Booms, Busts, and the PhD Pipeline: The Career Impacts of Biomedical Science Funding Volatility

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Abstract

Federal funding is the primary source of support for basic science, sustaining the laboratories where research is conducted and new scientists are trained, but it is often subject to significant variability. This study examines the effects of two types of funding shocks on research outcomes: system-wide shocks, such as the NIH doubling, and idiosyncratic shocks, such as when individual investigators unexpectedly secure multiple grants simultaneously. Using combined data at the investigator level on publications, theses, and grants, we find that while Principal Investigators (PIs) produce more papers during funding surges, these papers are of lower average quality and produced in labs that expand and reorganize their teams. PhD students who trained under PIs exposed to system-wide funding surges face weaker job-market outcomes and lower subsequent research productivity, consistent with graduating into a labor market temporarily flooded with new scientists. Idiosyncratic funding shocks, which do not create such market-level supply pressures, nonetheless also reduce the research productivity of affected mentees after graduation. Together, these patterns suggest that rapid expansions in funding can dilute effective supervision and strain mentoring capacity, lowering the quality of both scientific output and the next generation of researchers.

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I Introduction

Federal funding is the primary source of support for basic science, sustaining the laboratories where research is conducted and new scientists are trained, but it is often subject to significant variability. Biomedical science is funded by NIH, primarily. Over the past half-century, NIH appropriations in real terms have followed a boom–bust path with sizable year-to-year swings—an unusually rapid doubling in 1998–2003, followed by a decade of erosion in purchasing power, a temporary ARRA spike in 2009–2010, and sequestration cuts that left real funding in the 2010s well below its early-2000s peak. At the lab level, the aggregate changes and the vagaries of NIH's adjudication process insure labs are subject to variation in funding. While there is growing work on the consequences of funding interruptions and shortfalls especially at the PI/lab level—Babina, He, Howell, Perlman and Staudt, 2023 show that large negative shocks to federal funding reduce high-quality publications and high-tech entrepreneurship, and Tham, Staudt, Perlman and Cheng (2024) document that gaps in NIH R01 support push personnel out of U.S. science and into lower-earnings trajectories—much less is known about how positive shocks, especially strong, sudden, and brief ones, affect the training pipeline and long-term career outcomes. Our work is also complementary to Sattari, Bae, Berkes and Weinberg (2022), who show that funding increases have large "ripple effects" on publications and team size but do not study how transient surges in funding affect the quality of training or the subsequent careers of trainees.

The effectiveness of a sudden funding increase depends on the elasticity of the labor supply. In the short to medium term, due to the lengthy training requirements to become a PI, the supply of principal investigators (PIs) is inelastic.[1] The impact of the funding surge may have been further limited if these additional resources were concentrated among a smaller, established pool of PIs rather than distributed more broadly to include newer investigators. Evidence suggests that this may have been the case: according to Larson, Ghaffarzadegan and Diaz, 2012, the average age of scientists receiving grants increased from 1998 to 2003. This pattern implies that NIH may have concentrated the funding on PIs with established track records—potentially as a safer investment—resulting in larger labs under well-funded PIs but possibly at the cost of training quality for graduate students and postdocs.

So difficult to scale up in face of funding windfall. Freeman and Van Reenen (2009) formalize this idea in a simple model of federal R&D spending. With convex adjustment costs, Freeman and Van Reenen (2009) argue that a gradual increase in budget is preferable to a sudden one. When research has convex adjustment costs, they show that a 'doubling-then-squeeze' pattern—rapid expansion followed by real cuts—delivers less research output than a smoother, more predictable funding path with the same present value of spending. This is so that human capital can have the time to develop. The paper also points out that directing funding to younger PIs is preferable as they would yield a greater return on investment, assuming they possess similar abilities to more

[1] Goolsbee (1998) finds that increased funding in the global R&D sector primarily led to higher wages for engineers and other researchers rather than a substantial increase in their numbers, as the lengthy and rigorous training required to enter these fields results in an inelastic labor supply in the short term.

experienced researchers. Similarly, Teitelbaum (2008) argues for a stable funding trajectory to avoid budget interruptions or sudden decreases in funding.

In contrast, the supply of PhD students and postdocs, the primary source of labor in the PI-led labs, is likely relatively elastic. This asymmetry creates a bottleneck: when funding suddenly increases, labs expand by hiring more junior researchers, but the number of experienced mentors remains fixed. Consequently, each PI must supervise more trainees, potentially reducing the quality of mentorship and training each trainee receives.

This dynamic creates a "boom and bust" cycle. The funding surge attracts large cohorts of PhD students and postdocs into expanded labs. However, by the time these researchers complete their training, the temporary funding expansion has ended. These enlarged cohorts then face a tight job market, as the number of available PI and therefore academic positions is constrained by the return to normal funding levels.

Our hypothesis is that sudden spikes in research funding lead to larger research teams and more publications, albeit with reduced average quality. Increased funding likely results in higher enrollment of students in labs, but with more junior researchers per lab, the level of PI supervision becomes strained, potentially weakening training outcomes. Since it can potentially impact the current state of research, with publications of lesser quality, and its future, with less able students becoming the next waves of PIs, understanding these effects is crucial for agencies deciding between large budget shocks versus smoother funding trajectories.

Sudden spikes will happen without aggregate budgetary shocks. NIH paylines for investigator-initiated grants are low, and PIs face substantial year-to-year uncertainty in their lab's funding. In response, PIs often submit multiple grant applications as a form of insurance rather than tailoring submissions to a single "desired" scale of operations (Freeman and Van Reenen, 2008, 2009). Qualitative accounts of the funding process and analyses of grant competitions emphasize this dynamic of hyper-competition and over-submission, with investigators explicitly describing strategies of submitting "many grants a year" and "taking as many shots as we can" to increase the chance that at least one proposal is funded (Webel, 2025; Drugmonkey, 2025; Dobson, 2025). Given this environment, it is inevitable that some investigators experience idiosyncratic funding windfalls: more of their proposals are funded in a short window than they planned for ex ante.

Our empirical strategy treats such idiosyncratic PI-level funding windfalls as shocks to lab scale that are plausibly distinct from the steady, anticipated component of a PI's funding stream. In contrast to the system-wide "doubling" and ARRA shocks, which change the aggregate environment, these PI-level shocks exploit the fact that grant competitions with low success rates, combined with multiple submissions per PI, occasionally generate large, discrete jumps in support for individual labs. We then ask how these windfalls reverberate through the careers of the next generation of scientists trained in those labs.

To test these hypotheses, we look at the National Institute of Health which saw its budget

double between 1998 and 2003, an increment that was short lived as the budget significantly dropped at the end of the period. Such event constitutes the perfect natural experiment to observe in depth the results of positive budget shocks onto research outputs regarding the biomedical fields. To do so, we draw from multiple large datasets: (1) NIH's Exporter database, providing grant data; (2) Medline, which includes information on publications and citations in the biomedical field; and (3) ProQuest's database, which enables tracking of advisor-advisee relationships through thesis registrations. Using a staggered difference-in-differences model and a panel event study, we find that while PI productivity (in terms of publications and trained students) increased during the NIH doubling, the quality of outputs decreased, evidenced by fewer forward citations. Additionally, students trained during this period demonstrated lower productivity in the early stages of their careers, with fewer forward citations and greater difficulty securing early-career funding.

II NIH through the doubling and ARRA eras

From 1998 to 2003, Congress doubled the budget of the National Institutes of Health, the primary federal source of funding for research in the biomedical fields, from $14 billion to $27 billion, significantly increasing both the quantity and size of research grants.

While this surge in funding was expected to enhance both the quantity and quality of research output, doubts have been raised about its effectiveness. For example, Freeman and Van Reenen (2008) argue that the period of doubling not only failed to deliver the anticipated boost in research production, but also had adverse impacts in subsequent years. As funding levels were not sustained, budget cuts and interrupted projects became more common, hampering research continuity and long-term progress.

Another positive shock happened in 2009 called the American Recovery and Reinvestment Act (ARRA). While not as extreme as the doubling, it can serve as a secondary shock for our preliminary study. However, the details of the two shocks differ greatly. Indeed, ARRA was the Obama government-led fiscal stimulus in response to the Great Recession. The ARRA stimulus was probably understood by PIs and research universities to be short-lived. The doubling, however, was not explicitly communicated as temporary when it was implemented. In fact, many universities and research institutions made long-term commitments and expansions based on an assumption that the higher funding levels would be maintained or continue growing at normal rates after 2003.

Using data from the Graduate Student Survey (GSS), the Integrated Postsecondary Education Data System (IPEDS), and NIH RePORTER data, we can make a preliminary assessment on the effect of budget increases on research inputs such as number of PhD students and number of PIs. We separate out the top 20 schools as schools which receive the most funding. We take the top 20 schools because the distribution of scientific activity is highly skewed with a relatively small number of schools responsible for most of the research grant money and most of the research output. Moreover, faculty researchers in the top 20 universities are intensely engaged in mentoring

PhD students and postdocs while simultaneously engaged in research.

