Central Bank Information or Neo-Fisher Effect?

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Abstract

The central bank information (CBI) effect and the neo-Fisher effect produce similar outcomes: under both, a monetary tightening triggers an increase in inflation and an expansion in real activity. Separate estimates of these effects run the risk of confounding one with the other. To disentangle these two channels, we introduce into a new-Keynesian model a permanent monetary shock that generates neo-Fisher effects and an aggregate demand shock to which the central bank responds that creates CBI effects. We estimate the model on U.S. data, including information on monetary shocks obtained from high frequency identification and on long-term yields. We find that the neo-Fisher shock is an important driver of inflation, while the CBI shock explains a significant fraction of movements in the nominal interest rate. The CBI shock explains little of inflation and output, but, through counterfactual exercises, we establish that this reflects the central bank's success in isolating the economy from aggregate demand disturbances. The contribution of the traditional monetary shock is found to be minor. These results are shown to hold under full and imperfect information.

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Keywords: Neo-Fisher Effect, Central Bank Information Effect, Inflation, Interest Rates, Aggregate Activity, High Frequency Monetary Shocks

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1 Introduction

Central bank information (CBI) and neo-Fisher effects can occur simultaneously and can generate similar outcomes. Both can give rise to short-run increases in inflation and aggregate activity in response to a surprise increase in the policy interest rate. A natural question is which of these two mechanisms plays a larger role in explaining why sometimes interest rates, prices, and quantities all move in the same direction after a monetary disturbance. In this paper, we provide an answer to this question from the perspective of a dynamic general equilibrium model estimated using postwar U.S. data, including information on monetary shocks obtained from high frequency identification and on long-term interest rates.

We construct a model with nominal and real rigidities driven by permanent and transitory monetary shocks, two preference shocks, and permanent and transitory productivity shocks. To create a central bank information channel, we assume an augmented Taylor rule whereby the central bank responds directly to one of the preference shocks in addition to the output gap and inflation. To create a neo-Fisher effect, we assume that the Taylor rule is also buffeted by permanent monetary shocks in addition to standard transitory monetary shocks.

We examine two polar information structures: in one, private agents observe all shocks individually, while in the other, agents observe only the sum of the two preference shocks and the stochastic component of the Taylor rule, but not the individual shocks that comprise them.

The main result of the paper is that both the neo-Fisher and central bank information effects are important drivers of macroeconomic indicators of interest. Permanent monetary shocks explain between 20 and 30 percent of the variance of inflation changes, and the preference shock to which the central bank responds in the CBI channel explains about 50 percent of movements in the interest rate. Output and inflation respond little to this preference shock, but this is a reflection of the success of the central bank at stabilizing the aggregate effects of this type of shock. When we counterfactually shut down the central bank's response to this shock by setting its coefficient in the Taylor rule equal to zero, we

find that it explains about more than half of movements in inflation and about 20 percent of movements in real activity. The standard monetary shock plays a minor role in accounting for movements in interest rates, inflation, and output.

These results hold under full and imperfect information. The reason for why imperfect information does not seem to be empirically important is that the CBI and neo-Fisher shocks, although capable of causing similar effects in the short run, are dynamically quite distinct, so agents can rapidly extract signals to tell them apart. The results are also robust to including high frequency monetary shocks or long-term yields as observables in the estimation.

To the best of our knowledge, this paper is the first attempt to evaluate jointly the contributions of the neo-Fisher and the central bank information channels. It is related to two strands of the recent monetary literature, one dedicated to the neo-Fisher effect and the other to the central bank information effect. The formulation and estimation of the neo-Fisher effect adopted in this study follows Uribe (2022). The neo-Fisher effect has been estimated in the context of empirical models among others by Uribe (2017, 2022) and Valle e Azevedo, Ritto, and Teles (2022) using data from advanced countries, García-Cicco, Goldstein, and Sturzenegger (2024) using data from advanced and emerging economies, and Schmitt-Grohé and Uribe (2022) and Carvalho, Valle e Azevedo, and Pires Ribeiro (2024) in open-economy settings. Lukmanova and Rabitsch (2023) estimate the neo-Fisher effect in the context of an equilibrium model with transitory but persistent monetary shocks and imperfect information. These studies find evidence of a significant neo-Fisher effect, that is, they find that a permanent or highly persistent innovation in the interest rate raises inflation and aggregate activity in the short run. Theoretical formulations of the neo-Fisher effect can be found in Schmitt-Grohé and Uribe (2010, 2012), Cochrane (2016), Williamson (2016), and Garín, Lester, and Sims (2018).

There is a large literature studying the central bank information effect. Romer and Romer (2000), Barakchian and Crowe (2013), and Campbell, Fisher, Justiniano, and Melosi (2017) provide early evidence of a central bank information advantage. Estimates of the central bank

information shock include Hansen and McMahon (2016) using language analysis of central bank announcements, Kerssenfischer (2022), Cieslak and Schrimpf (2019), and Jarociński and Karadi (2020) using high frequency movements in stock prices, Miranda-Agrippino (2016) and Campbell, Fisher, Justiniano, and Melosi (2017) using forecast differentials between the private sector and the Federal Reserve, Acosta (2023) using textual analysis of newspaper articles around FOMC announcements, and García-Schmidt (2024) using data from forecasts of inflation and output growth from emerging countries. The macroeconomic effects of CBI shocks have been studied in Melosi (2017), Nakamura and Steinsson (2018), and Jarociński and Karadi (2020). Finally, there are studies that have raised questions about the neo-Fisher and CBI channels. See García-Schmidt and Woodford (2019) and Bouakez and Kano (2024) for papers questioning the neo-Fisher effect and Faust, Swanson, and Wright (2004) and Bauer and Swanson (2023) for papers challenging the CBI channel.

The present study connects the two branches of the monetary economics literature just described by assessing jointly the contributions of shocks that generate neo-Fisher and CBI effects. This is important because ignoring the simultaneous presence of neo-Fisher and CBI effects can lead to overestimating the influence of each channel individually.

The remainder of the paper is organized as follows. Section 2 presents the proposed model. Section 3 characterizes the signal extraction problem under imperfect information. Section 4 discusses the econometric estimation. Section 5 presents the results. Finally, section 6 presents concluding remarks.

2 The Model

The model economy features sticky prices, habit formation, permanent and transitory monetary shocks, permanent and transitory productivity shocks, and preference shocks.

2.1 Households

The economy is populated by households with preferences defined over streams of consumption and labor effort and exhibiting external habit formation. The household's lifetime utility function is

$$E_0 \sum_{t=0}^{\infty} \beta^t e^{\xi_t} \left\{ \frac{\left[(C_t - \delta \widetilde{C}_{t-1})(1 - e^{\theta} h_t)^{\chi} \right]^{1-\sigma} - 1}{1 - \sigma} \right\}, \tag{1}$$

where C_t denotes consumption, $\widetilde{C}_t$ denotes the cross-sectional average of consumption, h_t denotes hours worked, ξ_t is a preference shifter, and $\beta, \delta \in (0,1)$ and $\sigma, \chi > 0$ and θ are parameters.

To model a central bank information channel, we assume that the preference shifter ξ_t has two components,

$$\xi_t = \xi_t^h + \xi_t^c, \tag{2}$$

which are both exogenous stochastic processes. The central bank is assumed to observe ξ_t^c and to respond to it in setting the nominal interest rate. Households and firms are assumed to observe the preference shifter ξ_t, but not its individual components ξ_t^h and ξ_t^c.

Households are subject to the budget constraint

$$P_t C_t + \frac{B_t}{1 + I_t} + T_t = B_{t-1} + W_t h_t + \Phi_t, \tag{3}$$

where P_t denotes the price level, T_t denotes nominal lump-sum taxes, W_t denotes the nominal wage rate, and Φ_t denotes nominal profits received from firms. The variable B_t denotes the units of a one-period nominal discount bond purchased in period t that pays the interest rate I_t.

The consumption good C_t is assumed to be a composite of a continuum of varieties C_{it}

indexed by $i \in [0,1]$ with aggregation technology

$$C_t = \left[\int_0^1 C_{it}^{1-1/\eta} di \right]^{\frac{1}{1-1/\eta}}, \tag{4}$$

where the parameter $\eta > 0$ denotes the elasticity of substitution across varieties.

Households choose processes $\{C_t, h_t, B_t\}_{t=0}^{\infty}$ to maximize the utility function (1) subject to the budget constraint (3) and to some borrowing limit that prevents them from engaging in Ponzi schemes. Letting $\beta^t \Lambda_t / P_t$ denote the Lagrange multiplier associated with the budget constraint (3), the first-order conditions of the household's optimization problem are

$$e^{\xi_t} (C_t - \delta \widetilde{C}_{t-1})^{-\sigma} (1 - e^{\theta} h_t)^{\chi(1-\sigma)} = \Lambda_t, \tag{5}$$

$$\frac{\chi e^{\theta} (C_t - \delta \widetilde{C}_{t-1})}{1 - e^{\theta} h_t} = \frac{W_t}{P_t}, \tag{6}$$

and

$$\Lambda_t = \beta (1 + I_t) E_t \left[\frac{\Lambda_{t+1}}{1 + \Pi_{t+1}} \right], \tag{7}$$

where $\Pi_t \equiv P_t / P_{t-1} - 1$ denotes the inflation rate in period t.

Given its desired level of consumption, C_t, the household chooses the consumption of varieties C_{it} to minimize total expenditure, $\int_0^1 P_{it} C_{it} di$, subject to the aggregation technology (4), where P_{it} denotes the nominal price of variety i. This problem delivers the following demand for individual varieties:

$$C_{it} = C_t \left(\frac{P_{it}}{P_t} \right)^{-\eta}, \tag{8}$$

where the price level P_t is given by

$$P_t \equiv \left[\int_0^1 P_{it}^{1-\eta} di \right]^{\frac{1}{1-\eta}} \tag{9}$$

and represents the minimum cost of one unit of the composite consumption good, C_t.

2.2 Firms

The firm producing variety i operates in a monopolistically competitive market and faces quadratic price adjustment costs à la Rotemberg (1982). The production technology uses labor and is buffeted by transitory and permanent productivity shocks. Specifically, output of variety i is given by

$$Y_{it} = e^{z_t} \Omega_t h_{it}^{\alpha}, \tag{10}$$

where Y_{it} denotes output of variety i in period t, h_{it} denotes labor input used in the production of variety i, z_t is a stationary productivity shock, and Ω_t is a nonstationary productivity shock. The growth rate of Ω_t is assumed to be a stationary random variable with mean g,

$$g_t = \ln\left(\frac{\Omega_t}{\Omega_{t-1}}\right) - g.$$

The expected present discounted value of real profits of the firm producing variety i expressed in units of the final good is given by

$$E_0 \sum_{t=0}^{\infty} q_t \left[\frac{P_{it}}{P_t} A_{it} - \frac{W_t}{P_t} h_{it} - \frac{\phi}{2} \Omega_t \left(\frac{P_{it}/P_{it-1}}{1 + \widetilde{\Pi}_t} - 1 \right)^2 \right], \tag{11}$$

where A_{it} denotes the total demand for good i, $\phi > 0$ is a parameter governing the degree of price rigidity, and $q_t \equiv \beta^t \Lambda_t / \Lambda_0$ denotes a pricing kernel reflecting the assumption that households are the owners of the firms. Price adjustment costs are defined in units of the final good and are scaled by the trend component of productivity Ω_t to keep nominal rigidity from vanishing along the balanced growth path. The total demand for good i, A_{it}, includes the demand stemming from households and the demand stemming from firms to cover their

price adjustment costs. Specifically, the demand for good i is given by

$$A_{it} = A_t \left(\frac{P_{it}}{P_t} \right)^{-\eta} \tag{12}$$

with

$$A_t = C_t + \frac{\phi}{2} \Omega_t \int_0^1 \left(\frac{P_{jt}/P_{jt-1}}{1 + \widetilde{\Pi}_t} - 1 \right)^2 dj. \tag{13}$$

The second term of this expression is nil up to first order. The variable $\widetilde{\Pi}_t$ denotes the average level of inflation around which price-adjustment costs are defined. It is predetermined in period t and is assumed to evolve over time according to

$$1 + \widetilde{\Pi}_{t+1} = (1 + \widetilde{\Pi}_t)^{\gamma_\pi} (1 + \Pi_t)^{1-\gamma_\pi}, \tag{14}$$

where the parameter $\gamma_\pi \in [0, 1]$ governs the backward-lookingness of price indexation.

