# Asset Pricing with Supply Shocks*

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**Abstract**

Exploiting the 2016 SEC Tick Size Pilot as a randomized controlled trial, we find that a 1% exogenous and uninformed reduction in share repurchases lowers stock prices by 2.6%. Although treatment and control firms announced similar repurchase programs, treatment firms repurchased 22% fewer shares, a difference unexpected by both managers and investors. This reduction stems from an unanticipated conflict between the pilot and repurchase regulations. Because control stocks are unshocked, they allow us to control for price spillovers and identify a relative price multiplier. We further show that under a demand system, the relative price multiplier empirically approximates the own-price multiplier at the individual stock level. Finally, we find that investment advisors, mutual funds, and households absorb most of the supply shocks, followed by banks and other 13F institutions, while pension funds and insurance companies do not respond. The sizable uninformed price impact constrains managers' ability to repurchase undervalued equity and provides a new perspective on the long-run abnormal returns following share repurchases.

**Keywords:** Exogenous supply shock, price multiplier, share repurchase, Tick Size Pilot, demand-based asset pricing, portfolio rebalancing
**JEL Classification:** G11, G12, G20, G35

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# 1. Introduction

A traditional interpretation of long-run abnormal returns following repurchase announcements is market underreaction to the information conveyed by repurchase programs (Ikenberry, Lakonishok, and Vermaelen, 1995). Recently, Gabaix and Koijen (2023) conjecture that repurchases may mechanically raise stock prices beyond any informational channel due to market inelasticity. However, isolating this uninformed price impact of share repurchases is empirically challenging because managers tend to time the market, rendering repurchases endogenous (Baker and Wurgler, 2013). In this paper, we exploit a rare exogenous supply shock and document an economically significant price multiplier: a 1% uninformed reduction in share repurchases decreases stock prices by 2.6%.

The supply shock comes from the SEC's 2016 Tick Size Pilot (TSP). As a randomized controlled trial (RCT) in the equity market, the TSP randomly assigned 1,200 stocks to treatment groups, increasing the minimum tick size from one cent to five cents, while the remaining 1,199 stocks served as a control group. Surprisingly, this minor adjustment in tick size reduced share repurchases by 22% for treatment firms, creating exogenous positive supply shocks.

Neither managers nor market participants anticipated the reduction. We find no significant changes in the announced repurchase dollar volume between treatment and control firms, suggesting that the TSP did not alter managers' intent to buy back shares. Similarly, announcement cumulative abnormal returns (CAR) show no significant differences, indicating the supply shocks induced by the TSP were uninformed and unanticipated to the market. *Ex post,* Li, Ye, and Zheng (2024) show that the reduction in share repurchase stems from the conflict between the TSP and the SEC Rule 10b-18, a regulation on open-market repurchases.[1]

The randomized design of the TSP allows us to explicitly control for spillover effects. Without this control, the price change of a treatment stock reflects not only its own supply

[1]To prevent price manipulation, Rule 10b-18 effectively requires that firms use limit orders to repurchase at the bid price. The TSP removed four-fifths of price grids and thus lengthened queues at remaining prices. Time-priority execution prevented firms from competing with high frequency traders, mechanically reducing the fill rate of repurchases for treatment firms. See more details in Section 2.

shock but also the collective shocks from 1,199 other treatment stocks. Because control stocks faced no supply shocks, we match each treatment stock to a paired control stock with similar observables. Since the matched treatment and control share the same cross-price multipliers with all other stocks outside the matched pair (Haddad, He, Huebner, Kondor, and Loualiche, 2025),[2] the double-differences in prices removes spillovers from the remaining 1,199 treatment stocks.

Our regression specification then follows the theoretical framework of Haddad et al. (2025). The dependent variable is the within-pair double differences in log prices, capturing cumulative return differences between matched firms; the key explanatory variable is the corresponding within-pair double differences in cumulative shares repurchased. This specification identifies the relative price multiplier (Haddad et al., 2025) with an empirical magnitude of 2.575. This estimate implies that, relative to its matched substitute, a 1% uninformed reduction in share repurchases lowers the treatment stock's price by 2.575%.

Corporate managers, however, value the *own*-price multiplier, as it captures its own stock's price change when the firm exogenously changes its supply. We show that the estimated *relative* price multiplier closely approximates the *own*-price multiplier in our sample. Intuitively, the relative price multiplier retains only the one-to-one spillover from a treatment stock to the matched control, and we find that this bilateral spillover is negligible relative to the direct effect of the treatment stock's own supply shock.

To formalize this intuition, we build a simple demand system in which asset payoffs follow a single-factor structure (Koijen, Richmond, and Yogo, 2023). Theoretically, the own-price multiplier consists of an idiosyncratic-risk and a systematic-risk-exposure component. Because the relative multiplier captures only the idiosyncratic component, it approximates the own-price multiplier when idiosyncratic risk dominates systematic exposure. Empirically, we find strong supportive evidence: the average idiosyncratic risk is 8.4 times larger than average systematic exposure, and a $t$-test indicates that systematic exposure is less than 7% of idiosyncratic risk (at the 1% significance level). In other words, the residual spillover effect is less than 7% of the direct effect in our relative multiplier estimate for an *individual*

[2]This parallels the assumption of homogeneous cross-elasticities conditional on observables in Haddad et al. (2025). We empirically validate it under a factor structure of asset payoffs in Section 5.2.

stock.[3]

A unique feature of the TSP further controls for liquidity effects on treatment prices. The TSP divided the 1,200 treatment stocks into three groups. The 400 stocks in test group 3 exogenously repurchased 14% fewer shares relative to the pre-treatment mean than groups 1 and 2 because of an unexpected restriction on the use of dark pools for share repurchases.[4] Importantly, the restriction did not affect overall market liquidity for group 3 stocks (Albuquerque, Song, and Yao, 2020). We first estimate price multipliers separately for groups 1 and 2 and for group 3, finding no statistically significant differences from the baseline estimate. We then directly match each treatment firm in group 3 with a treatment firm in group 1 or 2. This additional layer of differencing rules out potential liquidity effects on treatment prices, as all three groups experienced the same decrease in liquidity. We re-estimate the price multiplier and again find no statistically significant differences from our baseline estimates. Taken together, our findings suggest that the three treatment groups have the same price multipliers, and group 3 stocks experienced a greater price decline than groups 1 and 2 primarily because of the additional reduction in share repurchases.

Finally, we examine how investors adjust their holdings in response to supply shocks. Intuitively, given the price change through the price multiplier, investors with higher demand elasticities absorb more supply shocks. We find that for each 1% additional exogenous supply, investment advisors (e.g., hedge funds) increase their holdings by 0.501%; mutual funds increase holdings by 0.249%; banks and other 13F institutions increase holdings by 0.102% and 0.104%, respectively. Insurance companies and pension funds do not actively respond to supply shocks. Though households do not change their holdings on average, we find that they actively absorb the supply shocks immediately after the TSP, but gradually decrease their holdings by selling shares to other investors, exaggerating the supply shocks alongside treatment firms.

**Contribution to Related Literature.** We contribute to the share repurchase liter-

[3]Nevertheless, this approximation does not imply that spillovers from all treatment firms are negligible. During the Tick Size Pilot, 1,200 treatment stocks experienced positively correlated supply shocks, so such aggregate spillovers can be substantial even when spillovers at the *individual*-stock level are small.

[4]Group 3 stocks were additionally subject to the "trade-at" rule, which required dark-pool executions to improve the NBBO by 2.5 cents, while Rule 10b-18 treated such price improvement for repurchase orders as an indicator of price manipulation. This unintended conflict effectively barred Group 3 firms from repurchasing shares through dark pools.

ature by showing that market inelasticity provides an additional explanation for why stock prices rise with share repurchases. A large body of corporate finance research attributes abnormal returns around repurchase announcement to managerial signaling of undervaluation (Dann, 1981; Vermaelen, 1981; Ofer and Thakor, 1987; Constantinides and Grundy, 1989; Billett and Yu, 2016). However, Ikenberry, Lakonishok, and Vermaelen (1995) argue that signaling theory alone cannot explain the post-announcement long-run abnormal returns;[5] instead, they interpret these returns as evidence of market underreaction. In contrast, our findings suggest a unified and consistent explanation: abnormal returns may arise mechanically from reductions in shares outstanding. If investors rationally anticipate such mechanical effects, these gains may be partially capitalized into announcement returns.[6] Furthermore, since a repurchase program typically takes years to complete (Stephens and Weisbach, 1998), the continued reduction in shares outstanding then contribute to the long-run abnormal returns after the announcement.

We also contribute to the behavioral corporate finance literature. An important strand of this literature examines how rational managers correct mispricings by issuing overvalued securities and repurchasing undervalued ones (Baker and Wurgler, 2013). To rationalize why managers do not repurchase entire undervalued firms, prior studies assume an exogenous price impact of repurchases (e.g., Stein, 1996). While Ye, Zhang, and Zou (2025) endogenize repurchase price impact through information asymmetry, we provide an alternative empirical foundation for this assumption by showing that market inelasticity alone generates economically significant price impact, even in the absence of asymmetric information.

Our paper complements a rich literature estimating asset price multipliers by providing the first estimate based on supply shocks. While economic principles suggest that identifying a demand curve requires exogenous supply shocks, reduced-form studies rely largely on demand shocks.[7]. Assuming a fixed supply, the demand-shock approach implicitly partitions

[5]The long-run abnormal returns are well documented over time and across different countries (e.g., Ikenberry, Lakonishok, and Vermaelen, 1995; Peyer and Vermaelen, 2009; Manconi, Peyer, and Vermaelen, 2019).

[6]Because repurchase program has various completion rates (Bonaimé, 2015), rational investors would not fully price in the announced volume at the announcement date.

[7]Typical demand shocks include index inclusions and deletions (Shleifer, 1986; Chang, Hong, and Liskovich, 2015; Pavlova and Sikorskaya, 2023), mutual fund flows (Coval and Stafford, 2007; Lou, 2012; Chaudhary, Fu, and Li, 2025), dividend reinvestment (Kvamvold and Lindset, 2018; Schmickler and Tremacoldi-Rossi, 2023; Chen, 2024; Hartzmark and Solomon, 2025), regulatory regime shifts (Cassella et al.,

investors into shocked and unshocked groups and identifies the residual price multiplier of unshocked investors.[8] In contrast, managers value the price multiplier across all investors. Under an exogenous supply shock, the shocked agent is the firm itself and unshocked agents are all investors; the supply shock therefore identifies the demand curve faced by the firm, yielding a price multiplier directly relevant to corporate decision-making.

On the structural side, Koijen and Yogo (2019) innovatively bypass the fixed-supply constraint by constructing and rotating hypothetical "shocks'' across investors via investment mandates. Their approach identifies *every* investor's demand curve and then aggregates them to form the stock-level demand curve. However, they impose downward-sloping demand curves on each investor to ensure a unique equilibrium. This restriction may top-code investors with negative demand elasticities (e.g., momentum traders), thereby attenuating the aggregate elasticity estimate (Binsbergen, David, and Opp, 2025). Our supply-shock approach complements theirs by directly identifying the demand curve faced by firms without parametric assumptions at the investor level.

The demand-shock approach does not study investors' equilibrium holdings because it relies on exogenous changes in the holdings of a subset of investors. In contrast, Koijen and Yogo (2019) explains cross-sectional investor holdings through a demand system. Koijen et al. (2021) study time-series portfolio rebalancing by bond market investors under quantitative easing by the European Central Bank. We provide the first equity market evidence of investors rebalancing their holdings in response to exogenous supply shifts, finding that investment advisors, mutual funds, and households absorb most of the supply shocks, followed by banks and other 13F institutions.

Our paper also contributes to the discussion of spillover effects in estimating asset price multiplier. Fuchs, Fukuda, and Neuhann (2025) theoretically demonstrate that spillover effects can bias the estimate; Chaudhary, Fu, and Li (2025) empirically show that ignoring spillovers leads to overestimation of individual bond price multipliers. Haddad et al. (2025) develop a theoretical framework for controlling spillover effects and identifying relative price

2025), and central bank asset purchase programs (Koijen et al., 2021; Haddad, Moreira, and Muir, 2025)

[8]Recent work also shows that firms may endogenously respond to demand shocks by adjusting equity supply, such as issuing shares following index inclusion (Sammon and Shim, 2024) or unexpected mutual fund inflows (Wang, 2025). As a result, estimates based on demand shocks may also confound investor demand elasticity with supply elasticity.

multipliers. We make two contributions to this line of research. First, our setting allows us to be the first to implement this framework empirically in the equity market: the randomized assignment of shocked and unshocked stocks under the TSP enables us to absorb spillover effects within a difference-in-differences design. Second, beyond estimating the relative price multiplier, we show both theoretically and empirically that when idiosyncratic risk dominates systematic risk exposure, the relative price multiplier closely approximates the own-price multiplier at the individual stock level. This result is particularly relevant for corporate managers, who cannot trade correlated stocks with similar characteristics and therefore care about own-price multipliers.

The rest of the paper is structured as follows. Section 2 describes the institutional background on the TSP. Section 3 establishes conceptual framework for our identification method. Section 4 outlines the sample construction and presents the effect of TSP on share supply, and the tests of exogeneity of the supply shock. Section 5 presents the estimation of relative price multiplier, shows the empirical approximation of relative price multiplier to the own-price multiplier, controls for the confounding liquidity effect, and investigates robustness checks. Section 6 reports the portfolio rebalancing behaviors by investors when faced supply changes. Finally, Section 7 concludes.

## 2. Tick Size Pilot Design and Institutional Background

In this section, we elaborate the institutional background of the unique RCT in U.S. equity market, the TSP. We further explain the mechanism under which the TSP induced the exogenous supply shocks.

In 2014, the SEC instructed the Financial Industry Regulatory Authority (FINRA) and National Securities Exchanges (NSEs) to develop a pilot program to test the effects of varying tick sizes on market performance. The TSP officially commenced on October 3, 2016, and concluded on October 1, 2018. The program selected 2,399 stocks from the Regulation National Market System (Reg NMS) that met specific criteria in the three months before the launch: a minimum share price of $1.50, an average daily trading volume of fewer than one million shares, and a market capitalization below $3 billion.

After selecting the eligible TSP stocks, FINRA and NSEs randomly assigned treatment and control groups. They first sorted all eligible stocks into 27 buckets based on three characteristics: (1) low, medium, or high share prices, (2) low, medium, or high volume, and (3) low, medium, or high market capitalization. They then randomly selected stocks from each of the 27 categories to form three test groups, each containing 400 stocks. The remaining 1,199 eligible stocks served as the control group.

The control group continued to trade and quote at the standard 1-cent tick size. Test group 1 stocks could only be quoted in $0.05 increments, but they could still be traded in $0.01 increments. Test group 2 stocks were both quoted and traded in $0.05 increments. Test group 3 followed the same rules as test group 2 but also fell under the Trade-at Rule,[9] which barred dark-pool executions unless they improved the NBBO by 2.5 cents.

Since repurchasing firms typically assign valuations well above the prevailing market price (Li, Ye, and Zheng, 2023), a modest increase in tick size should have limited direct impact on the intent of share repurchases. Surprisingly, Li, Ye, and Zheng (2024) find that the TSP significantly reduced share repurchases for treatment firms. This unexpected outcome stems primarily from an unintended conflict between the TSP and an existing regulation: SEC Rule 10b-18, which regulates open-market share repurchase. Rule 10b-18 imposes a price ceiling on share repurchase to prevent firms from manipulating their stock prices. The price ceiling prohibits firms from bidding above the highest independent bid or the last transaction price, effectively encouraging firms to place buy limit orders at the bid price on repurchase transactions. The TSP removes four out of every five quote levels, forcing investors who previously quoted at finer intervals to cluster their bids at coarser levels. Because executions of limit orders quoted at the same price follow the time-priority rule, faster HFTs may outpace firms, reducing firms' order execution probability and thus lowering their realized repurchase volume *ex post*. As we report later, the effect of TSP on repurchase was unanticipated by firms or market participants, making the resulting decline in repurchases an exogenous supply shock to treatment firms.

Notably, firms often buy back shares using dark pools for two reasons. First, dark

[9]More details of the TSP are in https://www.govinfo.gov/content/pkg/FR-2015-05-13/pdf/2015-11425.pdf and https://www.sec.gov/files/rules/sro/nms/2015/34-74892-exa.pdf

pools match orders passively based on reference prices from public exchanges, which aligns naturally with the requirements of Rule 10b-18. For instance, when a dark pool uses the bid price to match buy and sell orders, it inherently satisfies the price constraint imposed by Rule 10b-18. Second, dark pools allow firms to bypass speed competition by not imposing the time priority in trade execution. However, the TSP introduced an additional Trade-at Rule specifically for test group 3, which restricted dark pool executions unless they improved the NBBO by more than 2.5 cents. Since Rule 10b-18 regards outbidding the best bid by 2.5 cents as a signal of price manipulation, it unintentionally blocked firms in test group 3 from repurchasing shares through dark pools. Consequently, firms in test group 3 faced more stringent constraints on share repurchases than test groups 1 and 2, and thus experienced larger decline in share repurchases.

# 3. Conceptual Framework

To motivate our empirical specification, we first develop a simple demand system with supply shocks. It builds on the static CARA-Normal framework with heterogeneous beliefs following Lintner (1969).[10] This model illustrates the identification of relative price multiplier, the control for spillover effects using the TSP's randomized control, the approximation of relative price multiplier to own-price multiplier, and the characterization of investors' portfolio rebalancing behaviors under supply shocks.

