# Two Centuries of U.S. Innovation: Firms' Internal Networks and Resilience to Disasters*

Mathias S. Kruttli, Noah Stoffman, Sumudu W. Watugala†

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**Abstract**

We construct a comprehensive database of the universe of approximately 12 million U.S. patents from 1836 to 2023 and analyze the resilience of innovation to disaster shocks using hurricane landfalls. Major hurricanes reduce local innovation for up to a decade and lead to permanent losses. We show that spatial spillovers along firms' internal innovation networks can be negative (via the capital channel) or positive (via the innovation productivity channel). Empirically, we find the latter dominates: multi-location firms reallocate resources and increase innovation in establishments outside the landfall region after a hurricane hit, boosting aggregate innovation in distant counties.

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†Kruttli: Kelley School of Business, Indiana University. Email: mkruttli@iu.edu. Stoffman: Kelley School of Business, Indiana University. Email: nstoffma@iu.edu. Watugala: Kelley School of Business, Indiana University. Email: sumudu@iu.edu.

# 1 Introduction

Innovation is a key driver of economic growth (Solow, 1957; Romer, 1990, 1994). The geography of innovation—its spatial distribution and clustering—plays a crucial role in how knowledge diffuses and evolves.[1] However, there is little understanding on the impact of destructive, local exogenous shocks like extreme weather events on aggregate innovation and its spatial distribution. The local vulnerability to such shocks and potential propagation across regions are important to assess the resilience of innovation and the economic costs of extreme weather events.

In this paper, we analyze the local impact of extreme weather shocks and these shocks propagate through multi-location firms. We present a conceptual framework highlighting two channels of shock propagation within firms' internal networks: (i) the capital channel and (ii) the innovation productivity channel. The first, in the style of Giroud and Mueller (2019), posits that an extreme weather shock negatively impacts the local (and hence aggregate) capital of a firm and predicts that innovation will decrease in establishments both inside and outside the disaster zone, generating negative spillovers.[2] The second channel predicts that firms will reallocate innovation activity to establishments outside the disaster zone because of higher relative productivity, generating positive spillovers. Which effect dominates is ultimately an empirical question.

To analyze these mechanisms, we first overcome two key empirical challenges. First, by harnessing recent advances in optical character recognition (OCR) and large language model (LLM) technologies, we obtain a reliable database of patent activity that captures, for the first time, a comprehensive view into inventor and assignee names, locations, and types for the universe of around 12 million U.S. patents from 1836. We harness these data to match establishments of the same firm to identify firms' internal "innovation" networks over this

[1] The externalities to the local distribution of people and firms have been long debated (Marshall, 1890; Jacobs, 1969) with implications for both regional and national place-based policy (Acemoglu, Alp, Bloom, and Kerr, 2018; Fajgelbaum and Gaubert, 2020).

[2] These spillovers can occur as corporate headquarters optimally allocate resources across different establishments (Stein, 1997; Gertner, Scharfstein, and Stein, 1994).

period. Second, we obtain a set of exogenous local shocks that impact different locations at different points in time by exploiting the quasi-exogenous spatial and temporal variation of natural disasters over a period spanning two centuries. Adapting the research design of Kruttli, Roth Tran, and Watugala (2025), we use hurricane landfalls as exogenous shocks to local innovation. Using a long time series gives us a sample of 64 major hurricanes, which helps reduce selection bias and allows for subsample analysis.

We first analyze the direct effects of hurricanes on county-level innovation. We assign patents generated each year to counties by inventor location to obtain a county-year time series and estimate a panel version of the local projections method developed by Jordà (2005). The coefficient estimates imply large and long-lasting negative effects on patent output in counties that are within the landfall region of a hurricane. For hit counties, the annual patent output relative to the pre-period is up to ten percentage points lower than for control counties (i.e., counties outside of the landfall region) for close to a decade after landfall. The reduction in the growth rate of patents is most pronounced for counties located closest to the eye of the hurricane at landfall, which likely experience the most intense weather. Importantly, this reduction in the growth rate of patents for hit counties is a shortfall that does not appear to be recouped by a higher patent growth rate once the hit county recovers.

Next, we analyze how these local hurricane shocks to innovation can spill over to other regions in the U.S. through firms' internal firm networks, testing the predictions of our conceptual framework. Multi-location firms can have plants and establishments that span different regions in the U.S. with corporate headquarters allocating capital across establishments and adjusting to shocks. We find strong evidence of positive spillovers. Firms have a decrease in patent output in their establishments within a landfall region, but increase patent output in establishments elsewhere. The estimated effect is not only statistically but also economically significant. If half of a firm's establishments that produce patents are hit by a hurricane, the firm increases patent output in the other half of the establishments by up to 8.5%. These results are robust to including county interacted with year fixed effects, which

control for all time-varying and time-invariant county-level characteristics—like population and GDP—and allows us to compare establishments from different firms in the same county.

Interestingly, the magnitude of the effects are larger when we analyze longer horizons around the hurricane landfall. This finding is in line with agglomeration benefits and efficiency gains stemming from the within-firm capital reallocation. When a firm is forced to reallocate resources and capital across regions, they will choose the region where they expect the largest agglomeration benefits and such benefits could be more likely to materialize at longer horizons.

These spillover effects hold for both the period from 1851 to 1975—the period for which we used OCR and LLM technologies to produce the novel patents data—and for the period from 1976 to 2023. Further, the effects are stronger when a hurricane makes landfall during a year in which the U.S. is in a recession, which suggests that reallocating resources out of the hurricane landfall region is even more important during recessionary periods. This result is consistent with theories on the high value of innovation during recessions due to creative destruction (Schumpeter, 1942).

These spillovers via firms' internal networks are not only important at the firm-level but also at the county-level. We find that counties outside the hurricane landfall region but with links to the hurricane landfall region through local firms' internal networks experience increases in patent growth post-landfall. These results hold when including state interacted with year fixed effects, which allows for the comparison of two counties within the same state and controls for all time-varying and time-invariant state-level characteristics.

There are two major aspects to our contributions to the literature in this paper. First, by analyzing hurricane landfalls from 1851 to 2023, we show that extreme weather shocks have a large negative effect on innovation output lasting up to a decade for counties hit by the event. Further, firms' internal networks are important for the propagation of such exogenous local shocks, impacting innovation across regions. Our conceptual framework formalizes two potential channels of within-firm spillovers, the capital and innovation pro-

ductivity channels, which we test empirically. Our findings highlight the importance of the innovation productivity channel, which shows that firms can compensate for some of the lost innovation by reallocating resources out of the landfall region and increasing innovation in other counties. Second, by developing a new methodology and constructing a new, highly accurate, comprehensive database of patent information (including inventors and their locations, technological characteristics, and assignee information), we enable future research that answers important questions regarding two centuries of U.S. innovation.

Our findings contribute to several strands of economic research. Our analysis adds to previous studies on how local shocks can propagate across a firm's establishments. Giroud and Mueller (2019) show that for firms in non-tradable industries (e.g., retail and gastronomy), local shocks originating from collapse in house prices during the great recession lead to lower employment at a firm's establishments close to the local shock but also in other regions in which the firm operates. In contrast, we exploit multiple exogenous shocks yielding temporal and spatial variation that allow us to analyze the impact of local shocks on regional and firm-level innovation. We formalize two potential channels of generating negative and positive spillovers within firms and find strong support for positive spillovers. Our setting is unique in the literature that examines agglomeration and local bias in knowledge spillovers and the constraints to reallocating the physical and human capital necessary for innovation across regions. The existing literature has shown that innovation and attendant knowledge spillovers can be highly localized and sticky to a place (see, for example, Jaffe, Trajtenberg, and Henderson, 1993; Audretsch and Feldman, 1996; Moretti, 2021; Atkin, Chen, and Popov, 2022). However, only a few recent papers investigate spatial distribution of innovation within a firm.[3] Giroud, Liu, and Mueller (2025) find that large tech clusters have high productivity both in local inventors and in distant inventors in plants linked via parent firms. Chikis, Kleinman, and Prato (2025) study the social optimum of spatial distribution

[3]Other work on internal firm networks is not focused on innovation and include, for example, Cravino and Levchenko (2017); Bena, Dinc, and Erel (2022); Biermann and Huber (2024) on international shock propagation via multinational firms.

of firms' innovation. In contrast, our study focuses on the resilience of innovation in firms' internal networks.

We add to the literature on the economic risks associated with climatic shocks. The risks are broadly categorized into two types: physical and transition risks (Carney, 2015; Giglio, Kelly, and Stroebel, 2021). Physical risks encompass the destruction from more intense and frequent extreme weather events like hurricanes and heat waves (e.g., Hong, Li, and Xu, 2019; Kruttli, Roth Tran, and Watugala, 2025) or sea-level rise. Transition risks refer to the risks associated with the transition to a low-carbon economy due to policies imposing a cost on firms and households (e.g., a cap-and-trade program). There exists an inherent trade-off between the two risks: imposing a price or a tax on greenhouse gas emissions increases transition risks but reduces greenhouse gas emissions and consequently future physical risks (e.g., Ivanov, Kruttli, and Watugala, 2024). Therefore, a precise estimate of the cost of physical risks is key to determining how high the price or tax on greenhouse gas emissions should be set.

Beginning with the seminal work of Nordhaus (1977), there is an extensive literature on climatic risks and resulting economic costs in the form of reduced economic growth. Studies on physical climate risks and economic growth have focused on the impact of temperature shocks (e.g., Dell, Jones, and Olken, 2012 and Burke, Hsieng, and Miguel, 2015) and other natural disasters (e.g., Deryugina, 2017; Boustan, Kahn, Rhode, and Yanguas, 2020; Roth Tran and Wilson, 2024) on countries and regions. Other work has shown that extreme weather events can propagate through (across-firm) supply-chain networks (e.g., Barrot and Sauvagnat, 2016; Pankratz and Schiller, 2024). Importantly, a key gap in this literature is an in-depth empirical assessment of the impact of physical risks on innovation. Examining this innovation channel is particularly important given the integral role for innovation and technological change in economic growth.

Separately, there is a literature on climate adaptation, green technology, and innovation Nanda, Younge, and Fleming, 2015; Cohen, Gurun, and Nguyen, 2020; Bena, Bian, and

Tang, 2023; Howell, 2024; Atta-Darkua, Glossner, Krueger, and Matos, 2025. Noy and Strobl (2023) analyze the impact of hurricanes on subsequent disaster-mitigating innovation using the patent data constructed by Petralia, Balland, and Rigby (2016). They find small, temporary increases in patents with the terms "hurricane" or "storm" in the text. In contrast, we focus on transmission via firms' internal networks and the substitution of innovation across regions in the aftermath of a hurricane.[4]

By analyzing an extended history of the universe of U.S. patents, we complement other papers that study innovation for years not covered by the NBER patent database (Hall, Jaffe, and Trajtenberg, 2001), which starts in 1976. For example, Kogan, Papanikolaou, Seru, and Stoffman (2017) process patent information for the subset of patents issued since 1926 for which the assignee is a publicly traded firm. Kelly, Papanikolaou, Seru, and Taddy (2021) measure the novelty of historical patents using textual analysis methods but do not extract information on the inventor and assignee names and location, and rely on the text provided by Google Patents. Petralia, Balland, and Rigby (2016) also provide a database of historical patents with inventor and assignee names and location. By exploiting recent improvements in OCR and LLM technology to extract information, we are able to construct a database that contains fewer errors and missing data than earlier databases.

The paper is organized as follows. We present the details of the conceptual framework in Appendix A and the empirical design in section 2. Section 3 describes the data, including the methodology for constructing the novel patent database covering two centuries. Section 4 presents the results. Section 5 concludes.

# 2 Research design

Our empirical strategy takes hurricanes as exogenous shocks to regions and is adapted from Kruttli, Roth Tran, and Watugala (2025). We construct the sample of hurricanes that made

[4]Our analysis of exogenous natural disaster shocks over two centuries also differs from papers like Nanda and Nicholas (2014); Babina, Bernstein, and Mezzanotti (2023); Mao and Wang (2023) that analyze the impact of historical bank access or financial crises on innovation.

landfall on the Atlantic and Gulf Coasts of the U.S. by processing data on hurricane paths that are available from 1851 onwards from NOAA. Hurricanes have made landfall over major population and economic centers in various states in this region. The landfall region of a hurricane typically spans several counties in one or more states. Figure 5 plots the landfall regions of four hurricanes in our sample.

Figure 6A presents a stylized example illustrating inventor locations within counties that are exposed and unexposed to a particular hurricane landfall. We consider counties located within a hurricane landfall region as treated and counties outside of it as controls. This spatial variation in the empirical design gives us cross-sectional variation. Because counties are hit by hurricanes infrequently, the time series variation allows us to analyze within county effects.

## 2.1 County analysis

Using the hurricane landfall regions and the location of the inventors for a given patent, we first estimate how a hurricane affects innovation activity in a county. We adapt the local projection estimator of Jordà (2005) for our panel of county-year observations.[5]

The local projection regression specification is:

$$\log\left(\frac{NPatents_{c,t+h}}{NPatents_{c,t-1}}\right) = \beta_1 Hit_{c,t} + \sum_{r=-5,r\neq 0}^{h} \rho_r Hit_{c,t+r} + \sum_{r=1}^{5} \kappa_r \log NPatents_{c,t-r} + \mu_c + \theta_t + \epsilon_{c,t}. \quad (1)$$

The dependent variable is the change in the log of the number of patents, $NPatents$, generated in a county $c$ from the year before the hurricane hit, $t-1$, to $h$ years after the hurricane hit, $t+h$. A patent is assigned to county $c$ if all the inventors of the patent are located in that county. A patent is assigned to the year in which the patent is issued.[6] The exceptions

[5] A similar regression specification is used by Roth Tran and Wilson (2024).