The problem of the firm producing variety i is to choose processes $\{P_{it}, A_{it}, Y_{it}, h_{it}\}_{t=0}^{\infty}$ to maximize (11) subject to the demand equation (12), the production technology (10), and the requirement that demand be satisfied at the price set by the firm,

$$Y_{it} \geq A_{it}, \tag{15}$$

taking as given the processes $\widetilde{\Pi}_t$, z_t, Ω_t, A_t, W_t, P_t, and q_t.

Letting $q_t mc_{it}$ be the Lagrange multiplier associated with the demand constraint (15), the first-order conditions associated with the firm's profit maximization problem are

$$mc_{it} = \frac{W_t/P_t}{\alpha e^{z_t} \Omega_t h_{it}^{\alpha-1}} \tag{16}$$

and

$$\eta A_{it}\left(\frac{\eta-1}{\eta}\frac{P_{it}}{P_t}-mc_{it}\right)=-\phi\Omega_t\frac{P_{it}/P_{it-1}}{1+\widetilde{\Pi}_t}\left(\frac{P_{it}/P_{it-1}}{1+\widetilde{\Pi}_t}-1\right)+\phi E_t\frac{q_{t+1}}{q_t}\Omega_{t+1}\frac{P_{it+1}/P_{it}}{1+\widetilde{\Pi}_{t+1}}\left(\frac{P_{it+1}/P_{it}}{1+\widetilde{\Pi}_{t+1}}-1\right). \tag{17}$$

The first of these optimality conditions says that the multiplier mc_{it} represents marginal cost. The second optimality condition says that, all other things equal, if marginal revenue, $\frac{\eta-1}{\eta}\frac{P_{it}}{P_t}$, exceeds marginal cost, mc_{it}, the firm will increase prices at a rate below normal, $P_{it}/P_{it-1}<1+\widetilde{\Pi}_t$.

2.3 Monetary and Fiscal Policy

The monetary authority follows a Taylor-type interest-rate feedback rule buffeted by a stationary monetary shock denoted z_t^m and a nonstationary monetary shock denoted X_t^m. The nonstationary monetary shock X_t^m is meant to capture the neo-Fisher effect (Uribe, 2022). The monetary rule also includes a central bank information channel. Specifically, the monetary authority is assumed to respond to the exogenous preference shifter ξ_t^c. Formally,

$$\frac{1+I_t}{X_t^m}=\left[\Gamma\left(\frac{1+\Pi_t}{X_t^m}\right)^{\alpha_\pi}\left(\frac{Y_t}{Y_t^n}\right)^{\alpha_y}\right]^{1-\gamma_I}\left(\frac{1+I_{t-1}}{X_{t-1}^m}\right)^{\gamma_I}e^{z_t^m+\alpha_\xi\xi_t^c}, \tag{18}$$

where Y_t denotes aggregate output, Y_t^n denotes the flexible-price or natural level of output, and Γ, α_π, α_y, α_ξ, and $\gamma_I\in[0,1)$ are parameters. The parameter α_ξ defines the central bank information channel. The growth rate of the nonstationary monetary shock,

$$g_t^m\equiv\ln\left(\frac{X_t^m}{X_{t-1}^m}\right),$$

is assumed to be stationary.

Households and firms are assumed to observe the nominal interest rate, I_t, inflation, Π_t, the output gap, Y_t/Y_t^n, and the exogenous stochastic component of the Taylor rule, which

we denote $\widetilde{\Omega}_t^m$, and is given by

$$\widetilde{\Omega}_t^m \equiv X_t^{m1-\alpha_\pi(1-\gamma_I)} {X_{t-1}^m}^{-\gamma_I} e^{z_t^m + \alpha_\xi \xi_t^c}. \tag{19}$$

However, households and firms are assumed to be unable to observe the individual components of $\widetilde{\Omega}_t^m$, namely, X_t^m, z_t^m, and ξ_t^c.

Government consumption is assumed to be nil at all times. The government's budget constraint is then given by

$$T_t + \frac{B_t}{1 + I_t} = B_{t-1}.$$

Fiscal policy is assumed to be Ricardian, that is, it guarantees the intertemporal solvency of the government independently of the path of the price level.

2.4 Driving Forces

The structural shocks driving the economy, ξ_t^h, ξ_t^c, z_t^m, g_t^m, g_t, and z_t are assumed to follow AR(1) processes of the form

$$x_t = \rho_x x_{t-1} + \sigma_x \epsilon_t^x, \tag{20}$$

where ϵ_t^x is an i.i.d. random disturbance distributed $N(0,1)$, $\rho_x \in [0,1)$ is the serial correlation of x_t, and $\sigma_x > 0$ is the standard deviation of the innovation of the process, for $x = \xi^h$, ξ^c, z^m, g^m, g, and z.

2.5 Market Clearing

Clearing of the labor market requires that the demand for labor by firms equal the household's supply of labor, that is,

$$\int_0^1 h_{it} di = h_t. \tag{21}$$

Because all households are identical, so are individual and aggregate consumption per

capita,

$$C_t = \widetilde{C}_t.$$

We focus attention on a symmetric equilibrium in which all firms charge the same nominal price and employ the same amount of labor, that is, an equilibrium in which h_{it} and P_{it} are the same for all $i \in [0,1]$. We then have from equations (8), (9), (10), (16), and (21) that $P_{it} = P_t$, $C_{it} = C_t$, $h_{it} = h_t$, $mc_{it} = mc_t$, and $Y_{it} = e^{z_t}\Omega_t h_t^{\alpha}$, for all i. Output measured in units of the final good is then given by $Y_t \equiv \left(\int_0^1 P_{it}Y_{it}di\right)/P_t = e^{z_t}\Omega_t h_t^{\alpha}$. As long as the nominal wage is positive, the firm will choose to satisfy the demand constraint (15) with equality, so that in equilibrium

$$Y_t = C_t + \frac{\phi}{2}\Omega_t \left(\frac{1+\Pi_t}{1+\widetilde{\Pi}_t} - 1\right)^2.$$

2.6 Equilibrium Conditions in Stationary Form

The model is driven, among other sources of variation, by two nonstationary shocks, the nonstationary productivity shock, Ω_t, and the nonstationary monetary shock, X_t^m. We express the model in terms of stationary variables by scaling all variables with stochastic trends by their respective permanent components. Under imperfect information, however, the choice of these trend components is nontrivial. The reason is that all conditional expectations in the model are taken given the private agents's information sets, which in period t include Ω_t, but not X_t^m. Thus, Ω_t can be used as a scaler to transform nonstationary real variables into stationary real variables, but X_t^m cannot be used in a similar way to transform nonstationary nominal variables into stationary ones.

To illustrate the problem with using X_t^m to transform nonstationary nominal variables into stationary ones, consider a generic expression of the form $N_t^1 = E_t N_{t+1}^2$, where N_t^1 and N_t^2 are two nominal variables that are cointegrated with X_t^m. Under full information we could use X_t^m as a scaler and define $n_t^1 = N_t^1/X_t^m$ and $n_t^2 = N_t^2/X_t^m$. Because under

full information $1/X_t^m E_t N_{t+1}^2 = E_t N_{t+1}^2/X_t^m$, we can write $n_t^1 = E_t\left[n_{t+1}^2 e^{g_{t+1}^m}\right]$. However, if X_t^m is not in private agents' period-t information set, then $1/X_t^m E_t N_{t+1}^2 \neq E_t N_{t+1}^2/X_t^m$, rendering this approach to induce stationarity infeasible.

Therefore, in order to transform nonstationary nominal variables into stationary ones, we must use an observable object that is cointegrated with X_t^m. To this end, note that the exponents of X_t^m and X_{t-1}^m in the definition of the observable object $\widetilde{\Omega}_t^m$ given in equation (19) add up to $(1-\alpha_\pi)(1-\gamma_I)$. We therefore define the observable object Ω_t^m as

$$\Omega_t^m \equiv \left(\widetilde{\Omega}_t^m\right)^{\frac{1}{(1-\alpha_\pi)(1-\gamma_I)}},$$

which is cointegrated with X_t^m.

Accordingly, we use the observable exogenous variables Ω_t and Ω_t^m as the scalers to convert real and nominal nonstationary variables into stationary variables. Specifically, we create the stationary variables $c_t \equiv C_t/\Omega_t$, $y_t \equiv Y_t/\Omega_t$, $y_t^n \equiv Y_t^n/\Omega_t$, $a_t \equiv A_t/\Omega_t$, $w_t \equiv W_t/(P_t\Omega_t)$, $\lambda_t \equiv \Lambda_t\Omega_t^\sigma$, $1+\pi_t \equiv (1+\Pi_t)/\Omega_t^m$, $1+i_t \equiv (1+I_t)/\Omega_t^m$, and $1+\widetilde{\pi}_t \equiv (1+\widetilde{\Pi}_t)/\Omega_{t-1}^m$. Let $g_t^{\omega m} \equiv \ln\left(\frac{\Omega_t^m}{\Omega_{t-1}^m}\right)$ be the growth rate of Ω_t^m, so that

$$g_t^{\omega m} = \frac{g_t^m[1-\alpha_\pi(1-\gamma_I)] - g_{t-1}^m\gamma_I + z_t^m - z_{t-1}^m + \alpha_\xi(\xi_t^c - \xi_{t-1}^c)}{(1-\alpha_\pi)(1-\gamma_I)}. \tag{22}$$

We can then write equilibrium conditions (5)–(7), (10), and (13)–(18), respectively, as

$$\lambda_t = e^{\xi_t}\left(c_t - \delta\frac{c_{t-1}}{e^{g_t+g}}\right)^{-\sigma}\left(1-e^\theta h_t\right)^{\chi(1-\sigma)}, \tag{23}$$

$$\frac{\chi e^\theta\left(c_t - \delta\frac{c_{t-1}}{e^{g_t+g}}\right)}{1-e^\theta h_t} = w_t, \tag{24}$$

$$\lambda_t = \beta(1+i_t)E_t\left[\frac{\lambda_{t+1}}{1+\pi_{t+1}}e^{-g_{t+1}^{\omega m}-\sigma(g_{t+1}+g)}\right], \tag{25}$$

$$y_t = e^{z_t} h_t^{\alpha}, \tag{26}$$

$$a_t = c_t + \frac{\phi}{2}\left(\frac{1+\pi_t}{1+\widetilde{\pi}_t} e^{g_t^{\omega m}} - 1\right)^2, \tag{27}$$

$$1 + \widetilde{\pi}_{t+1} = [(1+\widetilde{\pi}_t) e^{-g_t^{\omega m}}]^{\gamma_m} (1+\pi_t)^{1-\gamma_m}, \tag{28}$$

$$y_t = a_t, \tag{29}$$

$$\mathrm{mc}_t = \frac{w_t}{\alpha e^{z_t} h_t^{\alpha-1}}, \tag{30}$$

$$\begin{aligned}\left(\frac{\eta-1}{\eta} - mc_t\right) a_t \;=\; & \frac{\phi}{\eta}\beta E_t e^{(1-\sigma)(g_{t+1}+g)} \frac{\lambda_{t+1}}{\lambda_t} \frac{1+\pi_{t+1}}{1+\widetilde{\pi}_{t+1}} e^{g_{t+1}^{\omega m}} \left(\frac{1+\pi_{t+1}}{1+\widetilde{\pi}_{t+1}} e^{g_{t+1}^{\omega m}} - 1\right) \\ & - \frac{\phi}{\eta} \frac{1+\pi_t}{1+\widetilde{\pi}_t} e^{g_t^{\omega m}} \left(\frac{1+\pi_t}{1+\widetilde{\pi}_t} e^{g_t^{\omega m}} - 1\right),\end{aligned} \tag{31}$$

and

$$1 + i_t = \left[\Gamma (1+\pi_t)^{\alpha_\pi} \left(\frac{y_t}{y_t^n}\right)^{\alpha_y}\right]^{1-\gamma_I} \left(\frac{1+i_{t-1}}{e^{g_t^{\omega m}}}\right)^{\gamma_I}. \tag{32}$$

The detrended flexible-price level of output, y_t^n, is given by the solution of equations (24), (26), and (30) evaluated at $c_t = y_t^n$ and $mc_t = (\eta-1)/\eta$, for all t. This yields y_t^n implicitly as

$$\frac{\chi e^{\theta}\left(y_t^n - \delta \frac{y_{t-1}^n}{e^{g_t+g}}\right)}{1 - e^{\theta}(y_t^n e^{-z_t})^{1/\alpha}} = \frac{\eta-1}{\eta} \alpha e^{z_t/\alpha} (y_t^n)^{(\alpha-1)/\alpha}. \tag{33}$$

Note that y_t^n depends on the productivity shocks z_t and g_t, but not on preference or monetary shocks, so it is observable even under imperfect information.

