous* shifts in borrowing capacity imply that economic shocks have a much larger effect on asset prices than they would in frictionless environments without collateralized financing restrictions. Both exogenous and endogenous components of time-varying credit availability to households are operative in the model of FLVNa.

While endogenous fluctuations in credit availability are clearly important in theory, it is unclear how quantitatively important they have been empirically, especially in the recent housing boom-bust episode. Some researchers have argued that credit availability is primarily driven by the political economy, and in particular by political constituencies that influence bank regulation related to credit availability (e.g., Rajan and Ramcharan (2011b)). Such a component to credit availability could in fact be independent of economic fundamentals, expectations of future fundamentals, and credit demand. Rice and Strahan (2010) construct

a time-varying index capturing state-level differences in regulatory constraints on interstate branching of banks and find that they correlate with the lobbying power of small banks relative to expansion-minded banks, but not with contemporaneous economic conditions.

Using observations on credit standards, capital flows, and interest rates for the U.S. and for a panel of 11 countries, we present evidence on how these variables are related to real house price movements in recent data. Our main measure of credit standards is compiled from quarterly bank surveys of senior loan officers, carried out by national central banks as part of their regulatory oversight. We consider this a summary indicator of fluctuations in the variables associated with a FML, as described above. The surveys specifically address changes in a bank's *supply* of credit, as distinct from changes in its perceived demand for credit. We find for the U.S. that this measure of credit supply, by itself, explains 53 percent of the quarterly variation in house price growth over the period 1992-2010, while it explains 66 percent over the period since 2000. By contrast, controlling for credit supply, various measures of capital flows, real interest rates, and aggregate activity–collectively–add less than 5% to the fraction of variation explained for these same movements in home values. Moreover, credit standards continues to be the most important variable related to *future* home price fluctuations even when it has been rendered statistically orthogonal to banks' perceptions of credit demand, and even when controlling for expected future economic growth and expected future real interest rates. Taken together, these findings suggest that a stark shift in bank lending practices–conspicuous in the FML and its reversal–were at the root of the housing crisis.

The rest of this paper is organized as follows. The next section discusses theoretical literature that has addressed the link between house prices, capital flows and/or credit supply. To provide a theoretical frame of reference, here we also describe in detail the predictions of FLVNa for house price movements. Section 3 turns to the data, presenting stylized facts on international capital flows, interest rates and credit standards. Section 4 presents an empirical analysis of the linkage between capital flows and house price fluctuations, controlling for measures of credit supply, economic activity, and real interest rates. Section 5 concludes. Section 6 is an Appendix that provides details on the data we use and on our estimation methodology.

2 Theories

A number of studies have addressed the link between house prices and capital flows, focusing on the recent boom period in housing. For brevity, we will refer to the period of rapid home price appreciation from 2000 to 2006 as the *boom* period in the U.S., and the period 2007 to present as the *bust*.

The *global savings glut* hypothesis Bernanke (2005), Mendoza, Quadrini, and Rios-Rull (2007), Bernanke (2008), Caballero, Fahri, and Gourinchas (2008), Caballero and Krishnamurthy (2009) contends that the excess savings of developing countries, notably China and emerging Asia, sought safe, high-quality financial assets that their own economies could not provide. Because of the depth, breadth, and safety of U.S. Treasury and Agency markets, those savings predominantly found their way to the United States. Some have directly linked these patterns to higher U.S. home prices, arguing that low interest rates (driven in part by the capital inflow) were a key determinant of higher house prices during the boom (e.g., Bernanke (2005), Himmelberg, Mayer, and Sinai (2005), Bernanke (2008), Taylor (2009), Adam, Marcet, and Kuang (2011)). In a similar spirit, Caballero and Krishnamurthy (2009) identify the start of the housing boom with the Asian financial crisis which fueled the demand for U.S. risk-free assets. In their model, Asian savers turn to U.S. assets, resulting in a net capital inflow for the U.S. Global interest rates then fall in their model because the U.S. economy is presumed to grow more slowly than the rest of the world.

Laibson and Mollerstrom (2010) have criticized the global savings glut hypothesis by noting that an increase in world-wide savings should have led to an investment boom in countries that were large importers of capital, notably the U.S. Instead, the U.S. experienced a consumption boom that accompanied the housing boom, suggesting that savings world-wide was not unusually high. Laibson and Mollerstrom (2010) present an alternative interpretation of the correlation between home values and capital flows during the boom based on asset bubbles. Assuming a bubble in the housing market, they argue that the rise in housing wealth generated by the bubble led to higher consumption, which in turn led to greater borrowing from abroad and a substantial net capital inflow to the U.S. A similar idea is presented in Ferrero (2011), but without the bubble. Ferrero studies a two-sector representative-agent model of international trade in which lower collateral requirements facilitate access to external funding and drive up house prices.

Others have argued that preference shocks and a desire for smooth (across goods) consumption can generate a correlation between house prices and capital inflows. Gete (2010)

shows that consumption smoothing across tradeable (non-housing) goods and nontradeable (housing) goods can lead to a positive correlation between houses and current account deficits. With an exogenous increase in the home country's preference for housing, productive inputs in the home country are reallocated toward housing production, so that housing consumption can rise. But with a preference for smooth consumption across goods, the tradeable non-housing good (presumed identical across countries) will then be imported from abroad, leading to capital inflows to the home country.

The theories above fall into two broad categories: those that rely on higher domestic demand to drive both house prices and capital inflows in the same direction (Gete (2010), Laibson and Mollerstrom (2010), Ferrero (2011)), and those that rely on capital inflow-driven low interest rates to drive up house prices (Bernanke (2005), Himmelberg, Mayer, and Sinai (2005), Bernanke (2008), Taylor (2009), Caballero and Krishnamurthy (2009), Adam, Marcet, and Kuang (2011)). While these papers were motivated by observations on housing and capital flows during the housing *boom*, they also have implications for the housing bust. The former imply that the housing bust should be associated with a reversal of domestic demand, leading to a capital outflow. The latter imply that the housing bust should be associated with a rise in real interest rates, in many cases correlated with a capital outflow.

As we show below, recent data pose number of challenges to these theories. First, while its true that real interest rates were low throughout the boom period, they have remained low in the bust period. Second, while capital certainly flowed into countries like the U.S. during the boom period, there is no evidence of a clear reversal in this trend during the bust period.[3] These observations suggest that the economic and political forces responsible for driving capital flows and house prices over the entire period were, to a large extent, distinct. Below we present empirical evidence that neither capital inflows nor real interest rates bear a strong relation to house prices in a sample that includes both the boom and the bust.

We interpret these recent events through the lens of the theoretical models in FLVNa and FLVNb, focusing specifically on the model in FLVNa in which a FML and its reversal are studied. Rather than reproducing the mathematical description of the model here, we simply describe it verbally and refer the reader to the original papers for details. Our focus here is on empirical evidence relating home prices to various indicators as a means of distinguishing

[3]Some empirical studies document a positive correlation between house prices and capital inflows to the U.S., but these studies typically have data samples that terminate at the end of the boom or shortly thereafter (e.g., Aizenman and Jinjarak (2009), Kole and Martin (2009)).

among theories. Next we describe the model in FLVNa, and explain how it differs from the theories above.

2.1 The Housing Boom-Bust: A Theory of Time-Varying Risk-Premia

FLVNa study a two-sector general equilibrium model of housing and non-housing production where heterogenous households face limited risk-sharing opportunities as a result of incomplete financial markets. A house in the model is a residential durable asset that provides utility to the household, is illiquid (expensive to trade), and can be used as collateral in debt obligations. The model economy is populated by a large number of overlapping generations of households who receive utility from both housing and nonhousing consumption and who face a stochastic life-cycle earnings profile. We introduce market incompleteness by modeling heterogeneous agents who face idiosyncratic and aggregate risks against which they cannot perfectly insure, and by imposing collateralized borrowing constraints on households.

Within the context of this model, FLVNa focus on the macroeconomic consequences of three systemic changes in housing finance, with an emphasis on how these factors effect risk premia in housing markets, and how risk premia in turn affect home prices. First, FLVNa investigate the impact of changes in housing collateral requirements.[4] Second, they investigate the impact of changes in housing transactions costs. Taken together, these two factors represent the theoretical counterpart to the real-world FML discussed above. Third, FLVNa investigate the impact of an influx of foreign capital into the domestic bond market. FLVNa argue that all three factors fluctuate over time and changed markedly during and preceding the period of rapid home price appreciation from 2000-2006, and the subsequent bust. In particular, the boom period was marked by a widespread relaxation of collateralized borrowing constraints and declining housing transactions costs (including costs associated with mortgage borrowing, home equity extraction, and refinance). The period was also marked by a sustained depression of long-term interest rates that coincided with a vast

[4]Ferrero (2011) also assumes a relaxation of credit constraints to explain the housing boom. A key distinction between his model and FLVNa, however, is that Ferrero studies a two-country representative agent model, so an increase in borrowing by the domestic agent is only possible with increase in lending from rest of the world, hence a higher current account deficit. By contrast, in FLVNa, borrowing and lending can happen within the domestic economy between heterogeneous agents, so housing finance need not be tied to foreign savings. Thus, a reversal of the FML in a setting like that of Ferrero's would necessitate a capital outflow, whereas in FLVNa it does not.

inflow of capital into U.S. safe bond markets. In the aftermath of the credit crisis that began in 2007, the erosion in credit standards and transactions costs has been sharply reversed.[5] We provide evidence on this below.

The main impetus for rising price-rent ratios in the model in the boom period is the simultaneous occurrence of positive economic shocks and a financial market liberalization, phenomena that generate an endogenous decline in risk premia on housing and equity assets. As risk premia fall, the aggregate house price index relative to aggregate rent, rises. A FML reduces risk premia for two reasons, both of which are related to the ability of heterogeneous households to insure against aggregate and idiosyncratic risks. First, lower collateral requirements directly increase access to credit, which acts as a buffer against unexpected income declines. Second, lower transactions costs reduce the expense of obtaining the collateral required to increase borrowing capacity and provide insurance. These factors lead to an increase in risk-sharing, or a decrease in the cross-sectional variance of marginal utility. The housing bust is caused by a reversal of the FML and of the positive economic shocks and an endogenous decrease in borrowing capacity as collateral values fall. These factors lead to an accompanying rise in housing risk premia, driving the house price-rent ratio lower. Almost all of the theories discussed above are silent on the role of housing risk premia in driving house price fluctuations.[6]

It is important to note that the rise in price-rent ratios caused by a financial market liberalization in FLVNa must be attributed to a decline in risk premia and not to a fall in interest rates. Indeed, the very changes in housing finance that accompany a financial market liberalization drive the endogenous interest rate up, rather than down. It follows that, if price-rent ratios rise after a financial market liberalization, it must be because the decline in risk premia more than offsets the rise in equilibrium interest rates that is attributable to the FML. This aspect of a FML underscores the importance of accounting properly for the role of foreign capital over the housing cycle. Without an infusion of foreign capital, any period of looser collateral requirements and lower housing transactions costs (such as that which characterized the housing boom) would be accompanied by an *increase* in equilibrium interest rates, as households endogenously respond to the improved risk-sharing opportunities afforded by a financial market liberalization by reducing precautionary saving.

[5] Streitfeld (2009) argues that, since the credit crisis, borrowing restrictions and credit constraints have become even more stringent than historical norms in the pre-boom period.

[6] An exception is Caballero and Krishnamurthy (2009), but they do not study housing or the FML and it's reversal.

To model capital inflows, FLVNa introduce foreign demand for the domestic riskless bond into the market clearing condition. This foreign capital inflow is modeled as driven by governmental holders who inelastically place all of their funds in domestic riskless bonds. Foreign governmental holders have a perfectly inelastic demand for safe securities and place all of their funds in those securities, regardless of their price relative to other assets. Below we discuss data on U.S. international capital flows that supports this specification of the net capital flows in the United States over the last 15 years.

The model in FLVNa implies that a rise in foreign purchases of domestic bonds, equal in magnitude to those observed in the data from 2000-2010, leads to a quantitatively large decline in the equilibrium real interest rate. An important feature of this model is that a capital inflow may not have a large effect on house prices *even if* it has a large effect on interest rates. By contrast, in models where risk premia are held fixed, a decline in the interest rate of this magnitude would be sufficient–by itself–to explain the rise in price-rent ratios observed from 2000-2006 under reasonable calibrations. But with time-varying housing risk premia, borrowed sums from the rest of the world can play at most a limited role in asset booms, despite their large depressing influence on interest rates. The reason is that foreign purchases of U.S. bonds crowd domestic savers out of the safe bond market, exposing them to greater systematic risk in equity and housing markets. In response, risk premia on housing and equity assets rise, substantially offsetting the lower interest rates and limiting the impact of foreign capital inflows on home prices.

It is useful to clarify the two opposing forces simultaneously acting on housing risk premia in the model of FLVNa. During the housing boom, there is both a FML and a capital inflow. As explained, the FML *lowers* risk premia, while foreign purchases of domestic safe assets *raise* risk premia. Under the calibration of the model, the decline in risk premia resulting from the FML during the boom period is far greater than the rise in risk premia resulting from the capital inflow. On the whole, therefore, risk premia on housing assets fall, and this is the most important contributing factor to the an increase in price-rent ratios during the boom. During the bust, modeled as a reversal of the FML but not the capital inflows, risk premia unambiguously rise even as interest rates remain low. The rise in risk premia drives the decline in house-price rent ratios.

These features of the model represent significant differences from other theories of capital flows and house prices. They permit the model to explain not just the housing boom, but also the housing bust, in which house price-rent ratios fell dramatically even though interest rates remained low and there has been no clear reversal in the trend toward capital inflows

into the U.S. bond market. Moreover, they underscore the importance of distinguishing between interest rate changes (which are endogenous) and credit supply. In the absence of a capital inflow, an expansion of credit supply in the form of lower collateral requirements and lower transactions costs should lead, in equilibrium, to higher interest rates, rather than lower, as households respond to the improved risk-sharing/insurance opportunities by reducing precautionary saving. Instead we observed low real interest rates, generated in the model of FLVNa by foreign capital inflows, but the inflows themselves are not they key factor behind the housing boom-bust.

To illustrate the independent role of house prices and capital inflows in the model, Figure 2 plots the transition dynamics for both the aggregate price-rent ratio and for foreign holdings of domestic assets over the period 2000-2010 from the model of FLVNa. The figure shows the dynamic behavior of the price-rent ratio in response to a series of shocks designed to mimic both the state of the economy and housing market conditions over the period 2000-2010. The economy begins in year 2000 the stochastic steady state of a world with "normal" collateral requirements (i.e., fraction of home value that must be held as collateral) and housing transactions costs calibrated to roughly match the data prior to the housing boom of 2000-2006. In 2001, the economy undergoes an unanticipated shift to a new steady state, in which there is a FML with lower collateral requirements and lower transactions costs, calibrated to match the changes in these variables during the boom period, as well as an unanticipated increase in foreign holdings of U.S. bonds from 0 to 18% of GDP. This 18% increase is calibrated to match the actual increase in net foreign holdings of U.S. securities over the period 2000-2010. Along the transition path, foreign holdings of bonds are increased linearly from 0% to 18% of GDP from 2000 to 2010. The adjustment to the new stochastic steady state is then traced out over the seven year period from 2001 to 2006, as the state variables evolve. Finally, starting in 2007 and continuing through 2010, the economy is presumed to undergo a surprise reversal of the financial market liberalization but not the foreign capital inflow.