[6] The results are qualitatively similar when using the application instead of the issue date. However, for patents before 1900, often no application date is recorded.

are patents issued between January and May. These patents count for the previous year to align patents to the correct event because Atlantic Coast hurricanes only make landfall during the hurricane season from June to November.[7] The variable $Hit_{c,t}$ is an indicator variable that takes the value one when a county $c$ is in the landfall region of a hurricane in year $t$ and zero otherwise. In addition to county and year fixed effects, we include several control variables.[8] First, we control for a county being hit by hurricanes that make landfall during the preceding 5-year period from $t-5$ to $t-1$. We also control for a county being hit by hurricanes that make landfall during the years from $t+1$ to $t+h$. These controls account for the possibility of staggered and multiple treatment of counties (Athey and Imbens, 2022; Baker, Larcker, and Wang, 2022; Dube, Girardi, Jorda, and Taylor, 2023). Second, we capture potential pre-trends in a county's innovation activity by including the log of $NPatents$ lagged up to five years.[9] The standard errors are double clustered at the county and year levels. Of all the states in the US, 32 have at one point been hit by a hurricane from 1851 to 2023 based on our landfall region estimations. We include counties from only these 32 states in our sample when estimating county-level impacts.

## 2.2 Firm's internal network analysis

The simple model that is presented in detail in Appendix A shows a firm with central headquarters that allocate resources across different establishments (e.g., Stein (1997); Giroud and Mueller (2019)) to maximize the firm's overall innovation. When a hurricane hits a firm's establishment, there are two channels for the propagation of this shock to other non-hit establishments of the same firm. First, the hurricane reduces the total capital of the firm and makes it poorer, because the establishment in the landfall region can experience damage to the facility, inventory, and reduce cash flow from the local operations. This capi-

[7] Omitting this adjustment leads to qualitatively similar results.

[8] The dependent variable is a difference between the pre- and post-period and omitting the county fixed effects leads to qualitatively similar estimates.

[9] The estimates are qualitatively similar when including additional lags. Including additional lags does not materially improve the fit of the regression model.

tal channel leads to a reduction in R&D at non-hit establishments as the firm equalizes the marginal innovation product across establishments. Second, the innovation productivity at the establishment in the landfall region decreases. The decrease in innovation productivity is motivated by, for example, damaged county infrastructure and a labor force that is preoccupied with the aftermath of the hurricane. This innovation productivity channel leads to an increase in R&D at non-hit establishments, as headquarters will allocate more resources to the relatively more productive establishments.

To estimate which of these two channels dominate, we analyze changes in innovation at the firm-county-year level. We employ a difference-in-differences regression framework. We collapse the pre- and the post-period around each hurricane (Bertrand, Duflo, and Mullainathan, 2004) and jointly estimate the difference-in-differences across all hurricanes:

$$\log\left(\frac{\overline{\text{NPatents}}_{i,c,t:t+h-1}}{\overline{\text{NPatents}}_{i,c,t-h:t-1}}\right) = \beta_1 \text{Hit}_{c,t} + \beta_2 \sum_{k \neq c} w_{i,k,t-h:t-1} \text{Hit}_{k,t} + \mu_c + \theta_t + \varepsilon_{i,c,t+h} \qquad (2)$$

The dependent variable is the change in the annual average number of patents of firm $i$ in county $c$ over $h$ years after a hurricane landfall relative to the annual average number of patents over $h$ years before landfall. The variable $h$ is set to one, three, five, and ten years, respectively. Using windows of multiple years around the landfall accounts for the fact that firms do not generate a patent in a location every year. The variable $w_{i,k,t-h:t-1}$ is the share of firm innovation in county $c$ over $h$ years before the hurricane hit. The variable $\sum_{k \neq c} w_{i,k,t-h:t-1} \text{Hit}_{k,t}$ measures the share of a firm's innovation locations other than county $c$ that are in the hurricane landfall region. For simplicity, we will denote this variable as $HitOther_{i,t}$. The weight is given by

$$w_{i,k,t-h:t-1} = \frac{\text{NPatents}_{i,k,t-h:t-1}}{\text{NPatents}_{i,t-h:t-1}}. \qquad (3)$$

Therefore, the variable $HitOther_{i,t}$ ranges from zero to up to but not including one.

We include county and year fixed effects.[10] Further, we estimate specification that include interacted county and year ($c \times t$) fixed effects, which allows us to compare the changes in innovation of one firm's establishment to the changes in innovation of another firm's establishment within the same county and time. These fixed effects control, for example, for differences in economic growth across regions. The standard errors are double clustered by firm and year. To ensure that a firm hit by one hurricane does not enter as a control for a different hurricane, we exclude firms that have more then 50% of their innovation establishments hit from the controls for ten years before and after the hurricane.

# 3 Data

## 3.1 Novel patents database

> *[The Congress shall have Power . . . ] To promote the Progress of Science and useful Arts, by securing for limited Times to Authors and Inventors the exclusive Right to their respective Writings and Discoveries.*
>
> —U.S. Constitution, art. I, §8, cl. 8.

Innovation has been constitutionally protected in the United States since 1789 with a recognition that innovation is fundamental to the advancement of science, health, prosperity, public welfare, and national defense. Over 12 million patents have since been granted, with the U.S. Patent and Trademark Office (USPTO).

To facilitate studying the longest possible time series of hurricanes, we construct a new patent database. The long time series is necessary to capture the long-lasting effects of hurricanes on innovation and helps minimize the potential bias from sampling just one or a set of hurricanes. The data include information on grant date and inventor location (city and state) spanning 1836 through 2023.

Modern patent documents include the title and abstract of the invention; application and grant dates; the name(s) and locations of residence for all inventors and any assignees; a

[10] The dependent variable is a difference between the pre- and post-period and omitting the county fixed effects leads to qualitatively similar estimates.

detailed description of the invention; drawings of the invention; and the claims, which define the scope of legal protection provided by the patent. An inventor must be a person, but patents can be assigned to a corporation or another person. Generally, when an invention is developed in a company, the inventor will assign the patent to their employer. Beginning in February 1947, patent documents began consistently including citations to existing patents as references to "prior art" (Nicholas, 2010). Older patents typically do not contain all of this information, although the grant date and inventor location have always been included.

While the layout of patent documents has changed over time, the required information has mostly remained constant. Since 1976, the USPTO has digitally recorded newly-granted patents and made machine-readable files publicly available. For the period prior to that, information must be extracted from scanned documents, an example of which is shown in Figure 1. We extract data directly from the PDFs provided by the USPTO by taking advantage of modern OCR algorithms and ChatGPT, as we explain below. Additional details about our data construction procedures are provided in Appendix B.

The launch of Google Patents in 2006 made it easier to search the text of the pre-1976 historical patent documents. These data serve as the source for several influential papers, including Kogan, Papanikolaou, Seru, and Stoffman (2017), who created a widely used database, and Kelly, Papanikolaou, Seru, and Taddy (2021). However, the quality of the text extracted from scanned documents using Optical Character Recognition (OCR) available in Google Patents is quite variable, especially for patents granted before 1950. The importance of the errors in these data depends greatly on how the text is being used. For our analysis, correctly identifying the inventor's location is of paramount importance.

As an example of the OCR errors that can be found in the Google data, Figure 2 shows how the text for the patent in Figure 1 is rendered by Google (left column). Note that this text does not include anything that resembles the name of the inventor, his city of residence, or the title of the invention. The entire first paragraph of the text is missing, as is the beginning of the next paragraph. Although not all of the OCR in Google Patents is of this

poor quality, many hundreds of thousands of documents have significant degradation in the extracted text. Even the use of exceptionally flexible textual analysis techniques cannot overcome the "garbage-in-garbage-out" problems associated with such poor OCR quality.

OCR algorithms have vastly improved in recent years (Correia and Luck, 2023) but the text available from Google Patents has not been updated. Advances in computer vision have given OCR algorithms the ability to identify the parts of documents, patterns in text layout, and words in ways that far surpass what is available in Google Patents. Moreover, LLMs can be a game changer when it comes to extracting information from text. Even in the presence of OCR errors, spelling mistakes, typos, or missing information, it can draw on its vast training data to predict answers to carefully-crafted questions about the (intended) meaning of text.

As a point of comparison with our novel data, we compared the locations of inventors in our data with those in the "HistPat" data created by Petralia, Balland, and Rigby (2016). Of the 3.9 million unique patents between 1836 and 1975, HistPat is missing over 600,000 patents, which appears to be mainly due to the exclusion of foreign inventors. Of the remaining 3.3 million patents, just over 400,000 have a discrepancy between inventor locations; this appears to be because the assignee's location is often misreported as the inventor's location in the HistPat data. Despite the requirement that all patents have an inventor, another 115,000 patents in HistPat have an assignee but no inventor. We also observe significant time series variation in differences between the two datasets. For example, the rate of missing patents in HistPat relative to our dataset is below 10% through 1880 before jumping to the teens for most of the period between 1880 and 1965. This missing rate increases to 19%, 24%, and 50% in 1965, 1970, and 1975, respectively.

Using our new inventor location data, aggregated to the county level, Figure 3 illustrates the expected westward shift in innovation over the last two hundred years.

## 3.2 Firms' internal networks

The raw patent data provide no identifier that links assignees across patents. An important part of our analysis relies on our ability to identify which patents are assigned to which firms over time—and to know that the *same* firm has been assigned different patents granted in different years to inventors in different locations. We accomplish this disambiguation by applying several machine learning techniques to the assignee names.

We begin by standardizing company names and extracting unique company names along with corresponding patent grant years. Our basic approach is to apply clustering techniques to these names, but we first preprocess the text, by standardizing special characters, merging single-letter sequences, and normalizing legal entity abbreviations such as Inc. or Corp. We then identify the 1,000 most common words across names to create a list of low-information tokens. We apply fuzzy matching techniques to correct misspelled legal entities, and correct inconsistent spacing, which can originate either in the original PDF document or from OCR transcription errors.

After producing this relatively clean list of unique company names, we divide the data into overlapping 40-year time windows for pairwise similarity scoring. We again apply a fuzzy matching algorithm to identify name similarities above a threshold that we determine from the data. (The pairwise nature of this procedure makes it extremely computationally intensive, but we are able to exploit parallel processing to complete the work in a reasonable time.) We use the *UnionFind* algorithm to merge related company names into clusters and ensure continuity across overlapping windows, so firms like General Electric that are continually granted patents over long spans of time are identified as one firm. To refine these clusters and address false positives, we use hybrid TF–IDF cosine similarity metrics to evaluate coherence, with low-quality clusters re-split using an agglomerative clustering algorithm. Finally, we used ChatGPT to review any large clusters (groups where at least 15 variations of a name are identified as belonging to one firm). This allows us to ensure

semantic consistency and correct any remaining misgroupings, resulting in a cleaned set of company name clusters.

Figure 4 illustrates the geographic distribution of patent generation for four of the the most innovative firms as of 1900.

### 3.3 Two centuries of hurricane landfalls

We obtain the hurricane path data for all Atlantic and Gulf Coast hurricanes making landfall from 1851 to 2023 from NOAA's Atlantic HURDAT2 database, which covers all known tropical cyclones and subtropical cyclones and is a part of the Re-analysis Project (Landsea and Franklin, 2013). These data record the latitude and longitude of the eye of a tropical cyclone at least every 6 hours before it dissipates. This allows us to calculate hurricane paths and landfall dates and times. We use these data to calculate the list of counties that are within 50, 100, and 200 miles of the landfall location of the eye of a hurricane. These radii line up with NOAA's measurements on the average distance from the hurricane eye within which wind speeds cause damage to infrastructure. Figure 5 shows the landfall regions for four hurricanes in our sample.

For our main analysis, we focus on the set of deadliest tropical storms with more than 25 fatalities during our sample period. We construct this set from NOAA's report on "The deadliest, costliest, and most intense United States tropical cyclones from 1851 to 2010 (and other frequently requested hurricane facts)" that covers the 1851-2010 period (Blake, Landsea, and Gibney, 2011). We manually augment these data for the more recent years to cover the full sample period from 1851 to 2023. The list of 64 deadliest hurricanes in our sample is shown in Table C.3. We focus on this set of storms for two key reasons. First, these are likely the hurricanes with the most reliable data in the early period of our sample. Second, these are likely the hurricanes that significantly impacted human and physical capital in the landfall regions. The number of fatalities is one reliable measure of the destructive impact of a hurricane, which does not rely on as many assumptions as, potentially, the monetary

value of damages. However, this set overlaps substantially with the set of hurricanes with the highest dollar damages.

Tables 1 and 2 present the summary statistics for the main variables used in our empirical analysis.

# 4 Results

This section presents the results of our empirical analysis. First, we discuss the baseline effects of hurricanes on local innovation and spillovers of these local innovation shocks through firms' internal innovation networks. Then we report robustness tests, time-series variation of the effects, and county-level spillovers.

## 4.1 Baseline effects

### 4.1.1 Local effects

We first estimate the regression in equation (1) to test if a hurricane hit adversely affects innovation in a county, how large the effects are, and whether the effects are transitory or permanent. When identifying hit counties, that is, the counties for which the variable $Hit_{c,t}$ is one, we measure landfall regions based on three radii (50, 100, and 200 miles) around the eye of the hurricane. We estimate the regression separately for each radius and horizon $h$. For all estimations, we exclude from the control group any counties located within the 200-mile landfall region.

Figure 7 plots the coefficient estimate of the variable $Hit_{c,t}$ for regressions separately run for different radii and horizons $h$. The negative effects on patents do not appear immediately after landfall. The drop in the number of patents generated in a hit county, relative to the year before landfall, begins three years after landfall. However, the effect intensifies as time passes and the decrease in innovation is most pronounced between five and eight years after

the hurricane landfall. Ten years after landfall, the coefficient estimates become statistically insignificant.

The effects are larger for counties closer to the eye of the hurricane at landfall. The magnitude of the coefficients for the 50-mile radius reaches as low as -0.094 in five years after landfall. This magnitude implies that the number of patents in a hit county is 9.4% lower in the fifth year after the hurricane hit than in the year before the hurricane hit. For the 200-mile radius, the coefficient magnitude decreases but still reaches -0.044. The lower coefficient magnitude is consistent with the idea that hurricanes are more destructive closer to the eye of the storm. The negative and significant coefficient estimates for several years after landfall show that a county in a hurricane landfall region experiences a substantial shortfall in patents.

Table C.1 in the appendix reports the estimates of the coefficients in equation (1). The first row of the table shows the coefficient of the independent variable of interest, $Hit_{c,t}$, which is plotted in Figure 7. The coefficient estimates on the lagged patent variables are highly significant. The coefficient on the first lag is negative with a magnitude between -0.68 and -0.77. This estimate implies that the time series of county-level patents is no longer explosive after differencing the dependent variable and is mean-reverting. We confirm the unit root in the (undifferenced) county-level patent output series with an augmented Dickey-Fuller test (unreported) for all U.S. patents from 1936 to 2023. Therefore, differencing the dependent variable is necessary when estimating regression models with county-level time series of patents. We also estimate a difference-in-differences "event study" regression as an alternative to the local projection model. The results are qualitatively similar and presented in Table C.2.