We are ready to define a competitive equilibrium under full information:

Definition 1 (Equilibrium Under Full Information) *An equilibrium under full information is a set of processes* c_t, h_t, λ_t, w_t, i_t, π_t, $\tilde{\pi}_t$, y_t, y_t^n, a_t, mc_t, $g_t^{\omega m}$, *and* ξ_t *satisfying* (2) *and* (22)–(33)*, given stochastic processes* ξ_t^h, ξ_t^c, z_t, g_t, z_t^m, *and* g_t^m*, and the initial conditions* c_{-1}, y_{-1}^n, $\tilde{\pi}_0$, i_{-1}, g_{-1}^m, z_{-1}^m, *and* ξ_{-1}^c.

We now turn to the characterization of the equilibrium under imperfect information.

3 Signal Extraction

Under imperfect information, agents must form conditional expectations about future values of endogenous and exogenous variables, based on the current position of the observable states of the economy and knowledge of the law of motion of the unobserved exogenous states ξ_t^h, ξ_t^c, z_t^m, and g_t^m. To solve this signal extraction problem, we proceed as follows. Let $\boldsymbol{y}_t = [h_t \;\; y_t \;\; y_t^n \;\; w_t \;\; mc_t \;\; \lambda_t \;\; \pi_t \;\; a_t \;\; i_t]'$ be the vector of controls, $\boldsymbol{v}_t = [c_{t-1} \;\; y_{t-1}^n \;\; \tilde{\pi}_t \;\; i_{t-1} \;\; z_t \;\; g_t\,]'$, the vector of endogenous states and observable exogenous states, and

$$\boldsymbol{o}_t = \begin{bmatrix} \xi_t \\ g_t^{\omega m} \end{bmatrix} \tag{34}$$

the vector of observable exogenous states used to extract information about the current and future expected positions of the unobservable exogenous states ξ_t^h, ξ_t^c, z_t^m, and g_t^m.

Let the lengths of $\boldsymbol{y}_t$, $\boldsymbol{v}_t$, and $\boldsymbol{o}_t$ be n_y, n_v and n_o. Linearize the equilibrium conditions (23)-(33). The resulting linearized expressions together with the AR(1) processes for the observed exogenous states included in $\boldsymbol{v}_t$ (i.e., z_t and g_t) can be written as

$$\mathcal{A}\, E_t \begin{bmatrix} \hat{\boldsymbol{y}}_{t+1} \\ \hat{\boldsymbol{v}}_{t+1} \\ \boldsymbol{o}_{t+1} \end{bmatrix} = \mathcal{B} \begin{bmatrix} \hat{\boldsymbol{y}}_t \\ \hat{\boldsymbol{v}}_t \\ \boldsymbol{o}_t \end{bmatrix}, \tag{35}$$

where hatted variables refer to deviations from steady-state values. (The vector $\boldsymbol{o}_t$ is not

hatted, as its steady-state value is zero, so it is already expressed in deviations from steady state.)

The system (35) has $n_y + n_v$ equations and $n_y + n_v + n_o$ variables. Let $\boldsymbol{u}_t \equiv [\xi_t^h \;\; \xi_t^c \;\; \xi_{t-1}^c \;\; z_t^m \;\; z_{t-1}^m \;\; g_t^m \; g_{t-1}^m]'$ be the vector of unobserved exogenous states with length n_u. Then, from the law of motion of the exogenous driving forces given in (20) we have that

$$\boldsymbol{u}_{t+1} = F\boldsymbol{u}_t + B\boldsymbol{\epsilon}_{t+1}^u \tag{36}$$

and from equations (2), (22), and (34) we have that

$$\boldsymbol{o}_t = H'\boldsymbol{u}_t \tag{37}$$

with

$$F = \begin{bmatrix} \rho_{\xi h} & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & \rho_{\xi c} & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & \rho_{zm} & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & \rho_{gm} & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 \end{bmatrix}, \quad B = \begin{bmatrix} \sigma_{\xi h} & 0 & 0 & 0 \\ 0 & \sigma_{\xi c} & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & \sigma_{zm} & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & \sigma_{gm} \\ 0 & 0 & 0 & 0 \end{bmatrix},$$

$$H' = \begin{bmatrix} 1 & 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & \frac{\alpha_\xi}{a} & -\frac{\alpha_\xi}{a} & \frac{1}{a} & -\frac{1}{a} & \frac{a+\gamma_I}{a} & \frac{-\gamma_I}{a} \end{bmatrix}, \quad \text{and} \quad \boldsymbol{\epsilon}_t^u = \begin{bmatrix} \epsilon_t^{\xi h} \\ \epsilon_t^{\xi c} \\ \epsilon_t^{zm} \\ \epsilon_t^{gm} \end{bmatrix},$$

where $a \equiv (1-\alpha_\pi)(1-\gamma_I)$. Let $\boldsymbol{\eta}_t \equiv E_{t-1}\boldsymbol{u}_t$ be the expected value of $\boldsymbol{u}_t$ conditional on $\boldsymbol{o}_{t-1}$. Then, applying the Kalman filter yields

$$\boldsymbol{\eta}_{t+1} = (F - KH')\boldsymbol{\eta}_t + K\boldsymbol{o}_t \tag{38}$$

and

$$E_t \boldsymbol{o}_{t+1} = H' \boldsymbol{\eta}_{t+1}, \tag{39}$$

where $K = FPH(H'PH)^{-1}$ is the Kalman gain, $P = F[P - PH(H'PH)^{-1}H'P]F' + Q$ is the steady-state mean square error of $\boldsymbol{\eta}_{t+1}$ expressed implicitly as a Ricatti equation, and $Q \equiv BB'$ is the variance-covariance matrix of the innovation $B\boldsymbol{\epsilon}_t^u$. Equation (38) gives the law of motion of agents' expectation of the position of the latent exogenous state vector $\boldsymbol{u}_t$. This expectation changes over time as information about the observable state $\boldsymbol{o}_t$ arrives. Equation (39) expresses the expected value of the observable state in period $t+1$, $E_t \boldsymbol{o}_{t+1}$, as a function of the expected position of the vector of unobservable states in $t+1$, $\boldsymbol{\eta}_{t+1}$.

The system consisting of (35), (38), and (39) has $n_y + n_v + n_o + n_u$ equations and also $n_y + n_v + n_o + n_u$ unknowns. Solving this system taking $\hat{\boldsymbol{v}}_t$, $\boldsymbol{\eta}_t$, and $\boldsymbol{o}_t$ as states and $\hat{\boldsymbol{y}}_t$ as controls yields

$$\boldsymbol{x}_{t+1} = h_x \boldsymbol{x}_t + \Sigma \boldsymbol{\epsilon}_{t+1} \tag{40}$$

and

$$\hat{\boldsymbol{y}}_t = g_x \boldsymbol{x}_t, \tag{41}$$

where $\boldsymbol{x}_t \equiv [\hat{\boldsymbol{v}}_t' \, \boldsymbol{\eta}_t' \, \boldsymbol{o}_t']'$, $\boldsymbol{\epsilon}_t = [\epsilon_t^z \;\; \epsilon_t^g]'$, and Σ contains the standard deviations of the elements of $\boldsymbol{\epsilon}_t$ in the corresponding positions. Let n_ϵ be the length of $\boldsymbol{\epsilon}_t$.

Finally, to express the evolution of this system in terms of its exogenous state $\boldsymbol{u}_t$, use equations (36) and (37) to eliminate $\boldsymbol{o}_t$ from (40) and (41). To this end, partition h_x, g_x, and Σ as

$$h_x = \begin{bmatrix} h_x^{11} & h_x^{12} \\ h_x^{21} & h_x^{22} \end{bmatrix}, \quad g_x = \begin{bmatrix} g_x^1 & g_x^2 \end{bmatrix}, \quad \text{and } \Sigma = \begin{bmatrix} \Sigma^1 \\ \Sigma^2 \end{bmatrix},$$

with h_x^{11} of size $n_v + n_u$ by $n_v + n_u$, g_x^1 of size n_y by $n_v + n_u$, and Σ^1 of size $n_v + n_u$ by n_ϵ. Then, the equilibrium evolution of all variables in the model can be written as

$$\boldsymbol{X}_{t+1} = H_x \boldsymbol{X}_t + C\boldsymbol{\mu}_{t+1}$$

Table 1: Calibrated Parameters

Parameter	Value	Description
g	0.0041	mean output growth rate
σ	2	inverse of intertemp. elast. subst.
β	0.9982	subjective discount factor
η	6	intratemporal elast. of subst.
α	0.75	labor semielast. of output
θ	0.4055	preference parameter

Note. The time unit is one quarter.

and

$$\hat{\boldsymbol{y}}_t = G_x \boldsymbol{X}_t,$$

where $\boldsymbol{X}_t \equiv [\hat{\boldsymbol{v}}_t' \, \boldsymbol{\eta}_t' \, \boldsymbol{u}_t']'$, $\boldsymbol{\mu}_t = [\boldsymbol{\epsilon}_t' \; \boldsymbol{\epsilon}_t^{w\prime}]'$, and

$$H_x = \begin{bmatrix} h_x^{11} & h_x^{12} H' \\ \emptyset & F \end{bmatrix}, \quad C = \begin{bmatrix} \Sigma^1 & \emptyset \\ \emptyset & B \end{bmatrix}, \quad \text{and} \quad G_x = \begin{bmatrix} g_x^1 & g_x^2 H' \end{bmatrix}.$$

We now proceed to estimate the model and characterize its equilibrium properties.

4 Estimation

We follow the standard practice of calibrating some parameters and estimating others. We estimate the parameters defining all stochastic processes, the parameters defining price stickiness and habit formation, and all policy parameters. The remaining parameters are calibrated.

Table 1 summarizes the calibration. The time unit is one quarter. The calibrated parameters take standard values in business-cycle analysis. We set the steady-state of the growth rate of the nonstationary productivity shock, g, to 0.0041, which is consistent with an average growth rate of output per capita of about 1.7 percent per year over the postwar period; the inverse of the intertemporal elasticity of consumption substitution, σ, to 2; the steady-state

real interest rate to 4 percent per year, which implies a growth-adjusted discount factor, $\beta e^{-\sigma g}$, of 0.99 per year and a value of β of 0.9982; the labor elasticity of output, α, to 0.75; the elasticity of substitution across good varieties, η, to 6; the steady-state share of time allocated to work, h, to 1/3; and the labor supply elasticity holding consumption constant, $(1-e^{\theta}h)/(e^{\theta}h)$, to 1. The last two restrictions imply that the preference parameter θ equals $-\ln(2h) = 0.4055$.