## 3.1 Model

We consider a two-period securities market with trading at $t = 0$ and payoffs realized at $t = 1$. The market consists of one riskless asset with a risk-free rate normalized to zero and $N$ risky assets indexed by $n = 1, \ldots, N$. Investors, indexed by $i = 1, \ldots, I$, allocate available funds across these assets. Each stock $n$ has an exogenously share supply $S(n)$ and delivers a per-share terminal payoff $d(n)$ at $t = 1$. Each investor $i$ holds heterogeneous beliefs over the

[10] We choose CARA-Normal framework because it captures the relationship between price and quantity in *levels*. Literature studying other changes in quantity level also applies this framework, such as Pavlova and Sikorskaya (2023) and Cassella et al. (2025).

payoff vector $\boldsymbol{d}_i = (d_i(1), \ldots, d_i(N))'$, following a joint-normal distribution with finite first and second moments, i.e., $\boldsymbol{d}_i \sim \mathcal{N}(\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i)$.[11]

Let $A_{i,0}$ denote investor $i$'s initial wealth. Let $\boldsymbol{P}$ and $\boldsymbol{q}_i$ be the endogenously determined, $N$-dimensional vectors of asset prices and the number of shares held by investor $i$ at $t = 0$, respectively. Investor $i$'s terminal wealth at $t = 1$ is

$$A_{i,1} = A_{i,0} + (\boldsymbol{d}_i - \boldsymbol{P})' \boldsymbol{q_i}$$

At $t = 0$, each investor $i$ chooses portfolio holdings $\boldsymbol{q}_i$ to maximize her expected constant absolute risk aversion (CARA) utility over the terminal wealth:

$$\max_{\boldsymbol{q}_i} \mathbb{E}_i \left[ -\exp\left(-\gamma_i A_{i,1}\right) \right]$$

where $\gamma_i > 0$ is investor $i$'s coefficient of absolute risk aversion. As shown in Appendix A.1, investor $i$'s optimal demand is

$$\boldsymbol{q}_i = \frac{1}{\gamma_i} \boldsymbol{\Sigma}_i^{-1} \left( \boldsymbol{\mu_i} - \boldsymbol{P} \right) \tag{1}$$

Let $\boldsymbol{S}$ be an $N$-dimensional vector of share supply at $t = 0$. By clearing markets for risky assets, $\sum_i^I \boldsymbol{q}_i = \boldsymbol{S}$, we obtain equilibrium asset prices:

$$\boldsymbol{P} = \boldsymbol{\mathcal{M}} \left( \sum_{i=1}^{I} \frac{1}{\gamma_i} \boldsymbol{\Sigma}_i^{-1} \boldsymbol{\mu}_i - \boldsymbol{S} \right), \text{ where } \boldsymbol{\mathcal{M}} \triangleq \left( \sum_{i=1}^{I} \frac{1}{\gamma_i} \boldsymbol{\Sigma}_i^{-1} \right)^{-1} \tag{2}$$

$\boldsymbol{\mathcal{M}} \in \mathbb{R}^{N \times N}$ is the price multiplier matrix, characterizing the equilibrium relationship between shifts in aggregate demand and asset prices. The diagonal elements $[\boldsymbol{\mathcal{M}}]_{n,n}$ represent the own-price multipliers, measuring how asset $n$'s price responds to a shift in its own aggregate demand, holding other prices fixed. The off-diagonal elements $[\boldsymbol{\mathcal{M}}]_{n,m}$ represent the cross-price multipliers, measuring how asset $n$'s price responds to a shift in aggregate

[11] The variance-covariance matrix, $\boldsymbol{\Sigma}_i \in \mathbb{R}^{N \times N}$, has full rank. In the CAPM, all investors hold homogeneous beliefs with constant $\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$.

demand for asset $m$. Formally,

$$[\boldsymbol{\mathcal{M}}]_{n,n} = -\frac{\partial P(n)}{\partial \sum_i^I q_i(n)} \quad \text{and} \quad [\boldsymbol{\mathcal{M}}]_{n,m} = -\frac{\partial P(n)}{\partial \sum_i^I q_i(m)}, \ \forall m \neq n$$

## 3.2 Identification with Supply Shocks

We now introduce supply shocks to our framework. Suppose the exogenous supply shocks shift the supply curve by a wedge of $\Delta \boldsymbol{S} = (\Delta S(1), ..., \Delta S(N))'$.[12] The total supply therefore changes to

$$\boldsymbol{S}^{\text{new}} = \boldsymbol{S} - \Delta \boldsymbol{S}$$

Thus, $\Delta S(n) > 0$ reflects a reduction in shares outstanding for stock $n$; $\Delta S(n) < 0$ indicates an increase in shares outstanding; and $\Delta S(n) = 0$ implies no supply shock. Equation (2) with $\boldsymbol{S}^{\text{new}}$ determines the new equilibrium price $\boldsymbol{P}^{\text{new}}$. The equilibrium price change $\Delta \boldsymbol{P} = \boldsymbol{P}^{\text{new}} - \boldsymbol{P}$ satisfies

$$\Delta \boldsymbol{P} = \boldsymbol{\mathcal{M}} \ \Delta \boldsymbol{S} \tag{3}$$

This transformation requires the assumption of uninformedness: supply shocks contain no information about assets' future payoffs (i.e., $\Delta \boldsymbol{\mu}_i = 0$ for all investors $i$). We later validate this assumption in our sample by showing no differences in CAR around the repurchase announcement during the TSP in Section 4. Appendix A.4 discusses the relaxation of this assumption and corresponding solutions.

### 3.2.1 Control for Spillover Effects

This section demonstrates how the randomized design of the TSP controls for spillover effects. Based on equation (3), the equilibrium price change for asset $n$ decomposes as direct and

[12] Here, "exogenous" means that $\Delta S(n)$ is independent from share price $P(n)$.

spillover effects:

$$\Delta P(n) = \underbrace{[\boldsymbol{\mathcal{M}}]_{n,n}\Delta S(n)}_{\text{direct effect}} + \underbrace{\sum_{m\neq n}[\boldsymbol{\mathcal{M}}]_{n,m}\Delta S(m)}_{\substack{\text{spillover effects from} \\ \text{other supply shocks}}} \tag{4}$$

The direct effect $[\boldsymbol{\mathcal{M}}]_{n,n}\Delta S(n)$ captures asset $n$'s price response to its own supply shock. Spillover effects characterize how a supply shocks to asset $m$ affects assent $n$'s price through the cross-price multiplier $[\boldsymbol{\mathcal{M}}]_{n,m}$.

Consider the TSP that induced simultaneous supply shocks to all 1,200 treatment firms. Estimating an unbiased own-price multiplier $[\boldsymbol{\mathcal{M}}]_{n,n}$ requires controlling for spillover effects from all other 1,199 treatment stocks.[13] Otherwise, price change induced by a treatment stock's own supply shock would be overestimated and partially reflect shocks to other treatment firms. We exploit the randomized assignment of treatment and control stocks in the TSP to difference out these spillover effects.

For each TSP treatment stock $n$, we match an unshocked TSP control stock $n'$ (i.e., $\Delta S(n') = 0$) with similar characteristics. From equation (4), we further express the equilibrium price change in the treatment stock $n$ as

$$\Delta P(n) = \underbrace{[\boldsymbol{\mathcal{M}}]_{n,n}\Delta S(n)}_{\text{direct effect}} + \underbrace{[\boldsymbol{\mathcal{M}}]_{n,n'}\Delta S(n')}_{\substack{\text{spillover effect from} \\ \text{the matched stock}}} + \underbrace{\sum\nolimits_{m\neq n,n'}[\boldsymbol{\mathcal{M}}]_{n,m}\Delta S(m)}_{\substack{\text{spillover effects from} \\ \text{other stocks}}} \tag{5}$$

where spillover effects are further decomposed into the effect from matched stock and the effect from all other stocks. Similarly, the price change in the matched control stock $n'$ is

$$\Delta P(n') = \underbrace{[\boldsymbol{\mathcal{M}}]_{n',n'}\Delta S(n')}_{\text{direct effect}} + \underbrace{[\boldsymbol{\mathcal{M}}]_{n',n}\Delta S(n)}_{\substack{\text{spillover effect from} \\ \text{the matched stock}}} + \underbrace{\sum\nolimits_{m\neq n,n'}[\boldsymbol{\mathcal{M}}]_{n',m}\Delta S(m)}_{\substack{\text{spillover effects from} \\ \text{other stocks}}} \tag{6}$$

Because stock $n$ and $n'$ have similar observables, we assume, following Haddad et al. (2025), that they share homogeneous cross-price multipliers with all other assets, i.e., $[\boldsymbol{\mathcal{M}}]_{n,m} =$

[13] Identification using demand-side instruments face the same challenge. Koijen and Yogo (2019) address it by imposing homogeneous substitution patterns implied by logit demand. Koijen and Yogo (2025) specify spillovers under a nested-logit structure. Chaudhary, Fu, and Li (2025) specify larger spillover effects for bonds within the same rating category as close substitutes, and weaker spillovers for bonds in different ratings.

$[\boldsymbol{\mathcal{M}}]_{n',m}$, $\forall m \neq n, n'$.[14] Under this assumption, differencing equations (5) and (6) yields

$$\Delta P(n) - \Delta P(n') = \left([\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n',n}\right) \Delta S(n) - \left([\boldsymbol{\mathcal{M}}]_{n',n'} - [\boldsymbol{\mathcal{M}}]_{n,n'}\right) \Delta S(n') \tag{7}$$

The double-difference in prices between the TSP treatment stock $n$ and its matched TSP control stock $n'$ eliminates spillover effects from all other 1,199 treatment stocks outside the pair. Because the matched TSP control stock $n'$ was unshocked ($\Delta S(n') = 0$), equation (7) further simplifies to

$$\Delta P(n) = \underbrace{\left([\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n',n}\right)}_{\substack{\text{relative price multiplier} \\ \mathcal{M}_{relative}(n,n')}} \Delta S(n) \tag{8}$$

The remaining bilateral spillover operates through the cross-price multiplier $[\boldsymbol{\mathcal{M}}]_{n',n}$. Therefore, the TSP supply shocks identify the *relative price multiplier* $\mathcal{M}_{relative}(n, n')$ —the difference between the own- and cross-price multipliers (Haddad et al., 2025). It measures how the price of a treatment stock changes relative to its comparable control stock when the relative supply (aggregate demand) shifts. $\mathcal{M}_{relative}(n, n')$ can be heterogeneous across matched pairs. Equation (8) motivates the following DID empirical specification:

$$\Delta p = \mathcal{M} \Delta s + \boldsymbol{\gamma}' \Delta \boldsymbol{x} + \varepsilon \tag{9}$$

where $\mathcal{M}$ identifies the relative price multiplier across treatment stocks. The variable $\Delta p = \Delta P(n) - \Delta P(n')$ and $\Delta s = \Delta S(n) - \Delta S(n')$ denote the double-difference in log stock price and repurchase payouts between a TSP treatment stock $n$ and its matched TSP control stock $n'$, respectively. We detail the variable construction in Section 5.1. As discussed in Appendix A.4, we further include differences in observables, such as liquidity, size, growth opportunities, profitability, dividend payout, and market beta to absorb confounding price effects.

[14] Section 3.3 provides a microfoundation for this assumption.

### 3.3 Approximation to Own-Price Multiplier

The relative price multiplier measures the price of stock relative to its matched substitute changes if the relative supply for these stocks shifts (Haddad et al., 2025), which is relevant for portfolio managers. In contrast, firm managers value the own-price multiplier, which measures a stock's price response to its own supply change. In this section, we derive empirical conditions under which the estimated relative price multiplier in equation (9) approximates the own-price multiplier. This approximation builds on the standard single-factor structure on asset payoffs as in Koijen, Richmond, and Yogo (2023).

**A Single-Factor Structure on Asset Payoffs.** Following Koijen, Richmond, and Yogo (2023), we assume each investor $i$'s belief on asset payoffs follows a single-factor structure:

$$\boldsymbol{d}_i = \boldsymbol{\mu}_i + \boldsymbol{b}_i F + \boldsymbol{\eta} \tag{10}$$

where $\boldsymbol{\mu}_i$ is an $N \times 1$ vector of investor $i$'s expected payoffs. $\boldsymbol{b}_i$ is an $N \times 1$ vector of investor-specific factor loadings, with its element $\boldsymbol{b}_i(n)$ capturing investor $i$'s perceived exposure of stock $n$ to the systematic risk factor $F$. The risk factor follows a standard normal distribution: $F \sim \mathcal{N}(0,1)$. The idiosyncratic shock $\boldsymbol{\eta} \sim \mathcal{N}(0, \sigma^2 \boldsymbol{I})$ is independent of $F$.[15]

Under assumption (10), the covariance matrix of investor $i$'s perceived payoffs is $\boldsymbol{\Sigma}_i = \boldsymbol{b}_i \boldsymbol{b}_i' + \sigma^2 \boldsymbol{I}$. This structure delivers a closed-form expression for the price multiplier matrix $\boldsymbol{\mathcal{M}}$ defined in equation (2), as stated in the following lemma.

**Lemma 1 (Price Multiplier under Factor Structure).** *Under the factor structure (10), the price multiplier matrix is*

$$\boldsymbol{\mathcal{M}} = \frac{\sigma^2}{\tau} \boldsymbol{I} + \frac{1}{\tau^2} \boldsymbol{\rho} \boldsymbol{L} \boldsymbol{\rho}' \tag{11}$$

*where*

$$\tau = \sum_{i=1}^{I} \frac{1}{\gamma_i}, \quad \boldsymbol{\rho} = \left( \frac{\boldsymbol{b}_1}{\sqrt{\gamma_1 (1 + \boldsymbol{b}_1' \boldsymbol{b}_1 / \sigma^2)}}, \cdots, \frac{\boldsymbol{b}_I}{\sqrt{\gamma_I (1 + \boldsymbol{b}_I' \boldsymbol{b}_I / \sigma^2)}} \right)_{N \times I}, \quad \boldsymbol{L} = \left( \boldsymbol{I} - \frac{1}{\tau \sigma^2} \boldsymbol{\rho}' \boldsymbol{\rho} \right)^{-1}$$

[15] Assuming homoskedastic idiosyncratic variance $\sigma^2$ simplifies exposition, though allowing $\sigma^2(n)$ heterogeneity leaves the core results unchanged.

*The corresponding own-price multiplier* $[\boldsymbol{\mathcal{M}}]_{n,n}$*, cross-price multiplier* $[\boldsymbol{\mathcal{M}}]_{n,m}$*, and the relative price multiplier* $\mathcal{M}_{relative}(n,m) = [\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n,m}$ *are*

$$
\begin{aligned}
[\boldsymbol{\mathcal{M}}]_{n,n} &= \frac{\sigma^2}{\tau} + \frac{1}{\tau^2}\,\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_n, \quad [\boldsymbol{\mathcal{M}}]_{n,m} = \frac{1}{\tau^2}\,\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_m \\
\mathcal{M}_{relative}(n,m) &= \frac{\sigma^2}{\tau} + \frac{1}{\tau^2}\,\boldsymbol{\rho}_n' \boldsymbol{L} (\boldsymbol{\rho}_n - \boldsymbol{\rho}_m) \qquad (12) \\
\text{where } \boldsymbol{\rho}_n &= \left( \frac{\boldsymbol{b}_1(n)}{\sqrt{\gamma_1(1 + \boldsymbol{b}_1'\boldsymbol{b}_1/\sigma^2)}}, \cdots, \frac{\boldsymbol{b}_I(n)}{\sqrt{\gamma_I(1 + \boldsymbol{b}_I'\boldsymbol{b}_I/\sigma^2)}} \right)'_{N \times 1}
\end{aligned}
$$

*Proof:* See Appendix A.2.

Intuitively, the own-price multiplier $[\boldsymbol{\mathcal{M}}]_{n,n}$ consists of an idiosyncratic component ($\sigma^2$), scaled by investors' aggregate risk-bearing capacities $\tau$, and a systematic component that aggregates across investors through $\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_n$. In contrast, the cross-price multiplier $[\boldsymbol{\mathcal{M}}]_{n,m}$ arises from systematic comovement between assets through $\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_m$. Hence, the relative price multiplier $\mathcal{M}_{relative}$ isolates the idiosyncratic component plus the difference in systematic exposure between the two assets. We later illustrate this intuition under the CAPM case.

**Constant Cross-Price Multipliers Outside the Pairs.** Our identification strategy follows Haddad et al. (2025) by assuming constant cross-price multipliers outside each matched pair $(n, n')$. Under the factor structure in Lemma 1, this condition becomes transparent. Since matched firms share similar characteristics, we assume they share similar factor exposures ($\boldsymbol{\rho}_n = \boldsymbol{\rho}_{n'}$).[16] Then, we have $[\boldsymbol{\mathcal{M}}]_{n,m} = [\boldsymbol{\mathcal{M}}]_{n',m}$ for all $m \neq n, n'$. Taking differences in price changes between $n$ and $n'$, therefore, cancels the spillover effects from all other shocked stocks as we show in equation (7).

For a given matched pair with $\boldsymbol{\rho}_n = \boldsymbol{\rho}_{n'}$, the relative price multiplier simplifies to $[\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n,n'} = \frac{\sigma^2}{\tau}$. Recalling that $[\boldsymbol{\mathcal{M}}]_{n,n} = \frac{\sigma^2}{\tau} + \frac{1}{\tau^2}\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_n$, the relative price multiplier closely approximates the own-price multiplier if the second term ($\frac{1}{\tau^2}\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_n$) is negligible. This occurs when the contribution of systematic risk exposure is sufficiently small relative to the idiosyncratic risk. The following proposition formalizes this intuition.

[16] One of the microfoundations of this assumption is to model the factor exposure as a linear function of characteristics, as in Koijen and Yogo (2019) and Koijen, Richmond, and Yogo (2023). Thus, similar characteristics between stocks imply similar factor exposures.

**Proposition 1 (Relative Price Multiplier Approximates Own-Price Multiplier).** *Under the single-factor payoff structure in equation 10, consider a matched pair $(n, n')$ such that $b_i(n) = b_i(n')$ for all investors $i$. Then:*

(i) ***(Identical cross-price multiplier)*** $[\boldsymbol{\mathcal{M}}]_{n,m} = [\boldsymbol{\mathcal{M}}]_{n',m}$ *for all* $m \neq n, n'$.

(ii) ***(Approximation to own-price multiplier)*** *Furthermore, if* $\dfrac{b_i^2(n)}{\sigma^2} \to 0$ *for all investor* $i$, $\dfrac{\mathcal{M}_{relative}(n, n')}{[\boldsymbol{\mathcal{M}}]_{n,n}} \to 1$.

*Proof:* See Appendix A.3.

Proposition 1 shows that when each investor's perceived systematic risk exposure for a stock $n$, $\boldsymbol{b}_i^2(n)$, is sufficiently small relative to its idiosyncratic risk $\sigma^2$ ($\boldsymbol{b}_i^2(n) \ll \sigma^2$), the systematic-driven term in the own-price multiplier is negligible. Consequently, the relative price multiplier $\mathcal{M}_{relative}(n, n')$ of the matched stocks $(n, n')$ provides a close approximation to stock $n$'s own-price multiplier.

**CAPM Illustration.** While formal proofs of Lemma 1 and Proposition 1 are respectively in Appendix A.2 and A.3, the intuition is clear under a representative-investor CAPM benchmark with $\gamma_i = \gamma_j = \gamma$, $\boldsymbol{b}_i = \boldsymbol{b}_j = \boldsymbol{b}$, and $I = 1$. In this case, the multiplier matrix simplifies to

$$\begin{aligned} [\boldsymbol{\mathcal{M}}]_{n,n} &= \gamma \left[\sigma^2 + \boldsymbol{b}^2(n)\right] \\ [\boldsymbol{\mathcal{M}}]_{n,m} &= \gamma\ \boldsymbol{b}(n)\boldsymbol{b}(m) \\ \mathcal{M}_{relative}(n, m) &= \gamma \left\{\sigma^2 + \ \boldsymbol{b}(n)\left[\boldsymbol{b}(n) - \boldsymbol{b}(m)\right]\right\} \end{aligned} \tag{13}$$

*Proof:* See Appendix A.2.1.

The own-price multiplier combines the idiosyncratic risk, $\sigma^2$, and the systematic risk exposure through the asset's beta, $\boldsymbol{b}^2(n)$. Cross-price multipliers scale with the product of two assets' betas, $\boldsymbol{b}(n)\boldsymbol{b}(m)$. When two assets have the same betas, i.e., $\boldsymbol{b}(n) = \boldsymbol{b}(m)$, they exhibit the same cross-price multipliers to other stocks. Furthermore, when $\boldsymbol{b}^2(n) \ll \sigma^2$, i.e., systematic risk exposure is small relative to idiosyncratic risk, the relative price multiplier approximates the own price multiplier.

**The Necessity of Eliminating Spillover Effects.** Proposition 1 does not imply that the spillover effects aggregating across all shocked assets in equation (4) are negligible. The

condition $\boldsymbol{b}_i^2(n) \ll \sigma^2$ ensures that systematic risk exposure is relatively small in individual asset. However, the spillover from all shocked stocks, $\sum_{m=1}^{N} \boldsymbol{b}_i(n)\boldsymbol{b}_i(m)\Delta S(m)$, can still be large when (1) the number of shocked assets $N$ is large, and (2) the supply shocks $\Delta S(m)$ are positively correlated. In the TSP, 1,200 treatment stocks experienced positively correlated supply shocks. Ignoring these spillovers would overstate the attribution of price changes to own shocks, thereby biasing estimates of the price multiplier.

