AI and Productivity: The Role of Innovation*

Viral V. Acharya† Xuehua Shao‡ Zhaoxia Xu§

Abstract

This paper examines how firms' adoption of artificial intelligence (AI) affects productivity via the innovation channel. Using a novel firm-level AI adoption measure based on the timing of AI product installations, we find that adopters subsequently increase patenting relative to non-adopters. Post-adoption patents receive more citations, include more claims, and contain a higher share of patents that build on the firm's existing technologies, while also exhibiting greater originality and generality. The effects are stronger for firms with a more focused business scope and are not concentrated among larger, financially unconstrained, or high-tech firms. Evidence on mechanisms points to AI improving efficiency, shifting inventor composition, and enhancing knowledge recombination. Additionally, adopters increase R&D but not capital expenditures, show no significant changes in operating costs, and exhibit higher productivity and market value. Quantitatively, our estimates imply that AI adoption increases aggregate value-added total factor productivity by approximately 0.61% in a representative post-adoption year. Taken together, these results indicate that AI functions as an innovation-enabling intangible investment that supports productivity growth.

Keywords: Artificial Intelligence, Productivity, Intangible investment, Knowledge recombination, R&D

JEL Classification: D22, D24, E22, L25, O31, O33

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†New York University Stern School of Business, NBER, ECGI, and CEPR. Email: vva1@stern.nyu.edu

‡University of New South Wales. Email: xuehua.shao@unsw.edu.au

§University of New South Wales. Corresponding author. Email: zhaoxia.xu@unsw.edu.au

1 Introduction

> "*As artificial intelligence evolves, we must remember that its power lies not in replacing human intelligence but in augmenting it. The true potential of AI lies in its ability to amplify human creativity and ingenuity.*"
>
> —Ginni Rometty, *Former CEO of IBM*

As artificial intelligence (AI) is increasingly adopted by firms, a growing literature debates whether and how it affects productivity growth. On the one hand, task-level evidence documents sizable productivity gains in specific workflows such as customer support (Brynjolfsson et al., 2025). On the other hand, AI adoption can require costly adjustment and complementary investments in processes and organization, generating J-curve dynamics in firm productivity (McElheran et al., 2025). Reflecting these mixed micro-level findings, macro aggregation exercises that translate task-level cost savings into aggregate total factor productivity (TFP) imply productivity gains that range from modest to moderate (Acemoglu, 2025; Aghion and Bunel, 2024).

An important unresolved question in this debate concerns the channels through which AI affects productivity. In task-based frameworks, AI can raise productivity directly by automating or augmenting existing tasks and lowering costs. However, AI may also raise productivity indirectly by increasing research productivity and reshaping the innovation process, consistent with the view of AI as a method of invention (Cockburn et al., 2018). Prior studies offer mixed predictions about how AI affects firm innovation. One view is that AI can strengthen innovation by improving strategic foresight and search and by using prediction to accelerate knowledge discovery and target high-value opportunities (Mühlroth and Grottke,

2020; Cockburn et al., 2018; Agrawal et al., 2019).[1] An alternative view emphasizes frictions and unintended consequences, including coordination failures and underinvestment in complements, concentration of innovative advantages due to scale and data, path dependence from prediction trained on historical data, and weaker incentives if AI-enabled patenting congests patent systems and erodes disclosure quality (Bresnahan and Trajtenberg, 2017; Cockburn et al., 2018; Agrawal et al., 2024; Ouellette et al., 2025). In light of these mixed predictions, it is important to systematically assess how AI adoption changes firm innovation.

One challenge in testing these competing hypotheses is measuring *firm-level* AI adoption and observing adoption timing (Raj and Seamans, 2019). Prior studies often use AI-related job postings as a proxy for AI investment (Babina et al., 2024). This approach may understate true adoption if firms do not publicly advertise AI positions or already employ workers with relevant skills. Another common approach constructs task-level exposure to AI using Occupational Information Network (O∗NET) data and aggregates them to the occupation-, firm-, or industry-level (Eloundou et al., 2024; Acemoglu, 2025; Eisfeldt et al., 2025; Felten et al., 2021). While informative, these proxies capture potential exposure rather than realized, firm-specific AI adoption and do not identify adoption timing. In departure from these studies, we leverage newly available historical data on IT product installations to construct a direct firm-level measure of AI adoption. This measure allows us to identify the timing of adoption and to examine its its subsequent effects on firm-level innovation and productivity. More broadly, our approach responds to recent calls in the literature for more direct measures of firm-level

[1] For instance, Microsoft employed Azure Quantum Elements, an AI-powered research platform, to accelerate scientific discovery in chemistry and materials science. To design a solid state battery electrolyte that outperforms those engineered decades ago, scientists face the challenge of screening a vast number of potential molecular combinations that can create an electrolyte (McPhee, 2024). AI models can help overcome this challenge by predicting the properties of electrolyte mixtures at scale. Scientists in Microsoft first used AI models to screen 32.6 million hypothetical materials and reduced the pool to about 500,000. Building on this initial screen, they coupled the models with physics based simulations, which narrowed the set to about 800 and then to about 150. In the final stage, AI property-prediction filters together with expert criteria selected 18 laboratory candidates, from which the lead electrolyte was synthesized and validated. This end-to-end discovery process, which would typically require years, was completed in less than nine months with the assistance of AI (McPhee, 2024; Microsoft, 2024). Microsoft subsequently filed a patent application covering the resulting electrolyte compositions (publication No. US2025/0105343 A1). By reducing the time and experimental inputs required to identify viable materials, AI increases the efficiency of the innovation process, thereby improving productivity in research and development.

AI adoption (Babina, 2026).

Given the staggered adoption of AI across firms, we implement a stacked difference-in-differences (DiD) design that compares changes in innovation outcomes between firms whose initial AI adoption occurs in a given calendar year (AI-adopting firms) and firms that have not yet adopted at that time (non-adopting firms).[2] For each adoption-year cohort, we include only not-yet-treated firms as clean controls, thereby mitigating bias from treatment effect heterogeneity under staggered adoption (Baker et al., 2022).

Using data on granted patents filed by publicly listed U.S. firms to measure innovation outcomes, we find that AI-adopting firms experience subsequent increases in both the quantity and quality of their patents relative to non-adopting firms. Patents produced by AI-adopting firms receive more citations after adoption than those of non-adopting firms, indicating higher scientific and technological value. Furthermore, AI adoption also influences firms' innovation characteristics along three key dimensions: novelty, direction, and reach. In terms of novelty, AI-adopting firms produce patents that exhibit greater originality and span more technologically distant fields, suggesting that AI enables firms to explore a broader and less familiar technological domains. Regarding innovation direction, we find that AI adoption shifts firms toward more exploitative innovation, which leverages their existing technological capabilities to generate incremental improvements. Finally, in relation to innovation reach, we observe that patents produced by AI-adopting firms demonstrate higher generality, implying broader applicability and influence across diverse technology areas. These firms also tend to file patents with a greater number of claims, suggesting a broader scope of the protected invention and greater legal protection.

A natural concern is that AI adoption may be an endogenous choice and that unobservable factors could simultaneously influence both a firm's decision to adopt AI and its innovation performance. For instance, firms that invest heavily in innovation might be more likely to adopt AI and tend to achieve better innovation outcomes regardless of AI adoption. To alleviate this

[2]We use the terms AI-adopting (non-adopting) firms, AI-using (non-using) firms, AI (non-AI) firms, and adopters (non-adopters) interchangeably throughout the paper.

concern, we first compare average R&D investment between AI-adopting and non-adopting firms and find that, on average, adopting firms spend significantly less on R&D than non-adopting firms. Moreover, our stacked DiD specification includes firm-by-cohort and year-by-cohort fixed effects, which help isolate the effect of AI adoption from unobservable time-invariant firm heterogeneity and cohort-specific shocks. We further estimate a cohort-specific event study with event-time leads and lags around the adoption year to trace the dynamic effects of AI adoption. The results show that innovation increases following adoption and does not precede it.[3]

To address remaining endogeneity concerns, we also employ a control function (CF) approach using firms' exposure to AI-strong universities developed by Babina et al. (2024) as an instrumental variable (IV) for AI adoption.[4] This instrument satisfies the relevance condition, as firms with well-established hiring networks with AI-strong universities are more likely to access AI-trained talent and benefit from AI-related knowledge spillovers, thereby increasing their likelihood of adopting AI technology. It plausibly satisfies the exclusion restriction because the instrument is measured in 2010 before the commercial boom in AI and is not simply capturing broader institutional prestige. To further disentangle the effect of AI-specific talent from general human capital quality, we control for firms' exposure to top-10 universities following Babina et al. (2024). Our results on innovation quantity, quality, and innovation characteristics remain robust after accounting for potential endogeneity in AI adoption.

We next examine the mechanisms through which AI adoption affects innovation. The evidence points to three complementary channels: improvements in innovation efficiency,

[3] Additionally, a placebo test in which AI adoption is randomly assigned shows that these fictitious adopters do not outperform non-adopting firms in innovation in the absence of actual AI adoption, indicating that our results are not driven by selection or unobserved firm characteristics correlated with AI adoption. To further mitigate the self-selection concern that AI adopters differ from non-adopters in observed characteristics, we implement entropy balancing to ensure that AI-adopting and non-adopting firms are comparable in terms of key pre-treatment firm characteristics. This reweighting approach enhances the validity of our comparisons by addressing observable differences that may simultaneously influence both AI adoption and innovation outcomes.

[4] An AI-strong university is defined as a university whose pre-2010 AI research strength ranks in the top 5% of all universities when comparing the number of AI researchers across universities (Babina et al., 2024).

shifts in inventor composition, and enhanced recombination of knowledge. Relative to non-adopting firms, AI adopters exhibit higher inventor productivity and greater innovation efficiency. Adopting firms expand their inventor workforce, hire more first-time inventors, and increase the share of inventors with AI-related expertise. In addition, AI adopters form larger and more distinctive inventor teams with broader core technical skills, consistent with an increased capacity to recombine knowledge across technological domains. These results indicate that AI adoption reshapes the innovation process in ways that enhance research productivity and strengthen firms' capacity to generate new technological knowledge.

Interestingly, cross-sectional analyses show that the impact of AI is generally more pronounced among focused firms. This result is contrary to our prior that AI would disproportionately benefit diversified firms operating in more complex information environments, given its capacity to process and learn from large volumes of data. One plausible explanation is that focused firms tend to generate high-quality, coherent, and consistent data from their operation, which enhances the effectiveness of AI applications. This finding is consistent with the view that the quality of firm-specific datasets plays a critical role in determining the performance of AI technologies (Brynjolfsson and McAfee, 2017; Brynjolfsson et al., 2019). Other factors that may enable focused firms to outperform include deeper domain expertise, less internal bargaining, and fewer coordination frictions.

Additionally, the innovation effects of AI adoption are not stronger for larger or financially unconstrained firms, even though these firms are typically viewed as better positioned to bear the fixed costs of AI adoption, build the necessary data and computing infrastructure, and exploit AI's returns to scale. We also show that the results are not driven by firms operating in high-tech sectors and that AI's impacts on innovation are not confined to AI-related patents, underscoring its broader economic significance.

Furthermore, we find that AI adopters spend more on R&D and have higher market value, while physical investment and operating costs do not exhibit significant changes relative to non-adopters. These patterns suggest that investors view AI primarily as innovation-

enabling intangible capital rather than as a driver of tangible capital deepening or immediate cost reductions. Consistent with this view, AI adopters experience an approximately 8.87% increase in value-added TFP after adoption relative to non-adopters. At the same time, non-adopters exhibit a decline in value-added TFP, consistent with spillover or competitive effects. Aggregating these firm-level productivity changes using a first-order Hulten-style aggregation (Hulten, 1978) implies a contribution of approximately 0.61% to aggregate value-added TFP growth in a representative post-adoption year. Taken together, these results are consistent with the innovation channel linking AI adoption to productivity growth.

2 Related Literature

A growing literature studies the economic implications of AI at the firm and aggregate levels. Prior work documents effects of AI on firm growth and product innovation (Babina et al., 2024), patent examination (Zheng, 2025), disclosure strategies (Cao et al., 2023), firm value (Eisfeldt et al., 2025), industry entry and exit dynamics (Lu et al., 2024), and employment (Chen and Wang, 2024). However, existing studies offer mixed predictions about AI's impacts on innovation, and comprehensive empirical evidence remains limited.[5] Using a novel firm-level measure that directly captures AI adoption timing, we show that adopters systematically reshape their innovation activities, increasing both the quantity and quality of patenting and shifting the characteristics of their innovations.

Our findings also contribute to a broader debate on whether AI fundamentally reshapes the innovation production function with implications for long-run economic and technological

[5] One perspective is that AI can raise innovation by making more information and knowledge usable in R&D through unstructured data processing (Ludwig and Mullainathan, 2024) and automated retrieval and combination of knowledge across technological domains (Agrawal et al., 2019), with related benefits from patent text and claim analytics (Setchi et al., 2021; Alderucci and Sicker, 2019). The other perspective is that AI can weaken innovation when complementarities and coordination frictions impede effective deployment (Bresnahan and Trajtenberg, 2017), when control of critical data and algorithms concentrates innovation among early leaders (Cockburn et al., 2018), and when AI-enabled prediction reinforce existing technological trajectories (Agrawal et al., 2024). AI-generated prior art may flood the system with superficially plausible but poorly disclosed documents that make otherwise patent-worthy inventions unpatentable and reduce firms' incentives to invest in innovation (Ouellette et al., 2025; Yordy, 2021).

trajectories. Some studies argue that advances in AI enable new approaches to scientific and technical research, effectively serving as a method of invention with economy-wide spillovers (Cockburn et al., 2018). Other work suggests that because recent AI progress has focused on automating, aggregate productivity gains are modest, and sustained improvements require redirecting AI innovation toward new tasks and idea generation (Acemoglu and Restrepo, 2020; Acemoglu, 2025). We find that AI adoption is associated with improved efficiency in the use of innovative inputs, as well as changes in inventor composition and more effective knowledge recombination within firms. These patterns are consistent with AI beginning to reshape elements of the innovation production function at the firm level.

Research on intangible capital shows that modern firms invest heavily in non-physical assets such as R&D, software, and organizational capabilities, and that capitalizing these expenditures makes intangibles a large and growing component of the capital stock and an important driver of growth and productivity (Corrado et al., 2009). A related literature shows that a large and growing share of the capital underlying firm valuation and asset pricing is intangible and that incorporating measures of intangible capital can materially affect risk and valuation (Crouzet et al., 2022; Eisfeldt et al., 2022; Kogan and Papanikolaou, 2019). Consistent with this view, our evidence suggests that the market values AI as innovation-enabling intangible capital.

The literature on the development and impact of AI technologies suggests that success relies on data availability, computing capacity, domain-specific expertise, and integration into core operations (Brynjolfsson and McAfee, 2017; Cockburn et al., 2018; Agrawal et al., 2018). Recent survey studies underscore that high-quality data for building machine learning models are critical to the performance of AI (Zha et al., 2025; Whang et al., 2023). Consistent with this view, we provide new evidence that focused firms realize larger innovation gains from AI adoption, as they tend to produce higher-quality, better-integrated, and more internally aligned data, accumulate deeper domain expertise, and face lower coordination and integration costs.

Finally, and importantly, our paper is also related to recent work investigating the impact of AI on aggregate and firm-level productivity. Existing studies typically extrapolate aggregate

effects either by calibrating a task-based framework using the GDP share of tasks impacted by AI and estimated average cost savings in those tasks or a historical analogy to earlier general-purpose technology waves. Under the task-based approach, Aghion and Bunel (2024) report implied increases in annual aggregate TFP growth ranging from 0.07 to 1.24 percentage points (pp), with a median estimate of 0.68 pp per year, while the preferred calibration in Acemoglu (2025) implies roughly 0.064 pp per year. Under the historical analogy, Aghion and Bunel (2024) benchmark AI against the European electricity wave and the US information and communication technology (ICT) wave, implying about 1.3 and 0.8 pp higher annual productivity growth, respectively, over the next decade. Using U.S. Census manufacturing data, Alderucci et al. (2024) document that firms with at least one AI-related patent have about 8% higher firm-level TFP. We find that AI-adopting firms experience about 8.87% higher value-added TFP on average after adoption than non-adopters, implying a contribution of approximately 0.61% to aggregate value-added TFP growth in a typical post-adoption year, after accounting for offsetting effects on non-adopters.

3 Data and Variable

3.1 Patent Data and Sample

To measure firms' innovation activities, we collect data on patents, citations, and technology classifications from the United States Patent and Trademark Office (USPTO), using the May 5, 2025 data release. Firm financial data are obtained from Compustat. We exclude firms in

financial industries (SIC code 6000 to 6999)[6] and firm-year observations with missing revenue data or less than $10 million in assets. Firms are required to have filed at least one patent during the sample period. For these firms, we retain firm-year observations with and without patent filings to mitigate concerns about selection bias.

To link patents to Compustat firms, we construct a comprehensive linking table by integrating link tables from Kogan et al. (2017), Wharton Research Data Services (WRDS), and Stoffman et al. (2022). We further supplement this table using fuzzy name matching and manual verification through SEC filings, Google searches, D&B Hoovers, PR Newswire, Wikipedia, and other sources. To enhance matching accuracy, we use additional firm-specific information, such as locations, prior M&A activities, and name change histories to identify the most likely firm match. We only retain those matches for which we have high confidence.

The final sample consists of 2,600 firms and 791,040 granted patents filed during the period 2011 to 2021. We stop the sample in 2021 to minimize truncation bias caused by the delay between patent application and grant, as patents usually take 2 to 3 years to be granted (Hall et al., 2001). All variables are winsorized at 1% and 99%. Following the innovation literature, missing values in innovation measures are set to zero to reflect a lack of output rather than missing data (Hall et al., 2001).

3.2 AI Adoption Data and Measure

We use detailed site-level IT installation records for U.S. firms from the Spiceworks Ziff Davis (SWZD) Historical Company Intelligence Database to measure firm's adoption of AI. This

[6]We exclude firms in financial industries because their business models and balance sheets differ systematically from those of non-financial firms, making it difficult to account for differences in observable characteristics. Moreover, our estimation of value-added TFP relies on a Cobb–Douglas production function, for which the conditions required to interpret the residual as technological efficiency are less likely to hold in the financial sector due to the nature of financial intermediation. In particular, measured value added reflects not only quantities of services but also but also variation in interest margins, risk premia, and asset prices, which embed pricing and risk-taking components, while balance sheet expansion and leverage can increase output without corresponding changes in productivity. In addition, key inputs such as capital are endogenously determined by regulatory constraints and risk choices, complicating the stable production mapping assumed by the model. Consequently, the residual may conflate productivity with risk-taking and financial structure, thereby limiting its economic interpretability.

database provides comprehensive information on IT products and their installation dates across firm sites. SWZD compiles data from multiple sources on firms' installations of software, hardware, networks, and telecommunications equipment. It collects information through interviews with senior IT executives and supplements these data with systematic collection methods, including client engagement, web scraping, job postings, document analysis, and purchasing customer lists from vendors. The data are updated annually and undergo extensive quality checks to ensure accuracy. According to SWZD, this database captures approximately 90% of global IT purchasing activity, offering a comprehensive view of technology adoption. The specificity and granularity of this data make it well-suited for tracking the timing of AI adoption across firms. We aggregate site-level technology installation data to the firm level and link them to historical company names from CRSP for Compustat firms using a fuzzy matching algorithm. All matches are manually reviewed to ensure accuracy. For cases where names do not match exactly, we conduct internet searches and include the observation only when we are confident in the validity of the match.