We estimate the remaining parameters under full and imperfect information using Bayesian techniques. The estimation uses quarterly U.S. data covering the period 1961:Q3 to 2019:Q4. We leave the Covid-19 years out of the sample as they were arguably driven by extraordinary shocks not included in the model. The data used in the estimation contains observations on the growth rate of real GDP per capita, $\Delta \ln Y_t$, the change in the nominal interest rate, proxied by the Federal Funds rate, ΔI_t, and the interest-rate inflation differential, $I_t - \Pi_t$, where Π_t is proxied by the growth rate of the implicit GDP deflator. In section 5.3, we expand the set of observables to include a monetary shock constructed from high frequency information. The results reported below are based on 1 million draws from the posterior distribution of the vector of estimated parameters using the Metropolis-Hastings sampler, after burning 50 million draws.

Table 2 provides a summary of the prior and posterior distributions of the estimated parameters. The serial correlations of all exogenous shocks are given beta prior distributions with mean 0.5 and standard deviation 0.2, except for those of the growth rates of the non-stationary shocks, g_t and g_t^m, which are given prior means of 0.3. The standard deviations of all shocks are given gamma distributions with mean and standard deviation 0.01 (or 1 percent), except for those of the monetary shocks z_t^m and g_t^m, which are given means and standard deviations that are 4 times smaller (0.0025), as they are frequency dependent.

Consider now the prior distributions of the parameters of the Taylor rule. The coefficients associated with inflation, the output-gap, and interest-rate smoothing are assigned standard values. Specifically, the inflation coefficient, α_π, has a gamma distribution with mean 1.5

and standard deviation 0.25. The output-gap coefficient, α_y, has a gamma distribution with mean 0.125 and standard deviation 0.1. And the interest-rate smoothing parameter, γ_I, has a uniform distribution with support [0 1]. The parameter governing the central bank information channel, α_ξ, has the same prior distribution as the output-gap coefficient, namely, a gamma distribution with mean 0.125 and standard deviation 0.1. This is motivated by the fact that, like α_y, the coefficient α_ξ represents the central bank's response to a measure of aggregate activity. Finally, the coefficient governing price stickiness, ϕ, has a gamma prior distribution with mean 50 and standard deviation 20, and the parameters governing habit formation, δ, and price indexation, γ_π, have uniform prior distributions with support [0 1].

Posterior moments are displayed in the last four columns of Table 2. As is common in estimated macroeconomic models, a number of parameters are imprecisely estimated. However, the estimation provides informative posteriors for the novel feature of the model, namely, the joint presence of a central bank information channel and a neo-Fisher effect. In particular, the estimation provides informative posteriors for the parameter governing the central bank's response to the unobserved component of the demand shock, α_ξ, the parameters defining the law of motion of this shock, ρ_{ξ^c} and σ_{ξ^c}, and the standard deviation of the permanent monetary shock, σ_{g^m}. The estimation also yields precise estimates for the parameters defining nominal and real frictions, particularly the price stickiness parameter, ϕ, and the habit formation parameter, δ, which govern the endogenous propagation of exogenous shocks.

Interestingly, it is evident from Table 2 that the posterior estimates are fairly similar whether the estimation is conducted under full or imperfect information. Thus, in the analysis that follows, differences in the model's predictions under these two information structures can largely be attributed to the impact of informational frictions on the behavior of private agents.

Table 2: Prior and Posterior Parameter Distributions

	Prior Distribution			Posterior Distribution			
				Full Information		Imperfect Information	
Parameter	Distribution	Mean	Std	Mean	Std	Mean	Std
ϕ	Gamma	50	20	96.1	25.9	96.7	26.1
α_π	Gamma	1.5	0.25	1.55	0.226	1.55	0.232
α_y	Gamma	0.125	0.1	0.0863	0.0629	0.0889	0.0681
α_ξ	Gamma	0.125	0.1	0.0657	0.0199	0.0736	0.0253
ρ_{ξ^h}	Beta	0.5	0.2	0.567	0.197	0.601	0.188
ρ_{ξ^c}	Beta	0.5	0.2	0.916	0.0289	0.907	0.0423
ρ_{z^m}	Beta	0.5	0.2	0.484	0.212	0.54	0.212
ρ_{g^m}	Beta	0.3	0.2	0.238	0.163	0.265	0.18
σ_{ξ^h}	Gamma	0.01	0.01	0.0101	0.00433	0.0111	0.00547
σ_{ξ^c}	Gamma	0.01	0.01	0.0248	0.00577	0.0234	0.00594
σ_{z^m}	Gamma	0.0025	0.0025	0.00061	0.000384	0.000631	0.000387
σ_{g^m}	Gamma	0.0025	0.0025	0.000889	0.000426	0.000842	0.000422
γ_π	Uniform	0.5	0.289	0.296	0.16	0.339	0.17
γ_I	Uniform	0.5	0.289	0.217	0.111	0.18	0.0834
δ	Uniform	0.5	0.289	0.353	0.081	0.333	0.0794
ρ_z	Beta	0.5	0.2	0.482	0.196	0.484	0.193
ρ_g	Beta	0.3	0.2	0.195	0.0968	0.19	0.0944
σ_z	Gamma	0.01	0.01	0.00162	0.00128	0.00159	0.00129
σ_g	Gamma	0.01	0.01	0.00791	0.00104	0.00795	0.00106
R_{11}	Uniform	3.16e-06	1.82e-06	3.92e-06	1.72e-06	3.95e-06	1.74e-06
R_{22}	Uniform	2.1e-06	1.21e-06	3.71e-06	3.06e-07	3.72e-06	3.02e-07
R_{33}	Uniform	2.49e-07	1.44e-07	2.32e-07	1.4e-07	2.34e-07	1.41e-07
χ				0.98	0.133	0.95	0.124

Notes. The time unit is one quarter. The posterior distribution of χ is derived from the corresponding distribution of δ. The parameters R_{ii}, for $i = 1, \ldots, 3$ are the diagonal elements of the variance-covariance matrix of measurement errors.

5 Results

This section characterizes the macroeconomic consequences of permanent monetary shocks and the central bank information channel. It begins with the full information environment and then turns to the one with imperfect information.

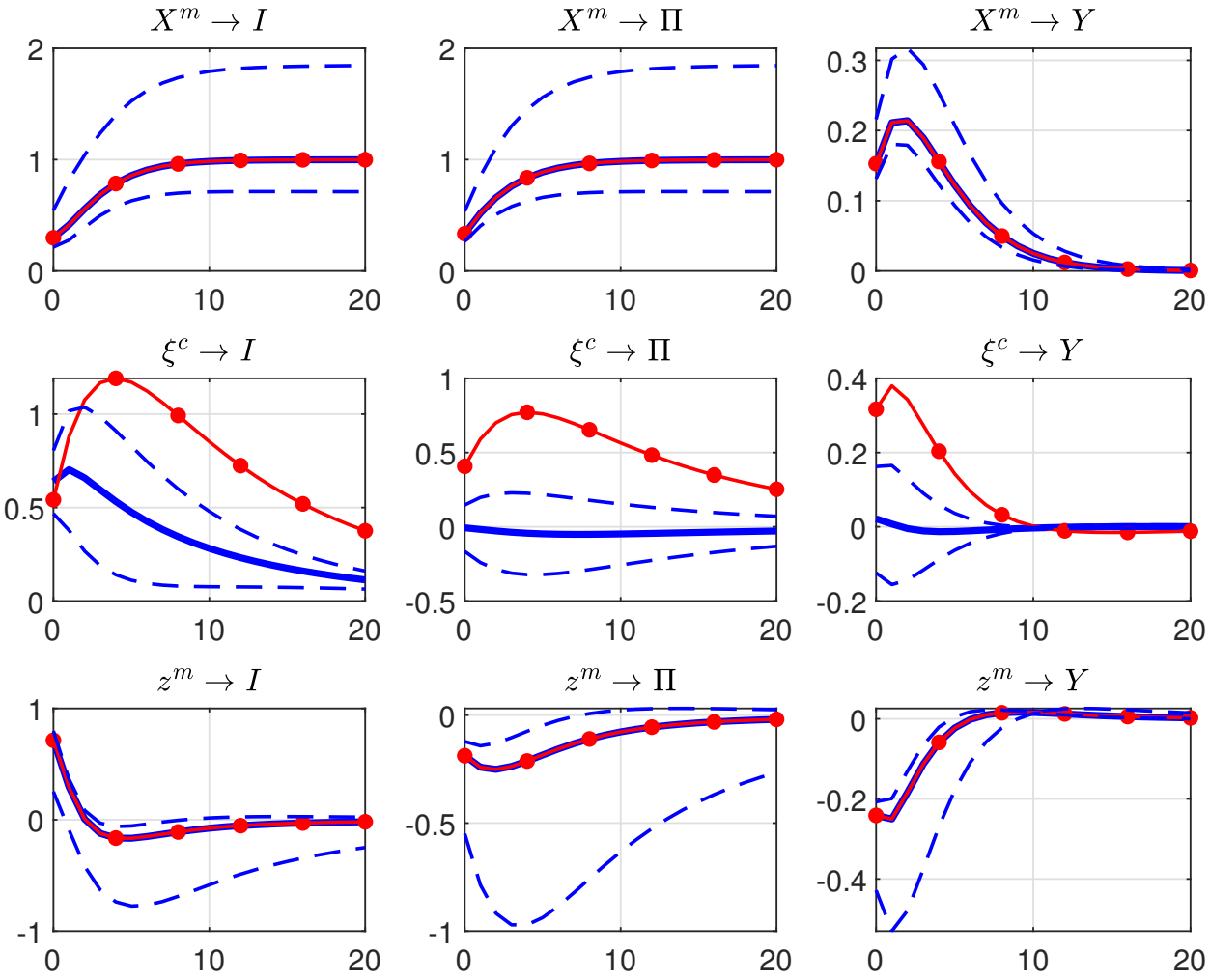
5.1 Full Information

Under full information, households and firms can observe separately all shocks buffeting the economy, that is, the stationary and nonstationary productivity shocks, z_t and g_t, the two demand shocks, ξ_t^h and ξ_t^c, and the stationary and nonstationary monetary shocks, z_t^m and X_t^m.

Figure 1 displays impulse response functions to X_t^m, ξ_t^c, and z_t^m under full information. The permanent monetary shock produces a significant neo-Fisher effect: In response to an innovation in X_t^m that increases the nominal interest rate and inflation by 1 annual percentage point in the long run, the interest rate, inflation, and output all increase in the short run. Just two quarters after the shock, inflation is already halfway to its higher long-run value. Output rises by 15 basis point on impact, reaches a peak of 20 basis points two quarters after the shock and then converges to its pre-shock level gradually over time. The intuition behind this result is as follows. When the innovation in X_t^m is revealed, firms know that inflation will be higher in the long run, as an increase in X_t^m is akin to an increase in the central bank's inflation target. Because firms face price adjustment costs, it pays for them to start increasing prices already in the short run. Inflation actually adjusts faster than the interest rate in the short run, which causes a fall in the real interest rate and consequently an expansion in aggregate activity.

The estimated model also predicts a significant central bank information channel of monetary policy, as the central bank actively stabilizes demand shocks: Figure 1 shows that in response to a one-standard-deviation innovation in ξ_t^c, the policy rate increases by 64 basis

Figure 1: Estimated Impulse Responses Under Full Information



Notes. The horizontal axes measure quarters after the shock. Solid lines are posterior means, dashed lines are 95% asymmetric error bands computed using the Sims-Zha (1999) method, and circled lines are posterior means restricting $\alpha_\xi = 0$ without reestimation. Inflation, Π_t, and the nominal interest rate, I_t, are deviations from pre-shock levels and are expressed in percentage points per year. Output, Y_t, is measured in percent deviations from trend. The size of the permanent monetary shock X_t^m is set so as to increase the nominal interest rate by 1 annual percentage point in the long run on average. The size of the transitory monetary shock z_t^m is 1 annual percentage point on impact. And the size of the demand shock ξ_t^c is one standard deviation.

points and reaches a peak of 70 basis points one quarter after the shock. Neither inflation nor output respond significantly to the demand shock. But this is a reflection of the central bank successfully stabilizing the economy in response to this type of disturbance ($\alpha_\xi > 0$). To see this, Figure 1 also displays with circled lines the economy's response to a ξ_t^c shock under the counterfactual that the central bank does not respond directly to this type of disturbance, that is, when α_ξ is set to zero. In this case, the demand shock causes a large increase in inflation and output. Interestingly, the interest rate increases almost twice as much in the absence of a central bank information channel than in its presence. This is because when $\alpha_\xi = 0$, the increases in inflation and output call for large increases in the policy rate through the standard systematic components of the Taylor rule.