Figure 2 shows that the house-price rent ratio rises by roughly 30% over the period 2000-2006 and then falls by three-quarters of this over the period 2007-2010. By contrast, foreign holdings of domestic riskless bonds, denoted B_t^F, rise at a constant rate throughout the boom-bust period. Although foreign holdings rise mechanically over time and are crudely calibrated to match the total increase in holdings over the entire 10-year period (rather than matching the year-by-year fluctuations), the figure shows clearly that capital flows are not a key determinant of the boom-bust pattern in the price-rent ratio in this model, despite the

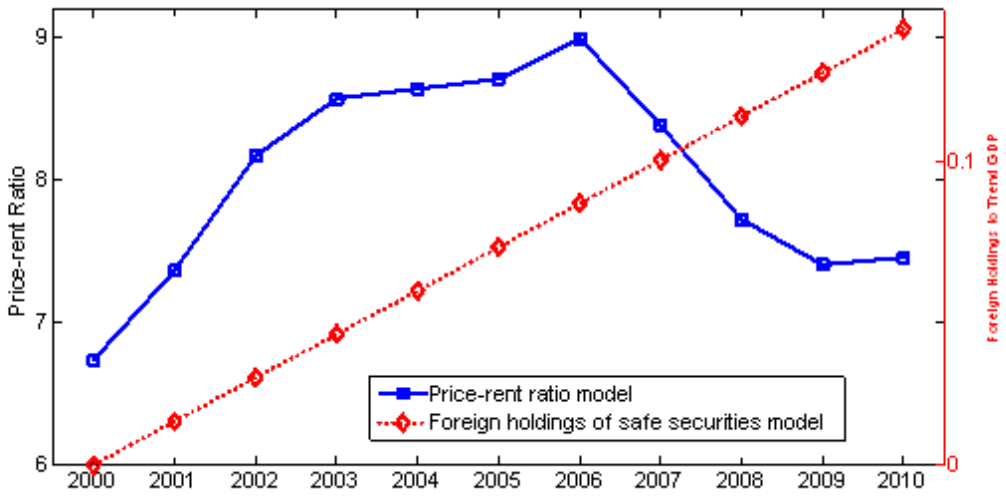


Figure 2: Price-Rent Ratio and Foreign Holdings in FLVNa. The figure plots the transition dynamics of the aggregate price-rent ratio and foreign holdings of the safe asset in the model of Favilukis, Ludvigson, and Van Nieuwerburgh (2009), for the period 2000-2010. The dynamics are driven by a sequence of aggregate economic shocks designed to mimic the business cycle over this period, a financial market liberalization in 2000, and a reversal of this liberalization in 2007, as well as the foreign flows depicted in the figure. Source: Favilukis, Ludvigson, and Van Nieuwerburgh (2009).

large decline in interest rates generated by these inflows.

The relationship between capital inflows and risk premia in FLVNa and FLVNb is worthy of emphasis. In equilibrium, higher capital inflows into the safe bond market *raise* risk premia on housing and equity, rather than lower them. This runs contrary to the argument, made by some, that the free flow of capital across borders should be associated with a reduction in risk premia (e.g., Geithner (2007)). Here, foreign purchases of the safe asset make both equity and housing assets *more risky*. Both the risk premium and Sharpe ratio for equity and housing rise when there is a capital inflow, for two reasons. First, the increase in foreign money forces domestic residents as a whole to take a leveraged position in the risky assets. This by itself increases the volatility of asset and housing returns, translating into higher risk premia. Second, domestic savers are crowded out of the bond market by foreign governmental holders who are willing to hold the safe asset at any price. As a result, they become more exposed to systematic risk in the equity and housing markets. This means that the equity and housing Sharpe ratios must rise, as domestic savers shift the composition of their financial wealth towards risky securities. In addition, the volatility of the stochastic

discount factor rises and there is a decrease in risk-sharing, as measured by the cross-sectional variance of marginal utility growth.

Of course, the affect of a capital inflow on house prices depends not only on the housing risk premium, but also on the risk-free interest rate. Although a capital inflow drives the housing risk premium up, in the model of FLVNa it drives the risk-free rate down by more, so a capital inflow still leads to a modest increase in the price-rent ratio.[7] In this model, an inflow of foreign capital calibrated to match the increase in foreign ownership of U.S. Treasuries and U.S. agency debt over the period 2000-2010 has a large downward impact on the equilibrium interest rate, which falls from 3.45% to 0.39%. The magnitude of this decline is close to the reduction in real rates observed in U.S. bond market data over the period 2000-2006.

With this discussion as theoretical background, we now turn to an analysis of the data on capital flows, interest rates, and credit standards over the boom-bust period.

3 Trends in Capital Flows, Interest Rates, Credit Supply

While the notion of a global savings glut is controversial, recent data clearly suggest a reallocation of savings away from the developed world, and toward the developing world, the so-called *global imbalances* phenomenon. Unlike any prior period, global financial integration allowed for the channeling of one country's excess savings towards another country's real estate boom. Such financing occurred directly, for example by German banks' purchases of U.S. subprime securities, but also indirectly through the U.S. Treasury and Agency bond markets. As the world's sole supplier of a global reserve currency, the U.S. experienced a surge in foreign ownership of U.S. Treasuries and Agency bonds. Agency bonds refers to the debt of the two government-sponsored enterprizes (GSEs) Freddie Mac and Fannie Mae, as well as to the mortgage-backed securities that they issue and guarantee. Due to their ambivalent private-public structure and their history as agencies of the federal government, private market investors (including foreign investors) have always assumed that the debt of Freddie Mac and Fannie Mae was implicitly backed by the U.S. Treasury. That implicit backing became an explicit backing in September 2008 when Freddie Mac and Fannie Mae

[7] Changes in expected future aggregate rent growth also can effect the price-rent ratio. The numbers here refer to a comparison of stochastic steady states, however, in which the expected rental growth rate is the same in both steady states (equal to the deterministic growth rate of the economy).

were taken into government conservatorship. See Acharya, Richardson, Van Nieuwerburgh, and White (2011) for details on the GSEs.

In this section, we discuss in detail data showing the trends in capital flows, U.S. real interest rates, and the relaxation and subsequent tightening of housing credit constraints and credit standards.

3.1 International Capital Flows

The Treasury International Capital Reporting System (TIC) is the official source of U.S. securities flows data. It reports monthly data (with a six week lag) on foreigners purchases and sales of all types of financial *securities* (equities, corporate, Agency, and Treasury bonds). We refer to these monthly transactions data as the *TIC flows data*. The TIC system also produces periodic benchmark surveys of the market value of foreigners' net *holdings*, or net asset *positions*, in U.S. securities. Unlike the flows data, these data take into account the net capital gains on gross foreign assets and liabilities. We refer to these as the *TIC holdings data*. The holdings data are collected in detailed surveys conducted in December of 1978, 1984, 1989, and 1994, in March 2000, and annually in June from 2002 to 2010. The survey data on holdings is thought to be of higher quality than the flows data because it more accurately accounts for valuation effects (Warnock and Warnock (2009)).[8]

The Bureau of Economic Analysis (BEA) in the U.S. Department of Commerce also provides annual estimates of the value of accumulated stocks (holdings) of U.S.-owned assets abroad and of foreign-owned assets in the United States. We will refer to these as the *BEA holdings data*. These include estimates of holdings of securities, based on the TICs data, as well as estimates of holdings of other assets such as foreign direct investment, U.S. official reserves and other U.S. government reserves. We refer to the sum of these other assets plus financial securities as *total assets*. In recent data, the main difference between the BEA estimate of net foreign holdings of total assets and its estimate of net foreign holdings of total securities is attributable to foreign direct investment (FDI), where, since 2006, the value of U.S. FDI abroad has exceeded the value of foreign FDI in the U.S.[9]

The BEA defines the U.S. *net international investment position* (NIIP) as the value of U.S.-owned assets abroad minus foreign-owned assets in the U.S. The overall change in the

[8] As explained in Warnock and Warnock (2009), reporting to the surveys is mandatory, with penalties for non-compliance, and the data are subjected to extensive analysis and editing. Data on foreign holdings of U.S. securities are available at http://www.treasury.gov/resource-center/data-chart-center/tic/Pages/index.aspx.

[9] These data are available at http://www.bea.gov/international/index.htm.

NIIP incorporates capital gains and losses on the prior stock of holdings of assets. Thus, the total change in U.S. gross foreign assets equals net purchases by U.S. residents plus any capital gains on the prior stock of gross foreign assets, while the total change in U.S. foreign liabilities equals net sales of assets to foreign residents plus any capital gains accrued to foreigners on their U.S. assets. The NIIP is the difference between the two. Capital gains are the most important component of valuation changes on the NIIP.

The BEA also collects quarterly and annual estimates of *transactions* with foreigners, including trade in goods and services, receipts and payments of income, transfers, and transactions in financial assets. We refer to these as the *BEA transactions data*. The transactions data measure the current account (CA). Since the CA transactions data only measure purchases and sales of assets, they do not adjust for valuation effects that must be taken into account in constructing the international investment *positions* (holdings) of the U.S., as just discussed.[10]

When thinking about the recent boom-bust period in residential real estate, a question arises as to which measure of capital flows to study. Obstfeld (2011) documents an increase in the sheer volume of financial trade across borders, and argues that it could be positively correlated with financial instability. Moreover, he shows that the amplitude of pure valuation changes in the NIIP has grown in tandem with the volumes of gross flows. Because the CA ignores such valuation changes, our preferred measure would therefore be a measure of total changes in net foreign *holdings* of assets rather than changes in net *transactions*. Unfortunately, data on net foreign asset holdings are only readily available in the U.S., and then only annually. (For the empirical work below, we construct our own quarterly estimate of these holdings for securities.) Outside the U.S., only the transactions-based CA data are available. Thus, when we use international data we use the CA as a measure of capital flows, bearing in mind the limitations of these data for measuring changes in actual asset holdings.

Since net foreign asset holdings data are available for the U.S., when working with U.S. data we focus most on net foreign holdings as a measure of capital flows (although for completeness we also present empirical results using the CA as a measure of capital flows). Within net foreign holdings, we focus on changes in holdings of financial *securities*, rather than changes in holdings of total *assets*. We argue that the former are far more relevant for residential real estate than the latter. Recall that the most important difference between the two, especially in recent data, results from flows in FDI. But it is unclear how relevant FDI is for the housing market. For example, during much of the housing boom, the value of net

[10] See the adjustments for valuations effects at http://www.bea.gov/international/xls/intinv10_t3.xls.

foreign holdings on FDI fell, implying a net capital *outflow* on those types of assets. This fact is hardly consistent with the notion that capital *inflows* to the U.S. helped finance the housing boom.

What flowed *in* during the housing boom was foreign capital directed at U.S. Treasuries and Agency securities. There are several reasons we expect these assets–unlike FDI–to be directly related to the U.S. housing market. First, foreign purchases of Agency securities allow U.S. banks to broaden their market for mortgage-backed securities (MBS) to international investors, instead of funding the mortgage investments themselves. Second, an inflow of capital into U.S. Agencies can in turn free up U.S. banks to fund additional, possibly riskier, mortgages. Third, because mortgage rates are often tied to Treasury rates, large foreign Treasury purchases could in principle directly affect house prices through their effect on interest rates. Fourth, low Treasury rates could lead U.S. banks in search of yield to undertake more risky mortgage investments (see Maddaloni and Peydro (2011) for evidence that banks do so). In summary, because the FDI streams are largely divorced from the U.S. housing market, the most appropriate measure capital flows for our purpose is not net foreign holdings of total assets but instead total securities.

Figure 3 shows the movement in various measures of international capital flows into the United States, relative to trend GDP, in annual data from 1976 to 2010. Plotted are the *change* in net foreign holdings of total assets, total securities, and in what we will call U.S. "safe" securities (defined as Treasuries and Agencies). We refer to a capital inflow as a positive change in holdings, and vice versa for a capital outflow. Also plotted is the current account deficit. 3 shows that there is considerable volatility in these measures during housing boom and the subsequent financial crisis, with particularly sharp increases in the change in net foreign holdings of U.S. assets from 2007 to 2008. This corresponds to a spike in the U.S. net foreign liability position in 2008. It then declines from 2008 to 2009 and increases again from 2009 to 2010. Comparing the end-points of these series in 2010 to their values in 2006, we see that–by any measure of assets–inflows (or the change in holdings) were higher at the end of the sample in 2010 than they were at the peak of the housing boom in 2006.

To get a better sense of the trends in these series, Figure 4 plots the same measures of international capital flows, but computed as four-year moving averages. The figure shows that changes in net foreign holdings of total assets show little trend movement over the entire boom-bust period 2000-2010, but if anything they trended *up* during the *bust* period from 2006 to 2010, while they trended *down* in the *boom* from 2002 to 2006. A similar pattern holds for net foreign holdings of total securities, except that here, inflows are much more

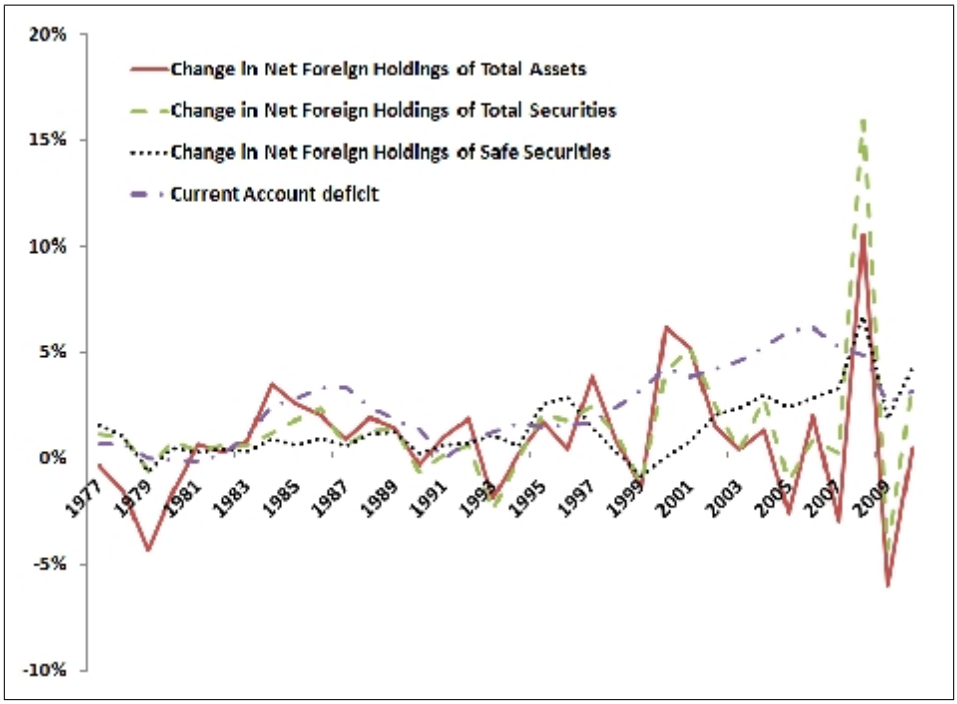


Figure 3: Measures of U.S. Capital Flows. Source: U.S. Department of Commerce, Bureau of Economic Analysis. Net foreign holdings of total assets is foreign-owned assets in the United States minus U.S.-owned assets abroad. Net foreign holdings of total securities is defined as foreign-owned U.S. government securities plus U.S. Treasury securities plus U.S. securities other than Treasury securities minus U.S.-owned foreign securities. Net foreign holdings of safe securities is defined as U.S. government securities plus Treasury securities. (We do not subtract off U.S. holdings of foreign government securities, since these carry at least exchange rate risk.) The current account deficit is imports of goods and services and income payments less exports of goods and services and income receipts less net unilateral current transfers. The sample is annual, 1976-2010.

sharply positive during the housing bust period. So where are the inflows during the housing boom? In U.S. safe securities.

The only assets for which we observe a significant increase in capital inflows during the boom period are those we defined above as U.S. safe securities, comprised of Treasury and Agency debt. We have argued above that these assets are likely to be the most relevant for housing markets, and indeed the change in net foreign holdings of these securities was positive and increasing throughout the boom period, from 2001 through 2006. At the same time, however, inflows into these securities, like the other categories of assets, continued to rise during the bust period, implying that the U.S. borrowing from abroad in these securities increased further from 2006 on, rather than declining.

The only measure in Figure 4 that suggests a decline in U.S. borrowing from abroad over

the whole period 2006-2010 is the current account deficit, pointing to a puzzling incongruence with the holdings data. Despite the volatility in the holdings data, the bust period still exhibited relatively high average inflows of foreign capital into all forms of U.S. securities, mirroring capital gains on U.S. Liabilities relative to U.S. assets abroad, as well as a net flight into U.S. safe securities, in 2008. The CA deficit (or, equivalently, the capital account) omits these significant valuation changes during the financial crisis. We view this as a serious shortcoming of the CA as a measure of international capital flows, since such valuation adjustments surely have wealth effects that in general equilibrium would influence the extent to which U.S. households can consume at rates that exceed domestic income. At the end of the sample in 2009, Figure 4 shows that there is a decline in the (moving average trend) inflows to total assets (this does not explain why the CA deficit declined starting in 2006, however). But this decline is barely discernable in total securities and not at all present in U.S. safe securities. The discrepancy is again net flows into FDI, which we've argued above are largely divorced from the housing market.