We construct a firm-level measure of AI adoption by identifying the earliest observed installation date of an AI product across the firm's sites.[7] The first installation year defines the onset of AI adoption. Because the measure is based on observed AI product installations, it provides a direct proxy for realized adoption and captures the deployment of AI technologies within the firm. In this sense, it reflects the introduction of AI capabilities into production and innovation processes. A firm is classified as an AI-adopting firm if it adopts any AI technology during the sample period and as a non-adopter otherwise.

[7]Examples of such products in our data include machine learning frameworks and platforms such as TensorFlow, scikit-learn, and H2O.ai, enterprise and cloud machine-learning platforms such as Microsoft Azure Machine Learning, Google Cloud Machine Learning Engine, Amazon Machine Learning, and the enterprise AI platform Dataiku, as well as more specialized decision-support systems such as Nara Logics and data-labelling and human-in-the-loop platforms such as CrowdFlower. These examples are illustrative rather than exhaustive. Since our sample ends in 2021, chatbot-based generative AI such as ChatGPT is not included. However, the machine learning frameworks and platforms covered in the database can be used to develop generative AI applications.

3.3 Innovation Measures

We measure innovation outcomes using patent-based innovation metrics. The filing date of a granted patent serves as the primary indicator of innovation activity, as it most closely reflects the timing of the underlying inventive efforts (Lerner and Seru, 2022).

3.3.1 Quantity and Quality

To capture the quantity of innovation outcomes, we count the number of patents filed by a firm in a given year. We proxy patent quality with the number of forward citations, which reflects a patent's technological impact (Hall et al., 2005). Given that the raw number of received citations is subject to truncation bias since citations take time to accrue, we follow Hall et al. (2001) and Lerner and Seru (2022) to adjust forward citation counts based on the patent's filing year and tech class. The metric is computed as follows:

$$Adjusted\ citation_p = \frac{Citation_p}{\sum_j^{N_{kt}} Citation_j / N_{kt}} \tag{1}$$

where $Citation_p$ is the number of citations received by patent p filed in year t, $Citation_j$ is the number of citations received by another patent j, filed in year t, within the same technology class k as patent p. N_{kt} denotes the total number of patents applied in year t within class k. This approach compares the number of citations a patent receives relative to other patents filed in the same year and technology class, thereby mitigating truncation bias.

3.3.2 Innovation Characteristics

3.3.2.1 Innovation Novelty: Originality and Distance

Originality captures the diversity of prior knowledge underpinning the invention (Acharya and Xu, 2017). Following Trajtenberg et al. (1997), we calculate originality as one minus the

Herfindahl–Hirschman Index of backward citations cited by a patent:

$$Originality_p = 1 - \sum_{k}^{n_p} B_{pk}^2 \tag{2}$$

where B_{pk} represents the ratio of backward citations made by patent p to patents in technology class k to the number of total backward citations made by patent p across all technology classes. n_p is the number of unique technology classes cited by patent p. A higher originality score indicates that a patent draws knowledge from a more diverse array of technological fields.

Distance measures the technology propinquity between a patent and the overall technology expertise of its owning firm (Akcigit et al., 2016a). Firstly, we compute the distance between technology classes, defined by the first two digits of the Cooperative Patent Classification (CPC) code. Specifically, the distance metric between technology classes X and Y is calculated based on the citations made to classes X and Y:

$$d(X,Y) \equiv 1 - \frac{\#(X \cap Y)}{\#(X \cup Y)} \tag{3}$$

where $\#(X \cap Y)$ denotes the number of patents that cite at least one patent from both class X and Y (intersection) and $\#(X \cup Y)$ denotes the number of patents that make citations to at least one patent from either class X or Y (union). If two classes are frequently cited together, they are considered technologically closer. Conversely, if they are rarely co-cited, the classes are deemed more distant.[8]

Next, we calculate the distance of a patent p to its filing firm i, proxied by the average distance between p and all other patents previously filed by firm i. The metric is computed as

[8]For robustness, we compute two versions of the distance metric for a given class pair in a given year: one based on citations in that year alone and another based on cumulative citations up to that year. The two approaches yield similar results. We report results using technology distances calculated from citations accumulated up to the given year.

follows:

$$d_t^{\iota}(p,i) \equiv \left[\frac{1}{\|\mathcal{P}_{it}\|} \sum_{p' \in \mathcal{P}_{it}} d(X_p, Y_{p'})^{\iota} \right]^{\frac{1}{\iota}} \tag{4}$$

where $\mathcal{P}_{it}$ denotes the set of all patents filed by firm i prior to the patent p in year t. The cardinality $\|\mathcal{P}_{it}\|$ represents the size of firm i's patent portfolio accumulated before filing patent p in year t. $d(X_p, Y_{p'})$ denotes the distance between a patent p in technology class X and another patent p' in technology class Y, where p' represents each other patent in firm i's portfolio. Following Akcigit et al. (2016a) and Ma et al. (2022), we set ι, the power parameter of the generalized mean operator, equal to 2/3. Here, $d_t^{\iota}(p,i)$ denotes the technology distance between patent p to its owning firm i in year t. A higher value indicates that the patent lies in a more technologically distant area relative to the firm's existing technological expertise.

3.3.2.2 Innovation Direction: Exploitative and Explorative

Exploitative patents refer to inventions that mainly build on a firm's existing knowledge base. Following Benner and Tushman (2002) and Gao et al. (2018), we define a firm's existing knowledge base for a given patent as (i) all patents filed by the firm and (ii) all citations made by the firm's patents from the beginning of the fifth calendar year prior to that patent's filing date to its filing date. We classify a patent as exploitative if at least 80% of its citations draw on this existing knowledge base (that is, either belong to the firm's own patent filings or were cited by the firm during this period). This measure captures the extent to which a firm's innovation improves along the existing technology trajectory (Benner and Tushman, 2002).

On the contrary, explorative patents refer to inventions that are mainly built outside of a firm's existing knowledge base. A patent is classified as explorative if at least 80% of its citations are outside the firm's existing knowledge base.[9] This measure gauges the extent to which a firm's innovation shifts to a new technology trajectory (Benner and Tushman, 2002).

[9] As a robustness check, we also use a 60% threshold to define exploitative and explorative patents and find similar results.

3.3.2.3 Innovation Reach: Generality and Patent Claim

Generality reflects how widely a patent's impact spreads across different technology classes (Hall et al., 2001). Following Trajtenberg et al. (1997), we calculate generality as one minus the Herfindahl–Hirschman Index of forward citations citing a patent:

$$Generality_p = 1 - \sum_{k}^{n_p} F_{pk}^2 \tag{5}$$

where F_{pk} represents the ratio of forward citations made to patent p from patents in technology class k to the number of total forward citations made to patent p across all technology classes. n_p is the number of unique technology classes citing patent p. A higher generality score indicates that a patent influences subsequent patents in a wider variety of technology domains.

Patent claims define the scope of a patentee's exclusive rights and are of significant economic and legal importance (Lerner, 1994; Lee, 2010). We follow Akcigit et al. (2016b) and Aghion et al. (2019) to count the number of claims contained in a patent application.

4 Empirical Framework

4.1 Summary Statistics

Table 1 reports innovation outcomes and firm characteristics for the full sample, AI adopters, and never-adopting firms. Out of 2,600 firms in our sample, 518 (20%) firms adopt AI technologies, while 2,082 (80%) firms never do so. Table A.2 reports the industry distribution of AI-adopting firms using the Fama-French 49 classification. The largest numbers of adopters are in computer software, electronic equipment, retail, business services, and pharmaceuticals, while the highest adoption rates are in transportation, entertainment, retail, utilities, and measuring and control equipment. Figure A.1 shows the geographic diffusion of AI adoption

across U.S. BEA Economic Areas from 2016 to 2021.[10] Adoption is initially concentrated in areas such as San Francisco, New York, Washington D.C., Chicago, and Los Angeles. Over time, these hubs deepen their adoption and additional areas, including several in Texas and the Midwest, begin to exhibit noticeable numbers of adopters. By 2021, high adoption remains clustered in major coastal and metropolitan areas, but medium levels of adoption are visible across much of the country, indicating gradual yet uneven spatial diffusion of AI.

AI-adopting firms produce more patents, and those patents, on average, receive more adjusted citations, exhibit greater originality and technology distance, are more exploitative, and contain more claims. AI-adopting firms exhibit higher labor and capital, yet lower value-added TFP on average compared with never-adopting firms.[11] It is worth noting that AI-using firms, on average, spend less on R&D than never-using firms, and the difference is both economically and statistically significant. This pattern alleviates the concern that our results are driven by firms that systematically invest more in R&D. We find that AI-adopting firms are larger, have higher *Tobin's q*, and hold less cash compared to never-adopting firms.

4.2 Empirical Analyses

4.2.1 Baseline Model

Since firms adopt AI technologies at different times, traditional two-way fixed effects models can suffer from the "bad comparisons" problem when treatment effects vary across cohorts or over time (Baker et al., 2022). To mitigate this bias, we follow Gormley and Matsa (2011, 2016) and use a stacked difference-in-differences approach to study the impacts of AI adoption on corporate innovation. We define cohorts based on the year of AI adoption, denoted as the event year. For each cohort, we construct an event window spanning five years before to five

[10] The earliest adoption year in our sample is 2016, consistent with evidence that AI adoption accelerated after 2015 (Acemoglu et al., 2022). This timing also coincides with major advances in widely used deep-learning tools and architectures, including TensorFlow's release in November 2015 and the Transformer architecture introduced in 2017.

[11] Value-added TFP is constructed as the residual from a log Cobb–Douglas value-added production function. Negative values indicate that, conditional on inputs, firms produce less value added than predicted by the estimated production function. Construction details are provided in Appendix Table A.1.

years after the adoption year. Treated firms are those that adopt AI in a given year, while the control group includes: (1) firms that never adopt AI during the sample period (never-treated), and (2) pre-adoption observations of firms that adopt AI later (not-yet-treated). We then stack all cohorts of treated and control firms to form the final stacked sample.

To empirically test how AI adoption affects innovation activities, we estimate the following model on the stacked sample, separately for each $h \in \{1, 2, 3\}$:

$$\mathbb{E}\left[Y_{i,j,c,t+h} \mid \cdot \right] = \exp\Big(\beta_h \, AI_{i,j,c,t} \;+\; \alpha_{i,c} \;+\; \psi_{t+h,c}\Big) \tag{6}$$

where i, j, c, t denote firm, industry, cohort, and year, respectively. $Y_{i,j,c,t+h}$ is the innovation measure for firm i in year t+h. $AI_{i,j,c,t}$ equals 1 if firm i in cohort c adopts AI in year t and 0 otherwise. $\alpha_{i,c}$ and $\psi_{t+h,c}$ are firm-by-cohort and year-by-cohort fixed effects. Industries are defined based on the Fama–French 49 classification. Standard errors are clustered at the firm level. Since the innovation outcomes are non-negative with many zeros, we estimate Equation (6) by Pseudo-Poisson Maximum Likelihood (PPML), following Cohn et al. (2022). Our parameters of interest are β_h, which capture differences in changes in patent quantity, quality, and innovation characteristics around AI adoption between adopting firms and the control groups at each horizon h.

4.2.2 Baseline Results

4.2.2.1 Innovation Quantity and Quality

We first examine the effect of AI adoption on innovation quantity, measured by the number of patents. Panel A of Table 2 reports positive and significant coefficients on AI, indicating that AI adopters increase post-adoption patenting relative to non-adopters. The coefficients are also economically significant. Three years after adoption, AI-adopting firms generate approximately 19.1% ($(e^{0.175} - 1) \times 100\%$) more patents than non-adopting firms.

We further examine how AI adoption influences the quality of innovation. Specifically, we

evaluate the scientific value of patents, measured by forward citations adjusted for technology class and application year. As shown in Panel B of Table 2, the coefficients on AI are all positive and significant, indicating that patents produced by AI-using firms receive more adjusted citations after AI adoption than those from non-adopting firms. The result suggests that AI-adopting firms develop patents of higher scientific value after using AI than non-adopting firms.

4.2.2.2 Innovation Characteristics

We further study whether the use of AI affects the innovation characteristics of AI-adopting firms. Table 3 Panel A reports the results for the impacts of AI on innovation novelty. In columns (1)-(3), we observe that AI-using firms experience a notable increase in the average originality of their patents following adoption, compared to non-adopting firms. This effect is both statistically and economically significant. After three years of adopting AI technology, AI-using firms experience 21.53% ($(e^{0.195} - 1) \times 100\%$) increase in originality, compared to non-using firms. Columns (4)-(6) show that firms experience a significant increase in technology distance following AI adoption. After three years of adopting AI technology, AI-using firms experience 13.31% ($(e^{0.125} - 1) \times 100\%$) increase in technology distance, compared to non-using firms. The results in Panel A of Table 3 indicate that AI-adopting firms subsequently produce patents that draw on a broader set of technology fields and are more technologically distant from their prior patent portfolios relative to non-adopting firms.

Panel B of Table 3 reports the results of how the use of AI is related to innovation direction. We present the results for exploitative and explorative patents in columns (1)-(3) and columns (4)-(6), respectively. Following AI adoption, firms generate more exploitative and explorative patents than non-adopting firms. In the third year after firms start using AI, exploitative and explorative patents increase by 16.65% ($(e^{0.154} - 1) \times 100\%$) and 21.04% ($(e^{0.191} - 1) \times 100\%$), respectively. These results suggest that the use of AI helps adopting firms not only deepen the utilisation of their existing technological base but also broaden it by exploring new

technological knowledge.

Table 3 Panel C reports the results of how AI affects innovation reach. Columns (1)-(3) show that after adopting AI technology, patents produced by AI-adopting firms exhibit significantly higher generality than patents produced by non-adopting firms. In columns (4)-(6), we observe that patents of AI-using firms have a higher number of claims after AI adoption, compared to patents of non-using firms. Three years post adoption, patents of AI-using firms have approximately 17.48% ($(e^{0.161}-1)\times 100\%$) more claims than patents of non-using firms. These findings indicate that firms adopting AI technologies produce patents with extended technological and legal scope, thereby shaping future innovation across a wider spectrum of technological fields.

4.2.2.3 Dynamic Effects

The baseline estimates indicate that AI adopters subsequently outperform non-adopters in innovation. These estimates may be biased by reverse causality, whereby firms with improving pre-adoption innovation are more likely to adopt AI, and by differential pre-trends if adopters and non-adopters follow distinct pre-adoption trajectories for reasons unrelated to AI. To address these concerns, we augment the stacked DiD model with a cohort-specific event study. For each adoption cohort, we estimate event-time coefficients (leads and lags relative to the year before adoption) to test pre-adoption parallel trends and to trace the timing and persistence of post-adoption effects on innovation.

Following Baker et al. (2022), let $t_{i,c}$ denote firm i's adoption year in cohort c and define the relative time $\ell \equiv t - t_{i,c}$. For each ℓ, the indicator $AI^{\ell}_{i,j,c,t} = \mathbf{1}\{t - t_{i,c} = \ell\}$, which equals 1 if t is ℓ years relative to adoption and 0 otherwise. We estimate dynamic effects using a cohort-specific event-study:

$$\mathbb{E}\big[\, Y_{i,j,c,t+1} \mid \cdot \,\big] = \exp\left(\sum_{\ell=-5}^{-2} \beta_\ell \, AI^{\ell}_{i,j,c,t} \;+\; \sum_{\ell=0}^{5} \beta_\ell \, AI^{\ell}_{i,j,c,t} \;+\; \alpha_{i,c} \;+\; \psi_{t+1,c} \right) \tag{7}$$

where, $Y_{i,j,c,t+1}$ is the innovation outcome for firm i in year t+1. We use $\ell = -1$ (one year before adoption) as the reference period. $\alpha_{i,c}$ and $\psi_{t+1,c}$ are firm-by-cohort and year-by-cohort fixed effects. Standard errors are clustered at the firm level. This model uses a set of relative-time indicators to capture the time periods before the treatment ("leads") and the time periods after treatment ("lags").

Figure 1 reports the coefficients from the lead-lag regressions from 3 years before the treatment to 4 years after treatment. Panel A shows that the difference in patent quantity and quality between AI-adopting and non-adopting firms increases significantly two years after AI adoption. Panels B, C, and D report differences in patent characteristics in terms of novelty, direction, and reach, respectively. In Panel B, we document that differences in patent originality and distance between AI-adopting and non-adopting firms become significant one year after adoption. Panel C suggests that AI-using firms produce notably more exploitative patents one year after AI adoption compared to non-using firms, while there are no significant changes in explorative patents. Panel D indicates that differences in patent generality become significant two years after AI adoption and the number of claims become significant one year after AI adoption. Furthermore, we find no significant pre-adoption differences in innovation activity between AI adopters and non-adopters, mitigating concerns that the results are driven by reverse causality or pre-existing trends.

Our results also show that the impact of AI adoption on corporate innovation emerges relatively quickly. This is consistent with our use of patent filing rather than grant dates to measure innovation output. Moreover, industry evidence documents that AI-enabled R&D projects often progress on the scale of months rather than years (Microsoft, 2024; Chu, 2025). Prior work also indicates that the transition from R&D to patenting is relatively rapid and that firms tend to apply for patents early in the R&D process rather than upon its completion (Hall et al. (1984); Griliches (1998); Moretti (2021); He and Qiu (2025)).

4.2.3 Addressing Endogeneity

The baseline estimates may be biased by endogeneity arising from non-random selection into AI adoption. The inclusion of firm–cohort fixed effects mitigates selection driven by time-invariant differences across firms within each cohort stack. However, endogeneity may persist if time-varying, unobserved changes, such as improvements in management quality or shifts toward a more innovation-oriented strategy, simultaneously increase firms' propensity to adopt AI and their innovation outcomes, even absent AI adoption. In this case, these unobserved determinants jointly affect both adoption and innovation, confounding the estimated effect of AI adoption on innovation. To alleviate this concern, we implement a control function regression approach using an instrumental variable to isolate exogenous variation in AI adoption.

A control function regression approach is a useful method to address endogeneity, especially in non-linear models (Wooldridge, 2015). In the first stage, it uses one or more instrumental variable to isolate exogenous variation in the endogenous variable, providing separate variation in the generalized residuals. In the second stage, these generalized residuals are included as a control function, so that the endogenous variable is conditionally exogenous. Compared with two-stage least squares (2SLS), the control function approach provides a more parsimonious way to handle non-linear models with endogenous explanatory variables (Wooldridge, 2015) and therefore is more suitable for our setting.

We follow Babina et al. (2024) to use *AI exposure* as an instrumental variable, which calculates a firm's exposure to AI-strong universities via the firm-university STEM worker hiring networks as of 2010. As noted by Fujii and Managi (2018), universities are key AI technology inventors in both the U.S. and China. Babina et al. (2023) find that firms that have higher initial shares of highly-educated and STEM employees are more likely to invest in AI. Moreover, Babina et al. (2024) document that the scarcity of well-trained AI talent is a major constraint on firms' ability to adopt AI technology. Therefore, our instrument satisfies the relevance condition: firms that maintain well-developed hiring pipelines with AI-

strong universities in 2010 are more likely to attract and access high-quality AI talent, and consequently adopt AI technologies. We thus expect a positive relationship between a firm's ex-ante exposure to AI-strong universities and its likelihood of adopting AI technology.