The estimated positive value of the policy coefficient α_ξ could be conducive to a more efficient allocation. To see this, note that it achieves a virtually complete stabilization of output and inflation in response to demand shocks, bringing the economy closer to the outcome that would obtain under full price flexibility (i.e., a mute output response). Responding directly to demand shocks allows the central bank to raise the real interest rate in line with the increase in the natural rate triggered by the demand shock, while keeping inflation and output close to their target values, something that would be difficult to achieve under a Taylor rule without the central bank information channel.

Under full information, a monetary tightening stemming from an increase in the transitory monetary shock z_t^m produces conventional effects, namely, a fall in inflation and output.

Table 3 displays the variance decomposition of the variables of interest. According to the estimated model, both the neo-Fisher effect and the central bank information effect are relevant drivers of the data. The permanent monetary shock, X_t^m, is important in explaining inflation, accounting for one third of the variance of changes in Π_t. The central bank information shock ξ_t^c plays a key role along a different dimension. It primarily drives movements in the nominal interest rate, explaining half of the variance of changes in I_t. This shock has a minor role in explaining movements in inflation and output. However, as

Table 3: Variance Decomposition Under Full Information

	α_ξ Estimated			$\alpha_\xi = 0$		
Shock	ΔI_t	$\Delta\Pi_t$	ΔY_t	ΔI_t	$\Delta\Pi_t$	ΔY_t
Permanent Monetary Shock, g_t^m	6.0	33.2	1.8	5.6	15.1	1.5
Central Bank Information Shock, ξ_t^c	52.9	6.8	1.5	56.1	57.2	17.9
Transitory Interest-Rate Shock, z_t^m	8.6	4.0	1.3	7.9	1.8	1.0
Preference Shock, ξ_t^h	25.4	44.0	25.5	23.8	20.8	21.8
Transitory Productivity Shock, z_t	4.7	7.8	1.6	4.4	3.3	1.3
Permanent Productivity Shock, g_t	2.5	4.3	68.3	2.3	1.9	56.5

Notes. Posterior means. Shares are expressed in percent. The variables ΔI_t, $\Delta\Pi_t$, and ΔY_t denote the change in the nominal interest rate, the change in the inflation rate, and the output growth rate. The last three columns display the variance decomposition under the counterfactual $\alpha_\xi = 0$ without reestimation.

explained in the analysis of impulse responses, this is due to the central bank's success in stabilizing the economy in response to demand shocks: The last three columns of Table 3 display the variance decomposition when α_ξ is restricted to be 0, that is, when the central bank does not respond directly to the demand shock ξ_t^c. In this case, ξ_t^c explains 57 percent of changes in inflation and 18 percent of changes in output.

Finally, Table 3 shows that once one allows for neo-Fisher and central bank information shocks, the standard monetary shock z_t^m plays a negligible role in explaining movements in output or inflation.

5.2 Imperfect Information

Under imperfect information, private agents observe the stochastic component of the Taylor rule, $\widetilde{\Omega}_t^m \equiv {X_t^m}^{1-\alpha_\pi(1-\gamma_I)} {X_{t-1}^m}^{-\gamma_I} e^{z_t^m + \alpha_\xi \xi_t^c}$, but not the individual shocks that drive it, namely, the two monetary shocks, X_t^m and z_t^m, and the demand shock ξ_t^c. Similarly, under imperfect information private agents observe the preference shifter ξ_t, but not separately its two components ξ_t^h and ξ_t^c. As it turns out, however, the equilibrium dynamics are not too different relative to the perfect information case.

Figure 2 displays impulse responses to X_t^m, ξ_t^c, and z_t^m under imperfect information.

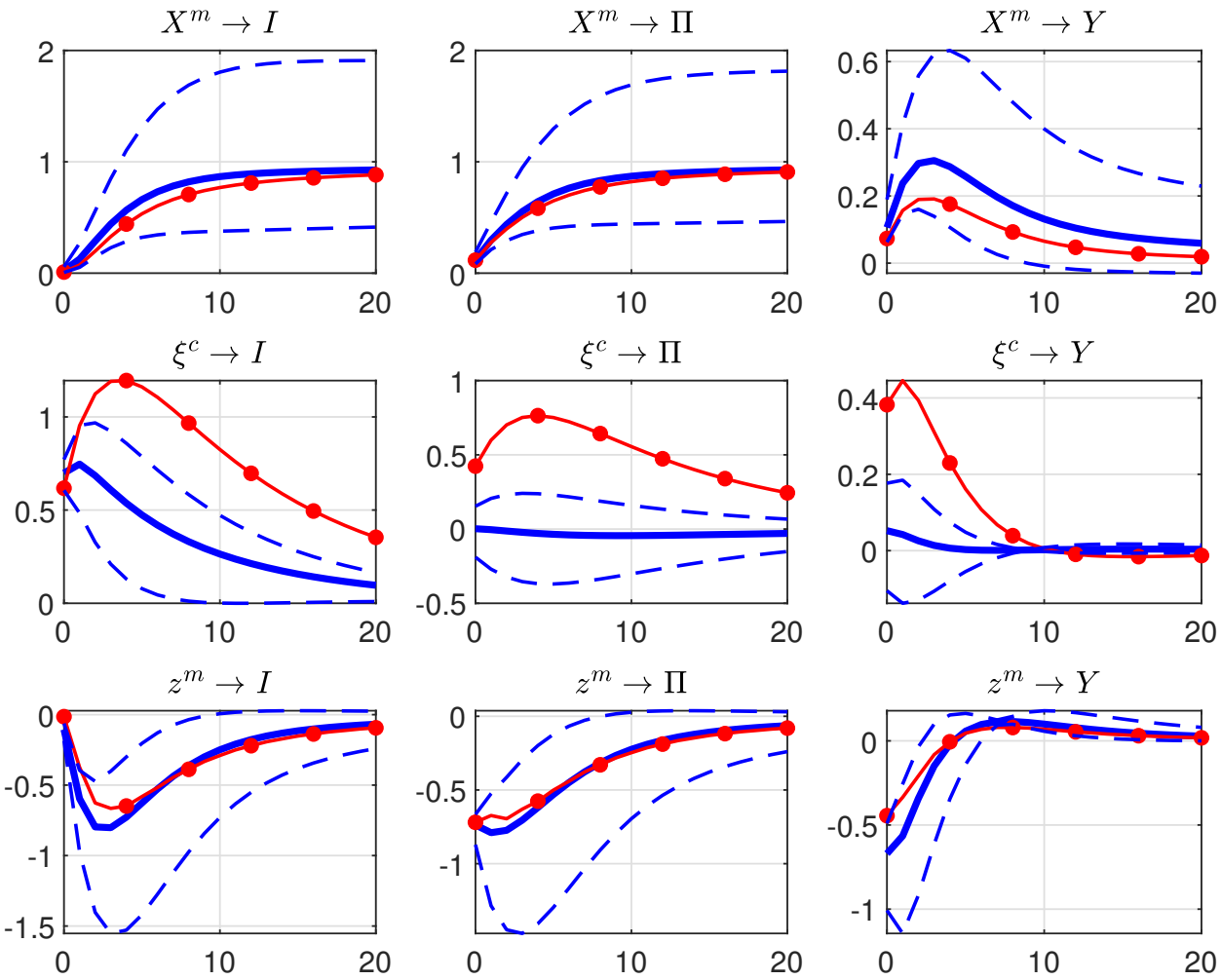
An increase in the permanent monetary shock continues to produce a significant neo-Fisher effect, as the interest rate, inflation, and output all increase in the short run in response to an increase in X_t^m. Quantitatively, the reaction of the nominal interest rate on impact is somewhat more muted relative to the full information case, but overall the economy continues to reach its long-run position relatively quickly and the expansionary effect on output is little changed.

A positive innovation in the demand shock ξ_t^c continues to produce a significant and persistent tightening, and the monetary authority achieves a virtually perfect stabilization of inflation and output. When we counterfactually shut down the central bank's reaction to ξ_t^c by setting α_ξ to zero, both inflation and output display large and significant increases, just as under full information.

One difference with the full information case is the response to a stationary monetary shock z_t^m, especially the response of the nominal interest rate. Under imperfect information, when z_t^m goes up, private agents know that the increase in the exogenous component of the Taylor rule is most likely not due to a demand shock, since they know that ξ_t did not move (and that a move in one of its components, ξ_t^h or ξ_t^c, perfectly offset by the other is highly unlikely). So they are quite sure that they are in the presence of a monetary shock. But they are not sure whether the monetary innovation is a stationary or a permanent one (z_t^m or X_t^m). Since the exogenous component of the Taylor rule, $\widetilde{\Omega}_t^m \equiv {X_t^m}^{1-\alpha_\pi(1-\gamma_I)} {X_{t-1}^m}^{-\gamma_I} e^{z_t^m + \alpha_\xi \xi_t^c}$, is increasing in z_t^m but decreasing in X_t^m, agents think that the economy could have been hit either with an increase in z_t^m or with a decrease in X_t^m. Both of these possibilities tend to produce falls in inflation and output. So the transition in response to an increase in z_t^m is characterized by paths of Π_t and Y_t below trend. The response of the interest rate is ambiguous, because an increase in z_t^m induces an increase in I_t, whereas a fall in X_t^m induces a decrease.

Table 4 displays a variance decomposition under imperfect information. Movements in X_t^m explain about 20 percent of the variance of inflation changes, suggesting that the

Figure 2: Estimated Impulse Responses Under Imperfect Information



Notes. The horizontal axes measure quarters after the shock. Solid lines are posterior means, dashed lines are 95% asymmetric error bands computed using the Sims-Zha (1999) method, and circled lines are posterior means restricting $\alpha_\xi = 0$ without reestimation. Inflation, Π_t, and the nominal interest rate, I_t, are deviations from pre-shock levels and are expressed in percentage points per year. Output, Y_t, is measured in percent deviations from trend. The size of the permanent monetary shock X_t^m is set so as to increase the nominal interest rate by 1 annual percentage point in the long run on average. The size of the transitory monetary shock z_t^m is 1 annual percentage point on impact. And the size of the demand shock ξ_t^c is one standard deviation.

Table 4: Variance Decomposition Under Imperfect information

	α_ξ Estimated			$\alpha_\xi = 0$		
Shock	ΔI_t	$\Delta\Pi_t$	ΔY_t	ΔI_t	$\Delta\Pi_t$	ΔY_t
Permanent Monetary Shock, g_t^m	3.8	19.6	2.3	3.7	9.7	1.1
Central Bank Information Shock, ξ_t^c	61.3	6.7	2.0	68.4	62.6	25.7
Transitory Interest-Rate Shock, z_t^m	5.3	23.7	6.7	5.3	8.2	2.4
Preference Shock, ξ_t^h	22.6	39.4	15.9	15.7	14.8	7.7
Transitory Productivity Shock, z_t	4.9	7.3	1.5	4.9	3.3	1.3
Permanent Productivity Shock, g_t	2.1	3.2	71.7	2.0	1.5	61.9

Notes. Posterior means, in percent. The variables ΔI_t, $\Delta\Pi_t$, and ΔY_t denote the change in the nominal interest rate, the change in the inflation rate, and the output growth rate.

permanent monetary shock continues to be an important driver of inflation, though less so than under full information. The central bank information channel is somewhat stronger under imperfect information. Now, the demand shock ξ_t^c explains 61 percent of movements in the nominal interest rate, up from 53 percent under full information. As under full information, this shock contributes relatively little to explaining movements in inflation and output, but this is again due to the central bank's success in isolating the economy from this type of disturbances. When we counterfactually shut down the central bank's response to demand shocks, this source of disturbance becomes the main driver of inflation and a significant driver of aggregate activity.