Figure 1 panel (a) shows NIH expenditures per year as derived from NIH RePorter data. It clearly shows a steady increase during the doubling period and the sharp, temporary increase under the ARRA stimulus. Notably, the data suggest that top 20 schools did not gain a disproportionate advantage during the doubling—in fact, their share of NIH funding declined slightly during this period. This pattern suggests that the increased funding was spread relatively broadly across institutions rather than being concentrated in the most well-funded universities. However, after the doubling ended, the funding distribution shifted, and by the time of ARRA, the top 20 schools had regained and even increased their share of NIH funding. This reversal suggests that while the doubling initially supported a more decentralized allocation of resources, subsequent funding cycles reinforced the concentration of research funding among well-established institutions.

The data suggest that the expansion of biomedical research during these funding surges was achieved primarily by increasing the number of PhD students, rather than by hiring more postdocs.Figure 1 panel (b) shows the number of full time PhD students in the life sciences increases shortly after the doubling begins, particularly in top 20 schools. This increase lags the budget increase slightly, consistent with the time required to recruit and train new students. We can also observe a spike in doctoral students during Covid probably caused by potential graduates delaying their graduation because of the state of the job market. In contrast, Panel (c) shows that the number of postdocs experienced no change in the trend during the doubling, suggesting that PIs primarily staffed the additional research with graduate students rather than postdoctoral researchers. However, during the ARRA period, postdoc appointments show a pronounced spike, consistent with the expectation that short-term funding increases are more likely to result in hiring postdocs or research staff under short-term contracts rather than committing to multi-year PHD training. PhD students not only require a longer commitment but are often ineligible for intensive research in a lab until after coursework.

Figure 1 Panel (d) shows a delayed increase in the number of distinct PIs receiving NIH funding during both the doubling and ARRA periods. The rise in distinct PIs, coupled with a decline in the proportion of PIs from top 20 programs, suggest that funding increases were not disproportionately allocated to well-established, well-funded researchers and institutions. This pattern runs counter to the expectation of a Matthew effect in funding allocation (Bol, De Vaan and van de Rijt (2018)). However, during and after ARRA period, the share of funding claimed by top 20 programs rose sharply, surpassing pre-doubling levels.

Figure 1 Panel (e) shows that the ratio of student, postdoc, or graduation per PI. The overall students-per-PI ratio shows a subdued but notable shift, with the decline in the ratio of doctoral students per NIH-funded PI reversed midway through the doubling, rising modestly through 2008. This suggests that while the increase in new PIs helped distribute the mentoring burden more broadly, established PIs still took on more students as the period progressed. At top universities, the pattern is similar but more pronounced. Panel (f) shows that at the top 20 institutions, the number

of students per PI also increased midway through the doubling, reversing the prior downward trend. However, these institutions exhibit a distinct response in their reliance on postdocs: top schools consistently have more postdocs per PI, and this trend accelerates sharply during the ARRA period.

The descriptive trends suggest that funding increases were accompanied by adjustments in the research workforce, with evidence that the mentoring burden on PIs increased. During the doubling, the number of PhD students per PI rose, reversing a prior decline, while postdoc numbers remained relatively stable. Under ARRA, postdoc appointments increased sharply, particularly at top universities, likely reflecting a preference for hiring more immediately productive researchers in response to short-term funding boosts.

These figures have limitations. Graduate student counts are not necessarily tied to NIH-funded PIs, as they come from university-level sources rather than NIH grant records, making direct comparisons between PIs and student or postdoc numbers imprecise. They imprecisely show how expansions play out in the lab size and composition. This pattern is consistent with evidence from Sattari et al., who find that additional funding tends to expand research teams by adding non-PhD student and -postdoc persons, such as professional research staff undergraduates more than postdocs and graduate students.

To better understand how funding shocks shaped research training and careers, we turn to PI-level data linking grants, publications, and student outcomes. Increased mentoring demands may have weakened training, leading to worse outcomes for doctoral students. Alternatively, expanded funding may have improved training by providing access to better equipment, larger-scale projects, or more research experience. However, it could also have shifted students toward marginal projects with lower long-term returns, reshaping the scientific workforce in unexpected ways.

Our multivariate analysis directly tests these possibilities by examining how funding shocks at the PI level influence research output, student outcomes, and career trajectories. We now turn to this analysis.

III Data

We are combining three main datasets as shown in Figure 2. First, we use a private disambiguated dataset from ProQuest obtained by Thorvik and Smalheiser containing thesis information such as publisher's name (advisor and advisee), university, field, and advisor with each individual being given an ID. This dataset is from 1867 to 2013 and contains 1.2 million observations. We drop non-doctorate diplomas and errors coming from the disambiguation. This yields a sample of 458,955 person-year observations, representing all individuals who appear in the ProQuest data as the primary advisor on a graduating PhD student's thesis at least once between 1867 and 2013.

Second, we extracted data from the NIH's ExPORTER about funding cost, sorted by Institute Code (IC), categorized by project number, and as giving information about PIs. The range is

from 1985 to 2008 and we obtain 1.5 million initial observations. After selecting only the relevant activity (R01 equivalent), we end up with 604 804 observations. We then link the two datasets using an intermediary set (Authorlink by Thorvik) containing both NIH's project numbers and the unique ID which allows for a smooth matching obtaining 832 495 observations.

Finally, we gathered data from Author-ity (Torvik, Weeber, Swanson and Smalheiser (2005), Torvik and Smalheiser (2009)) containing publication information for the biomedical fields for a total of 40 million publication between 1945 and 2013. We also combine these data with Yu, Marschke, Ross, Staudt and Weinberg (2023)'s data to be able to use their metrics of disciplinary diversity for both citations and references, as well as a metric that detects whether a paper mentions old or new concepts. We will be using those metric as additional measure of the quality of research, in addition of the sheer number of forward citations.

After collapsing at the individual-year level, we can then merge this dataset with our two previous ones obtaining a final dataset of 21.7 million observations. Moreover, by combining the data we have from ProQuest and Author-ity we can obtain publishing information on advisees per advisors. We then only keep observation within the common range (1985 to 2008) and drop individuals who have never had advisors or publications (we keep individuals who have never received grants) while expanding for missing years within an individual's career and account for inactivity at the beginning and at the end of each advisor's career. Additionally, we drop individuals with production (advisee, publication or grant) on fewer than 10 years during our sample. We end up with 1,135,640 individual-year observations.

The descriptive statistics are detailed in Table 1. All the observations are at the advisor-year level and a short description of the variables can be found in Table 2. Forward citations are measured five years after publications to avoid any truncation problems (the dataset containing publication information, Author-ity, goes to 2013). Fractioned variables are the ratio of the variable of choice with the team size working on the project. Student quality observations are measured at the advisor level from the time of graduation to five years afterward, e.g., a student graduating in 2000 will appear on his advisor's observations for the year 2000 and will count all the papers published by said student between 2000 and 2005. Mean team size is the average of authors per paper per year for each advisor, while lab size represents the total amount of collaborator per year for each advisor (a distinction is that lab size do abstraction of multiple collaboration with the same author throughout a year). As we can see, the number of observations greatly decreases for our study of student's quality which could lead to a low-level of predictive power for our analysis.

We also repeat the same process keeping only the 20 universities which received the most funding. The top 20 universities are: Johns Hopkins University, University of California San Francisco, University of Michigan, University of Pittsburgh, University of Pennsylvania, Columbia University, Yale University, Washington University, Stanford University, University of California San Diego, University of Washington, Duke University, New York University, University of North Carolina Chapel Hill, University of California Los Angeles, Emory University, University of Minnesota,

University of Wisconsin Madison, Northwestern University, Harvard University. ProQuest has information about the institution of the advisor of a thesis. As such, we can identify advisors who works for such establishments.

IV Model

We investigate how unexpected increases in PI funding affect (1) PI research output and labor organization, including the number of publications, team size, and collaboration structure, and (2) the development and career outcomes of PhD students trained under PIs experiencing funding shocks. These two dimensions are interrelated: how PIs scale up their lab operations in response to additional funding affects the training environment for PhD students.

In response to a positive shock, we expect PIs to increase research output by publishing more papers, expand their research teams by hiring more PhD students and postdocs, and collaborate with more co-authors. If these additional publications involve newly recruited PhDs, postdocs, or more senior collaborators—or if new PhDs and postdocs are not yet fully integrated into the research process—publication quality may decline

There is a timing issue: PIs cannot immediately bring on new research personnel, as funding must first be secured, and then time is required to recruit and integrate them into the lab. For example, PhD students must complete coursework before becoming fully engaged in research, and postdoc hiring requires identifying and vetting qualified candidates. In the meantime, the PI's research output may change in ways that affect quality. This could happen because the PI is stretched across more papers, or because the marginal research projects funded under the shock are of lower quality. We measure these effects using total publications per PI and total citation-weighted publications (overall productivity), number of co-authors per PI (team or lab size), and co-authors per paper (how collaboration is structured). This allows us to confirm that the funding shock expands research activity while also examining whether it reshapes how research is structured and produced in the lab, before turning to its effects on student careers.

