# Risk and Segmentation in Covered-Interest Parity Arbitrage

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**Abstract**

Prevailing theories of financial intermediation struggle to explain a striking feature of bank-intermediated arbitrages: spreads on these trades exhibit substantial cross-sectional variation in sign and magnitude. We use confidential supervisory data—covering $25 trillion in daily notional exposures on average—to study covered-interest parity (CIP) deviations in currency markets. We uncover three novel forces important for explaining cross-sectional variation in CIP deviations: foreign safe asset scarcity, which makes CIP arbitrage imperfect and risky; market segmentation, with banks specializing in different markets; and concentration of demand. Our findings highlight the presence of risk in ostensibly riskless arbitrage and the importance of segmentation and search frictions in even the most liquid markets.

**JEL Codes:** F3, F31, F65, G1, G13, G15, G2, G23

**Keywords:** basis, covered-interest parity deviation, foreign exchange, safe assets

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# 1 Introduction

Spreads on bank-intermediated arbitrage trades, called bases, have persisted since the 2008 financial crisis, attracting substantial attention from academics and market participants alike. The existence of bases is the clearest evidence that financial intermediaries are not simply a veil, as assumed in classical theories, and that the litany of constraints that intermediaries face are important determinants of asset prices. However, prevailing theories of intermediary asset pricing face a challenge in explaining a striking feature of the data—substantial cross-sectional variation in the signs and magnitude of different basis trades, even within the same asset class. For example, the covered-interest parity (CIP) basis for 1-year currency forwards versus the US dollar averages an annualized $-50$ basis points (bps) for the Japanese Yen but $+6$ bps for the Australian Dollar.[1]

We use granular confidential supervisory data to study the cross-section of covered-interest parity bases to better understand the role of intermediaries in asset prices.[2] We find that three novel forces are especially important for the cross-section of CIP bases. First, intermediaries face search costs to find foreign safe cash bonds in scarce supply, a phenomenon we call *foreign safe asset scarcity*. Since CIP arbitrage requires a position in foreign safe assets, safe asset scarcity means that banks can only imperfectly execute CIP arbitrage, and their CIP trades have meaningful risk. Second, markets are segmented, with banks specializing in different currencies and tenors, so bases reflect their specializing banks' constraints. Segmentation reduces the elasticity of bases with respect to demand by reducing risk sharing. Third, in some markets, intermediaries face concentrated demand from certain types of counterparties, and require compensation for the associated risk. Our results highlight the presence and importance of segmentation and search frictions in even the largest and most liquid markets.

We organize our empirical investigation with a stylized model where intermediaries satisfy customers' demand for dollars in exchange for foreign currency by engaging in basis

[1] Covered-interest parity arbitrage is a basis trade that compares the returns of two trades, one lending at domestic interest rates and the other lending at foreign interest rates while carefully hedging exchange rate risk. Without frictions, arbitrageurs should push the returns on the two to be identical, implying the bases are zero across all currency pairs.

[2] There are several types of basis trades beyond covered-interest parity, including the equity index futures/cash basis (Hazelkorn et al., 2022), the Treasury on-the-run/off-the-run spread (Krishnamurthy, 2002), the Treasury cash/futures basis (Barth and Kahn, 2021), the Treasury cash/swap basis (J Jermann, 2020; Boyarchenko et al., 2018b), the bond/CDS (Bai and Collin-Dufresne, 2019), and the CDX/CDS basis (Boyarchenko et al., 2018c).

trades. Intermediaries face several frictions in executing basis trades. These include costs to expand their balance sheets, search costs (it is costly to locate scarce foreign safe bonds), heterogeneous expertise in investing in risky assets, and counterparty limits.

The model reveals how each friction contributes to cross-sectional differences in bases and helps motivate empirical tests. Balance sheet costs drive a common component of bases across currencies. Search costs for scarce foreign safe bonds make CIP arbitrage imperfect since intermediaries hold risky bonds, rather than safe bonds, in their trades. These frictions lead to differences in bases across currencies based on the amount of synthetic dollar borrowing demand from those currencies. Heterogeneous expertise in investing in risky assets leads intermediaries to specialize in the markets where they hold the highest expertise, and makes segmentation a driver of the elasticity of the basis due to impeded risk sharing. Counterparty limits mean that the composition of demand also contributes to the inelasticity of bases. While prior work primarily focuses on frictions akin to leverage constraints, we examine and present evidence on the importance of the other frictions.