The reason why the full and imperfect information cases deliver similar equilibrium dynamics is that the three shocks that are imperfectly observed, namely, the permanent monetary shock, the demand shock to which the central bank responds, and the standard monetary shock, produce starkly different dynamics after the period of impact, allowing private agents to learn quickly about the nature of the disturbances buffeting the economy. The overall takeaway of this and the preceding subsections is that in negotiating the tradeoff between complexity of the model and economic insight, the full information benchmark is a reasonable framework to interpret the data.

We also consider an intermediate case between the full information and the imperfect

information specifications in which private agents can observe ξ_t and $\widetilde{\Omega}_t^m$, as in the imperfect information case, and also noisy signals of the individual shocks g_t^m, ξ_t^c, and z_t^m. The result of this robustness analysis, which in the interest of space we do not present, is that the noise components are imprecisely estimated and the predictions of the model are not significantly different from those of the imperfect information specification. This should not come as a surprise, given that the full and imperfect information cases were found to deliver similar results and the noisy-information case is an intermediate one.

5.3 High Frequency Information

We now estimate the model with an additional observation equation that incorporates a measure of monetary policy shocks derived from high frequency information. In the context of the present model, high frequency monetary shocks capture unforecastable changes in the exogenous stochastic component of the Taylor rule, namely, $\ln \widetilde{\Omega}_t^m - E_{t-1} \ln \widetilde{\Omega}_t^m$, where $\widetilde{\Omega}_t^m$ is defined in equation (19). This is because within the short time span used to estimate surprise changes in interest rates, typically a 30-minute window around FOMC announcements, one can safely assume that inflation and output—the two macroeconomic indicators entering the Taylor rule—are unchanged.

Letting hf_t be the high frequency monetary shock, we assume that it is linked to the forecast error of the exogenous component of the Taylor rule by the expression

$$hf_t = \alpha_{hf}(\ln \widetilde{\Omega}_t^m - E_{t-1} \ln \widetilde{\Omega}_t^m) + \mu_t^{hf},$$

where α_{hf} is a scaling parameter and μ_t^{hf} is an exogenous i.i.d. disturbance with mean zero and standard deviation $\sigma_{\mu^{hf}}$, capturing measurement error. From the definition of $\widetilde{\Omega}_t^m$, we can write hf_t in terms of the innovations of the individual shocks entering the interest-rate

feedback rule. This gives

$$hf_t = \alpha_{hf}\left[\sigma_{zm}\epsilon_t^{zm} + \alpha_\xi \sigma_{\xi^c}\epsilon_t^{\xi^c} + (1-\alpha_\pi(1-\gamma_I))\sigma_{g^m}\epsilon_t^{g^m}\right] + \mu_t^{hf}, \tag{42}$$

where the disturbances ϵ_t^x, for $x = z^m, \xi^c$, and g^m are defined in equation (20). The high frequency shock is driven by innovations in the traditional monetary shock, $\epsilon_t^{z^m}$, the central bank information shock, $\epsilon_t^{\xi^c}$, and the neo-Fisher shock, $\epsilon_t^{g^m}$.

We estimate the full-information model of section 5.1 incorporating equation (42) as an observation equation. We measure hf_t using the time series NS provided by Acosta (2023), which he constructs following the methodology employed by Nakamura and Steinsson (2018). For each quarter, hf_t is the sum of the high frequency monetary shocks NS occurring at the FOMC announcements taking place in that particular quarter. The NS shock is an index with mean 0 and unit standard deviation. Without loss of generality, we normalize its standard deviation to 0.0025, or 1 percent per year. The time series of hf_t starts in 1995:Q1, whereas the time series of the other observables used in the estimation of the model (changes in the policy rate, changes in inflation, and the output growth rate) start in 1961:Q3. We therefore treat values of hf_t prior to 1995:Q1 as missing observations (Harvey, 1989).

The present formulation features two additional parameters vis-à-vis the baseline model of section 5.1: α_{hf} and $\sigma^2_{\mu^{hf}}$. We re-estimate all parameters of the model using Bayesian methods. We assume that $\sigma^2_{\mu^{hf}}$ has a gamma prior distribution with mean and standard deviation equal to 10 percent of the observed variance of hf_t. To obtain a reference value for the magnitude of the scaling parameter α_{hf}, we start with the expression $\text{var}(hf_t) = \alpha^2_{hf}\text{var}(\ln\widetilde{\Omega}^m_t - E_{t-1}\ln\widetilde{\Omega}^m_t) + \sigma^2_{\mu^{hf}}$, which follows directly from equation (42). The posterior mean of $\text{std}(\ln\widetilde{\Omega}^m_t - E_{t-1}\ln\widetilde{\Omega}^m_t)$ in the full-information estimation of section 5.1 is 0.0017, or 0.68 percent per year. Evaluating the above expression at this value, setting the variance of hf_t at its normalized value, and the variance of μ_t^{hf} at its prior mean, gives $0.0025^2 = \alpha^2_{hf} \times 0.0017^2 + 0.1 \times 0.0025^2$, which implies a reference value for α_{hf} of 1.4. We assume a

uniform prior distribution for α_{hf} with lower bound 0 and upper bound 4. This upper bound is about 3 times the reference value of α_{hf}. The prior distributions of all other estimated parameters are as shown in Table 2.

Table A1 in the appendix displays prior and posterior parameter distributions. The posterior mean of α_{hf} is 2.1, which is similar to its prior mean of 2. However, its posterior standard deviation is 0.54, which is about half as large as its prior counterpart of 1.1, suggesting that the scaling parameter is identified. The posterior variance of the measurement error μ_t^{hf} is 34 percent of the variance of the observable hf_t (more than three times as large as its prior value), indicating that according to the model the time series hf_t contains substantial measurement error.

The parameters that are common to the baseline model of section 5.1 overall have similar posterior distributions, with two exceptions. One is the parameters defining the law of motion of the neo-Fisher shock, σ_{g^m} and ρ_{g^m}, which now imply a standard deviation of g_t^m that is 73 percent larger than its counterpart in the model estimated without high frequency shocks. The other is the parameter α_ξ governing the responsiveness of the central bank to demand shocks, which falls by about half relative to the baseline estimation. Thus, including information from high frequency monetary shocks appears to elevate the estimated importance of both the neo-Fisher shock and the central bank information shock. To understand the latter effect, recall that as α_ξ becomes smaller, the central bank is less responsive to demand disturbances and therefore ξ_t^c has a larger impact on inflation and output.

Table 5 confirms this interpretation. It displays the predicted variance decomposition of changes in the policy rate, inflation, and output. The neo-Fisher shock, g_t^m now explains more than 70 percent of the variance of inflation changes, more than twice its share in the estimation without high frequency information. Much of this increase in the importance of the neo-Fisher shock comes at the expense of a reduction in the importance of ξ_t^h, the demand shock to which the central bank does not respond, whose share falls from 44 percent to 6 percent. The central bank information effect also becomes more important than in the

Table 5: Estimation with High Frequency Shocks: Variance Decomposition

	α_ξ Estimated			$\alpha_\xi = 0$		
Shock	ΔI_t	$\Delta\Pi_t$	ΔY_t	ΔI_t	$\Delta\Pi_t$	ΔY_t
Permanent Monetary Shock, g_t^m	20.8	72.2	3.8	15.5	21.8	2.5
Central Bank Information Shock, ξ_t^c	61.9	10.7	5.4	71.0	73.5	37.6
Transitory Interest-Rate Shock, z_t^m	11.0	2.3	1.7	8.8	0.6	1.1
Preference Shock, ξ_t^h	2.6	6.3	5.5	2.0	1.8	3.8
Transitory Productivity Shock, z_t	2.8	6.4	1.4	2.1	1.8	0.9
Permanent Productivity Shock, g_t	0.8	2.0	82.2	0.6	0.5	54.2

Notes. Posterior means, in percent. The variables ΔI_t, $\Delta\Pi_t$, and ΔY_t denote the change in the nominal interest rate, the change in the inflation rate, and the output growth rate.

baseline estimation. It now accounts for more than 60 percent of the variance of changes in the policy rate, up from 53 percent in the baseline model. Further, when we counterfactually shut down the central bank information effect by setting α_ξ equal to zero, the demand shock ξ_t^c now explains a larger fraction of changes in inflation and output (73 and 38 percent, respectively, compared to 57 and 18 percent in the baseline model). Finally, as in the baseline estimation, the traditional monetary policy shock, z_t^m, is estimated to play a minor role in explaining movements in the interest rate, inflation, or output.

Figure A1 in the appendix displays the impulse responses to the neo-Fisher shock, the central-bank information shock, and the standard monetary shock. The results are in line with those obtained when the estimation does not include high-frequency shocks except that now output is somewhat more responsive to tightenings originated in central-bank information shocks, which is consistent with the smaller estimated value of α_ξ.

Overall, the findings of this section suggest that including information on high frequency monetary shocks in the estimation of the model significantly increases the importance of the neo-Fisher shock and to a lesser extend also that of the central bank information shock.

5.4 Longer Maturity Yields

Incorporating information about long-maturity bonds in the model estimation is of interest because different shocks can imply distinct predictions for long-term yields. For instance, Figure 1 illustrates that the short rate converges more or less monotonically to its long-run level in response to both the permanent monetary shock, X_t^m, and the demand shock to which the central bank reacts, ξ_t^c. However, the permanent monetary shock has a lasting effect on the short rate, whereas the demand shock does not. Since long rates are approximately moving averages of short rates, one would expect long rates to rise more than short rates in the short run following a positive innovation in X_t^m, but to rise less following a positive innovation in ξ_t^c. Therefore, including both short- and long-maturity yields in the estimation should improve the identification of the neo-Fisher and central bank information effects.

Accordingly, we now assume that households can hold bonds of different maturities. With this asset structure, the household budget constraint can be expressed as

$$P_t C_t + \sum_{h=1}^{H} Q_{ht} B_{ht} + T_t = \sum_{h=1}^{H} Q_{h-1t} B_{ht-1} + W_t h_t + \Phi_t, \tag{3}$$

where the variable B_{ht} denotes a zero-coupon bond trading at price Q_{ht} in period t and paying 1 dollar in period $t+h$, with $Q_{0t} \equiv 1$. The zero-coupon nominal yield on an h-period bond, denoted I_t^h, is defined as

$$1 + I_t^h = Q_{ht}^{-\frac{1}{h}},$$

with $I_t \equiv I_t^1$. The first-order conditions of the household's optimization problem with respect to bond holdings are

$$\Lambda_t Q_{ht} = \beta E_t \left[\frac{\Lambda_{t+1}}{1 + \Pi_{t+1}} Q_{h-1t+1} \right],$$

for $h = 1, 2, \ldots, H$. All other aspects of the model are as in the baseline formulation of section 2.

In the estimation of the model we include as an observable, in addition to changes in

the short-run nominal rate, the short-run interest rate-inflation differential, and changes in output, the change in the 12-quarter zero-coupon yield, ΔI_t^{12}. For I_t^{12} we use the estimate of the 3-year zero coupon Treasury yield produced by Gürkaynak, Sack, and Wright (2006).[1] To preserve comparability with the baseline results of section 5.1, the estimation does not include the high-frequency monetary shock as an observable. The results presented below are robust to including this variable as a fifth observable.