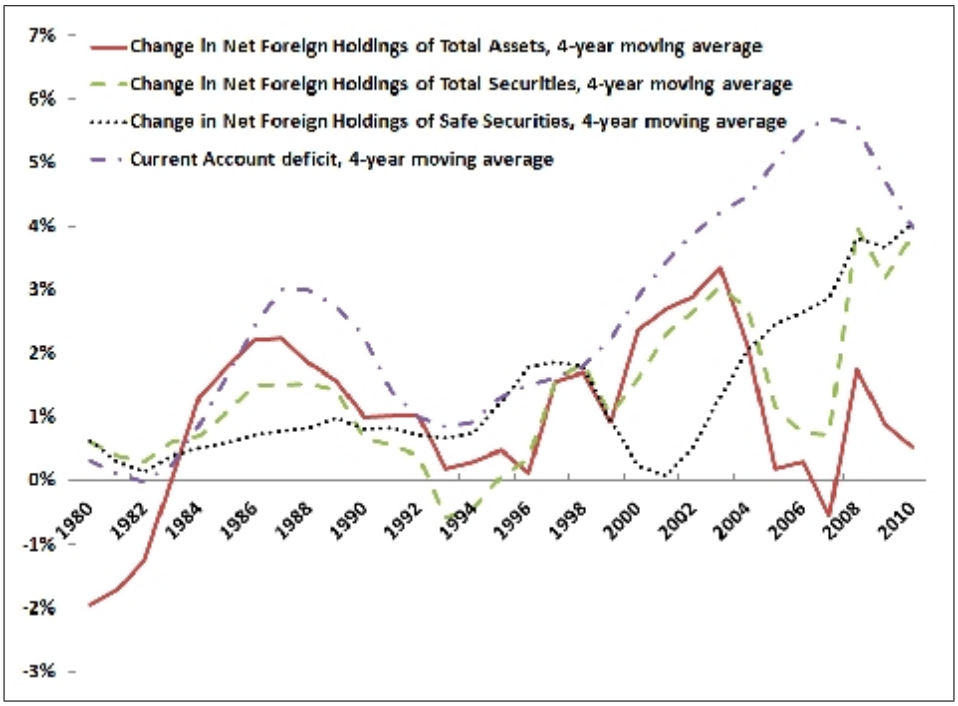


Figure 4: Measures of U.S. Capital Flows, 4-Year Moving Average. Notes: see Figure 3. Source: U.S. Department of Commerce, Bureau of Economic Analysis. The sample is annual, 1980-2010.

To summarize, during the housing boom, only U.S. capital inflows on *securities* (equities, corporate, U.S. Agency and Treasury bills and bonds) show any discernable upward trend.

Among securities, the upward trend has been driven almost entirely by an increase in net foreign holdings of U.S. safe assets, specifically U.S. Treasury and Agency debt. Yet net inflows on these securities, rather than declining during the housing bust, have continued to increase.

We now provide more detail on the flows to U.S. safe securities. To get a better sense of the quantitative importance of these flows to U.S. safe assets, the solid line of Figure 5, measured against the left axis, plots the combined foreign holdings in billions of U.S. dollars of short-term and long-term U.S. Treasuries and Agencies. The dashed line, measured against the right axis, shows long-term (not short-term) Treasuries and Agencies relative to the amount of long-term marketable debt outstanding. Figure 6 plots total foreign holdings of U.S. treasuries and Agencies relative to the size of the U.S. economy, measured as trend GDP.

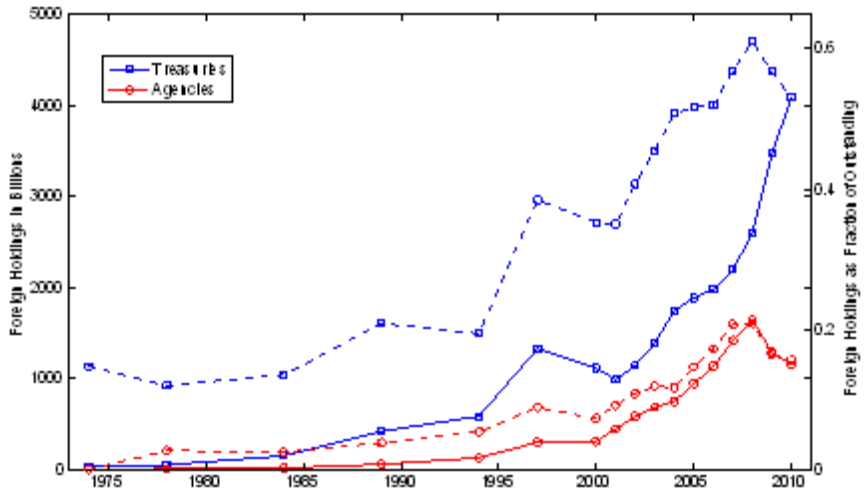


Figure 5: Foreign Holdings of U.S. Safe Assets. The figure plots foreign holdings of U.S. Treasuries (squares) and U.S. Agencies (circles). U.S. Agencies denotes both the corporate bonds issued by the Government Sponsored Enterprizes and the mortgage-backed securities guaranteed by them. The solid lines denote the amount of long-term and short-term holdings, in billions of U.S. dollars, as measured against the left axis. The dashed lines denote the long-term foreign holdings relative to the total amount of outstanding long-term (marketable) debt. Source: U.S. Treasury International Capital System's annual survey of foreign portfolio holdings of U.S. securities. The foreign holdings data are available for December 1974, 1978, 1984, 1989, 1994, 1997, March 2000, and annually from June 2002 until June 2010.

The figure shows that foreign holdings of U.S. Treasuries were modest until the mid 1990s.

In December 1994, foreign holdings of long-term Treasuries were $464bn, which amounted to 19.4% of marketable Treasuries outstanding and to 6.4% of U.S. trend GDP. Foreign holdings of long-term Agencies were $121bn, which amounted to 5.4% of outstanding Agencies and 1.5% of trend GDP. Over the course of the Asian financial crisis, these holdings doubled. By March 2000, towards the end of the crisis, foreign holdings of long-term Treasuries and Agencies were $884bn and $261bn, respectively, corresponding to 35.3% and 7.3% of the amounts outstanding. Total foreign holdings of Treasuries and Agencies increased from 9.8% to 14.8% of trend GDP. Caballero, Fahri, and Gourinchas (2008) argue that the Asian financial crisis represented a negative shock to the supply of (investable/pledgeable) assets in East Asia, and led their investors to increase their investments in U.S. bonds, one of the scarce risk-free assets available worldwide.

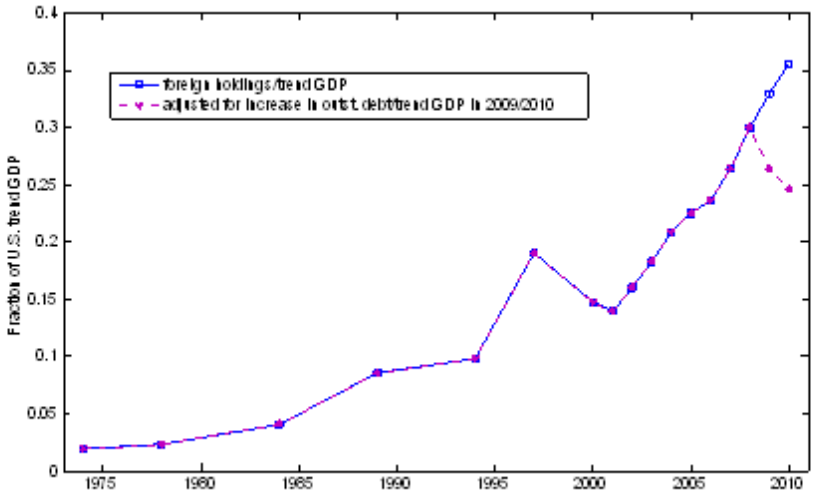


Figure 6: Foreign Holdings Relative to GDP. The solid line denotes foreign holdings of U.S. Treasuries and Agencies relative to U.S. trend GDP (squares). Trend GDP is computed with a Hodrick-Prescott filter (Hodrick and Prescott (1997)). The dashed line (stars) asks what the foreign holdings relative to trend GDP would have been if the foreign holdings relative to the amount of debt outstanding declined the amount they did, but the amount of debt outstanding relative to trend GDP was held at 2008 values for the years 2009 and 2010. The foreign holdings data are available for December 1974, 1978, 1984, 1989, 1994, 1997, March 2000, and annually from June 2002 until June 2010.

During housing boom from 2000-2006, the increase in foreign holdings continued at an even more rapid pace. Total foreign holdings of Treasuries and Agencies more than doubled

from $1,418bn in March 2000 to $3,112 in June 2006. Foreign holdings of long-term Treasuries went from 35.3% to 52.0% of the total amount of Treasuries outstanding, while holdings of long-term Agencies went from 7.3% to 17.2%. Most of the rise in foreign holdings of Treasuries took place by 2004, while most of the rise in Agencies took place from 2004 to 2006. Total foreign holdings of Treasuries and Agencies increased from 14.8% to 23.7% of trend GDP. The boom in U.S. house prices, which started at the end of 1994 and accelerated after 2000, coincided with a massive inflow of foreign capital in safe U.S. assets.

At the same time, however, capital inflows in the U.S. safe assets continued to rise during the housing bust and financial crisis. Figure 5 shows that between June 2006 and June 2010, total foreign holdings of Treasuries and Agencies rose from $3,112bn to $5,232bn, or from 23.7% to 35.5% of trend GDP. The share of outstanding long-term Treasuries held by foreign investors also increased from 52.0% in 2006 to 61.1% in 2008 before falling back to 53.0% in 2010. The reduction in 2010 is attributable to large increase in the total quantity of marketable Treasuries outstanding in 2009 and 2010 (which rose from 33.2% of trend GDP in 2008 to 54.9% in 2010), rather than to a reduction in nominal foreign holdings. The latter actually continued to increase rapidly from $2,211bn in 2008 to $3,343bn in 2010. The dashed line in Figure 6 is a foreign holdings-to-trend GDP series that we have adjusted in 2009 and 2010 to reflect the large increase in the quantity of Treasury debt outstanding that occurred in 2009 and 2010. The adjusted series equals the level of foreign holdings as a fraction of trend GDP that *would have* occurred in 2009 and 2010 had Treasury debt outstanding as a fraction of trend GDP been fixed at its 2008 level. The dashed line shows that the increase in foreign holdings of U.S. Treasuries in 2009 and 2010 is less than proportional to the increase in outstanding Treasuries over those years. In this relative sense, therefore, foreigners have become less willing to hold U.S. Treasuries. According to the adjusted series there is a reduction in foreign holdings as a fraction of trend GDP, from 30.0% of trend GDP in 2008, to 24.6% in 2010, suggesting that a substantial "unwind" of foreign positions in U.S. Treasuries may be underway, at least relative to the total amount of U.S. debt being issued.

Although there has so far been no reduction in nominal foreign holdings of U.S. Treasuries during the housing bust, the financial crisis did lead to a substantial reduction in nominal foreign holdings of U.S. Agencies. While foreign holdings of Agencies still rose from 17.2% in 2006 to 20.8% of the amount outstanding in 2008, they fell back sharply to 15.6% of the amount outstanding in 2010 even as the amount outstanding remained flat.

Foreign Official Holdings An important aspect of recent patterns in international capital flows is that foreign demand for U.S. Treasury securities is dominated by Foreign Official Institutions. Krishnamurthy and Vissing-Jorgensen (2010) find that demand for U.S. Treasury securities by governmental holders is extremely inelastic, implying that when these holders receive funds to invest they buy U.S. Treasuries, regardless of their price. As explained in Kohn (2002), government entities have specific regulatory and reserve currency motives for holding U.S. Treasuries and face both legal and political restrictions on the type of assets that can be held, forcing them into safe securities.

Data from the TIC system breaks out what share of foreign holdings of U.S. Treasuries is attributable to Foreign Official Institutions, which are government entities, mostly central banks. Foreign Official Institutions own the vast majority of U.S. Treasuries in recent data: in June 2010 Foreign Official Institutions held 75% of all foreign holdings of U.S. Treasuries. That share has always been high and has risen from 58% in March 2000 to 75% in June 2010. Indeed, 75% represents a *lower bound* on the fraction of such securities held by Foreign Official Institutions, since some prominent foreign governments purchase U.S. securities through offshore centers and third-country intermediaries, purchases that would not be attributed to Foreign Official entities by the TIC system–see Warnock and Warnock (2009). Foreign Official Institutions also accounted for 64% of the foreign holdings of Agencies in June 2010.

Asian central banks (China, Japan, Korea) have acquired massive U.S. dollar reserves in the process of stabilizing their exchange rate. The share of foreign holdings is higher for long-term than for short-term securities. The left panel of Figure 7 shows the foreign holdings as of December 2010 by country groups. China excludes Hong Kong, which is part of Rest of Asia. Banking centers consist of the United Kingdom, the Caribbean, Luxembourg, Belgium, and Ireland. It is widely believed that China holds a non-trivial fraction of its safe dollar assets through financial intermediaries in the U.K. and in other banking centers Warnock (2010). The graph then suggests that as much as two-thirds of safe U.S. assets is held by Asian countries. China (narrowly defined) held nearly $1,500bn in Treasuries and Agencies in June 2010; Japan held nearly $1,000bn.

The right panel of Figure 7 shows the share of foreign holdings of U.S. Treasuries and Agencies for different maturity ranges, expressed in years. The data are for June 2009. As the caption explains, the maturity of the Agency holdings is adjusted to account for the prepayment option embedded in mortgage-backed securities. Total long-term and short-term foreign holdings have a weighted average maturity of 4.6 years. About a quarter of foreign holdings have a maturity of one year or less. Fully half of all holdings have a maturity below

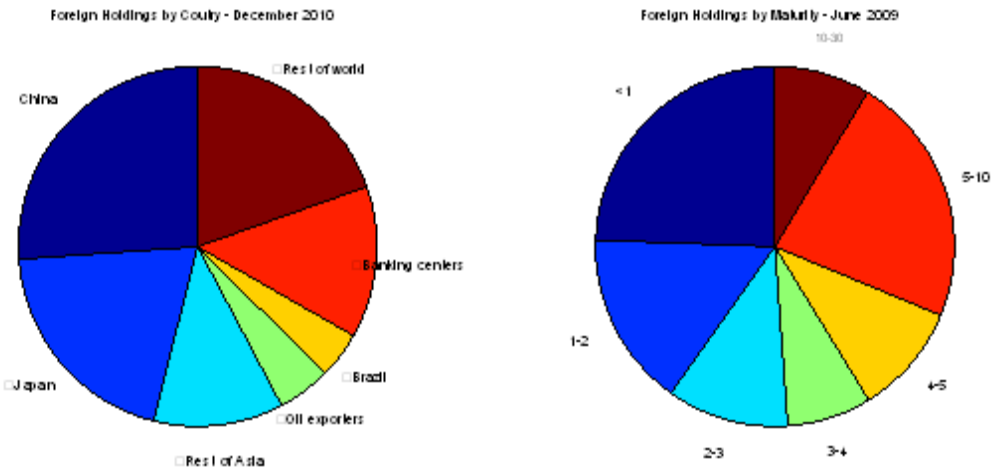


Figure 7: Foreign Holdings By Maturity and By Country. The left panel shows the share of foreign holdings of U.S. Treasuries and Agencies for different maturity ranges, expressed in years. The data are for June 2009 (Table 14a of the 2009 TICS report). 51% of the Agency holdings are reported to be of maturity 25-30 years. These are 30-year mortgage backed securities, which have a nominal maturity of 30 years but an effective maturity of about 7 years because of prepayment. We reallocate these 30-year Agency holdings across maturities so that they have a weighted average maturity (WAM) of 7 years. The resulting series for the maturity of foreign holdings of U.S. Agencies has a WAM of 6.4 years instead of the original 17.7 years. Total long-term foreign holdings have a WAM of 5.2 years. Once we add the short-term Treasury holdings, which are 1-year or less in maturity, the WAM drops to 4.6 years. The right panel shows the foreign holdings as of December 2010 by country groups. China excludes Hong Kong, which is part of Rest of Asia. Banking centers consist of the United Kingdom, the Caribbean, Luxembourg, Belgium, and Ireland.

3 years. This suggests that a substantial reduction in foreign holdings of U.S. safe assets could occur over a relatively short period without an outright fire-sale of long-term bonds, if current holders simply stopped rolling over existing positions.