#### 4.1.2 Spillovers via firms' internal innovation networks

Given the results that hurricanes constitute large negative effects on local innovation, we investigate how such shocks propagate in firms' internal networks and test the two channels

described by our conceptual framework. On the one hand, negative shocks to one location might lead to firms reducing R&D investment in other regions to better absorb the financial shock—that is, the capital channel. On the other hand, firms might be able to move some of their innovation capacity to counties outside the hurricane landfall region and prevent the loss of innovation at least to some extent—that is, the innovation productivity channel.

The estimates of the regression model in equation (2) are presented in Table 3. We estimate the regression model separately for landfall regions based on the 100-mile and 200-mile radii, respectively.[11] For estimations that rely on the 100-mile radius landfall region, we exclude from the control group any observations in counties located within the 200-mile but not the 100-mile landfall region, respectively.

The coefficient estimates of the variable $Hit_{c,t}$ are negative and significant. These estimates imply that when a county is in a hurricane landfall region, a firms' patent output is up to 4% lower than for counties not in the landfall region. These estimates are consistent with the county-level local innovation impact discussed in Section 4.1.1.

Interestingly, the coefficient estimate of the variable $HitOther_{i,c,t}$ is positive and strongly significant across all specifications. These results show that firms increase innovation in establishments located in other counties when part of their operations are hit by a hurricane. Across the different specifications, the coefficient estimates range from 0.05 to 0.17. These magnitudes imply that a firm with nearly all of its establishments in a landfall region will increase innovation in the remaining establishments located elsewhere by 5% to 17%.

These estimates are consistent with the innovation productivity channel dominating the capital channel. A firm increases R&D output at other establishments when some of its establishments are hit by a hurricane. This result indicates that firms' internal innovation networks are surprisingly resilient.

The magnitude of the coefficient estimates is larger for longer horizons around the hurricane landfall. For both landfall region radii, the specification with the largest estimate is

[11] We focus on the larger radii in this analysis to ensure a sufficient number of hit establishments for more precise estimates.

the one where we compare ten years after to ten years before landfall. The larger increase at longer horizons could be explained by agglomeration benefits and efficiency gains that come with the relocation. When a firm is forced to reallocate resources and capital to other regions, management will choose the region where they expect the agglomeration benefits to be largest and such benefits are likely to materialize more at longer horizons.

Panel B shows the results when including county times year fixed effects. Including the interacted fixed effects strengthens the identification. We can now compare the change in innovation of two establishments of two different firms within the same county and time. The fact that the coefficient estimates remain strongly significant and largely unchanged in magnitude confirms that our results hold after controlling for both time-invariant and time-varying unobservable county characteristics such as economic growth.[12]

Based solely on these regressions, it is unclear if firms simply reallocate resources within counties in the hurricane landfall region or to counties elsewhere. Reallocation within the landfall region is possible because the damage to counties within the landfall region can vary. To address this point, Table 4 presents the results for the specification with the county times year fixed effects, but where we exclude counties in the hurricane landfall region from the sample. The magnitude and significance of the coefficient estimates remains qualitatively the same. These results confirm that hurricane shocks create spillovers via firms' internal networks, and this leads to innovation growth in distant firm establishments outside the landfall region.

## 4.2 Effects across time

The results presented in the previous sections are for our total time series of 173 years. A natural question is whether these baseline estimates on the propagation of hurricane shocks via firms' internal innovation networks change through time. Particularly, are the results different when using the patents data up to 1975, which we generated using our novel

[12]This specification is particularly useful because county-level economic data and other statistics are not available for the full sample period. For example, county-level GDP growth is only available from the 1950s.

methodology, compared to when using the digitized patents data available directly from USPTO starting in 1976?

To address this question, we adapt the regression specification in equation (2) and include an indicator term that takes the value one for hurricanes that make landfall between 1851 and 1975 and zero otherwise. We interact this indicator variable with our variable of interest $HitOther_{i,c,t}$. Table 5 reports the results. The magnitude of the coefficient estimate on the unconditional variable $HitOther_{i,c,t}$ remains strongly significant and positive for all specifications. The magnitude decreases slightly compared to the estimates in Table 3. However, the coefficient estimates for $HitOther_{i,c,t} \times I_{t \in 1851-1975}$ are also positive, strongly significant, and larger in magnitude.

These estimates show that the mechanism we document is present across both patent datasets. However, the magnitude of the effects are larger in the sample up to 1975. A potential reason could be that infrastructure became more resilient and rebuilding efforts quicker, and therefore, the need to move innovation activity into other regions became somewhat mitigated. Improved and faster rebuilding efforts could have been partly driven by the Federal government. The Federal Disaster Relief Act of 1950 started efforts for centrally organized and improved disaster relief.

Our second analysis on the time-series variation of the baseline effects focuses on U.S. recessions. On the one hand, recessions constrain firms' operations and can limit their innovation capacity (Hall, 2015). Such constraints could tighten even further when a firm is hit by a hurricane, which leads to a stronger capital channel. On the other hand, recessions could present opportunities for innovation and accelerate creative destruction (Schumpeter, 1942). These opportunities would incentivize firms to keep innovating and enhancing productivity, which could lead to a stronger innovation productivity channel with greater reallocation of R&D resources to more productive establishments.

To test these hypotheses, we estimate the same regression as presented in Table 5, but we replace the time indicator variable with an indicator variable that takes the value of

one if the year of the hurricane landfall is in a recession year. To identify recession years, we use the NBER recession data to identify years for which at least six months were in a recession. Table 6 presents the results. The unconditional regression estimate remains positive and strongly significant, but the spillovers within firms' internal networks are more pronounced when a hurricane makes landfall in a recession year. The coefficient estimates for $HitOther_{i,c,t} \times I_{t \in Recession}$ are always positive and mostly strongly significant. These estimates are consistent with the innovation productivity channel becoming stronger during recessions.

## 4.3 Extensive margin

In the regression in equation (2), the dependent variable is the log change in the average annual number of patents of a firm in a county from $h$ years pre- to $h$ years post-hurricane landfall. Using the log change as the dependent variable ensures that the variable is well behaved despite large differences in the number of patents across firms, and we control for unobservable firm-county level characteristics. However, the estimates do not capture if a firm completely stops innovation in a particular county in the post-hurricane landfall period or starts innovating in a county in which it had no innovation in the pre-hurricane landfall period. In this section, we analyze these potential extensive margin effects.

To address the first case, we estimate a logit regression to analyze the likelihood that a firm continues to innovate in a county after a hurricane hit. Here, the dependent variable takes the value of one if the firm innovates in the county in the post-hurricane landfall period and zero otherwise. The regression specification is given by

$$\begin{aligned} logit(P(\overline{\text{NPatents}}_{i,c,t:t+h-1} > 0 | \overline{\text{NPatents}}_{i,c,t-h:t-1} > 0)) = & \beta_1 HitOther_{i,c,t} \\ & + \psi_{c,t} + \varepsilon_{i,c,t+h}. \end{aligned} \tag{4}$$

Table 7 presents the estimates. A firm whose establishments elsewhere are in counties that are in a hurricane landfall region is much more likely to keep innovating in an existing establishment in a county. The coefficient estimate for $HitOther_{i,c,t}$ is strongly significant and positive for all specifications. The coefficient magnitude is between 0.4 and 0.5, which implies a marginal effect between 8% and 12%. This indicates that the probability of a firm retaining innovation activity in a county is 8% to 12% higher if the $HitOther_{i,c,t}$ variable increases by one unit. Compared to the baseline estimates in Table 3, the magnitude of the coefficient estimates do not increase for longer horizons. This finding is intuitive, as the dependent variable in the logit regression is capped at one and does not measure agglomeration benefits.

To address the second case—whether firms with exposure to a hurricane are more likely to start generating patents in counties in which they did not innovate before—we estimate the firm-year level regression given by

$$\frac{NrNewCounties_{i,t:t+h-1}}{NrTotalCounties_{i,t-h:t-1}} = \beta_1 SharePatentsHit_{i,t} + \theta_t + \varepsilon_{i,t+h}. \tag{5}$$

The dependent variable measures the number of counties in which the firm generated patents in the post-landfall period but not the pre-landfall period (i.e., innovation in counties "new" to the firm) relative to the total number of counties in which the firm generated patents in the pre-landfall period. We include year fixed effects and the standard errors are double clustered by year and firm. The variable $SharePatentsHit_{i,t}$ measures the share of a firms' innovation locations within the landfall region of the hurricane, which can range from zero to one:

$$\text{Hit}_{i,t} = \frac{\sum_c \left( \text{NPatents}_{i,c,t-h:t-1} \times \text{Hit}_{c,t} \right)}{\sum_c \text{NPatents}_{i,c,t-h:t-1}}. \tag{6}$$

Table 8 reports the results. The coefficient estimate of $SharePatentsHit_{i,t}$ is positive for all eight specifications and strongly significant for six of the eight specifications. The magnitude of the coefficient estimate goes up to 0.04, which suggests that a firm that had all

of its establishments in the landfall region adds innovation to 4% new counties. This number is relatively small. The results suggest that while firms do shift resources and innovation activity into new counties post-hurricane landfall, our baseline regression specification in equation (2) captures the major effects by analyzing the spillovers within a firm's existing internal innovation network.

## 4.4 Firm-level substitution

Our results raise the question of whether firms can perfectly substitute the impact from a hurricane on their innovation output by moving resources to other locations. To test whether there is perfect substitution, we estimate a firm-level regression that is similar to equation (5):

$$\log\left(\frac{\overline{\text{NPatents}}_{i,t:t+h-1}}{\overline{\text{NPatents}}_{i,t-h:t-1}}\right) = \beta_1 SharePatentsHit_{i,t} + \theta_t + \varepsilon_{i,t+h}. \tag{7}$$

The dependent variable is the change in the average annual number of patents pre- and post-hurricane. We again include a year fixed effects and double cluster the standard errors by year and firm.

The results in Table 9 show that firms can almost perfectly substitute the effects from a hurricane. The coefficient estimates are negative and significant for some specification but the magnitude is small. The magnitude ranges from 0 to -0.05, which implies that a firm that has all of its establishments hit by a hurricane experiences a decline in its total number of patents output at the firm-level of only up to 5%.

## 4.5 County-level spillovers

When firms reallocate resources in their internal networks in response to a hurricane shock, there may be observable changes in innovation output at the aggregate county level. A natural question that follows is if a county is located outside of the hurricane landfall region

but the firms with establishments in that county have other establishments in the hurricane landfall region, does the county see an increase in innovation? This question is important as it allows us to examine whether the shock propagation through firms' internal networks is large enough to affect aggregate local innovation output.

To test such county-level spillovers we estimate the regression model given by:

$$\log\left(\frac{\overline{\text{NPatents}}_{c,t:t+h-1}}{\text{NPatents}_{c,t-1}}\right) = \beta_1 CountyHitOther_{c,t} + \mu_c + \phi_{s,t} + \epsilon_{c,t+h}. \tag{8}$$

The dependent variable is the change in the average annual number of patents in county $c$ over $h$ years after the hurricane landfall relative to the number of patents in the year before landfall. We include county fixed effects and state interacted with year fixed effects. The standard errors are clustered by county and year. Counties in the hurricane landfall region are excluded from the regression. The independent variable $CountyHitOther_{c,t}$ measures the share of patents for firms with establishments in county $c$ that are hit by a hurricane in counties other than county $c$:

$$CountyHitOther_{c,t} = \frac{\sum_{i \in c} \sum_{k \neq c} \text{NPatentsFirm}_{i,k,t-5:t-1} \times \text{Hit}_{k,t}}{\sum_{i \in c} \sum_{k} \text{NPatentsFirm}_{i,k,t-5:t-1}} \tag{9}$$

Table 10 presents the results. Panel A shows the estimates with county and year fixed effects. The estimates are positive for all specifications. The magnitude of the estimates are increasing for longer horizons $h$ and are strongly significant for all specifications where $h$ is larger than one year. The magnitude increases from 0.05 at the one year horizon to 0.18 at the ten year horizon. This increase is in line with the increasing estimates at the firm-county level (e.g., Table 3) and suggest that agglomeration benefits also occur at the county-level.

Panel B includes state interacted with year fixed effects in the regression model. These fixed effects allow us to compare counties within the same state and year, but with different

exposure to the hurricane landfall region through the firms' internal networks. The results are largely unchanged in terms of magnitude and significance.

# 5 Conclusion

Innovation supports continued economic growth and is an essential driver of an economy maintaining a competitive edge. In this paper, we analyze how quasi-exogenous local extreme weather shocks impact the spatial distribution of innovation. To do so, we first construct a comprehensive, nearly-error-free database of all U.S. patents from 1836 to 2023 using advanced OCR and LLM technologies. We then analyze how hurricanes affect local innovation output and lead to spillovers across the U.S. through firms' internal innovation networks.

We find that major hurricanes destroy local innovative capacity for up to a decade following landfall and lead to permanent counterfactual losses for the landfall region. Firms that produce patents in the region before landfall, reallocate resources and increase their innovation output in other counties. These result are in line with a simple model where the innovation productivity channel dominates the capital channel. Our results are robust to different time-periods and become stronger during recessions. Further, the spillovers along the firms' internal networks lead to increases in aggregate innovation output for counties located outside of the hurricane landfall region. Our findings suggest potential efficient reallocation and agglomeration benefits following destructive natural disasters.

Harnessing almost two centuries of data, we reveal a mechanism that can generate divergence in regional economic prosperity. This mechanism is important for local and national policymakers to factor in when choosing policies to promote economic growth and prosperity, especially those focused on protecting and promoting innovative activity. Our results on how local natural disaster shocks have long-lasting effects on innovation are of particular impor-

tance to current debates regarding how to ensure the competitiveness of the U.S. economy while keeping it resilient to unexpected disasters and shocks.

Further, if extreme weather events like hurricanes become more damaging in the future, academics and policymakers alike will need to better understand the impact of extreme weather events on innovation to comprehensively assess the economic costs of these events. Our findings will be crucial for policymakers who seek to balance the costs and benefits of adaptation and resilience to extreme weather events.