When funding increases, a PI must expand their lab, often by hiring more graduate students or postdoctoral researchers. This creates a potential mentorship trade-off. More PhD students and postdocs mean mentorship is spread thinner, reducing per-student investment and leading to lower-quality training. If PIs need to scale up quickly, they may lower admission standards, bringing in weaker students, which could reduce research output. If the PI expands into lower-value projects, students may work on less impactful research, affecting their long-term success. We test these hypotheses by examining the number of PhD students graduating under the PI, student publication and citation records and post-graduation funding success (early-career productivity).

To test those hypothesis, we estimate the effect of a positive funding shock on research outcomes using a difference-in-difference model with PI and year fixed effects:

$$y_{i,t+n_j} = \beta \text{Shock}_{i,t} + \gamma_i + \delta_t + \epsilon_{i,t} \tag{1}$$

where i indexes the PI, the period (whether yearly or ranges of years) is denoted by t, and y represents PI research output or PhD student outcomes. Fixed effects, γ and δ account for PI-specific ability and broader funding trends, respectively. Since it will take time from the moment a laboratory receives funding and when the output is effectively produced, we include an effect n_j which will lag the dependent variable, with $j \in \{1, 2\}$, $n_1 = 1$, and $n_2 = 2$, differentiating between short-term and slightly delayed effects of funding shocks. Variables related to PI production (e.g., publications, citations) are expected to respond more immediately.[2] Student productivity outcomes take longer to manifest because newly recruited PhD students typically take several years to complete their research and graduate. As a result, we expect the effects of a funding shock on student career outcomes to emerge only after a lag of a few years. We run this analysis for both the overall sample and within the top 20 universities only.

The key independent variable, Shock, is an indicator of whether the PI receives an unexpected, substantial increase in funding. We use a binary indicator instead of the level of funding to capture cases where the PI must rapidly scale up their lab in a way that is not easily adjustable. If we used funding levels as a continuous variable, the model would capture gradual adaptations, which PIs can often manage smoothly. Instead, we focus on cases where the scale of the PI's operation must expand quickly, forcing them to hire additional personnel and restructure their research efforts—exactly the conditions under which the constraints we describe above should bind. We now need to define what a shock is.

1. NIH Shock: NIH's Institute Centers (ICs) all saw significant increase in their fundings around the same period, some ICs getting shocked slightly earlier and some later. Instead of sorting between ICs receiving different levels of raises, we separate the control group as individuals who do not receive any sort of funding or receive most of the funding before getting shocked (the current threshold is 75%) and the treatment group as researchers who receive a significant amount of their funding while the IC is getting shocked.

2. Individual Shock: PIs need sustained grant funding to run their labs, but because NIH proposals are often rejected, they submit more proposals than necessary. Occasionally, multiple proposals are unexpectedly funded, creating a positive funding shock that forces rapid expansion. At the individual level, a researcher is considered shocked if the moving average of his funding increases by more than 80% from one year to the next. Therefore, the control group is made of individuals who have never had big burst in budget, but more likely received

[2]When a PI experiences a funding shock, there is typically a delay before new PhD students are recruited and trained. During this transition period, research output may still increase as the PI may shift their publishing strategy to increase publication counts, for example, by splitting the same research findings into several smaller papers of lesser quality. This would allow them to increase their publication rate even before new students are in place and productive, consistent with incentives created by funding agencies.

their fundings more regularly.

The time distribution of the first period of shock greatly differs between our two definitions as shown in Figure 3. Since IC shocks mostly happen during the 1998-2003 period, initial period of shocks will be condensed through the same period. For individual shocks, it seems to follow a normal distribution centered around 2000, since individual shocks will also arise based on career experience even without budget increase from the funding agency.

We also want to test whether an individual being shocked affect the output of the rest of its department, granted they have not been shocked yet. We define a department as a combination of field and university. For example, a researcher publishing from Boston University in the field of Biology is considered to belong in that specific department. Due to that broad definition, one individual can belong into more than one department. Then, a department is considered shocked if at least one individual in the department has been shocked, following the previous definitions of shocks. We only use our IC definition of shock for this analysis. For an individual i who has not been shocked yet and a department k, we have:

$$y_{i,t+n_j,k} = \beta \text{Shock}_{k,t} + \gamma_i + \delta_t + \epsilon_{i,t,k} \tag{2}$$

To better understand the dynamics before and after such shocks, we also conduct an event study analysis:

$$y_{i,t+n_j} = \sum_{\tau=-5}^{5} \beta_\tau \mathbb{I}\{t - T_i = \tau\} + \gamma_i + \delta_t + \epsilon_{i,t} \tag{3}$$

The indicator $\mathbb{I}$ serves as a dummy around the first period an individual i is being shocked T_i. This allows us to observe the behavior of the dependent variables around the shock, 5 years prior and 5 years later. With an event analysis a relevant lag selection has a lesser importance, but helps move the window to obtain a frame that is interesting to observe.

V Results

The findings from equation (1) are presented in Table 3. Notably, although the two shock definitions yield the same directional effects across all variables, their statistical significance differs. Consistent with previous literature, a positive funding shock appears to correlate with increased production. Specifically, advisors experience a rise in publication rate—approximately 0.09 additional publications per year—and an increase in the graduation rate of their advisees by about 0.02. With a higher budget, it is plausible that advisors are better equipped to undertake additional projects concurrently. However, the dynamics of team composition also change: the average team size per paper decreases by 0.29 individuals, which suggests that principal investigators (PIs) may be redistributing their available workforce across a greater number of projects, rather than merely expanding the size of individual projects. Concurrently, the overall lab size increases by

roughly 1.22 individuals, indicating a broader search for diverse collaborative opportunities among researchers.

A critical concern associated with this production increase is the apparent decline in the quality of output. Both forward citations per paper and fractioned citations per paper drop—by 1.71 and 0.67 citations respectively. This decline suggests that while increased funding enables PIs to produce more publications, the attention each project receives may be diluted. One possible interpretation is that the elevated production pace leaves less time for refining each project, or alternatively, projects might be subdivided into smaller, less impactful units. Moreover, the reduction in the average fractioned publications per paper reinforces the idea that, despite larger lab teams, the per-project quality declines. Similarly, the use of new concepts appears to be declining after the doubling, to the benefit of older concepts. This is an indication of less innovative research. However, the doubling appears to increase the diversification of the citations, with references more narrow and forward citations spreading out to other fields.

Turning to Table 4, the effects on students are also noteworthy. Although students publish more and graduate at higher rates during periods of increased funding, the average impact of their publications—as measured by forward citations—drops by approximately 2.46 citations per paper. Same as their advisors, students formed during the doubling tend to use fewer new concepts and cite less diverse research. This downward trend in quality is further supported by other indicators: both the percentage of students publishing and the percentage receiving funding decline by 4% and 0.9% respectively. These findings resonate with Tham et al. (2024), who show that funding gaps push researchers—especially early-career personnel—out of U.S. science and depress their earnings, and with Babina et al. (2023) who find that negative funding shocks reduce high-quality, basic publications. Our contribution is to show that even positive funding shocks, which expand labs and temporarily increase output, can have adverse consequences for the next generation of scientists trained in those labs.

One plausible explanation for the poorer student outcomes is that the increased workload imposed on advisors leaves less time for individualized mentorship. Alternatively, it could be that an influx of applicants, driven by a broader recruitment push in response to funding increases, results in the admission of students with relatively lower qualifications (as noted by Akcigit, Pearce and Prato (2024)). If the decline in quality is due to a limited pool of high-ability young researchers, then other departments that do not experience such funding surges should also exhibit similar reductions in student performance. In contrast, if the quality loss results primarily from diminished advisor engagement, then departments unaffected by funding shocks should maintain their student outcomes. To resolve this ambiguity, we conduct a departmental-level analysis of the impact of funding shocks on student performance.

Table 5 summarizes the results of this departmental analysis that results from equation (2). In general, the effects are statistically insignificant, which suggests that having another researcher from the same department receive a significant increase in funding does not lead to having students

of worse quality. This outcome points toward a scenario where the loss in performance is predominantly attributable to the dispersion of advisors' attention over more tasks, rather than an inherent drop in the quality of the students recruited. A potential decline in student funding, aligning with the observations reported by Larson et al. that younger researchers face more challenges in securing federal funding, was also observed. However, the limited size of the control group in this analysis—primarily due to the prevalence of funding increases in most departments—complicates the interpretation, particularly as departments loosely connected to the biomedical field (e.g., psychology) remain less affected, possibly introducing endogeneity issues. Expanding this analysis to include more departments within the same field could yield further insights.

A similar pattern emerges when we focus on the top 20 universities, as shown in Tables 6 and 7. Although the overall trends mirror those observed in the full sample, statistical significance diminishes—likely a result of the smaller sample size and the increased data requirements of our methodological approach.

Given the variability in lag structures depending on the dependent variable under consideration—and the absence of a universally accepted method to determine an optimal lag—it is standard practice to complement regression analyses with a panel event study. Such a study examines the effect on dependent variables over a continuous range of time surrounding the event, thereby providing a more nuanced view of whether the observed effects are attributable to a specific event or merely represent a trend.