Longer Term Trends in Net Foreign Holdings of Securities We have emphasized the special relevance for the U.S. housing boom-bust cycle of U.S. securities considered to be safe stores-of-value (i.e., U.S. Treasury and Agency debt). But is worth emphasizing that, even over a longer period of time, foreign holdings of these securities behave similarly to total net foreign holdings of all securities. The reason is that foreign holdings of U.S. securities other than Treasuries and Agencies are roughly equal in magnitude to U.S. holdings of securities abroad. Figure 8 makes this point visually; net foreign holdings of all securities other than

U.S. Treasury and Agency debt as a fraction of U.S. Trend GDP have hovered close to zero since 1994, even as net foreign holdings of safe securities have soared. This shows that all of the long term upward trend in net foreign holdings of U.S. securities since 1994 has been the result of an upward trend in net foreign holdings of U.S. *safe* securities; net foreign holdings of other securities are almost exactly zero in June of 2010. Thus the long term downward trend in the U.S. net international investment position is well described by the foreign holdings of U.S. safe assets. Of the safe assets foreigners hold, 70%, on average over the period 1994-2010, are held in U.S. Treasuries. The large U.S. current account deficits in the boom period are, to a large extent, the mirror image of the increase in foreign holdings of U.S. safe assets.[11] This is true in the aggregate net flows to the U.S., but also for China. China's cumulative current account surplus between 2003 and 2007 largely matches up with its acquisitions of U.S. Treasuries and Agencies (Bernanke (2011)).

Risky Mortgage Holdings Although net flows into securities other than Treasuries and Agencies have hovered around zero, there were substantial *gross* flows across borders into private-label products such as mortgage-backed securities (MBS), collateralized debt obligations (CDO), and credit default swaps (CDS) with non-prime residential or commercial real estate as the underlying or as the reference entity. Because an average of 80% of such private-label MBS principal received a AAA rating from the credit ratings agencies and earned yields above those of Treasuries, large foreign (as well as domestic) institutional investors were able and willing to hold these assets on their books. The TICS data indicate that foreigners held $594bn of non-agency mortgage-backed securities in June 2007. By June 2009, these holdings more than halved to $266bn, after which they stabilized at $257bn in June 2010. Less than 10% of these are held by foreign official institutions.

Bernanke (2011) shows interesting cross-country differences in the composition of countries' U.S. investment portfolio. China and emerging Asia held three-quarters of their U.S. investments in the form of Treasuries and Agencies in 2007. Their share of all AAA-rated securities was 77.5%, while the AAA-rated share of all U.S. securities outstanding was only 36%. European (as well as domestic) investors held only about one-third of their U.S. portfolio in the form of AAA-rated assets. Not only do Europeans invest in non-AAA corporate debt, they accumulated $500bn in U.S. asset-backed (largely mortgage-backed) securities

[11] Though, as discussed above, an important discrepancy between the current account data, based on transactions, and the net foreign assets holdings data, is that the former do not fully adjust for valuation effects that are captured in the international holdings data.

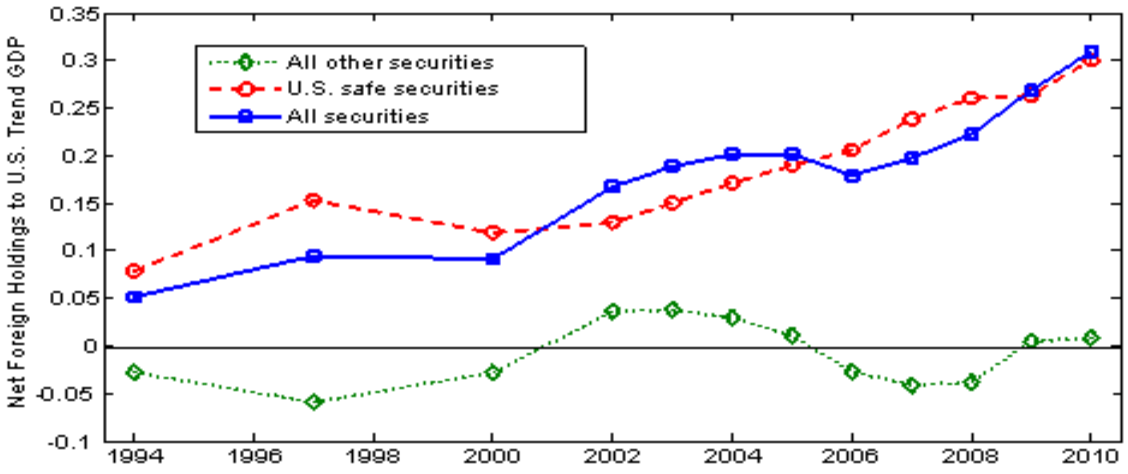


Figure 8: Net Foreign Holdings Relative to U.S. Trend GDP. The solid line (squares) denotes total net foreign holdings of long-term securities (the net foreign liability position of the U.S.) relative to U.S. trend GDP. Net foreign holdings are defined as foreign holdings of U.S. securities minus U.S. holdings of foreign securities. We define as net foreign holdings of safe securities to be foreign holdings of U.S. Treasuries and Agencies. (We do not subtract off U.S. holdings of foreign government securities, since these carry at least exchange rate risk.) The dotted line (stars) denotes the thus constructed net foreign holdings in safe securities, while the dashed line (circles) denotes the net foreign holdings in all other securities. The data are from the U.S. Treasury International Capital System's annual survey of foreign portfolio holdings of U.S. securities. The data are available for December 1994, December 1997, March 2000, and annually from June 2002 until June 2010.

between 2003 and 2007.

In addition to their different risk profiles over the housing cycle, Europe and Asia differ by their current account positions. While the Asian economies ran a large current account surplus, financing the purchases of U.S. assets with large trade surpluses, Europe had a balanced current account over this period. It financed the purchases of risky U.S. assets by issuing external liabilities, mostly equity, sovereign debt, and asset-backed commercial paper (ABCP). A prototypical example of European holdings were AAA-rated tranches of subprime MBS held by large banks through lightly-regulated off-balance sheet vehicles, and financed with ABCP (Acharya, Schnabl, and Suarez (2010)).

3.2 U.S. Interest Rates

We have seen that the long term upward trend in net foreign holdings of U.S. securities since 1994 has been the result of an upward trend in net foreign holdings of U.S. safe secu-

rities. The rise in net holdings by foreigners over time has coincided with downward trend in interest rates. Both 30-year fixed rate mortgages and the 10-year Treasury bond yield trended downward, as illustrated in Figure 9. The decline in nominal rates was not merely attributable to a decline in inflation: the real annual interest rate on the ten-year Treasury bond fell from 3.5% at the start of 2000 to -0.2% at the end of 2008 using the consumer price index as a measure of inflation.[12] It subsequently increased to 1.6% by the end of 2010. We also investigate 10-year Treasury Inflation Protected (TIPS) rates, which are available from 2003 onward from FRED II and from McCulloch from 1997 to 2003. Using the same 12 month moving average, we find that real 10-year TIPS yields were 4.0% at the start of 2000 and fall to 1.8% by the end of 2008. They fall further to 1.2% by the end of 2010. In summary, real long-term yields fall more than 3 percentage points between 2000 and 2008, and continue to fall during and after the financial crisis. Since 2009, there is evidence of an uptick in real rates according to the series in (9), but not in the TIPs data, which shows a further decline.

Empirically, Bernanke, Reinhart, and Sack (2004) find evidence for lower Treasury yields around periods of Japanese purchases of U.S Treasuries in the 2000-2004 period, while Warnock and Warnock (2009) estimate that 12-month flows equal to one percent of GDP are associated with a 19 basis point reduction in long rates. They also find U.S. mortgage rates to be affected. The effects are large. Had the twelve months ending in May 2005 seen zero foreign official purchases of U.S. Treasury and agency bonds, their results suggest that, ceteris paribus, U.S. long rates would have been about 80 basis points higher.

3.3 Financial Market Liberalization

While there is little doubt that positive inflows into U.S. Treasury and Agency markets have lowered long-term Treasury and mortgage rates, there is no direct evidence that they have played an important role in raising house prices during the boom. We argued above that there are good theoretical reasons to doubt the hypothesis that lower interest rates had a quantitatively large effect on house prices during the boom. Empirically, Glaeser, Gottlieb, and Gyourko (2010) concur and find that even when the house price impact of lower interest

[12] To compute the real interest rate, we use the 10-year constant maturity Treasury rate and subtract realized inflation over the preceding 12 months. We then take a 12-month moving average of this series to smooth it out.

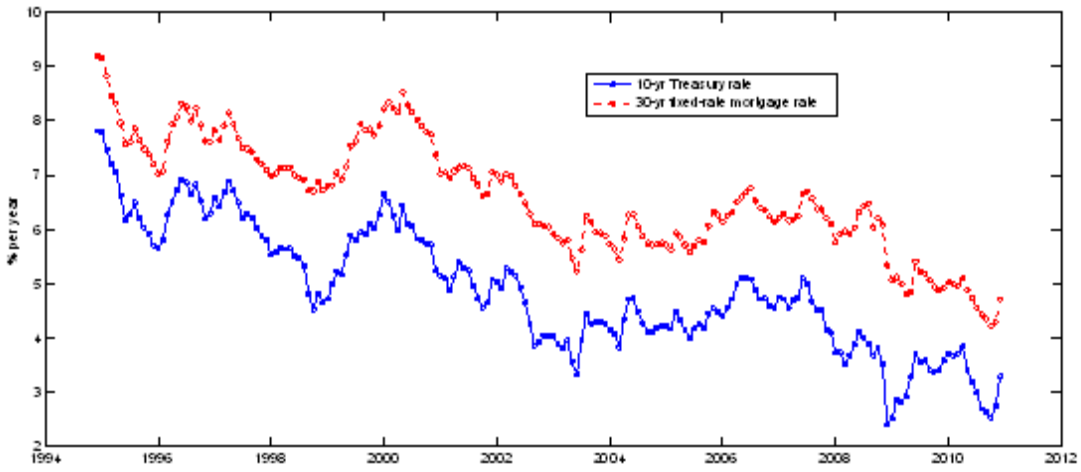


Figure 9: Fixed-rate Mortgage Rate and Ten-Year Constant Maturity Treasury Rate The solid line (squares) plots the ten-year Constant Maturity Treasury Yield (CMT). The dashed line (circles) plots 30-year Fixed-Rate Mortgage rate (FRM). The FRM data are from Freddie Mac's Primary Mortgage Market Survey. They are average contract rates on conventional conforming 30-year fixed-rate mortgages. The CMT yield data are from the St.-Louis Federal reserve Bank (FRED). The data are monthly from December 1994 until December 2010.

rates should be stronger (at a low initial rate), they account for at most 1/5 of the observed change in housing prices. We present additional evidence on this below.

What then accounts for the dramatic rise in U.S. house prices during the boom if not low interest rates? A key missing element in this scenario is the shift in credit standards and housing transactions costs, summarized above as a FML and its reversal. The widespread relaxation of credit standards is well documented (see Nichols, Pennington-Cross, and Yezer (2005), Favilukis, Ludvigson, and Van Nieuwerburgh (2009), Ferrero (2011) for more details). Moreover, a growing body of empirical evidence directly links measures that identify changes in credit *supply* (as opposed to changes in demand) to movements in asset prices.

Adelino, Schoar, and Severino (2011) exploit exogenous variation in the government-controlled conforming loan limit (CLL) as an instrument for changing credit supply. The CLL determines the maximum size of a mortgage that can be purchased or securitized by Fannie Mae or Freddie Mac. Because these loans were widely understood to have the implicit (and since 2008 explicit) backing of the U.S. government, borrowers in the market for loans that fall below the CLL have easier access to credit at less costly terms. Changes in the CLL are set annually and depend on the previous year's limit plus the change in the median national house price. These movements are clearly exogenous to individual mortgage transactions,

local housing markets, and the local economy. Using data on single-family house purchases in 10 metropolitan statistical areas between 1998 and 2006, Adelino et. al show that houses that became newly eligible for a conforming loan just after an increase in the CLL saw significant price increases relative to similar houses that were already below the limit before the CLL increase.

Rajan and Ramcharan (2011a) exploit interstate banking restrictions to study the effect of credit supply on land prices in the early 20th century U.S. Regulations at the time stipulated that banks could not lend across state borders. They argue that the number of banks in this era proxied for credit supply, with more banks indicating higher supply. They show that the number of banks in a county positively predicts land prices independently of fundamentals likely to move credit demand for land (commodity prices). They also find that the number of banks in neighboring in-state counties affects land prices more than the number of banks in equidistant counties out-of-state. Since banks were prohibited from lending across state borders, it is difficult to form a coherent story for this latter fact that does not involve credit supply.

In a similar spirit, Favara and Imbs (2011) identify movements credit supply in more recent data (since 1994) by studying bank branching restrictions. Even though interstate banking (i.e., cross-state ownership of banks) was made legal after the passage of the Interstate Banking and Branching Efficiency Act of 1994, U.S. states retained the right to erect barriers to interstate branching. They study branching deregulations since 1994 and show that they significantly affect the supply of mortgage credit. With deregulation, the number and volume of originated mortgage loans rise, while denial rates fall, echoing evidence in Mian and Sufi (2009), discussed below. This deregulation has no effect on a placebo sample, formed of independent mortgage companies that should not be affected by the regulatory changes. Deregulation leads to greater supply of mortgage credit, which they find leads to significantly higher house prices.

Our main measure of credit availability is based on quarterly bank lending surveys for countries in the Euro area and the U.S. For the U.S., we use the Senior Loan Officer Opinion Survey on Bank Lending Practices (SLO), collected by the Federal Reserve. An important aspect of this survey is that it asks banks to explicitly distinguish between changes in the *supply* of credit as distinct from the *demand* for credit, on bank loans to businesses and households over the past three months. Thus in principal, answers to the appropriate questions are able to identify a movement in supply separately from a movement in demand. We focus on questions related to mortgage credit supply to households. The detailed information

is considered highly reliable because the surveys are carried out by central banks which are also bank regulators with access to a large amount of information about a bank's operations, including those reflected in loan applications and balance sheet data.

Data for other countries are from bank lending surveys conducted by national central banks, and the European Central Bank. The survey questions are modeled after the U.S. Survey of Senior Loan Officers. (See the Appendix for data sources.) We use these data below in our empirical analysis.

For the U.S. SLO survey, banks indicate easing, tightening, or no change in lending standards over the previous three months. We use the net percentage of banks that have eased their lending standards on mortgage loans as a measure of credit supply. This is the difference between the percentage of banks reporting easing and the percentage of banks reporting tightening, thus a positive figure indicates a net easing of lending standards, considering all bank respondents.

Figure 10 reports the net percentage of banks easing over time. According to this measure, there was a significant easing of standards from 2002-2006, though not unprecedented, and a very sharp tightening afterwards. Notice that this measure does not weight banks by their relative importance in the mortgage market. Moreover, until 2006, the survey did not distinguish between prime and subprime mortgages. Thus the full easing of standards between 2002 and 2006 may not be reflected in this measure of credit supply. As a supplement, we therefore report several other measures of credit supply in Figure 11, using data collected by Mian and Sufi (2009). These data stop in 2007, but the four measures show a significant increase in credit availability over the period 2000-2006, and the beginnings of a sharp decline thereafter from 2006 to 2007. These measures include the cumulative growth in the number of mortgages originated (subprime versus prime), the cumulative growth in the amount of mortgages originated (subprime versus prime), the difference in applicant denial rate (subprime versus prime), and the share of subprime mortgages securitized (as measured by share sold to non-GSE investors). As Mian and Sufi (2009) argue, these measures are all likely related to credit availability, via increases in maximal loan-to-value ratios, increases in the availability of new mortgage contracts, reductions in documentation requirements (asset and income verification), increases in the use of private-label securitization, and reductions in fees associated with refinancing a mortgage or getting a home equity line of credit. Mian and Sufi (2009) show that these mortgage credit expansion measures are also statistically related to house price changes (in the same direction) in those same areas.