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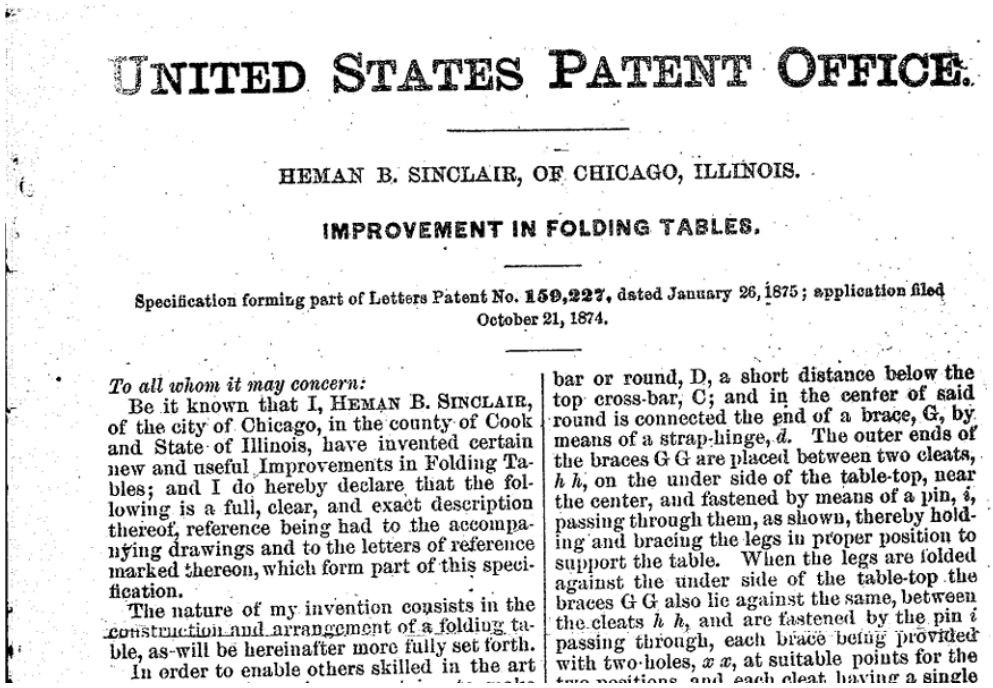
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# UNITED STATES PATENT OFFICE.

HEMAN B. SINCLAIR, OF CHICAGO, ILLINOIS.

## IMPROVEMENT IN FOLDING TABLES.

Specification forming part of Letters Patent No. **159,227**, dated January 26, 1875; application filed October 21, 1874.

*To all whom it may concern:*

Be it known that I, HEMAN B. SINCLAIR, of the city of Chicago, in the county of Cook and State of Illinois, have invented certain new and useful Improvements in Folding Tables; and I do hereby declare that the following is a full, clear, and exact description thereof, reference being had to the accompanying drawings and to the letters of reference marked thereon, which form part of this specification.

The nature of my invention consists in the construction and arrangement of a folding table, as will be hereinafter more fully set forth.

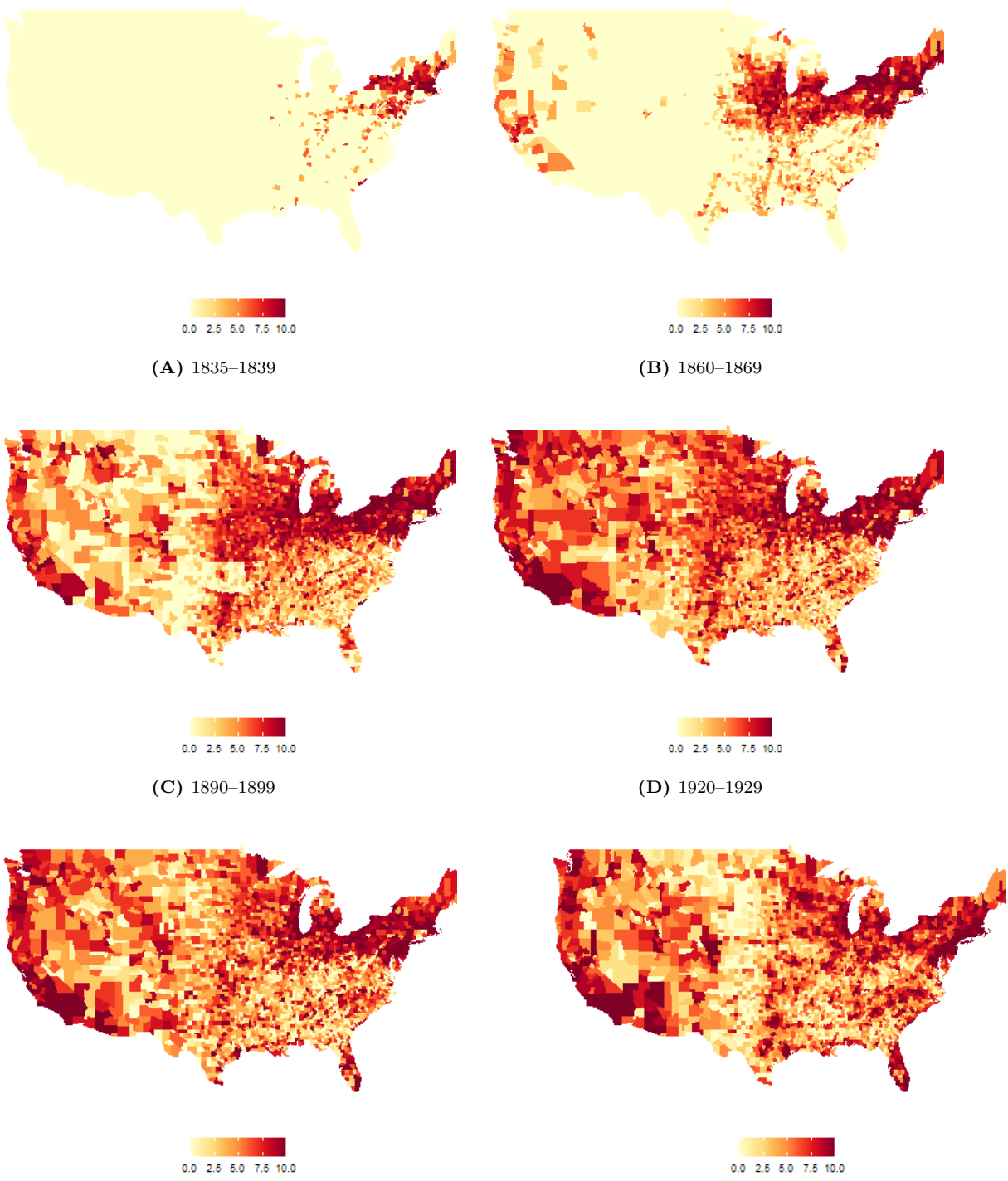
In order to enable others skilled in the art

bar or round, D, a short distance below the top cross-bar, C; and in the center of said round is connected the end of a brace, G, by means of a strap-hinge, *d*. The outer ends of the braces G G are placed between two cleats, *h h*, on the under side of the table-top, near the center, and fastened by means of a pin, *i*, passing through them, as shown, thereby holding and bracing the legs in proper position to support the table. When the legs are folded against the under side of the table-top the braces G G also lie against the same, between the cleats *h h*, and are fastened by the pin *i* passing through, each brace being provided with two holes, *x x*, at suitable points for the

**Figure 1: Example USPTO scan of a patent (number 159,227)**

| **Google Patents OCR** | **Our New OCR** |
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| Specification formingpart of Letters Patent | Specification forming part of Letters Patent |
| No. 159,227. dated January 26, li875; | No. 159,227, dated January 26, 1875 |
| application-filed October 21, 1874. f f | application filed October 21, 1874. |
| runstnnctionjindrrangement ofraQldiUE. tn- | To all whom it may concern: |
| | Be it known that I, HEMAN B. SINCLAIR, |
| | of the city of Chicago, in the county of Cook |
| | and State of Illinois, have invented certain |
| | new and useful Improvements in Folding Ta- |
| | bles; and I do hereby declare that the fol- |
| | lowing is a full, clear, and exact description |
| | thereof, reference being had to the accompa- |
| | nying drawings and to the letters of reference |
| | marked thereon, which form part of this speci- |
| | fication. |
| | The nature of my invention consists in the |
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**Figure 2: Comparison of OCR rendering of patent number 159,227**



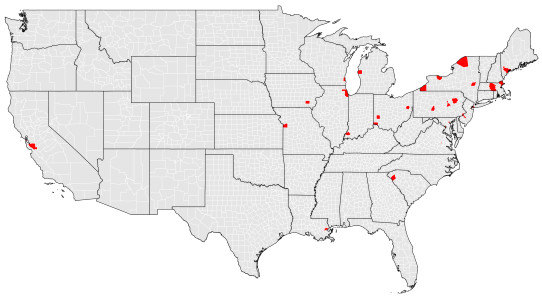

**(A)** 1835–1839 **(B)** 1860–1869

**(C)** 1890–1899 **(D)** 1920–1929

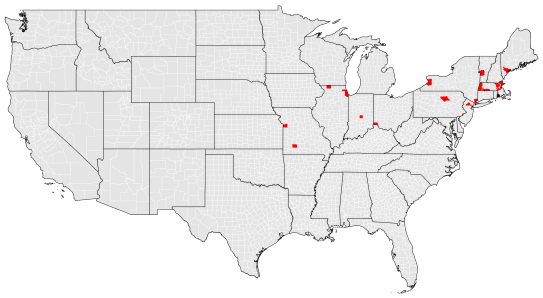
**(E)** 1950–1959 **(F)** 2010–2019

## Figure 3: Locations of inventors

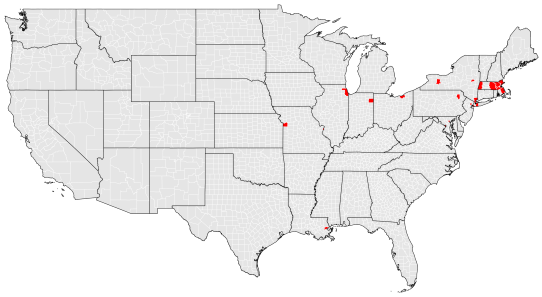
This figure shows the location of inventors who have patents granted during different periods over the last two centuries. Inventor locations are aggregated to counties that are then sorted into deciles in each period based on the number of inventors residing in each county. The darker the shade, the greater the number of active inventors in a county.



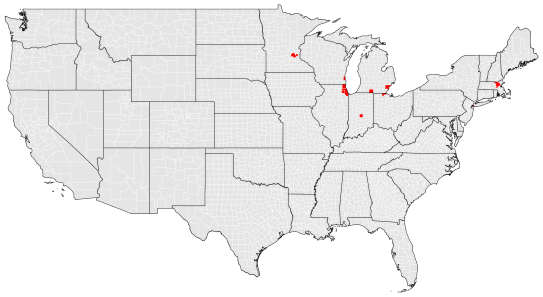
**(A)** SS White Dental Manufacturing Company



**(B)** The American Bell Telephone Company



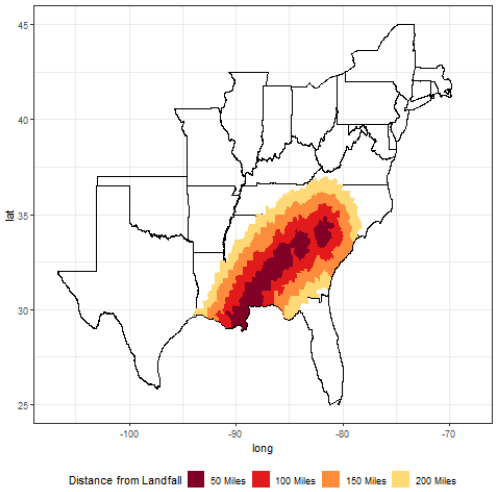
**(C)** General Electric Company



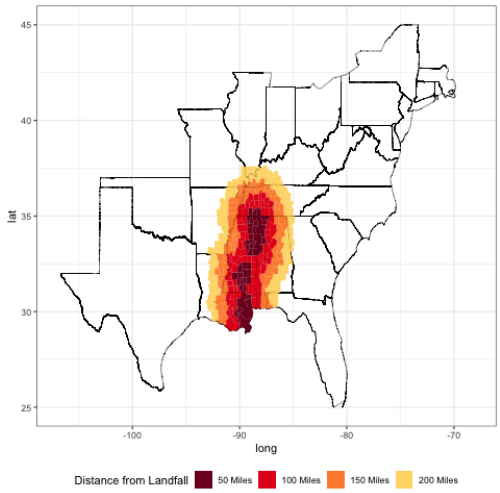
**(D)** Western Electric Company

**Figure 4: Firms' internal innovation networks as of 1900**

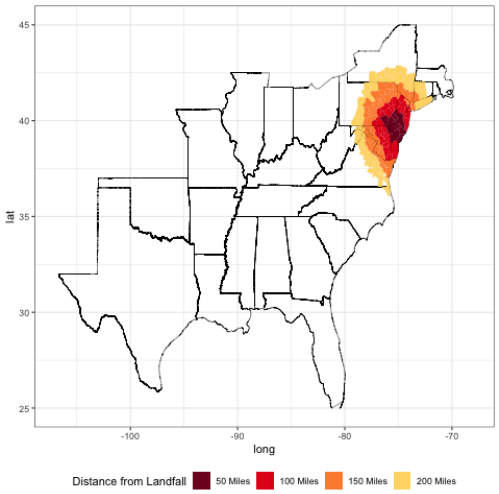
This figure shows for each firm the counties where they generated patents up to 1900. These four firms are in the top five in generated patents up to 1900: SS White Dental Manufacturing Company (216 cumulative patents), The American Bell Telephone Company (266 cumulative patents), General Electric Company 581 (cumulative patents), and Western Electric Company (587 cumulative patents).