### 3.4 Investor Response to Supply Shocks

This section studies the changes in equilibrium holdings across investor types in response to supply shocks. Since the demand system of Koijen and Yogo (2019) explains the cross-sectional stock holdings under fixed supply, a natural question is how investors adjust their portfolios when share supply changes exogenously. Starting from investor $i$'s demand function (1) and substituting the equilibrium price (2), we have:

$$\boldsymbol{q_i} = \frac{1}{\gamma_i}\boldsymbol{\Sigma}_i^{-1}\left(\boldsymbol{\mu_i} - \boldsymbol{\mathcal{M}}\left(\sum_{i=1}^{I}\frac{1}{\boldsymbol{\gamma}_i}\boldsymbol{\mu}_i - \boldsymbol{S}\right)\right)$$

The change in investor $i$'s holdings, $\Delta\boldsymbol{q}_i$, in response to a supply shock $\Delta\boldsymbol{S}$, is thus

$$\Delta\boldsymbol{q}_i = \boldsymbol{\mathcal{A}}_i \times \Delta\boldsymbol{S}, \quad \text{where } \boldsymbol{\mathcal{A}}_i \triangleq \underbrace{\frac{1}{\gamma_i}\boldsymbol{\Sigma}_i^{-1}}_{\text{investor } i\text{'s elasticity}} \times \underbrace{\boldsymbol{\mathcal{M}}}_{\text{aggregate price multiplier}} \tag{14}$$

We refer to $\boldsymbol{\mathcal{A}}_i \in \mathbb{R}^{N\times N}$ as investor $i$'s absorption matrix, defined as the product of investor $i$'s demand elasticity matrix and the aggregate price multiplier matrix. Intuitively, $\boldsymbol{\mathcal{A}}_i$ describes how investor $i$ adjusts her portfolio given a supply shock: a supply shock alters equilibrium prices through $\boldsymbol{\mathcal{M}}$, and each investor rebalances according to her own elasticity matrix $\frac{1}{\gamma_i}\boldsymbol{\Sigma}_i^{-1}$. Investors with higher demand elasticity absorb a larger share of the shock.

In Koijen and Yogo (2019), the assumption of downward-sloping individual demand curves implies that an increase in shares outstanding, and the resulting decrease in price, leads every investor to increase their holdings. In contrast, equation (14) does not impose

this assumption and therefore allows some investors to decrease their holdings when share supply increases.

Because an investor chooses a portfolio of assets, a supply shock to asset $m$ changes prices of all other assets through the price multiplier matrix $\boldsymbol{\mathcal{M}}$, leading to the investor's rebalancing across multiple assets. Thus, investor $i$'s absorption matrix includes both own-asset rebalancing terms $[\boldsymbol{\mathcal{A}}_i]_{n,n}$ and cross-asset rebalancing terms $[\boldsymbol{\mathcal{A}}_i]_{n,m}$ for $n \neq m$. In particular, for a matched stock pair $(n, n')$, investor $i$'s holding changes are

$$\Delta \boldsymbol{q}_i(n) = [\boldsymbol{\mathcal{A}}_i]_{n,n}\, \Delta S(n) + [\boldsymbol{\mathcal{A}}_i]_{n,n'}\, \Delta S(n') + \sum\nolimits_{m \neq n,n'} [\boldsymbol{\mathcal{A}}_{\boldsymbol{i}}]_{n,m}\, \Delta S(m)$$
$$\Delta \boldsymbol{q}_i(n') = [\boldsymbol{\mathcal{A}}_{\boldsymbol{i}}]_{n',n'}\, \Delta S(n') + [\boldsymbol{\mathcal{A}}_i]_{n',n}\, \Delta S(n) + \sum\nolimits_{m \neq n,n'} [\boldsymbol{\mathcal{A}}_i]_{n',m}\, \Delta S(m)$$

Similar to the spillover effects discussed in Section 3.2, the observed holding changes reflect both the own-asset rebalancing effect $[\boldsymbol{\mathcal{A}}_i]_{n,n}\Delta S(n)$ and cross-asset substitution effects $[\boldsymbol{\mathcal{A}}_i]_{n,m}\Delta \boldsymbol{S}(m), \forall m \neq n$, that is, a supply shock to asset $m$ changes its price, leading the position adjustment in asset $n$. To eliminate cross-asset substitution effects, we rely on the TSP's randomized-controlled design, where control stocks face no supply shocks ($\Delta S(n') = 0$). Taking the difference in investor $i$'s holding changes between matched treatment and control stocks yields[17]

$$\Delta \boldsymbol{q}_i(n) - \Delta \boldsymbol{q}_i(n') = \left( [\boldsymbol{\mathcal{A}}_i]_{n,n} - [\boldsymbol{\mathcal{A}}_i]_{n',n} \right) \Delta \boldsymbol{S}(n) \tag{15}$$

Equation (15) thus characterizes how investor $i$ rebalances her holdings in stock $n$ relative to its matched stock $n'$ when stock $n$ experiences an exogenous supply shock. Similar to the relative price multiplier, the coefficient $\left( [\boldsymbol{\mathcal{A}}_i]_{n,n} - [\boldsymbol{\mathcal{A}}_i]_{n',n} \right)$ captures the relative adjustment between treatment stock $n$ and its matched control stock $n'$. We provide corresponding empirical evidence in Section 6.

[17] This simplification also relies the factor structure in Section 3.3. Assuming that matched stocks share similar factor exposures, we have $\boldsymbol{\Sigma}_i^{-1}\boldsymbol{\mathcal{M}} = \frac{1}{\gamma_i \sigma^2}(\boldsymbol{\mathcal{M}} - \frac{\boldsymbol{b}_i \boldsymbol{b}_i' \boldsymbol{\mathcal{M}}}{\sigma^2 + \boldsymbol{b}_i' \boldsymbol{b}_i})$. Defining $\boldsymbol{v}_i = \boldsymbol{b}_i' \boldsymbol{\mathcal{M}}$, this implies $[\boldsymbol{\Sigma}_i^{-1}\boldsymbol{\mathcal{M}}]_{n,m} = \frac{1}{\gamma_i \sigma^2}\left([\boldsymbol{\mathcal{M}}]_{n,m} - \frac{\boldsymbol{b}_i(n)\boldsymbol{v}_i'(m)}{\sigma^2 + \boldsymbol{b}_i' \boldsymbol{b}_i}\right)$. For a matched pair $(n, n')$ with $\boldsymbol{b}_i(n) = \boldsymbol{b}_i(n')$, we have $[\boldsymbol{\Sigma}_i^{-1}\boldsymbol{\mathcal{M}}]_{n,m} = [\boldsymbol{\Sigma}_i^{-1}\boldsymbol{\mathcal{M}}]_{n',m}$ for $m \neq n, n'$.

# 4. Exogenous Supply Shock

This section begins by outlining the research specification, data sources, and sample construction. We show that the TSP significantly reduced share repurchases among treatment firms relative to control firms. We then show that this reduction is exogenous, as neither corporate managers nor market participants anticipated the TSP's impact on repurchase activity.

## 4.1 Specification, Data, and Sample Selection

We begin by estimating the effect of TSP on share repurchases. Following Li, Ye, and Zheng (2024), we implement the following DID specification:

$$Y_{n,t} = \beta \times Treatment_n \times Post_t + \boldsymbol{\gamma}' \boldsymbol{X}_{n,t} + u_t + u_n + \epsilon_{n,t} \tag{16}$$

where $n$ indexes firms and $t$ indexes year-quarters. $Y_{n,t}$ is the dependent variable of interest, including percentage of shares repurchased ($Repurchase_{n,t}$), defined by the number of common shares repurchased, divided by the number of common shares outstanding at the beginning of the quarter; equity issuance ($Issuance_{n,t}$), defined as the sale of common and preferred stocks minus any increase in preferred stocks divided by beginning-of-quarter total assets; and dividend payouts ($Dividend_{n,t}$), defined as common stock dividends divided by beginning-of-quarter total assets. $Treatment_n$ is an indicator that equals one if a firm is in one of three test groups and zero if it is in the control group. $Post_t$ is an indicator that equals one if an observation is from the post-treatment period (2016Q4–2018Q3) and zero if it is from the pre-treatment period (2014Q4–2016Q3). Controls include size, profitability, and growth opportunity (Fama and French, 2001). We include firm fixed effects ($u_n$) to control time-invariant heterogeneity across firms, and year-quarter fixed effect ($u_t$) to control time-varying market-wise shocks. The main coefficient of interest is $\beta$, which estimates the average treatment effects of TSP on share repurchase. All continuous and unbounded variables are winsorized at the 1% and 99% percentiles. The standard errors are robust to heteroskedasticity and are clustered at the firm level.

Table 1
Summary Statistics

| Variables | N | Mean | *SD* | Q1 | Median | Q3 |
|---|---|---|---|---|---|---|
| Repurchase | 19,590 | 0.259 | 0.715 | 0.000 | 0.000 | 0.052 |
| Issuance | 19,590 | 2.532 | 16.832 | 0.000 | 0.016 | 0.190 |
| Dividend | 19,590 | 0.229 | 0.549 | 0.000 | 0.000 | 0.243 |
| Price ($) | 19,586 | 27.724 | 33.668 | 7.920 | 17.950 | 36.600 |
| Shares (million) | 19,590 | 32.443 | 30.028 | 12.404 | 26.431 | 43.917 |
| Size | 19,590 | 5.850 | 1.423 | 4.844 | 5.928 | 6.942 |
| Profitability | 19,590 | 0.090 | 0.261 | 0.008 | 0.045 | 0.095 |
| Growth | 19,590 | 2.627 | 15.408 | −2.979 | 0.329 | 3.618 |
| Quoted Spread | 19,553 | 0.849 | 1.079 | 0.229 | 0.406 | 0.953 |
| Beta | 19,544 | 1.227 | 1.098 | 0.644 | 1.166 | 1.721 |

Notes. This table presents the summary statistics for the main variables in our empirical specifications. Variable definitions can be found in Appendix B. To mitigate the potential influence of extreme values, the continuous variables are winsorized at the 1st and 99th percentiles.

We construct our sample by using all common stocks (CRSP share codes 10 and 11) in the TSP program, utilizing the Pilot Securities and Change files maintained by FINRA.[18] The sample period spans from 2014Q4 to 2018Q3, allowing for eight quarters of data both before and after the TSP implementation. Following conventions in the payout literature (Fama and French, 2001), we exclude firms in the utility (SIC codes 4200–4299) and financial (SIC codes 6000–6999) sectors, as their payout activities are subject to unique regulatory constraints. Furthermore, consistent with Li, Ye, and Zheng (2024), we require the sample to have observations with non-missing values for all covariates specified in Equation (16). The final sample comprises 611 firms across the three treatment groups and 653 firms in the control group. We collect firm-level financial information from Compustat's North America Fundamentals Quarterly files; daily and monthly stock data from CRSP; percentage quoted spread data from Daily Trade and Quote (TAQ); firms' repurchase announcement dates and values from the SDC Platinum; and institutional investor holdings from the LSEG Institutional (13F) Holdings dataset.

[18]The TSP Pilot Securities and Change files are accessible via the FINRA website: https://www.finra.org/rules-guidance/key-topics/tick-size-pilot-program/data-collection-securities-and-pilot-securities-files.

Table 1 reports summary statistics for the main variables used in our analysis. To mitigate the potential influence of extreme values, the continuous variables are winsorized at the top and bottom 1% percentile. On average, firms repurchase 0.259% of their outstanding common shares per quarter. Equity issuance and common stock dividends amount to 2.532% and 0.229% of total assets, respectively. The average stock price is \$27.724, and the average number of shares outstanding is 32.443 million. The average market beta and percent quoted spread is 1.227 and 0.849%, respectively.

## 4.2 Reduction in Share Repurchase

Table 2 presents the results from estimating equation (16). Column (1) shows that treatment firms repurchase 0.077% less outstanding shares per quarter relative to controls, representing 22% of the pre-treatment sample mean. Though this quarterly repurchase effect is still modest relative to the number of common shares outstanding, its cumulative impact is economically meaningful over multiple quarters.

To validate our identification strategy, we replace $Post_t$ in equation (16) with event-time dummy variables, defined relative to the implementation date of TSP. Displayed in Figure 1 are the coefficients of event quarters intersected with $Treatment_i$, along with their respective 95% confidence intervals. In line with the parallel-trend assumption, no significant trend is observed prior to the implementation, indicating that the treatment and control groups exhibit similar trends in share repurchases prior to the TSP. Following the shock, we observe a notable divergence, indicating a significant decremental effect in share repurchase (incremental effect in share supply). To mitigate weak-instrument concerns and ensure sufficient variation, we aggregate repurchase reductions over the eight-quarter post-treatment period.

The significant decrease in share repurchase is consistent with Li, Ye, and Zheng (2024): the price ceiling under Rule 10b-18 incentivizes firms to submit limit orders for repurchases rather than market orders. However, the wider tick increments introduced by the TSP intensify speed competition between firms and other investors at the price ceiling. Since firms are generally slower than HFTs, they are at a disadvantage in this speed competition. Consequently, their limit orders suffer from reduced execution probabilities, resulting in lower overall repurchase volumes.

Table 2
Supply Shock

| | $Repurchase_{n,t}$ | $Issuance_{n,t}$ | $Dividend_{n,t}$ |
|---|---|---|---|
| | (1) | (2) | (3) |
| $\boldsymbol{Treatment_n \times Post_t}$ | **−0.077*** ** | **−0.219** | **−0.009** |
| | **(0.03)** | **(0.36)** | **(0.01)** |
| $Size_{n,t}$ | 0.035* | −3.819*** | −0.051*** |
| | (0.02) | (0.68) | (0.02) |
| $Profitability_{n,t}$ | 0.036 | −1.702*** | 0.012 |
| | (0.03) | (0.57) | (0.02) |
| $Growth_{n,t}$ | −0.002*** | 0.635*** | 0.000 |
| | (0.00) | (0.04) | (0.00) |
| Firm FE | Yes | Yes | Yes |
| Year-quarter FE | Yes | Yes | Yes |
| Obs. | 19,589 | 19,589 | 19,589 |
| Adj. $R^2$ | 0.286 | 0.431 | 0.702 |

Notes. This table presents the effect of 2016 Tick Size Pilot on repurchase, equity issuance, and dividend payouts. *Treatment* is an indicator that equals one if a stock is in one of the three test groups and zero if it is in the control group. *Post* is an indicator that equals one if a year-quarter falls into the post-treatment period (2016Q4–2018Q3) and zero if it falls into the pre-treatment period (2014Q4–2016Q3). Controls include size, profitability, and growth opportunity. The robust standard errors clustered at the firm level are shown in parentheses. ***, **, and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively.

The insignificant coefficient on $\beta$ in column (2) suggests that the TSP affected share supply through constraints on share repurchases, rather than equity issuance. Similarly, column (3) shows no significant effect of the TSP on dividend payouts. These results are consistent with the institutional features of equity issuance and dividend payments, which take place outside the secondary market and are therefore unaffected by the TSP or the price ceiling rule. The negligible effect on dividend payouts also provides indirect evidence of the exogeneity of supply shocks, as managers do not adjust dividend policy to compensate for the reduction in share repurchases.

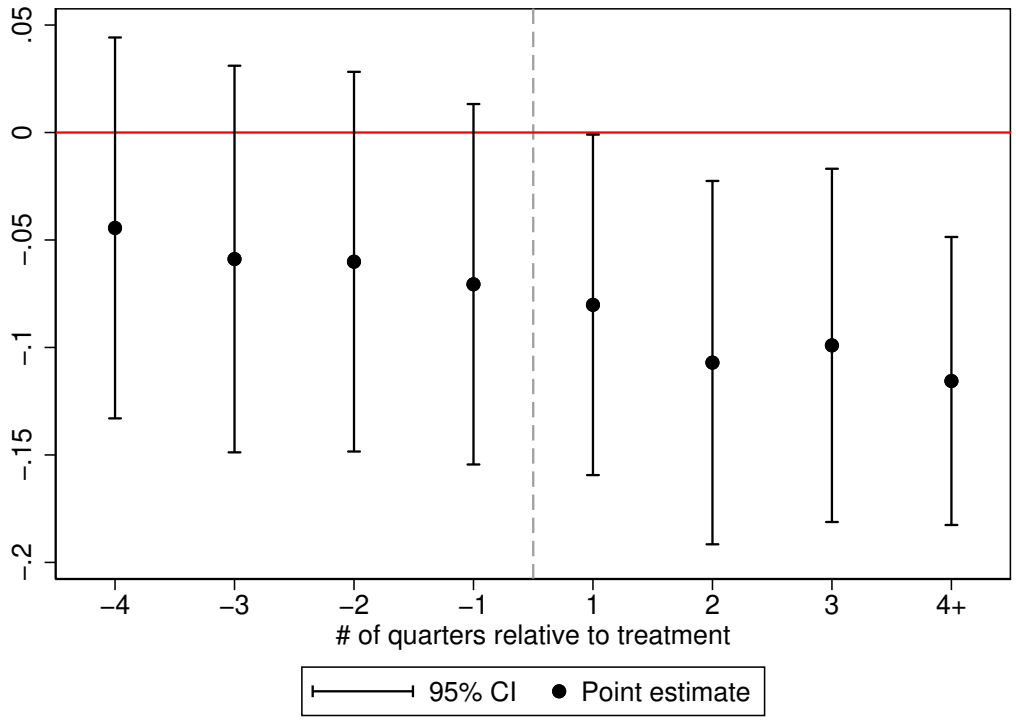



Figure 1. Event study. This figure presents an event-time analysis of the effect of TSP on share repurchases. We re-estimate the equation (16) by replacing the dummy variable $Post_t$ with a set of indicators for the quarters around the implementation of TSP for each firm in our sample. Quarter $-h$ ($h$) denotes the $h$-th quarter before (after) implementation, and Quarter 4+ aggregates all observations occurring four or more quarters after implementation. The dots report the coefficient estimates of $Treatment_i \times Post_h$, and the capped vertical lines report the corresponding 95% confidence intervals.

## 4.3 Repurchase Announcement

We next test whether the effect of the TSP on share repurchases is exogenous and unexpected by both firm managers and market participants. We modify equation (16) by replacing the dependent variables with (i) the intent of firm managers to buyback shares, and (ii) the response of investors to a repurchase announcement.

We use three measures to gauge managers' intent to repurchase: first, the announced repurchase dollar volume in a new repurchase program ($Volume_{n,t}$), reflecting the maximum amount a firm can repurchase within the program; second, the number of repurchase announcements in each quarter ($\#_{n,t}$); and third, an indicator equals one if firm $i$ initiates a new repurchase program in quarter $t$ ($\mathbb{1}_{n,t}$). Panel A of Table 3 shows no statistically sig-

Table 3
Effect of the TSP on Share Repurchase Announcement

| Panel A: Repurchase Announcement | $Volume_{n,t}$ | $\#_{n,t}$ | $\mathbb{1}_{n,t}$ |
|---|---|---|---|
| | (1) | (2) | (3) |
| $\boldsymbol{Treatment_n \times Post_t}$ | **−0.106** | **−0.001** | **−0.000** |
| | **(0.20)** | **(0.01)** | **(0.50)** |
| Pre-treatment mean of LHS | 10.376 | 0.048 | 4.602% |
| Controls | Yes | Yes | Yes |
| Firm FE | Yes | Yes | Yes |
| Time FE | Yes | Yes | Yes |
| Obs. | 579 | 19,589 | 19,589 |
| Adj. $R^2$ | 0.619 | 0.074 | 0.071 |

| Panel B: Repurchase Announcement CAR | $CAR_{n,t}$ | | |
|---|---|---|---|
| | CAPM | FF3 | FF5 |
| | (1) | (2) | (3) |
| $\boldsymbol{Treatment_n \times Post_t}$ | **−0.543** | **−0.408** | **−0.309** |
| | **(2.33)** | **(2.27)** | **(2.20)** |
| Pre-treatment mean of LHS | 1.040% | 1.126% | 1.093% |
| Controls | Yes | Yes | Yes |
| Firm FE | Yes | Yes | Yes |
| Announcement quarter FE | Yes | Yes | Yes |
| Obs. | 571 | 571 | 571 |
| Adj. $R^2$ | 0.124 | 0.138 | 0.139 |

Notes. This table presents the effects of the 2016 Tick Size Pilot on planned share repurchases. In Panel A, the dependent variables are the natural logarithm of announced repurchased dollar volume, the number of repurchase announcements per quarter, and an indicator of announcing a repurchase program at a quarter. *Treatment* is a dummy variable that equals one if a stock is in one of the three test groups and zero if it is in the control group. *Post* is an indicator variable that equals one if the year-quarter falls into the post-treatment period (2016Q4–2018Q3) and zero if it falls into the pre-treatment period (2014Q4–2016Q3). The robust standard errors clustered at the firm level are in parentheses. Panel B presents the effects of the 2016 Tick Size Pilot program on announcement CAR. In columns (1), (2), and (3), the dependent variable is the cumulative abnormal return (%) around [−1,+1] event window based on CAPM, Fama and French (1993) three-, and Fama and French (2015) five-factor model, respectively. The robust standard errors clustered at the firm and announcement quarter level are in parentheses. ***, **, and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively.

nificant results for any of the three measures, suggesting that managers of treatment firms do not alter the intensity and the probability of announcing repurchase during the TSP. Therefore, the observed decline in realized share repurchases is not driven by changes in managerial intent but rather by an exogenous shock that lowered the completion rate of repurchase programs.