In the empirical work below, we will focus on the quarterly loan survey data on mortgage

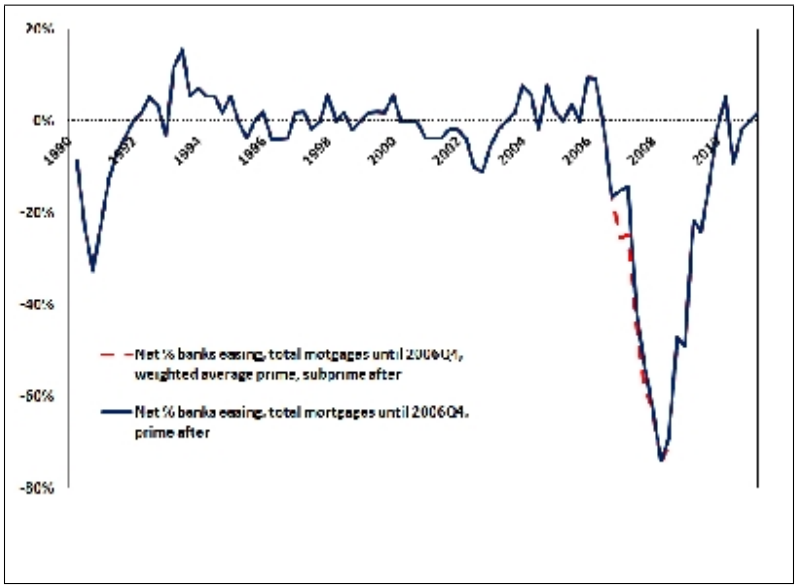


Figure 10: Net Percentage of U.S. Banks Reporting Easier Credit Standards. The figure reports the percentage of banks that reported easing of credit standards on mortgages, less the percentage that reported tightening os standards. A positive number indicates that more banks report easing than tightening. A negative number indicates the opposite (more banks tightening than easing). Until 2006, surveys did not distinguish between prime and subprime mortgages. Beginning in 2007:Q1, the figure shows the net percentage easing for two series: prime mortgages only, and a weighted average of prime and subprime mortgages, where weights are computed based on the shares of prime and subprime in total mortgages, as described in the appendix. Source: U.S. Survey of Senior Loan Officers, 1990Q2-2011Q2.

credit standards as a measure of credit supply, since we have far more observations (at higher frequency and over a longer time period) than we have with the Mian and Sufi data. In thinking about these data, it is instructive to consider how they may relate to the notion of credit availability in FLVNa. In that model, a FML involves both a reduction in transactions costs associated with buying and selling the housing asset, and a change in collateralized borrowing constraints. Consider the borrowing constraint component, which takes the form:

$$-B^i_{t+1} \leq (1-\varpi)\, p^H_t\, H^i_{t+1}, \qquad \forall a, t \tag{1}$$

where B^i_{t+1} is the amount of bonds household i owns at the beginning of period $t+1$, p^H_t denotes the relative price of housing in units of the non-housing consumption good (p^H_t is the time t price of a unit of housing of fixed quality and quantity), and H^i_t is the housing stock owned by household i and the beginning of period $t+1$. A negative value for B^i indicates a borrowing position. This equation represents the collateral constraint in the model, where $0 \leq \varpi \leq 1$. It says that households may borrow no more than a fraction $(1-\varpi)$ of the

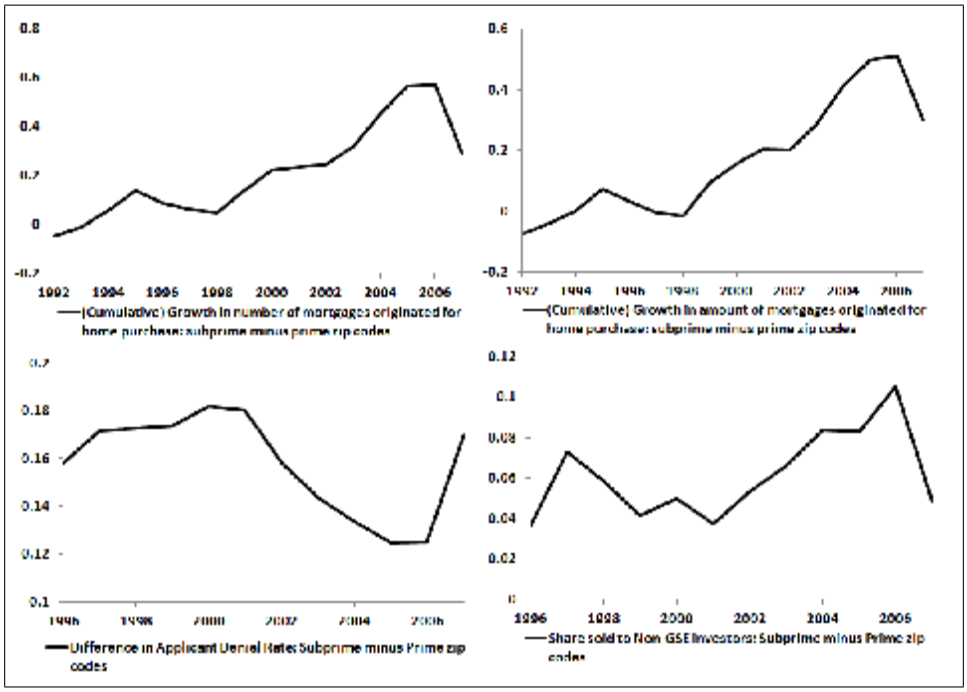


Figure 11: Measures of Credit Supply (Mian and Sufi). Source: Mian and Sufi (2009).

value of housing, implying that they must post collateral equal to a fraction ϖ of the value of the house. This constraint can be thought of as a down-payment constraint for new home purchases, but it also encompasses collateral requirements for home equity borrowing against existing homes. The constraint gives the maximum combined loan-to-value (LTV) ratio for first and second mortgages and home equity withdrawal.

FLVNa ask how a plausibly calibrated change in its value (along with a calibrated change in housing transactions costs) changes the equilibrium outcome. Thus, one way credit supply can change is via a change in the fraction $(1 - \varpi)$ of the home's value that must be held as collateral. But, as discussed above, borrowing capacity will fluctuate endogenously with the collateral value $p_t^H H_{t+1}^i$ even if that fraction remains unchanged. This represents an endogenous change in borrowing capacity, driven by economic shocks and accompanied by revisions in expectations about future economic conditions. These factors get immediately reflected in house prices, which affects borrowing capacity and the tightness of constraints. We argue that *either* of these represent a change in credit supply in the sense that they are related to a borrower's access to funds via her credit constraint. Moreover, the two could be correlated (expectations of a decline in economic activity could lead to an increase in ϖ), as they are in the transition dynamics studied by FLVNa and displayed in Figure 2 above.

4 Empirical Evidence on House Price Fluctuations

We have seen above that the U.S. experienced large capital inflows and commensurate current account deficits at the same time that it experienced strong growth in house prices. The same is true for countries such as Iceland, Ireland, and Spain. In fact, during the boom, there is a positive cross-country correlation of current account deficits with house price growth on the one hand (Ferrero (2011), Laibson and Mollerstrom (2010), Gete (2010)) and with value added and the labor share of the construction industry on the other hand (Gete (2010)). Using data that ends before the bust, Aizenman and Jinjarak (2009) provide a precise estimate of the relationship between house prices and external imbalances: a one standard deviation increase in lagged current account deficits is associated with a 10% appreciation of real estate prices.

Panel A of Table 1 replicates the flavor of these results. It reports real house price growth (deflated by CPI), cumulative current account deficits, and cumulative residential investment over the period 2000:Q1 to 2006:Q4, with the last two variables measured relative to GDP. Countries, such as Germany, Switzerland, China, Japan, Austria that accumulated large current surpluses exhibited slow house price growth and modest residential investment, while countries such as the U.S., Ireland, Spain, the U.K., Portugal, and Australia attracted lots of external capital and exhibited large rises in house prices and large residential investment booms. In the boom period, there is a positive (on average) cross-country correlation between house price changes and current account deficits equal to 23%. There is also a negative cross-country correlation between residential investment an house price growth: countries with more residential investment experienced lower house price growth, consistent with the idea that residential investment drives up the expected housing stock and puts downward pressure on house prices.

Table 1

	Panel A 2000Q1-2006Q4			Panel B 2006Q4-2010Q4		
	real HP gr. (% change)	CA def. (cum.) / GDP2006	Res. Inv. (cum.) /GDP2006	real HP gr. (% change)	CA def. (cum.) / GDP2006	Res. Inv. (cum.) /GDP2006
Australia	55%	24%	35%	17%	23%	27%
Austria	1%	-8%	29%	20%	-17%	19%
Belgium	18%	-17%	32%	10%	-1%	26%
Canada	46%	-10%	35%	10%	6%	30%
Czech Republic	20%	19%	18%	4%	14%	16%
Denmark	64%	-16%	32%	-20%	-15%	22%
Estonia	387%	47%	19%	-47%	26%	21%
Finland	37%	-35%	39%	8%	-15%	28%
France	85%	-3%	36%	1%	7%	22%
Germany	-16%	-17%	38%	-3%	-29%	23%
Greece	50%	39%	42%	-22%	62%	23%
Hungary	40%	39%	25%	-27%	15%	14%
Iceland	64%	57%	28%	-28%	62%	21%
Ireland	60%	8%	57%	-40%	15%	26%
Israel	-16%	-6%	27%	34%	-14%	22%
Italy	35%	7%	29%	-2%	12%	21%
Korea	25%	-12%	29%	-4%	-9%	20%
Luxembourg	71%	-51%	13%	-3%	-38%	14%
Netherlands	28%	-31%	37%	-7%	-27%	26%
New Zealand	73%	30%	35%	-10%	29%	23%
Norway	46%	-73%	21%	9%	-68%	16%
Poland	-2%	18%	16%	33%	29%	11%
Portugal	-6%	51%	42%	2%	51%	18%
Slovenia	46%	8%	11%	1%	18%	10%
Spain	87%	28%	45%	-16%	35%	30%
Sweden	61%	-35%	17%	15%	-36%	15%
Switzerland	12%	-75%	28%	13%	-40%	15%
United Kingdom	78%	13%	21%	-6%	9%	14%
United States	64%	30%	32%	-36%	17%	13%
Russia	157%	-39%	8%	10%	-30%	12%
China	-1%	-22%	38%	-6%	-50%	63%
Euro Area	32%	0.04%		-3%	2.4%	
Corr. CA	0.23	1.00	0.22	-0.38	1.00	-0.14
Corr. HP gr.	1.00	0.23	-0.25	1.00	-0.38	-0.09

Notes: House prices are deflated by CPI. The data is from different national sources (See below), mostly

quarterly, except for Germany (annual), Luxembourg (annual, and until 2009), Italy (semi-annual) and Japan (semi-annual). CPI is collected by EIU from national sources. For Slovenia series begins in 2003Q1; for Russia in 2001Q1. CA deficit data is from IMF, and SAFE for China. CA balances are accumulated and deflated by 2006 GDP (collected by EIU), all in current US dollars. For Belgium, CA data 2000Q1-2001Q4 is from NBB, via OECD Residential Investment is from Eurostat and National sources. Residential investment is accumulated and deflated by 2006 GDP, all in current national currency. For France, Hungary, Poland, Switzerland and Russia, data are available only until 2009.

It is tempting to conclude that the excess savings of the first group of countries found its way to the real estate industry in the second group of countries and fueled the housing boom there. However, as argued above, general equilibrium considerations suggest that large inflows into safe assets need not lead to large house price booms because the effect of lower interest rates is offset by a rise in risk premia and an expected increase in the housing stock from higher residential investment. This may help explain why the current account patterns from the boom period persisted in many countries during the housing bust, while house prices and residential investment patterns obviously did not. Panel B of Table 1 shows that the cross-country correlation between house price changes and current account dynamics reverses in the bust sample from 2006:Q4 to 2010:Q4. The cross-country correlation is now -38%. We now turn to a statistical analysis of the relation between house price changes (or changes in price-rent ratios), measures of capital flows, credit standards and interest rates.

4.1 Regression Analysis

In this section we undertake a basic empirical analysis of correlations among house prices and other variables. We emphasize that, in presenting these next results, we do not make claims about causality. Later, focusing on the correlation between credit standards and house prices, we will provide some additional discussion and evidence on this question.

A few words about the data are in order. First, with regard to an international panel of data, we are limited to far fewer observations given the availability of bank lending survey data for non-U.S. countries. These data extend only from 2002:Q2 to 2010:Q4. For this reason we also look separately at regressions for the U.S. alone, where data reaches back much further, starting in 1991:Q4. Second, as explained in the Appendix, the data on bank lending standards differs somewhat by country. When we analyze the U.S. alone, we use the *net percentage easing* indicator plotted above. In the other countries the surveys are modeled after the U.S. SLO survey, but the way the survey results are aggregated can differ.

For 9 of the 11 countries for which we have data on credit standards (Austria, Belgium, Euro Area, France, Korea, Portugal, Spain, United States, Ireland) we have available, or can construct, a diffusion index of credit standards or a scale transformation thereof with the information reported by the Central Banks. This index, however, is not a scale transformation of the net percentage indicator discussed above and there are two (Canada and Netherlands) that only report the net percentage indicator. For the panel regressions that we report on here, we simply use all these data together in one regression, even though the credit standards measure for two countries (Canada and Netherlands) are not a scale transformation of the other countries' measures. Results available on request show that the findings are virtually unchanged if we exclude these two countries. We standardize these bank lending survey indexes, country by country, by subtracting the mean and dividing by the standard deviation, for the period 2002Q4-2010Q4. The Appendix provides more details on the data construction.

Third, we use a measure of net foreign holdings of U.S. securities as our main measure of international capital flows into the U.S. (Although for completeness, we also report results using the CA deficit.) Recall that annual net foreign holdings estimates are compiled by the BEA in their international investment data, year-end positions.[13] Yet all the rest of our data are quarterly. Instead of limiting ourselves to annual observations, we therefore form a simple estimate of the *quarter*-end net foreign liability position of the U.S. in total securities, by employing a methodology to interpolate between the year-end positions, taking into account the quarterly *transactions* data in these same securities.[14] The procedure is as follows.

Let nh_{Q4} be the value of net foreign holdings of total securities observed at the end of quarter $Q4$ of a given year, where net foreign holdings are defined as foreign holdings of U.S. securities minus U.S. holdings of foreign securities. These data are available from the BEA year-end positions table. Let $\widehat{nh}_q$ be an estimate that we will form of the value of these net foreign holdings at the end of quarter q in that same year. Let nt_q be net transactions in those securities during that quarter, where net transactions are net foreign purchases (gross foreign purchases less gross foreign sales) of U.S. owned securities minus U.S. net purchases (gross U.S. purchases less gross U.S. sales) of foreign owned securities. These data are available from the BEA international transactions table. To obtain estimates of quarterly holdings

[13] The data are located at http://www.bea.gov/international/xls/intinv10_t2.xls

[14] The data on international transactions in financial securities are in the balance of payments dataset, found at http://www.bea.gov/international/xls/table1.xls.

for the three quarters within a year, we accumulate according to

$$\widehat{nh}_q = \widehat{nh}_{q-1} + nt_q + adj_q$$

where

$$\begin{aligned} adj_q &= gap_{Q4} * w_{q,Q4} \\ w_{q,Q4} &\equiv \frac{|nt_q|}{\sum_{k=0}^{3} |nt_{Q4-k}|} \\ gap_{Q4} &\equiv (nh_{Q4} - nh_{Q4-4}) - \sum_{k=0}^{3} nt_{Q4-k}. \end{aligned}$$

The above recursion ensures that our estimate of the holdings at the end of Q4 of a given year, $\widehat{nh}_{Q4}$, is equal to the recorded value from the annual holdings data, nh_{Q4}. For all other quarters within a year, the above recursion forms an estimate of holdings at the end of the quarter, which is equal to the estimated net holdings from last quarter, plus the net transactions in that quarter, plus an adjustment. The adjustment is equal to the *gap* between the change in measured holdings from the year in which the quarter resides and the previous year and the cumulation of all the quarterly transactions over the year, times a weight, where the weight is given by that quarter's volume of net transactions relative to the volume over the entire year. Thus, quarters for which net transactions were higher in absolute value receive a greater weight in the adjustment. Notice that, in the absence of any valuation adjustments, the cumulation of all the quarterly transactions over the year would equal the total change in net foreign holdings or year-end positions. The observed change in year-end positions takes into account the valuation changes, and gap_{Q4} is the difference between the observed change in year-end positions and the cumulation of the quarterly transactions. Thus, roughly speaking, the adjustment adj_q captures the pure valuation effects that are not reflected in the cumulation of transactions but are reflected the total change in net foreign holdings. The weights $w_{q,Q4}$ give quarters with a larger volume of transactions more weight in the adjustment. We now present results from a regression analysis using these and other data.

We begin with evidence from the panel of 11 countries mentioned above (Austria, Belgium, Canada, Euro Area, France, Korea, Netherlands Portugal, Spain, United States, Ireland). For these countries, we have quarterly observations from 2002:Q4 to 2010:Q4 on real house price growth, the current account deficit and credit standards. The variable CS

in these regressions is a index measure that, when increased, indicates an *easing* of credit standards. A detailed description of all of these data, including data sources, is given in the Appendix. Table 2 reports the results of a panel fixed-effects regression of real house price growth on the current account, CS, and interactions of these variables. The variable CA/GDP is the current account *deficit,* divided by the country's GDP.