The exclusion restriction may be challenged if firms' exposure to AI-strong universities in 2010 influences firms' innovative activities in ways other than AI adoption. However, we argue that this is less of a concern in our setting for several reasons. First, *AI exposure* is an ex-ante measure built on firms' hiring connections to these universities as of 2010 when the commercial use of AI had not yet become widespread (Babina et al., 2024). In addition, Babina et al. (2024) document that from 2005 to 2010, the proportion of graduates hired from AI-strong universities remains stable among firms connected to those institutions. This helps alleviate the endogenous concern that it is the anticipation of future needs for AI talent that encourages the hiring of AI-trained graduates in 2010. Second, while the large-scale commercial deployment of AI has only surged in recent years, academic research on artificial intelligence can be traced back to the 20th century. Given that AI-strong universities are identified using academic publications prior to 2010, it is relatively unlikely that it is the demand of AI-adopting firms that reversely promotes AI research in universities prior to 2010. Third, access to AI talent and access to high-quality human capital do not necessarily overlap: of the 44 AI-strong universities, 11 (25) are in the U.S. News & World Report top 20 (top 50). To further address the possibility that access to top-tier human capital confounds the results, we include a control variable, *Top 10 exposure*, which measures the firm's exposure to top 10 universities via the firm-university STEM worker hiring networks as of 2010 (Babina et al., 2024).

To implement the control-function approach, we first estimate a probit model on the stacked panel:

$$\begin{aligned} AI^{*}_{i,j,c,t} &= \beta\, AI\ exposure_{i,j,c} \;+\; \theta\, Top\,10\, exposure_{i,j,c} \;+\; \gamma_{j,c} \;+\; \tau_{t,c} \;+\; u_{i,j,c,t}, \\ AI_{i,j,c,t} &= \mathbf{1}\{AI^{*}_{i,j,c,t} > 0\}, \qquad \Pr\big(AI_{i,j,c,t} = 1 \mid \cdot\big) = \Phi\big(z'_{i,j,c,t}\boldsymbol{\delta}\big) \end{aligned} \tag{8}$$

where $AI\ exposure_{i,j,c}$ is exposure to AI-strong universities via firm–university STEM hiring

networks as of 2010 (Babina et al., 2024); $Top\,10\,exposure_{i,j,c}$ is exposure to the top-10 universities (2010 U.S. News & World Report) via the same networks. $\gamma_{j,c}$ and $\tau_{t,c}$ are industry-by-cohort and year-by-cohort dummies. We do not include firm fixed effects because the instrumental variable varies cross-sectionally and would be absorbed by them. Let $z_{i,j,c,t}$ stack the regressors and dummies, and let $\boldsymbol{\delta}$ denote the corresponding coefficient vector.

We compute the generalized residuals from the probit model as:

$$\widehat{Residual}_{i,j,c,t} = AI_{i,j,c,t}\,\frac{\phi\left(z'_{i,j,c,t}\hat{\boldsymbol{\delta}}\right)}{\Phi\left(z'_{i,j,c,t}\hat{\boldsymbol{\delta}}\right)} \;-\; \left(1 - AI_{i,j,c,t}\right)\frac{\phi\left(z'_{i,j,c,t}\hat{\boldsymbol{\delta}}\right)}{1-\Phi\left(z'_{i,j,c,t}\hat{\boldsymbol{\delta}}\right)} \tag{9}$$

where $\phi(\cdot)$ and $\Phi(\cdot)$ denote the standard normal PDF and CDF, and $\hat{\boldsymbol{\delta}}$ is the first-stage estimate.

In the second stage, we estimate the PPML model including the generalized residuals:

$$\mathbb{E}[Y_{i,j,c,t+h} \mid \cdot] = \exp\left(\beta\, AI_{i,j,c,t} + \theta\, Top10\, exposure_{i,j,c} + \kappa\, \widehat{Residual}_{i,j,c,t} + \gamma_{j,c} + \psi_{t+h,c}\right) \tag{10}$$

where $Y_{i,j,c,t+h}$ is the innovation outcome for firm i in year $t+h$. $\gamma_{j,c}$ and $\psi_{t+h,c}$ are industry-by-cohort and year-by-cohort fixed effects, respectively. Standard errors are computed using a firm-level cluster bootstrap with 1,000 replications.

Table 4 reports the control function regression results for innovation quantity, quality, characteristics. For brevity, we report the estimates in year $t+3$ since the results for other years are similar.[12] Column (1) presents the results of the first-stage regression. The estimated coefficient on *AI exposure* is positive and significant at the 1% level, indicating that exposure to AI-strong universities is positively associated with firms' likelihood of adopting AI technology. The large and highly significant Wald χ^2 statistic ($\chi^2 = 541.82$, $p < 0.001$) indicates that the instrument strongly predicts the endogenous variable in the first stage, supporting the relevance condition. Columns (2)-(9) present the second-stage regression results. We find

[12]The probit model in the first stage is estimated by maximum likelihood, with convergence attained at the maximum of the log-likelihood function. Estimation may fail when the binary outcome is perfectly or nearly perfectly predicted by the covariates, or when the within-firm sample size is insufficient for reliable identification. Consequently, firms with very few observations are excluded from the estimation sample, which reduces the number of observations available for inference.

that the positive effects of AI adoption persist across various innovation outcomes, including *Patent count*, *Adjusted citations*, *Originality*, *Distance*, *Exploitative*, *Generality*, and *Claim*. The negative coefficients on the generalized residuals indicate negative selection, i.e., factors increasing the propensity to adopt AI are associated with lower innovation outcomes, biasing baseline estimates downward. Accounting for this selection leads to larger estimates relative to the baseline. Overall, these results suggest that differences in innovation quantity, quality, and characteristics between AI-adopting and non-adopting firms remain robust after accounting for the endogeneity of AI adoption.

5 Channels

The results thus far suggest that AI adoption significantly affects adopters' innovation outcomes. In this section, we examine the potential channels through which AI technology influences innovation.

5.1 Efficiency Channel

5.1.1 Inventor Productivity

We begin by examining an efficiency channel through which AI can raise corporate innovation by increasing patent output generated from a given set of inventive inputs. We first focus on inventor productivity and discuss three related mechanisms through which AI can increase patent output per inventor. First, AI can automate routine, time-intensive tasks and process technical and alternative data at scale. Setchi et al. (2021) show that AI can reduce the time and cost of prior-art searches by efficiently screening large volumes of patent documents. Second, AI can lower the costs of accessing and internalizing relevant knowledge within the firm by codifying and disseminating effective practices, which may particularly benefit less-experienced inventors. Brynjolfsson et al. (2025) find that a generative AI-based conversational assistant increases the productivity of novice and lower-skilled customer-support agents

by summarizing top performers' practices and diffusing them across workers. Third, AI can accelerate knowledge search and recombination by helping inventors retrieve, synthesize, and connect disparate information, thereby increasing the rate at which they generate patentable ideas. For instance, AI can help identify emerging technological trends (Mühlroth and Grottke, 2020) and support hypothesis generation (Ludwig and Mullainathan, 2024).

To explore this channel empirically, we define *Inventor productivity* as the average number of granted patents filed by inventors of firm i in year t. We first calculate the sum of all patents filed by an inventor at the inventor-year-level. For patents with multiple inventors, we assign each inventor an equal fraction of the patent following Moretti (2021). We then aggregate the inventor–year patent counts to the firm–year level.[13] Using Equation (6), we compare inventor productivity in AI-adopting firms before and after AI adoption with that of non-adopting firms. Columns (1)-(3) in Panel A of Table 5 reports the results. We find that inventor productivity in AI-adopting firms increases significantly following AI adoption, relative to that in non-AI firms. Compared with non-using firms, AI-using firms experience a 17.59% ($(e^{0.162} - 1) \times 100\%$) increase in inventor productivity three years after adoption. Our results suggest that AI promotes corporate innovation through increased inventor productivity.

5.1.2 Innovation Efficiency

AI may also affect corporate innovation through innovation efficiency, the effectiveness with which firms convert R&D spending into patents. AI can enhance innovation efficiency by improving the screening of research opportunities and increasing the expected discoveries per experiment. For example, using AI to predict the bioactivity of candidate molecules substantially improves the productivity of early-stage drug screening (Cockburn et al., 2018). Merchant et al. (2023) argue that AI can accelerate material discovery by allowing out-of-distribution exploration in previously unseen chemical spaces and by improving predictive performance on downstream tasks. Bennett et al. (2025) claim that AI models can substantially

[13] As a robustness check, we compute *Inventor productivity* as the ratio of the number of patents filed by firm i in year t to the number of inventors who file patents in firm i in year t. Our results remain similar.

improve antibody discovery and development by enabling faster and more cost-effective identification of promising candidates compared with approaches such as animal immunization and random library screening.[14]

By reducing information frictions in project selection and enabling better allocation of R&D resources, AI can help firms direct investment toward higher-potential projects and generate more patent output per dollar spent. Yoo et al. (2023) show that using machine learning models to predict the performance of R&D projects improves the efficiency of funding decisions. Efficiency gains from better project selection and fewer avoidable failure modes can lead to higher patent quality. Finally, AI may enable more efficient R&D experimentation. By automating screening, candidate generation, and experiment design, AI lowers the marginal cost of search and expands the feasible search space for experimentation, raising innovation efficiency (Cockburn et al., 2018).

To investigate this channel, we follow Gao and Chou (2015) to define *Innovation efficiency* as the firm-year ratio of the number of patents to R&D expenditure. We use Equation (6) to compare the innovation efficiency in AI-using firms before and after using AI technology with that of non-using firms. Columns (4)-(6) in Panel A of Table 5 shows that AI-using firms experience a significant increase in innovation efficiency relative to non-adopting firms. Compared with non-adopting firms, innovation efficiency in AI-adopting firms rises by 18.07% ($(e^{0.166}-1)\times 100\%$) three years after adoption. Our results indicate that AI promotes corporate innovation through a higher rate of conversion from inventive effort to patentable outcomes.

5.2 Inventor Composition Channel

AI adoption may affect not only inventor productivity, but also the firm's optimal choice of inventive labor inputs along both the extensive margin (scale) and the intensive margin

[14] McKinsey & Company (2023) suggests that AI models can accelerate research operations by both speeding up each simulation and enabling multiple tests simultaneously. Based on a survey of biopharma R&D executives, Accenture (2023) finds that AI-led discovery strategies can help firms shorten discovery cycle times by up to two-thirds.

(composition of skills and experience). An important insight from the technology and labor literature is that information technology complements workplace reorganization and skilled labor, so adopting new technologies is often accompanied by changes in workforce composition (Bresnahan et al., 2002). The direction of these adjustments depends on whether a technology primarily replaces human task performance (substitution) or raises the productivity of humans performing those tasks (complementarity).

This perspective implies that AI can reshape the innovation process by reallocating tasks within R&D and, in turn, shifting demand across inventive skills. In the task-based framework of technological change, computerization substitutes for routine, codifiable tasks while complementing non-routine problem solving and complex communication (Autor et al., 2003). To the extent that AI expands the set of activities that can be automated or reliably augmented, firms have incentives to reorganize inventive work by shifting routine components from labor to capital and reallocating human effort toward tasks where researchers retain comparative advantage. For example, if AI automates or substantially reduces the time cost of routine R&D inputs, such as documentation, code templating, prior-art and literature search, and basic experimentation, then inventors can be redeployed toward higher-value activities, including framing research questions, generating and testing hypotheses, debugging complex designs, and coordinating across technical domains.

These adjustments could lead to changes in the composition of the inventor workforce. First, by lowering fixed costs of producing and codifying inventions, AI may increase the size of the inventor pool and encourage entry of first-time inventors. Second, if AI raises the productivity of non-routine analytic and evaluative tasks, the marginal product of skilled inventive labor increases, leading firms to tilt team composition toward inventors who can effectively deploy, supervise, and validate AI-augmented workflows. Third, if AI automates routine components of invention while complementing evaluation and integration, the internal division of labor in R&D may shift, changing field specialization across inventors and teams.

To explore this channel, we construct four variables to measure inventor composition. We

define *Inventor pool* as the total number of unique inventors who file at least one patent with firm i in year t. For firm i in year t, *First-time inventor* is the average (across the firm's patents filed in year t) of the patent-level share of inventors in each patent's inventor team who have no prior USPTO patent filings before year t. *AI-experienced inventor* for firm i in year t is measured as the average (across the firm's patents filed in year t) of the patent-level share of inventors in each patent's inventor team who are AI-experienced. An inventor is classified as AI-experienced if she has filed at least one patent that is predicted by Giczy et al. (2022) to contain AI technology up to year *t-1*. *Specialization* for firm i in year t is computed as the average (across the firm's patents filed in year t) of the patent-level average Specialization scores of inventors in each patent's inventor team. An inventor's Specialization score in year t is the Herfindahl–Hirschman Index computed from the distribution of her granted patents across CPC technology classes at the three-digit level, using granted patents applied for up to year t−1 (Li and Wang, 2023). A higher *Specialization* score indicates that the inventor is more specialized in a narrower set of technology classes.

Using these measures as dependent variables, we re-estimate Equation (6) to compare changes in inventor composition in AI-adopting firms before and after adoption relative to non-adopting firms. The estimations include year–cohort fixed effects, which control for aggregate trends in inventors and common calendar-year shocks, such as overall trends and macroeconomic conditions, within each cohort-specific event-time window, as well as firm–cohort fixed effects, which control for unobserved, time-invariant firm characteristics, such as persistent management quality, corporate culture, and technology level.

Panel B of Table 5 shows that AI-using firms experience significant increases in *Inventor pool*, *First-time inventor*, *AI-experienced inventor*, and *Specialization* compared with non-using firms. These results suggest that AI influences corporate innovation through changes in inventor composition. Specifically, in comparison with non-adopters, adopters employ more inventors after AI adoption, including inventors with no prior patenting experience as well as those with AI expertise. Moreover, AI adoption is associated with a relative

increase in inventors' technological specialization among AI-adopting firms compared with non-adopting firms, indicating that AI adopters tilt innovation toward exploiting the firm's existing technological strengths.

5.3 Knowledge Recombination Channel

Another mechanism emphasized in the innovation literature is that invention is often recombinant: new patents frequently arise from reconfiguring existing ideas or technological components (Weitzman, 1998; Fleming, 2001). In this view, the production of patents depends on access to relevant prior knowledge and the ability to search over and evaluate novel combinations under technological uncertainty (Fleming, 2001). As the knowledge frontier expands, inventive activity increasingly requires coordination among specialized contributors, contributing to the long-run shift toward team-based patenting (Jones, 2009; Wuchty et al., 2007).

If AI reduces the cost of searching, evaluating, and integrating interdependent technological components, it should facilitate recombinant invention by expanding the set of knowledge combinations that can be feasibly explored and validated. As the number of relevant components grows, the combinatorial search space expands, increasing the marginal cost of generating, screening, and validating candidate solutions. AI can relax this constraint by accelerating recombination search, including the generation and prioritization of candidate combinations, and improving screening and validation under uncertainty.[15]

This recombination channel has implications for inventive labor inputs. By expanding the scope of feasible combinations, AI can increase the marginal value of assembling complementary expertise within a project by enabling teams to cover more distinct knowledge domains. At the same time, by lowering cross-domain coordination and knowledge-transfer frictions, AI

[15]For example, Exscientia reports using an AI-enabled discovery platform within a design–make–test–learn workflow in the discovery of clinical candidate EXS21546. A related patent disclosure includes bioactivity data for 46 exemplified molecules, which Exscientia reports as a subset of 163 molecules synthesized and assayed during discovery, illustrating how AI can help generate and prioritize candidate recombinations for iterative experimental validation (Wills, 2022).

can raise the returns to maintaining broader technical breadth within teams. Taken together, these forces predict changes in team size and patent knowledge inputs following AI adoption.

We measure knowledge recombination from four perspectives. We define *Team size* as the firm-year average number of inventors per patent, and *Unique inventor team* as the number of distinct inventor teams, defined by unique inventor combinations, that file at least one patent with firm i in year t. For firm i in year t, *Core skill breadth* is defined as the average (across the firm's patents filed in year t) of the patent-level number of distinct core skills possessed by inventors in each patent's inventor team. Following Li and Wang (2023), we define an inventor's core skill in year t as the first three digits of CPC technology class in which she has the largest number of granted patents, based on her patent applications filed up to year *t-1*. *New cite ratio* for firm i in year t is calculated as the is the average of the patent-level *New cite ratio* across the firm's patents filed in year t. *New cite ratio* for patent p filed by firm i in year t is defined as the share of backward citations made to new knowledge, that is, knowledge outside the inventor team's existing knowledge, relative to the total number of backward citations to U.S. patents, foreign patents, and other references (Fitzgerald and Liu, 2020). The inventor team's new knowledge consists of citations to U.S. patents, foreign patents, or other references that have neither been previously cited nor produced by any member of the inventor team prior to the filing date of patent p. The inventor team's existing knowledge comprises all U.S. patents, foreign patents, and other references that have been previously cited or produced by at least one team member before the filing date of patent p.

Using these measures as dependent variables, we re-estimate Equation (6) to compare changes in knowledge recombination in AI-adopting firms before and after AI adoption relative to non-adopting firms. Panel C of Table 5 indicates that, relative to non-adopting firms, AI-adopting firms show significant increases in *Team size*, *Unique inventor team*, *Core skill breadth*, and *New cite ratio*. These results indicate that AI affects corporate innovation through the channel of knowledge recombination. Following AI adoption, inventors in AI-using firms form larger and a greater number of unique inventor teams and cite knowledge that is new to

all team members, compared with inventors in non-using firms. In addition, inventor teams in AI-adopting firms exhibit broader core technical skill coverage after adoption relative to non-adopting firms, consistent with greater cross-domain capability and higher capacity to integrate knowledge inputs within the team.

6 Cross-Sectional Heterogeneity

6.1 Business Focus

A growing literature emphasizes that AI and other data-intensive technologies are particularly valuable in data-rich and complex information environments. Theoretical models that view AI as an improvement in prediction show that the gains from better prediction are larger when decisions are made in complex environments with many potential states and abundant data (Agrawal et al., 2018). Mihet and Philippon (2019) argue that big data and AI technologies can be viewed as intangible assets, search and matching technologies, and forecasting technologies, in ways that strengthen the advantages of data-rich firms. Consistent with this view, Begenau et al. (2018) show that big-data technologies disproportionately benefit firms with more economic activity and longer histories, because they generate more data that can be processed and exploited. Accordingly, we conjecture that AI adoption benefits diversified firms operating in richer data and more complex information environments.

To test this hypothesis, we examine whether the effect of AI on innovation differs between focused and diversified firms. Following Comment and Jarrell (1995), we use the number of business segments to measure a firm's business focus at the firm-year level. Specifically, we construct an indicator variable called *Focused firm*, which equals 1 if firm i has only one business segment and no additional segments in 2015, and 0 otherwise. The measure is constructed using 2015 values, one year before the earliest adoption date in the sample, to preclude contamination from changes in business focus induced by AI. We include an interaction term between *AI* and *Focused firm* in Equation (6). The estimated coefficients on

this interaction term capture the differential impacts of AI on corporate innovation for firms with different business focus. To alleviate the concern that time-invariant firm characteristics and/or time trends may differ systematically between firms with different business focus, we include firm-cohort-focused firm and year-cohort-focused firm fixed effects.