Further, we examine the effect of TSP on investors' response to repurchase announcements, proxied by the three-day cumulative abnormal returns (CAR) around the [−1,+1] announcement window. Panel B of Table 3 reports the results. Columns (1), (2), and (3) present CAR estimated using CAPM, Fama and French (1993) three-factor, and Fama and French (2015) five-factor model, respectively. None of these models reveals a significant change in CAR around repurchase announcements for treatment firms during the TSP. These results indicate that market participants do not anticipate the adverse effect of the TSP on share repurchases and thus do not adjust their response to new repurchase announcements. These results also imply that the supply shocks induced by the TSP are uninformed.

Taken together, the decline in repurchases among treatment firms during the TSP reflects an exogenous, unexpected shift in the supply curve. Relative to controls, treatment firms experienced lower *ex post* completion rates of repurchase programs. These results extend Li, Ye, and Zheng (2024) by showing that the TSP constitutes an exogenous and unexpected supply shock.

# 5. Estimation of Price Multiplier

In this section, we first outline our identification strategy and present baseline estimates of the relative price multiplier. We then empirically show that the estimated relative price multiplier closely approximates the own-price multiplier. Finally, we exploit the regulatory dark-pool constraint, which applies exclusively to test group 3. This restriction induced an additional decline in share repurchases and stock prices for test group 3, while leaving overall liquidity conditions unchanged. This unique setting allows us to isolate the price impact of supply shocks from broader liquidity effects.

## 5.1 Estimation of Relative Price Multiplier

After validating the TSP as an exogenous and uninformed supply shock, we now estimate the relative price multiplier using the specification developed in Section 3.2. To control for spillover effects, we exploit the TSP's random assignment of treatment and control stocks and match each treatment stock with a closest control stock.

Specifically, we match each TSP-treatment firm to a TSP-control firm within the same Fama-French 48 industry based on shares repurchased and profitability. We implement matching via the nearest-neighbor propensity score method without replacement. We construct the matching covariates using average pre-treatment characteristics over 2014Q4–2016Q3, ensuring comparability while avoiding conditioning on outcomes potentially affected by the TSP. We further impose a caliper of 0.05 to restrict matches to the region of common support. The final sample consists of 524 treatment–control pairs. Table 4 reports the pre-treatment mean of various characteristics for treatment firms and their matched controls. It shows no statistically significant differences across observable dimensions, such as stock price, total shares outstanding, and percentage quoted spread.

**Construction of Supply Shocks.** To ensure the supply shocks are of sufficient magnitude for identification, we cumulate the quarterly reduction in shares repurchased over eight post-treatment quarters. Specifically, we construct the corresponding supply shock as the cumulative double-difference in share repurchases within the matched pair at the end of the TSP (2018Q3):

$$\Delta s_j = \sum_{t=1}^{8} \left[ \left( Repurchase_{j,t}^{T} - \overline{Repurchase_j^{T}} \right) - \left( Repurchase_{j,t}^{C} - \overline{Repurchase_j^{C}} \right) \right] \tag{17}$$

where $Repurchase_{j,t}^{T}$ and $Repurchase_{j,t}^{C}$ denote the percentage of shares repurchased in quarter $t$ by the treatment firm and its matched control, respectively; $\overline{Repurchase}j^{T}$ and $\overline{Repurchase}j^{C}$ denote their corresponding pre-treatment eight-quarter averages. The first difference removes firm-specific trends in repurchase between the pre- and post-TSP periods. The second difference (between treatment and control firms) isolates the exogenous decline in repurchase attributable to the TSP. This double-difference measure captures the exogenous

Table 4
Test of Balance on Covariates

| | Treatment | Control | $t$-statistics | $p$-value |
|---|---|---|---|---|
| | (1) | (2) | (3) | (4) |
| Repurchase | 0.275 | 0.257 | 0.606 | 0.544 |
| Size | 5.788 | 5.738 | 0.575 | 0.566 |
| Profitability | 7.807 | 7.774 | 0.031 | 0.975 |
| Growth | 2.466 | 2.201 | 0.676 | 0.499 |
| Price ($) | 24.348 | 24.634 | −0.166 | 0.868 |
| Shares Outstanding ($m) | 33.218 | 31.410 | 1.150 | 0.250 |
| Dividend | 0.220 | 0.217 | 0.082 | 0.935 |
| Beta | 1.287 | 1.222 | 1.056 | 0.291 |
| Quoted Spread | 0.867 | 0.935 | −0.975 | 0.330 |

Notes. This table examines the test of balance on observables after the propensity score matching. Matching is conducted one-to-one within the same Fama-French 48 industry using the nearest-neighbor method, based on *ex ante* firm characteristics. Columns (1) and (2) report the pre-treatment average of observables (2014Q4–2016Q3) for matched treatment and control stocks, respectively. Column (3) reports the $t$-statistics for the null hypothesis of equal means between treatment and control groups, and column (4) reports the corresponding $p$-value.

shift in the share supply faced by treatment firms during the TSP.

Figure 2 plots the trend of our constructed supply shocks, confirming a stable reduction pattern as in Figure 1. The magnitude of the cumulative supply shock at the end of TSP is −0.25%, −0.09%, and −0.62% for full treatment stocks, stocks in treatment group 1&2, and stocks in treatment group 3, respectively. These magnitudes are comparable to those found in prior literature. For example, Shleifer (1986) reports an average abnormal volume of 0.34% on index-addition announcement dates; Li and Lin (2026) document an absolute quarterly flow order imbalance of 0.58% relative to total shares outstanding.[19]

**Regression Specification.** Motivated by equation (9) and Haddad et al. (2025), we

[19]The relatively moderate size of our supply shocks also mitigates concerns that investor demand might endogenously become more elastic in response to large shocks (Chaudhry and Li, 2026). Even if the demand function is non-linear, the small shock size justifies a local approximation of the demand curve.

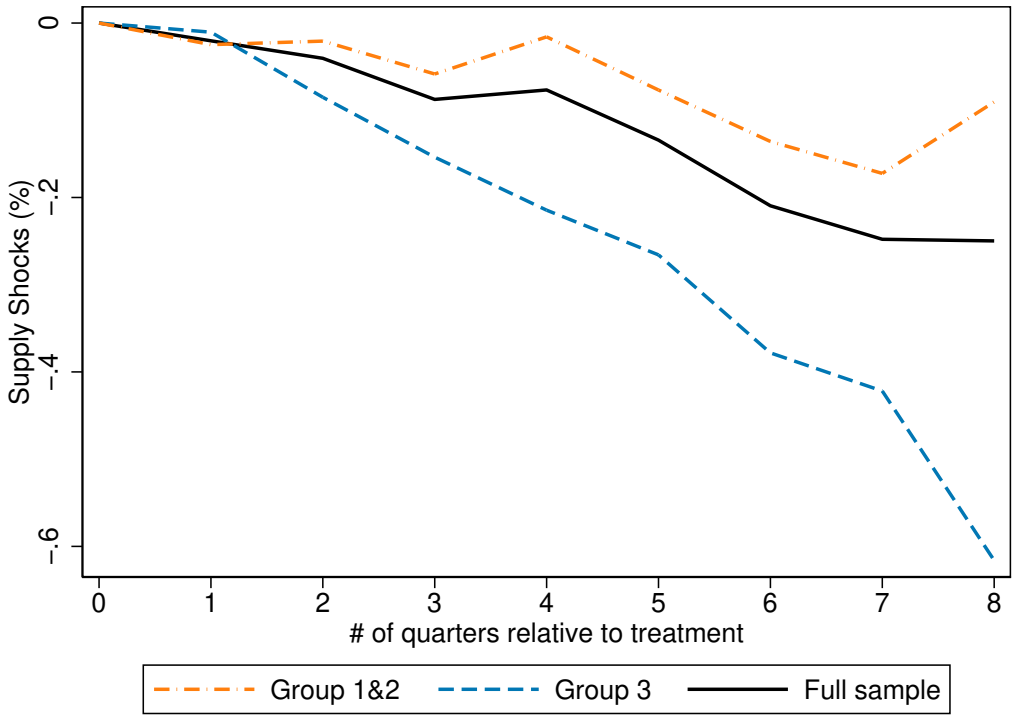



Figure 2. Dynamics of supply shocks. This figure presents the dynamics of constructed supply shocks over time. The lines represent the full sample (solid), pairs where the treatment stock is in Groups 1 & 2 (dash-dotted), and pairs where the treatment stock is in Group 3 (dashed). The horizontal axis denotes the number of quarters following the implementation of the TSP.

estimate the following regression:

$$\Delta p_j = \mathcal{M}\Delta s_j + \boldsymbol{\gamma}'\Delta\overline{\boldsymbol{x}}_j + u_g + \epsilon_j \tag{18}$$

where $j$ indexes matched pairs and $g$ denotes industries. The dependent variable $\Delta p_j$ is the double differences in log stock prices between each matched pair $j$ before and after the TSP, capturing the eight-quarter cumulative return difference induced by the TSP:

$$\Delta p_j = 100\% \times \left[\left(\log Price_{j,8}^T - \overline{\log Price_j^T}\right) - \left(\log Price_{j,8}^C - \overline{\log Price_j^C}\right)\right] \tag{19}$$

where $\log Price_{j,8}^T$ and $\log Price_{j,8}^C$ denote the log quarter-end closing price of the treatment and control stocks in the matched pair $j$ at the post-treatment quarter 8 (2018Q3), respectively. $\overline{\log Price_j^T}$ and $\overline{\log Price_j^C}$ denote the pre-treatment average log quarter-end prices of

the treatment and control stock of pair $j$, respectively. Intuitively, $\Delta p_j$ corresponds to the cumulative return on a long–short portfolio that takes a long position in the treatment stock and a short position in its matched control stock.

The coefficient $\mathcal{M}$ in regression (9) identifies the *relative* price multiplier: it measures the relative percentage change in the stock price associated with a 1% change in the relative share supply between treatment and control firms (Haddad et al., 2025). Following Cassella et al. (2025), the control vector $\Delta\overline{\boldsymbol{x}}_j$ includes pairwise differences in pre-treatment firm characteristics known to explain cross-sectional stock returns, including firm size (logarithm of total assets), profitability, growth opportunities, dividend payout, market beta, and liquidity (percentage quoted spread). In an alternative specification, we include pairwise post-treatment differences in firm characteristics, $\Delta\boldsymbol{x}_j$, to account for potential price effects arising from changes in observables during the TSP. We also control for industry fixed effects ($u_g$) to absorb common unobservables within the same industry. These controls absorb the potential price changes induced by changes in fundamentals or liquidity during the TSP (see Appendix A.4).

Table 5 presents the estimates of $\mathcal{M}$ under different specifications. In column (1), we estimate equation (18) without control variables or industry fixed effects. The estimated relative price multiplier, $\mathcal{M} = 2.578$, is statistically significant at the 1% significance level. In column (2), we include pre-treatment differences in firm fundamentals and industry fixed effects. The estimated $\mathcal{M}$ remains stable at 2.575 with a substantial increase in explanatory power. This magnitude implies that relative to its matched control, a 1% decrease in shares repurchased lowers the treatment stock price by 2.575%, to which we refer as the baseline estimate of the relative price multiplier. In column (3), we use post-treatment differences in firm fundamentals and industry fixed effects as the alternative set of control variables. The estimated $\mathcal{M}$ has a similar magnitude of 2.359. The stability of $\mathcal{M}$ across columns (1) to (3) suggests that the matching procedure effectively removes most confounding price effects from fundamentals, isolating the component attributable to supply shocks.[20]

[20]Measurement errors in supply shocks induced by scaling, or firms' endogenous responses to price declines, could bias the price multiplier. To address these concerns, we instrument the quarterly percentage of shares repurchased and re-estimate the model using a two-stage least squares (2SLS) approach. As reported in Table C.1, we find statistically similar results. However, the IV regression may under-control for spillover effects relative to our baseline. We provide a detailed discussion in Appendix C.1.

Table 5
Estimation of Price Multiplier

| | $\Delta p_j$ | | |
|---|---|---|---|
| | (1) | (2) | (3) |
| $\boldsymbol{\Delta s_j}$ | **2.578*** ** | **2.575** ** | **2.359** ** |
| | **(0.90)** | **(0.97)** | **(0.94)** |
| Control | No | $\Delta \overline{\boldsymbol{x}}_j$ | $\Delta \boldsymbol{x}_j$ |
| Industry FE | No | Yes | Yes |
| Obs. | 508 | 501 | 504 |
| Adj. $R^2$ | 0.007 | 0.122 | 0.306 |

Notes. This table presents the estimates of relative price multipliers. The dependent variable is $\Delta p_j$. For the post-treatment quarter 2018Q3, $\Delta p_j$ is calculated as the difference-in-differences of log quarter-end stock prices and pre-treatment average log prices (2014Q4-2016Q3), between treatment and control stock in a matched pair $j$. The main independent variable is denoted by $\Delta s_j$, estimated by the difference-in-differences of cumulative shares repurchased of matched pair $j$. Column (2) controls the industry fixed effects and the pre-treatment differences ($\Delta \overline{\boldsymbol{x}}_j = \bar{x}_j^T - \bar{x}_j^C$) in firm fundamentals, including size, profitability, growth opportunity, dividend payout, market beta, and percentage quoted spread. Column (3) alternatively controls for post-treatment differences ($\Delta \boldsymbol{x_j} = (x_{j,8}^T - \bar{x}_j^T) - (x_{j,8}^C - \bar{x}_j^C)$) in firm fundamentals. The robust standard errors clustered at the industry level are in parentheses. The robust standard errors clustered at the industry level are in parentheses. All continuous and unbounded variables are winsorized at the 1% and 99% levels. ***, **, and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively.

While our baseline specification (18) estimates the relative price multiplier using the supply shocks cumulated over eight quarters of the TSP effective period, we also examine how the estimate evolves over time. To do so, we re-estimate the price multiplier using supply shocks cumulated to different horizons:

$$\Delta p_{j,h} = \mathcal{M}_h \Delta s_{j,h} + \boldsymbol{\gamma}' \Delta \overline{\boldsymbol{x}}_j + u_g + \epsilon_{j,h} \tag{20}$$

where quarter $h$ refers to the $h$-th quarter after the TSP is implemented. Figure 3 shows that the estimated relative price multiplier is statistically insignificant over the first four post-treatment quarters but becomes significantly positive and remains stable thereafter. This pattern is consistent with Figure 1, which shows limited change in repurchases in the first post-treatment quarter but gradually larger effects over time. Figure 2 further shows

that the quarterly reduction in share repurchases cumulates steadily, with no sign of reversal before the end of the TSP. Together, these patterns indicate that it takes several quarters for repurchases to cumulate sufficiently to generate a statistically substantial price impact.

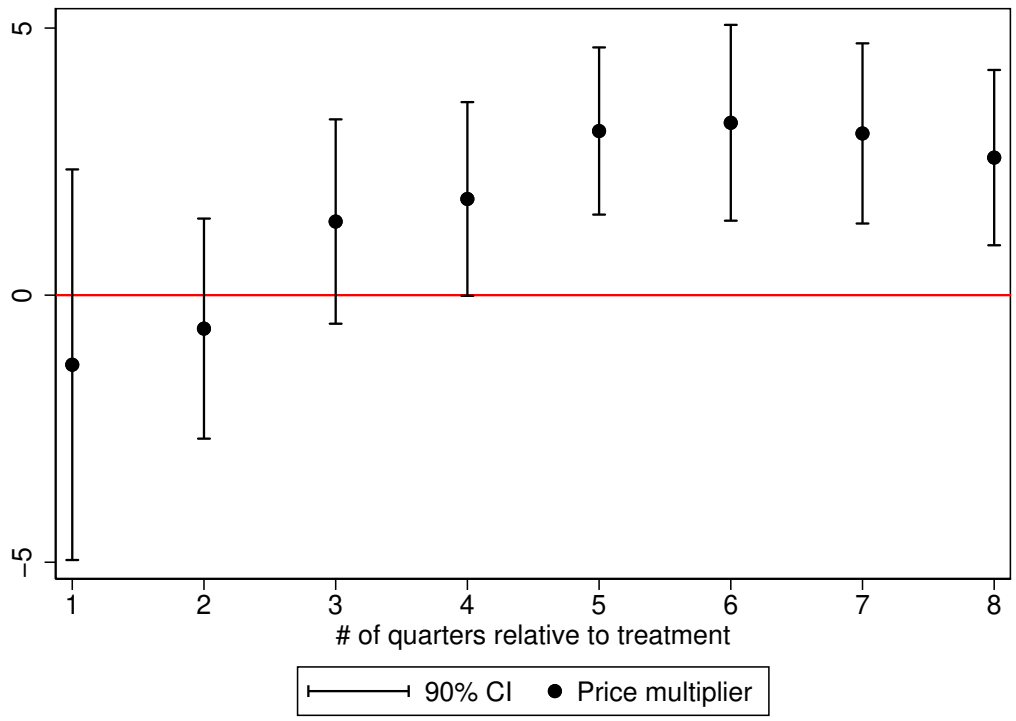



Figure 3. Dynamic estimates of the relative price multiplier. This figure plots the evolution of the relative price multiplier estimated using Equation (20) over varying time horizons. The dots represent point estimates, while the capped vertical lines indicate 90% confidence intervals. The horizontal axis denotes the number of quarters following the implementation of the TSP.

## 5.2 Relative Price Multiplier Approximates Own-Price Multiplier

As discussed in Section 3, equation (18) identifies the relative price multiplier à la Haddad et al. (2025). However, the own-price multiplier is more informative for firm managers, who are primarily concerned with the effect of repurchasing their own equity. Building on Proposition 1, we conduct two empirical tests to assess whether the estimated relative price multiplier empirically approximates the own-price multiplier at the stock level. Specifically, we verify two conditions: (1) the matched treatment and control stock pair $(n, n')$ exhibit the same systematic risk exposure, i.e., $\boldsymbol{b}(n) = \boldsymbol{b}(n')$; and (2) systematic risk exposure is

small enough relative to idiosyncratic risk, i.e., $\boldsymbol{b}^2(n) \ll \sigma^2$.[21]

### 5.2.1 Test of Balance on Systematic Risk Exposure

For each matched pair, we construct a portfolio by going long the treatment stock and shorting the control stock. We use the market excess return as the systematic risk factor in our factor structure assumption (10). For each year-month, we regress the daily excess returns of the long-short portfolio on the daily market excess return $r_{m,t}$ over the past 12 months to estimate the factor loading $\beta_j$:

$$r_{j,t}^T - r_{j,t}^C = \alpha_j + \beta_j r_{m,t} + \nu_j \tag{21}$$

Let $n$ and $n'$ denote the treatment and control stocks within matched pair $j$. Since we normalize the variance of the risk factor to one, we have that

$$\beta_j = \frac{\boldsymbol{b}_j(n) - \boldsymbol{b}_j(n')}{\sigma(r_m)}$$

in which $\sigma(r_m)$ is the standard deviation of the market excess return. Therefore, $\beta_j$ captures the difference in the systematic risk exposure between the matched treatment and control stocks. Figure 4 plots the cross-sectional average of $\hat{\beta}_j$ across matched pairs and the corresponding $t$-statistics from October 2014 to September 2018. We find that the average coefficients fluctuate around zero and are statistically indistinguishable at 5% significance level for the entire 48 months. This finding suggests that the treatment and control stocks within each pair exhibit the same factor exposures, i.e., $\boldsymbol{b}(n) = \boldsymbol{b}(n')$.