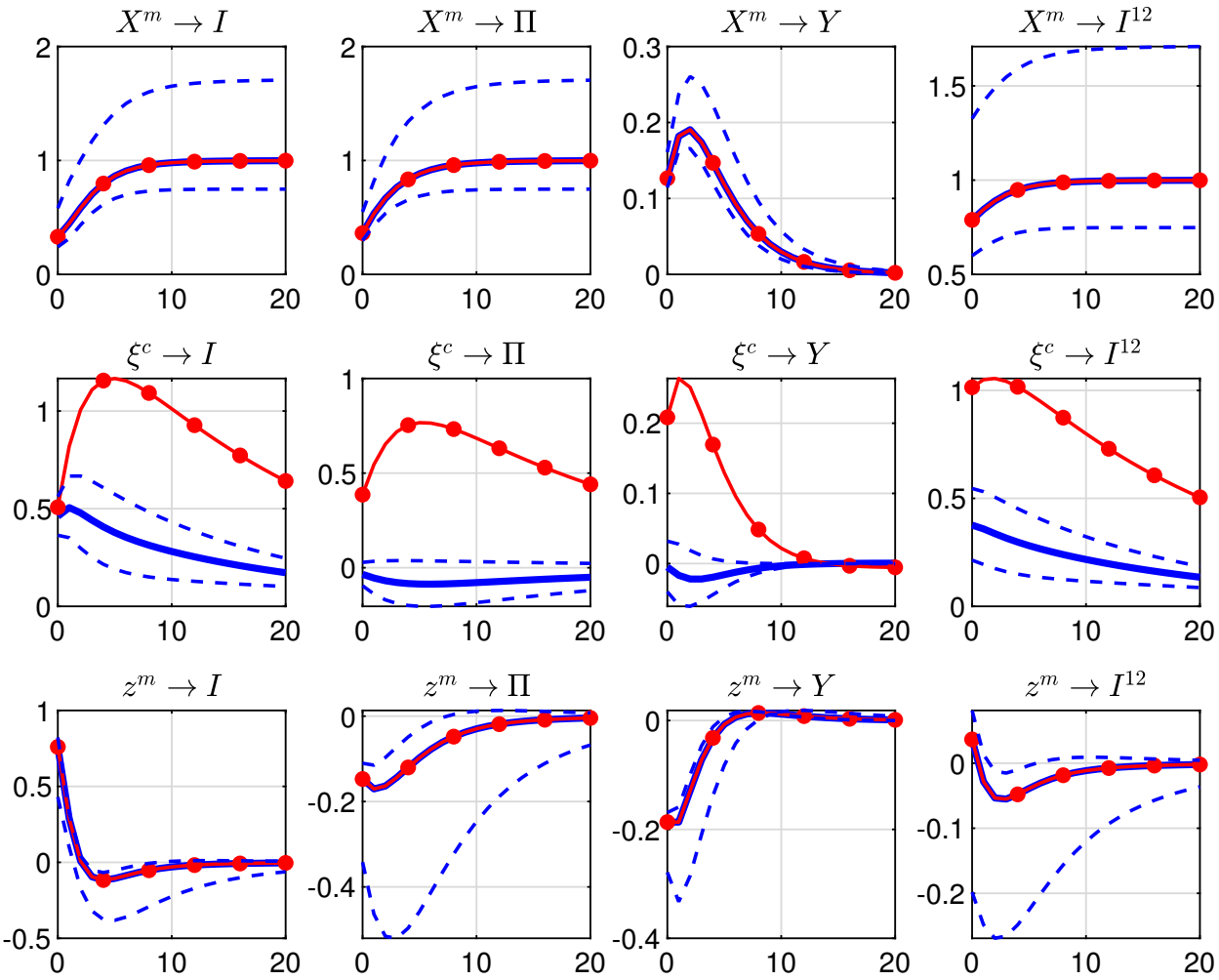
Table A2 in the appendix displays posterior moments of the estimated parameters. As conjectured, the inclusion of information on longer maturity yields produces parameter estimates with smaller posterior standard deviations, indicating that the new observable is informative. The parameter estimates are in general stable relative to the baseline estimation with three observables (Table 2).

Figure 3 displays posterior impulse responses when the model is estimated with the 12-quarter zero-coupon yield as an additional observable. The predicted dynamics are in general similar to the baseline case. The error bands are now narrower, confirming that the inclusion of the long-term yield helps to identify the parameters of the model. The neo-Fisher shock continues to produce, in response to an unexpected tightening, a significant increase in inflation and output in the short run. The central bank information shock does not have a significant effect on these two variables. But under the counterfactual $\alpha_\xi = 0$, an unexpected increase in the demand shock ξ_t^c continues to cause sizable increases in inflation and output in the short run. The response of the twelve-quarter yield is a forward moving average of that of the short-run interest rate, and as a result it inherits much of its qualitative characteristics.

Table A3 in the appendix displays the variance decomposition of endogenous variables of interest. The neo-Fisher shock X_t^m continues to be the second most important driver of inflation changes, after the demand shock to which the central bank does not respond, ξ_t^h.

[1]Specifically, we use their series SVNY03, which is publicly available at the Board of Governors of the Federal Reserve System's website, www.federalreserve.gov/data/nominal-yield-curve.htm. The raw data is daily and we transform it to a quarterly frequency by simple averaging. It is available for our sample period, 1961Q3 to 2019Q4.

Figure 3: Estimation with the Twelve-Quarter Yield: Impulse Responses



Notes. The horizontal axes measure quarters after the shock. Solid lines are posterior means, dashed lines are 95% asymmetric error bands computed using the Sims-Zha (1999) method, and circled lines are posterior means restricting $\alpha_\xi = 0$ without reestimation. Inflation, Π_t, the nominal interest rate, I_t, and the twelve-quarter yield, I_t^{12}, are deviations from pre-shock levels and are expressed in percentage points per year. Output, Y_t, is measured in percent deviations from trend. The size of the permanent monetary shock X_t^m is set so as to increase the nominal interest rate by 1 annual percentage point in the long run on average. The size of the transitory monetary shock z_t^m is 1 annual percentage point on impact. And the size of the demand shock ξ_t^c is one standard deviation.

It explains 19.2 percent of the variance of $\Delta\Pi_t$, which is a sizeable fraction, but smaller than the one obtained in the baseline estimation (33 percent). It also continues to be the case that the CBI shock explains a significant fraction of the variance of changes in the policy rate 25 percent, though less than without using the 12-quarter yield as observable, 53 percent. This shock continues to explain little of changes in inflation and output, but under the counterfactual $\alpha_\xi = 0$ it is a key driver of inflation changes, with a share of 52 percent. The standard monetary shock z_t^m now explains a larger fraction of changes in the short term interest rate, 25 percent versus 8.6 percent in the estimation with three observables. This is due to a significantly higher estimated standard deviation of the innovation z_t^m. However, it explains a small fraction of movements in the 12-quarter zero-coupon rate and continues to play a negligible role in explaining movements in inflation and output. The preference shock to which the central bank does not respond, ξ_t^h becomes more important, now explaining 58 percent of the variance of inflation changes and 41 percent of that of output changes.

6 Conclusion

The neo-Fisher effect and the central bank information effect produce similar outcomes: under both, a monetary tightening triggers an increase in inflation and an expansion in real activity. Separate estimates of these effects run the risk of confounding one with the other. In this paper we present the first attempt to jointly estimate the two effects. To this end, we incorporate into a dynamic general equilibrium model with nominal and real rigidities permanent monetary shocks—the neo-Fisher channel—and an interest-rate feedback rule that responds directly to an aggregate demand shock in addition to inflation and output—the central bank information channel. We estimate the model using postwar U.S. data, including information on high-frequency monetary shocks and long-term yields.

The estimated model suggests that both the neo-Fisher shock and the central bank information shock are relevant drivers of the U.S. economy. Permanent monetary shocks

explain a significant fraction of movements in inflation, and the demand shock to which the monetary authority responds explains a large share of variations in the nominal interest rate. In equilibrium, this demand shock does not explain much of the volatility of inflation or output. However, through counterfactual exercises we establish that this is a reflection of the CBI channel's success in stabilizing the aggregate effects of this type of disturbance.

The estimated model predicts that through its direct response to aggregate shocks, the central bank brings the real allocation closer to the one that would arise under flexible prices. This finding suggests that the central bank information channel could be conducive to a more efficient outcome. A promising line of investigation is ascertaining precisely how much closer the direct response of central banks to aggregate disturbances—of which the current formulation of the CBI channel is just one example—can bring the economy to the efficiency frontier. We leave this task for future work.

References

Acosta, Miguel, "The Perceived Causes of Monetary Policy Surprises," Federal Reserve Board, working paper, February 26, 2023.

Barakchian, S. Mahdi, and Christopher Crowe, "Monetary policy matters: Evidence from new shocks data," *Journal of Monetary Economics 60,* November 2013, 950–966.

Bauer, Michael D., and Eric T. Swanson, "An Alternative Explanation for the 'Fed Information Effect'," *American Economic Review 113,* 2023, 664–700.

Bouakez, Hafedh, and Takashi Kano, "Deciphering the Neo-Fisherian Effect," working paper, HEC Montreal, June 2024.

Campbell, Jeffrey R., Jonas D.M. Fisher, Alejandro Justiniano, and Leonardo Melosi, "Forward Guidance and Macroeconomic Outcomes since the Financial Crisis," in Martin Eichenbaum and Jonathan A. Parker (eds.), *NBER Macroeconomics Annual 31,* The University of Chicago Press, Chicago, 2017, 283–357.

Carvalho, Alexandre, João Valle e Azevedo, and Pedro Pires Ribeiro, "Permanent and Temporary Monetary Policy Shocks and the Dynamics of Exchange Rates," *Journal of International Economics 147,* 2024, 103871.

Cieslak, Anna, and Andreas Schrimpf, "Non-monetary News in Central Bank Communication," *Journal of International Economics 118,* 2019, 293–315.

Cochrane, John, "Do Higher Interest Rates Raise or Lower Inflation?," working paper, The University of Chicago, 2016.

Faust, Jon, Eric T. Swanson, and Jonathan H. Wright, "Do Federal Reserve Policy Surprises Reveal Superior Information about the Economy?," *Contributions to Macroeconomics 4,* 2004, Article 10.

García-Cicco, Javier, Patricio Goldstein, and Federico Sturzenegger, "Permanent and Transitory Monetary Shocks Around the World," working paper, Universidad de San Andrés, June, 2024.

García-Schmidt, Mariana, "Is the Information Channel of Monetary Policy Alive in Emerg-

ing Markets?," working paper 1017, Central Bank of Chile, June 2024.

García-Schmidt, Mariana, and Michael Woodford, "Are Low Interest Rates Deflationary? A Paradox of Perfect-Foresight Analysis," *American Economic Review 109,* 2019, 86–120.

Garín, Julio, Robert Lester, and Eric Sims, "Raise Rates to Raise Inflation? Neo-Fisherianism in the New Keynesian Model," *Journal of Money, Credit and Banking 50,* February 2018, 243–259.

Gürkaynak, Refet S., Brian Sack, and Jonathan H. Wright, "The U.S. Treasury Yield Curve: 1961 to the Present," Finance and Economics Discussion Series working paper 2006-28, June 2006.

Hansen, Stephen, and Michael McMahon, "Shocking Language: Understanding the Macroeconomic Effects of Central Bank Communication," *Journal of International Economics 99,* 2016, S114–33.

Harvey, Andrew C., *Forecasting, structural time series models and the Kalman filter,* Cambridge, UK: Cambridge University Press, 1989.

Jarociński, Marek, and Peter Karadi, "Deconstructing Monetary Policy Surprises—The Role of Information Shocks," *American Economic Journal: Macroeconomics 12,* 2020, 1–43.

Kerssenfischer, Mark, "Information effects of euro area monetary policy," *Economics Letters 216,* 2022, 110570.

Lukmanova, Elizaveta, and Katrin Rabitsch, "Evidence on monetary transmission and the role of imperfect information: Interest rate versus inflation target shocks," *European Economic Review 158,* 2023, 104557.

Melosi, Leonardo, "Signaling Effects of Monetary Policy," *Review of Economic Studies 84,* 2017, 853–84.

Miranda-Agrippino, Silvia, "Unsurprising Shocks: Information, Premia, and the Monetary Transmission," Centre for Macroeconomics, Discussion Paper 2016-13, 2016.

Nakamura, Emi, and Jón Steinsson, "High Frequency Identification of Monetary Non-Neutrality," *Quarterly Journal of Economics 133,* August 2018, 1283–1330.