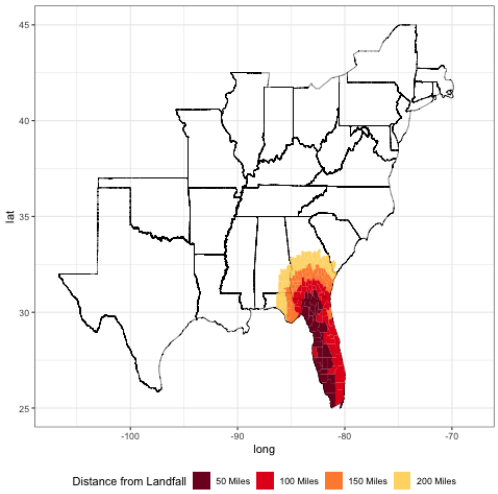

**(A)** 1893 Louisiana Hurricane




**(B)** 2005 Katrina



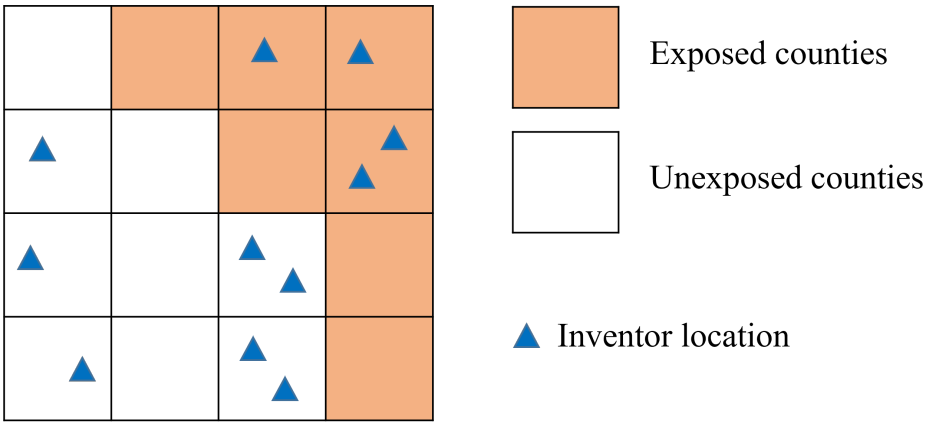

**(C)** 2012 Sandy



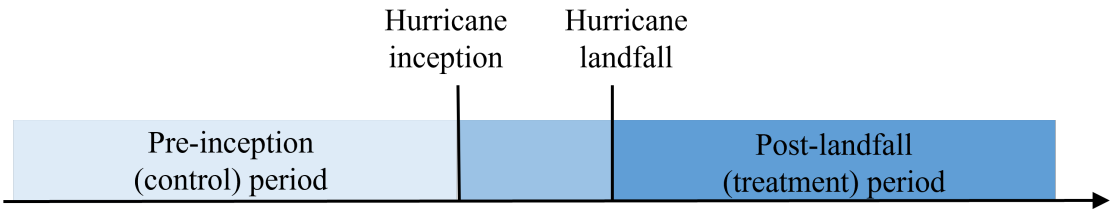

**(D)** 2017 Irma

## Figure 5: Counties within a hurricane landfall region

This figure shows the counties within 50, 100, 150, and 200 miles of the hurricane eye for four hurricanes in our sample from 1851-2023.



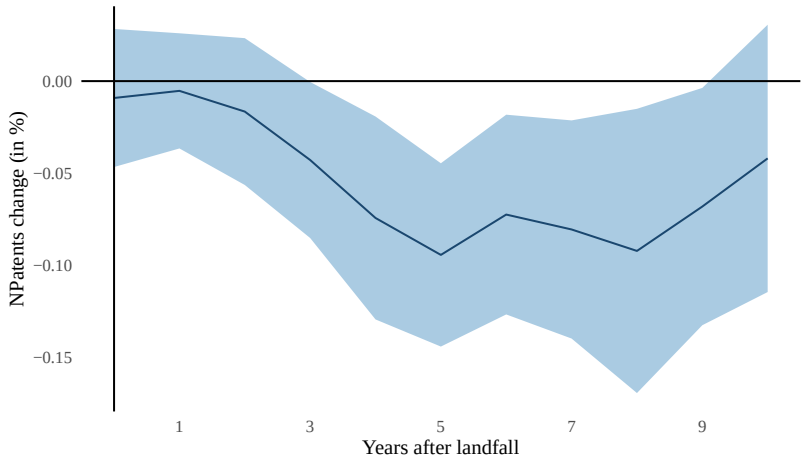

**(A)** Stylized example of spatial exposure (across variation)



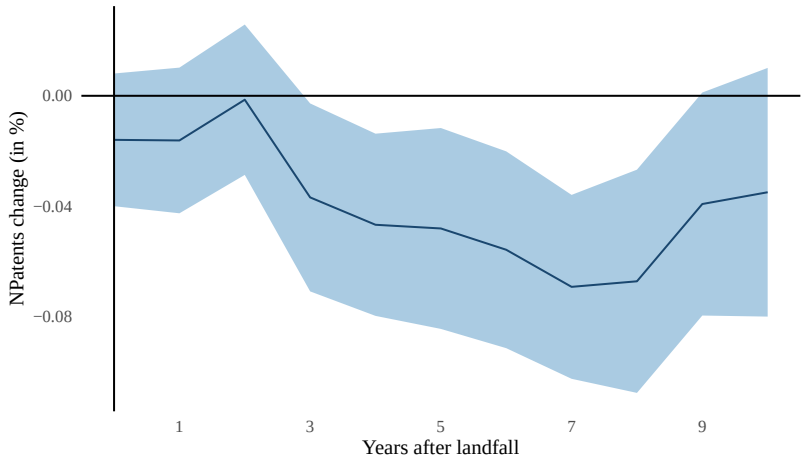

**(B)** The timeline of a hurricane (within variation)

**Figure 6: Identification strategy**

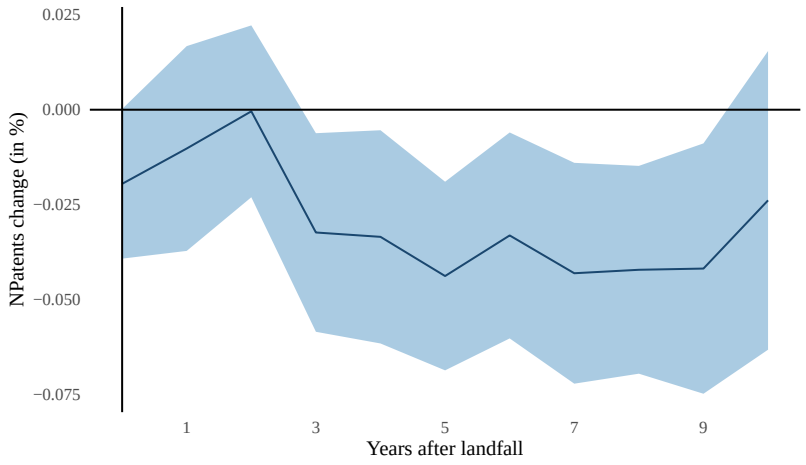
Panel A illustrates a stylized example of inventor locations and county exposure to a hurricane landfall region. Panel B illustrates the timeline of a hurricane.




**(A)** Counties within 50-mile radius




**(B)** Counties within 100-mile radius




**(C)** Counties within 200-mile radius

## Figure 7: Change in local innovation following landfall

This figure presents the $\beta_1$ estimate and its 95% confidence interval from estimating the local projection estimation described in equation (1) for horizons $h = 0, 1, 2, \ldots, 9, 10$. Panels A, B, and C show the estimates with landfall regions determined at 50-, 100- and 200-miles radii, respectively, from the eye of the hurricane.

## Table 1: Patent Summary Statistics

This table shows the summary statistics for the main variables used in the paper. Panel A presents the data from 1851 to 2023 used in the county-year level local projection model given in equation (1). Panel B presents the summary statistics for the firm-county-year level event study regression given in equation (2). Panel C reports the summary statistics for the firm-level event study regression given in equation (7). Panel B and C show summary statistics for the landfall region based on a 100-mile radius around the eye of the hurricane and five years pre- and post-hurricane landfall. All variables are described in Appendix Table C.4.

**Panel A: County-Year Level**

| | Observations | Avg. | St. dev. | $10^{\text{th}}$ | $25^{\text{th}}$ | $50^{\text{th}}$ | $75^{\text{th}}$ | $90^{\text{th}}$ |
|---|---|---|---|---|---|---|---|---|
| $NPatents_{c,t}$ | 514,753 | 19.337 | 185.338 | 0.000 | 0.000 | 1.000 | 4.000 | 17.000 |
| $\Delta log(NPatents_{c,t})$ | 226,493 | 0.014 | 0.633 | -0.693 | -0.336 | 0.000 | 0.375 | 0.693 |

**Panel B: Firm-County-Year Level**

| | | | | | | | | |
|---|---|---|---|---|---|---|---|---|
| $\overline{NPatents}_{i,c,t:t+4}$ | 394,053 | 2.035 | 12.915 | 0.200 | 0.200 | 0.400 | 1.200 | 3.200 |
| $\overline{NPatents}_{i,c,t-5:t-1}$ | 394,053 | 1.721 | 10.520 | 0.200 | 0.200 | 0.400 | 1.000 | 2.600 |
| $log(\overline{NPatents}_{i,c,t:t+4}/\overline{NPatents}_{i,c,t-5:t-1})$ | 394,053 | 0.079 | 0.961 | -1.099 | -0.560 | 0.000 | 0.693 | 1.322 |
| $Hit_{c,t}$ | 394,053 | 0.093 | 0.290 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| $HitOther_{i,c,t}$ | 394,053 | 0.051 | 0.162 | 0.000 | 0.000 | 0.000 | 0.001 | 0.100 |

**Panel C: Firm-Year Level**

| | | | | | | | | |
|---|---|---|---|---|---|---|---|---|
| $\overline{NPatents}_{i,t:t+4}$ | 184,163 | 5.030 | 55.539 | 0.200 | 0.400 | 0.600 | 1.800 | 5.400 |
| $\overline{NPatents}_{i,t-5:t-1}$ | 184,163 | 4.181 | 45.370 | 0.200 | 0.200 | 0.600 | 1.600 | 4.600 |
| $log(\overline{NPatents}_{i,t:t+4}/\overline{NPatents}_{i,c,t-5:t-1})$ | 184,163 | 0.095 | 1.044 | -1.099 | -0.651 | 0.000 | 0.693 | 1.386 |
| $SharePatentsHit_{i,t}$ | 184,163 | 0.082 | 0.248 | 0.000 | 0.000 | 0.000 | 0.000 | 0.227 |

## Table 2: Hurricane Summary Statistics

This table shows summary statistics for the number of counties that lie in a hurricane landfall region. Panel A presents the summary statistics for the full time series of hurricanes from 1851 to 2023. Panel B presents the summary statistics for the hurricanes from 1851 to 1975—the time series for which we construct the patent data. Panel C presents the summary statistics for hurricanes from 1976 to 2023—the time series for which the USPTO has digitized patent data. The number of counties in the landfall region are shown based on different radii from the eye of the hurricane.

**Panel A: Hurricanes from 1851 to 2023**

| | | # counties in hurricane landfall region | | | | | | |
|---|---|---|---|---|---|---|---|---|
| Landfall radius | Hurricane-years | Avg. | St. dev. | $10^{th}$ | $25^{th}$ | $50^{th}$ | $75^{th}$ | $90^{th}$ |
| 50 | 49.000 | 63.469 | 32.220 | 21.800 | 44.000 | 60.000 | 77.000 | 93.000 |
| 100 | 50.000 | 163.540 | 88.751 | 77.000 | 108.000 | 155.000 | 213.250 | 229.900 |
| 200 | 50.000 | 375.680 | 156.437 | 201.300 | 267.500 | 353.500 | 477.250 | 536.300 |

**Panel B: Hurricanes from 1851 to 1975**

| | | | | | | | | |
|---|---|---|---|---|---|---|---|---|
| 50 | 36.000 | 63.361 | 35.090 | 21.500 | 43.750 | 60.500 | 77.000 | 92.000 |
| 100 | 37.000 | 161.892 | 97.231 | 73.400 | 99.000 | 155.000 | 217.000 | 230.800 |
| 200 | 37.000 | 373.324 | 171.990 | 172.600 | 252.000 | 351.000 | 483.000 | 544.400 |

**Panel C: Hurricanes from 1976 to 2023**

| | | | | | | | | |
|---|---|---|---|---|---|---|---|---|
| 50 | 13.000 | 63.769 | 23.686 | 41.600 | 52.000 | 60.000 | 77.000 | 91.600 |
| 100 | 13.000 | 168.231 | 61.398 | 102.000 | 129.000 | 170.000 | 195.000 | 226.600 |
| 200 | 13.000 | 382.385 | 105.458 | 263.200 | 280.000 | 387.000 | 440.000 | 528.600 |

## Table 3: Spillovers via Firms' Internal Networks

This table presents results from estimating the regression specification given in equation (2). The dependent variable $\log\left(\frac{\overline{NPatents}_{i,c,t:t+h-1}}{\overline{NPatents}_{i,c,t-h:t-1}}\right)$ is at the firm-county-year level. The numerator (denominator) is the annual average number of patents over $h$ years post (pre) hurricane landfall. The first independent variable is an indicator variable that takes a value of 1 if a county was in the landfall region of a hurricane in a given year ($Hit_{c,t}$). The dependent variable ($HitOther_{i,c,t}$) measures the share of a firm's patents in counties other than county $c$ that were in the hurricane landfall region. The data span from 1851 to 2023. The specifications include year and county fixed effects (Panel A) and year interacted with county fixed effects (Panel B). Standard errors are double clustered by year and firm and shown in parentheses. The significance of each coefficient estimate is indicated by * for $p < 0.10$, ** for $p < 0.05$, and *** for $p < 0.01$.