Therefore, under the single-factor structure, treatment and control stocks share similar cross-price multipliers with respect to other stocks in the market. This result supports the validity of our empirical strategy, which relies on differencing price changes between matched stocks to remove spillover effects.[22]

[21] Empirically, identifying individual investor $i$'s systematic risk exposure to stock $n$, $\boldsymbol{b}_i(n)$, requires additional assumptions. We therefore assume homogeneous (or average) factor exposures across investors and suppress the subscript $i$ in our empirical implementation. We also assume constant idiosyncratic risk across stocks; empirical support for this assumption is provided in Appendix C.2 and Figure C.1.

[22] Recall from Lemma 1 that the cross-price multiplier between stocks $n$ and $m$ is given by $[\boldsymbol{\mathcal{M}}]_{n,m} =$

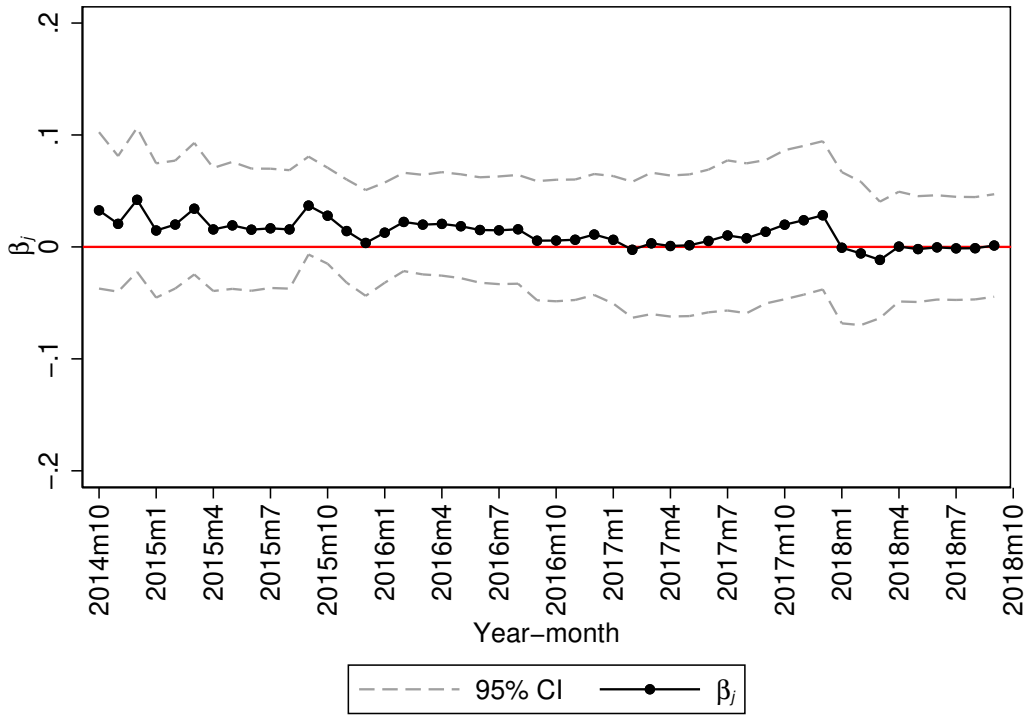



Figure 4. Test of balance on systematic risk exposure. This figure plots the average systematic exposure of long-short portfolio returns between matched treatment and control stocks to the excess stock market returns. At each month, we obtain the market beta of each matched pair, and then calculate the cross-sectional mean across matched pairs and the corresponding 95% confidence interval. The time range is from October 2014 to September 2018. Estimates are winsorized at the 1% and 99% levels.

### 5.2.2 Relatively Large Idiosyncratic Risk

We use the market excess return as the systematic risk factor and estimate the time-series regression of stock $n$'s daily excess return:[23]

$$r_t(n) = \alpha(n) + \beta(n) r_{m,t} + \varepsilon_t(n) \tag{22}$$

where $r_t(n)$ is the daily excess return of stock $n$ on day $t$, and $r_{m,t}$ is the daily market excess return. For each year-month, we compute the idiosyncratic risk $\sigma^2(n)$ as the variance of the residuals $\varepsilon_t(n)$ from equation (22) over the past 12 months. Since our model normalizes the

$\frac{\sigma^2}{\tau^2}\boldsymbol{\rho}_{n'}\boldsymbol{L}\boldsymbol{\rho}_m$. If $\boldsymbol{b}(n) = \boldsymbol{b}(n')$, then $\boldsymbol{\rho}_n = \boldsymbol{\rho}_{n'}$, implying $[\boldsymbol{\mathcal{M}}]_{n,m} = [\boldsymbol{\mathcal{M}}]_{n',m}$ for all $m \neq n, n'$.

[23] We also report the idiosyncratic risk calculated from Fama-French three factors, Carhart four factors, and Fama-French five factors in Table C.2 of Appendix C.2 and they all exhibit the same magnitudes.

variance of the systematic risk factor to one, the systematic risk exposure $\boldsymbol{b}_i^2(n)$ equals the square of the estimated loading $\beta^2(n)$ multiplied by the market factor variance:

$$\boldsymbol{b}_i^2(n) = \hat{\beta}^2(n)\sigma^2(r_m)$$

We then test the null hypothesis that the systematic risk exposure is no smaller than $c$ fraction of the idiosyncratic risk:

$$H_0 : \ln\left(\frac{\hat{\beta}^2(n)\sigma^2(r_m)}{\sigma^2(n)}\right) \geq \ln(c) \tag{23}$$

Panel A of Table 6 reports summary statistics for annualized idiosyncratic risk and systematic risk exposure for our sample stocks during the two-year TSP periods. The average idiosyncratic risk (5.704 b.p.) is 8.4 times larger than the average systematic risk exposure (0.676 b.p.).[24]

Panel B reports one-sided $p$-values from the test in (23) for various thresholds $c$. We reject the null hypothesis at the $c = 0.07$ ($c = 0.059$) threshold, corresponding to the 1% (10%) significance level. It indicates that, on average, the systematic risk exposure is less than 7% (5.9%) of idiosyncratic risk at these respective significance levels. These results confirm that systematic exposure risk is substantially smaller than idiosyncratic risk, thus satisfying the condition $\boldsymbol{b}^2(n) \ll \sigma^2$ stated in the CAPM illustration of Proposition 1.

While the relatively large idiosyncratic risk implies that the own-price multiplier is much larger than a singular cross-price multiplier, this finding does not imply that spillover effects aggregating across all other shocked stocks can be ignored. To demonstrate this, consider the special case of CAPM, and we can express the spillover effects to a TSP treatment stock $n$ as follows:

$$\text{Spillover Effects to Stock } n = \sum_{m=1}^{1,199} \gamma \boldsymbol{b}(n)\boldsymbol{b}(m)\Delta S(m)$$

[24]This relative size matches with the previous empirical findings. For example, Campbell et al. (2001) show that the idiosyncratic risks are 4.2 (6.2) times larger than the market (industrial) risks on average from July 1962 to December 1997. Goyal and Santa-Clara (2003) extrapolate that the idiosyncratic risks are 6.1 times larger than the systematic risks. Table 2 of Bartram, Brown, and Stulz (2019) report that the average of $\sigma^2/\sigma^2(r_m)$ (assuming average the $\boldsymbol{b}(n)$ of 1) from 1961 to 2017 ranges from 4.60 to 10.41 under different CRSP Index Volatility periods.

Table 6
The Comparison of Idiosyncratic Risk and Systematic Risk Exposure

| | N | Mean | $SD$ | Q1 | Median | Q3 |
|---|---|---|---|---|---|---|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Panel A: Summary statistics | | | | | | |
| Idiosyncratic ($\sigma^2(n) \times 10^{-4}$) | 24,937 | 5.704 | 3.924 | 2.737 | 4.474 | 7.778 |
| Systematic ($\hat{\beta}^2(n)\sigma^2(r_m) \times 10^{-4}$) | 24,937 | 0.676 | 1.407 | 0.140 | 0.320 | 0.637 |
| Panel B: Statistical tests | | | | | | |
| | $c = 0.10$ | $c = 0.09$ | $c = 0.08$ | $c = 0.07$ | $c = 0.06$ | $c = 0.059$ |
| $p-$value | 0.000*** | 0.000*** | 0.000*** | 0.000*** | 0.040** | 0.082* |

Notes. This table compares the idiosyncratic risk and systematic risk exposure. Panel A presents the summary statistics of estimated annualized idiosyncratic risk and systematic risk exposure. Idiosyncratic risk is calculated as the variance of the daily regression residual from equation (22) multiplied by 252. Systematic risk exposure is calculated as the product of estimated betas from equation (22) and the variance of daily market factor multiplied by 252. Panel B reports the one-sided $p$-value under a different threshold level $c$ and the null hypothesis specified in equation (23). The sample period is from October 2014 to September 2018. ***, **, and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively.

Since $N$=1,200 stocks were simultaneously shocked in TSP, the spillover effects arise from the other 1,199 treatment stocks. Moreover, all treatment stocks experienced a reduction in share repurchases ($\Delta S(m) < 0,\ m = 1, \cdots, 1{,}200$). Therefore, the spillover effects from all other shocked stocks are non-negligible. Failing to account for them would overestimate the relative price multiplier, as discussed in Section 3.2.

In summary, the estimated relative price multiplier in our sample approximates the own-price multiplier according to Proposition 1. At the same time, our results underscore the necessity of controlling for spillover effects.

## 5.3 Controlling for Liquidity Effect on Price

In this section, we provide more evidence that our estimated relative price multiplier in Table 5 identifies the price effect of supply shock, instead of liquidity. We exploit a unique feature of the TSP: the stocks in test group 3 (G3) were subject to additional "trade-at" rule.

This rule required that orders be executed in lit venues unless dark venues could improve

the NBBO by more than 2.5 cents. However, under Rule 10b-18, dark orders that improve the best bid by 2.5 cents are suspected as "manipulative". As a result, the "trade-at" rule, in conjunction with Rule 10b-18, effectively discouraged G3 firms from repurchasing shares in dark pools. Li, Ye, and Zheng (2024) document that this regulatory constraint led G3 stocks to reduce repurchase volume more than G1&2 stocks. We validate this channel by replicating Li, Ye, and Zheng (2024). We estimate equation (16) separately for G1&2 and G3 stocks, respectively. Panel A of Table 7 shows that G3 stocks reduce repurchase by 0.116% per quarter (33% of the pre-treatment mean), while G1&2 stocks reduce repurchase by 0.066% per quarter (19% of the pre-treatment mean). The difference is both economically (14%) and statistically significant at 10% level.

Albuquerque, Song, and Yao (2020) document an empirical puzzle: while the "trade-at" rule had the same negative impact on market liquidity across three groups, G3 stocks experienced larger price decline than G1&2 stocks. This setting provides an ideal environment to examine the behavior of price multiplier estimates when liquidity is held constant. As the liquidity is dampened by the same magnitude, the price discrepancy between G1&2 and G3 should be largely attributable to the differential reduction in share repurchases. Accordingly, the relative price multipliers for G1&2 and G3 should be similar with each other. Furthermore, if our identification captures the price effect mainly driven by share repurchases instead of liquidity, both estimates should have similar magnitude to the baseline estimate.

We then re-estimate relative price multipliers for G1&2 and G3 using their corresponding matched control firms. Panel B of Table 7 reports the results. Column (1) shows that the estimated price multiplier for G1&2 and G3 is 2.695 and 3.186, respectively. The difference between the two estimates is not statistically significant ($p$-value = 0.391), indicating that the relative price multipliers of G1&2 and G3 are statistically indistinguishable. The row "Test of = baseline" reports the $p$-value of the difference test between each estimate with respect to the baseline relative price multiplier of 2.575 reported in column (2) of Table 5. Neither estimates are statistically different from the baseline magnitude.

Using equation (20), Figure 5 formally visualizes this comparison by plotting the estimated relative price multipliers at each post-treatment horizon for three samples: the full set of treatment stocks (baseline), treatment stocks in G1&2, and treatment stocks in G3.

Table 7
G1&2 v.s. G3 Stocks

| Panel A: Effect of TSP on share repurchase | | | |
|---|---|---|---|
| | $Repurchase_{n,t}$ | | |
| | G1&2 | | G3 |
| | (1) | | (2) |
| $Treatment_n \times Post_t$ | −0.066** | | −0.116** |
| | (0.03) | | (0.05) |
| **Test of (1) = (2)** | | **0.060*** | |
| Control | Yes | | Yes |
| Firm FE | Yes | | Yes |
| Year-quarter FE | Yes | | Yes |
| Obs. | 12,197 | | 5,381 |
| Adj. $R^2$ | 0.290 | | 0.261 |

| Panel B: Price multiplier of G1&2 and G3 | | | |
|---|---|---|---|
| | $\Delta p_j$ | | |
| | G1&2 | G3 | Long G1&2 Short G3 |
| | (1) | (2) | (3) |
| $\Delta s_j$ | 2.695** | 3.186* | 3.358* |
| | (1.12) | (1.65) | (1.73) |
| **Test of difference** | **0.391** | | |
| **Test of = baseline** | **0.460** | **0.389** | **0.476** |
| Control | Yes | Yes | Yes |
| Industry FE | Yes | Yes | Yes |
| Obs. | 346 | 144 | 142 |
| Adj. $R^2$ | 0.132 | 0.132 | 0.735 |

Notes. This table presents the subgroup analyses that isolates price changes from general liquidity influences. Panel A re-estimates equation (16) separately for G1&2 and G3. Controls include size, profitability and growth opportunity. Firm and year-quarter fixed effects are included. The robust standard errors clustered at firm level are in parentheses. Columns (1) and (2) in Panel B re-estimate equation (18) separately for G1&2 and G3. In column (3), we match each pair whose treatment stock is assigned to G3 to a pair whose treatment stock is in G1&2, and then differentiate the double-differences of all variables, and re-estimate equation (18). The robust standard errors clustered at industry level are in parentheses. The reported $p$-value for difference is based on 1,000 bootstrap iterations. ***, **, and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively.

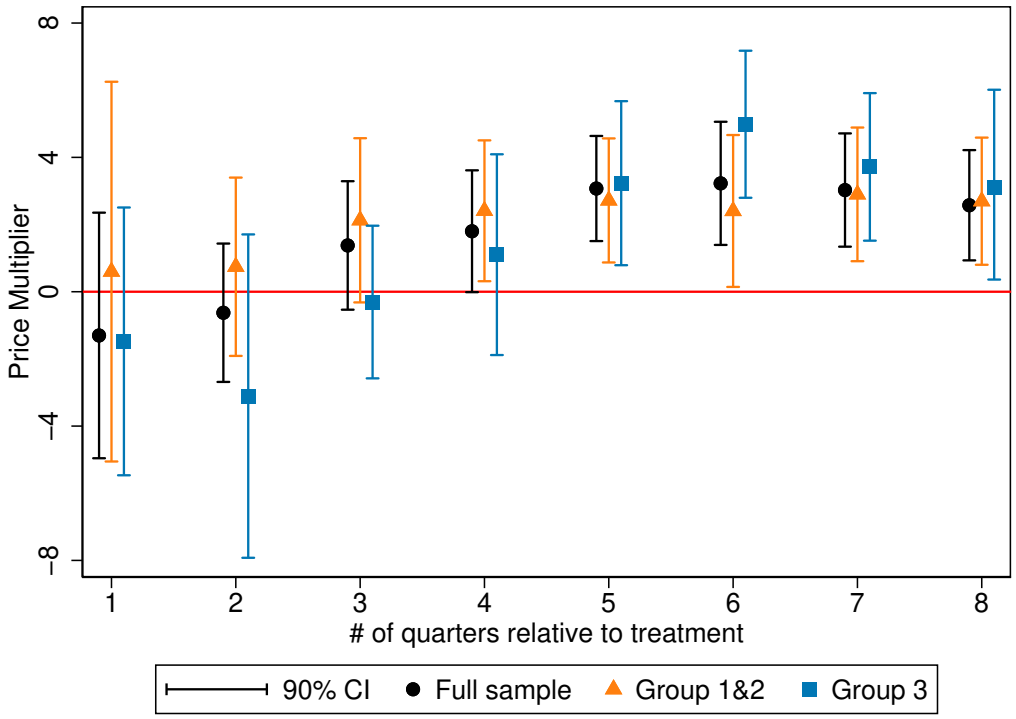



Figure 5. Estimates of relative price multiplier by treatment groups. This figure presents the estimates of relative price multiplier by horizons (quarters) using equation (20). The dots, triangles, and squares show the estimates of relative price multiplier using treatment stocks in the full sample, group 1&2, and group 3, respectively. The capped vertical lines report the 90% confidence intervals. Quarter $h$ refers to the $h$-th quarter after the TSP is implemented.

Across all three samples, the point estimates are close to zero and statistically insignificant over the first four post-treatment quarters, consistent with the gradual accumulation of repurchase-driven supply shocks. From the fourth quarter onward, the estimates turn significantly positive and remain stable. Importantly, the point estimates across the three samples are indistinguishable from one another at every horizon, and their 90% confidence intervals overlap throughout. This pattern reinforces that the estimated price multipliers are driven primarily by supply shocks rather than by a general liquidity effect.

Finally, we directly remove the liquidity effect by forming "long-G1&2-and-short-G3" pairs. We match each pair whose treatment stock is assigned to G3 to a pair whose treatment stock is in G1&2, and then differentiate the double-differences of all variables, including stock prices and shares repurchased.[25] Since G1&2 and G3 stocks exhibit the same liquidity

[25]For example, for a stock in G3 and its matched stock from G1&2, $\Delta s_j = \Delta s_{G1\&2} - \Delta s_{G3}$, in which

conditions, this additional layer of differentiation cancels out common liquidity effects on prices. Column (3) reports the resulting estimate of 3.358. The sharp increase in adjusted $R^2$ suggests that the supply-side effect explains most of the variation in the additional price decrease of G3 compared to G1&2 stocks. The row "Test of = baseline" shows that this estimate is not statistically different from the baseline magnitude.

Taken together, these results help reconcile the puzzle documented in Albuquerque, Song, and Yao (2020): the larger price decline observed for G3 relative to G1&2 is primarily driven by a larger reduction in share repurchases rather than liquidity.

## 5.4 Robustness Tests

We conduct a series of robustness checks in this section. First, we address the concern that firms that never repurchased shares could bias the estimate due to systematic differences in unobserved characteristics correlated to stock prices. We re-match treatment and control firms conditional on firms that repurchased at least one quarter during the pre-TSP period and then re-estimate equation (18) using this restricted sample. Column (1) of Table 8 reports the results. The estimated relative price multiplier remains positive and statistically significant, with a magnitude of 2.407. This estimate is insignificantly different from the baseline estimate of 2.575, which confirms that the baseline effect is not driven by firms that do not participate in repurchases.