Figure 4 and 5 depicts the event study results, from equation (3). For advisees, a moderate pre-event trend is evident; however, following the funding shock, there is a marked increase in advisee count followed by a gradual reversion to baseline levels. A parallel pattern is observed for advisors' publication counts. In contrast, forward citations per PI paper initially show an inconsistent pattern but transition to a pronounced downward trend post-event, with no observable rebound. Mean team size and lab size display distinct trends—downward and upward respectively—with both metrics beginning to converge towards pre-shock values approximately five years after the funding event.

For student-specific metrics, the analysis is more challenging due to the smaller sample size and the inherent variability in graduation timelines, which complicates the estimation of an appropriate lag. Despite these limitations, a delayed downward trend is observable across the student variables, implying that our lag selection may be underestimating the time required for the full effect to manifest.

VI Conclusion

The conventional wisdom that "the more, the better" does not always apply to research funding. Between 1998 and 2003, the National Institutes of Health (NIH) underwent a significant funding increase, an expansion that many anticipated would translate into improved research outcomes in

the biomedical sciences—namely, higher-quality publications and stronger mentorship for emerging researchers. However, our analysis challenges this assumption.

By combining three large national dataset- ProQuest, Author-ity, and ExPorter - we assessed the impact of NIH funding increases through a staggered difference-in-differences model. This method allowed us to evaluate a broad spectrum of outcomes, including principal investigators' (PIs) publication productivity and citation impact, changes in team composition, and the training and early career trajectories of their mentees. Unlike prior studies that have predominantly concentrated on the direct effects of funding surges on PI output, our research uniquely examines how these fiscal shocks influence the formation and subsequent development of PhD students who are being trained during periods of expansive funding.

Our results reveal that, although an increase in funding corresponds with higher publication counts and a greater number of graduations, these gains come at a cost. Specifically, the publications produced in this environment tend to attract fewer citations, suggesting a compromise in research impact or visibility. Moreover, the early career prospects for new researchers suffer. PhD graduates who trained under PIs benefitting from system-wide funding surges tend to face a more competitive and saturated job market, which subsequently dampens their research productivity. Even when funding shocks occur on an idiosyncratic basis—without generating broad market-wide pressures—the negative effect on post-graduation productivity persists, potentially due to factors such as reduced mentorship quality.

In summary, our findings indicate that funding volatility—whether experienced across the system or at the level of individual PIs—can have pronounced adverse effects on the training and career development of early-career researchers. Our results complement recent evidence that funding interruptions and negative shocks can permanently alter who stays in science and what kind of science is produced (Babina et al.; Tham et al.), by showing that large, temporary increases in funding can also degrade training quality and early-career outcomes. These results carry significant policy implications for funding agencies. Specifically, they underscore the potential benefits of adopting measures that favor more stable and predictable funding trajectories. Reducing extreme fluctuations in financial resources could help mitigate disruptions, thereby supporting both the quality of research outputs and the professional growth of the next generation of scientists.

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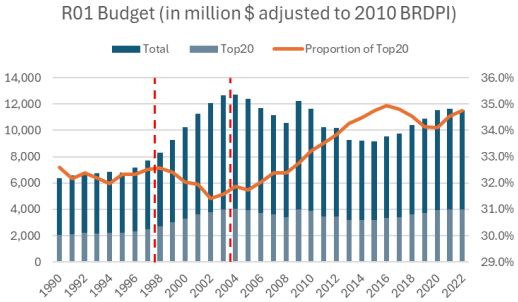
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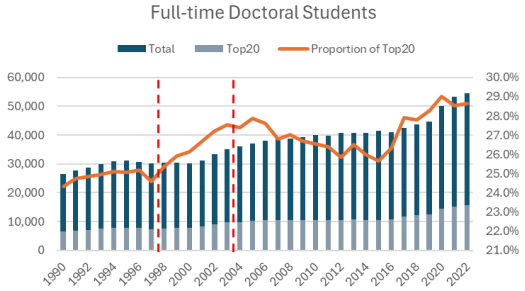
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Figure 1: Budget and Labor Supply

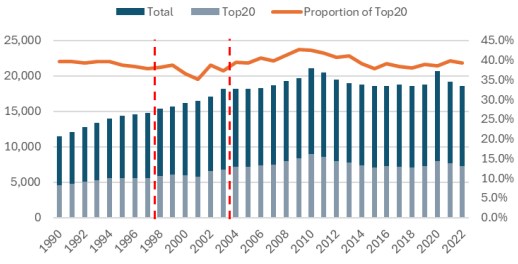
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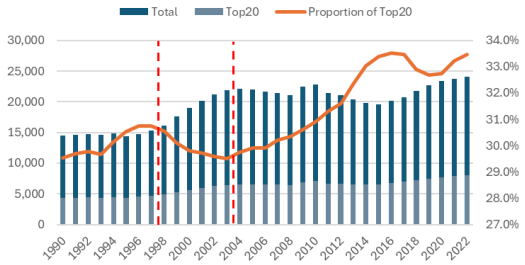
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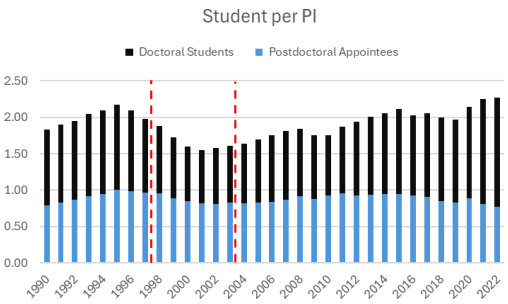
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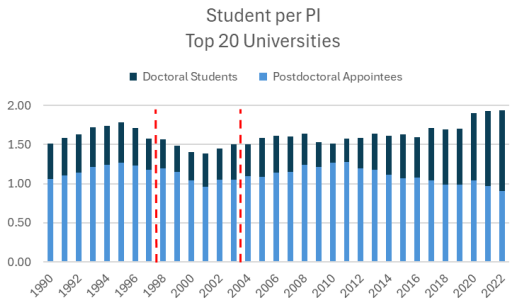


Figure 2: Data Combination Flow Chart

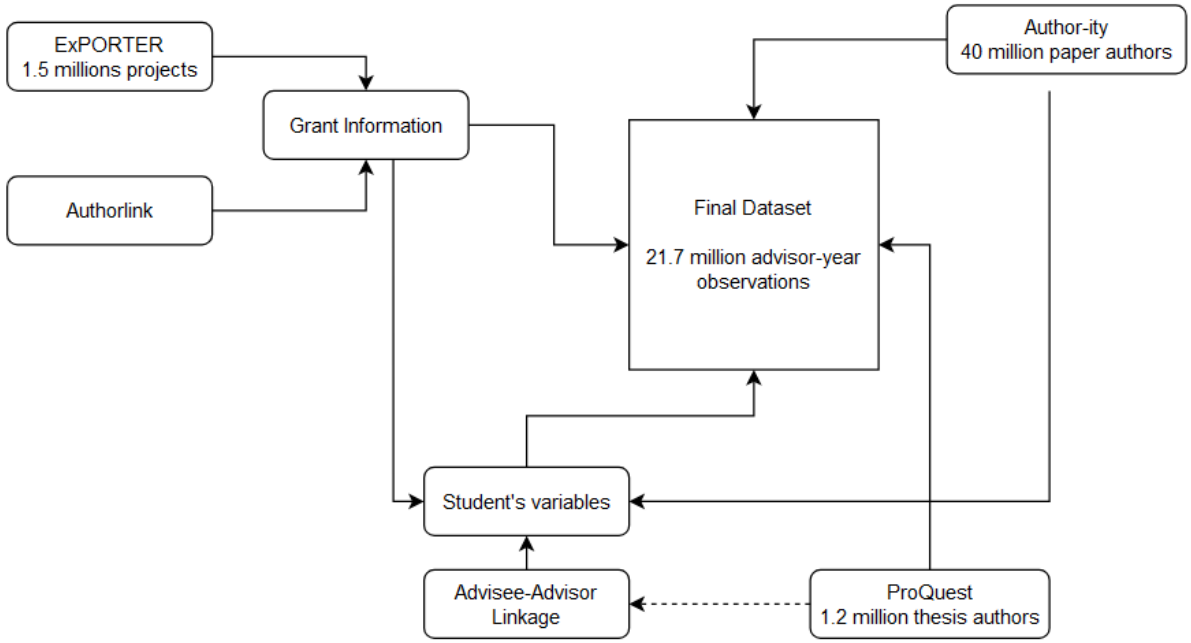
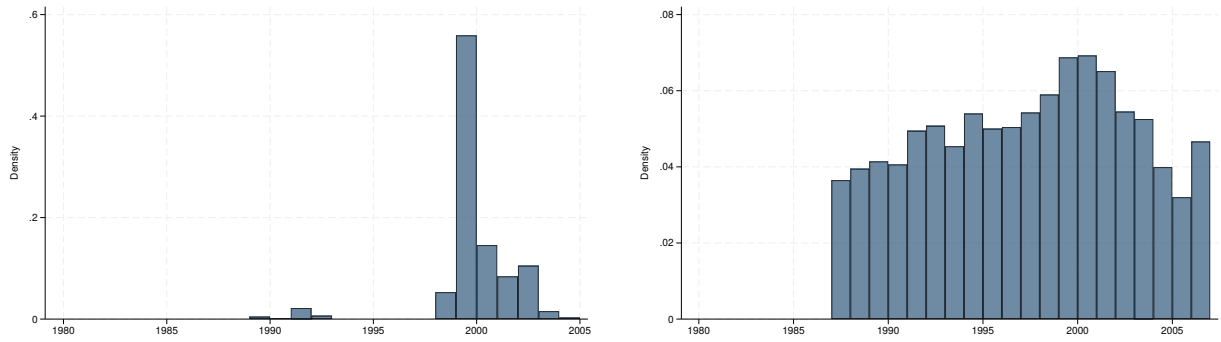