With the model predictions in hand, we turn to our empirical analysis. Our main data source is the Federal Reserve's FR2052a *Complex Institution Liquidity Monitoring Report*, which provides granular, high-frequency data on the assets and liabilities of the largest banks in the US. The data cover $25 trillion of daily notional exposure on average. We show that banks lend the most dollars in the same markets where the basis indicates dollar funding is the most expensive. We argue that this fact is consistent with the banking sector facing increasing marginal costs to meet dollar demand from each currency. Guided by our model, we study how the outlined frictions contribute to these increasing marginal costs.

First, we find foreign safe asset scarcity is an important driver of CIP bases. To do the CIP basis arbitrage, an intermediary must hold the equivalent of $1 of maturity-matched foreign safe assets for every $1 lent; otherwise, they cannot earn the foreign risk-free rate. In practice, banks hold $0.05 per dollar lent. Even when employing a generous definition of maturity-matched safe assets that ignores counterparty risk and permits small maturity mismatches, banks hold only $0.48 of foreign "safe" assets per $1 lent. Rather than holding maturity-matched foreign safe assets, intermediaries hedge their currency exposure partially with maturity-mismatched safe assets and partially with risky assets. The ability to execute only imperfect CIP arbitrage naturally explains our result that dollar funding is most expensive in markets where banks are doing the most dollar lending because those are the markets

where intermediaries are taking the most risk.

We find that a one standard deviation change in dollar borrowing demand increases the magnitude of the basis by 4 to 9 bps without accounting for systematic cross-currency differences in banks' abilities to access maturity-matched foreign safe assets. Moreover, there are differences in the cost of locating maturity-matched safe assets across currencies, as evidenced by cross-currency variation in the number of dollars of foreign safe assets held per dollar of swap exposure, which we call the *safe asset ratio*. In a regression, we find that a one standard deviation difference in the safe asset ratio corresponds to a 7 to 9 bps larger basis, further highlighting the importance of safe asset scarcity.

Second, we find that currency market segmentation contributes to CIP bases. We define a *market* as a tenor by currency pair, so the 1-month EURUSD, 1-year EURUSD, and 1-year JPYUSD are all distinct markets. For each market, we calculate a Herfindahl-Hirschman Index (HHI), which measures segmentation in that market. Analyzing the cross-sectional relationship between bases and segmentation, we find a strong relationship between more concentrated markets and larger basis dislocations: a one standard deviation more segmented market has a 10 to 14 bps larger basis. Interpreted alongside our safe asset scarcity result, the relationship between segmentation and the magnitude of bases highlights the importance of risk—and intermediaries' (in)ability to share it—as drivers of CIP deviations.

We can also observe segmentation by the extent to which bank-specific constraints are reflected in bases. We use the March 2023 banking turmoil as a natural experiment to test the relationship between bank-specific shocks and the basis. There was a notable shift of deposits toward the largest US banks following the turmoil. We show that currency markets intermediated by banks with comparatively large deposit inflows during this period had smaller basis dislocations. The result is consistent with the model's prediction about supply segmentation where specializing banks' constraints transmit to the markets in which those banks specialize.

The presence and importance of segmentation are surprising in our setting—global currency markets are among the largest and most liquid markets, without obvious barriers to entry. But we provide evidence that segmentation is a real feature of the data. Market segmentation is persistent. More segmented currency markets remain so over time, with intermediaries holding persistent market shares in a particular market. In our model, we ascribe segmentation to heterogeneous expertise in executing arbitrage across markets, with

banks specializing in markets where they have the highest expertise. We provide support for this mechanism by showing that banks tend to specialize in counterparty segments. For example, one bank might specialize in Canadian insurance counterparties, while another may cater more to Asian sovereign wealth funds. Even more, we show that banks tend to hold more loans in the currencies in which they have large currency market shares in FX swap markets. The evidence indicates that specializing banks likely have market-specific expertise and a more readily available set of counterparties in the markets they specialize in.

Third, we find that some currency markets have concentrated demand. In addition to foreign safe asset scarcity and segmentation, markets with a less diverse mix of counterparties have larger basis dislocations. The composition of demand affects bases, with more concentrated demand corresponding with larger bases, consistent with banks managing counterparty risk. We find that one standard deviation more concentrated market has a 7 bps larger basis.

Our analysis of the cross-section of CIP deviations reveals several novel and important frictions at play. In particular, the evidence indicates that foreign safe asset scarcity and segmentation are key features of global currency markets, which we interpret through the lens of heterogeneous intermediaries facing search costs to locate foreign safe bonds across markets.