Table 2 shows that CA/GDP bears a *negative* relation to contemporaneous house price growth, though it is not statistically significant.[15] The credit standard measure CS is statistically significant, and an increase in CS leads to an increase in real house price growth (row 2). Row 3 shows that CS remains the only significant determinant of house price growth when both variables are included, while rows 4 and 5 document some interaction effects: countries and time periods in which there was an increase in credit supply experienced a larger increase in house price growth if they also ran current account deficits, but, controlling for this, CA/GDP has a statistically insignificant marginal effect on house price growth (row 4), while CS by itself has a strongly significant marginal effect (row 5). The R^2 statistics range from 6 to 9 percent whenever CS_t is included in the regression, either by itself or interactively with CA/GDP. These results provide little support for the hypothesis that capital inflows played an important role in driving the volatility in house prices internationally over the recent boom-bust period.

[15] We use the Driscol and Kraay (1998) consistent covariance-matrix estimates to produce heteroscedasticity consistent standard errors that are robust to general forms of autocorrelation and cross-sectional (spatial) correlation between the residuals.

Table 2: Quarterly Panel Regressions (2002Q4-2010Q4)
11 Countries

	real House prices growth on				
Regression	*Cons*	CA/GDP	CS	$(CA/GDP)xCS$	R^2
1	0.005	-0.055			0.01
	(1.52)	(-0.73)			
2	0.005		0.005		0.06
	(1.69)		(3.24)**		
3	0.005	-0.018	0.005		0.07
	(1.62)	(-0.29)	(3.26)**		
4	0.005	-0.009		0.083	0.05
	(1.58)	(-0.14)		(5.34)**	
5	0.005		0.004	0.060	0.09
	(1.96)		(3.20)**	(6.61)**	

Notes: Panel data estimation with fixed effects. (CA/GDP) is current account deficit divided by the country's GDP. CS is net percentage of banks reporting easing of credit; a positive value indicates more banks eased than tightened credit conditions with respect to previous quarter. The column labeled *Cons* gives the coefficient on the regression constant. Credit conditions have been standardized country by country. Driscoll-Kraay corrected t-statistics in parentheses (lags *max{H-1,3}*). * significative at 5% .** significative at 1%. 11 countries included are Austria, Belgium, Canada, Euro Area, France, Ireland, Korea, Netherlands, Portugal, Spain, U.S. 363 Observations in total.

To investigate a longer time frame, we now turn to an analysis of U.S. time series data. Table 3 presents results from regressions using the same variables as in Table 2, but this time only for the U.S. The U.S. data are quarterly and span the period 1990:Q2-2010:Q4. Row 1 shows that CA/GDP has virtually no effect, by itself, on real house price growth. This variable explains two percent of the quarterly variation in real U.S. house price growth. By contrast, CS is strongly statistically significant, and by itself explains 53% of the quarterly variation in house price growth. When we include both CA/GDP and CS in the regression, the current account has a statistically significant and positive effect, and this adds to the regression model's ability to explain the data: the adjusted R^2 rises by 9 percentage points

to 62%. This happens because, over this sample, the CS and CA/GDP are negatively correlated. Since credit supply is so strongly positively related to house price growth, removing its effects by including it in the regression along with CA/GDP allows the regression to distinguish a modest positive role for the current account. Rows 4 and 5 show that an increase in CA/GDP is associated with a larger positive change in house price growth in periods during which there was an expansion of credit supply, or an increase in CS, but row 5 shows that the current account deficit does not influence the way CS is related to house price growth: the raw series and the interacted variable together explain the same amount of variation in house price growth as does CS alone. While this evidence is more supportive of a positive role for capital flows in affecting house prices, even here the results suggest that credit supply is the most quantitatively important variable related to future home prices.

Table 3: Quarterly Regressions for US (1990Q2-2010Q4)

	real House prices growth on				
Regression	*Cons*	*CA/GDP*	*CS*	(*CA/GDP*)* *CS*	Adj. R^2
1	-0.006	0.207			0.02
	(-1.35)	(0.92)			
2	0.001		0.016		0.53
	(0.27)		(9.94)**		
3	-0.011	0.365	0.017		0.62
	(-2.68)**	(2.54)*	(10.32)**		
4	-0.008	0.322		0.385	0.61
	(-2.33)*	(2.31)*		(10.74)**	
5	0.001		0.007	0.221	0.55
	(0.52)		(1.26)	(0.88)	

Notes: See Table 2. *CA* is current account deficit, GDP is real gross domestic product, both from the U.S. Department of Commerce, BEA. *CS* is a measure of credit standards from the SLO survey that gives the net percentage of banks that reported easier credit conditions. A positive value for this variable therefore indicates an easing of credit conditions, while a negative value indicates a tightening. Newey and West (1987) corrected t-statistics in parentheses. *significative at 5%, **significative at 1%.

For comparison, Table 4 shows output from the same regressions but over the subsample that only includes the recent boom-bust period: 2000:Q1 to 2010:Q4. The results with respect to the relation between CA/GDP and house price growth are little changed: even in this subsample the variable explains none of the variation in the growth of residential real estate

prices in a univariate regression. But the table shows that credit standards explains a much larger fraction of the variation in house price growth in this sample: CS now explains 66% of the quarterly variation in house price growth (row 2). Moreover, unlike the results for the full sample, in this subsample covering the boom-bust, the CA deficit adds nothing to our understanding of house price growth volatility, even when included in the regression along with CS. CA/GDP is statistically unrelated to house price growth in all of the regression specifications over this subperiod. To summarize, to the extent that the current account bears any relation to U.S. house price growth, it does so only in samples *prior to* the recent housing boom-bust. There is no relation between these variables in the recent boom-bust cycle.

Table 4: Quarterly Regressions for US (2000:Q1-2010:Q4)

	real House prices growth on				
Regression	*Cons*	*CA/GDP*	*CS*	*(CA/GDP)* CS*	Adj. R^2
1	-0.018	0.435			0.01
	(-0.96)	(1.02)			
2	0.002		0.023		0.66
	(0.43)		(11.43)**		
3	-0.008	0.214	0.023		0.66
	(-0.48)	(0.57)	(11.89)**		
4	-0.008	0.189		0.465	0.62
	(-0.40)	(0.45)		(10.60)**	
5	0.002		0.031	-0.162	0.65
	(0.46)		(1.84)	(-0.46)	

Notes: See Table 3.

Returning to the full sample, Table 5 shows the same regressions when we replace CA/GDP with our quarterly measure of the change in the net foreign holdings of total securities (the change in the U.S. net foreign liability position in securities), divided by trend GDP. We denote this variable ΔNFL_t. This variable has no effect on U.S. house price growth, whether it is included in the regression by itself or jointly with CS_t. The interaction of this variable with CS_t has a positive effect (row 4) but only if CS itself is excluded. Row 5 shows that, once we control for CS itself, the interaction variable has no marginal affect on house price growth. Otherwise the quantitative results are similar to those using the CA as a measure of capital flows.

Table 5: Quarterly Regressions for US (1990Q2-2010Q4)

	real house prices growth on				
Regression	*Cons*	ΔNFL_t	*CS*	ΔNFL_tx*CS*	Adj. R^2
1	0.003	-0.142			0.06
	(0.76)	(-1.46)			
2	0.001		0.016		0.53
	(0.27)		(9.94)**		
3	0.000	0.036	0.016		0.53
	(0.06)	(0.89)	(8.75)**		
4	0.001	0.143		0.135	0.25
	(0.18)	(1.54)		(4.94)**	
5	0.001		0.016	0.002	0.53
	(0.28)		(5.99)**	(0.15)	

Notes: See Table 3. ΔNFL_t is change in total net foreign U.S. Liabilities in total securities (where quarter-end positions have been estimated as described in the paper), divided by trend GDP. The trend is measured using a Hodrick and Prescott (1997) filter.

So far we have investigated only contemporaneous correlations between house prices, capital flows and credit standards. Tables 6 and 7 show results from *forecasting* regressions of real house price growth, $\Delta\ln(P)$ (Table 6) and the growth in the price-rent ratio, $\Delta\ln(P/R)$ (Table 7) in period $t+H$, on variables known at time t. We report results from long-horizon forecasts of these house price variables on CS_t by itself (row 1), on CS_t, ΔNFL_t and the real 10-year T-bond rate, r_t^{10} (row 2), and on CS_t, ΔNFL_t, r_t^{10}, and the growth in real GDP, ΔGDP_t (row 4). The first column of results reports results from the $H = 0$ ahead forecasts (contemporaneous correlations), and the following four columns report results from the 1, 2, 3, and 4 quarter ahead forecasts of these house price measures.

Table 6 shows that credit standards are strongly significantly related to the change in the log of real house prices at all future horizons. Indeed, this variable explains more than 47% of the one- and two-quarter ahead variation, and 41% of the three-quarter ahead variation. Only when explaining the four-quarter ahead variation in $\Delta\ln(P)$ does the explanatory power drop off, but even here credit standards explain over 20% of the one-year ahead variation. In general, the other explanatory variables add very little to the explanatory power of the regression, except at the four-quarter ahead horizon. For example, at $H = 0, 1$, and 2, none of ΔNFL_t, r_t^{10}, and ΔGDP_t are individually significant, and the adjusted R^2 is little changed from the univariate regression using CS_t as the sole predictor variable. The change in U.S. net foreign liabilities ΔNFL_t has a *negative* effect on house price growth, one quarter

ahead. The real interest rate does have a statistically significant negative effect on house price growth in the three- and four-quarter ahead regressions; this adds substantially to the R^2 only at $H = 4$. The coefficient on the net foreign liabilities indicator ΔNFL_t is statistically insignificant throughout.

Table 6: Regressions of $\Delta\ln(P)$ on CS, covariates (1991Q4-2010Q4)

U.S. Data

		Dependent Variable: $\Delta\ln P_{t+H}$				
		Forecast Horizon H				
Row	Regressors	0	1	2	3	4
1	CS_t	0.015	0.015	0.028	0.041	0.050
		(9.63)**	(7.00)**	(5.46)**	(4.76)**	(4.09)**
		[0.52]	[0.47]	[0.47]	[0.41]	[0.21]
2	CS_t	0.018	0.016	0.032	0.054	0.071
		(6.29)**	(4.11)**	(3.87)**	(4.71)**	(4.57)**
	ΔNFL_t	0.036	-0.026	0.018	0.218	0.435
		(0.79)	(-0.40)	(0.12)	(1.04)	(1.62)
	r_t^{10}	-0.004	-0.003	-0.009	-0.019	-0.027
		(-1.10)	(-0.72)	(-1.24)	(-2.33)*	(-2.31)*
		[0.53]	[0.48]	[0.49]	[0.53]	[0.50]
3	CS_t	0.017	0.016	0.033	0.054	0.070
		(6.19)**	(4.30)**	(4.25)**	(5.00)**	(4.67)**
	ΔNFL_t	0.058	-0.024	0.012	0.216	0.456
		(1.10)	(-0.36)	(0.07)	(0.87)	(1.37)
	r_t^{10}	-0.005	-0.003	-0.008	-0.019	-0.028
		(-1.23)	(-0.63)	(-0.99)	(-2.00)*	(-2.20)*
	ΔGDP_t	0.568	0.036	-0.153	-0.032	0.449
		(1.60)	(0.09)	(-0.19)	(-0.03)	(0.27)
		[0.54]	[0.47]	[0.48]	[0.53]	[0.49]

Notes: P is CoreLogic National House Price Index. CS is net percentage of banks reporting easing of credit; a positive value indicates more banks eased than tightened credit conditions with respect to previous quarter. Credit conditions have been standardized. ΔNFL is change in total net foreign holdings of securities, adjusted as described in the paper, divided by trend GDP. r^{10} is real 10-year bond yield: 10-year constant maturity rate in last month of quarter minus 10-year ahead inflation forecast: median response, from Survey of Professional Forecasters. ΔGDP is real GDP growth. Newey-West corrected t-statistic in parenthesis (lags $max\{H-1, 4\}$). Adjusted R^2 in brackets. * Significative at 5%. ** Significative at 1%.77 quarterly

observations.

The results in Table 7 for forecasts of the change in the log house-price rent ratio are similar. One difference is that CS explains more of the four-quarter ahead variation in this measure than it did of the four-quarter ahead variation in the change in log real house prices. In this case, even a year out, CS alone explains 46% of the variation in $\Delta\ln(P/R)$ and including the predictors ΔNFL_t, r_t^{10}, and ΔGDP_t allows the statistical model to explain an only an additional 9% of the variation. In these regressions ΔGDP_t helps explain the $H = 0$ (only) variation in price-rent changes, and the real 10-year T-bond rate is again significant at horizons $H = 3$ and 4.

Table 7: Regressions of $\Delta\ln(P/R)$ on CS, covariates (1991Q4-2010Q4)
U.S. Data

		Dependent variable: $\Delta\ln(P/R)_{t+H}$ Forecast Horizon H				
Row	Regressors	0	1	2	3	4
1	CS_t	0.015	0.015	0.043	0.055	0.064
		(8.36)**	(7.54)**	(7.38)**	(6.26)**	(5.08)**
		[0.49]	[0.46]	[0.52]	[0.51]	[0.46]
2	CS_t	0.017	0.015	0.049	0.070	0.087
		(6.08)**	(4.02)**	(5.46)**	(6.15)**	(5.49)**
	ΔNFL_t	0.013	-0.075	-0.056	0.097	0.312
		(0.29)	(-0.99)	(-0.31)	(0.46)	(1.14)
	r_t^{10}	-0.003	-0.003	-0.017	-0.029	-0.037
		(-0.95)	(-0.87)	(-1.80)	(-2.68)**	(-2.65)**
		[0.50]	[0.48]	[0.56]	[0.58]	[0.55]
3	CS_t	0.015	0.014	0.047	0.069	0.085
		(5.65)**	(4.04)**	(5.22)**	(5.93)**	(5.29)**
	ΔNFL_t	0.051	-0.068	-0.010	0.147	0.385
		(1.00)	(-0.83)	(-0.05)	(0.56)	(1.09)
	r_t^{10}	-0.006	-0.004	-0.019	-0.032	-0.041
		(-1.39)	(-0.85)	(-1.85)	(-2.79)**	(-2.90)**
	ΔGDP_t	0.977	0.181	1.088	1.142	1.601
		(2.36)*	(0.48)	(1.25)	(0.93)	(0.94)
		[0.54]	[0.48]	[0.57]	[0.58]	[0.55]

Notes: See Table 6. P/R is the aggregate house price index divided by rent as described in the notes to Figure 1.

To summarize, the evidence discussed above suggests that bank loan officers' accounts of their willingness to supply more mortgage credit, as distinct from their perceptions of the demand for credit, are strongly statistically related to house price movements, both and especially in U.S. data, as well as international data. By contrast, data on capital flows, real interest rates and GDP growth at best add modestly to explanatory power of these statistical models, and most of the time they are found to add nothing.

While this evidence strongly suggests that bank credit standards and credit availability matter for home values, with increases in credit supply associated with higher home prices, there is nothing in the evidence to suggest that such movements in standards are exogenous to the state of the economy or to expectations about future economic conditions, including the direction of future home prices. Nor is there any theoretical reason to expect them to be. As in classic financial accelerator models, endogenous shifts in borrowing capacity through fluctuations in collateral values imply that economic shocks have a much larger effect on asset prices than they would in frictionless environments without collateralized financing restrictions. Bank loan surveys on credit standards could in principle elicit information on either or both forms of a borrower's access to funds, and we have no way of knowing from the survey questions how much of any given change in standards is represented by one or the other. But we argue that this distinction is less relevant than it may at first seem. In our view, *either* of these represents a movement in credit availability, and both could be important for home prices.

Still, don't endogenous movements in credit supply raise a question of causality? If credit standards move in response to changing economic conditions, which in turn alter expectations of future home price movements, then how do we rule out the possibility that home prices predict credit availability rather than the other way around? The answer, we argue, is that we don't but we also shouldn't, because this issue is not really central for the question of whether credit availability plays a role in driving asset price fluctuations. A natural benchmark is a complete markets environment, where borrowing constraints and transactions costs play no role in the equilibrium allocations. Indeed, it's hard to understand why credit standards would be correlated at all with asset values in such an environment. But with incomplete markets, credit availability can have a large dynamic impact on asset prices even if fluctuations in that availability are completely endogenous. From this perspective, as long as we have a clean measure of credit supply, as distinct from credit demand, any correlation of credit supply with asset prices is of relevance for the question of whether credit supply matters for asset prices.

That said, we do want to distinguish movements in credit supply from movements in credit *demand*. The SLO survey explicitly asks banks to do so, but their may be some residual correlation between the two. Moreover, it is of some interest to ask whether, as a factual matter, the credit supply measure that we study has an exogenous component that still effects house prices once we control for expectations about future economic conditions.