Figure 3: Distribution of shocks



(a) NIH Shocks

(b) Individual Shocks

Table 1: Descriptive Statistics

Panel A: PI-level Variables	**Obs.**	**Mean**	**Std. Dev.**	**Min**	**Max**
Advisee Count	1,135,640	0.288	0.645	0	25
Publication Count	1,135,640	2.018	2.993	0	142
Fractioned Publications	1,135,640	0.567	0.835	0	21.9996
Fractioned Publications per Paper	707,798	0.300	0.162	0.001	1
Forward Citations	707,798	83.140	180.440	1	11,199
Forward Citations per Paper	707,798	21.161	35.761	1	5,157
Fractioned Forward Citations	707,798	20.516	45.886	0.002	3,518.617
Fractioned Forward Citations per Paper	707,798	5.535	9.231	0.002	820.167
Maximum Forward Citations	707,798	40.594	88.604	1	10,967
New Concepts	615,941	1.405	3.580	0	113
Average New Concepts	615,941	0.363	0.797	0	18
Old Concepts	615,941	56.893	69.427	0	1,721
Average Old Concepts	615,941	16.074	10.199	0	98
Diversity of References	601,074	951.286	32.482	0	987.578
Diversity of Citations	611,441	951.840	26.037	0	989.196
Mean Team Size	707,798	5.861	22.584	1	743
Lab Size	707,798	12.985	30.346	1	1,340
Panel B: Student-level Variables					
Publication Count	245,759	3.031	5.762	0	295
Publications per Student	245,759	2.388	4.463	0	284
Fractioned Publications	245,759	0.757	1.360	0	56.002
Fractioned Publications per Student	245,759	0.599	1.068	0	56.002
Forward Citations	139,037	134.063	283.986	1	10,369
Forward Citations per Paper per Student	139,037	17.604	26.874	0.050	1,789
Fractioned Forward Citations	245,759	15.966	42.871	0	2,536.402
Fractioned Forward Citations per Student	245,759	2.241	4.502	0	397.5
Maximum Citations	139,037	48.744	104.777	1	7,261
New Concepts	127,631	2.156	5.283	0	146
Average New Concepts	127,631	0.336	0.667	0	16
Old Concepts	127,631	96.902	139.682	0	5,732
Average Old Concepts	127,631	17.064	10.318	0	94
Diversity of References	123,939	951.265	34.645	0	986.305
Diversity of Citations	126,284	949.814	28.030	0	986.361
Amount of Publishing Students	245,759	0.679	0.708	0	8
Fraction of Publishing Students	245,759	0.530	0.482	0	1
Amount of Funding from Students	245,759	6,697.535	127,454.1	0	13,600,000
Fraction of Students Receiving Funding	245,759	0.006	0.070	0	1

Table 2: Variables

Index	**Variable Name**	**Description**	**Formula**
A	Advisee Count	Total number of students graduating per PI each year, measured by counting the number of doctoral dissertations (Masters excluded) for which the PI is listed as advisor.	
B	Publication Count	Total number of publications per PI per year for which the PI appears as any author.	
C	Fractioned Publications	Total amount of fractioned publications per PI per year. A fraction publication is obtained by dividing 1 by the number of authors for each paper the PI has appeared as an author. (Example: A PI who publishes two papers, one with 4 authors and another with 2 authors, will have $\frac{1}{4} + \frac{1}{2} = 0.75$).	
D	Fractioned Publications Per Paper	The average fractioned publications per paper for each PI and each year.	C/B
E	Forward Citations	Total amount of forward citations per PI per year. To avoid recency bias, forward citations are registered 5 years after publication.	
F	Forward Citations Per Paper	The average forward citations per paper published for each PI and each year.	E/B
G	Fractioned Citations	Total amount of fractioned citations per PI per year. A fraction citation is obtained by dividing the number of forward citations (5 years after publication) by the number of authors for each paper the PI has appeared as an author. (Example: A paper with 10 citations and 4 authors contributes $\frac{10}{4} = 2.5$).	
H	Fractioned Citations Per Paper	The average fractioned forward citations per paper published for each PI and each year.	G/B
I	New Concepts	The number of new concepts used by a given paper using text analysis. This variable is then aggregated for each advisor-year.	
K	Average New Concepts	The number of new concepts averaged across the number of paper published for each advisor-year.	I/B

J	Old Concepts	The number of old concepts used by a given paper using text analysis. This variable is then aggregated for each advisor-year.	
L	Average Old Concepts	The number of old concepts averaged across the number of paper published for each advisor-year.	J/B
M	Diversity of References	A metric to measure how diverse the papers being cited are for a given paper. The metric is then averaged across publications for each advisor-year.	
N	Diversity of Citations	A metric to measure how diverse the forward citations are for a given paper. The metric is then averaged across publications for each advisor-year.	
O	Mean Team Size	The average of the number of authors over the number of publications per PI for a given year.	
P	Lab Size	Total number of unique collaborators per PI per year measured by using common authorship of publications.	
Q	Amount of Publishing Students	The number of students who are counted as graduating under a PI (per A) who also publishes within 5 years after graduation. Observations are made at the PI-year level (the year being the year of graduation, not publication).	
R	Fraction of Publishing Students	The fraction of publishing students per PI per year.	Q/A
S	Amount of Funding from Students	Total amount of funding (in adjusted \$) received by students graduating under a PI (per A) within 5 years of graduation. Observations are made at the PI-year level (the year being the year of graduation, not publication).	
T	Fraction of Students Receiving Funding	The fraction of students who receive funding within 5 years of graduation.	S/A
V	Student's Forward Citations Per Student Per Paper	The average of forward citations for students who publish within 5 years of graduation. Observations are at the PI-year level and averaged over publications and students.	

Table 3: Regression Results for PI variables

	(1) IC		(2) Individual	
Variables	Coeff. (S.E.)	N° obs.	Coeff. (S.E.)	N° obs.
Advisee Count	0.023***	1,135,640	0.065***	1,135,769
	(0.006)		(0.005)	
Publication Count	0.094***	1,135,640	0.487***	1,135,769
	(0.026)		(0.025)	
Fractioned Publications per Paper	-0.003**	707,798	-0.005***	707,892
	(0.001)		(0.002)	
Forward Citations per Paper	-1.712***	707,798	-2.808***	707,892
	(0.333)		(0.410)	
Fractioned Forward Citations per Paper	-0.673***	707,798	-0.893***	707,892
	(0.098)		(0.113)	
Maximum Forward Citations	0.198	707,798	1.465	707,798
	(0.972)		(0.916)	
New Concepts	-0.497***	615,941	-0.054	615,995
	(0.053)		(0.049)	
Average New Concepts	-0.099***	615,941	-0.112***	615,995
	(0.010)		(0.011)	
Old Concepts	4.208***	615,941	12.938***	615,995
	(0.798)		(0.752)	
Average Old Concepts	-0.152*	615,941	0.007	615,995
	(0.091)		(0.090)	
Diversity of References	-1.628***	601,074	-1.446***	601,128
	(0.178)		(0.170)	
Diversity of Citations	0.845***	611,441	-0.030	611,495
	(0.154)		(0.145)	

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Regression Results for PI variables (continued)

	(1) IC		(2) Individual	
Variables	Coeff. (S.E.)	N° obs.	Coeff. (S.E.)	N° obs.
Mean Team Size	-0.286***	707,798	-0.136***	707,892
	(0.053)		(0.035)	
Lab Size	1.220***	707,798	2.221***	707,892
	(0.175)		(0.158)	
Individuals in Control Group		49,134		45,999
Individuals in Treatment Group		12,914		16,058

Regression Results: Contains the ATT estimates of the stacked difference-in-difference for our dependent variables of choice regarding PIs. Column (1) uses our IC definition of shock, i.e., the control group contains individuals that either receive no funding or receive funding from IC when they are not being shocked. Column (2) uses our definition of shock at the individual level, among the productive researchers of 20 years or more of experience. A shock is defined as a growth of 50% from its moving average of 3 years. Only PIs with at least one publication during a year are counted for variables related to publications, explaining the decline in the number of observations. Robust standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Table 4: Regression Results for Student variables