**Related Literature** Our work is most closely related to work on CIP deviations (Du et al. (2018), Iida et al. (2018), Cenedese et al. (2021), Wallen (2022), Du and Schreger (2022), Augustin et al. (2022)) and bank-intermediated arbitrage spreads (e.g., Garleanu and Pedersen (2011), Pasquariello (2014), Boyarchenko et al. (2018a), Andersen et al. (2019), Anderson et al. (2021), Foley-Fisher et al. (2020)). Prior work primarily focuses on increased bank borrowing costs that give rise to CIP deviations following the 2008 financial crisis, for example, due to bank regulation (Du et al. (2018)) or debt overhang frictions associated with the expansion of bank balance sheets (Andersen et al. (2019)).

In contrast, our work shines a light on the asset side of intermediaries' basis trades. It indicates that the difficulty of locating foreign safe assets for use in CIP trades makes CIP arbitrage risky, with intermediaries resorting to holding riskier securities in the foreign legs of their basis trades. The result complements the findings of Diamond and Van Tassel (2021), who suggest that convenience yields on foreign safe assets may help explain CIP bases. This result is also related to Du et al. (2023a), who find that the security holdings of the banking

and insurance sectors in the Euro-area far exceed the amount of government debt, with institutions tilting their portfolios towards risky corporate debt. Moreover, given the real costs of CIP deviations (Du and Huber (2023)), our results that safe asset scarcity helps drive CIP deviations also contribute to the literature on the real effects of safe asset scarcity (e.g., Caballero (2006), Caballero et al. (2017), Caballero and Farhi (2018)).

Our paper is unique in its focus on intermediation frictions that give rise to cross-sectional variation in bases.[3] This focus leads us to conclude that intermediary heterogeneity (Kargar (2021)) and the accompanying segmentation of intermediaries into different markets are important drivers of bases. Using supervisory regulatory data, we provide direct evidence of segmentation's impact in transmitting idiosyncratic, bank-specific constraints into asset prices. Moreover, relative to previous work indicating segmentation in basis arbitrage trading (e.g., Rime et al. (2022); Siriwardane et al. (2022); Kloks et al. (2023)), because we find that CIP arbitrage is risky, our results highlight another channel through which segmentation affects bases—reduced risk sharing. Our empirical evidence sheds light on the source of segmentation. It is consistent with intermediaries having market-specific expertise that lowers their costs and risks of operating within the markets they specialize in, adding evidence that investors specialize in "complex asset markets" (e.g., Eisfeldt et al. (2023), Bryzgalova et al. (2023)). Our results particularly align with the perspective of Eisfeldt et al. (2023), emphasizing that the riskiness of arbitrage—and differences in intermediaries' abilities to reduce that risk—are key to understanding segmented arbitrage.

Lastly, our work contributes to the literature on intermediary asset pricing (Brunnermeier and Pedersen (2009), He and Krishnamurthy (2013), Adrian et al. (2014), He et al. (2017), and Du et al. (2023b)). The literature emphasizes that, due to households' limited participation in asset markets, the intermediary sector's marginal utility is a state variable that determines the level of risk premia. Our work indicates that limited participation is present even within the intermediary sector and suggests that the marginal utility of the specializing intermediaries determines a market's risk premium.

[3]An ingredient required of our explanation is that there are differences in dollar demand via currency forwards, which, for example, may arise from differences in currency hedging demand across currencies (e.g., Liao and Zhang (2021) and Du et al. (2023b)). Similar heterogeneity in demand also exists in other markets, for example, equity index futures markets (Hazelkorn et al. (2022)).

# 2 Model

We present a stylized model to organize our empirical investigation. The model yields a set of predictions on the cross-sectional drivers of CIP bases and illustrates how different frictions give rise to different features of bases.

## 2.1 Model Setup

There are $N_k$ foreign currencies (against the USD), indexed by $k = \{1, \ldots, N_k\}$. For simplicity, each currency faces a unitary exchange rate versus the USD. There are two types of investors: $N_i$ financial intermediaries, indexed by $i = \{1, 2, \ldots, N_i\}$ and $N_c$ customers, indexed by $c = \{1, 2, \ldots, N_c\}$. There are two periods, $t = 1, 2$. All investors invest in period 1 and realize payoffs in period 2.

The U.S. offers safe bonds in perfectly elastic supply, with expected return normalized to zero. Each foreign currency features three types of assets: one-period currency forwards, safe bonds, and risky bonds. All foreign bonds also have zero expected returns. Currency forwards are in zero net supply, and their price is endogenously determined at a level that also corresponds with the level of the basis.

**Customer Demand for Currency Forwards**. Customers only transact in currency forwards. In period 1, customer $c$ exogenously demands to synthetically swap $X_{c,k}$ dollars from currency $k$ into USD via forwards. In aggregate, synthetic dollar demand to swap currency $k$ for dollars is given by $X_k = \sum_{c=1}^{N_c} X_{c,k}$. For simplicity, we assume that customers' trades are netted out, such that $\text{sign}(X_k) = \text{sign}(X_{c,k}), \forall c, k$.