Table 8 presents some evidence related to these questions. Instead of using the raw credit supply CS_t series, for these results we first regress it on the SLO survey's measure of credit demand (the net percentage of banks reporting higher credit demand), and take the *residual* of this regression $\epsilon_{CD,t}$, as a measure of credit supply. We also replace current GDP growth from the regressions above, with the *expected* GDP growth rate for the year ahead, $\Delta GDP_{t\rightarrow t+4}$, as measured by the Survey of Professional Forecasters (SPF) median forecast. In these regressions, we continue to use our measure of capital inflows ΔNFL_t and the real 10-year T-bond rate r_t^{10} as additional explanatory variables. Notice that r_t^{10} is itself a forward looking variable since it equals the nominal 10-year T-bond rate minus the *expected* 10-year inflation rate (also from the SPF, median forecast). Once we include these forecasts of future economic activity as additional predictor variables, any remaining role for our residual credit supply measure $\epsilon_{CD,t}$ in explaining house price movements must be independent of such expectations.

Table 8: Regressions of $\Delta\ln(P)$ on CS, covariates (1991Q4-2010Q4)

U.S. Data

		Dependent Variable: $\Delta\ln P_{t+H}$				
		Forecast Horizon H				
Row	Regressors	0	1	2	3	4
1	$\epsilon_{CD,t}$	0.085	0.080	0.148	0.211	0.259
		(7.20)**	(5.50)**	(4.39)**	(3.94)**	(3.54)**
		[0.48]	[0.43]	[0.41]	[0.40]	[0.35]
2	$\epsilon_{CD,t}$	0.098	0.085	0.172	0.284	0.376
		(5.28)**	(3.50)**	(3.22)**	(3.67)**	(3.58)**
	ΔNFL_t	0.023	-0.038	-0.012	0.164	0.368
		(0.51)	(-0.55)	(-0.08)	(0.71)	(1.23)
	r_t^{10}	-0.004	-0.003	-0.009	-0.020	-0.028
		(-1.13)	(-0.76)	(-1.22)	(-2.10)*	(-2.08)*
		[0.50]	[0.44]	[0.43]	[0.46]	[0.43]
3	$\epsilon_{CD,t}$	0.074	0.067	0.153	0.278	0.384
		(2.95)**	(2.11)*	(2.30)*	(3.01)**	(3.10)**
	ΔNFL_t	0.017	-0.044	-0.019	0.162	0.372
		(0.38)	(-0.61)	(-0.12)	(0.70)	(1.24)
	r_t^{10}	-0.002	-0.002	-0.007	-0.019	-0.028
		(-0.51)	(-0.36)	(-0.95)	(-1.99)	(-2.11)*
	$\Delta GDP_{t\to t+4}$	0.011	0.008	0.008	0.002	-0.003
		(2.32)*	(1.33)	(0.69)	(0.15)	(-0.15)
		[0.55]	[0.46]	[0.43]	[0.45]	[0.42]

Notes: P is CoreLogic National House Price Index. $\epsilon_{CD,t}$ is the residual from regressing CS (net percentage of banks reporting easing of mortgage credit) on a constant and the variable CD (net percentage of banks reporting higher demand for mortgage credit). ΔNFL is change in total net foreign holdings of securities, adjusted as described in the paper, divided by trend GDP. r^{10} is real 10-year bond yield: 10-year constant maturity rate in last month of quarter minus 10-year ahead inflation forecast: median response, from Survey of Professional Forecasters. $\Delta GDP_{t\to t+4}$ is median forecasted real GDP growth between periods t and $t+4$, from the Survey of Professional Forecasters. Newey-West corrected t-statistic in parenthesis (lags $max\{H-1,4\}$). Adjusted R^2 in brackets. * Significative at 5%. ** Significative at 1%.77 quarterly observations.

Table 8 shows that the residual credit supply measure $\epsilon_{CD,t}$ by itself explains about the same amount of variation in house price growth as does the raw series CS_t, and this is true no matter what other variables we include as additional regressors. The table also

shows that expected future GDP growth $\Delta GDP_{t\to t+4}$ has significant explanatory power for house price growth contemporaneously but does not help predict *future* house price growth, consistent with the notion that such expectations are reflected immediately in asset prices and collateral values. Even so, the residual supply measure $\epsilon_{CD,t}$ maintains its marginal explanatory power for even for contemporaneous movements in $\Delta \ln(P)$. When it comes to forecasting *future* house price changes, only the residual credit supply measure $\epsilon_{CD,t}$ displays any clear predictive power: expectations of future GDP growth, real interest rates, and the change in U.S. net foreign liabilities, all have a statistically negligible effect of $\Delta \ln(P_t)$. The change in U.S. net foreign liabilities ΔNFL_t again has a negative (but statistically insignificant) effect on house price growth, one and two quarters ahead. Moreover, the forecasting regressions for horizons ranging from one to four quarters ahead using all four predictor variables explains about the same amount of variation in future house price growth as does the univariate forecasting regression in row 1 using the residual credit supply measure alone. The results for changes in the log price-rent ratio (not reported) are very similar.

To summarize, even when we control for expectations about future economic conditions and purge the credit standards measure of any residual demand effects, we still find that credit supply has a statistically significant effect on house price movements, and is the most important economic determinant of future house price movements from one to four quarters ahead. In short, a tightening of credit supply today has an immediate and long-lasting effect on home prices that cannot be readily explained by changing expectations of future economic activity or real interest rates.

To form a crude basis for theoretical comparison, we close this section by presenting some implications from the model in FLVNa for regressions of this type. Recall that FLVNa explored the transition dynamics of the model over 10 years, as displayed in Figure 2 and described there, where the transition was calibrated to the boom-bust years from 2000-2010. Here we simply use the 10 years of data from that period to linearly relate $\Delta \ln(P/R)_t$ or $\Delta \ln(P_t)$ to the change in net foreign asset holdings over this period, and to a dummy variable indicating whether or not the economy is in a FML state. The 10 years of data are also influenced by the sequence of aggregate productivity shocks, which are chosen to mimic the business cycle and the relative performance of the housing and non-housing sectors over this period (see FLVNa for details). We define ΔNFL_t as the *change* in foreign holdings of domestic assets divided by the gross interest rate (i.e., multiplied by bond price) and divided by trend GDP. We define a FML dummy, CS, to be zero in period 1 (the year 2000), unity in periods 2 to 7 (2000-2006), and zero in periods 8-10 (2007-2010).

We then run "regressions" of $\Delta \ln (P/R)_t$ or $\Delta \ln (P_t)$ on CS and ΔNFL_t using the data from this transition, and report coefficients, t-statistics and R^2 statistics. We use quotations around the word 'regressions' here because this exercise (using only 10 observations) is best thought of as a convenient way of summarizing the marginal contributions of credit supply and capital flows for explaining house price changes over the period the model was calibrated to explain, rather than an actual regression analysis. In these "regressions," the only reason there is a residual at all (or an R^2 that differs from unity) is that the true relationship between these variables in the model is nonlinear, and because we have omitted the aggregate productivity shocks as explanatory variables. Table 9 presents the results.

Table 9: Regressions Based On Model Simulated Data

Panel A				Panel B			
	House P growth on				House P/R growth on		
Regression	ΔNFL	CS	R^2	Regression	ΔNFL	CS	R^2
1	0.1295		0.008	1	0.4177		0.067
	(0.291)				(0.748)		
2		0.063	0.370	2		0.095	0.626
		(2.21)				(3.618)	
3	-0.5605	0.0875	0.476	3	-0.505	0.117	0.690
	(-6.01)	(3.60)			(-4.009)	(5.724)	

Notes: The regression includes a constant even though it is not reported. ΔNFL is change in foreign holdings divided by the gross interest rate (i.e., multiplied by bond price) and divided by trend GDP. We multiply this number by 100 to express it as a percent. CS is a FML dummy, zero in period 1, 1 in periods 2 to 7, and zero in periods 8-10. r^f is risk-free rate. t-statistics are reported in parentheses. Data simulated from model of FLVNa, as described in the text.

Not surprisingly, given the transition displayed in Figure 2, the model implies that capital flows by themselves explain none of the variation in house price changes or price-rent changes. By contrast, the credit supply measure CS, which is positively related to house price fluctuations, explains quite a lot (37 percent of house price growth and 63% of the growth in the aggregate price-rent ratio). When both are included as explanatory variables, the measure of capital inflows ΔNFL_t becomes *negatively* related to house price growth, echoing some of the coefficient estimates from the analysis of historical data above. In the

model this occurs because capital inflows steadily rise throughout the boom period, when credit supply also rises. Once we control for credit supply, however (the more important factor driving the boom), the regression simply picks up the fact that capital inflows are negatively correlated with both credit supply and house price growth during the bust.

5 Conclusion

In this paper we have studied the correlations between house price changes and international capital flows, over the boom-bust period in the housing market from 2000-2010. We have argued that foreign capital flows into safe U.S. securities–U.S. Treasury and Agency bonds–played an important role in understanding the low interest rates in the last decade and quantitatively account for all of the upward long-term trend in the U.S. net foreign liability position over this period. Many countries that saw large housing booms and busts attracted foreign capital, as witnessed by their large current account deficits. Much of this capital seems to have found its way into residential investment and mortgage credit extension.

Despite these stylized facts, we have also argued that the same capital inflows that lowered interest rates and supported mortgage borrowing over this period had only a small impact on house prices. Although the housing boom was characterized by sharp increases in the rate at which capital flowed into the U.S., the bust occurred with no clear reversal in the trend toward healthy capital inflows into U.S. assets considered to be safe stores-of-value and integral to housing finance. While U.S. borrowing from abroad may ultimately have to decline, home values have not waited to do so, having already given up almost all of their gains during the boom years. This simple observation is reflected in our statistical analysis of the relationship between home prices, capital flows, credit standards, and interest rates, not only in the U.S., but also internationally: capital flows have little if any explanatory power for residential real estate fluctuations in samples that include both the boom and the bust. We have argued that capital inflows need not have a large affects on real estate prices if they simultaneously push up the housing risk premium and the expected stock of future housing.

A quantitatively meaningful account of the massive boom and bust in house prices must therefore rely on (at least one) additional argument. We argue here that the missing element is the financial market liberalization in the mortgage space (and its subsequent reversal), which made it easier and cheaper during the boom period for homeowners to purchase a house or to borrow against a existing home equity. The relaxation of credit constraints, by

itself, is a powerful force for higher house prices. Easier access to mortgage credit increases households' ability to withstand income shocks, and it reduces the risk premium households require to invest in risky assets like houses. In addition, lower transaction costs associated with new or refinanced home mortgages and home equity lines of credit raise the liquidity of houses, and therefore their price.

We have presented evidence that these mechanisms appear to have operated in the U.S., but also in countries other than the U.S. Using observations on credit standards, capital flows, interest rates and house prices, we find that, of these variables, it is credit standards and credit standards alone that explains home price fluctuations, with our measure of credit supply explaining 53% of the quarterly variation in U.S. house price growth over the period 1992-2010, and 66% over the boom-bust period from 2000 to 2010. By contrast, the other variables combined add less than 5% to the fraction of quarterly variation in house price changes explained, once we control for credit standards.

There are several interesting questions for future research. First, why is capital flowing from relatively productive economies, like China, Germany, Japan, Switzerland, etc. to relatively unproductive economies like Spain, the United States, Greece, and Italy? Second, why is capital flowing into safe assets like U.S. Treasuries? We have argued here that purchases of U.S. safe assets appear to be driven by reserve currency motives and political constraints by governmental holders in the source countries, but more research is needed on this issue. Third, why were the capital inflows directed towards housing (and mortgage-related products whose cash-flows directly depend on house prices), a relatively unproductive investment technology? We suspect that a relaxation of mortgage credit standards in some countries (the U.S., Spain, Ireland) and not in others (Germany, China, Japan) triggered the capital flow pattern. The financial market liberalization in the surplus countries itself may have originated from regulatory reform and the growth in securitization technology that differentially affected these countries. Finally, we have studied the relationship between *net* capital flows and house prices. Other researchers (notably Obstfeld (2011)), have argued that *gross* flows in international financial assets may lead to financial market instability. Future work should investigate the link between gross flows and prices of all kinds of assets, including real estate, equity, and bond markets.

6 Appendix

6.1 House Price Data

Data on house prices are deflated by a consumer price index (CPI). The data are from different national sources (See below), mostly quarterly, except for Germany (annual), Luxembourg (annual, and until 2009), Italy (semi-annual) and Japan (semi-annual). CPI is collected by Economist Intelligence Unit (EIU) and from national sources. For Slovenia series begins in 2003Q1; for Russia in 2001Q1.

6.1.1 United States

For regressions involving the house prices in the U.S., we use the Core Logic measure of house prices. This is a repeat-sales price index that is based on the universe of mortgages (conforming and non-conforming).[16] House prices are deflated using consumer price index (All Urban Consumers, U.S. city average, All items) from the Bureau of Labor and Statistics $http://www.bls.gov/cpi/$. The monthly data are averaged over the quarter and rebased at 2005=100, so real house price are in 2005 U.S. dollars. Regressions of growth rates use log changes, $log(HP_t) - log(HP_{t-1})$ for contemporaneous changes, and $log(HP_{t+H}) - log(HP_t)$ for H horizon changes.

For regressions using the aggregate price-rent ratio, we construct an index by combining a measure of rent, for primary residences, constructed from the Shelter component of the Consumer Price Index for all urban consumers, SA, last month of each quarter, with a measure of housing prices. Data for rent are available from the Bureau of Economic Analysis of the U.S. Department of Commerce. House price is the CoreLogic National House Price Index (SA, Jan.2000=100). The price-rent ratio has been normalized to equal the level, in 1975:Q4 of the quarterly Price-Rent ratio constructed from the flow of funds housing wealth and National Income and Products data on housing consumption.

6.1.2 International Data

House prices are deflated using consumer price indices for each country, from national sources. For the Euro Area we deflate with the Harmonized Index of Consumer Prices for the Euro

[16] Other indexes are available only for conforming mortgages. For example, the Federal Housing Finance Administration (FHFA) measure is based only on conforming mortgages and therefore misses price changes associated with non-conforming mortgages. Like the Core Logic measure, the Case-Shiller measure is also based on the universe of mortgages, but it has substantially smaller geographic coverage than the Core Logic measure.

Area 17, all-items. Data sources for residential real estate are given in the table below.

Australia	Australian Bureau of Statistics. Exist. dwellings(8 CITIES),PER DWEL.,Q-ALL NSA. (1)
Austria	Oesterreichische National Bank. All dwellings(VIENNA),PER SQ.M.,Q-ALL NSA. (1)
Belgium	EXISTING DWELLINGS, PER DWEL.,Q-ALL NSA. (1)
Canada	Teranet- National Bank of Canada. Comp. 6 Cities (monthly, averaged to quarterly using sales pair co
Czech Republic	Czech Statistical Office. House prices.
Denmark	Statistics Denmark. ALL SINGLE-FAMILY HOUSE,PER DWEL,Q-ALL NSA.
Estonia	Statistical Office of Estonia. Av. price per sq.m., 2-rooms and kitchen, Tallinn (1). 2009 onwards, 55-7
Euro Area	European Central Bank. Euro area 17 (fixed composition); New and existing dwellings; Not S.A.
Finland	Statistics Finland. EXISTING HOUSES,PER SQ.M,Q-ALL NSA. (1)
France	National Institute of Statistics and Economic Studies. Existing Dwellings, Q-ALL NSA. (1)
Germany	Deutsche Bundesbank, based on data provided by BulwienGesa AG. Existing Dwellings, Y-ALL NSA.
Greece	Bank of Greece. ALL DWELL.(URBAN GREECE EX.ATHENS),PER SQ.M, NSA. (1)
Hungary	FHB Bank. FHB House Price Index (actual buying and selling transaction data of residential real esta
Iceland	Icelandic Property Registry. ALL DWELLINGS(GR. REYKJAVK),PER SQ.M,M-ALL NSA. (1)
Ireland	Economic and Social Research Institute (ESRI). ALL DWELLINGS, PER DWELLING., Q-ALL NSA.
Israel	The Central Bureau of Statistics. Prices of Dwellings (Until September 2010 - Owner Occupied Dwelli
Italy	Bank of Italy. ALL DWELLINGS, PER SQUARE M., H-ALL NSA. (1)
Korea	Kookmin Bank in Korea. ALL DWELLINGS, M-ALL NSA. (1)
Luxembourg	Central Bank of Luxembourg. ALL DWELLINGS, Y-ALL NSA. (1)
Netherlands	Statistics Netherlands. Existing own homes. Dwellings: all. Price index purchase prices.
New Zealand	The Reserve Bank of New Zealand. All dwellings, Q-ALL NSA. QVL. (1)
Norway	Statistics Norway. All dwellings, Q-AVG,NSA. (1)
Poland	Central Statistics Office. Price of a square meter of usable floor space of a residential building.
Portugal	Inteligência de Imobiliário. All dwellings, PER SQUARE METER,M-ALL NSA. (1)
Slovenia	Statistical Office of The Republic of Slovenia. Existing Dwellings,Q-ALL NSA. (1)
Spain	Bank of Spain. All dwellings, PER SQUARE M., Q-ALL NSA. (1)
Sweden	Statistics Sweden. ALL OWNER-OCCUPIED DWELLINGS,PER DWEL.,Q-ALL NSA. (1)
Switzerland	Swiss National Bank. ALL 1-FAMILY HOUSES,PER DWELLING,Q-ALL NSA.(1)
United Kingdom	Office for National Statistics. All dwellings (ONS), PER DWEL.,M,Q-ALL NSA.(1)
United States	Federal Housing Finance Agency. Family Houses, Q-ALL NSA (all transactions, FHFA).(1)
Russia	Federal State Statistics Service. EXISTING DWELLINGS,PER SQUARE M,Q-ALL NSA.(1)
China	National Bureau of Statistics of China. Land prices, Resid. and Commercial, Q-ALL NSA. (1)

Notes: Series (1) can be found at the Bank of International Settlements website, *http* : *//www.bis.org/statistics/pp.htm*

6.2 Current Account Data

Current Account is measured as the current account *deficit.* Data are available from the International Monetary Fund for all countries except Ching; for China data are from State

Administration of Foreign Exchange, SAFE. CA balances are accumulated and deflated by 2006 GDP (collected by Economist Intelligence Unit, EIU), in current US dollars. For Belgium, CA data 2000Q1-2001Q4 are from the National Bank of Belgium, provided by the Organization of Economic Cooperation and Development (OECD).