**Panel A: With County and Year Fixed Effects**

| Landfall radius | 100 | | | | 200 | | | |
|---|---|---|---|---|---|---|---|---|
| Window length in years ($h$) | 1 | 3 | 5 | 10 | 1 | 3 | 5 | 10 |
| $Hit_{c,t}$ | -0.024*** | -0.028*** | -0.031*** | -0.027 | -0.023*** | -0.030*** | -0.038*** | -0.039* |
| | (0.006) | (0.008) | (0.011) | (0.019) | (0.005) | (0.008) | (0.012) | (0.022) |
| $HitOther_{i,c,t}$ | 0.054*** | 0.076*** | 0.101*** | 0.162*** | 0.058*** | 0.064*** | 0.100*** | 0.172*** |
| | (0.019) | (0.021) | (0.025) | (0.040) | (0.015) | (0.019) | (0.022) | (0.036) |
| Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| County FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 232,955 | 406,105 | 451,265 | 384,621 | 210,375 | 407,701 | 409,129 | 346,958 |
| $R^2$ | 0.012 | 0.015 | 0.017 | 0.023 | 0.013 | 0.015 | 0.017 | 0.024 |

**Panel B: With County×Year Fixed Effects**

| | | | | | | | | |
|---|---|---|---|---|---|---|---|---|
| $HitOther_{i,c,t}$ | 0.059*** | 0.072*** | 0.095*** | 0.154*** | 0.063*** | 0.069*** | 0.104*** | 0.168*** |
| | (0.020) | (0.022) | (0.027) | (0.043) | (0.016) | (0.020) | (0.023) | (0.038) |
| County × Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 232,955 | 406,105 | 451,265 | 384,621 | 210,375 | 407,701 | 409,129 | 346,958 |
| $R^2$ | 0.094 | 0.071 | 0.072 | 0.083 | 0.104 | 0.075 | 0.074 | 0.086 |

## Table 4: Spillovers via Firms' Internal Networks (Excluding Hit Counties)

This table presents results from estimating the regression specification given in equation (2) but excluding hit counties from the sample. The dependent variable $\log\left(\frac{\overline{NPatents}_{i,c,t:t+h-1}}{\overline{NPatents}_{i,c,t-h:t-1}}\right)$ is at the firm-county-year level. The numerator (denominator) is the annual average number of patents over $h$ years post (pre) hurricane landfall. The dependent variable ($HitOther_{i,c,t}$) measures the share of a firm's patents in counties other than county $c$ that were in the hurricane landfall region. The data span from 1851 to 2023. The specifications include county times year fixed effects. Standard errors are double clustered by year and firm and shown in parentheses. The significance of each coefficient estimate is indicated by * for $p < 0.10$, ** for $p < 0.05$, and *** for $p < 0.01$.

| Landfall radius | 100 | | | | 200 | | | |
|---|---|---|---|---|---|---|---|---|
| Window length in years ($h$) | 1 | 3 | 5 | 10 | 1 | 3 | 5 | 10 |
| $HitOther_{i,c,t}$ | 0.054** | 0.066** | 0.087** | 0.124** | 0.052** | 0.051* | 0.087*** | 0.128*** |
| | (0.023) | (0.032) | (0.035) | (0.047) | (0.022) | (0.028) | (0.028) | (0.041) |
| County × Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 192,735 | 369,023 | 357,472 | 295,421 | 185,009 | 355,942 | 346,952 | 284,358 |
| $R^2$ | 0.100 | 0.072 | 0.073 | 0.086 | 0.102 | 0.074 | 0.075 | 0.089 |

## Table 5: Spillovers via Firms' Internal Networks (Pre-1976 Period)

This table presents results from estimating the regression specification given in equation (2) but interacting the independent variable with an indicator that takes the value of one for the years from 1851 to 1975 and zero otherwise. The dependent variable $\log\left(\frac{\overline{NPatents}_{i,c,t:t+h-1}}{\overline{NPatents}_{i,c,t-h:t-1}}\right)$ is at the firm-county-year level. The numerator (denominator) is the annual average number of patents over $h$ years post (pre) hurricane landfall. The dependent variable ($HitOther_{i,c,t}$) measures the share of a firm's patents in counties other than county $c$ that were in the hurricane landfall region. The data span from 1851 to 2023. The specifications include county times year fixed effects. Standard errors are double clustered by year and firm and shown in parentheses. The significance of each coefficient estimate is indicated by * for $p < 0.10$, ** for $p < 0.05$, and *** for $p < 0.01$.

| Landfall radius | 100 | | | | 200 | | | |
|---|---|---|---|---|---|---|---|---|
| Window length in years ($h$) | 1 | 3 | 5 | 10 | 1 | 3 | 5 | 10 |
| $HitOther_{i,c,t}$ | 0.041** | 0.059** | 0.070** | 0.111** | 0.049*** | 0.053** | 0.079*** | 0.115*** |
| | (0.018) | (0.025) | (0.029) | (0.048) | (0.017) | (0.023) | (0.024) | (0.038) |
| $HitOther_{i,c,t} \times I_{t\in 1851-1975}$ | 0.128*** | 0.079** | 0.134*** | 0.177*** | 0.095*** | 0.086** | 0.127*** | 0.191*** |
| | (0.025) | (0.036) | (0.036) | (0.058) | (0.033) | (0.035) | (0.039) | (0.054) |
| County × Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 207,623 | 398,787 | 394,053 | 333,573 | 210,375 | 407,701 | 409,129 | 346,958 |
| $R^2$ | 0.010 | 0.071 | 0.072 | 0.082 | 0.104 | 0.074 | 0.074 | 0.085 |

## Table 6: Spillovers via Firms' Internal Networks (during US Recessions)

This table presents results from estimating the regression specification given in equation (2) but interacting the independent variable with an indicator that takes the value of one if the year was a US recession year and zero otherwise. A year is classified as a recession year if at least six months were in a recession based on NBER data. The dependent variable $\log\left(\frac{\overline{NPatents}_{i,c,t:t+h-1}}{\overline{NPatents}_{i,c,t-h:t-1}}\right)$ is at the firm-county-year level. The numerator (denominator) is the annual average number of patents over $h$ years post (pre) hurricane landfall. The dependent variable ($HitOther_{i,c,t}$) measures the share of a firm's patents in counties other than county $c$ that were in the hurricane landfall region. The data span from 1851 to 2023. The specifications include county times year fixed effects. Standard errors are double clustered by year and firm and shown in parentheses. The significance of each coefficient estimate is indicated by * for $p < 0.10$, ** for $p < 0.05$, and *** for $p < 0.01$.

| Landfall radius | 100 | | | | 200 | | | |
|---|---|---|---|---|---|---|---|---|
| Window length in years ($h$) | 1 | 3 | 5 | 10 | 1 | 3 | 5 | 10 |
| $HitOther_{i,c,t}$ | 0.049** | 0.067*** | 0.084*** | 0.135*** | 0.059*** | 0.063*** | 0.095*** | 0.146*** |
| | (0.019) | (0.023) | (0.028) | (0.045) | (0.016) | (0.021) | (0.023) | (0.036) |
| $HitOther_{i,c,t} \times I_{t \in Recession}$ | 0.110*** | 0.064 | 0.102*** | 0.147*** | 0.069* | 0.084*** | 0.121*** | 0.199*** |
| | (0.032) | (0.047) | (0.033) | (0.042) | (0.035) | (0.029) | (0.023) | (0.033) |
| County × Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 207,623 | 398,697 | 394,029 | 333,595 | 210,375 | 407,645 | 409,171 | 346,958 |
| $R^2$ | 0.099 | 0.072 | 0.072 | 0.083 | 0.104 | 0.075 | 0.075 | 0.086 |

## Table 7: Probability of Firm Keeping Existing Innovation Location

This table presents results from estimating the logit regression specification given in equation (4). The dependent variable takes a value of one if the firm has patents in a county over $h$ years post-hurricane landfall and zero otherwise. The sample includes only firm-county-year observations where the number of patents was positive over $h$ years pre-hurricane. The dependent variable ($HitOther_{i,c,t}$) measures the share of a firm's patents in counties other than county $c$ that were in the hurricane landfall region. The data span from 1851 to 2023. The specifications include county times year fixed effects. Standard errors are double clustered by year and firm and shown in parentheses. The significance of each coefficient estimate is indicated by * for $p < 0.10$, ** for $p < 0.05$, and *** for $p < 0.01$.

| Landfall radius | 100 | | | | 200 | | | |
|---|---|---|---|---|---|---|---|---|
| Window length in years ($h$) | 1 | 3 | 5 | 10 | 1 | 3 | 5 | 10 |
| $HitOther_{i,c,t}$ | 0.523*** | 0.426*** | 0.400*** | 0.403*** | 0.534*** | 0.435*** | 0.408*** | 0.403*** |
| | (0.048) | (0.054) | (0.059) | (0.063) | (0.054) | (0.055) | (0.056) | (0.060) |
| Marginal effect | 0.118 | 0.097 | 0.087 | 0.079 | 0.120 | 0.099 | 0.089 | 0.079 |
| County × Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 541,060 | 1,037,433 | 1,157,467 | 1,185,779 | 553,164 | 1,060,368 | 1,192,796 | 1,211,923 |
| Pseudo $R^2$ | 0.040 | 0.036 | 0.035 | 0.032 | 0.041 | 0.037 | 0.035 | 0.033 |

## Table 8: Firms' Expansion Into New Counties

This table presents results from estimating the firm-level regression specification given in equation (5). The dependent variable is the number of new counties in which the firm generated patents over $h$ years post-hurricane landfall divided by the total number of counties in which the firm generated patents over $h$ years pre-hurricane landfall. The dependent variable ($SharePatentsHit_{i,t}$) measures the share of a firm's patents in counties that are in the hurricane landfall region. The data span from 1851 to 2023. The specifications include year fixed effects. Standard errors are double clustered by year and firm and shown in parentheses. The significance of each coefficient estimate is indicated by * for $p < 0.10$, ** for $p < 0.05$, and *** for $p < 0.01$.

| Landfall radius | 100 | | | | 200 | | | |
|---|---|---|---|---|---|---|---|---|
| Window length in years ($h$) | 1 | 3 | 5 | 10 | 1 | 3 | 5 | 10 |
| $SharePatentsHit_{i,t}$ | 0.024** | 0.021** | 0.029** | 0.005 | 0.033*** | 0.031*** | 0.036*** | 0.013 |
| | (0.010) | (0.009) | (0.012) | (0.010) | (0.009) | (0.009) | (0.010) | (0.008) |
| Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 232,955 | 406,105 | 451,265 | 384,621 | 210,375 | 407,701 | 409,129 | 346,958 |
| $R^2$ | 0.094 | 0.071 | 0.072 | 0.083 | 0.104 | 0.075 | 0.074 | 0.086 |

## Table 9: Within-Firm Substitution After Hurricane Shock

This table presents results from estimating the regression specification given in equation (7). The dependent variable $\log\left(\frac{\overline{NPatents}_{i,t:t+h-1}}{\overline{NPatents}_{i,t-h:t-1}}\right)$ is at the firm-year level. The numerator (denominator) is the annual average number of patents over $h$ years post (pre) hurricane landfall. The dependent variable ($SharePatentsHit_{i,t}$) measures the share of a firm's patents in counties that are in the hurricane landfall region. The data span from 1851 to 2023. The specifications include year fixed effects. Standard errors are double clustered by year and firm and shown in parentheses. The significance of each coefficient estimate is indicated by * for $p < 0.10$, ** for $p < 0.05$, and *** for $p < 0.01$.

| Landfall radius | 100 | | | | 200 | | | |
|---|---|---|---|---|---|---|---|---|
| Window length in years ($h$) | 1 | 3 | 5 | 10 | 1 | 3 | 5 | 10 |
| $SharePatentsHit_{i,t}$ | -0.020* | -0.029*** | -0.029** | -0.050*** | -0.002 | -0.012 | -0.005 | -0.031*** |
| | (0.012) | (0.010) | (0.011) | (0.010) | (0.008) | (0.011) | (0.011) | (0.010) |
| Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 93,582 | 176,664 | 184,163 | 178,620 | 98,079 | 183,568 | 192,548 | 185,843 |
| $R^2$ | 0.0063 | 0.0101 | 0.0127 | 0.0177 | 0.0060 | 0.0103 | 0.0133 | 0.0173 |

## Table 10: Aggregate County Patents Spillovers via Firms' Internal Networks

This table presents results from estimating the regression specification given in equation (8). The dependent variable $\log\left(\frac{\overline{NPatents}_{c,t:t+h-1}}{NPatents_{c,t-1}}\right)$ is at the firm-year level. The numerator is the annual average number of patents over $h$ years post-hurricane landfall. The denominator is the number of patents in the year before landfall. The dependent variable ($CountyHitOther_{c,t}$) measures for firms located in county $c$ the share of their patents that are outside county $c$ and in the hurricane landfall region given in equation (8). Only counties outside the landfall region are included in the regression. The data span from 1851 to 2023. The specifications include county, year, and year interacted with state fixed effects. Standard errors are double clustered by year and county and shown in parentheses. The significance of each coefficient estimate is indicated by * for $p < 0.10$, ** for $p < 0.05$, and *** for $p < 0.01$.

**Panel A: With County and Year Fixed Effects**

| Landfall radius | 100 | | | | 200 | | | |
|---|---|---|---|---|---|---|---|---|
| Window length in years ($h$) | 1 | 3 | 5 | 10 | 1 | 3 | 5 | 10 |
| $CountyHitOther_{c,t}$ | 0.047 | 0.138*** | 0.139*** | 0.176*** | 0.045 | 0.123*** | 0.119*** | 0.150*** |
| | (0.041) | (0.043) | (0.038) | (0.049) | (0.032) | (0.029) | (0.029) | (0.037) |
| Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| County FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 35,821 | 36,128 | 34,174 | 30,731 | 30,460 | 30,818 | 29,520 | 26,466 |
| $R^2$ | 0.371 | 0.489 | 0.548 | 0.614 | 0.376 | 0.498 | 0.558 | 0.628 |

**Panel B: With County and State×Year Fixed Effects**

| | | | | | | | | |
|---|---|---|---|---|---|---|---|---|
| $CountyHitOther_{c,t}$ | 0.026 | 0.098** | 0.119*** | 0.169*** | 0.032 | 0.105*** | 0.122*** | 0.147*** |
| | (0.039) | (0.047) | (0.042) | (0.037) | (0.033) | (0.033) | (0.031) | (0.032) |
| Year×State FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| County FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Controls | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 35,821 | 36,128 | 34,174 | 30,731 | 30,460 | 30,818 | 29,520 | 26,466 |
| $R^2$ | 0.418 | 0.539 | 0.598 | 0.666 | 0.426 | 0.550 | 0.610 | 0.680 |

# Appendix A Conceptual framework

Consider a firm operating $J$ establishments indexed by $j = 1, \ldots, J$. Each establishment can spend on R&D. The variable $\varphi_j \geq 0$ is the level of R&D inputs at an establishment, for example, labor, plants, and equipment. Each establishment takes the price of the R&D input, $p_j > 0$, as given. Whether or not an establishment $j$ is in a disaster-hit region is captured by the parameter $\theta_j$. An establishment hit by a disaster has $\theta_j > 0$, otherwise $\theta_j = 0$.