**Foreign Bonds and Safe Asset Scarcity**. There is a sufficient supply of safe bonds to hedge all currency trades, but each intermediary $i$ faces a total search cost of $\frac{1}{2}\lambda_{s,k} s_{i,k}^2$ to locate safe bonds in currency $k$, where $s_{i,k}$ is the safe bond position of intermediary $i$ in currency $k$ and $\lambda_{s,k}$ is a coefficient that captures how quickly search costs increase in currency $k$. That is, safe bonds in currency $k$ become increasingly difficult to locate as demand for them increases.

Intermediaries may also take positions in risky bonds of each currency, in perfectly elastic supply. Intermediary $i$ faces an idiosyncratic payoff variance of $\sigma_{i,k}^2$ when purchasing risky

bonds in currency $k$. Intermediaries with lower $\sigma^2_{i,k}$ for a given market face less risk in their basis trades when substituting away from safe bonds in the cash legs of their basis trades. This can be interpreted as intermediaries having market-specific expertise in substituting away from safe assets. As in Eisfeldt et al. (2023), the heterogeneity in risk faced by different intermediaries may reflect, for example, differences in technologies across intermediaries in producing information about issuers in a given market, or alternatively differences in trade execution.

**Intermediary Hedging and Safe Asset Choice**. Each intermediary $i$ maximizes a mean-variance objective function, $\mathbb{E}(W_{i,2}) - \frac{\gamma_i}{2}\mathbb{V}(W_{i,2})$, where $W_{i,2}$ is the terminal wealth of intermediary $i$, $\mathbb{E}(\cdot)$ and $\mathbb{V}(\cdot)$ are the expectations and variance operators, and $\gamma_i$ is a coefficient that captures the risk-bearing capacity of intermediary $i$. Intermediaries take the other side of customer demand in synthetic funding markets. Intermediary $i$ takes a position of $Z_{i,k}$ in currency forward $k$. Forward market clearing is given by $\sum_i Z_{i,k} + \sum_c X_{c,k} = 0, \forall k$.

Intermediaries are constrained to fully hedge their currency exposure from meeting forward demand via cash bonds. To satisfy their first-order conditions, intermediaries' allocations to safe and risky bonds in currency $k$ must make them indifferent to obtaining marginal hedge positions in either. This means that given their total position of $-Z_{i,k}$ in cash bonds of currency $k$, intermediary $i$ allocates a proportion $\alpha_{i,k} \equiv \frac{\lambda_{s,k}}{\gamma_i \sigma^2_{i,k} + \lambda_{s,k}}$ to risky bonds and allocates the remainder of their hedge positions to the safe bond.

Each intermediary also faces a set of constraints that we detail below.[4]

**Balance Sheet Cost**: Intermediary $i$ pays increasing marginal costs to expand their balance sheet (e.g., because of regulation or debt overhang). This is captured by each intermediary facing a cost of the form

$$\frac{1}{2}\lambda_{BS}\left(\sum_k |Z_{i,k}|\right)^2. \tag{1}$$

**Counterparty Constraints**: Intermediary $i$ pays increasing marginal costs to meet demand from specific counterparties; all else equal, intermediaries prefer to equalize positions across

[4]Conceptually, our results only depend on the fact that constraints are (weakly) convex. For example, explicit counterparty limits are convex constraints that would yield the same qualitative predictions as the quadratic counterparty costs that we assume.

different counterparties (e.g., for diversification purposes).

Here, we assume that for each currency $k$ and customer $c$, intermediaries' counterparty position is directly proportional to their holding in currency $k$, i.e.,

$$Z_{i,k,c} = Z_{i,k}\frac{X_{c,k}}{X_k} \tag{2}$$

where $Z_{i,k,c}$ is defined as intermediary $i$'s position in currency $k$ opposite customer $c$. For each counterparty $c$, intermediary $i$ faces a cost of the form

$$\frac{1}{2}\lambda_{CP}\left(\sum_k |Z_{i,k,c}|\right)^2. \tag{3}$$

**Fixed Participation Costs**: Each intermediary pays a fixed participation cost of $\lambda_{PC,k}$ to trade in currency $k$. This can be thought of, for example, as the cost of setting up a trading desk for currency $k$.