United States: $(CA/GDP)_t$ is current account *deficit* over nominal GDP, at current market prices. Balance of current account is from the Bureau of Economic Analysis, U.S. International Transactions Accounts Data, in millions of dollars, not seasonally adjusted. GDP data is from the Bureau of Economic Analysis, National Economic Accounts, in billions of current dollars, seasonally adjusted at annual rates; we transform it to quarterly rates dividing by four.

International Data: $(CA/GDP)_t$ is current account *deficit* over nominal GDP, at current market prices. Balance of current account data is from International Monetary Fund, in millions of US dollars, and GDP data are collected by the Economist Intelligence Unit on nominal GDP ($US\$$), quarterly.

6.3 Residential Investment

Residential Investment data are from Eurostat and national sources, as indicated in the table below. Residential investment is accumulated and deflated by 2006 GDP, all in current national currency. For France, Hungary, Poland, Switzerland and Russia, data is available only until 2009.

Australia	Australian Bureau of Statistics. Private Gross fixed capital formation, Dwellings, Total , $ millions
Austria	Eurostat.
Belgium	National Bank of Belgium. Gross fixed capital formation, dwellings, current prices, millions of Euros.
Canada	Statistics Canada, Capital formation in residential structures, current prices, national currency, NSA.
Czech Republic	Eurostat.
Denmark	Eurostat.
Estonia	Eurostat.
Euro Area	
Finland	Eurostat.
France	Eurostat.
Germany	Eurostat.
Greece	Eurostat.
Hungary	Eurostat.
Iceland	Eurostat.
Ireland	Central Statistics Office Ireland. Gross Dom. Physical Cap. Formation at current prices. Fixed capital
Israel	Central Bureau of Statistics. Gross Domestic Capital Formation. Residential Buildings, current prices.
Italy	Eurostat.
Japan	Cabinet Office, Gov. of Japan, Billions of Yen, nominal Private residential investment, not SA.
Korea	Bank of Korea, Gross fixed capital formation in residential buildings, current prices, Bil. Won.
Luxembourg	Eurostat.
Netherlands	Eurostat.
New Zealand	Statistics New Zealand. Gross Fixed Capital Formation, Residential buildings, current prices, $ Million
Norway	Norway Statistics. Gross fixed capital formation, Dwellings (households). Current prices (mill. NOK)
Poland	Eurostat.
Portugal	Eurostat.
Slovenia	Eurostat.
Spain	Eurostat.
Sweden	Eurostat.
Switzerland	Eurostat.
United Kingdom	Eurostat.
United States	US Bureau of Economic Analysis. Fixed Investment, Residential, Billions of dollars.
Russia	Federal State Statistics Service. Investment in fixed capital, in residential houses, current prices, billion
China	National Bureau of Statistics China (NBSC). Total Inv. Residential Buildings in the whole country, mi

Notes: Most of the series are from national sources via Eurostat. Gross fixed capital formation, in construction work: housing, millions of national currency, current prices.

6.4 Data on Credit Standards

6.4.1 United States

The variable CS_t is a net percentage index that indicates the percentage of banks relaxing credit standards for mortgage loans (both *'considerably'* and *'somewhat'*), with respect to the previous quarter, minus the percentage of banks tightening credit standards (both *'considerably'* and *'somewhat'*). This indicator is taken from the Senior Loan Officer Opinion Survey on Bank Lending Practices for the US, published by The Federal Reserve. They report the net percentage of banks *tightening* standards. The negative of this is the net percentage of banks easing standards, which we use in our empirical work. This is a quarterly survey of approximately sixty large domestic banks and twenty-four U.S. branches and agencies of foreign banks. Questions cover changes in the standards and terms of the banks' lending and the state of business and household demand for loans. These data are available since May 1990, when the survey then began including approximately 20 questions designed to measure changes in credit standards and terms on bank loans and perceived changes in the demand for bank credit. See $http://www.federalreserve.gov/boarddocs/snloansurvey/$. We focus on the question that asks about residential mortgage loans at each bank. From 1990Q2 (beginning of the Survey) to 2006Q4, the question is about residential mortgage loans in general: *'Over the past three months, how have your bank's credit standards for approving applications from individuals for mortgage loans to purchase homes changed?'*. The recommendations for answering this question state *'...If your bank's credit standards have not changed over the relevant period, please report them as unchanged even if the standards are either restrictive or accommodative relative to longer-term norms. If your bank's credit standards have tightened or eased over the relevant period, please so report them regardless of how they stand relative to longer-term norms. Also, please report changes in enforcement of existing standards as changes in standards.'* See $http://www.federalreserve.gov/boarddocs/snloansurvey/$.

From 2007Q1 onwards, the question is asked for each of three categories of residential mortgage loans: prime residential mortgages, nontraditional residential mortgages, and subprime residential mortgages. The recommendations state *'For the purposes of this survey, please use the following definitions of these loan categories (note that the loan categories are not mutually exclusive) and include first-lien loans only: The prime category of residential mortgages includes loans made to borrowers that typically had relatively strong, well-documented credit histories, relatively high credit scores, and relatively low debt-to-income ratios at the time of origination. This would include fully amortizing loans that have a*

fixed rate, a standard adjustable rate, or a common hybrid adjustable rate—those for which the interest rate is initially fixed for a multi-year period and subsequently adjusts more frequently. The nontraditional category of residential mortgages includes, but is not limited to, adjustable-rate mortgages with multiple payment options, interest-only mortgages, and "Alt-A" products such as mortgages with limited income verification and mortgages secured by non-owner-occupied properties. (Please exclude standard adjustable-rate mortgages and common hybrid adjustable-rate mortgages.) The subprime category of residential mortgages typically includes loans made to borrowers that displayed one or more of the following characteristics at the time of origination: weakened credit histories that include payment delinquencies, chargeoffs, judgments, and/or bankruptcies; reduced repayment capacity as measured by credit scores or debt-to-income ratios; or incomplete credit histories.'

The answer to this question can be one of the following: tightened considerably, tightened somewhat, remained basically unchanged, eased somewhat, eased considerably. Responses are grouped in 'Large Banks' and 'Other Banks'. The index, however, is calculated using information of "all respondents". Given that the question is referenced to the past three months, we date the index with respect to the quarter when changes to lending standards occurred (as opposed to when the responses are collected, i.e., net percentage reported in July 2011 is the net percentage for 2011Q2). In the report beginning in 2007:Q1, a distinction is made between prime and subprime mortgages in the survey. Figure 10 reports the net percentage for prime only and also for a weighted average of prime and subprime mortgages: (net percentage easing on prime)*weight + (net percentage easing on subprime)*(1-weight), where weight is 0.75 for 2007 and 0.95 for 2008. After that weight =1, because no bank reported that they originated sub-prime mortgages. The earlier weights are based on the paper

$http://www.jchs.harvard.edu/publications/finance/UBB10-1.pdf$, page 85, figure 1-3. These numbers are approximately the average share of banks that originated subprime residential mortgages, according to the Survey (23% for 2007 and 8% for 2008). Results are not sensitive to using one or the other number.

We standardize the net percentage indicator by subtracting the mean and dividing by the standard deviation.

6.4.2 International Data

CS_t stands for credit standards for housing loans. Data are from bank lending surveys conducted by national central banks, and the European Central Bank. The survey questions are

modeled after the U.S. Survey of Senior Loan Officers. Central Banks report the information in different ways. Some report net percentages, some report diffusion indices, and some report mean values. *Net percentage* is the percentage of banks loosening (or tightening) credit standards with respect to the previous quarter minus the percentage of banks tightening them (or relaxing). *Diffusion index* is the percentage of banks loosening (or tightening) credit standards "considerably" with respect to the previous quarter multiplied by 1 plus the percentage of banks loosening (or tightening) credit standards "somewhat" multiplied by 0.5 minus the percentage of banks tightening (or relaxing) "somewhat" times 0.5, minus the percentage of banks tightening (or relaxing) "considerably" times 1. *Mean values*: each answer receives a value from 1 to 5 (where for 5, the bank reported that relaxed the credit standards "considerably", 3 didn't change them, and 1 is a "considerable" tightening), and the mean value for each quarter is reported.

Because mean values are a scale transformation of the diffusion index, but the net percentage indicator is not, we report results for a subset of 9 countries for which we can construct either the diffusion index (Austria, Belgium, Euro Area, France, Korea, Portugal, Spain, United States) or a mean value (Ireland) with the information reported by the Central Banks. We also report results for a larger set of 11 countries for which we have only information on the net percentage (Canada and Netherlands). We standardize these indices, country by country, by subtracting the mean and dividing by the standard deviation, for the period 2002Q4-2010Q4.

A positive value for CS_t reflects easing credit conditions with respect to previous quarter, and units are in terms of standard deviations. For Austria, Belgium, Euro Area, France, Ireland, Netherlands, Portugal, Spain, the survey is based on the Bank Lending Survey conducted by the European Central Bank. See $http://www.ecb.int/stats/money/surveys/lend/html/index.$

Results for each country are not publicly available, but each of the countries above publishes an indicator (net percentage, mean index, diffusion index) that reflects the change in credit conditions in the country.

The European Central Bank's website states "*...the survey addresses issues such as credit standards for approving loans as well as credit terms and conditions applied to enterprizes and households. It also asks for an assessment of the conditions affecting credit demand."*

"The survey is addressed to senior loan officers of a representative sample of euro area banks and will be conducted four times a year. The sample group participating in the survey comprises around 90 banks from all euro area countries and takes into account the characteristics of their respective national banking structures."

We focus on question 8: *"Over the past three months, how have your bank's credit standards as applied to the approval of loans to households changed?"*, Item 8.1. Loans for house purchase. Respondents can reply one of the following answers: tighten considerably, tighten somewhat, basically unchanged, ease somewhat, ease considerably.

For the United States, we construct a diffusion index from the data reported by the Senior Loan Officer Opinion Survey on Bank Lending Practices for the US, published by The Federal Reserve. Data for 2007 onwards is weighted average of prime and sub-prime mortgages, with the following weights for prime mortgages: 2007, 0.75; 2008, 0.95; 2009 and 2010, 0. See above.

For Korea, we use the diffusion index reported by The Bank of Korea, Financial System Review, which conducts a Survey on Bank Lending Practices since 2002. See $http://ecos.bok.or.kr/EIndex_en.jsp$ We use the information from the Lending attitude diffusion index for Household housing.

For Canada the data is for general credit conditions (not specifically housing). The information is from the Overall Balance of Opinion reported by the Bank of Canada, in its Senior Loan Officer Survey.For more info about the Survey see $http://www.bankofcanada.ca/wp-content/uploads/2011/07/senior_loan_officer_survey_backgrounder.pdf$

Data sources for each country are given the table below.

Australia	
Austria	Oesterreichische Nationalbank. Bank Lending Survey. Diffusion Index.
Belgium	Nationale Bank van Belgie. Bank Lending Survey. Diffusion Index.
Canada	Bank of Canada. Senior Loan Officer Survey. Lending conditions: Balance of Opinion, net percenta
Czech Republic	
Denmark	Danmarks Nationalbank. BLS. Lending to Households - existing customers (All Banks). Diff. Inde
Estonia	
Euro Area	European Central Bank.
Finland	
France	Banque de France. Bank Lending Survey. Net percentage.
Germany	Deutsche Bundesbank. Bank Lending Survey. Diffusion Index.
Greece	
Hungary	Magyar Nemzeti Bank. BLS. Question 3/A - Credit Standards. Net percentage.
Iceland	
Ireland	Central Bank of Ireland. Bank Lending Survey. Mean.
Israel	
Italy	Banca D'Italia. Bank Lending Survey. Diffusion Index.
Japan	Bank of Japan. Senior Loan Officer Opinion Survey Bank Lending Practices at Large Banks. Diff.
Korea	Bank of Korea, Financial System Review. Survey Bank Lending Practices. Lending attitude. Diff.
Luxembourg	Banque centrale du Luxembourg. BLS. Net percentages from seven-bank sample. Unweighted.
Netherlands	De Nederlandsche Bank. Bank Lending Survey. Net percentage.
New Zealand	
Norway	Norges Bank. Survey of Bank Lending. Lending to households, Loans secured on dwellings, diffusio
Poland	National Bank of Poland. Senior Loan officer opinion. Net percentage.
Portugal	Banco de Portugal. Bank Lending Survey. Diffusion index.
Slovenia	
Spain	
Sweden	
Switzerland	
United Kingdom	Bank of England. Credit Conditions Survey. Secured credit to households, for house purchase. Diff
United States	Federal Reserve. Senior Loan Officer Opinion Survey. Net percentage.
Russia	
China	

Notes: Most of the surveys follow the BLS conducted by the European Central Bank. In that survey, the questions attained are Q8.1 and Q13.1. For the Euro Area, the results come from net percentages.

6.5 Estimation Details

6.5.1 U.S. Regressions

For the contemporaneous quarterly regressions we report Newey-West corrected standard errors (and t-statistics) using 4 lags. For the long horizon quarterly regressions we use lags $= max\{H-1, 4\}$, to take into consideration the use of overlapping data.

6.5.2 Panel Regressions

We use a balanced panel from 2002Q4 to 2010Q4, for 10 countries plus the Euro-Area: Austria, Belgium, Canada, Euro Area, France, Ireland, Korea, Netherlands, Portugal, Spain, United States. The choice of sample period is determined by the availability of a balanced panel for data on credit standards (European Central Bank conducts the Bank Lending Survey since 2002Q4), and quarterly house prices (For Italy and Germany we only have annual data on house prices, for Hungary only semi-annual data on credit conditions, and for Poland we only have data on credit conditions since 2003Q4). We also use a subsample of 9 countries where we drop Canada and Netherlands (see information on credit standards, above). The Euro Area consists of 17 countries: Austria, Belgium, Cyprus, Estonia, Finland, France, Germany, Greece, Ireland, Italy, Luxembourg, Malta, the Netherlands, Portugal, Slovakia, Slovenia, and Spain. We construct log changes, $log(HP_t) - log(HP_{t-1})$ for contemporaneous changes, and $log(HP_{t+H}) - log(HP_t)$ for H horizon changes.

For the contemporaneous quarterly regressions, we report the robust standard errors (and t-statistics) using the Driscoll-Kraay statistic, with lags=3 (default). For the quarterly long horizon regressions, we use instead number of lags $= max\{H-1, 3\}$ to take into account the use of overlapping data.

For more on Driscoll-Kraay statistic, see :

"Robust Standard Errors for Panel Regressions with Cross-Sectional Dependence", by Daniel Hoechle, http://fmwww.bc.edu/repec/bocode/x/xtscc_paper.pdf.

and Driscol and Kraay (1998).

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