Each establishment $j$ has a capital endowment:

$$K_j(\theta_j) = \bar{K}_j \, e^{-\xi\theta_j}, \qquad \bar{K}_j > 0, \ \xi > 0. \tag{10}$$

The capital endowment decreases when an establishment is in the hurricane landfall region. The interpretation is that the establishment can experience physical damages from the hurricane, which reduces the firm's overall capital. The hurricane can destroy, for example, facilities, equipment, and inventory. The parameter $\xi$ determines the sensitivity of an establishment's capital to the hurricane. If $\xi$ is large, the capital is more vulnerable to the hurricane hit, for example, the firm has no insurance or has less resilient facilities. Its derivative with respect to the shock is

$$K_j'(\theta_j) = -\xi \, \bar{K}_j e^{-\xi\theta_j} < 0. \tag{11}$$

We call this effect from the hurricane the *capital channel*.

Given a level of R&D input $\varphi_j$ and input productivity $\gamma$, the innovation output of establishment $j$, $\pi_j$, is given by

$$\pi_j = e^{-\theta_j}\varphi_j^{\gamma}, \qquad 0 < \gamma < 1. \tag{12}$$

The derivatives of $\pi$ are

$$\frac{\partial \pi}{\partial \varphi} = e^{-\theta_j}\gamma\varphi_j^{\gamma-1} > 0, \tag{13}$$

$$\frac{\partial^2 \pi}{\partial \varphi^2} = e^{-\theta_j}\gamma(\gamma-1)\varphi_j^{\gamma-2} < 0, \tag{14}$$

$$\frac{\partial^2 \pi}{\partial \varphi \, \partial \theta} = -e^{-\theta_j}\gamma\varphi_j^{\gamma-1} < 0. \tag{15}$$

The sign of these derivatives show that higher R&D input for establishment $j$ increases innovation (13) but at a decreasing rate (14). If the establishment is hit by a hurricane, the marginal product of an additional unit of R&D input is lower (15). In other words,

establishments become less productive at innovation following a hurricane landfall. The worsening productivity is motivated by the large decrease in patent generation for counties in a hurricane landfall region reported in Section 4.1.1. There is less innovation activity in these counties because of, for example, damaged county-level infrastructure or a labor force that is pre-occupied with the aftermath of the hurricane, which lowers the productivity of establishments located there. We call this the *innovation productivity channel*.

Importantly, the firm's headquarters make the decisions on factor input choices and funding for all establishments. The headquarters' goal is to maximize overall firm value as modeled by, for example, Williamson (1975); Gertner, Scharfstein, and Stein (1994); Stein (1997). As in Giroud and Mueller (2019), the headquarters pool the endowments of all the establishments to form the total capital endowment for the firm.

$$\bar{K}(\theta_1, \ldots, \theta_J) = \sum_{j=1}^{J} K_j(\theta_j) = \sum_{j=1}^{J} \bar{K}_j e^{-\xi\theta_j}. \tag{16}$$

The budget constraint for R&D expenditures is:

$$\sum_{j=1}^{J} p_j \, \varphi_j = \bar{K}(\theta_1, \ldots, \theta_J). \tag{17}$$

The firm allocates its capital endowment to maximize total innovation:[13]

$$\max_{\varphi_1,\ldots,\varphi_J \geq 0} \sum_{j=1}^{J} e^{-\theta_j} \varphi_j^{\gamma} + \lambda \Big( \sum_{j=1}^{J} \bar{K}_j e^{-\xi\theta_j} - \sum_{j=1}^{J} p_j \varphi_j \Big). \tag{18}$$

The first order conditions are

$$e^{-\theta_j} \gamma \varphi_j^{\gamma-1} = \lambda p_j, \ \forall j, \tag{19}$$

and the budget constraint binds:

$$\sum_{j=1}^{J} p_j \varphi_j = \sum_{j=1}^{J} \bar{K}_j e^{-\xi\theta_j}. \tag{20}$$

Because of the binding budget constraint, the shadow value of capital, $\lambda$, is greater than zero.

[13]The firm allocates all capital to R&D and is only concerned with how much capital to allocate to each establishment. This choice keeps the framework tractable. An extension where the firm also decides how much capital to allocate to R&D in general leads to qualitatively similar predictions.

Based on equation (19), the following equality has to hold for any two establishments $i$ and $j$:

$$\frac{e^{-\theta_i}\gamma\varphi_i^{\gamma-1}}{p_i} = \lambda = \frac{e^{-\theta_j}\gamma\varphi_j^{\gamma-1}}{p_j}. \tag{21}$$

This equation implies that the marginal product of R&D input per dollar of price must be the same across the establishments. We want to determine how R&D input and consequently innovation output changes at establishment $j$ in response to a hurricane making landfall in the location of establishment $k$. The derivative of the R&D input at establishment $j$, $\varphi_j$, with respect to the hurricane hit of establishment $k$ is given by

$$\frac{\partial\varphi_j}{\partial\theta_k} = \frac{p_j\, e^{\theta_j}\, \varphi_j^{2-\gamma}}{\sum_{m=1}^{J} p_m^2\, e^{\theta_m}\, \varphi_m^{2-\gamma}} \left[-\xi\, \bar{K}_k e^{-\xi\theta_k} + \frac{p_k\, \varphi_k}{1-\gamma}\right], \qquad (j \neq k). \tag{22}$$

Because of the two channels: capital and innovation productivity, the sign of the derivative can be negative or positive. In other words, establishment $k$ being hit by the hurricane can increase or decrease innovation at establishment $j$. The reason for this can be seen from the term in brackets in equation (22). The capital endowment channel is captured by the first bracket term $-\xi\, \bar{K}_k e^{-\xi\theta_k} < 0$. The hurricane shrinks the total R&D budget of the firm, which reduces innovation across the establishments. This mechanism is consistent with the theoretical and empirical results in Giroud and Mueller (2019), who find that regional shocks from the house price collapses during the Great Recession decrease employment of establishments in this region but also for the same firms' establishments in other regions. Importantly, the innovation productivity channel is captured by the second bracket term $(p_k\, \varphi_k)/(1-\gamma) > 0$. The hurricane makes establishment $k$ a less productive establishment to innovate, which leads the firm to reallocate R&D resources to other establishments.[14]

The term outside of the bracket is a weight that captures how sensitive innovation at establishment $j$ is to establishment $k$ being hit by the hurricane. The numerator is the relative steepness of each establishment's marginal innovation productivity. When establishment $k$ becomes less productive, the firm must reallocate capital so that the marginal productivity is the same across all plants. This effect operates through a drop in the shadow value of capital, $\lambda$, based on equation (22). Plants that operate at steeper marginal productivity curves will see larger adjustments to their R&D inputs. Similarly, when the capital at establishment $k$

[14] In the model, the lower innovation productivity at establishment $k$ because of the hurricane does not affect the price of R&D input, $p_k$. This choice is motivated by sticky prices but also by the fact that a hurricane is only a regional event and part of the R&D input, for example, equipment is tradable across regions. Empirically, if the lower productivity led to a reduction in $p_k$, the capital channel would be more likely to dominate the innovation productivity channel.

decreases, the shadow value of capital increases. Establishments with a high R&D input, $\varphi$, and R&D input price, $p$, will experience a larger decrease in $\varphi$ to achieve a sufficiently high marginal productivity.

Whether the capital or the innovation productivity dominates—whether the firm reduces R&D input at all establishments or reallocates R&D input to establishments not hit by the hurricane—is ultimately an empirical question.

## A.1 Derivative with respect to hurricane shock

In this section, we show how to derive equation (22). For $j \neq k$, differentiating (21) with respect to $\theta_k$ results in

$$e^{-\theta_j}\gamma(\gamma-1)\varphi_j^{\gamma-2}\frac{\partial \varphi_j}{\partial \theta_k} = p_j\frac{\partial \lambda}{\partial \theta_k}, \tag{23}$$

hence

$$\frac{\partial \varphi_j}{\partial \theta_k} = \frac{p_j e^{\theta_j}}{\gamma(\gamma-1)}\varphi_j^{2-\gamma}\frac{\partial \lambda}{\partial \theta_k}. \tag{24}$$

For $j = k$, differentiating (21) with respect to $\theta_k$ results in

$$e^{-\theta_k}\gamma(\gamma-1)\varphi_k^{\gamma-2}\frac{\partial \varphi_k}{\partial \theta_k} - e^{-\theta_k}\gamma\varphi_k^{\gamma-1} = p_k\frac{\partial \lambda}{\partial \theta_k}. \tag{25}$$

Next, differentiate the budget constraint with respect to $\theta_k$:

$$\sum_{j=1}^{J} p_j\frac{\partial \varphi_j}{\partial \theta_k} = -\xi\bar{K}_k e^{-\xi\theta_k}. \tag{26}$$

Substitute in (24) and (25) and solve for $\partial\lambda/\partial\theta_k$:

$$\frac{\partial \lambda}{\partial \theta_k} = \frac{-\xi\bar{K}_k e^{-\xi\theta_k} + \frac{p_k\varphi_k}{1-\gamma}}{\sum_{m=1}^{J} p_m^2 e^{\theta_m}\varphi_m^{2-\gamma}/[\gamma(\gamma-1)]}. \tag{27}$$

Using $\partial\lambda/\partial\theta_k$ in (24) and simplifying yields the derivative in (22).

# Appendix B Constructing the patents database

We extract information from pre-1976 patents in three steps. First, we identify the title page of each patent document. Second, we perform OCR (optimal character recognition) on the title page of each patent to obtain the text of each title page. Third, we use ChatGPT to extract inventor and assignee names and locations.

Combining OCR and LLM technologies, we construct a new database of patents.[15] Our process is as follows:

1. Identify the title-page in each patent document. Beginning in the early 1970s, the first page of a patent document is the title page and contains relevant biographical and other information. Prior to that, the initial pages of a patent were often one or more pages of drawings, so the first page of text appears later.
2. Apply state-of-the-art OCR on the relevant PDF pages. We extract the layout of the document, making it possible to, for example, read columns of text in the correct order.
3. Use ChatGPT to extract information from the OCR text. To date, we have extracted inventor and assignee names and locations. We do this using a fine-tuned version of the GPT model and use OpenAI's Python API to submit millions of batch requests.

Consider again the patent shown in Figure 1; the OCR output we obtain is shown in the right column of Table 2. The quality improvement over the Google Patents version of the text is readily apparent; this rendering perfectly matches the original patent document.

Next, we use ChatGPT to extract the relevant information from the enhanced text. In particular, we use automated queries with questions about the text, as shown here along with ChatGPT's corresponding answers:

| **Input question** | **ChatGPT response** |
|---|---|
| 1) What is the title of this patent document? | IMPROVEMENT IN FOLDING TABLES |
| 2) How many inventors are listed? | 1 |
| 3) What is the name of each inventor? | Heman B. Sinclair |
| 4) What city/state is each inventor from? | Chicago, IL |
| 5) If the text mentions an "assignor" or "assignee," to what person or company was it assigned? | null |
| 6) And if it was assigned, list any corresponding cities. | null |

Given the improved quality of the OCR, one might wonder whether ChatGPT provides a significant benefit over simpler text analysis approaches. Determining the name and

[15] To date, we have extracted information only for the title page of patent documents, which covers the data required for this paper. We are in the process of expanding these data to create a complete and highly accurate historical database.

city/state of the inventor from this high-quality OCR text may appear to be quite straightforward using simple regular expression pattern matching. But ChatGPT can seamlessly handle far more complicated situations involving multiple inventors from different cities, multiple assignees, significant changes in formatting, and other edge cases that are hard to anticipate across millions of patents.

Our results were obtained using a recent ChatGPT model, `gpt-4o-mini`, which OpenAI describes as an "affordable and intelligent small model for fast, lightweight tasks." Out-of-the-box, this model performed quite well on our questions. We are able to elicit even higher-quality responses over a range of input types with two approaches. First, we adopt recent advances in prompt engineering. For example, we instruct the model to "take it step by step" before answering. Despite its apparent simplicity, this instruction has been shown to provide a significant improvement of an LLM's ability to "reason" through certain types of questions (Kojima, Gu, Reid, Matsuo, and Iwasawa, 2024). This approach is especially useful in patents with multiple inventors; in these patents we see an improvement in ChatGPT's answers once we require it *first* to count how many inventors there are, and *then* to name them and identify their locations.

Second, we further improve the output by "fine-tuning" the model to our particular needs. This is done by providing the model with additional training examples of questions along with our desired output; the model then learns to adjust its output to match the target. Fine-tuning alters the learning environment from "zero-shot" to "few-shot" by showing the LLM what an appropriate response looks like. After providing 100 fine-tuning examples, we verify that our model performs exceptionally well. We manually checked the accuracy of the extracted inventor and assignee names and locations of a random sample of 500 patents by comparing them to the patent documents. Only one patent had a minor discrepancy, indicating an error rate of 0.2%, which can likely be reduced further with additional fine-tuning of the model.

# Appendix C Additional tables

**Table C.1: Baseline Local Effects - Local Projection Estimation**

This table presents results from estimating the regression specification given in equation (1) for horizons $h = 0, 1, 2, \ldots, 9, 10$. Panels A, B, C show the estimates with landfall regions determined at 50-, 100- and 200-miles, respectively, from the radius of the eye of a hurricane. The data span from 1851 to 2023. Standard errors are double clustered by year and county and shown in parentheses. The specifications include county and time fixed effects. The significance of each coefficient estimate is indicated by * for $p < 0.10$, ** for $p < 0.05$, and *** for $p < 0.01$.