## 2.2 Model Predictions

We can write financial intermediary $i$'s problem as

$$\begin{aligned}
\max_{Z_{i,k},\alpha_{i,k}\forall k} & \sum_k -\text{Basis}_k Z_{i,k} \\
& -\sum_k \frac{\gamma_i}{2}(\alpha_{i,k}Z_{i,k})^2\sigma_{i,k}^2 && \text{(Risk)} \\
& -\sum_k \frac{1}{2}\lambda_{s,k}\left((1-\alpha_{i,k})Z_{i,k}\right)^2 && \text{(Safe Asset Scarcity)} \\
& -\sum_k \lambda_{PC,k}\mathbb{1}_{\{Z_{i,k}\neq 0\}} && \text{(Fixed Participation Costs)} \\
& -\frac{1}{2}\lambda_{BS}\left(\sum_k |Z_{i,k}|\right)^2 && \text{(Balance Sheet Costs)} \\
& -\frac{1}{2}\lambda_{CP}\sum_c\left(\sum_k |Z_{i,k,c}|\right)^2. && \text{(Counterparty Costs)}
\end{aligned}$$

Our first set of predictions does not rely on intermediary heterogeneity, so for clarity of exposition, we (1) set fixed participation costs to zero; and (2) assume there is no heterogeneity

across intermediaries in risk-bearing capacity or in risk faced: ($\sigma_{i,k} = \sigma_k$ for some value $\sigma_k$ and $\gamma_i = \gamma$ for some value $\gamma$, $\forall i$ and $\forall k$).[5]

Taking the first order condition with respect to $Z_{i,k}$, and averaging across intermediaries, we get that

$$\begin{aligned}
\text{Basis}_k = & \\
& \frac{X_k}{N_i} \frac{\lambda_{s,k}\gamma\sigma_k^2}{\lambda_{s,k} + \gamma\sigma_k^2} && \text{(Risk \& Safe Asset Scarcity)} \\
& + \lambda_{BS} \frac{\text{Sign}(X_k)}{N_i} \sum_{k'} |X_{k'}| && \text{(Balance Sheet Costs)} \\
& + \lambda_{CP} \frac{\text{Sign}(X_k)}{N_i} \left( \underbrace{|X_k| \sum_c \frac{X_{c,k}^2}{X_k^2}}_{\text{Demand Concentration in } k} + \sum_c \frac{X_{c,k}}{X_k} \sum_{k' \neq k} |X_{c,k'}| \right). && \text{(Counterparty Costs)}
\end{aligned}$$

This expression for the basis immediately yields a number of predictions.

**Prediction 1** (Sign of the Basis)**.** *The sign of synthetic demand for dollars from currency $k$ determines the sign of the basis.*

**Prediction 2** (Balance Sheet Costs)**.** *The magnitude of the basis is increasing in the total balance sheet usage of intermediaries across basis trades, $\sum_{k'} |X_{k'}|$.*[6]

Predictions 1 and 2 are predictions of standard explanations for CIP deviations. The basis for each currency contains a common balance sheet cost component that reflects the marginal cost that the intermediation sector faces in expanding its balance sheet. The post-GFC increase in the magnitude of CIP bases reflects the increase in the magnitude of balance sheet costs (e.g., an increase in $\lambda_{BS}$).

**Prediction 3** (Safe Asset Scarcity, Bases, and Net Lending)**.** *The magnitude of the basis for currency $k$ is increasing in synthetic demand to borrow dollars from currency $k$.*

[5] In the appendix, we present an expression for the basis without this simplification, which results in a more complicated expression that yields the same predictions as presented here.

[6] In principle, this reflects the balance sheet usage of the financial system.

**Corollary 1** (Heterogeneity in Safe Asset Scarcity)**.** *All else equal, currencies where foreign safe assets are harder to find (search costs are higher; $\lambda_{s,k}$ is high) have larger magnitude bases than currencies where foreign safe assets are easier to find.*

Prediction 3 is a novel, cross-sectional prediction that arises from the risk and safe asset scarcity terms in the basis. The basis embeds the search cost that intermediaries pay to locate foreign safe bonds. The marginal search cost (and distaste for risk) is increasing in the amount of demand banks must intermediate. Hence, the basis in a market increases in the amount of demand that intermediaries face in the market.

Corollary 1 extends this prediction to consider the possibility that search costs for foreign safe assets may systematically vary across currencies. For currencies that have systematically higher search costs, we expect bases to be larger per unit of demand, reflecting these costs and the additional risky substitutions that intermediaries must make.

**Prediction 4** (Demand Concentration)**.** *The magnitude of the basis for currency $k$ is increasing in the concentration of demand to swap foreign currency for dollars across counterparties $c$.*

Prediction 4 reflects the fact that intermediaries are sensitive when demand is concentrated from particular counterparties and demand larger compensation for concentrated demand.