To put the price multiplier estimates in different contexts, we re-estimate the price multiplier by changing the definitions of prices and quantities. In column (2) of Table 8, we redefine $\Delta s_j$ as the double-differences of cumulative total expenditures in common stock repurchases (CSHOPQ×PRCRAQ), scaled by lagged total assets (ATQ). The estimate of 2.946 indicates that relative to its closest substitute, an exogenous 1% increase in uninformed share supply (scaled by total assets) decreases share prices by 2.946% on average. In column (3), we define $\Delta p_j$ as the double-differences in log market capitalization following Koijen and Yogo (2019). The estimated relative multiplier of 2.073 suggests that, relative to its closest substitute, a 1% exogenous increase in share supply leads to an average of 2.073% decline in firm's market value. In column (4), we re-calculate $\Delta p_j$ and $\Delta s_j$ in the post-treatment

$\Delta s_{G3}$ and $\Delta s_{G1\&2}$ are constructed separately from equation (17).

Table 8
Robustness Checks

| | $\Delta p_j$ | | | |
|---|---|---|---|---|
| | Repurchasing firms only (1) | Alternative quantity (2) | Alternative price (3) | Alternative quarter (4) |
| $\Delta s_j$ | 2.407** | 2.946*** | 2.073*** | 3.073*** |
| | (1.07) | (0.72) | (0.70) | (0.93) |
| Control | Yes | Yes | Yes | Yes |
| Industry FE | Yes | Yes | Yes | Yes |
| Obs. | 270 | 501 | 501 | 507 |
| Adj. $R^2$ | 0.082 | 0.135 | 0.166 | 0.109 |

Notes. This table presents the robustness checks related to our estimate of price multiplier. In column (1), we repeat the estimation using repurchasing firms only. The repurchasing firms are defined as firms ever engaged in share repurchase in the pre-treatment period. In column (2), $\Delta s_j$ denotes the double-differences of cumulative total expenditures in common stock repurchase relative to lagged total assets. In column (3), $\Delta p_j$ denotes the double-differences of log market capitalization. We remove the total assets from the control variables. In column (4), we re-calculate $\Delta p_j$ and $\Delta s_j$ in the post-treatment quarter 5 (2017Q4). The robust standard errors clustered at industry level are in parentheses. ***, **, and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively.

quarter 5 (2017Q4), and we estimate a significant relative price multiplier of 3.073.

# 6. Investor Responses to Supply Shocks

In this section, we examine portfolio rebalancing behavior across investor types during the TSP to identify which investors absorb the exogenous supply shocks. Guided by equation (15) in Section 3.4, we estimate the panel regression:

$$\Delta q_{i,j,t} = \bar{\mathcal{A}}_i \Delta s_{j,t} + \boldsymbol{\gamma}_i' \Delta \boldsymbol{x}_{j,t} + u_t \times u_g + \epsilon_{i,j,t} \tag{24}$$

where $\Delta s_{j,t}$ denotes the double-difference in percentage of shares repurchased for matched pair $j$ in quarter $t$. Investor types $i$ include banks, insurance companies, investment advisors (e.g., hedge funds), mutual funds, pension funds, other 13F institutions (e.g., endowments,

foundations, and non-financial corporations), and households following the taxonomy in Koijen and Yogo (2019). We include the same control variables and industry–time fixed effects as in equation (18). $\Delta q_{i,j,t}$ is the double difference in investor $i$'s quarterly holding change as of each TSP quarter $t$:

$$\Delta q_{i,j,t} = (q^T_{i,j,t} - \bar{q}^T_{i,j}) - (q^C_{i,j,t} - \bar{q}^C_{i,j}) \tag{25}$$

$q^T_{i,j,t}$ ($q^C_{i,j,t}$) denotes investor $i$'s holding changes in treatment (control) stock in pair $j$ at quarter $t$, scaled by lagged total shares outstanding. Accordingly, the regression coefficient $\overline{\mathcal{A}}_i$ in equation (24) captures the average percentage change in holdings of investor type $i$ in response to every 1% exogenous shock in supply $\Delta s_{j,t}$, relative to the matched control stock. In the case of an exogenous increase in share supply like the TSP (where $\Delta s < 0$), a negative coefficient implies a relative increase in the holdings of treatment stocks versus control stocks, thereby indicating absorption of the supply shock.

Table 9 reports the estimation results. Investment advisors (e.g., hedge funds), mutual funds, banks, and other 13F institutions significantly increase their holdings of treatment stocks relative to control stocks. When treatment firms exogenously increase share supply by 1%, relative to control firms, investment advisors and mutual funds on average purchase 0.501% and 0.249%, respectively. Banks and other 13F institutions purchase 0.102% and 0.104%, respectively. We do not find statistically significant changes in the relative holdings of insurance companies, pension funds, and households.

Equation (24) estimates the average change in holdings for each investor sector in response to the supply shocks. To examine the dynamics of portfolio rebalancing, we re-estimate equation (24) separately for each post-treatment quarter $t$ and investor type $i$, yielding the coefficient $\overline{\mathcal{A}}_{i,t}$; we then cumulate investor $i$'s responses over different post-treatment horizons by calculating $\sum_{l=1}^{t} \overline{\mathcal{A}}_{i,l}$.

Figure 6 displays the cumulative portfolio adjustments ($\sum_{l=1}^{t} \overline{\mathcal{A}}_{i,l}$) of each investor sector $i$. Consistent with Table 9, investment advisers and mutual funds exhibit the strongest responses. For example, assuming an exogenously increased cumulative share supply of 8% over eight-quarter TSP period (1% per quarter), investment advisers and mutual funds cu-

Table 9
Estimation of Investor Responses

| | $\Delta q_{i,j,t}$ | | | | | | |
|---|---|---|---|---|---|---|---|
| | Banks | Insurance companies | Investment advisors | Mutual funds | Pension funds | Other 13F institutions | Households |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
| $\boldsymbol{\Delta s_{j,t}}$ | **−0.102**** | −0.027 | **−0.501***** | **−0.249**** | −0.003 | **−0.104**** | −0.025 |
| | **(0.04)** | (0.03) | **(0.13)** | **(0.11)** | (0.01) | **(0.05)** | (0.20) |
| Control | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| I×T FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Obs. | 4,085 | 4,085 | 4,085 | 4,085 | 4,085 | 4,085 | 4,085 |
| Adj. $R^2$ | 0.032 | 0.003 | 0.071 | 0.052 | −0.010 | −0.011 | 0.001 |

Notes. This table estimates the relative holding changes in each investor type during 2016 Tick Size Pilot, specified in equation (24). The dependent variable is $\Delta q_{i,j,t}$, which is the double differences in quarterly holding change of investor type $i$ between treatment and control stocks, before and after the shock. We consider seven types of investor: banks, insurance companies, investment advisors, mutual funds, pension funds, other 13F institutions, and households. The definition of investor sectors follows Koijen and Yogo (2019). The main independent variable is denoted by $\Delta s_{j,t}$, the double-differences of quarterly shares repurchased of matched pair $j$. Control variables includes double-differences of size, profitability, growth opportunity, dividend payout, market beta, and percentage quoted spread. The robust standard errors clustered at industry-by-time level are in parentheses. ***, **, and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively.

mulatively absorb around 4% and 2% of the additional shares, respectively. Banks and other 13F institutions respond more modestly by absorbing 1%, while insurance companies and pension funds show little change in their holdings. Importantly, the insignificant coefficient for households in column (7) of Table 9 does not imply inelastic demand. Instead, households initially purchase shares following the implementation of the TSP and subsequently unwind their positions gradually as shares are reallocated to other investor sectors.

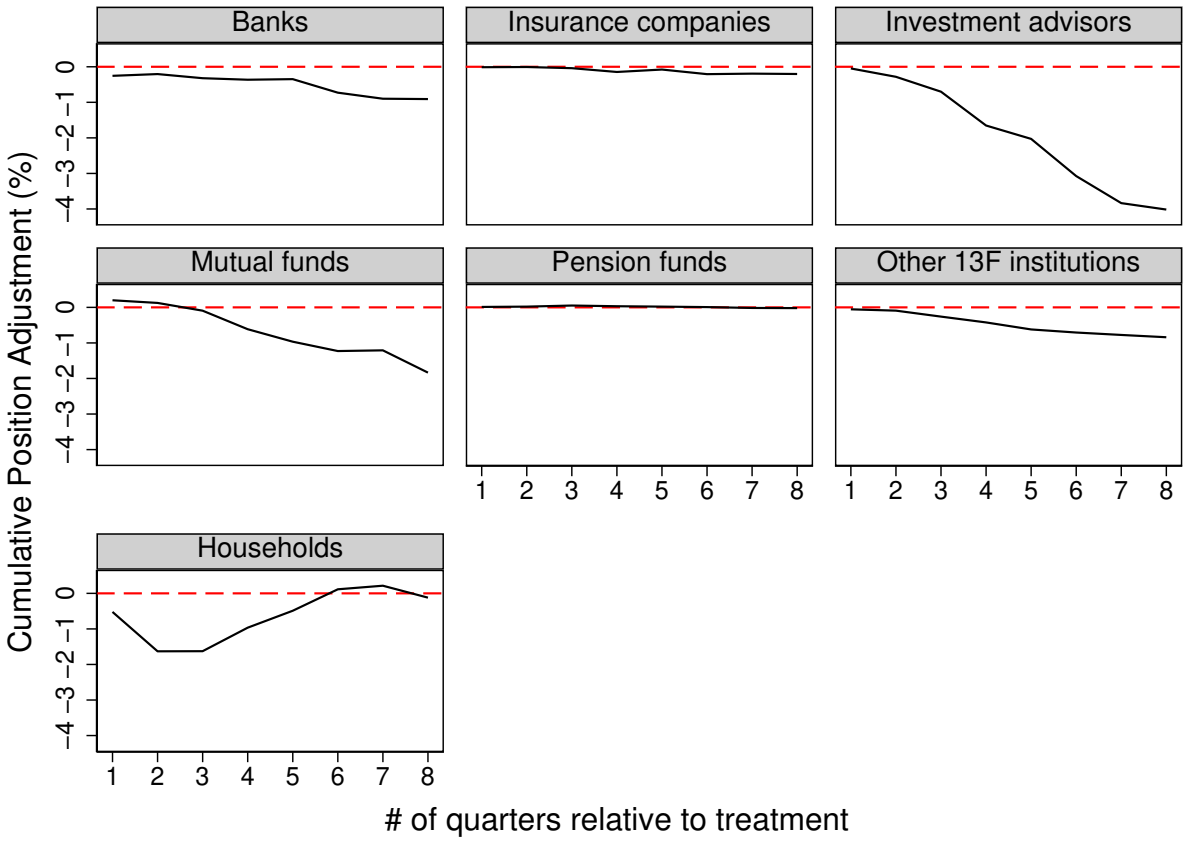



Figure 6. Cumulative portfolio adjustments by investor sectors. This figure depicts the cumulative relative holding changes in each investor type during 2016 Tick Size Pilot. At each post-treatment quarter (2016Q4-2018Q3), we estimate $\bar{\mathcal{A}}_i$ using equation (24), and then cumulative the coefficient across quarters. The definition of investor sectors follows Koijen and Yogo (2019).

# 7. Conclusions

We develop the first supply-based approach to estimate the stock-level price multiplier. We exploit the SEC 2016 TSP as exogenous and uninformed supply shocks that reduced treatment firms' share repurchases by 22%.

Taking the double differences in prices between treatment and control stocks within each matched pair controls for spillover effects from 1,198 treatment stocks outside the pair. We then estimate a relative price multiplier of 2.575. We theoretically derive and empirically validate the conditions such that our estimated relative price multiplier approximates

the own-price multiplier. In addition, because G1&2 and G3 stocks experience identical liquidity changes but different supply shocks, we estimate price multipliers separately for the two groups and obtain similar results, thereby ruling out liquidity effects on the multiplier estimates.

Exogenous supply shocks also allow us to track investors' portfolio rebalancing in response to supply shifts. We find investment advisors, mutual funds, and households primarily absorb the supply shocks, followed by banks and other 13F institutions. Insurance companies and pension funds exhibit little response. These patterns underscore the role of different investor types in responding corporate events like share repurchases.

Beyond the specific context of the Tick Size Pilot, our findings have several broader implications. First, we provide empirical support for the assumption of repurchase price impact in behavioral corporate finance. Second, we offer an alternative interpretation of abnormal returns following share repurchases: reductions in shares outstanding may mechanically raise stock prices, beyond informational channels such as signaling. Finally, our results suggest that the secular rise in U.S. share repurchases (Kahle and Stulz, 2021) may have mechanically contributed to the increase in aggregate stock valuations.

# References


Albuquerque, R., S. Song, and C. Yao (2020). The price effects of liquidity shocks: A study of the SEC's tick size experiment. *Journal of Financial Economics 138*(3), 700–724.

Baker, M. and J. Wurgler (2013). Behavioral corporate finance: An updated survey. In *Handbook of the Economics of Finance*, Volume 2, pp. 357–424. Elsevier.

Bartram, S. M., G. W. Brown, and R. M. Stulz (2019). Why is there a secular decline in idiosyncratic risk in the 2000s? *Working Paper*.

Billett, M. T. and M. Yu (2016). Asymmetric information, financial reporting, and open-market share repurchases. *Journal of Financial and Quantitative Analysis 51*(4), 1165–1192.

Binsbergen, J. H. v., B. David, and C. C. Opp (2025). How (not) to identify demand elasticities in dynamic asset markets. *Working Paper*.

Bonaimé, A. (2015). Mandatory disclosure and firm behavior: Evidence from share repurchases. *The Accounting Review 90*(4), 1333–1362.

Campbell, J. Y., M. Lettau, B. G. Malkiel, and Y. Xu (2001). Have individual stocks become more volatile? an empirical exploration of idiosyncratic risk. *Journal of Finance 56*(1), 1–43.

Carhart, M. M. (1997). On persistence in mutual fund performance. *Journal of Finance 52*(1), 57–82.

Cassella, S., A. Rizzo, O. Spalt, and L. Zimmerer (2025). Constrained by law: The impact of fiduciary duties on portfolios and prices in us equity markets. *Journal of Financial Economics*, forthcoming.

Chang, Y., H. Hong, and I. Liskovich (2015). Regression discontinuity and the price effects of stock market indexing. *Review of Financial Studies 28*(1), 212–246.

Chaudhary, M., Z. Fu, and J. Li (2025). Corporate bond multipliers: Substitutes matter. *Working Paper*.

Chaudhry, A. and J. Li (2026). Endogenous elasticities and self-stabilizing markets. *Working Paper*.

Chen, H. (2024). Cash induced demand. *Journal of Financial and Quantitative Analysis 59*(1), 195–220.

Constantinides, G. M. and B. D. Grundy (1989). Optimal investment with stock repurchase and financing as signals. *Review of Financial Studies 2*(4), 445–465.

Coval, J. and E. Stafford (2007). Asset fire sales (and purchases) in equity markets. *Journal of Financial Economics 86*(2), 479–512.

Dann, L. Y. (1981). Common stock repurchases: An analysis of returns to bondholders and stockholders. *Journal of Financial Economics 9*(2), 113–138.

Fama, E. F. and K. R. French (1993). Common risk factors in the returns on stocks and bonds. *Journal of Financial Economics 33*(1), 3–56.

Fama, E. F. and K. R. French (2001). Disappearing dividends: changing firm characteristics or lower propensity to pay? *Journal of Financial Economics 60*(1), 3–43.

Fama, E. F. and K. R. French (2015). A five-factor asset pricing model. *Journal of Financial Economics 116*(1), 1–22.

Fuchs, W., S. Fukuda, and D. Neuhann (2025). Demand-system asset pricing: Theoretical foundations. *Working Paper*.

Gabaix, X. and R. S. Koijen (2023). In search of the origins of financial fluctuations: The inelastic markets hypothesis. *Working Paper*.

Goyal, A. and P. Santa-Clara (2003). Idiosyncratic risk matters! *Journal of Finance 58*(3), 975–1007.

Haddad, V., Z. He, P. Huebner, P. Kondor, and E. Loualiche (2025). Causal inference for asset pricing. *Working Paper*.

Haddad, V., A. Moreira, and T. Muir (2025). Asset purchase rules: How QE transformed the bond market. *Working Paper*.

Hartzmark, S. and D. Solomon (2025). Marketwide predictable price pressure. *American Economic Review 115*(9), 3171–3213.

Ikenberry, D., J. Lakonishok, and T. Vermaelen (1995). Market underreaction to open market share repurchases. *Journal of Financial Economics 39*(2-3), 181–208.

Kahle, K. and R. M. Stulz (2021). Why are corporate payouts so high in the 2000s? *Journal of Financial Economics 142*(3), 1359–1380.

Koijen, R. S., F. Koulischer, B. Nguyen, and M. Yogo (2021). Inspecting the mechanism of quantitative easing in the euro area. *Journal of Financial Economics 140*(1), 1–20.

Koijen, R. S., R. J. Richmond, and M. Yogo (2023). Which investors matter for equity valuations and expected returns? *Review of Economic Studies 91*(4), 2387–2424.

Koijen, R. S. and M. Yogo (2019). A demand system approach to asset pricing. *Journal of Political Economy 127*(4), 1475–1515.

Koijen, R. S. and M. Yogo (2025). Exchange rates and asset prices in a global demand system. *Journal of Political Economy*, forthcoming.

Kvamvold, J. and S. Lindset (2018). Do dividend flows affect stock returns? *Journal of Financial Research 41*(1), 149–174.

Li, J. and Z. Lin (2026). Price multipliers are larger for less diversifiable order flows. *Working Paper*.

Li, S., M. Ye, and M. Zheng (2023). Refusing the best price? *Journal of Financial Economics 147*(2), 317–337.

Li, X., M. Ye, and M. Zheng (2024). Price ceilings, market structure, and payout policies. *Journal of Financial Economics 155*, 103818.

Lintner, J. (1969). The aggregation of investor's diverse judgments and preferences in purely competitive security markets. *Journal of Financial and Quantitative Analysis 4*(4), 347–400.

Lou, D. (2012). A flow-based explanation for return predictability. *Review of Financial Studies 25*(12), 3457–3489.

Manconi, A., U. Peyer, and T. Vermaelen (2019). Are buybacks good for long-term shareholder value? evidence from buybacks around the world. *Journal of Financial and Quantitative Analysis 54*(5), 1899–1935.

Ofer, A. R. and A. V. Thakor (1987). A theory of stock price responses to alternative corporate cash disbursement methods: Stock repurchases and dividends. *Journal of Finance 42*(2), 365–394.

Pavlova, A. and T. Sikorskaya (2023). Benchmarking intensity. *Review of Financial Studies 36*(3), 859–903.

Peyer, U. and T. Vermaelen (2009). The nature and persistence of buyback anomalies. *Review of Financial Studies 22*(4), 1693–1745.

Sammon, M. and J. J. Shim (2024). Who clears the market when passive investors trade. *Working Paper*.

Schmickler, S. and P. Tremacoldi-Rossi (2023). Spillover effects of payouts on asset prices and real investment. *Working Paper*.

Shleifer, A. (1986). Do demand curves for stocks slope down? *Journal of Finance 41*(3), 579–590.

Stein, J. C. (1996). Rational capital budgeting in an irrational world. *Journal of Business 69*(4), 429–455.

Stephens, C. P. and M. S. Weisbach (1998). Actual share reacquisitions in open-market repurchase programs. *Journal of Finance 53*(1), 313–333.

Stock, J. H. and M. Yogo (2005). Testing for weak instruments in linear iv regression. In *Identification and Inference for Econometric Models: Essays in Honor of Thomas Rothenberg*, pp. 80–108. Cambridge: Cambridge Univ. Press.

Vermaelen, T. (1981). Common stock repurchases and market signalling: An empirical study. *Journal of Financial Economics 9*(2), 139–183.

Wang, Y. (2025). Estimating investor demand elasticity with endogenous firm responses: Evidence from mutual fund flow shocks. *Working Paper*.