Romer, Christina D., and David H. Romer, "Federal Reserve Information and the Behavior of Interest Rates," *American Economic Review 90,* June 2000, 429–457.

Rotemberg, Julio J., "Sticky Prices in the United States," *Journal of Political Economy 90,* December 1982, 1187–1211.

Schmitt-Grohé, Stephanie, and Martín Uribe, "Liquidity Traps: An Interest-Rate-Based Exit Strategy," NBER Working Paper 16514, 2010.

Schmitt-Grohé, Stephanie, and Martín Uribe, "The Making of a Great Contraction with a Liquidity Trap and a Jobless Recovery," NBER Working Paper 18544, 2012.

Schmitt-Grohé, Stephanie, and Martín Uribe, "The Effects of Permanent Monetary Shocks on Exchange Rates and Uncovered Interest Differentials," *Journal of International Economics 135,* March 2022, 103560.

Sims, Christopher, and Tao Zha, "Error Bands for Impulse Responses," *Econometrica 67,* 1999, 1113–1156.

Uribe, Martín, "The Neo-Fisher Effect in the United States and Japan," NBER working paper 23977, October 2017.

Uribe, Martín, "The Neo-Fisher Effect: Econometric Evidence from Empirical and Optimizing Models," *American Economic Journal: Macroeconomics 14,* July 2022, 133–62.

Valle e Azevedo, João, João Ritto, and Pedro Teles, "The Neutrality of Nominal Rates: How Long is the Long Run?," *International Economic Review 63,* November 2022, 1745–1777.

Williamson, Stephen, "Neo-Fisherism: A Radical Idea, or the Most Obvious Solution to the Low-Inflation Problem," *The Regional Economist 24,* July 2016, 5–9.

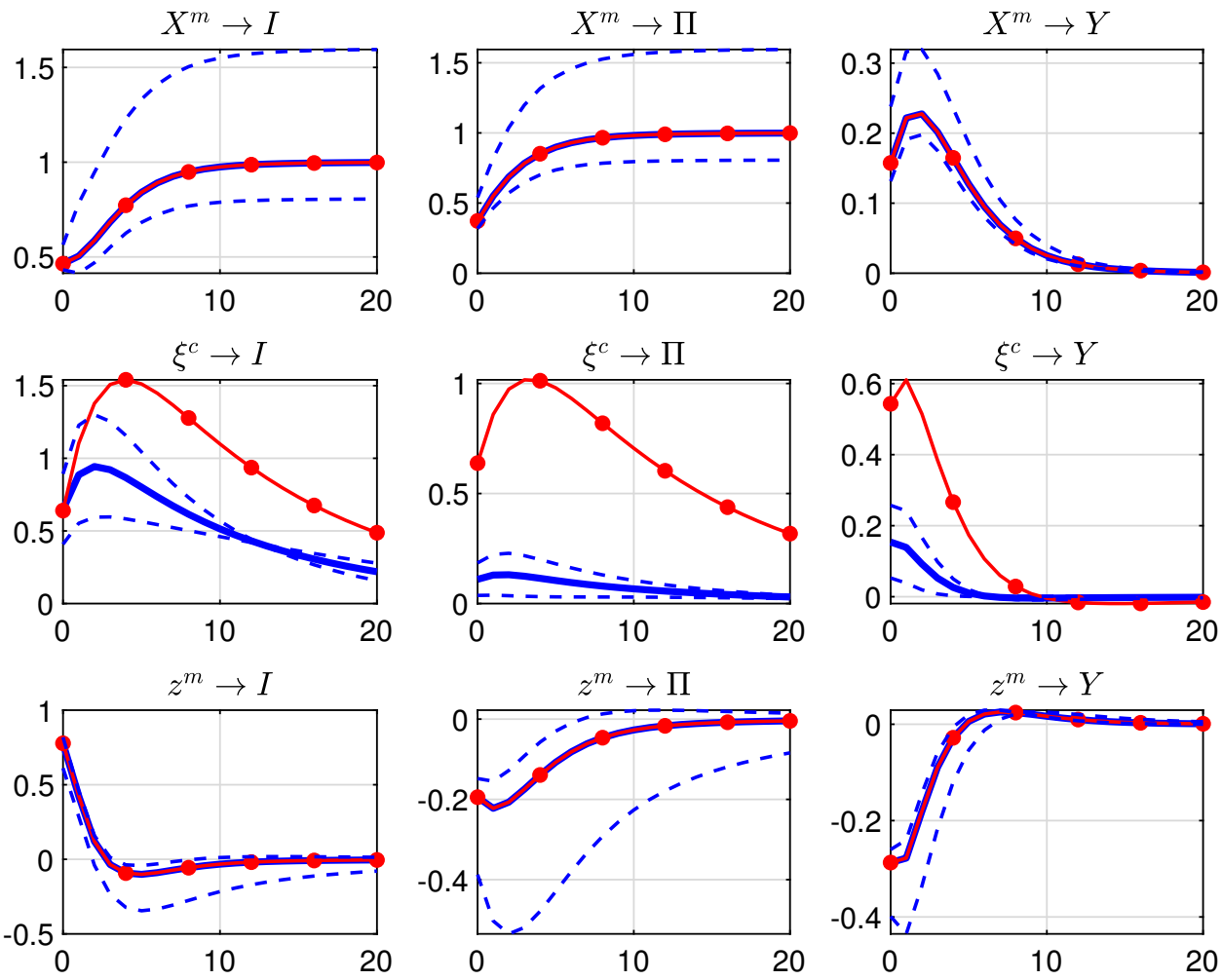
A Appendix

Table A1: Estimation with High Frequency Shocks: Prior and Posterior Parameter Distributions

	Prior Distribution			Posterior Distribution	
Parameter	Distribution	Mean	Std	Mean	Std
ϕ	Gamma	50	20	90	25.3
α_π	Gamma	1.5	0.25	1.53	0.204
α_y	Gamma	0.125	0.1	0.0979	0.0815
α_{hf}	Uniform	2	1.15	2.13	0.541
α_ξ	Gamma	0.125	0.1	0.0388	0.0106
ρ_{ξ^h}	Beta	0.5	0.2	0.499	0.21
ρ_{ξ^c}	Beta	0.5	0.2	0.916	0.0219
ρ_{z^m}	Beta	0.5	0.2	0.335	0.165
ρ_{g^m}	Beta	0.3	0.2	0.165	0.129
σ_{ξ^h}	Gamma	0.01	0.01	0.00322	0.00262
σ_{ξ^c}	Gamma	0.01	0.01	0.0334	0.00715
σ_{z^m}	Gamma	0.0025	0.0025	0.000714	0.00028
σ_{g^m}	Gamma	0.0025	0.0025	0.00156	0.000377
γ_π	Uniform	0.5	0.289	0.369	0.17
γ_I	Uniform	0.5	0.289	0.424	0.112
δ	Uniform	0.5	0.289	0.294	0.0617
ρ_z	Beta	0.5	0.2	0.487	0.201
ρ_g	Beta	0.3	0.2	0.175	0.0849
σ_z	Gamma	0.01	0.01	0.00152	0.00114
σ_g	Gamma	0.01	0.01	0.00822	0.000998
R_{11}	Uniform	3.16e-06	1.82e-06	3.73e-06	1.77e-06
R_{22}	Uniform	2.1e-06	1.21e-06	3.75e-06	2.95e-07
R_{33}	Uniform	2.49e-07	1.44e-07	2.49e-07	1.41e-07
$\sigma^2_{\mu^{hf}}$	Gamma	6.25e-07	6.25e-07	2.15e-06	1.29e-06
χ				0.891	0.0819

Notes. The time unit is one quarter. The posterior distribution of χ is derived from the corresponding distribution of δ. The parameters R_{ii}, for $i = 1, \ldots, 3$ are the diagonal elements of the variance-covariance matrix of measurement errors.

Figure A1: Estimation with High Frequency Shocks: Impulse Responses



Notes. The horizontal axes measure quarters after the shock. Solid lines are posterior means, dashed lines are 95% asymmetric error bands computed using the Sims-Zha (1999) method, and circled lines are posterior means restricting $\alpha_\xi = 0$ without reestimation. Inflation, Π_t, and the nominal interest rate, I_t, are deviations from pre-shock levels and are expressed in percentage points per year. Output, Y_t, is measured in percent deviations from trend. The size of the permanent monetary shock X_t^m is set so as to increase the nominal interest rate by 1 annual percentage point in the long run on average. The size of the transitory monetary shock z_t^m is 1 annual percentage point on impact. And the size of the demand shock ξ_t^c is one standard deviation.

Table A2: Estimation with the Twelve-Quarter Yield: Prior and Posterior Parameter Distributions

	Prior Distribution			Posterior Distribution	
Parameter	Distribution	Mean	Std	Mean	Std
ϕ	Gamma	50	20	84	21.4
α_π	Gamma	1.5	0.25	1.53	0.207
α_y	Gamma	0.125	0.1	0.112	0.0591
α_ξ	Gamma	0.125	0.1	0.0396	0.0123
ρ_{ξ^h}	Beta	0.5	0.2	0.431	0.1
ρ_{ξ^c}	Beta	0.5	0.2	0.954	0.0121
ρ_{z^m}	Beta	0.5	0.2	0.384	0.177
ρ_{g^m}	Beta	0.3	0.2	0.216	0.15
σ_{ξ^h}	Gamma	0.01	0.01	0.0121	0.00221
σ_{ξ^c}	Gamma	0.01	0.01	0.0318	0.00783
σ_{z^m}	Gamma	0.0025	0.0025	0.00114	0.000262
σ_{g^m}	Gamma	0.0025	0.0025	0.000734	0.000193
γ_π	Uniform	0.5	0.289	0.368	0.194
γ_I	Uniform	0.5	0.289	0.221	0.102
δ	Uniform	0.5	0.289	0.388	0.0558
ρ_z	Beta	0.5	0.2	0.46	0.182
ρ_g	Beta	0.3	0.2	0.178	0.0805
σ_z	Gamma	0.01	0.01	0.00167	0.00149
σ_g	Gamma	0.01	0.01	0.00768	0.000907
R_{11}	Uniform	3.16e-06	1.82e-06	3.78e-06	1.75e-06
R_{22}	Uniform	2.1e-06	1.21e-06	3.77e-06	2.89e-07
R_{33}	Uniform	2.49e-07	1.44e-07	3.86e-07	9.42e-08
R_{44}	Uniform	9.68e-08	5.59e-08	1.36e-07	4.63e-08
χ				1.03	0.0938

Notes. The time unit is one quarter. The posterior distribution of χ is derived from the corresponding distribution of δ. The parameters R_{ii}, for $i = 1, \ldots, 4$ are the diagonal elements of the variance-covariance matrix of measurement errors.

Table A3: Estimation with the Twelve-Quarter Yield: Variance Decomposition

	α_ξ Estimated				$\alpha_\xi = 0$			
Shock	ΔI_t	$\Delta\Pi_t$	ΔY_t	ΔI_t^{12}	ΔI_t	$\Delta\Pi_t$	ΔY_t	ΔI_t^{12}
Permanent Monetary Shock, g_t^m	3.6	19.2	0.7	29.3	2.9	9.9	0.6	9.0
Central Bank Information Shock, ξ_t^c	25.4	2.3	0.1	42.6	39.4	52.1	8.3	82.4
Transitory Interest-Rate Shock, z_t^m	24.6	4.8	1.9	0.9	19.7	2.3	1.7	0.3
Preference Shock, ξ_t^h	35.8	57.5	40.6	22.0	29.4	27.9	37.2	6.7
Transitory Productivity Shock, z_t	6.4	9.6	1.9	2.8	5.2	4.6	1.7	0.8
Permanent Productivity Shock, g_t	4.2	6.6	54.8	2.4	3.4	3.2	50.4	0.7

Notes. Posterior means. Shares are expressed in percent. The variables ΔI_t, $\Delta\Pi_t$, and ΔY_t denote the change in the nominal interest rate, the change in the inflation rate, and the output growth rate. The last three columns display the variance decomposition under the counterfactual $\alpha_\xi = 0$ without reestimation.