Table 6 Panel A reports the results on innovation quantity, quality, and characteristics in the third year following AI adoption. Column (1) presents estimates for innovation quantity, column (2) for innovation quality, and columns (3) through (6) for various innovation characteristics measures. The positive and statistically significant coefficients on *Originality*, *Distance*, *Exploitative*, and *Claim* indicate that AI adoption has a particularly pronounced effect on these aspects of innovation among firms with a focused business scope. These results are inconsistent with the hypothesis that AI disproportionately benefits diversified firms with richer and more complex data by enabling them to process and exploit large volumes of complex information.

Several factors may help explain why firms with a more focused business scope are better positioned to utilize AI to generate innovation than diversified firms. First, focused firms are more likely to generate higher-quality and more coherent internal data. Because data quality is central to AI performance, such high-quality and internally consistent data enhance AI effectiveness by reducing noise and inconsistencies in model training and fine-tuning (Zha et al., 2025; Whang et al., 2023; Brynjolfsson and McAfee, 2017). Second, the productivity of AI depends on complementary assets such as domain expertise (Brynjolfsson et al., 2019; Azoulay et al., 2025). Business focus deepens domain expertise and tacit knowledge, improving problem formulation, labeling, and feature engineering, thereby increasing the productivity of AI in invention. Third, AI adoption requires complementary fixed investments in data infrastructure. With fewer competing units, focused firms face less internal bargaining and capital-allocation distortion, enabling more targeted and scalable deployment of these complements (Rajan et al., 2000). Fourth, effective AI implementation requires coordination across engineering, product, legal, and compliance. Focused firms face fewer coordination

frictions and lower integration costs, allowing faster and more complete embedding of AI in the invention process. Overall, these mechanisms suggest that focused firms can capture greater innovative returns from AI by combining higher-quality data, more targeted complementary investments, and smoother organizational integration.

6.2 Firm Size

A natural concern is that the estimated effect of AI adoption on corporate innovation is confounded by firm size. Larger firms systematically differ from smaller ones, both in their propensity to adopt AI and in the benefits they can extract from it, for reasons unrelated to AI itself. In particular, large firms have greater financial resources and richer data assets, which can make them more willing and able to adopt AI technologies, and prior work shows that such firms benefit more from AI adoption (Babina et al., 2024; Lu et al., 2024). At the same time, large firms may inherently generate more innovation regardless of AI, owing to structural advantages such as economies of scale, greater market influence, and broader patent portfolios. If this is the case, our baseline estimates could simply capture the behavior of large firms.

To examine whether our results are driven by firm size, we augment Equation (6) with an interaction term between *AI* and an indicator for *Big firm*, which equals 1 if the total assets of a firm in 2015 are at least as large as the industry average value in 2015, and 0 otherwise. Similarly, we include firm-cohort-big firm and year-cohort-big firm fixed effects to account for time-invariant firm characteristics and time-varying trends that differ across firms of different sizes. Panel B of Table 6 reports the coefficients on this interaction, which capture the differential impact of AI on corporate innovation for large versus smaller firms. The interaction term is statistically insignificant for all innovation outcomes except *Adjusted citation*. These results indicate that AI's effects on innovation quantity and innovation characteristics do not differ systematically across firm sizes. For innovation quality, as captured by *Adjusted citation*, the interaction coefficient is only weakly significant, indicating at most weak evidence of modestly stronger gains among larger adopters. Overall, these results suggest that our main

findings are not primarily driven by big firms.

6.3 Financial Constraint

Prior research shows that financial constraints shape firms' investment decisions (Almeida and Campello, 2007; Li, 2011). AI requires large complementary investments in software, data infrastructure, organizational change, and human capital, which entail substantial up-front and recurring costs (Brynjolfsson et al., 2021). Financial constraints can prevent firms from making the substantial investments needed for effective AI adoption. Such constraints may also depress innovation by limiting R&D spending and hindering talent acquisition. As a result, observed differences between AI adopters and non-adopters may reflect both technology choices and underlying financial capacity.

To examine whether financing frictions explain weaker innovation among non-adopters compared with AI adopters, we introduce an interaction term between *AI* and *Unconstrained firm* in Equation (6). Using the financial constraint measure developed by Hoberg and Maksimovic (2015), we define *Unconstrained firm* as an indicator variable equal to 1 if a firm's risk of delaying investments due to liquidity issues is no greater than the 2015 industry average, and 0 otherwise. We control for firm-cohort-unconstrained firm and year-cohort-unconstrained firm fixed effects to absorb differences related to firms' financing constraints across cohorts and over time. In Table 6 Panel C, we present estimates for the interaction term that captures the differential impact of AI on corporate innovation for firms facing different financial constraints. Our results show that the coefficients on *AI*×*Unconstrained firm* are insignificant across all innovation measures. These results are inconsistent with financial constraints being the driver of weaker innovation among non-adopters compared with AI adopters.

6.4 Technological Capabilities

Firms differ in their technological capabilities. In particular, high-tech firms with more developed R&D infrastructure and inventor expertise, might be more likely to adopt AI

and inherently more innovative. Consistent with this, Webb et al. (2018) document a rapid increase in patenting among U.S. technology firms in recent years, especially in software, cloud computing, and artificial intelligence. This raises the concern that our results could reflect preexisting differences rather than the effect of AI adoption itself.

To alleviate the concern that AI-adopting and non-adopting firms differ systematically in their underlying technological capabilities, we augment Equation (6) with an interaction term between *AI* and *High-tech firm* to examine whether the effect of AI adoption on innovation differs between high-tech firms and other firms. *High-tech firm* is an indicator variable that equals 1 if firm i operates in high-tech sectors in 2015, and 0 otherwise. We classify high-tech sectors as 2-digit NAICS sectors 51 ("Information") and 54 ("Professional, Scientific, and Technical Services") following Babina et al. (2024). We further include firm-cohort-high-tech firm and year-cohort-high-tech firm fixed effects to mitigate concerns that firms with different technological capabilities may systematically differ in their time-invariant characteristics and/or exhibit distinct time-varying trends. Panel D of Table 6 reports that the coefficients on *AI*×*High-tech firm* are mostly insignificant, indicating that we do not find evidence that high-tech firms experience a stronger effect of AI adoption on innovation compared to firms in non-high-tech industries. These results ease the concern that our results are driven by pre-existing differences in technological capabilities between AI-adopting and non-adopting firms rather than by AI adoption. To further assess whether the effects of AI on innovation are driven by technology-intensive industries, we exclude firms operating in tech sectors in 2015 from our baseline sample and re-estimate Equation (6). Table A.3 shows that the impacts of AI on innovation remain significant after excluding high-tech firms. Our results are consistent with Babina et al. (2024), which document that the benefits of AI adoption extend across a

broad range of economic sectors, not only those traditionally classified as high-tech.[16]

7 Productivity and Economic Outcomes

The evidence so far indicates that AI adoption promotes innovation by raising efficiency, altering the composition of inventors, and enabling richer knowledge recombination. We now investigate whether these innovation effects are accompanied by changes in firms' investment policies, operating costs, market value, and productivity. This analysis helps distinguish whether AI primarily operates as innovation-enabling intangible capital with productivity implications or instead functions like physical capital investment or an immediate cost-cutting technology.

7.1 Intangible versus Tangible Investment

We first examine whether AI adoption is associated primarily with greater intangible investment, measured by R&D, or with higher physical capital investment, measured by capital expenditures. If AI operates mainly by enhancing firms' inventive capabilities and their stock of knowledge capital, we expect to observe stronger adjustments along the R&D margin than in physical investment. If instead AI operates primarily as a physical capital–embodied technology that induces higher physical capital investment, we should also observe a response in capital expenditures.

To examine how AI adoption affects the investment policies of AI-using firms relative to non-using firms, we re-estimate Equation (6) using *R&D* and *Capital expenditures* as alterna-

[16] We also examine whether product market competition shapes the impact of AI on innovation. Table A.4 shows that competitive pressure amplifies the effect of AI adoption on *Originality* and *Exploitative*, but does not significantly affect other innovation outcomes. These results suggest that under intense competition, firms use AI to generate more novel inventions and to strengthen existing technological capabilities. In highly competitive markets, firms face pressure to differentiate and respond quickly (Arts et al., 2025). Such pressure can shift innovation toward exploitative activities that rely on established capabilities, entail lower uncertainty, and yield more immediate returns. Jansen et al. (2006) show that exploitative innovation is more closely associated with short-term financial performance in competitive environments, making it a strategic priority when market pressure is high.

tive dependent variables. Table 7 Panel A shows that AI adopters increase investment in R&D following adoption, whereas their capital expenditures do not rise relative to non-adopters. This pattern is consistent with AI operating primarily as innovation-enabling intangible capital: adopters allocate additional resources to R&D, the main input into innovation, rather than expanding physical capital. Together with the patent evidence and the efficiency, inventor-composition, and knowledge recombination mechanisms documented above, these results suggest that AI raises the level of innovation investment and enhances its effectiveness, without a corresponding expansion in tangible capital.

7.2 Operating Costs

We next examine whether AI adoption operates as a cost-cutting technology, as would be the case if firms deploy AI primarily to automate production or lower operating costs. Under this hypothesis, we would expect to observe declines in operating costs. By contrast, if AI operates mainly as innovation-enabling intangible capital, operating costs would remain largely unchanged in the short run, with resources instead reallocated toward R&D.

To examine these predictions, we proxy operating costs using three variables: *Production cost*, *Non-production cost*, and *Total operating cost*. *Production cost* is defined as cost of goods sold scaled by sales. *Non-production cost* is based on main SG&A, defined as selling, general and administrative expenses minus R&D expenses and advertising expenses, scaled by sales.[17] *Total operating cost* is measured as total operating expenses scaled by total sales.[18] Using these measures as dependent variables, we re-estimate Equation (6) to examine how AI adoption affects the operating costs of adopting firms relative to non-adopting firms.

Panel B Table 7 shows that the estimated coefficients on AI are statistically insignificant for *Production cost*, *Non-production cost*, and *Total operating cost*, indicating that AI-using

[17] According to Compustat, selling, general and administrative expenses (xsga) include R&D expenses (xrd) and advertising expenses (xad). Following Banker et al. (2019) and Ang et al. (2022), we subtract xad and xrd from xsga to measure non-production, non-R&D, non-advertising operating costs.

[18] As a robustness check, we follow Enache and Srivastava (2018) and scale production costs (cogs), non-production costs (xsga-xrd-xad), and total operating costs (xopr) by total assets. The results are similar.

firms do not experience significant changes in operating-cost ratios following adoption relative to non-adopting firms. These results do not support the view that AI lowers operating costs. Combined with the R&D and patent evidence, these findings suggest that AI operates mainly through innovation-related intangible investment and the reorganization of inventive activity, rather than through changes in operating costs.

7.3 Market Value

This innovation-based, intangible-capital view of AI adoption has direct implications for market value. If investors view AI adoption as strengthening firms' inventive capabilities and accelerating the accumulation of productivity-enhancing intangible capital, they should assign higher valuations to adopters relative to non-adopters, reflecting higher expected future cash flows and growth opportunities. By contrast, if investors perceive AI primarily as affecting current operating costs and routine capital expenditures without materially altering the trajectory of innovation or productivity, market value should change little around adoption.

To study the valuation impact of AI adoption, we use *Tobin's q* and *Total q* as proxies for market value. *Tobin's q* is defined as the market value of assets divided by the book value of assets. *Total q*, developed by Peters and Taylor (2017), is the firm's market value divided by the sum of its physical and intangible capital stocks. By explicitly incorporating an estimate of intangible capital, *Total q* refines the standard *Tobin's q* measure and better captures firms' investment opportunities, particularly for innovative firms.

Using the same stacked DiD framework, we re-estimate Equation (6) with *Tobin's q*, and *Total q* as alternative dependent variables to compare changes in market value for AI adopters before and after adoption relative to non-adopters. Panel C in Table 7 shows that the estimated coefficients on AI are positive and significant. Firms adopting AI exhibit significant increases in both *Tobin's q* and *Total q* compared with non-adopting firms, implying that the market assigns higher valuations to adopters following AI adoption. The valuation response is consistent with capital markets capitalizing expected future cash flows and growth

opportunities associated with AI-enabled innovation.

7.4 Firm-Level Productivity

Furthermore, we investigate how AI adoption affects firm-level productivity. In innovation and intangible-capital frameworks, AI can be viewed as a technology that raises the productivity of innovative effort rather than as another unit of physical capital. By improving prediction, search, and experimentation in R&D, AI can enhance innovation which is embodied in new products and processes and thereby raises firm productivity (Cockburn et al., 2018). These effects operate through the accumulation and deployment of knowledge and related organizational assets, and therefore require complementary intangible investment such as R&D and organizational capital (Corrado et al., 2009; Peters and Taylor, 2017). If AI mainly functions as innovation-enabling intangible capital in this way, firm productivity should improve after adoption.

To examine the impact of AI adoption on firm productivity, we measure firm-level productivity using value-added total factor productivity (*Value-added TFP*). We focus on value-added TFP because it more closely captures underlying technical efficiency. Revenue-based TFP, by contrast, incorporates variation in output prices, demand conditions, and markups, thereby conflating production efficiency with firm-specific market power and demand shocks (Foster et al., 2008; De Loecker, 2011). We obtain *Value-added TFP* as the residual from a log value-added Cobb–Douglas production function estimated with the generalized method of moments (GMM) approach of Wooldridge (2009).[19] The measurement of value added, capital, and labor inputs in the Cobb–Douglas production function follows Bennett et al. (2020), and we use intermediate inputs, measured as the log of real materials, as a proxy

[19] We adopt the generalized method of moments (GMM) framework of Wooldridge (2009), which implements the proxy-variable approaches of Levinsohn and Petrin (2003) as a system of moment conditions estimated jointly in a single step. This unified GMM formulation preserves the identification of the original Levinsohn–Petrin estimators and accommodates heteroskedasticity and serial correlation through a robust weighting matrix when estimating the production-function coefficients used to compute firm-level TFP. Our results are robust to alternative TFP estimates using the Levinsohn–Petrin method with the Ackerberg, Caves and Frazer (2015) correction.

for unobserved productivity shocks as in Levinsohn and Petrin (2003). Real materials are constructed by deflating materials using the GDP implicit price deflator. Prior to estimation, we remove common industry-time variation by demeaning value added, intermediate inputs, capital, and labor using Fama–French 49 industry-year fixed effects, following İmrohoroğlu and Tüzel (2014). Firm-level variables, even after deflation, may still reflect differences in prices or markups rather than pure quantities. The inclusion of industry-year fixed effects absorbs components that are common within an industry-year, including shared movements in prices and demand. As a result, the estimation of firm-level TFP relies on within-industry-year variation. We then estimate *Value-added TFP* using these residualized variables. Additional details are provided in Appendix Table A.1. As a widely used measure of productivity, value-added TFP can be interpreted as capturing differences in technological and organizational efficiency in transforming capital and labor into value added in the production process (Bennett et al., 2020).

We re-estimate Equation (6) with *Value-added TFP* as the dependent variable to compare changes in productivity for AI adopters before and after adoption relative to non-adopters. Columns (1)–(3) of Panel A in Table 8 show that the estimated coefficients on *AI* are positive and significant, indicating that, relative to non-adopters, AI adopters experience higher productivity and use capital and labor more efficiently in the production process. The impact of AI adoption on firm-level productivity is economically meaningful. The coefficient on *AI* in value-added TFP is 0.082 in year $t + 1$, implying that adopters experience about 8.55% ($e^{0.082} - 1 \approx 0.0855$) higher value-added TFP one year after adoption relative to non-adopters.

To shed light on the value-added TFP response, we examine post-adoption effects on production inputs, focusing on capital and labor. These dynamics indicate whether AI adoption is accompanied by tangible-capital deepening or employment expansion, helping separate immediate scale adjustment from mechanisms that rely more heavily on intangible and organizational complements. This distinction is particularly relevant for the innovation channel, in which AI raises productivity by improving research efficiency and reshaping

the innovation process. Under this channel, the key complements are largely intangible, so tangible capital is expected to respond weakly in the near term even as productivity rises (Cockburn et al., 2018; Brynjolfsson et al., 2021). Labor demand may either increase or decrease, depending on whether innovation and AI deployment function primarily as labor-augmenting technologies that require complementary human input in R&D and organizational implementation, or as labor-displacing technologies that substitute for existing tasks.

To test whether AI adoption affects labor input, we re-estimate Equation (6) with the log number of employees as the dependent variable. Columns (4)–(6) of Table 8 Panel A show that the coefficients on *AI* are insignificant, indicating that adopters do not experience significant changes in employment after adoption relative to non-adopters. Relatedly, Panel B of Table 5, reported earlier, shows that AI adopters exhibit larger post-adoption increases in the *Inventor pool* than non-adopters, pointing to an expansion of innovative labor inputs despite stable overall headcount. The increase in R&D spending following adoption (Panel A of Table 7) is also consistent with greater allocation of resources to innovative activities. Taken together, the absence of employment growth alongside a larger inventor pool and higher R&D spending suggests that AI adoption is associated with a shift in workforce composition toward inventive and research roles rather than broad-based hiring. These findings are consistent with Hampole et al. (2025) who document modest net employment effects of workers' exposure to AI, reflecting offsetting forces of displacement and reallocation.

Repeating the analysis for capital yields no significant post-adoption differences between adopters and non-adopters as shown in columns (7)–(9). Combined with the evidence that AI adoption strengthens innovative output and efficiency, the increase in value-added TFP without significant growth in employment or tangible capital is consistent with the innovation channel, whereby AI raises the productivity of innovative inputs and the resulting innovations

are embodied in firm-level efficiency gains.[20]

These findings are consistent with endogenous growth models in which productivity growth arises from the production of new ideas through R&D activities (Romer, 1990; Aghion and Howitt, 1992). Within this framework, AI can raise productivity by improving the efficiency of the innovation process. The increases in inventor productivity and innovation efficiency documented in Table 5 Panel A provide evidence consistent with this mechanism. By enhancing the productivity of innovative effort, AI enables firms to generate and implement new ideas more efficiently, which can raise firm-level value-added TFP even without significant increases in physical capital or employment.

7.5 Contribution to Aggregate Productivity

Our results show that AI adopters experience higher value-added TFP following adoption than non-adopters. To gauge the potential aggregate effects, we map the estimated firm-level effects of AI adoption into an approximate contribution to aggregate value-added productivity using a first-order Hulten-style aggregation (Hulten, 1978). In the canonical Hulten-Domar framework, producer-level productivity shocks are aggregated using gross-output-based Domar weights. Because our firm-level productivity measure is value-added TFP, we implement the aggregation using firms' value-added shares, so the resulting measure should be interpreted as an approximation to aggregate value-added productivity that abstracts from intermediate-input propagation. A first-order approximation implies that the change in aggregate productivity can be written as the weighted sum of firm-level productivity changes:

$$d \ln A_h^{\text{agg}} = \sum_{i \in S} w_i \, d \ln A_{ih} + \sum_{j \in N} w_j \, d \ln A_{jh} \tag{11}$$

[20] We also conduct cohort-specific event studies with lead–lag specifications to assess pre-adoption parallel trends and to trace the dynamics of post-adoption effects on productivity and other economic outcomes. As shown in Figure A.2, there is no evidence of pre-trends, and increases in *R&D*, *Tobin's q*, *Total q*, and *Value-added TFP* for adopters relative to non-adopters begin after AI adoption. Additionally, we implement a control function approach to address endogeneity concerns, using *Value-added TFP* and other economic outcomes as dependent variables. The results are robust, as reported in Table A.5 and Table A.6.

where S denotes AI adopters, N denotes non-adopters, $d \ln A_{ih}$ ($d \ln A_{jh}$) is the horizon-h change in firm-level value-added TFP for adopter i (non-adopter j), and w_i (w_j) denotes the firm's Domar-style weight, defined as its value-added share of GDP.