Ye, M., F. Zhang, and J. Zou (2025). Disclosing share repurchases more frequently? *Working Paper*.

# A. Supplementary Proofs and Discussions

## A.1 Proofs of Optimal Portfolio Choice and Equilibrium Prices

Using our assumption on the future asset payoff, $\boldsymbol{d}_i \sim \mathcal{N}(\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i)$, and the expectation properties of log-normal distribution, the investor's objective function can be expressed as

$$\begin{aligned}\mathbb{E}_i\left[-\exp\left(-\gamma_i A_{i,1}\right)\right] &= -\exp\left(-\gamma_i\left(\mathbb{E}_i(A_{i,1}) + \frac{1}{2}\mathbb{V}\text{ar}_i(A_{i,1})\right)\right) \\ &= -\exp\left(-\gamma_i\left(A_{i,0} + \left(\boldsymbol{\mu_i} - \boldsymbol{P}\right)'\boldsymbol{q}_i\right) + \frac{\gamma_i^2}{2}\left(\boldsymbol{q}_i'\boldsymbol{\Sigma}_i\boldsymbol{q}_i\right)\right)\end{aligned}$$

The first-order condition of the investor's portfolio choice is

$$-(\boldsymbol{\mu_i} - \boldsymbol{P}) + \gamma_i(\boldsymbol{\Sigma}_i)\boldsymbol{q}_i = 0$$

It solves the optimal portfolio choice in equation (1):

$$\boldsymbol{q_i} = \frac{1}{\gamma_i}\boldsymbol{\Sigma}_i^{-1}\left(\boldsymbol{\mu_i} - \boldsymbol{P}\right)$$

From equation (1), we can explicitly express market clearing condition as

$$\sum_{i=1}^{I}\frac{1}{\gamma_i}\boldsymbol{\Sigma}_i^{-1}\Big(\boldsymbol{\mu}_i - \boldsymbol{P}\Big) = \boldsymbol{S} \tag{A1}$$

Rearranging equation (A1) yields the equilibrium price vector:

$$\boldsymbol{P} = \boldsymbol{\mathcal{M}}\left(\sum_{i=1}^{I}\frac{1}{\gamma_i}\boldsymbol{\Sigma}_i^{-1}\boldsymbol{\mu}_i - \boldsymbol{S}\right) \tag{A2}$$

where $\boldsymbol{\mathcal{M}} \triangleq \left(\sum_{i=1}^{I}\frac{1}{\gamma_i}\boldsymbol{\Sigma}_i^{-1}\right)^{-1}$ is the price multiplier matrix. ■

## A.2 Proof of Lemma 1

Note that under the factor structure (10), we have $\boldsymbol{\Sigma}_i = \mathbb{V}\text{ar}(\boldsymbol{d}_i) = \sigma^2\boldsymbol{I} + \boldsymbol{b}_i\boldsymbol{b}_i'$. The Woodbury matrix identity yields $\boldsymbol{\Sigma}_i^{-1} = \frac{1}{\sigma^2}\Big(\boldsymbol{I} - \frac{\boldsymbol{b}_i\boldsymbol{b}_i'}{\sigma^2 + \boldsymbol{b}_i'\boldsymbol{b}_i}\Big)$. Applying it to $\boldsymbol{\mathcal{M}} = \Big(\sum_{i=1}^I \frac{1}{\gamma_i}\boldsymbol{\Sigma}_i^{-1}\Big)^{-1}$, we have

$$\boldsymbol{\mathcal{M}} = \left(\sum_{i=1}^I \frac{1}{\gamma_i\sigma^2}\Big(\boldsymbol{I} - \frac{\boldsymbol{b}_i\boldsymbol{b}_i'}{\sigma^2 + \boldsymbol{b}_i'\boldsymbol{b}_i}\Big)\right)^{-1} = \left(\sum_{i=1}^I \frac{1}{\gamma_i\sigma^2}\boldsymbol{I} - \sum_{i=1}^I \frac{\boldsymbol{b}_i\boldsymbol{b}_i'}{\gamma_i\sigma^2(\sigma^2 + \boldsymbol{b}_i'\boldsymbol{b}_i)}\right)^{-1}$$

For tractability, denote a scalar $\tau = \sum_{i=1}^I \frac{1}{\gamma_i}$ and an $N \times I$ matrix

$$\boldsymbol{\rho} = \left[\frac{\boldsymbol{b}_1}{\sqrt{\gamma_1(1 + \boldsymbol{b}_1'\boldsymbol{b}_1/\sigma^2)}}, \cdots, \frac{\boldsymbol{b}_I}{\sqrt{\gamma_I(1 + \boldsymbol{b}_I'\boldsymbol{b}_I/\sigma^2)}}\right]_{N\times I}$$

so that the we can express the price multiplier matrix $\boldsymbol{\mathcal{M}}$ as:

$$\boldsymbol{\mathcal{M}} = \left(\frac{\tau}{\sigma^2}\boldsymbol{I} - \frac{1}{\sigma^4}\boldsymbol{\rho}\boldsymbol{\rho}'\right)^{-1}$$

Applying the Woodbury matrix identity, we can express the matrix inverse as

$$\left(\frac{\tau}{\sigma^2}\boldsymbol{I} - \frac{1}{\sigma^4}\boldsymbol{\rho}\boldsymbol{\rho}'\right)^{-1} = \frac{\sigma^2}{\tau}\boldsymbol{I} + \frac{1}{\tau^2}\boldsymbol{\rho}\Big(\boldsymbol{I} - \frac{1}{\tau\sigma^2}\boldsymbol{\rho}'\boldsymbol{\rho}\Big)^{-1}\boldsymbol{\rho}'$$

Finally, defining $\boldsymbol{L} = \Big(\boldsymbol{I} - \frac{1}{\tau\sigma^2}\boldsymbol{\rho}'\boldsymbol{\rho}\Big)^{-1}$, we can simplify the matrix $\boldsymbol{\mathcal{M}}$ as

$$\boldsymbol{\mathcal{M}} = \frac{\sigma^2}{\tau}\boldsymbol{I} + \frac{1}{\tau^2}\boldsymbol{\rho}\boldsymbol{L}\boldsymbol{\rho}' \tag{A3}$$

■

### A.2.1 Special Case of CAPM

We derive the special case of CAPM to reveal the intuition behind the price multiplier matrix based on Lemma 1. The representative-investor CAPM implies that $\gamma_i = \gamma_j = \gamma$, $\boldsymbol{b}_i = \boldsymbol{b}_j = \boldsymbol{b}$, and $I = 1$. Then, $\tau = 1/\gamma$. The $N \times I$ matrix $\boldsymbol{\rho}$ simplifies to an $N \times 1$ vector:

$$\boldsymbol{\rho} = \left[ \frac{b(1)}{\sqrt{\gamma \left(1 + \sum_{m=1}^{N} b(m)^2/\sigma^2\right)}}, ..., \frac{b(N)}{\sqrt{\gamma \left(1 + \sum_{m=1}^{N} b(m)^2/\sigma^2\right)}} \right]'_{N \times 1}$$

and the $I \times I$ matrix $\boldsymbol{L}$ simplifies to a scalar $L$:

$$L = \left(1 - \frac{\gamma}{\sigma^2} \sum_{m=1}^{N} \frac{b(m)^2}{\gamma \left(1 + \sum_{m=1}^{N} b(m)^2/\sigma^2\right)}\right)^{-1} = \left(\frac{\sigma^2}{\sigma^2 + \sum_{m=1}^{N} b(m)^2}\right)^{-1}$$

Since $\boldsymbol{L}$ becomes a scalar, $\boldsymbol{\rho L \rho'}$ simplifies to:

$$\boldsymbol{\rho L \rho'} = L \boldsymbol{\rho \rho'} = L \frac{\boldsymbol{b b'}}{\gamma \left(1 + \sum_{m=1}^{N} b^2(m)/\sigma^2\right)} = L \frac{\sigma^2 \boldsymbol{b b'}}{\gamma \left(\sigma^2 + \sum_{m=1}^{N} b^2(m)\right)}$$

Therefore, the price multiplier matrix becomes:

$$\begin{aligned} \boldsymbol{\mathcal{M}} &= \gamma \sigma^2 \boldsymbol{I} + \gamma^2 \left(\frac{\sigma^2 + \sum_{m=1}^{N} b(m)^2}{\sigma^2}\right) \frac{\sigma^2 \boldsymbol{b b'}}{\gamma \left(\sigma^2 + \sum_{m=1}^{N} b^2(m)\right)} \\ &= \gamma \sigma^2 \boldsymbol{I} + \gamma \boldsymbol{b b'} \end{aligned}$$

The corresponding own-price, cross-price, and relative price multipliers are:

$$\begin{aligned} [\boldsymbol{\mathcal{M}}]_{n,n} &= \gamma \left[\sigma^2 + \boldsymbol{b}^2(n)\right] \\ [\boldsymbol{\mathcal{M}}]_{n,m} &= \gamma \ \boldsymbol{b}(n) \boldsymbol{b}(m) \\ \mathcal{M}_{relative}(n,m) &= \gamma \left\{\sigma^2 + \ \boldsymbol{b}(n) \left[\boldsymbol{b}(n) - \boldsymbol{b}(m)\right]\right\} \end{aligned}$$

■

## A.3 Proof of Proposition 1

Given equation (A3) and $b_i(n) = b_i(n')$ for all investor $i$, the relative price multiplier and the $(n, n)$ element of matrix $\boldsymbol{\mathcal{M}}$ are

$$\mathcal{M}_{relative}(n, n') = \frac{\sigma^2}{\tau}$$
$$[\boldsymbol{\mathcal{M}}]_{n,n} = \frac{\sigma^2}{\tau} + \frac{1}{\tau^2}\,\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_n$$

where $\boldsymbol{\rho}_n = \left( \frac{\boldsymbol{b}_1(n)}{\sqrt{\gamma_1(1 + \boldsymbol{b}_1'\boldsymbol{b}_1/\sigma^2)}}, \cdots, \frac{\boldsymbol{b}_I(n)}{\sqrt{\gamma_I(1 + \boldsymbol{b}_I'\boldsymbol{b}_I/\sigma^2)}} \right)'$ is the $n-$th column of the matrix $\boldsymbol{\rho}$. Note that

$$\frac{\mathcal{M}_{relative}(n, n')}{[\boldsymbol{\mathcal{M}}]_{n,n}} = \frac{\frac{\sigma^2}{\tau}}{\frac{\sigma^2}{\tau} + \frac{1}{\tau^2}\,\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_n} = \frac{1}{1 + \frac{1}{\tau\sigma^2}\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_n}$$

Define the ratio

$$R_n = \frac{1}{\tau\sigma^2}\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_n$$

Achieving $\dfrac{\mathcal{M}_{relative}(n, n')}{[\boldsymbol{\mathcal{M}}]_{n,n}} \to 1$ requires $R_n \ll 1$. Denote $\lambda_{\max}(\boldsymbol{X})$ and $\lambda_{\min}(\boldsymbol{X})$ as the largest and smallest eigenvalues of matrix $\boldsymbol{X}$, the upper bound of this quadratic form is

$$\begin{aligned} R_n = \frac{1}{\tau\sigma^2}\boldsymbol{\rho}_n' \boldsymbol{L} \boldsymbol{\rho}_n &\leq \frac{1}{\tau\sigma^2} \cdot \lambda_{\max}(\boldsymbol{L}) \cdot \|\boldsymbol{\rho}_n\|^2 \\ &= \frac{1}{\tau\sigma^2 - \lambda_{\max}(\boldsymbol{\rho}'\boldsymbol{\rho})} \cdot \|\boldsymbol{\rho}_n\|^2 \end{aligned}$$

The first inequality follows the Rayleigh-Ritz theorem and the second line uses the fact that $\lambda_{\max}(\boldsymbol{L}) = 1/\lambda_{\min}(\boldsymbol{I} - \frac{1}{\tau\sigma^2}\boldsymbol{\rho}'\boldsymbol{\rho}) = 1/\left(1 - \frac{1}{\tau\sigma^2}\lambda_{\max}(\boldsymbol{\rho}'\boldsymbol{\rho})\right)$. Therefore, a sufficient

condition for $R_n \ll 1$ is

$$\begin{aligned} &\frac{1}{\tau\sigma^2 - \lambda_{\max}\left(\boldsymbol{\rho}'\boldsymbol{\rho}\right)} \cdot \|\boldsymbol{\rho}_n\|^2 \ll 1 \\ \Rightarrow &\|\boldsymbol{\rho}_n\|^2 \ll \tau\sigma^2 - \lambda_{\max}\left(\boldsymbol{\rho}'\boldsymbol{\rho}\right) \end{aligned} \tag{A4}$$

Because $\lambda_{\max}(\boldsymbol{X}) \leq \text{trace}(\boldsymbol{X})$, to satisfy equation (A4), it is sufficient to show that

$$\|\boldsymbol{\rho}_n\|^2 \ll \tau\sigma^2 - \text{trace}\left(\boldsymbol{\rho}'\boldsymbol{\rho}\right)$$

Finally, note that

$$\text{trace}\left(\boldsymbol{\rho}'\boldsymbol{\rho}\right) = \sum_{i=1}^{I} \frac{\boldsymbol{b}_i'\boldsymbol{b}_i}{\gamma_i(1 + \boldsymbol{b}_i'\boldsymbol{b}_i/\sigma^2)} = \sum_{i=1}^{I} \frac{\sigma^2\boldsymbol{b}_i'\boldsymbol{b}_i}{\gamma_i(\sigma^2 + \boldsymbol{b}_i'\boldsymbol{b}_i)}$$

Therefore, the sufficient condition for $R_n \ll 1$ is

$$\|\boldsymbol{\rho}_n\|^2 = \sum_{i=1}^{I} \frac{\sigma^2\ \boldsymbol{b}_i^2(n)}{\gamma_i(\sigma^2 + \boldsymbol{b}_i'\boldsymbol{b}_i)} \ll \tau\sigma^2 - \sum_{i=1}^{I} \frac{\sigma^2\ \boldsymbol{b}_i'\boldsymbol{b}_i}{\gamma_i(\sigma^2 + \boldsymbol{b}_i'\boldsymbol{b}_i)} \tag{A5}$$

Since $\sigma^2 > 0$ and $\tau = \sum_{i=1}^{I} \frac{1}{\gamma_i}$, rearranging equation (A5) yields:

$$\sum_{i=1}^{I} \frac{\boldsymbol{b}_i^2(n) + \boldsymbol{b}_i'\boldsymbol{b}_i}{\gamma_i(\sigma^2 + \boldsymbol{b}_i'\boldsymbol{b}_i)} \ll \sum_{i=1}^{I} \frac{1}{\gamma_i} \tag{A6}$$

A sufficient condition for equation (A6), and also $R_n \ll 1$, is

$$\begin{aligned} &\frac{\boldsymbol{b}_i^2(n) + \boldsymbol{b}_i'\boldsymbol{b}_i}{\sigma^2 + \boldsymbol{b}_i'\boldsymbol{b}_i} \ll 1 \\ \Rightarrow\ &\boldsymbol{b}_i^2(n) \ll \sigma^2 \text{ or } \frac{b_i^2(n)}{\sigma^2} \to 0 \end{aligned} \tag{A7}$$

■

## A.4 Addressing the Confounding Effects

This section further addresses the concern that reduced share repurchases during the TSP may change firm observables (such as firm size), causing confounding price effects. We show that controlling for changes in observables in the regression yields an unbiased estimate of the relative price multiplier. To formalize this intuition, we first assume investors perceive the assets' expected future payoffs as a linear combination of firms' observables, following Koijen, Richmond, and Yogo (2023):

$$\boldsymbol{\mu}_i(n) = \boldsymbol{\gamma}_i' \boldsymbol{X}(n) + \varepsilon_i(n) \tag{A8}$$

where $\boldsymbol{X}(n) \in \mathbb{R}^{L\times 1}$ represents the firm $n$'s $L$ observables such as liquidity, size, growth opportunity, profitability, dividend payouts, market's beta, etc.[26] $\boldsymbol{\gamma}_i' \in \mathbb{R}^{1\times L}$ and $\varepsilon_i(n) \in \mathbb{R}$ represent investor specific preference on observables and individual demand shifter on asset $n$, respectively. We also assume investor-specific demand shifter is uncorrelated with firms' observables, such that $\text{Cov}\left(\boldsymbol{X}(n), \varepsilon_i(n)\right) = 0$.

In our main context, identifying the relative price multiplier requires investors do not change their expectations on assets' future payoffs (i.e., $\Delta\boldsymbol{\mu}_i = 0$). With assumption (A8), we can relax this condition and express the equilibrium price from equation (2) as a total differential:

$$\Delta P(n) = \underbrace{\sum_{m=1}^{N} [\boldsymbol{\mathcal{M}}]_{n,m} \Delta S(m)}_{\text{Supply Effects}} + \underbrace{\sum_{m=1}^{N} [\boldsymbol{\mathcal{M}}]_{n,m} \left( \sum_{i=1}^{I} \frac{1}{\gamma_i} \sum_{k=1}^{N} [\boldsymbol{\Sigma}_i^{-1}]_{m,k} \Delta\boldsymbol{\mu}_i(k) \right)}_{\text{Confounding Effects}} \tag{A9}$$

where $\Delta S(m) > 0$ is the exogenous amount of share repurchases, and $\Delta S(m) < 0$ is the amount of increased share supplies. Intuitively, investors may perceive different asset payoffs after observing changes in firms' observables during the supply shocks. In this case, $\Delta\boldsymbol{\mu}_i \neq 0$; the price change reflects not only the supply effects documented in our main context but also confounding effects from changing observables coinciding with the supply shocks, such

[26] We can also add industry fixed effects as control variables and repeat the deviation to get the regression with fixed effects.

as the increase in cash holding (firm size) from the reduced share repurchases. The double differences in prices between the matched stocks $(n, n')$ is therefore

$$\Delta P(n) - \Delta P(n') = \sum_{m=1}^{N} \left([\boldsymbol{\mathcal{M}}]_{n,m} - [\boldsymbol{\mathcal{M}}]_{n',m}\right) \Delta S(m) + \sum_{m=1}^{N} \left([\boldsymbol{\mathcal{M}}]_{n,m} - [\boldsymbol{\mathcal{M}}]_{n',m}\right) \left( \sum_{i=1}^{I} \frac{1}{\gamma_i} \sum_{k=1}^{N} [\boldsymbol{\Sigma}_i^{-1}]_{m,k} \Delta \boldsymbol{\mu}_i(k) \right) \tag{A10}$$

For tha matched pair $(n, n')$, we assume that $[\boldsymbol{\mathcal{M}}]_{n,m} = [\boldsymbol{\mathcal{M}}]_{n',m}, [\boldsymbol{\Sigma}_i^{-1}]_{n,m} = [\boldsymbol{\Sigma}_i^{-1}]_{n',m}$ for any $m \neq n, n'$, and that $[\boldsymbol{\mathcal{M}}]_{n,n} = [\boldsymbol{\mathcal{M}}]_{n',n'}, [\boldsymbol{\Sigma}_i^{-1}]_{n,n} = [\boldsymbol{\Sigma}_i^{-1}]_{n',n'}$.[27] Then, we can simplify equation (A10) as

$$\Delta P(n) - \Delta P(n') = \left([\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n',n}\right) \Delta S(n) + \left([\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n',n}\right) \left\{ \sum_{i=1}^{I} \frac{1}{\gamma_i} \left[ \sum_{k=1, k\neq n'}^{N} \left( \left[\boldsymbol{\Sigma}_i^{-1}\right]_{n,k} - \left[\boldsymbol{\Sigma}_i^{-1}\right]_{n',k} \right) \left(\Delta \boldsymbol{\mu}_i(n) - \Delta \boldsymbol{\mu}_i(n')\right) \right] \right\} \tag{A11}$$

$$= \left([\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n',n}\right) \Delta S(n) + \left([\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n',n}\right) \left( \sum_{i=1}^{I} \frac{1}{\gamma_i} \left( \left[\boldsymbol{\Sigma}_i^{-1}\right]_{n,n} - \left[\boldsymbol{\Sigma}_i^{-1}\right]_{n',n} \right) \left(\Delta \boldsymbol{\mu}_i(n) - \Delta \boldsymbol{\mu}_i(n')\right) \right) \tag{A12}$$

$\Delta S(n') = 0$, $[\boldsymbol{\mathcal{M}}]_{n,n} = [\boldsymbol{\mathcal{M}}]_{n',n'}$, $[\boldsymbol{\mathcal{M}}]_{n,m} = [\boldsymbol{\mathcal{M}}]_{n',m}, m \neq n, n'$, and $[\boldsymbol{\Sigma}_i^{-1}]_{n,n} = [\boldsymbol{\Sigma}_i^{-1}]_{n',n'}$ help to simplify equation (A10) to equation (A11); $[\boldsymbol{\Sigma}_i^{-1}]_{n,m} = [\boldsymbol{\Sigma}_i^{-1}]_{n',m}, m \neq n, n'$ further simplifies it to equation (A12).