**Panel A: Country treatment within 50-mile radius**

| Years since hit | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
|---|---|---|---|---|---|---|---|---|---|---|---|
| $Hit_{c,t}$ | -0.009 | -0.005 | -0.017 | -0.043** | -0.074*** | -0.094*** | -0.072*** | -0.081*** | -0.092** | -0.068** | -0.042 |
| | (0.019) | (0.016) | (0.020) | (0.022) | (0.028) | (0.025) | (0.028) | (0.030) | (0.039) | (0.033) | (0.037) |
| $NPatents_{c,t-1}$ | -0.677*** | -0.700*** | -0.721*** | -0.730*** | -0.737*** | -0.750*** | -0.749*** | -0.765*** | -0.760*** | -0.771*** | -0.774*** |
| | (0.005) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) | (0.007) | (0.007) | (0.007) | (0.007) |
| $NPatents_{c,t-2}$ | 0.206*** | 0.190*** | 0.190*** | 0.185*** | 0.174*** | 0.178*** | 0.169*** | 0.172*** | 0.160*** | 0.159*** | 0.155*** |
| | (0.004) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) |
| $NPatents_{c,t-3}$ | 0.138*** | 0.143*** | 0.141*** | 0.134*** | 0.136*** | 0.130*** | 0.131*** | 0.126*** | 0.124*** | 0.122*** | 0.114*** |
| | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) |
| $NPatents_{c,t-4}$ | 0.112*** | 0.120*** | 0.113*** | 0.118*** | 0.111*** | 0.116*** | 0.106*** | 0.110*** | 0.105*** | 0.098*** | 0.103*** |
| | (0.004) | (0.004) | (0.004) | (0.004) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) |
| $NPatents_{c,t-5}$ | 0.109*** | 0.109*** | 0.118*** | 0.113*** | 0.115*** | 0.109*** | 0.106*** | 0.104*** | 0.102*** | 0.104*** | 0.096*** |
| | (0.004) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.006) | (0.006) | (0.006) |
| Other controls | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| County FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 80,813 | 79,263 | 77,797 | 76,288 | 74,993 | 73,589 | 72,388 | 71,197 | 69,936 | 68,531 | 67,264 |
| R2 | 0.349 | 0.368 | 0.387 | 0.403 | 0.413 | 0.421 | 0.422 | 0.432 | 0.433 | 0.442 | 0.452 |

**Panel B: Landfall Region Based on 100-Mile Radius Around Hurricane Eye**

| Years since hit | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
|---|---|---|---|---|---|---|---|---|---|---|---|
| $Hit_{c,t}$ | -0.016 | -0.016 | -0.001 | -0.037** | -0.047*** | -0.048** | -0.056*** | -0.069*** | -0.067*** | -0.039* | -0.035 |
| | (0.012) | (0.013) | (0.014) | (0.017) | (0.017) | (0.019) | (0.018) | (0.017) | (0.021) | (0.021) | (0.023) |
| $NPatents_{c,t-1}$ | -0.678*** | -0.705*** | -0.724*** | -0.734*** | -0.741*** | -0.752*** | -0.755*** | -0.767*** | -0.764*** | -0.774*** | -0.778*** |
| | (0.005) | (0.005) | (0.006) | (0.006) | (0.005) | (0.006) | (0.006) | (0.006) | (0.007) | (0.006) | (0.007) |
| $NPatents_{c,t-2}$ | 0.203*** | 0.190*** | 0.187*** | 0.183*** | 0.173*** | 0.174*** | 0.169*** | 0.170*** | 0.160*** | 0.158*** | 0.153*** |
| | (0.004) | (0.005) | (0.004) | (0.004) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) |
| $NPatents_{c,t-3}$ | 0.138*** | 0.142*** | 0.143*** | 0.136*** | 0.134*** | 0.131*** | 0.131*** | 0.128*** | 0.124*** | 0.120*** | 0.115*** |
| | (0.005) | (0.004) | (0.004) | (0.004) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) |
| $NPatents_{c,t-4}$ | 0.112*** | 0.119*** | 0.114*** | 0.116*** | 0.113*** | 0.116*** | 0.107*** | 0.109*** | 0.105*** | 0.100*** | 0.103*** |
| | (0.004) | (0.004) | (0.004) | (0.004) | (0.005) | (0.005) | (0.004) | (0.005) | (0.005) | (0.005) | (0.004) |
| $NPatents_{c,t-5}$ | 0.109*** | 0.111*** | 0.117*** | 0.116*** | 0.117*** | 0.109*** | 0.108*** | 0.105*** | 0.101*** | 0.102*** | 0.096*** |
| | (0.004) | (0.004) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) |
| Other controls | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| County FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 92,189 | 90,415 | 88,710 | 86,985 | 85,477 | 83,918 | 82,510 | 81,154 | 79,764 | 78,213 | 76,766 |
| R2 | 0.348 | 0.369 | 0.387 | 0.403 | 0.413 | 0.419 | 0.422 | 0.430 | 0.432 | 0.440 | 0.450 |

**Panel B: Landfall Region Based on 200-Mile Radius Around Hurricane Eye**

| Years since hit | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
|---|---|---|---|---|---|---|---|---|---|---|---|
| $Hit_{c,t}$ | -0.019* | -0.010 | -0.000 | -0.032** | -0.034** | -0.044*** | -0.033** | -0.043*** | -0.042*** | -0.042** | -0.024 |
| | (0.010) | (0.014) | (0.011) | (0.013) | (0.014) | (0.013) | (0.014) | (0.015) | (0.014) | (0.017) | (0.020) |
| $NPatents_{c,t-1}$ | -0.679*** | -0.705*** | -0.724*** | -0.735*** | -0.743*** | -0.751*** | -0.756*** | -0.766*** | -0.763*** | -0.774*** | -0.777*** |
| | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.006) | (0.006) | (0.006) | (0.006) | (0.006) |
| $NPatents_{c,t-2}$ | 0.202*** | 0.192*** | 0.187*** | 0.183*** | 0.175*** | 0.173*** | 0.169*** | 0.170*** | 0.160*** | 0.158*** | 0.153*** |
| | (0.004) | (0.004) | (0.004) | (0.004) | (0.005) | (0.004) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) |
| $NPatents_{c,t-3}$ | 0.139*** | 0.141*** | 0.142*** | 0.138*** | 0.134*** | 0.131*** | 0.133*** | 0.127*** | 0.123*** | 0.119*** | 0.116*** |
| | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.005) | (0.005) | (0.004) | (0.004) | (0.004) | (0.005) |
| $NPatents_{c,t-4}$ | 0.113*** | 0.119*** | 0.115*** | 0.115*** | 0.113*** | 0.117*** | 0.108*** | 0.110*** | 0.105*** | 0.101*** | 0.103*** |
| | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.004) | (0.005) | (0.005) | (0.005) | (0.004) |
| $NPatents_{c,t-5}$ | 0.109*** | 0.112*** | 0.117*** | 0.116*** | 0.118*** | 0.108*** | 0.106*** | 0.103*** | 0.101*** | 0.101*** | 0.093*** |
| | (0.004) | (0.004) | (0.004) | (0.004) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) | (0.005) |
| Other controls | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Year FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| County FE | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 100,938 | 98,997 | 97,075 | 95,156 | 93,443 | 91,772 | 90,203 | 88,688 | 87,185 | 85,533 | 83,818 |
| R2 | 0.348 | 0.369 | 0.386 | 0.403 | 0.414 | 0.418 | 0.421 | 0.428 | 0.430 | 0.439 | 0.448 |

## Table C.2: Baseline Local Effects - Event Study Specification

This table presents results from estimating a difference-in-differences regression specification to examine the impact of hurricanes on county-level patent output. This specification is an alternative to the local projection estimation shown in equation 1. The dependent variable is $\log\left(\frac{\overline{NPatents}_{c,t+h}}{\text{NPatents}_{c,t-1}}\right)$, where the numerator is the average number of patents from 5 to 7 and 5 to 9 years, respectively, post hurricane hit. The main independent variable on interest is $Hit_{c,t}$. The data span from 1851 to 2023. Standard errors are clustered by year and county and shown in parentheses. The specifications include county and time fixed effects. The significance of each coefficient estimate is indicated by * for $p < 0.10$, ** for $p < 0.05$, and *** for $p < 0.01$.

| Landfall radius | 50 | | 100 | | 200 | |
|---|---|---|---|---|---|---|
| Post-hit window length | 3 | 5 | 3 | 5 | 3 | 5 |
| $Hit_{c,t}$ | -0.084** | -0.086** | -0.069** | -0.067** | -0.056** | -0.061** |
| | (0.034) | (0.035) | (0.029) | (0.029) | (0.023) | (0.025) |
| $logNPatents_{c,t-1}$ | -0.735*** | -0.752*** | -0.739*** | -0.756*** | -0.740*** | -0.747*** |
| | (0.010) | (0.010) | (0.011) | (0.011) | (0.010) | (0.011) |
| $logNPatents_{c,t-2}$ | 0.182*** | 0.180*** | 0.183*** | 0.175*** | 0.175*** | 0.169*** |
| | (0.009) | (0.009) | (0.009) | (0.009) | (0.010) | (0.009) |
| $logNPatents_{c,t-3}$ | 0.123*** | 0.128*** | 0.134*** | 0.132*** | 0.140*** | 0.137*** |
| | (0.011) | (0.009) | (0.011) | (0.009) | (0.011) | (0.010) |
| $logNPatents_{c,t-4}$ | 0.132*** | 0.125*** | 0.115*** | 0.119*** | 0.133*** | 0.132*** |
| | (0.011) | (0.010) | (0.011) | (0.011) | (0.011) | (0.011) |
| $logNPatents_{c,t-5}$ | 0.100*** | 0.103*** | 0.097*** | 0.101*** | 0.097*** | 0.098*** |
| | (0.009) | (0.009) | (0.009) | (0.009) | (0.010) | (0.009) |
| Year FE | Yes | Yes | Yes | Yes | Yes | Yes |
| County FE | Yes | Yes | Yes | Yes | Yes | Yes |
| Observations | 17,287 | 16,911 | 18,297 | 17,999 | 17,816 | 17,574 |
| R2 | 0.523 | 0.568 | 0.526 | 0.570 | 0.531 | 0.570 |

## Table C.3: List of Deadliest Hurricanes

This table presents the list of deadliest Atlantic and Gulf Coast hurricanes included in our sample based on data from NOAA. The first column indicates the main states of impact and/or, from 1954, the hurricane names.

| Hurricane | Year | Fatalities |
|---|---|---|
| LA (Last Island) | 1856 | 400 |
| LA | 1860 | 47 |
| TX | 1875 | 176 |
| New England | 1878 | 27 |
| NC, VA | 1879 | 46 |
| GA/SC | 1881 | 700 |
| NC | 1883 | 53 |
| TX (Indianola) | 1886 | 150 |
| Texas | 1886 | 27 |
| Mid-Atlantic | 1889 | 40 |
| LA (Cheniere Caminanda) | 1893 | 1100-1400 |
| SC/GA (Sea Islands) | 1893 | 1000-2000 |
| SC, FL | 1893 | 28 |
| FL, GA, SC | 1896 | 130 |
| GA, SC, NC | 1898 | 179 |
| NC, SC | 1899 | 50 |
| TX (Galveston) | 1900 | 8000 |
| SE FL | 1906 | 164 |
| MS/AL/Pensacola | 1906 | 134 |
| LA (Grand Isle) | 1909 | 350 |
| TX (Velasco) | 1909 | 41 |
| SW FL | 1910 | 30 |
| LA (New Orleans) | 1915 | 275 |
| TX (Galveston) | 1915 | 275 |
| SW LA/Upper TX | 1918 | 34 |
| FL (Keys)/S TX | 1919 | 287 |
| FL (Miami)/MS/AL/Pensacola | 1926 | 372 |
| LA | 1926 | 25 |
| FL (SE/Lake Okeechobee) | 1928 | 2500 |
| TX (Freeport) | 1932 | 40 |
| S TX | 1933 | 40 |
| FL (Keys) | 1935 | 408 |
| New England | 1938 | 256 |
| GA/SC/NC | 1940 | 50 |
| Northeast U.S. | 1944 | 64 |
| SE FL/SE LA/MS | 1947 | 51 |
| Hazel (SC/NC) | 1954 | 95 |
| Carol (NE U.S.) | 1954 | 60 |
| Diane (NE U.S.) | 1955 | 184 |
| Connie (NC) | 1955 | 25 |
| Audrey (SW LA/N TX) | 1957 | 416 |
| Donna (FL/Eastern U.S.) | 1960 | 50 |
| Carla (N & Central TX) | 1961 | 46 |
| Hilda (LA) | 1964 | 38 |
| Betsy (SE FL/SE LA) | 1965 | 75 |
| Camille (MS/SE LA/VA) | 1969 | 256 |
| Agnes (FL/NE U.S.) | 1972 | 122 |
| Andrew (S FL, LA) | 1992 | 26 |
| Alberto (NW FL, GA, AL) | 1994 | 30 |



*Continued on the next page.*

**Table C.3: List of Deadliest Hurricanes (continued)**

| Hurricane | Year | Fatalities |
|---|---|---|
| Fran (NC) | 1996 | 26 |
| Floyd (Mid Atlantic & NE U.S.) | 1999 | 56 |
| Allison (SE TX) | 2001 | 41 |
| Ivan (NW FL, AL) | 2004 | 25 |
| Katrina (SE LA/MS) | 2005 | 1200 |
| Irene | 2011 | 48 |
| Sandy | 2012 | 160 |
| Matthew | 2016 | 47 |
| Harvey | 2017 | 106 |
| Irma | 2017 | 96 |
| Florence | 2018 | 54 |
| Michael | 2018 | 59 |
| Laura | 2020 | 41 |
| Ida | 2021 | 92 |
| Ian | 2022 | 156 |

## Table C.4: Variable Definitions

This table presents definitions of the main variables. The first column gives the variable name. The second column includes a short description.

| Variable Name | Description | Source |
|---|---|---|
| $CountyHitOther_{c,t}$ | This variable measures for firms located in county $c$ the share of their patents over five years before hurricane landfall from counties in the hurricane landfall region and but not from county $c$. | NOAA, USTPO, Census |
| $Hit_{c,t}$ | This is an indicator variable that takes a vlue of one if a county $c$ is in the landfall region of a hurricane in year $t$ and zero otherwise. | NOAA, Census |
| $HitOther_{i,c,t}$ | This variable measures the share of firm $i$'s patents over $h$ years before hurricane landfall that are from counties in the landfall region but not from county $c$. | NOAA, USPTO, Census |
| $\text{NPatents}_{c,t}$ | This variable counts the number of patents issued in year $t$ for which all inventors reside in county $c$. | USPTO, Census |
| $\overline{\text{NPatents}}_{i,c,t:t+h-1}$ | This variable is the average annual number of patents issued for firm $i$ in county $c$ over $h$ years starting in year $t$. | USPTO, Census |
| $SharePatentsHit_{i,t}$ | This variable measures the share of firm $i$'s patents over $h$ years before hurricane landfall from counties in the hurricane landfall region. | NOAA, USPTO, Census |