Let $w^S \equiv \sum_{i \in S} w_i$ and $w^N \equiv \sum_{j \in N} w_j$, and define the corresponding weighted-average value-added TFP changes for adopters and non-adopters as

$$g_h^S \equiv \frac{\sum_{i \in S} w_i d \ln A_{ih}}{w^S}, \qquad g_h^N \equiv \frac{\sum_{j \in N} w_j d \ln A_{jh}}{w^N}$$

Equation (11) can then be written as

$$d \ln A_h^{\text{agg}} = w^S g_h^S + w^N g_h^N \tag{12}$$

In our implementation, w^S is computed as AI adopters' aggregate value-added share of GDP in 2015 levels, thereby holding base-period shares fixed and abstracting from endogenous reallocation.[21] Empirically, the stacked DiD estimate $\hat{\beta}_h$ captures the average post-adoption change in value-added TFP for adopters relative to non-adopters at horizon h. We use $\hat{\beta}_h$ as an approximation to the difference in weighted-average value-added TFP changes between adopters and non-adopters:

$$\hat{\beta}_h \approx g_h^S - g_h^N \tag{13}$$

To capture potential spillover and competitive effects on non-adopters, we estimate an auxiliary specification to obtain the estimate of $\hat{\beta}_h^N$, which we use to approximate the average change in value-added TFP among non-adopters:

$$\hat{\beta}_h^N \approx g_h^N \tag{14}$$

[21] In Domar (1961) and Hulten (1978), when intermediate inputs matter, the first-order effect of a producer-level productivity change on aggregate output is weighted by the producer's gross output relative to aggregate final demand, so the canonical Domar weights are sales based. As an implementation choice, we instead use value-added shares to align the aggregation with our firm-level measure of productivity. As a result, the aggregate contribution should be interpreted as an approximation to aggregate value-added productivity that abstracts from input-output propagation.

Substituting Equation (13) and Equation (14) into Equation (12) yields:

$$d \ln A_h^{\text{agg}} \approx w^S \hat{\beta}_h + (w^S + w^N)\hat{\beta}_h^N \tag{15}$$

When non-adopters do not experience significant changes in value-added TFP around the adoption year ($\hat{\beta}_h^N = 0$), Equation (15) simplifies to $d \ln A_h^{\text{agg}} \approx w^S \hat{\beta}_h$. More generally, when non-adopters are also affected, for example through spillover or competitive effects, the aggregate effect includes their contribution through the second term in Equation (15).

Table 8 Panel A shows that AI adopters experience an average increase in firm-level value-added TFP of 0.085 log points relative to non-adopters in a typical post-adoption year, where $0.085 = (0.082 + 0.08 + 0.093)/3$ is the average of estimated effects in years $t + 1$, $t + 2$, and $t + 3$ after adoption. This corresponds to an increase of approximately 8.87% ($e^{0.085} - 1$). Given that AI adopters' value added accounts for 8.72% of GDP, their contribution to aggregate value-added TFP is approximately 0.0074 log points (0.0872×0.085), which is about 0.74% ($e^{0.0074} - 1$).

To assess whether non-adopters experience changes in value-added TFP around the adoption year, we re-estimate the stacked DiD specification by replacing the indicator *AI* with its components. Let *Post* equals 1 in the post-adoption year relative to a cohort's adoption year, and 0 otherwise. Let *Adopter* equals 1 for adopters in that cohort, and 0 otherwise. By construction, $AI \equiv Post \times Adopter$. We then estimate a stacked DiD regression of firm-level value-added TFP on *Post* and $Post \times Adopter$ with firm-by-cohort fixed effects that absorb all time-invariant firm heterogeneity within each cohort, including the effect of *Adopter*, and calendar year fixed effects that absorb shocks common to all firms in a given year. Standard errors are clustered at the firm level to account for serial correlation and repeated appearances of firms across cohorts. In this specification, the coefficient on *Post* captures the average change in value-added TFP around adoption for non-adopters, while the coefficient on $Post \times Adopter$ captures the incremental change for AI-adopters relative to non-adopters.

As shown in Table 8 Panel B, the coefficients of *Post* are negative and statistically significant. The average of estimated effects on firm-level value-added TFP of non-adopter is -0.007 log points $(-(0.009 + 0.007 + 0.005)/3)$ in a typical post-adoption year. Since non-adopters account for 9.89% of GDP in 2015 levels, their direct contribution is approximately −0.0007 log points (−0.007×0.0989), corresponding to about −0.07% ($e^{-0.0007} - 1$). This indicates a negative contribution of non-adopters to aggregate value-added TFP, consistent with spillover or competitive effects that partially offset aggregate gains from AI adoption. Applying Equation (15), the implied aggregate effect is 0.0061 log points (0.0872×0.085+(0.0872+0.0989)×(−0.007)). This corresponds to an increase of roughly 0.61% ($e^{0.0061} - 1$) in aggregate value-added TFP in a typical post-adoption year, indicating that aggregate gains from AI adoption remain positive but are partially offset by negative effects on non-adopters. This estimate is on the conservative side, as it abstracts from input–output propagation and general equilibrium reallocation.

8 Robustness Checks

8.1 Entropy Balancing

One potential concern is that AI-adopting firms may differ systematically from non-adopting firms in observable characteristics, which could drive the observed effects. To mitigate this concern, we implement entropy balancing following Hainmueller (2012) to achieve covariate balance between treated (AI-using) and control (non-using) firms. Entropy balancing reweights the control group to match the treatment group on specified moments of selected covariates. Specifically, we balance on the means, variances, and skewnesses of *Ln(Total assets)*, *Cash*, and *Tobin's q*, using firm-level averages from financial years prior to 2015. The resulting weights are then applied to the full sample to ensure a balanced comparison across the treatment and control groups. Using the weighted sample, we re-estimate the baseline stacked DiD model (Equation (6)) and report the results in Panel A of Table 9. The estimates are consistent with the

baseline results, indicating that our findings are not driven by observable differences between AI-adopting and non-adopting firms.

8.2 Placebo Test

To further mitigate the concern that our results may be driven by pre-existing differences between AI-adopting and non-adopting firms, or some unobserved shocks coinciding with adoption that directly affect firms' innovation, we conduct a placebo test in which AI adoption is randomly assigned to firms that never adopt AI. For each year, we preserve the actual number of adopters observed in the original data and then randomly assign the same number of placebo adopters among the never-adopting firms. The assignment is performed without replacement, so each placebo-treated firm can be assigned at most one fictitious adoption year. We then construct a placebo treatment indicator based on these assignments and re-estimate Equation (6) using this placebo sample. Panel B of Table 9 reports the estimation results. Across all specifications, the coefficients on AI are significantly negative or statistically insignificant. These results indicate that our results are unlikely to be driven by pre-existing differences between adopting and non-adopting firms or other confounding effects.

8.3 Alternative Control Group

In our baseline analysis, the control group includes both firms that never adopt AI technology (never-treated) and the pre-adoption observations of AI-adopting firms (not-yet-treated). A potential concern is that including not-yet-treated firms in the control group might introduce bias if their innovation behavior starts changing in anticipation of AI adoption. To assess the sensitivity of our results to the composition of the control group, we re-estimate the baseline stacked DiD model (Equation (6)) using only never-treated firms as the control group. Panel C of Table 9 shows that the coefficients on AI are all positive and significant, which are similar to the baseline estimates. This robustness check strengthens our identification strategy by showing that the baseline results are not driven by pre-treatment contamination or anticipatory

behavior among future adopters.

8.4 Excluding AI Patents

The innovation measures in our baseline analysis are constructed from USPTO utility patents and include patents that explicitly utilise AI-related technologies. If AI adoption has a meaningful impact on firm-level innovation beyond AI-specific technologies, the estimated effects should remain when AI patents are excluded from the construction of these measures. To examine whether the effect of AI adoption is limited to AI-related invention, we exclude AI patents identified by Giczy et al. (2022) when constructing the innovation measures and re-estimate Equation (6). As shown in Panel D of Table 9, the results are similar to our baseline estimates, supporting the interpretation that AI adoption enhances broader corporate innovation capabilities rather than merely stimulating activity within AI-specific technological domains.

9 Conclusions

This paper investigates how AI adoption affects productivity growth through the quantity, quality, and direction of corporate innovation. Using a novel firm-level, usage-based measure of staggered AI adoption and a stacked DiD design, we find that adopters generate more patents and that their post-adoption patents receive more citations than those of non-adopters. AI adoption also reshapes innovation characteristics: post-adoption patents have more claims, higher originality and generality, span more technologically distant classes, and adopters shift toward more exploitative patents that build on their existing technologies. These results are robust to a control-function approach that addresses endogenous AI adoption and to a wide range of additional robustness checks.

Channel tests highlight three mechanisms: improved efficiency, shifts in inventor composition, and enhanced knowledge recombination. We also find that the impact of AI adoption

is more pronounced among firms with a focused business scope, consistent with the idea that AI technologies achieve better performance when fueled by high-quality and consistent data. Overall, our findings indicate that AI enhances firms' ability to process and apply knowledge, thereby increasing the quantity, improving the quality, and refining the characteristics of corporate innovation.

Furthermore, we examine how AI adoption relates to firms' productivity and other economic outcomes. Following adoption, AI-adopting firms exhibit an approximately 8.87% increase in value-added TFP and higher investment in R&D, with little evidence of a corresponding expansion in tangible capital expenditure or a reduction in operating costs relative to non-adopters. This pattern is more consistent with productivity gains operating through an innovation channel, in which AI raises research efficiency and reshapes the innovation process, than with immediate cost cutting or physical-capital driven scale expansion. Consistent with this interpretation, equity markets assign higher valuations to AI adopters after adoption, suggesting that investors capitalize expected future cash flows and growth opportunities from AI-enabled innovation. To gauge the potential aggregate implications of the firm-level TFP gains, we apply a first-order Hulten-style aggregation and estimate a contribution of approximately 0.61% to aggregate value-added TFP growth in a representative post-adoption year. This implied aggregate contribution is a first-order, partial-equilibrium approximation that abstracts from the pace of diffusion and from general equilibrium forces such as input-output linkages, endogenous reallocation, and spillovers across firms and sectors. To the extent that mechanisms such as intermediate-input propagation and factor reallocation amplify productivity gains, these estimates should be interpreted as conservative. Quantifying these forces is an important direction for future research.

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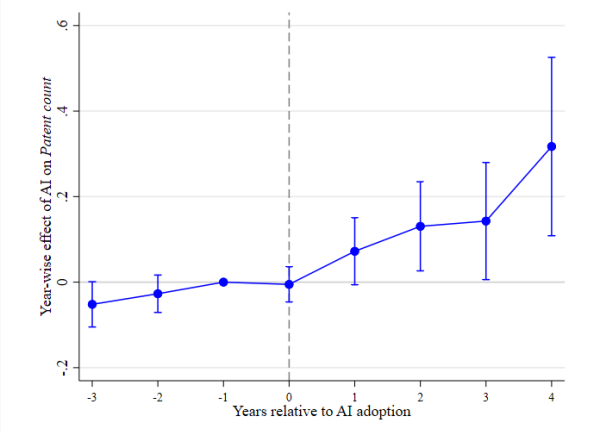
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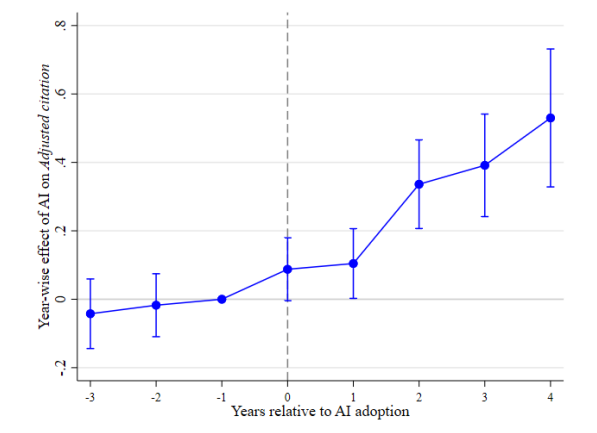
Figure 1
Dynamic Effects of AI Adoption

This figure reports the estimated coefficients from the lead-lag model, capturing the dynamic effects of AI adoption. The horizontal axis denotes years relative to AI adoption. The horizontal axis varies from year -3 to year 4, with year 0 representing the year of adoption within each cohort. We use year -1 as the reference point. Error bars indicate 95% confidence intervals. In Panel A, Figure (a) shows the dynamic treatment effects of AI adoption on innovation quantity measured by *Patent count*. Figure (b) show the dynamic treatment effects of AI adoption on innovation quality measured by *Adjusted citation*. Panel B, C, and D show the dynamic treatment effects of AI adoption on innovation characteristics in terms of innovation novelty, innovation direction, and innovation reach, respectively. In Panel B, Figure (c) and (d) present the dynamic treatment effects of AI adoption on innovation novelty measured by *Originality* and *Distance*. In Panel C, Figure (e) and (f) present the dynamic treatment effects of AI adoption on innovation direction measured by *Exploitative* and *Explorative*. In Panel D, Figure (g) and (h) present the dynamic treatment effects of AI adoption on innovation reach measured by *Generality* and *Claim*. In all panels, we use stacked difference-in-differences (DiD) sample and apply the Pseudo-Poisson Maximum Likelihood estimation method. We use one year before adoption as the reference period. All specifications include Year-Cohort FE and Firm-Cohort FE. Standard errors are clustered at the firm level. Variable definitions and sources are provided in Appendix A.1.

Panel A: Innovation Quantity and Quality



(a) *Patent count*

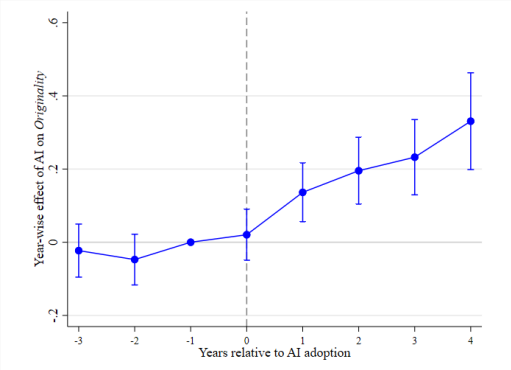


(b) *Adjusted citation*

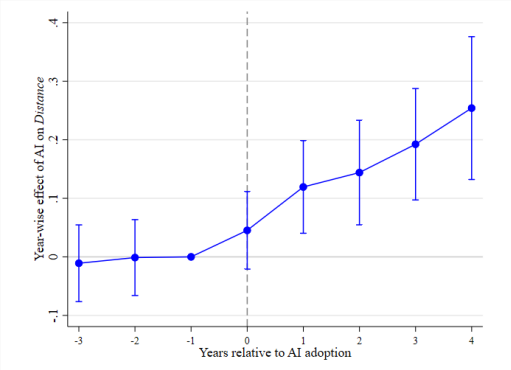
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Figure 1 (continued)

Panel B: Innovation Novelty

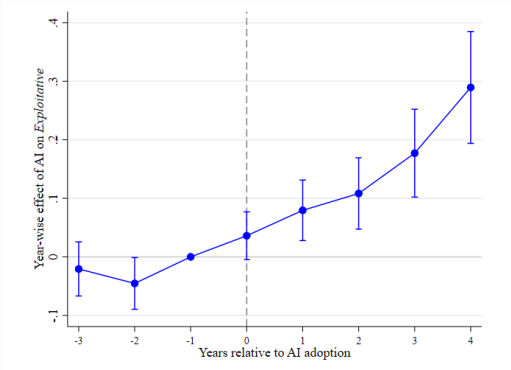


(c) *Originality*

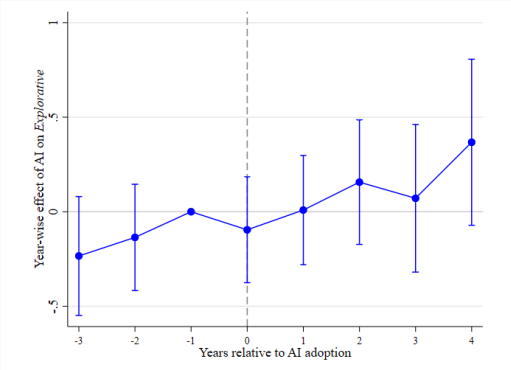


(d) *Distance*

Panel C: Innovation Direction

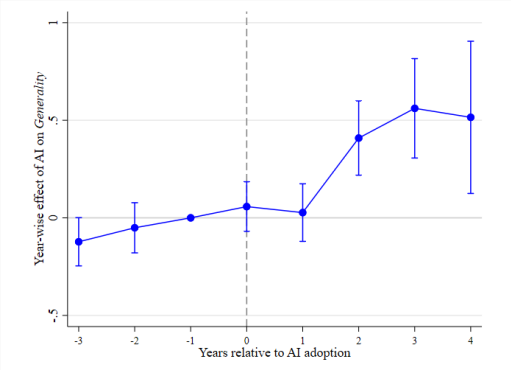


(e) *Exploitative*

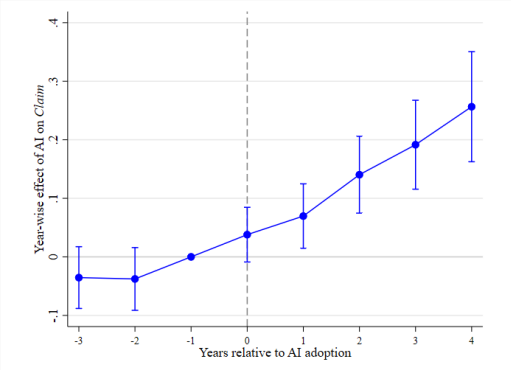


(f) *Explorative*

Panel D: Innovation Reach



(g) *Generality*



(h) *Claim*

Table 1
Summary Statistics

This table reports the summary statistics for all firms in the firm-year panel sample from 2011 to 2021. *All firms* refer to all firms in the sample. *AI-adopting firms* refer to firms that ever adopts AI technology from 2011 to 2021. *Never-adopting firms* refer to firms that never adopt AI technology from 2011 to 2021. Columns (1)-(3) report the summary statistics for all firms, columns (4)-(6) report the summary statistics for AI-adopting firms, and columns (7)-(9) report the summary statistics for never-adopting firms. Variable definitions and sources are provided in Appendix Table A.1.