For our last predictions, we focus on heterogeneity across intermediaries. We re-introduce non-zero fixed participation costs, heterogeneous risk in substituting away from safe bonds, and heterogeneous risk-bearing capacity, and for clarity of exposition, set $\lambda_{BS} = \lambda_{CP} = 0$.

Given the fixed participation costs, as we derive in the appendix, intermediary $i$ participates in currency $k$ in equilibrium only when they face sufficiently low risk in substituting away from risk-free assets in currency $k$, i.e., $\sigma_{i,k}$ is sufficiently low.[7] The basis can be expressed as

$$\text{Basis}_k = \frac{X_k}{\frac{N_{p,k}}{\lambda_{s,k}} + \sum_i^{N_{p,k}} \frac{1}{\gamma_i \sigma_{i,k}^2}},$$

where $N_{p,k} \leq N_i$ is the number of participating intermediaries. Segmentation arises because some intermediaries cannot justify the fixed participation costs given the risk they face in

[7] In the appendix, we also show that intermediaries that have $\sigma_{i,k}$ that is too low will also avoid participation. Their capacity to execute the trade at low risk may drive basis profits to be sufficiently low as to not justify participation.

substituting away from currency $k$ safe assets.

**Prediction 5** (Supply Segmentation and Size of the Basis)**.** *The basis for currency $k$ is larger in magnitude when intermediary supply for currency $k$ is more segmented (smaller $N_{p,k}$).*

Prediction 5 arises from a combination of foreign safe asset scarcity, heterogeneity in the risk faced by intermediaries in different markets, and fixed participation costs. It is a distinctive prediction from other work on bases that suggests that funding markets may be segmented (e.g., Rime et al. (2022); Siriwardane et al. (2022)) and arises in our setting because CIP arbitrage is imperfect, and intermediaries take risk in their positions. With segmentation, intermediaries are less able to share risk and accordingly demand larger compensation for meeting customer forward demand.[8]

While we do not micro-found the sources of heterogeneity in our stylized model, we discuss these sources in our empirical analysis. We document that the intermediaries that hold large market shares in a particular currency market also, for example, tend to hold larger loan portfolios in that market. The evidence indicates that lower risk in substituting away from safe assets may be driven by informational advantages, for example, coming from greater familiarity with a given market and a more easily accessible set of counterparties.[9]

**Prediction 6** (Specializing Intermediaries' Risk-Bearing Capacity)**.** *The basis for currency $k$ reflects the risk-bearing capacity of intermediaries that specialize in currency $k$.*

Prediction 6 is straightforward. When intermediaries specialize in different markets, the risk-bearing capacity of the *specializing* intermediaries in a market is the relevant driver of bases in that market, not the risk-bearing capacity of the overall intermediary sector. Siriwardane et al. (2022) present evidence consistent with this prediction by examining the correlations of different arbitrage strategies across asset classes; we present direct evidence

[8]This prediction is related to an implication of Eisfeldt et al. (2023), that participation should be lower in markets that have higher Sharpe ratios. An alternative perspective is that segmentation may lead to larger bases via a market power channel, as argued by Wallen (2022) in the context of CIP. While we do not argue that market power is not important, we present empirical evidence consistent with market-specific expertise playing an important role.

[9]Bryzgalova et al. (2023) similarly argue that intermediaries tend to specialize in options markets that are their 'natural markets.' They emphasize low fixed costs of entry due to related business areas or economies of scale.

that this prediction holds even within an asset class, using direct data on intermediaries' participation across markets.

With our empirical predictions in hand, we next turn to discussing our data sources and our empirical analysis.

# 3 Data

We collect information on bank-specific FX positions from the FR 2052a *Complex Institution Liquidity Monitoring Report* and use Bloomberg for exchange rates, prices, and interest rates.

## 3.1 FR 2052a Complex Institution Liquidity Monitoring Report

The Federal Reserve collects granular data on banks' liquidity in the FR 2052a as part of its capital adequacy framework as required by the Dodd-Frank Act and implemented by the Federal Reserve's regulation YY. The data are confidential and not publicly available. The data provide information by asset class, outstanding balance, and purpose, each reported by maturity and date. The data cover large U.S. and foreign banks. Global systemically important banks (G-SIBs), category II, and category III banks with a weighted average short-term funding of $75 billion must file the report each business day. Smaller banks report data monthly.[10] Banks have a strong incentive to report data fully and truthfully because of the possible consequences of making misrepresentations to government authorities, which could result in enforcement actions. Unlike other public companies, regulators closely scrutinize banks, so any misreporting would likely be identified and corrected quickly. Researchers have recently begun using this dataset to study bank behavior: for example, Infante and Saravay (2020), Cooperman et al. (2023), and—most related to our focus—Correa et al. (2020).