Since we assume $\boldsymbol{\mu}_i(n) = \boldsymbol{\gamma}_i' \boldsymbol{X}(n) + \varepsilon_i(n)$, we have that $\Delta \boldsymbol{\mu}_i(n) - \Delta \boldsymbol{\mu}_i(n') = \boldsymbol{\gamma}_i' \Delta \boldsymbol{X}(n, n') + \Delta \varepsilon_i(n, n')$, where $\Delta \boldsymbol{X}(n, n')$ represents the double differences in observables; $\Delta \varepsilon_i(n, n')$ is the difference in investor $i$'s demand shifters between stocks $n$ and $n'$. With this assumption,

[27] Such conditions can be satisfied based on our microfoundation of price multipliers in Section 3.3.

we can further simplify equation (A12) as

$$\Delta P(n) - \Delta P(n') = \underbrace{([\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n',n})}_{\bar{\mathcal{M}}} \Delta S(n) + \underbrace{([\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n',n}) \left\{ \sum_{i=1}^{I} \frac{1}{\gamma_i} \left( \left[\boldsymbol{\Sigma}_i^{-1}\right]_{n,n} - \left[\boldsymbol{\Sigma}_i^{-1}\right]_{n',n} \right) \boldsymbol{\gamma}_i' \right\}}_{\boldsymbol{\gamma}'} \Delta \boldsymbol{X}(n,n') + \underbrace{([\boldsymbol{\mathcal{M}}]_{n,n} - [\boldsymbol{\mathcal{M}}]_{n',n}) \left\{ \sum_{i=1}^{I} \frac{1}{\gamma_i} \left( \left[\boldsymbol{\Sigma}_i^{-1}\right]_{n,n} - \left[\boldsymbol{\Sigma}_i^{-1}\right]_{n',n} \right) \Delta \varepsilon_i(n,n') \right\}}_{\varepsilon} \quad \text{(A13)}$$

In other words, even if investors perceive different asset payoffs with changing observables, we can control the potentially confounding price effects by including changes in observables as control variables. This expression aligns with our regression specification (9). Empirically, when we compare regression results before and after adding observable controls, we find identical estimates. This implies that the relative price multiplier captures the causal effect of supply shocks on prices, rather than being confounded by contemporaneous changes in fundamentals or liquidity conditions.

# B. Variable Definition

| Variable | Definition |
|---|---|
| Repurchase | Shares repurchased, estimated by the number of common stock repurchases (CSHOPQ), divided by lagged total number of common shares outstanding (CSHOQ). The number of shares is adjusted by stock split (AJEXQ). It is expressed in percentage point. |
| Issuance | Equity issuance, estimated by the sales of common and preferred stocks (SSTKQ) minus any increase in preferred stocks, scaled by lagged total assets (ATQ). Preferred stocks depend on the availability in the order of redemption (PSTKRQ) and capital (PSTKQ). It is expressed in percentage point. |
| Dividend | Dividend payouts, estimated by common stock dividends (DVQ - DVPQ), divided by lagged total assets (ATQ). Compustat quarterly data only provides the year-to-date dividends (DVY). The quarterly dividend (DVQ) is computed as DVY (in the first fiscal quarter) or the change in DVY (in the second, third, and fourth fiscal quarter). It is expressed in percentage point. |
| Treatment | An indicator that equals one if a stock is in one of the three test groups and zero if it is in the control group. |
| Post | An indicator that equals one if a year-quarter falls into the post-treatment period (2016Q4–2018Q3) and zero if it falls into the pre-treatment period (2014Q4–2016Q3). |
| Size | Natural logarithm of total assets (ATQ). |
| Growth | Growth opportunity, estimated by the quarterly growth of total assets (ATQ). It is expressed in percentage point. |

| | |
|---|---|
| Profitability | Profitability, estimated by the quarterly revenues (REVTQ) minus the sum of cost of goods sold (COGSQ) and selling, general, and administrative expenses (XSGAQ) and interest and related expenses (XINTQ), scaled by lagged book equity. Book equity is defined as the book value of stockholders' equity (SEQQ), plus balance sheet deferred taxes and investment tax credit (TXDITCQ), minus the book value of preferred stock. |
| Volume | The natural logarithm of announced repurchased dollar volume. |
| # | The number of repurchase announcements each quarter. |
| $\mathbb{1}$ | An indicator that equals one if a stock announces a repurchase announcement at the quarter and zero otherwise. |
| CAR | The cumulative abnormal return around $[-1, +1]$ event window based on CAPM, Fama and French (1993) three-, or Fama and French (2015) five-factor model. It is expressed in percentage point. |
| Quoted Spread | The average daily percentage quoted spread over a quarter. Percentage quoted spread is the difference between the national best ask and the national best bid (NBBO) at any time interval divided by the midpoint of the two. The daily percentage quoted spread is the time weighted average percentage quoted spread computed during market time. It is expressed in percentage point. |
| Beta | Market beta. For each stock-quarter, it is estimated from a regression of monthly stock excess returns over the risk-free rate, onto excess market returns using the past 60-month (five-year) observations. |
| Price | The quarter-end stock prices (PRCCQ). When calculating price differences, it is adjusted for stock splits and stock dividends by dividing to adjust factor (AJEXQ). |
| Shares | The quarter-end number of total common shares outstanding (CSHOQ). When calculating share differences, it is adjusted for stock splits and stock dividends by multiplying adjust factor (AJEXQ). |

| | |
|---|---|
| Return | The quarterly stock returns. Raw return is estimated by the first-order difference of log quarter-end stock prices (PRCCQ). The quarterly abnormal return is estimated by continuously compounding daily abnormal returns using CAPM or Fama and French (1993) three-factor model. It is expressed in percentage point. |
| Inverse Price | The inverse of quarter-end stock prices (PRCCQ). |
| Turnover | The average daily share turnover over a quarter. The share turnover is defined as the stock trading volume (VOL) to the number of total shares outstanding (SHROUT). |
| Q | Tobin's Q, calculated as the market value of equity (PRICE×CSHOQ) plus the book value of assets (ATQ) minus the book value of equity (CEQQ), scaled by total assets (ATQ). |
| $\text{Return}_{t-1}$ | The cumulative stock returns at quarter $t-1$. It is estimated by continuously compounding daily stock returns. |
| $\text{Return}_{t-2}$ | The cumulative stock returns at quarter $t-2$. It is estimated by continuously compounding daily stock returns. |
| $\text{Return}_{t-3,t-4}$ | The cumulative stock returns from quarter $t-4$ to $t-3$. It is estimated by continuously compounding daily stock returns. |
| $\text{Return}_{-n}$ | The average quarterly stock returns of stocks in the same Fama-French 48 industry but excluding stock $n$ per se. |

# C. Additional Empirical Results

## C.1 Two Stage Least Squares (2SLS) Regression

In our baseline specification, we construct supply shocks as the double difference in share repurchases between matched treatment and control firms. Although the reduction in repurchases stems from the unexpected interaction between the TSP and Rule 10b-18, a remaining concern is that such changes may still reflect firms' endogenous responses to price declines (as we cannot exhaust all firm characteristics). To further mitigate this concern, we instrument the quarterly percentage of shares repurchased ($\Delta s_{n,t}$) using treatment assignment and its interaction with the post-TSP indicator, and estimate the following 2SLS regression:

$$\begin{aligned} \text{First stage: } \Delta s_{n,t} =& \boldsymbol{\beta}' \begin{pmatrix} Treatment_n \\ Treatment_n \times Post_t \end{pmatrix} + \boldsymbol{\gamma}' \boldsymbol{X_{n,t}} + u_g \times u_t + \epsilon_{n,t} \\ \text{Second stage: } \Delta p_{n,t} =& \mathcal{M} \times \widehat{\Delta s}_{n,t} + \boldsymbol{\gamma}' \boldsymbol{X_{n,t}} + u_g \times u_t + \varepsilon_{n,t} \end{aligned} \tag{C1}$$

where $\Delta p_{n,t}$ denotes quarterly abnormal returns adjusted using the Fama and French (1993) three-factor model.[28] The control vector $\boldsymbol{X}_{n,t}$ includes the inverse of stock price, turnover, percentage quoted spread, firm size, Tobin's Q, profitability, growth opportunities, dividend payout, and market beta. We also control for past returns over quarter $t-1$, $t-2$, and $t-3$ to $t-4$ to account for return momentum and reversal, and industry-by-time fixed effects to absorb common shocks within each industry-quarter (Cassella et al., 2025). Additionally, following Chaudhary, Fu, and Li (2025), we include the average return of peer firms in the same Fama–French 48 industry to control for spillover effects from a stock's close substitutes.

Table C.1 reports the 2SLS estimates. The first-stage $F$-statistic of 13.23 exceeds conventional weak-instrument thresholds (Stock and Yogo, 2005), supporting instrument relevance. The second stage yields a statistically significant relative price multiplier of 3.460, implying that, relative to TSP control firms, a 1% exogenous reduction in shares repurchased lowers abnormal returns of TSP treatment firms by 3.460%. Although this estimate is not statis-

[28] The FF3-adjusted abnormal return removes common factor-driven variation in returns.

Table C.1
Estimation of Price Multiplier: a Two Stage Least Squares Approach

| | $\Delta p_{n,t}$ |
|---|---|
| | (1) |
| $\widehat{\boldsymbol{\Delta s_{n,t}}}$ | **3.460*** |
| | **(1.81)** |
| $F$-stat. | 13.23 |
| Control | Yes |
| Industry *times* Time FE | Yes |
| Obs. | 19,341 |
| Adj. $R^2$ | 0.255 |

Notes. This table reports the 2SLS IV regression results. The dependent variable is the quarterly abnormal returns using Fama and French (1993) three-factor model. $\widehat{\Delta s}_{n,t}$ is the instrumented quarterly shares repurchased estimated by regressing shares repurchased on treatment and DID dummies. Control variables consist of inverse of stock price, turnover, percentage quoted spread, size, profitability, growth opportunity, dividend payout, market beta, Tobin's Q, past stock returns, and average return of peer stocks in the same Fama-French 48 industry. The robust standard errors clustered at year-quarter level in column (4). All continuous and unbounded variables are winsorized at the 1% and 99% levels. ***, **, and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively.

tically different from the baseline,[29] its larger magnitude reflects the stronger substitution structure implicit in the 2SLS specification. In particular, the IV approach assumes that stocks within the same industry are close substitutes and share identical cross-price multipliers with respect to stocks outside the industry (Chaudhary, Fu, and Li, 2025). By contrast, our DID baseline requires only that each matched treatment-control pair constitute close substitutes, without imposing homogeneity of cross-price multipliers at the industry level. Under this more restrictive notion of close substitutes, the baseline estimate provides a more conservative measure of the relative price multiplier.

## C.2 Test of Balance on Idiosyncratic Volatility

We test whether the idiosyncratic volatility is similar between treatment and control stocks. For each year-month, we respectively regress the daily excess stock returns of treatment and

[29] A bootstrap test with 1,000 repetitions under the null that the IV and baseline estimates are equal yields a $p$-value of 0.871.

control stocks, on the market excess return factor over the past 12 months. Next, we calculate the equal-weighted average idiosyncratic volatility as the standard deviation of residuals and annualize it. Consistent with our assumption of constant idiosyncratic volatility, Figure C.1 shows the nearly identical patterns of idiosyncratic volatility ($t$-statistics = 0.624) between matched treatment and control stocks conditional on observables.

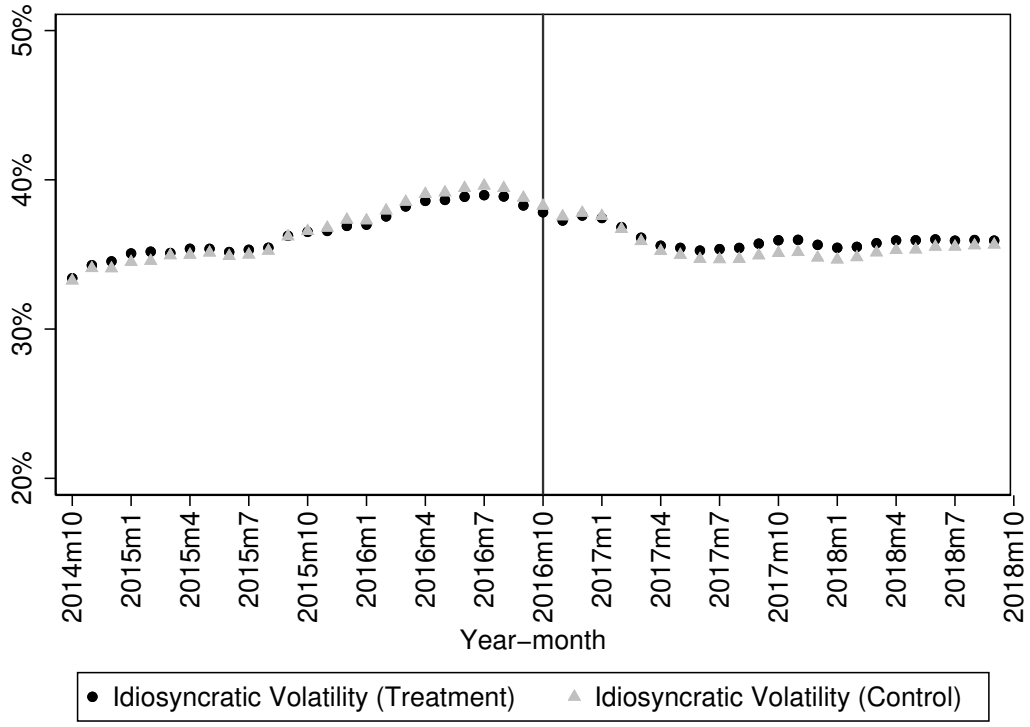



Figure C.1. Balance on variances: average annualized idiosyncratic volatility. This figure testifies our assumption of constant volatility, by comparing the average idiosyncratic volatility between treatment and control stocks. At each month, we calculate the equal-weighted average idiosyncratic volatility by regressing daily excess stock returns on market excessive return. The sample period is from October 2014 to September 2018.

Table C.2 reports $t$-tests under alternative factor structures. In addition to the market factor, we compute idiosyncratic volatility using Fama and French (1993) three-factor model, Carhart (1997) four-factor model, and Fama and French (2015) five-factor model. Across all specifications, we find no statistically significant differences in idiosyncratic volatility between treatment and control firms in our sample.

Table C.2
The $t$-test of Difference in Annualized Idiosyncratic Volatility

| Idiosyncratic Volatility | Treatment | Control | $t$-statistics | $p$-value |
|---|---|---|---|---|
| Market Factor | 36.279% | 36.087% | 0.624 | 0.534 |
| Fama-French Three Factors | 34.851% | 34.588% | 0.951 | 0.344 |
| Carhart Four Factors | 34.578% | 34.318% | 0.992 | 0.324 |
| Fama-French Five Factors | 34.429% | 34.144% | 1.034 | 0.304 |

Notes. This table reports the $t$-test for the idiosyncratic volatility calculated from different factor models. We consider Fama and French (1993) three-factor model, Carhart (1997) four-factor model, and Fama and French (2015) five-factor model. Idiosyncratic volatility is calculated as the equal-weighted average variance of residuals from the daily regression of stock returns on factors.

## C.3 Subsample Analyses

This section presents two subsample analyses. First, to examine whether the effect of the TSP on repurchase announcements remains insignificant over time, we split the eight post-treatment quarters into two equal subsamples: the first four quarters and the subsequent four quarters. We re-estimate the announcement effect separately for each period. Columns (1) and (2) of Table C.3 show that the effect of the TSP on repurchase announcements remains insignificant in both the early and later post-TSP periods. Second, we conduct the same analysis separately for Groups 1&2 and Group 3. Columns (3) and (4) indicate that the insignificance of the TSP effect persists across these subsamples as well.

## C.4 Placebo Tests

This section conducts placebo tests by changing the time of the TSP implementation. If no exogenous supply shock exists between treatment and control, we should obtain no significant variation in stock prices or share repurchases, and subsequently, an insignificant estimate of the price multiplier. In column (1) of Table C.4, we use the two years before the Pilot implementation as a placebo shock. The corresponding pre-treatment and post-treatment periods are 2013Q4–2014Q3 and 2014Q4–2015Q3, respectively. we report the estimate using the last post-treatment quarter 2015Q3. In column (2), we repeat the estimation of equation

Table C.3
Subsample Analyses

| | $CAR_{n,t}$ | | | |
|---|---|---|---|---|
| Subsample = | Earlier | Later | G1&2 | G3 |
| | (1) | (2) | (3) | (4) |
| $\boldsymbol{Treatment_n \times Post_t}$ | **3.772** | **−2.677** | **−1.598** | **1.621** |
| | **(2.41)** | **(2.83)** | **(2.67)** | **(2.18)** |
| Controls | Yes | Yes | Yes | Yes |
| Firm FE | Yes | Yes | Yes | Yes |
| Announcement quarter FE | Yes | Yes | Yes | Yes |
| Obs. | 443 | 498 | 506 | 398 |
| Adj. $R^2$ | 0.052 | 0.075 | 0.059 | 0.090 |

Notes. This table presents subsample analyses of the effects of TSP on repurchase announcement. The dependent variable is CAR (%) around [−1,+1] event window based on CAPM. In column (1) and (2), the post-treatment period is split to "Earlier" (the first four quarters after the shock) and "Later" (the second four quarters after the shock) period. In column (3) and (4), we repeat the analysis for G1&2 and G3, respectively. The robust standard errors clustered at firm and announcement quarter level are in parentheses. ***, **, and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively.

Table C.4
Placebo Tests

| | $\Delta p_j$ | |
|---|---|---|
| | Placebo | Pre-treatment |
| | (1) | (2) |
| $\boldsymbol{\Delta s_j}$ | **−0.155** | **0.641** |
| | **(0.59)** | **(2.94)** |
| Control | Yes | Yes |
| Industry FE | Yes | Yes |
| Obs. | 441 | 514 |
| Adj. $R^2$ | 0.165 | −0.006 |

Notes. This table presents the placebo tests. In column (1), we use a counterfactual shock, with pre- and post-treatment period as 2013Q4–2014Q3 and 2014Q4–2015Q3, respectively. The estimation uses the last post-treatment quarter 2015Q3. In column (2), we estimate the relative price multiplier using the pre-treamt quarter 2016Q3 only. The robust standard errors clustered at industry level are in parentheses. ***, **, and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively.

(18) using the pre-treatment quarter 2016Q3 only. We do not find significant estimate of price multiplier in either case. These results confirm that our estimates of relative price multiplier is mainly driven by the supply shocks induced by TSP.