Variable	All firms			AI-adopting firms			Never-adopting firms		
	Mean	Median	S.D.	Mean	Median	S.D.	Mean	Median	S.D.
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Patent count	24.685	2.000	74.205	65.971	9.000	123.191	9.308	1.000	32.161
Adjusted citation	0.521	0.248	0.753	0.672	0.529	0.760	0.465	0.127	0.743
Originality	0.271	0.265	0.254	0.291	0.306	0.223	0.264	0.233	0.264
Distance	0.242	0.143	0.269	0.295	0.266	0.258	0.222	0.068	0.271
Exploitative	0.499	0.636	0.458	0.643	0.929	0.438	0.446	0.400	0.454
Explorative	0.140	0.000	0.299	0.103	0.000	0.253	0.153	0.000	0.313
Generality	0.093	0.000	0.142	0.109	0.071	0.130	0.087	0.000	0.146
Claim	12.262	15.500	9.463	14.079	17.060	8.432	11.585	14.326	9.734
Value-added TFP	0.001	-0.049	0.670	-0.129	-0.173	0.606	0.051	0.004	0.686
Labor	0.791	0.924	2.075	2.358	2.405	1.649	0.192	0.289	1.901
Capital	5.741	5.750	2.461	5.111	5.095	2.284	6.669	6.819	2.025
Ln(total assets)	7.003	6.987	2.087	8.661	8.693	1.812	6.386	6.338	1.830
Cash	0.248	0.153	0.250	0.184	0.127	0.178	0.272	0.171	0.268
Tobin's q	2.414	1.762	1.855	2.481	1.856	1.797	2.390	1.721	1.876
R&D	0.078	0.021	0.127	0.047	0.017	0.071	0.090	0.023	0.141
Number of firms		2,600			518			2,082	

Table 2
AI and Innovation: Quantity and Quality

This table presents results from stacked DiD regressions estimating the impact of AI adoption on innovation outcomes. All models are estimated using Pseudo-Poisson Maximum Likelihood. Panel A presents results on patent quantity measured by *Patent count* from $t + 1$ to $t + 3$ years following AI adoption in year t. Panel B presents results on patent quality measured by *Adjusted citation* from $t + 1$ to $t + 3$ years following AI adoption in year t. *AI* is an indicator variable that equals 1 if firm i in cohort c uses AI technology in year t, and 0 otherwise. All specifications include Year-Cohort FE and Firm-Cohort FE. Standard errors are clustered at the firm level. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

Panel A: Innovation Quantity			
	Patent count		
Variable	t+1	t+2	t+3
	(1)	(2)	(3)
AI	0.131***	0.149***	0.175***
	(3.02)	(3.60)	(4.56)
Year-Cohort FE	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes
Pseudo R^2	0.928	0.933	0.938
Observations	60,738	48,606	38,245
Panel B: Innovation Quality			
	Adjusted citation		
Variable	t+1	t+2	t+3
	(1)	(2)	(3)
AI	0.265***	0.277***	0.331***
	(6.05)	(5.71)	(5.96)
Year-Cohort FE	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes
Pseudo R^2	0.261	0.255	0.248
Observations	53,623	42,190	32,551

Table 3
AI and Innovation: Characteristics

This table presents results from stacked DiD regressions estimating the impact of AI adoption on innovation outcomes. In Panel A, we present results on innovation novelty measured by *Originality* and *Distance* from $t+1$ to $t+3$ years following AI adoption in year t. In Panel B, we present results on innovation direction measured by *Exploitative* and *Explorative* from $t+1$ to $t+3$ years following AI adoption in year t. In Panel C, we present results on innovation reach measured by *Generality* and *Claim* from $t+1$ to $t+3$ years following AI adoption in year t. All models are estimated using Pseudo-Poisson Maximum Likelihood. *AI* is an indicator variable that equals 1 if firm i in cohort c uses AI technology in year t, and 0 otherwise. All specifications include Year-Cohort FE and Firm-Cohort FE. Standard errors are clustered at the firm level. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

Panel A: Innovation Novelty						
	Originality			*Distance*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)
AI	0.169***	0.214***	0.195***	0.137***	0.162***	0.125***
	(5.55)	(6.52)	(5.20)	(4.54)	(4.98)	(3.51)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.0885	0.0846	0.0805	0.116	0.117	0.118
Observations	56,947	45,518	35,653	51,366	42,174	33,865
Panel B: Innovation Direction						
	Exploitative			*Explorative*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)
AI	0.147***	0.162***	0.154***	0.185**	0.204**	0.191*
	(6.50)	(6.59)	(5.69)	(2.12)	(1.97)	(1.69)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.109	0.104	0.0988	0.139	0.144	0.152
Observations	48,478	38,951	30,639	51,702	39,704	30,331
Panel C: Innovation Reach						
	Generality			*Claim*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)
AI	0.298***	0.317***	0.449***	0.142***	0.155***	0.161***
	(5.33)	(4.93)	(5.34)	(6.01)	(6.13)	(5.69))
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.115	0.112	0.109	0.308	0.304	0.299
Observations	40,887	31,240	23,334	60,678	48,574	38,215

Table 4
The Impact of AI on Innovation: Control Function Approach

This table presents the results of Control Function (CF) regressions on the effect of AI on corporate innovation using the stacked DiD sample. The first-stage result on *AI exposure* is reported in column (1). The second-stage results on corporate innovation measures in year $t+3$ following AI adoption in year t are reported in columns (2)-(9). Column (2) presents results on innovation quantity, column (3) presents results on innovation quality, and columns (4)-(9) present results on innovation characteristics. We estimate the probit model in the first stage. In the second stage, we use Pseudo-Poisson Maximum Likelihood method. *AI exposure* is firm *i*'s exposure to AI-strong universities via the firm-university STEM worker hiring networks as of 2010 (Babina et al., 2024). *Top 10 exposure* is firm *i*'s exposure to top 10 universities ranked by 2010 U.S. News & World Report via the firm-university STEM worker hiring networks as of 2010 (Babina et al., 2024). *Residual* is the regression residuals from the first-stage estimation. All regressions include Year-Cohort FE and Industry-Cohort FE. Industries follow the Fama-French 49 Industry Classification. Standard errors are computed using a firm-level cluster bootstrap with 1,000 replications. *t*-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

	1st Stage	2nd Stage							
Variable	*AI*	*Patent count*	*Adjusted citation*	*Originality*	*Distance*	*Exploitative*	*Explorative*	*Generality*	*Claim*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
AI exposure	0.572***								
	(4.68)								
AI		6.118***	1.485***	1.071***	0.832**	1.664***	-0.811	1.822**	1.348***
		(4.36)	(2.17)	(2.89)	(1.96)	(4.42)	(-0.93)	(2.03)	(3.68)
Top 10 exposure	-0.685	0.477	1.061**	0.274	0.367	0.388	0.730	1.266**	0.469**
	(-1.54)	(0.74)	(2.41)	(0.93)	(1.00)	(1.60)	(0.52)	(2.11)	(2.03)
Residual		-2.240***	-0.473	-0.394**	-0.194	-0.608***	0.311	-0.646	-0.479***
		(-2.95)	(-1.26)	(-1.99)	(-0.86)	(-3.00)	(0.65)	(-1.30)	(-2.61)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Industry-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Wald χ^2	541.82								
Pseudo R^2	0.1149	0.427	0.0753	0.0464	0.0560	0.0565	0.0310	0.0466	0.0969
Observations	19,918	6,271	6,199	6,271	6,271	6,213	6,243	5,939	6,271

Table 5
The Impact of AI on Innovation: Channel

This table presents results from stacked DiD regressions examining the channels through which AI affects corporate innovation. Panel A reports results for the efficiency channel, measured by *Inventor productivity* and *Innovation efficiency*. Panel B presents results for the inventor composition channel, measured by *Inventor pool*, *First-time inventor*, *AI-experienced inventor*, and *Specialization*. Panel C reports results for the knowledge recombination channel, measured by *Team size*, *Unique inventor team*, *Core skill breadth*, and *New cite ratio*. In all panels, we present Pseudo-Poisson Maximum Likelihood results from $t+1$ to $t+3$ years following AI adoption in year t. *AI* is an indicator variable that equals 1 if firm i in cohort c uses AI technology in year t, and 0 otherwise. All specifications include Year-Cohort FE and Firm-Cohort FE. Standard errors are clustered at the firm level. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

Panel A: Efficiency Channel

	Inventor productivity			*Innovation efficiency*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)
AI	0.144***	0.166***	0.162***	0.161***	0.149***	0.166***
	(4.73)	(5.24)	(4.32)	(5.02)	(4.22)	(4.27)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.142	0.139	0.136	0.122	0.121	0.120
Observations	60,738	48,606	38,245	58,601	46,813	36,746

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Table 5 (continued)

Panel B: Inventor Composition												
	Inventor pool			*First-time inventor*			*AI-experienced inventor*			*Specialization*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
AI	0.077*	0.094**	0.127***	0.164***	0.214***	0.199***	0.064*	0.070*	0.037	0.148***	0.144***	0.148***
	(1.65)	(2.16)	(3.43)	(2.93)	(3.52)	(2.80)	(1.69)	(1.75)	(0.87)	(6.37)	(5.32)	(5.02)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.948	0.952	0.956	0.0930	0.0932	0.0955	0.177	0.172	0.167	0.0927	0.0886	0.0844
Observations	60,738	48,606	38,245	51,044	40,567	31,523	46,628	37,336	29,300	59,355	47,522	37,417
Panel C: Knowledge Recombination												
	Team size			*Unique inventor team*			*Core skill breadth*			*New cite ratio*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
AI	0.142***	0.150***	0.154***	0.080*	0.106**	0.159***	0.136***	0.144***	0.142***	0.160***	0.163***	0.173***
	(5.69)	(5.43)	(5.08)	(1.67)	(2.35)	(4.17)	(5.75)	(5.44)	(5.00)	(6.37)	(5.93)	(5.68)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.237	0.231	0.227	0.939	0.944	0.948	0.142	0.138	0.133	0.0994	0.0970	0.0950
Observations	60,738	48,606	38,245	60,738	48,606	38,245	59,355	47,522	37,417	60,283	48,203	37,910

Table 6
The Impact of AI on Innovation: Heterogeneity

This table presents results from stacked DiDiD regressions examining cross-sectional heterogeneity in the impacts of AI on corporate innovation. Panels A–D present results on the differential effects of AI adoption across firms with different business focus, size, financial constraints, and technological capabilities. In all panels, we present Pseudo-Poisson Maximum Likelihood results in year $t + 3$ following AI adoption in year t. *AI* is an indicator variable that equals 1 if firm i in cohort c uses AI technology in year t, and 0 otherwise. Across Panels A–E, all specifications include Year–Cohort and Firm–Cohort fixed effects, as well as subgroup-specific Year–Cohort and Firm–Cohort fixed effects obtained by interacting these fixed effects with indicator variables for business focus, firm size, financial constraints, and technological capabilities, respectively. Standard errors are clustered at the firm level. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

Panel A: Business Focus

Variable	*Patent count*	*Adjusted citation*	*Originality*	*Distance*	*Exploitative*	*Explorative*	*Generality*	*Claim*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI	0.175***	0.315***	0.178***	0.119***	0.148***	0.144	0.420***	0.134***
	(4.41)	(5.52)	(4.50)	(3.20)	(5.08)	(1.21)	(4.70)	(4.52)
AI x *Focused firm*	0.445	0.368	0.411***	0.370**	0.251**	0.374	0.592	0.348***
	(1.63)	(1.16)	(2.98)	(2.12)	(2.37)	(0.88)	(1.37)	(2.84)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year-Cohort-Focused firm FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort-Focused firm FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.938	0.251	0.0812	0.118	0.100	0.142	0.108	0.301
Observations	36,792	39,862	34,292	32,743	29,514	37,663	22,786	36,762

Panel B: Firm Size

Variable	*Patent count*	*Adjusted citation*	*Originality*	*Distance*	*Exploitative*	*Explorative*	*Generality*	*Claim*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI	0.524**	-0.017	0.281**	-0.026	0.036	-0.057	0.071	0.057
	(2.08)	(-0.12)	(2.34)	(-0.21)	(0.37)	(-0.29)	(0.25)	(0.74)
AI x *Big firm*	-0.360	0.295*	-0.131	0.166	0.134	0.302	0.361	0.087
	(-1.41)	(1.83)	(-1.02)	(1.28)	(1.30)	(1.22)	(1.20)	(1.04)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year-Cohort-Big firm FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort-Big firm FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.937	0.247	0.0811	0.117	0.101	0.150	0.109	0.302
Observations	36,215	30,979	33,729	32,171	29,009	28,977	22,333	36,185

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Table 6 (continued)

Panel C: Financial Constraints								
Variable	*Patent count*	*Adjusted citation*	*Originality*	*Distance*	*Exploitative*	*Explorative*	*Generality*	*Claim*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI	0.177***	0.374***	0.205***	0.151***	0.164***	0.208	0.510***	0.169***
	(3.73)	(5.93)	(4.94)	(3.70)	(5.26)	(1.54)	(5.00)	(5.18)
AI x *Unconstrained firm*	-0.018	-0.187	-0.026	-0.076	-0.024	-0.094	-0.256	-0.045
	(-0.24)	(-1.37)	(-0.28)	(-0.89)	(-0.36)	(-0.37)	(-1.38)	(-0.67)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year-Cohort-Unconstrained firm FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort-Unconstrained firm FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.938	0.245	0.0811	0.118	0.100	0.150	0.108	0.301
Observations	36,792	31,509	34,292	32,726	29,514	29,465	22,786	36,762
Panel D: Technological Capabilities								
Variable	*Patent count*	*Adjusted citation*	*Originality*	*Distance*	*Exploitative*	*Explorative*	*Generality*	*Claim*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI	0.138***	0.341***	0.219***	0.135***	0.190***	0.067	0.482***	0.170***
	(3.46)	(5.62)	(5.28)	(3.44)	(6.46)	(0.51)	(5.01)	(5.33)
AI x *High-tech firm*	0.149	-0.029	-0.094	-0.010	-0.165*	0.418	-0.296	-0.045
	(1.11)	(-0.19)	(-0.85)	(-0.09)	(-1.87)	(1.48)	(-1.28)	(-0.57)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year-Cohort-High-tech firm FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort-High-tech firm FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.938	0.245	0.0812	0.118	0.100	0.150	0.108	0.301
Observations	36,792	31,509	34,292	32,743	29,514	29,465	22,786	36,762

Table 7
The Impact of AI on Economic Outcomes

This table presents results from stacked DiD regressions estimating the impact of AI adoption on intangible versus tangible investment, operations, and market value. Panel A reports results on intangible versus tangible investment, measure by *R&D* and *Capital expenditure* respectively. Panel B reports results on operation costs, measured by *Production cost*, *Non-production cost*, *Total operating cost*. Panel C reports results on market value measured by *Tobin's q* and *Total q*. In all panels, we present linear regression results from $t+1$ to $t+3$ following AI adoption in year t. *AI* is an indicator variable that equals 1 if firm i in cohort c uses AI technology in year t, and 0 otherwise. All specifications include Year-Cohort FE and Firm-Cohort FE. Standard errors are clustered at the firm level. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

Panel A: Intangible versus Tangible Investment

	R&D			*Capital expenditure*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)
AI	0.006***	0.006***	0.006***	-0.000	-0.001	-0.001
	(3.19)	(3.09)	(2.92)	(-0.12)	(-0.71)	(-0.58)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Adjusted R^2	0.847	0.857	0.863	0.650	0.663	0.676
Observations	65,054	54,278	44,353	65,018	54,250	44,332

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Table 7 (continued)

Panel B: Operating Costs									
	Production cost			*Non-production cost*			*Total operating cost*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
AI	0.145	0.154	0.121	-0.001	-0.013	-0.017	0.132	0.063	0.000
	(1.04)	(1.14)	(1.01)	(-0.07)	(-0.91)	(-1.30)	(0.61)	(0.30)	(0.00)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Adjusted R^2	0.617	0.619	0.639	0.761	0.787	0.816	0.623	0.632	0.661
Observations	65,054	54,278	44,353	56,546	47,539	39,179	65,054	54,278	44,353

Panel C: Market Value						
	Tobin's q			*Total q*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)
AI	0.273***	0.246***	0.163**	0.329***	0.337***	0.242***
	(4.27)	(3.59)	(2.40)	(3.72)	(3.59)	(2.63)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Adjusted R^2	0.713	0.728	0.752	0.715	0.714	0.710
Observations	64,938	54,174	44,260	64,620	53,917	44,081

Table 8
The Impact of AI on Firm-Level Productivity

This table presents results from stacked DiD regressions that examine the impact of AI adoption on firm-level productivity, labor, and capital input. Panel A reports results on the impact of AI adoption on firm-level productivity, labor, and capital input for AI-adopting firms relative to non-adopting firms. Panel B presents results on whether non-adopters exhibit effects in firm-level productivity, labor, and capital input around cohort AI adoption years. In all panels, we present linear regression results from $t + 1$ to $t + 3$ following AI adoption in year t. *AI* is an indicator variable that equals 1 if firm i in cohort c uses AI technology in year t, and 0 otherwise. *Post* is an indicator variable that equals 1 in post-adoption years t relative to the adoption year of cohort c, and 0 otherwise. *Adopter* is an indicator variable that equals 1 for adopters in cohort c, and 0 otherwise. In Panel A, specifications include Year-Cohort FE and Firm-Cohort FE. In Panel B, specifications include Year FE and Firm-Cohort FE. Industries follow the Fama-French 49 Industry Classification. Standard errors are clustered at the firm level. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

Panel A: Adopters vs. Non-Adopters

	Value-added TFP			*Labor*			*Capital*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
AI	0.082***	0.080***	0.093***	0.019	0.024	0.029	-0.009	0.001	0.007
	(4.73)	(4.44)	(4.67)	(0.86)	(1.11)	(1.39)	(-0.37)	(0.02)	(0.27)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Adjusted R^2	0.878	0.887	0.893	0.977	0.981	0.983	0.979	0.982	0.984
Observations	59,981	49,987	40,841	59,981	49,987	40,841	59,981	49,987	40,841

Panel B: Non-adopter Post Effect

	Value-added TFP			*Labor*			*Capital*		
Variable	t+1	t+2	t+3	t+1	t+2	t+3	t+1	t+2	t+3
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Post	-0.009***	-0.007***	-0.005***	0.002	0.002	0.002	0.003	0.002	0.001
	(-5.59)	(-4.26)	(-3.48)	(0.96)	(1.17)	(1.14)	(1.52)	(0.87)	(0.64)
Adopter x *Post*	0.084***	0.083***	0.094***	0.020	0.025	0.030	-0.009	0.001	0.007
	(4.84)	(4.53)	(4.72)	(0.89)	(1.15)	(1.42)	(-0.36)	(0.05)	(0.28)
Year FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Adjusted R^2	0.878	0.887	0.893	0.977	0.981	0.983	0.979	0.982	0.984
Observations	59,981	49,987	40,841	59,981	49,987	40,841	59,981	49,987	40,841

Table 9
AI and Innovation: Robustness Tests

This table presents the results of stacked DiD regressions on the effect of AI on corporate innovation in different robustness tests. Panel A reports the results of stacked DiD regressions using an entropy-balanced sample. The firm-year panel sample is re-weighted based on the averages of *Ln(total assets)*, *Cash*, and *Tobin's q* in financial years before 2015 so that the means, variances, and skewnesses of the control group match the values of the same moments in the treated group between 2011 to 2015. Panel B reports the results of stacked DiD regressions from a placebo test in which AI adoption is randomly assigned to firms that never adopt AI. For each year, we preserve the actual number of adopters observed in the original data, switch off the treatment status from true adopters, and then randomly assign the same number of "placebo adopters'' among the never-adopting firms. The assignment is performed without replacement, ensuring that each placebo-treated firm can be assigned an adoption year at most once. Panel C presents the results of stacked DiD regressions using the never-treated firms (i.e., firms that never adopt AI technology from 2011 to 2021) as the control group. Panel D reports the stacked DiD estimates of the effect of AI on non-AI innovation. We exclude AI patents identified by Giczy et al. (2022) from the patent-level sample and re-construct corporate innovation measures without AI patents. All models are estimated using Pseudo-Poisson Maximum Likelihood. In all panels, we present the results on corporate innovation measures in year $t+3$ following AI adoption in year t. Column (1) presents the results on innovation quantity, column (2) presents the results on innovation quality, and columns (3)-(8) present results on innovation characteristics. *AI* is an indicator variable that equals 1 if firm i in cohort c uses AI technology in year t, and 0 otherwise. All specifications include Year-Cohort FE and Firm-Cohort FE. Standard errors are clustered at the firm level. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

Panel A: Entropy Balancing

Variable	*Patent count*	*Adjusted citation*	*Originality*	*Distance*	*Exploitative*	*Explorative*	*Generality*	*Claim*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI	0.129*	0.153**	0.111***	0.043	0.083***	0.183	0.240**	0.100***
	(1.87)	(2.18)	(2.70)	(0.96)	(2.75)	(1.35)	(2.23)	(3.17)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R-squared	0.944	0.209	0.0710	0.100	0.0781	0.181	0.0874	0.291
Observations	36,096	30,860	33,685	32,273	28,987	29,025	22,386	36,066

(continued on next page)

Table 9 (continued)

Panel B: Placebo Test								
Variable	*Patent count*	*Adjusted citation*	*Originality*	*Distance*	*Exploitative*	*Explorative*	*Generality*	*Claim*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI	-0.211***	-0.274***	-0.086	-0.069	-0.123***	0.084	-0.348**	-0.084*
	(-2.83)	(-3.12)	(-1.57)	(-1.26)	(-2.61)	(0.63)	(-2.50)	(-1.85)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.945	0.229	0.077	0.114	0.091	0.165	0.100	0.305
Observations	34,993	30,227	32,672	31,273	28,429	28,937	22,629	34,976
Panel C: Never-Treated Firms as the Control Group								
Variable	*Patent count*	*Adjusted citation*	*Originality*	*Distance*	*Exploitative*	*Explorative*	*Generality*	*Claim*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI	0.253***	0.376***	0.225***	0.163***	0.189***	0.201*	0.500***	0.187***
	(4.56)	(6.17)	(5.66)	(4.29)	(6.28)	(1.74)	(5.42)	(6.16)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.918	0.257	0.0852	0.121	0.104	0.138	0.119	0.300
Observations	31,495	26,424	29,313	27,505	24,854	24,607	18,202	31,465
Panel D: Excluding AI Patents								
Variable	Patent count	Adjusted citation	Originality	Distance	Exploitative	Explorative	Generality	Claim
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI	0.212***	0.340***	0.186***	0.119***	0.150***	0.295**	0.448***	0.169***
	(4.86)	(6.25)	(4.95)	(3.37)	(5.62)	(2.44)	(5.30)	(6.05)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.930	0.245	0.0804	0.116	0.0995	0.151	0.116	0.298
Observations	37,694	31,747	34,882	33,117	30,268	29,036	22,354	37,672

Appendix A.