We use the data on banks' foreign exchange swaps and forwards.[11] Our sample includes daily observations from January 2016 through March 2023, spanning over-the-counter (OTC) and centrally-cleared transactions. Roughly 85 percent of the gross notional positions in

[10]For details on regulation YY and FR 2052, see `https://www.federalreserve.gov/supervisionreg/reglisting.htm`. The reporting instructions for FR 2052a are available at `https://www.federalreserve.gov/apps/reportingforms/Report/Index/FR_2052a`. Additional details on which large financial institutions must report daily are at `https://www.federalreserve.gov/supervisionreg/large-institution-supervision.htm`.

[11]The data on forwards includes forwards and futures. We will refer to them as forwards for brevity.

our sample are settled bilaterally on a value-weighted basis, while the rest clear centrally. Firms report FX transactions for eight currencies: AUD, CAD, CHF, EUR, GBP, JPY, USD, and other. The data cover cash-settled transactions settled with the physical exchange of currency, so it does not include contracts for difference or other non-deliverable transactions. The swaps in our sample include both FX forward swaps and cross-currency swaps, where the latter involve periodic interest payments in addition to the exchange of notional currencies at the beginning and end of the transaction. The data report maturities at daily increments up to 60 days, weekly increments from 61 days to 90, monthly increments to 180 days, 6-month increments to 1 year, and yearly increments beyond that. We discuss additional data cleaning in Appendix A.2.

We focus on the subset of the data with transactions in the seven currencies against the dollar. The sample averages $25 trillion in gross notional daily across foreign exchange swaps and forwards. We plot the daily sample average by currency in Figure 1. The sample is large and represents a material slice of the foreign exchange derivative market. While not an apples-to-apples comparison, Bank of International Settlements (2022) estimated the total notional amount of OTC foreign exchange derivatives at $110 trillion in 2022. Euro contracts are the largest ($9 trillion), and Swiss Franc contracts are the smallest ($1 trillion). Generally speaking, the tenors with the largest notional amounts are at the weekly increments—7, 14, 21, 28 days—and steadily grow in total beyond beginning with 6-month tenors. We limit our sample to maturities with at most 4-year maturity since the 5-year bucket contains all maturities at five years and beyond.

We also collect data on banks' safe assets. We focus on unencumbered assets (assets that face no restriction on use as collateral), assets pledged to central banks against which the bank could borrow, unrestricted central bank reserve balances, unsettled asset purchases, and encumbered assets (assets restricted from use as collateral). Encumbered assets are available only from mid-2022. Across these categories, we measure banks' safe assets as the subset of the level 1 high-quality liquid assets (HQLAs). HQLAs are the securities broadly considered the closest proxy for risk-free securities, so long as banks hold them to maturity. While the data do not provide the asset CUSIPs, the data provides collateral categories. In some instances, we include a broader set of assets beyond just level 1 HQLA, which we describe later.

## 3.2 Covered-Interest Parity Violations

We calculate covered-interest parity violations using interest rates, spot exchange rates, forward points, and forward maturity dates from Bloomberg. We use OIS interest rates across the curve from 1 week to 4 years. We provide details on cleaning the OIS interest rates in the online appendix.[12] We calculate CIP violations following Du et al. (2018). Define $s_t$ as the spot exchange rate in units of foreign currency per US dollar available at date $t$, $y^{\$}_{t,t+n}$ as the dollar interest rate available on date $t$ and maturing at $t+n$, and $f_{t,t+n}$ as the $n$-period outright forward exchange rate in foreign currency per USD. Then, the CIP basis is:

$$\text{Basis}_{t,t+n} = y^{\$}_{t,t+n} - \left( y_{t,t+n} - \frac{1}{n}(f_{t,t+n} - s_t) \right). \tag{4}$$

When the basis is negative, $\text{Basis}_{t,t+n} < 0$, dollar arbitrageurs can profit by borrowing at USD interest rates, simultaneously converting their USD to foreign currency at $s_t$, buying a forward $f_{t,t+n}$ to exchange that foreign currency back into dollars at maturity, and investing abroad at the foreign interest rate. Intuitively, an investor should be indifferent between holding USD to earn $y^{\$}_{t,t+n}$ and exchanging their USD for foreign currency, investing at the foreign risk-free rate, and converting the foreign currency back to USD by buying a forward.