	(1) IC		(2) Individual	
Variables	Coeff. (S.E.)	N° obs.	Coeff. (S.E.)	N° obs.
Publications per Student	-0.695***	245,759	-0.316***	245,759
	(0.114)		(0.099)	
Fractioned Publications per Student	-0.190***	245,759	-0.119***	245,759
	(0.028)		(0.027)	
Forward Citations per Paper per Student	-2.464***	139,037	-0.712	139,037
	(0.773)		(0.736)	
Fractioned Forward Citations per Student	-0.795***	245,759	-0.372***	245,759
	(0.151)		(0.137)	
Maximum Citations	-6.837*	139,037	-2.884	139,037
	(3.962)		(5.275)	
New Concepts	-0.641***	127,631	-0.193	127,631
	(0.170)		(0.214)	
Average New Concepts	-0.058***	127,631	-0.043*	127,631
	(0.019)		(0.024)	
Old Concepts	-9.841**	127,631	1.765	127,631
	(4.622)		(4.341)	
Average Old Concepts	-0.422*	127,631	0.376	127,631
	(0.216)		(0.242)	
Diversity of References	-2.513***	123,939	-0.736	123,939
	(0.829)		(0.996)	
Diversity of Citations	0.682	126,284	0.559	126,284
	(0.604)		(0.452)	

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Regression Results for Student variables (continued)

	(1) IC		(2) Individual	
Variables	Coeff. (S.E.)	N° obs.	Coeff. (S.E.)	N° obs.
Amount of Publishing Students	-0.040**	245,759	-0.001	245,759
	(0.019)		(0.020)	
Fraction of Publishing Students	-0.040***	245,759	-0.030***	245,759
	(0.008)		(0.010)	
Amount of Funding from Students	-12520**	245,759	-9201*	245,759
	(5378)		(5501)	
Fraction of Students Receiving Funding	-0.009***	245,759	-0.007**	245,759
	(0.003)		(0.003)	
Individuals in Control Group		49,134		45,999
Individuals in Treatment Group		12,914		16,058

Regression Results: Contains the ATT estimates of the stacked difference-in-difference for our dependent variables of choice regarding students. A student's production is counted within 5 years of graduation and is accounted at the PI level the year of the student's graduation. Column (1) uses our IC definition of shock, i.e., the control group contains PIs that either receive no funding or receive funding from IC when they are not being shocked. Column (2) uses our definition of shock at the PI level, among the productive researchers of 20 years or more of experience. A shock is defined as a growth of 50% from its moving average of 3 years. Only PIs with at least one student during a year are counted for variables related to publications, explaining the decline in the number of observations. Only PIs with at least 1 student who published within 5 years of graduation are counted for the average of student's quality. Robust standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Table 5: Departmental Shock Effects on Students

	(1) IC		(2) Individual	
Variables	Coeff. (S.E.)	N° obs.	Coeff. (S.E.)	N° obs.
Publications per Student	0.063	207,253	-0.015	195,983
	(0.080)		(0.024)	
Fractioned Publications per Student	-2.932*	207,253	-0.310**	195,983
	(1.595)		(0.153)	
Forward Citations per Paper per Student	-5.690	108,394	-0.240	100,330
	(4.022)		(0.244)	
Fractioned Forward Citations per Student	-0.019	207,253	14.157*	195,983
	(0.034)		(7.557)	
Maximum Citations	-0.745**	108,394	-4.925	100,330
	(0.371)		(4.104)	
New Concepts	0.570	98,388	0.032	90,693
	(1.763)		(0.019)	
Average New Concepts	0.012	98,388	1697.510	90,693
	(0.012)		(2781.843)	
Old Concepts	0.002	98,388	0.0016	90,693
	(0.002)		(0.002)	
Average Old Concepts	0.063	98,388	-0.015	90,693
	(0.080)		(0.024)	
Diversity of References	-2.932*	94,756	-0.310**	87,080
	(1.595)		(0.153)	
Diversity of Citations	-5.690	97,070	-0.240	89,382
	(4.022)		(0.244)	
Amount of Publishing Students	-0.005	207,253	0.032	195,983
	(0.013)		(0.019)	
Fraction of Publishing Students	-0.011	207,253	0.012	195,983
	(0.008)		(0.012)	
Amount of Funding from Students	-9351*	207,253	1698	195,983
	(5173)		(2789)	
Fraction of Students Receiving Funding	-0.004***	207,253	0.002	195,983
	(0.001)		(0.002)	

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Departmental Shock Effects on Students (continued)

	(1) IC		(2) Individual	
Variables	Coeff. (S.E.)	N° obs.	Coeff. (S.E.)	N° obs.
Individuals in Control Group		14,572		10,335
Individuals in Treatment Group		44,892		49,742

Table contains the ATT estimates of the stacked difference-in-difference for our dependent variables on student outcomes. A department is considered shocked if at least one individual during the year experienced a shock. Robust standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Table 6: Regression Results for PI variables, Top 20 Universities

	(1) IC		(2) Individual	
Variables	Coeff. (S.E.)	N° obs.	Coeff. (S.E.)	N° obs.
Advisee Count	-0.005	213,501	0.084***	213,501
	(0.013)		(0.009)	
Publication Count	0.113**	213,501	0.680***	213,501
	(0.055)		(0.057)	
Fractioned Publications per Paper	-0.004	150,302	-0.002	150,302
	(0.003)		(0.003)	
Forward Citations per Paper	-1.987**	150,302	-2.917***	150,302
	(0.797)		(1.022)	
Fractioned Forward Citations per Paper	-0.613***	150,302	-0.717***	150,302
	(0.221)		(0.267)	
Maximum Forward Citations	-1.662	150,302	5.342**	150,302
	(2.180)		(2.258)	
New Concepts	-0.602***	133,214	0.198*	133,214
	(0.119)		(0.120)	
Average New Concepts	-0.093***	133,214	-0.106***	133,214
	(0.018)		(0.021)	
Old Concepts	5.267***	133,214	15.994***	133,214
	(1.728)		(1.682)	
Average Old Concepts	-0.109	133,214	-0.055	133,214
	(0.174)		(0.166)	
Diversity of References	-1.438***	130,862	-0.930***	130,862
	(0.346)		(0.281)	
Diversity of Citations	0.677**	132,491	0.461*	132,491
	(0.287)		(0.235)	

Continued on next page

Regression Results for PI variables, Top 20 Universities (continued)

	(1) IC		(2) Individual	
Variables	Coeff. (S.E.)	N° obs.	Coeff. (S.E.)	N° obs.
Mean Team Size	-0.473***	150,302	-0.329***	150,302
	(0.176)		(0.102)	
Lab Size	1.031**	150,302	2.342***	150,302
	(0.411)		(0.387)	
Individuals in Control Group		7,059		6,607
Individuals in Treatment Group		4,138		4,591

Same as Table 3, but only keeping PIs from universities in our list of most productive establishments. Robust standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Table 7: Regression Results for Student variables, Top 20 Universities

	(1) IC		(2) Individual	
Variables	Coeff. (S.E.)	N° obs.	Coeff. (S.E.)	N° obs.
Publications per Student	-0.606***	49,321	-0.293***	49,321
	(0.137)		(0.106)	
Fractioned Publications per Student	-0.161***	49,321	-0.085***	49,321
	(0.036)		(0.032)	
Forward Citations per Paper per Student	-1.016***	31,447	-0.612	31,447
	(0.383)		(0.471)	
Fractioned Forward Citations per Student	-0.476***	49,321	-0.259**	49,321
	(0.104)		(0.112)	
Maximum Citations	-4.286	31,447	-1.918	31,447
	(2.794)		(3.025)	
New Concepts	-0.389***	29,049	-0.127	29,049
	(0.096)		(0.112)	
Average New Concepts	-0.048***	29,049	-0.031*	29,049
	(0.013)		(0.018)	
Old Concepts	-5.927**	29,049	0.894	29,049
	(2.763)		(2.511)	
Average Old Concepts	-0.311*	29,049	0.162	29,049
	(0.172)		(0.199)	
Diversity of References	-1.952***	28,317	-0.578	28,317
	(0.573)		(0.652)	
Diversity of Citations	0.454	28,783	0.283	28,783
	(0.398)		(0.327)	

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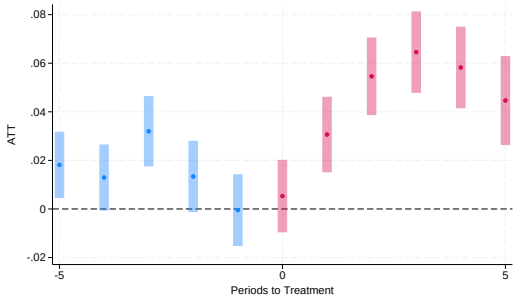
Regression Results for Student variables, Top 20 Universities (continued)