Figure A.1
Diffusion of AI Adoption across U.S. Geographies

This figure presents a heat map of the diffusion of AI adoption across the continental United States from 2016 to 2021. The color intensity reflects the number of AI-adopting firms within each U.S. Bureau of Economic Analysis Economic Area in a given year, with darker colors indicating more adopters. Puerto Rico, Alaska, and Hawaii are excluded from the figure. Grey areas denote regions with no available data.

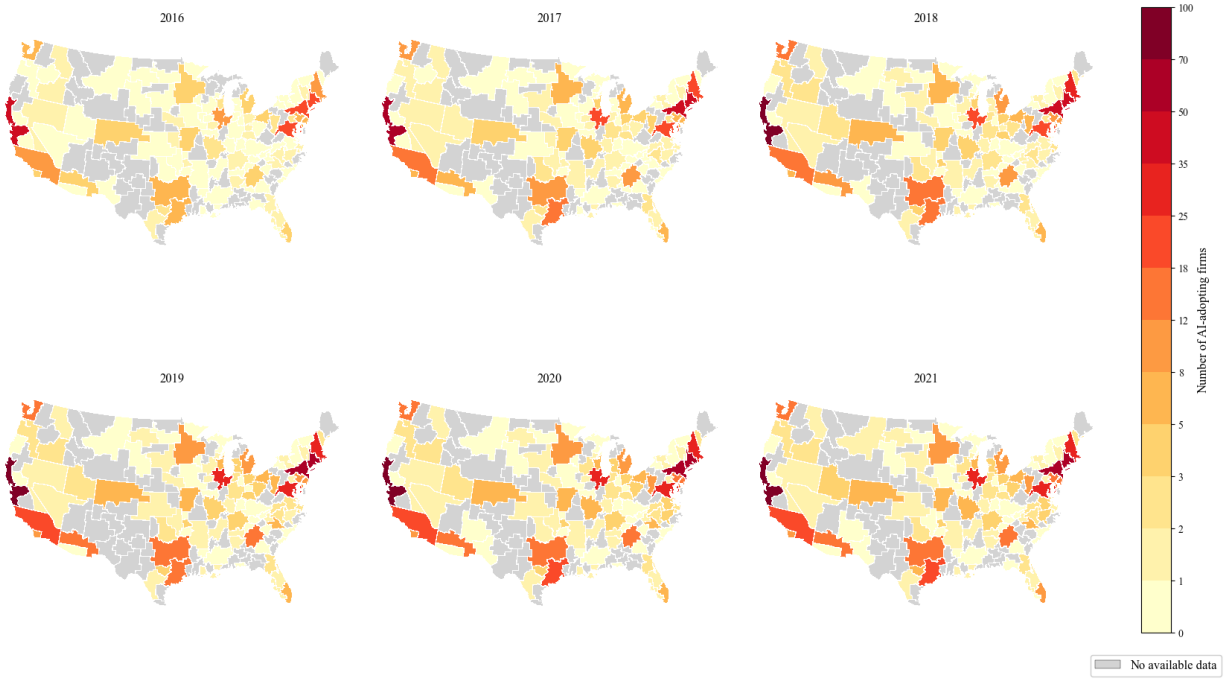
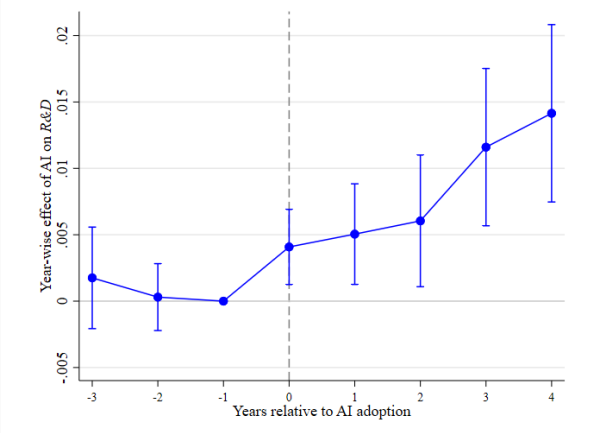


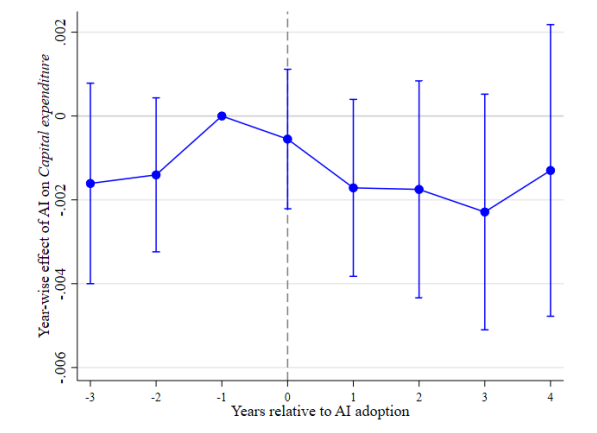
Figure A.2
Dynamic Effects of AI Adoption: Productivity and Economic Outcomes

This figure reports the estimated coefficients from the lead-lag model, capturing the dynamic effects of AI adoption on productivity and economic variables. The horizontal axis denotes years relative to AI adoption. The horizontal axis varies from year -3 to year 4, with year 0 representing the year of adoption within each cohort. We use year -1 as the reference point. Error bars indicate 95% confidence intervals. Panel A, B, C, and D show the dynamic treatment effects of AI adoption on intangible versus tangible investment, operating costs, market value, and firm-level productivity, respectively. In Panel A, Figure (a) and (b) show the dynamic treatment effects of AI adoption on *R&D* and *Capital expenditure*, respectively. In Panel B, Figure (c), (d), and (e) present the dynamic treatment effects of AI adoption on *Production cost*, *Non-production cost*, and *Total operating cost*, respectively. In Panel C, Figure (f) and (g) present the dynamic treatment effects of AI adoption on *Tobin's q* and *Total q*, respectively. In Panel D, Figure (h), (i), and (j) present the dynamic treatment effects of AI adoption on *Value-added TFP*, *Labor*, and *Capital*, respectively. In all panels, we use stacked difference-in-differences (DiD) sample and report linear regression results. We use one year before adoption as the reference period. All specifications include Year-Cohort FE and Firm-Cohort FE. Standard errors are clustered at the firm level. Variable definitions and sources are provided in Appendix A.1.

Panel A: Intangible versus Tangible Investment



(a) *R&D*

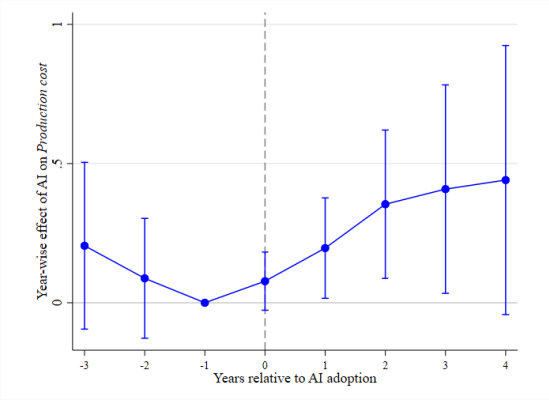


(b) *Capital expenditure*

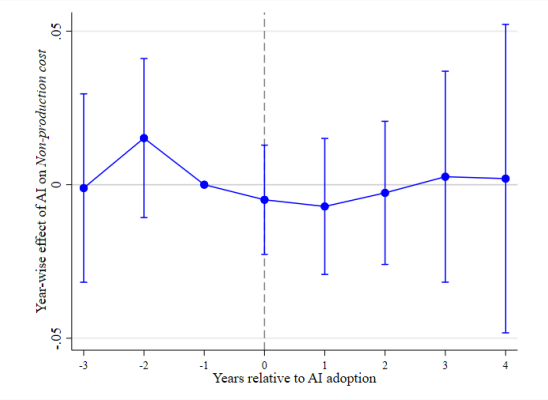
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Figure A.2 (continued)

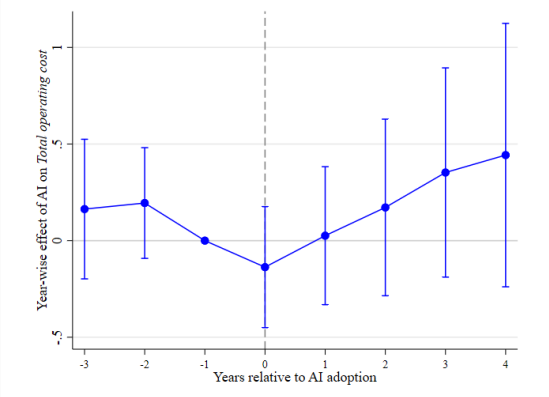
Panel B: Operating Costs



(c) *Production cost*

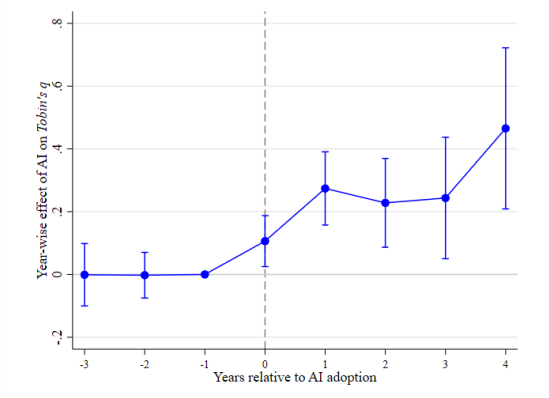


(d) *Non-production cost*

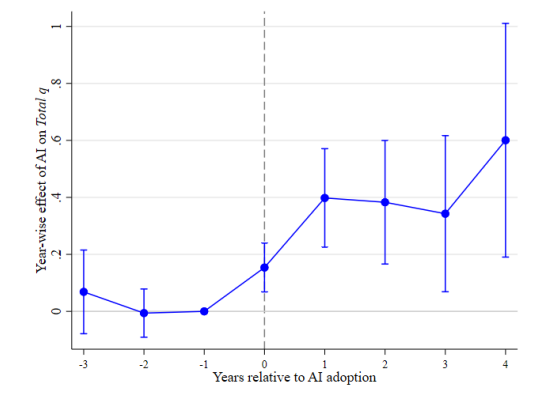


(e) *Total operating cost*

Panel C: Market Value



(f) *Tobin's q*

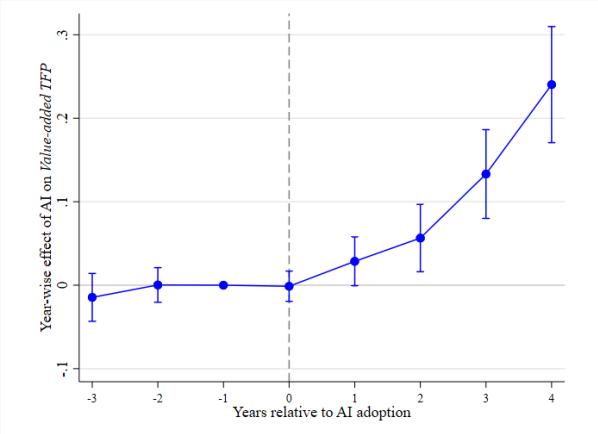


(g) *Total q*

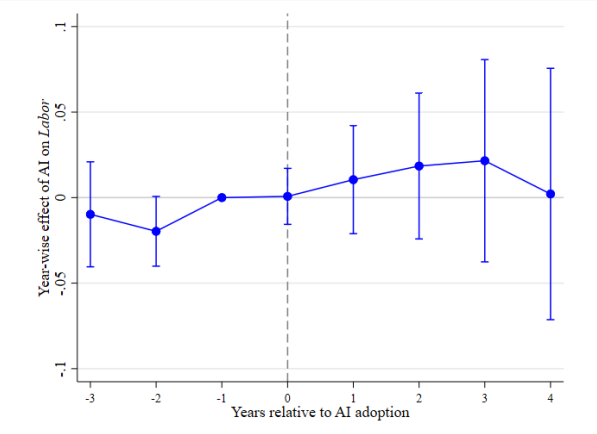
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Figure A.2 (continued)

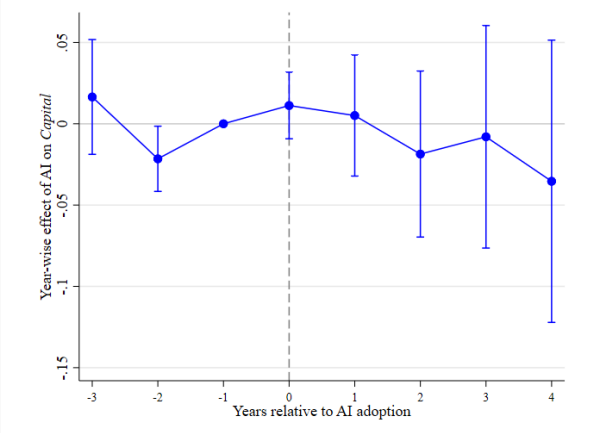
Panel D: Productivity



(h) *Value-added TFP*



(i) *Labor*



(j) *Capital*

Table A.1
Variable Definitions

Variable	Description	Source
	A. Innovation variables	
Patent count	The sum of granted patents filed by firm i in year t.	USPTO
Adjusted citation	The average of the time and technology class adjusted forward citations received by firm i to granted patents applied for in year t. Specifically, adjusted forward citations are the number of citations received by patent p applied by firm i in year t in class k scaled by the average number of citations received by all granted patents applied in year t in class k.	USPTO
Originality	One minus the average Herfindahl–Hirschman Index of cited patents for granted patents filed by firm i in year t.	USPTO
Distance	The average technology distance between patent p and all patents previously filed by firm i up to year t.	USPTO
Exploitative	The proportion of exploitative patents filed by firm i in year t. Patent p is categorized as exploitative if 80% or more of its backward citations are based on firm i's existing expertise. Existing expertise is defined as the set of firm i's granted patents filed from the start of the fifth calendar year prior to the filing date of patent p up to the filing date of patent p, together with the patents cited by those granted patents during the same period.	USPTO
Explorative	The proportion of explorative patents filed by firm i in year t. Patent p is classified as explorative if 80% or more of its backward citations are not based on firm i's existing expertise. Existing expertise is defined as the set of firm i's granted patents filed from the start of the fifth calendar year prior to the filing date of patent p up to the filing date of patent p, together with the patents cited by those granted patents during the same period.	USPTO
Generality	One minus the average Herfindahl–Hirschman Index of citing patents for granted patents filed by firm i in year t.	USPTO
Claim	The average number of claims made by granted patents filed by firm i in year t.	USPTO

(continued on next page)

Table A.1 (continued)

Variable	Description	Source
	B. AI variables	
AI	An indicator variable that equals 1 if firm i in cohort c has adopted AI technology by year t, and 0 otherwise.	SWZD
AI patent	Patent p filed by firm i in year t that is predicted by Giczy et al. (2022) to contain AI technology in any of the eight technology components (machine learning, natural language processing, computer vision, speech, knowledge processing, AI hardware, evolutionary computation, and planning and control) based on a 50% threshold.	Artificial Intelligence Patent Dataset
	C. Instrumental and related variables	
AI exposure	Firm i's exposure to AI-strong universities via the firm-university STEM worker hiring networks as of 2010.	Babina et al. (2024)
Top 10 exposure	Firm i's exposure to top 10 universities ranked by 2010 U.S. News & World Report via the firm-university STEM worker hiring networks as of 2010.	Babina et al. (2024)
	D. Financial variables	
Ln(Total assets)	The natural logarithm of the total assets (at).	Compustat
Leverage	The sum of long-term debt (dltt) and debt in current liabilities (dlc), divided by total assets (at).	Compustat
Cash	The sum of cash and short-term investments (che), divided by total assets (at).	Compustat
Tobin's q	Total assets (at) plus market value of equity (csho x prcc_f) minus book value of equity (ceq), divided by total assets (at). Missing prcc_f values are replaced with prcc_c.	Compustat
Total q	Market value of equity (csho x prcc_f) plus long-term debt (dltt) and debt in current liabilities (dlc), minus current assets (act), divided by the sum of total property, plant and equipment (ppegt) and replacement cost of intangible capital (Peters and Taylor, 2017). Replacement cost of intangible capital is the sum of a firm's externally purchased and internally created intangible capital. Externally purchased intangible capital is total intangible assets (intan), or 0 if the intangible assets are missing. A firm's internal intangible capital is the sum of knowledge capital and organization capital estimated using the perpetual inventory method. The stock of knowledge capital for firm i in year t is estimated by accumulating past R&D spending using the perpetual inventory method: $G_{i,t} = (1 - \delta_{R\&D})G_{i,t-1} + R\&D_{i,t}$, where $\delta_{R\&D}$ is the U.S. Bureau of Economic Analysis (BEA)'s industry-specific R&D depreciation rates and $R\&D_{i,t}$ is R&D expenses (xrd), or 0 if R&D expenses are missing. The stock of organization capital for firm i in year t is estimated by accumulating a fraction of past SG&A spending using the perpetual inventory method. SG&A is measured as selling, general and administrative expenses (xsga) minus the sum of R&D expenses (xrd) and in-process R&D expenses (rdip). Missing selling, general and administrative expenses (xsga), R&D expenses (xrd), and in-process R&D expenses (rdip) are replaced with 0. 30% of SG&A that is assumed to present an investment in intangible capital. The depreciation rate of SG&A is 20%.	Peters and Taylor (2017)

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Table A.1 (continued)

Variable	Description	Source
	D. Financial variables (continued)	
Value-added TFP	Value-added total factor productivity. We estimate firm-level value-added TFP using a log Cobb–Douglas production function and the single-step GMM estimator of Wooldridge (2009). The baseline specification is: $y_{i,t} = \beta_0 + \beta_k k_{i,t} + \beta_l l_{i,t} + \epsilon_{i,t}$, where $y_{i,t}$ is the log of the value added of firm i in year t, $k_{i,t}$ is the log value of capital (K), $l_{i,t}$ is the log value of labor (L). We follow Bennett et al. (2020) in defining value added, capital (K), and labor (L) variables with a minor modification to labor expenses due to limited wage data. Value added is sales minus materials, deflated by the gross domestic product (GDP) implicit price deflator from Federal Reserve Economic Data (FRED). Sales are Compustat revenue (revt). Materials are computed as total expenses minus labor expenses. Total expenses are revenue (revt) less operating income before depreciation and amortization (oibdp). Labor expenses are non-missing wages (xlr). If xlr is missing, we use the non-missing wages (xlr) to calculate the average wage for each Fama-French 49 industry in year t. For firm-years with missing wages, we impute labor expenses as number of employees (emp) × industry-year average wage. Capital (K) is measured as gross plant, property, and equipment (ppegt) deflated by the price index for private fixed investment from Bureau of Economic Analysis (BEA), following the average age adjustment of İmrohoroğlu and Tüzel (2014), Hall (1990), and Brynjolfsson and Hitt (2003). Given that investment is made at different times in the past, we assume capital (K) was installed in a vintage year (calculated as year t minus the average age of capital; values larger or equal to 0.5 are rounded upward and values smaller than 0.5 are rounded downward), and apply the price index for private fixed investment in the corresponding vintage year as the deflator. We estimate the average age of capital for firm i in year t as accumulated depreciation (dpact; if dpact is missing, we use ppegt-ppent) divided by current depreciation (dp), then smooth this average age using a three-year moving average (or available years if fewer than three). Labor (L) is the number of employees (emp). We use intermediate inputs, measured as the log of real materials, as a proxy for unobserved productivity shocks following Levinsohn and Petrin (2003), implemented within the GMM framework of Wooldridge (2009). Real materials are constructed by deflating materials using the GDP implicit price deflator. Prior to estimation, we remove common industry-time variation by demeaning $y_{i,t}$, $k_{i,t}$, $l_{i,t}$, and the log of real materials using Fama–French 49 industry-year fixed effects, following İmrohoroğlu and Tüzel (2014). This transformation absorbs shocks that are common to firms within an industry-year, including shared movements in prices and demand. Firm-level TFP is thus estimated using within-industry-year variation. We then estimate the baseline specification using these residualized variables. *Value-added TFP* is computed as $y^{res}_{i,t} - \hat{\beta}_k k^{res}_{i,t} - \hat{\beta}_l l^{res}_{i,t}$, where $y^{res}_{i,t}$, $k^{res}_{i,t}$, $l^{res}_{i,t}$ denote the residuals from the industry-year fixed effects regressions.	Compustat, FRED, BEA

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Table A.1 (continued)

Variable	Description	Source
	D. Financial variables (continued)	
Labor	The natural logarithm of labor (L). For details of labor (L) please see the definition of *Value-added TFP*.	Compustat
Capital	The natural logarithm of capital (K). For details of capital (K) please see the definition of *Value-added TFP*.	Compustat, BEA
R&D	R&D expenses (xrd) divided by total assets (at), or 0 if R&D expenses are negative or missing.	Compustat
Capital expenditure	Capital expenditures (capx) divided by total assets (at).	Compustat
Production cost	Cost of goods sold (cogs) divided by total sales (sale).	Compustat
Non-production cost	Selling, general and administrative expenses (xsga) minus R&D expenses (xrd) and advertising expenses (xad), divided by total sales (sale). Missing R&D expenses (xrd) and advertising expenses (xad) are replaced with 0.	Compustat
Total operating cost	Total operating expenses (xopr) divided by total sales (sale).	Compustat
	E. Channel variables	
Inventor productivity	The average number of granted patents filed by inventors of firm i in year t. The number of patents created by inventor m is calculated as the sum of all patents filed by inventor m in year t. For patents with multiple inventors, we follow Moretti (2021) to assign equally-weighted fractions of the patent to each of its inventors. We then aggregate the inventor–year patent counts to the firm–year level.	USPTO
Innovation efficiency	The ratio of the number of granted patents filed by firm i in year t to R&D expenditure in firm i in year t. The ratio is set to 0 if R&D expenses is reported as negative or missing (Gao and Chou, 2015).	USPTO
Inventor pool	The total number of unique inventors who file at least one patent with firm i in year t.	USPTO
First-time inventor	For firm i in year t, *first-time inventor* is the average (across the firm's patents filed in year t) of the patent-level share of inventors in each patent's inventor team who have no prior USPTO patent filings before year t.	USPTO
AI-experienced inventor	For firm i in year t, *AI-experienced inventor* is the average (across the firm's patents filed in year t) of the patent-level share of inventors in each patent's inventor team who are AI-experienced. An inventor is classified as AI-experienced if she has filed at least one patent that is predicted by Giczy et al. (2022) to contain AI technology in any of the eight technology components (machine learning, natural language processing, computer vision, speech, knowledge processing, AI hardware, evolutionary computation, and planning and control) based on a 50% threshold up to year *t-1*.	USPTO, Artificial Intelligence Patent Dataset

(continued on next page)

Table A.1 (continued)

Variable	Description	Source
	E. Channel variables (continued)	
Specialization	For firm *i* in year *t*, *Specialization* is the average (across the firm's patents filed in year *t*) of the patent-level average *Specialization* scores of inventors in each patent's inventor team. An inventor's *Specialization* score in year *t* is the Herfindahl–Hirschman Index computed from the distribution of her granted patents across the Cooperative Patent Classification (CPC) technology classes at the 3-digit level, using her granted patents applied for up to year *t−1* (Li and Wang, 2023). A higher *Specialization* score implies that the inventor is more specialized in her technology classes.	USPTO
Team size	The average number of inventors per patent *p* filed by firm *i* in year *t*.	USPTO
Unique inventor team	The number of distinct inventor teams, made of unique combinations of inventors, that file at least one patent with firm *i* in year *t*.	USPTO
Core skill breadth	For firm *i* in year *t*, *Core skill breadth* is the average (across the firm's patents filed in year *t*) of the patent-level number of distinct core skills possessed by inventors in each patent's inventor team. An inventor's core skill in year *t* is defined as the first three digits of the Cooperative Patent Classification (CPC) technology class in which she has the largest number of granted patents, based on her patent applications filed up to year *t-1* (Li and Wang, 2023).	USPTO
New cite ratio	For firm *i* in year *t*, *New cite ratio* is the average of the patent-level new cite ratio across the firm's patents filed in year *t*. The new cite ratio for patent *p* filed by firm *i* in year *t* is defined as the share of backward citations made to new knowledge, that is, knowledge outside the inventor team's existing knowledge, relative to the total number of backward citations to U.S. patents, foreign patents, and other references (Fitzgerald and Liu, 2020). The inventor team's new knowledge consists of citations to U.S. patents, foreign patents, or other references that have neither been previously cited nor produced by any member of the inventor team prior to the filing date of patent *p*. The inventor team's existing knowledge comprises all U.S. patents, foreign patents, and other references that have been previously cited or produced by at least one team member before the filing date of patent *p*.	USPTO

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Table A.1 (continued)

Variable	Description	Source
	F. Heterogeneity variables	
Focused firm	An indicator variable that equals 1 if firm i has only one business segment and no additional segments in 2015, and 0 otherwise.	Compustat Segments
Big firm	An indicator variable that equals 1 if the total assets of firm i in 2015 are equal to or larger than the industry average value in 2015, and 0 otherwise. Industries follow the Fama-French 49 Industry Classification.	Compustat
Unconstrained firm	An indicator variable that equals 1 if the risk of firm i to delay investments due to issues with liquidity is less than or equal to the 2015 industry average, and 0 otherwise. The risk of delaying investments due to liquidity issues is a text-based measure of financial constraints developed by Hoberg and Maksimovic (2015). Industries follow the Fama-French 49 Industry Classification.	Hoberg and Maksimovic (2015)
High-tech firm	An indicator variable that equals 1 if firm i operates in high-tech sectors in 2015, and 0 otherwise. We classify high-tech sectors as 2-digit NAICS sectors 51 ("Information") and 54 ("Professional, Scientific, and Technical Services") following Babina et al. (2024).	Compustat

Table A.2
Industry Distribution

This table presents the industry distribution of all firms in the firm-year panel sample from 2011 to 2021. *All firms* refer to all unique firms in the sample. *AI-adopting firms* refer to firms that ever adopts AI technology from 2011 to 2021. *Never-adopting firm* refers to firms that never adopt AI technology from 2011 to 2021. Industries follow the Fama-French 49 Industry Classification.

	All firms	AI-adopting firms		Never-adopting firms	
Fama-French 49 Industry Classification	Count	Count	Percentage	Count	Percentage
Transportation	44	20	45.45	24	54.55
Entertainment	18	7	38.89	11	61.11
Retail	75	28	37.33	47	62.67
Utilities	51	18	35.29	33	64.71
Measuring and Control Equipment	63	22	34.92	41	65.08
Printing and Publishing	9	3	33.33	6	66.67
Aircraft	19	6	31.58	13	68.42
Wholesale	59	18	30.51	41	69.49
Business Services	95	28	29.47	67	70.53
Communication	62	18	29.03	44	70.97
Restaurants, Hotels, Motels	14	4	28.57	10	71.43
Tobacco Products	7	2	28.57	5	71.43
Computer Software	354	101	28.53	253	71.47
Healthcare	32	9	28.12	23	71.88
Consumer Goods	40	11	27.50	29	72.50
Personal Services	17	4	23.53	13	76.47
Computer Hardware	65	15	23.08	50	76.92
Petroleum and Natural Gas	79	18	22.78	61	77.22
Food Products	31	7	22.58	24	77.42
Recreation	23	5	21.74	18	78.26
Automobiles and Trucks	58	12	20.69	46	79.31
Apparel	20	4	20.00	16	80.00
Defense	10	2	20.00	8	80.00
Shipbuilding, Railroad Equipment	10	2	20.00	8	80.00
Candy and Soda	5	1	20.00	4	80.00
Electronic Equipment	218	42	19.27	176	80.73
Business Supplies	24	4	16.67	20	83.33
Beer and Liquor	6	1	16.67	5	83.33
Construction Materials	51	8	15.69	43	84.31
Steel Works Etc	26	4	15.38	22	84.62
Machinery	113	17	15.04	96	84.96
Rubber and Plastic Products	14	2	14.29	12	85.71
Chemicals	72	9	12.50	63	87.50
Medical Equipment	169	20	11.83	149	88.17
Almost Nothing	18	2	11.11	16	88.89
Shipping Containers	9	1	11.11	8	88.89
Electrical Equipment	55	6	10.91	49	89.09
Construction	27	2	7.41	25	92.59
Pharmaceutical Products	470	23	4.89	477	95.11
Agriculture	7	0	0.00	7	100.00
Coal	3	0	0.00	3	100.00
Fabricated Products	4	0	0.00	4	100.00
Non-Metallic and Industrial Metal Mining	10	0	0.00	10	100.00
Precious Metals	3	0	0.00	3	100.00
Textiles	4	0	0.00	4	100.00

Table A.3
AI and Innovation: Non-High-Tech Firms

This table presents the results of stacked DiD regressions on the effect of AI on corporate innovation for non-high-tech firms. We exclude firms that operate in high-tech sectors in 2015. We classify high-tech sectors as 2-digit NAICS sectors 51 ("Information") and 54 ("Professional, Scientific, and Technical Services") following Babina et al. (2024). We present Pseudo-Poisson Maximum Likelihood results in year $t + 3$ following AI adoption in year t. Column (1) presents the results on innovation quantity, column (2) presents the results on innovation quality, and columns (3)-(8) present results on innovation characteristics. *AI* is an indicator variable that equals 1 if firm i in cohort c uses AI technology in year t, and 0 otherwise. All specifications include Year-Cohort FE and Firm-Cohort FE. Standard errors are clustered at the firm level. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

Variable	*Patent count*	*Adjusted citation*	*Originality*	*Distance*	*Exploitative*	*Explorative*	*Generality*	*Claim*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI	0.138***	0.341***	0.219***	0.135***	0.190***	0.067	0.482***	0.170***
	(3.46)	(5.62)	(5.28)	(3.44)	(6.46)	(0.51)	(5.01)	(5.32)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.938	0.249	0.0808	0.115	0.103	0.147	0.111	0.299
Observations	31,679	27,109	29,771	28,344	25,535	25,215	19,381	31,666

Table A.4
The Heterogeneous Impact of AI on Innovation: Product Market Competition

This table presents results from stacked DiDiD regressions examining cross-sectional heterogeneity in the impacts of AI on corporate innovation across firms with different product market competition. We present Pseudo-Poisson Maximum Likelihood results in year $t + 3$ following AI adoption in year t. Column (1) presents the results on innovation quantity, column (2) presents the results on innovation quality, and columns (3)-(8) present results on innovation characteristics. *AI* is an indicator variable that equals 1 if firm i in cohort c uses AI technology in year t, and 0 otherwise. *High competition* is an indicator variable that equals 1 if firm i's TNIC-based Herfindahl-Hirschman Index (HHI) is equal to or below the 2015 average, and 0 otherwise. The TNIC-based HHI is a text-based measure of product market competition developed by Hoberg and Phillips (2016). All specifications include Year–Cohort and Firm–Cohort fixed effects, as well as subgroup-specific Year–Cohort and Firm–Cohort fixed effects obtained by interacting these fixed effects with the indicator variable for product market competition. Standard errors are clustered at the firm level. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

Variable	*Patent count*	*Adjusted citation*	*Originality*	*Distance*	*Exploitative*	*Explorative*	*Generality*	*Claim*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI	0.180***	0.358***	0.104*	0.083	0.077	0.212	0.518***	0.077*
	(3.21)	(4.19)	(1.78)	(1.54)	(1.62)	(1.12)	(3.44)	(1.87)
AI x *High competition*	-0.011	-0.071	0.155**	0.083	0.120**	-0.018	-0.128	0.095
	(-0.15)	(-0.63)	(2.02)	(1.14)	(2.04)	(-0.07)	(-0.70)	(1.57)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year-Cohort-High competition FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Firm-Cohort-High competition FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Pseudo R^2	0.938	0.251	0.0811	0.118	0.100	0.150	0.108	0.307
Observations	36,792	39,862	34,292	32,743	29,514	29,465	22,786	36,762

Table A.5
The Impact of AI on Firm-Level Productivity: Control Function Approach

This table presents the results of Control Function (CF) regressions on the effect of AI on productivity, labor and capital input using the stacked DiD sample. The first-stage result on *AI exposure* is reported in column (1). The second-stage results on corporate productivity, labor, and capital input measures in year $t + 3$ following AI adoption in year t are reported in columns (2)-(4), respectively. Firm-level productivity, labor and capital input are measured by *Value-added TFP*, *Labor*, *Capital*, respectively. We estimate the probit model in the first stage and use linear regressions in the second stage. *AI exposure* is firm i's exposure to AI-strong universities via the firm-university STEM worker hiring networks as of 2010 (Babina et al., 2024). *Top 10 exposure* is firm i's exposure to top 10 universities ranked by 2010 U.S. News & World Report via the firm-university STEM worker hiring networks as of 2010 (Babina et al., 2024). *Residual* is the regression residuals from the first-stage estimation. All regressions include Year-Cohort FE and Industry-Cohort FE. Industries follow the Fama-French 49 Industry Classification. Standard errors are computed using a firm-level cluster bootstrap with 1,000 replications. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

	1st Stage	2nd Stage		
Variable	*AI*	*Value-added TFP*	*Labor*	*Capital*
	(1)	(2)	(3)	(4)
AI exposure	0.572***			
	(4.68)			
AI		0.692**	0.836	1.337
		(2.00)	(0.97)	(1.31)
Top 10 exposure	-0.685	0.114	-1.299*	-1.141
	(-1.54)	(0.27)	(-1.86)	(-1.51)
Residual		-0.418**	0.589	0.397
		(-2.21)	(1.24)	(0.71)
Year-Cohort FE	Yes	Yes	Yes	Yes
Industry-Cohort FE	Yes	Yes	Yes	Yes
Wald χ^2	541.82			
Pseudo R^2	0.1149			
Adjusted R^2		0.0121	0.396	0.443
Observations	19,918	6,089	6,089	6,089

Table A.6
The Impact of AI on Economic Outcomes: Control Function Approach

This table presents the results of Control Function (CF) regressions on intangible versus tangible investment, operations, and market value using the stacked DiD sample. The first-stage result on *AI exposure* is reported in column (1). The second-stage results on economic outcome measures in year $t + 3$ following AI adoption in year t are reported in columns (2)-(8). Columns (2)-(3) presents results on intangible versus tangible investment, columns (4)-(6) presents results on operating costs, and columns (7)-(8) present results on market value. We estimate the probit model in the first stage and use linear regressions in the second stage. *AI exposure* is firm i's exposure to AI-strong universities via the firm-university STEM worker hiring networks as of 2010 (Babina et al., 2024). *Top 10 exposure* is firm i's exposure to top 10 universities ranked by 2010 U.S. News & World Report via the firm-university STEM worker hiring networks as of 2010 (Babina et al., 2024). *Residual* is the regression residuals from the first-stage estimation. All regressions include Year-Cohort FE and Industry-Cohort FE. Industries follow the Fama-French 49 Industry Classification. Standard errors are computed using a firm-level cluster bootstrap with 1,000 replications. t-statistics are in parentheses. ***, ** and * indicate statistical significance at the 1%, 5%, and 10% level, respectively. Variable definitions and sources are provided in Appendix Table A.1.

	1st Stage	2nd Stage						
Variable	*AI*	*R&D*	*Capital expenditure*	*Production cost*	*Non-production cost*	*Total operating cost*	*Tobin's q*	*Total q*
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
AI exposure	0.572***							
	(4.68)							
AI		0.137***	-0.012	1.281	0.149	1.002	3.997***	5.276**
		(4.10)	(-0.97)	(1.06)	(0.54)	(0.54)	(3.76)	(2.48)
Top 10 exposure	-0.685	0.173**	-0.014**	9.213	0.802	11.291	1.760	2.326
	(-1.54)	(2.54)	(-2.34)	(1.54)	(0.88)	(1.47)	(1.61)	(1.77)
Residual		-0.085***	0.006	-0.916	-0.167	-0.952	-2.137***	-2.881**
		(-4.55)	(0.97)	(-1.38)	(-1.12)	(-0.93)	(-3.62)	(-2.36)
Year-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Industry-Cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Wald χ^2	541.82							
Pseudo R^2	0.1149							
Adjusted R^2		0.432	0.190	0.102	0.104	0.0938	0.158	0.0911
Observations	19,918	6,284	6,283	6,284	5,503	6,284	6,265	6,239