	(1) IC		(2) Individual	
Variables	Coeff. (S.E.)	N° obs.	Coeff. (S.E.)	N° obs.
Amount of Publishing Students	-0.086**	49,321	0.026	49,321
	(0.036)		(0.041)	
Fraction of Publishing Students	-0.042***	49,321	-0.003	49,321
	(0.015)		(0.020)	
Amount of Funding from Students	-14630	49,321	-15632	49,321
	(11614)		(14668)	
Fraction of Students Receiving Funding	-0.002	49,321	-0.004	49,321
	(0.005)		(0.006)	
Individuals in Control Group		7,059		6,607
Individuals in Treatment Group		4,138		4,591

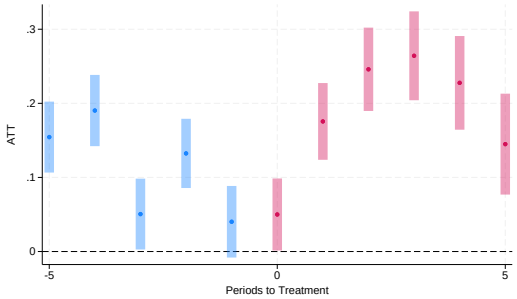
Same as Table 4, but only keeping PIs from universities in our list of most productive establishments. Robust standard errors in parentheses. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Figure 4: Panel Event Study

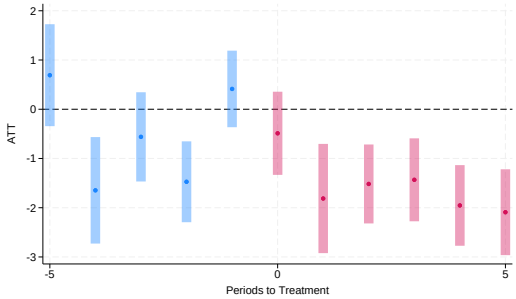
(a) Advisee



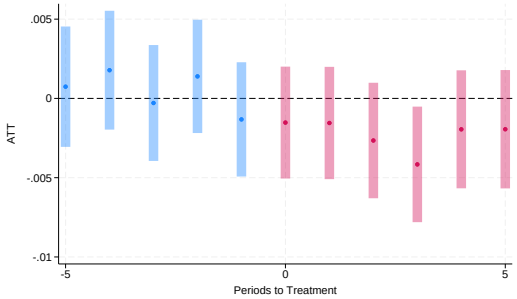
(b) Annual Publication Count



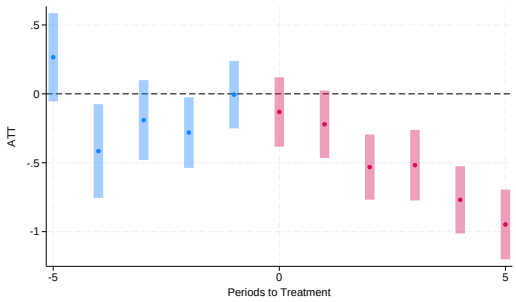
(c) Forward Citations per Paper



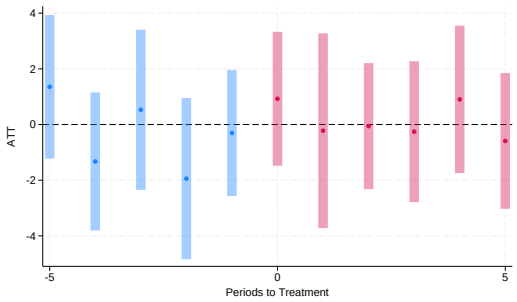
(d) Fractioned Publications per Paper



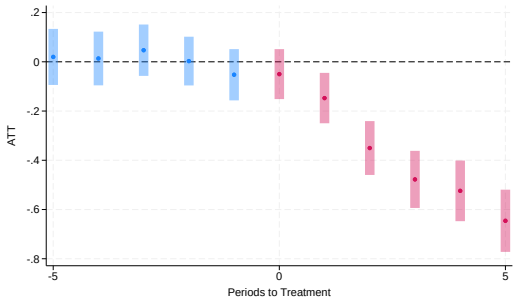
(e) Fractioned Citations per Paper



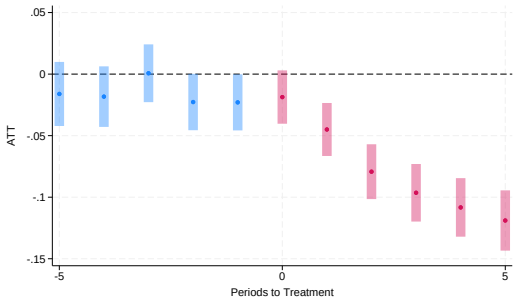
(f) Maximum Citations



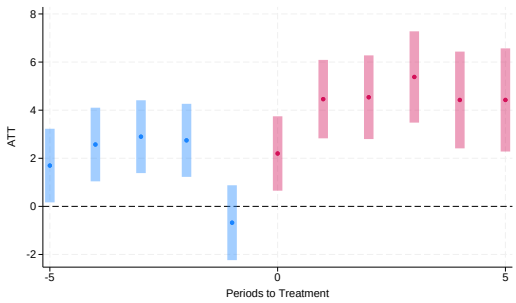
(g) New Concepts



(h) Average New Concepts



(i) Old Concepts



(j) Average Old Concepts

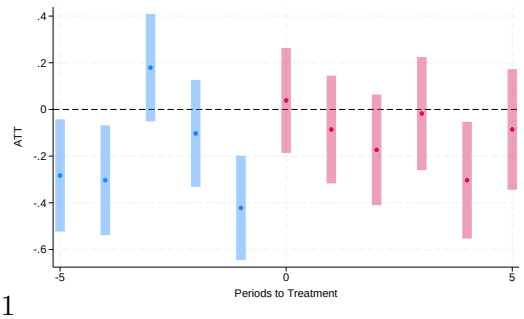
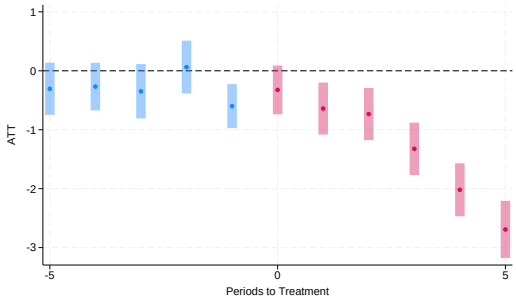
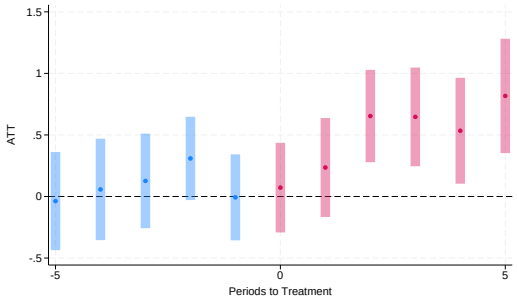


Figure 4: Panel Event Study (cont.)

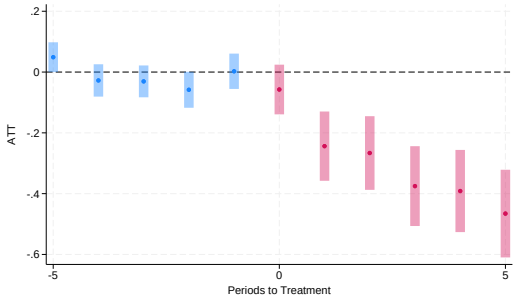
(k) Diversity of References



(l) Diversity of Citations



(m) Mean Team Size



(n) Lab Size

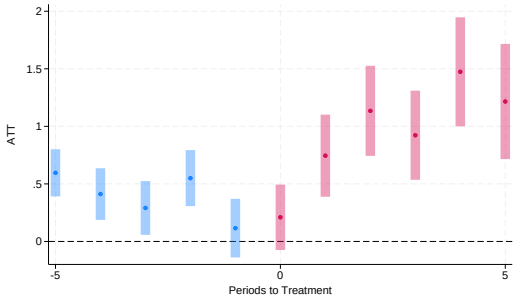
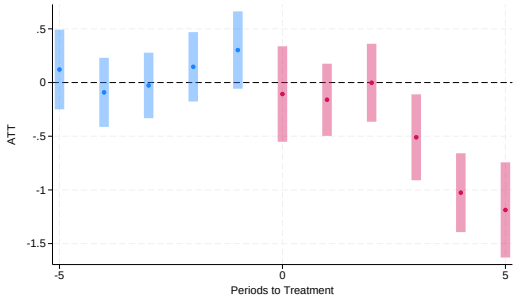
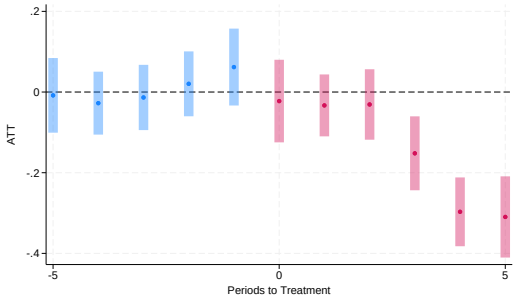


Figure 5: Panel Event Study (Students' Variables)

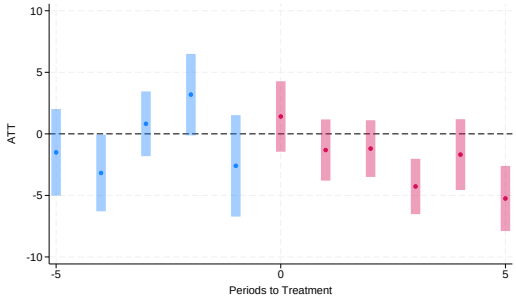
(a) Publication per Student



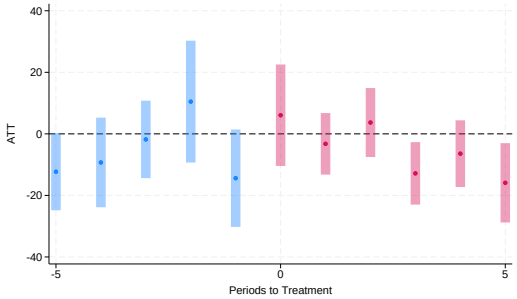
(b) Fractioned Publications per Student



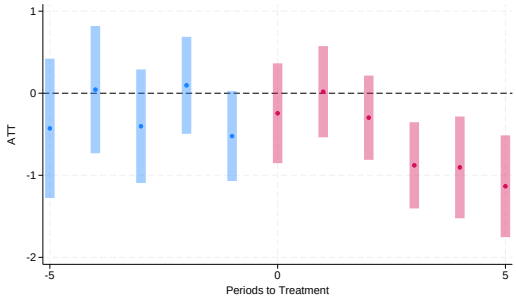
(c) Forward Citations per Paper per Student



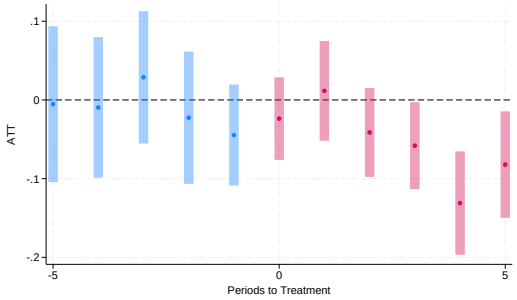
(d) Maximum Citations



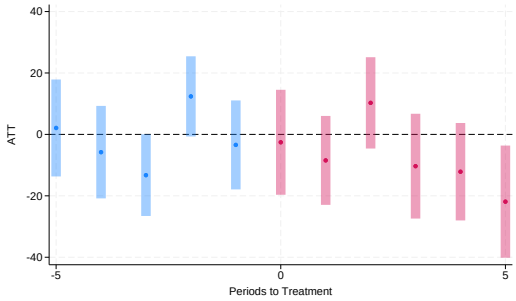
(e) New Concepts



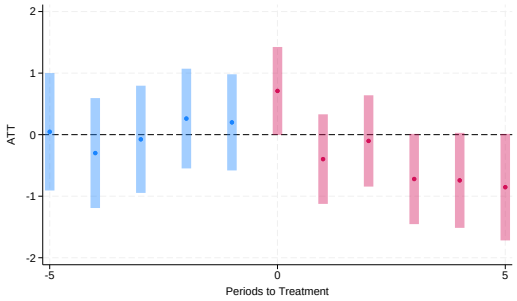
(f) Average New Concepts



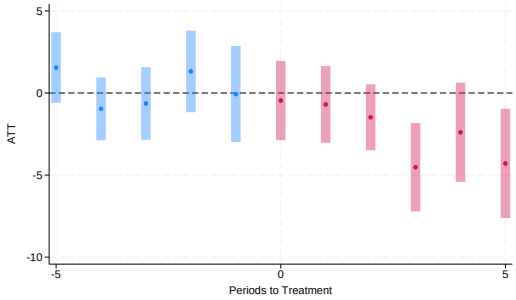
(g) Old Concepts



(h) Average Old Concepts



(i) Diversity of References



(j) Diversity of Citations

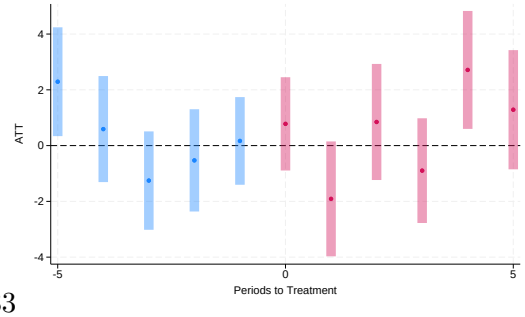
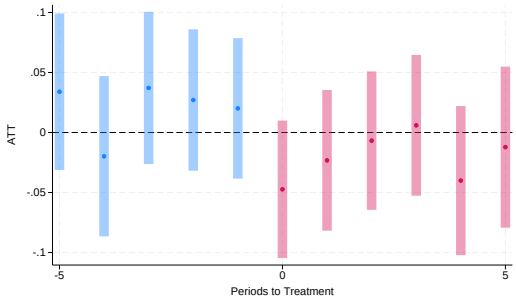
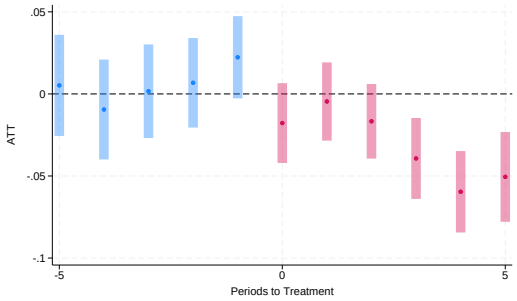


Figure 5: Panel Event Study (Students' Variables, cont.)

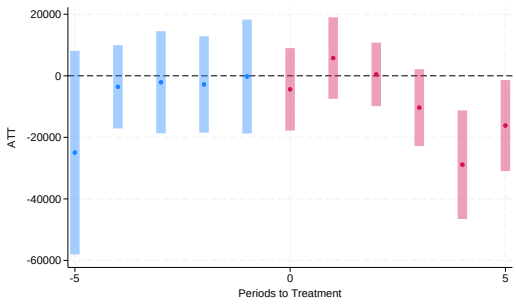
(k) Amount of Publishing Students



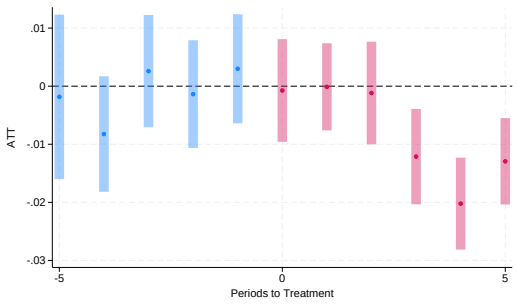
(l) Fraction of Publishing Students



(m) Amount of Funding from Students



(n) Fraction of Students Receiving Funding



Appendix 1: Dataset Construction

Throughout our data, the main individual identification is called "auid". The authorlink dataset is a convenient disambiguated dataset that combines project number (for grants), other project information, and auid. We first only keep R01-equivalent grants (R01, DP1, DP2, DP5, R37, R56, RF1, RL1, U01, R35), and then merge the file with an NIH ExPORTER file containing data about NIH grants and their PIs from 1985 to 2008, creating a link between auid and grant received.

Simultaneously, we use a disambiguated file from ProQuest which links auid with thesis information (advisor name, advisee name, degree, school, year of graduation). We drop master's degrees, and individuals with too many graduations (caused by disambiguation errors) and obtain a graduation date for individuals we have information on their doctorate thesis. For those we do not, we estimate a graduation date 16 years prior to their first advisee. The graduation year for individuals with more than one thesis is their second graduation, as long as this one is no more than 4 years after the first.

Additionally, we use an Author-ity file that contains data on all authorship and citations for papers published in the biomedical fields. We construct an indicator that only look at forward citations five years after publication to prevent recent papers from having less citations. We create a lab size variable that prevent the same author to be counted more than once as a collaborator for each author and each year. We also construct a mean size variable by averaging the total authors (here duplicates are counted) by the number of papers, for each year. We combine these data with the metrics obtained by Yu et al. (2023), namely their variables about the use of new concepts in a paper, old concept, and diversity of both references and citations. Yu's data does not contain information on all the authors that our dataset possesses and therefore lack of information will be treated as missing values as opposed to bull values.

Student's variables are obtained by first creating a linkage between advisor's auid and advisee's auid. Combining it with the three previous dataset we can obtain information on students such as grants received, paper published, and forward citations. For students with an ambiguous graduation date (more than one graduation), we take the average. This way we lose precision but keep observations.

We then combine all 4 datasets together, linking them through auid and years. Finally, we expand the dataset to obtain "0s" – for missing years therefore production is equal to 0 – discard observations due to inactivity both at the beginning and at the end of researchers' career. Then the dataset is merged with an indicator of shock, depending on the definition being used.

For the top 20 universities, we follow the same steps but only keeping observations from ProQuest where the university is in our top 20.