Figure 2 plots our estimates of CIP bases at the 1-week and 1-year tenor. Dislocations are apparent during the 2008 financial crisis and the early stages of the Covid pandemic. We also provide the average and standard deviation of the 1-year CIP bases in Table 1. The basis averages $-24$ bps across all currencies, although it ranges from 6 for AUD to $-50$ for JPY. The first and second moments of the bases do not change much if we restrict the sample to begin in 2016, the period of our bank data sample.

We merge the Bloomberg basis panel with the FR 2052a panel using the days to maturity. For contracts with at least a month of maturity, we merge the panels using the commonly reported days to maturity in the FR 2052a data.[13]

Table 2 provides the daily average and standard deviation by currency and by tenor after

[12] The specific tenors we use are 1w, 2w, 3w, 1m, 2m, 3m, 4m, 5m, 6m, 1y, 2y, 3y, and 4y. We exclude the 5y tenor because the longest maturity category in the FR2052a data is five years and greater, so its average maturity is likely much more than five years.

[13] Specifically, 1m is 28 days; 2m is 61 days; 3m is 83 days if the days to maturity are between 83 and 90, inclusive, and 91 days if the days to maturity are greater than 90 days but less than 120 days; 4m is 121 days; 5m is 151 days; 6m is 181 days; 9m is 271 days; 1 year is 366 days; 2 years is 731 days; 3y is 1096 days; 4y is 1461 days.

merging with our estimates of CIP deviations using Bloomberg data—which limits the sample to the tenors of the OIS contracts. The daily average gross notional of the merged sample is $10.7 trillion. Forwards are larger at maturities less than six months, while swaps are larger beyond that. Shorter tenors also have more volatility in volumes. For example, the 1-week swap averages $91 billion with a standard deviation of $78 billion.

# 4 Empirical Strategy and Results

We begin by showing and discussing cross-sectional variation in currencies' bases. After constructing a measure of dollar lending against different currencies, we show that banks lend the most dollars in the currency and maturity markets with the largest CIP deviations. We interpret this result as showing that banks face increasing marginal costs to meet currency- and maturity-specific demand for dollar borrowing, consistent with our model.

We use our model to guide us in understanding the role of safe asset scarcity, market segmentation, and counterparty concentration in contributing to intermediaries' marginal costs. Finally, we provide additional evidence on the impact of segmentation through an event study surrounding the March 2023 turmoil to show how shocks trace through net dollar lending and its effect on the bases.

## 4.1 Signed Time-series Variation of Bases

Figure 3 plots the cross-sectional standard deviation across currencies on a given day. The variance across bases increases in tenor. The cross-sectional dispersion varies over time in a non-obvious pattern outside the financial crisis and Covid panic. Although an aggregate intermediary leverage constraint is important, the fact that we observe bases with different magnitudes and signs is *prima facie* evidence that forces beyond an aggregate intermediary constraint are major drivers of the basis cross-section.

We turn to our model predictions to help explore these other drivers. Prediction 1 notes that the sign of demand for dollars from a currency determines the sign of the basis. The summary statistics in Table 1 support the prediction. The AUD basis is positive while the others are negative. The rank ordering is consistent across tenors: AUD is typically the largest, CAD the second largest, and JPY the smallest. Prediction 1 explains this rank order:

AUD and CAD have the least dollar demand and JPY the most. That different foreign economies have different dollar demands and are willing to pay different prices for dollars, is intuitive. For example, countries with large commodity exports invoiced in dollars, like Australia and Canada, will have different demand for dollars than countries without similar commodity-related dollar inflows.[14]

Prediction 2 states the size of the basis—in absolute value—is decreasing in the banking system's balance sheet capacity across all basis trades. The size of the basis reflects dislocations that prevent the banking system from pushing the basis back toward zero. Such a prediction is consistent with work that studies the important role the aggregate intermediary sector plays in basis dynamics, for example Du et al. (2023b).

Predictions 1 and 2 reflect (signed) time-series variation in bases, the primary focus of earlier work. We calculate the first principal component of the bases which we use to proxy for an intermediary-wide factor. Figure A1 plots the variance across the bases in each tenor that can be explained by the first principal component. The figure shows that the common comovement across all the bases is important but leaves variation to be explained. The focus of our paper—cross-sectional variation in the size of bases—accounts for the remaining variation in bases.

## 4.2 Net Dollar Lending

Our remaining predictions each depend upon the intermediaries' dollar lending across currencies. Accordingly, we begin by constructing a measure of net dollar lending for each market. We discuss the construction and details of the variable below.

### 4.2.1 Construction

We use the FR 2052a data to calculate banks' net lending position for each market, where we define a market as a specific currency × tenor. We calculate banks' net USD supplied from date $t$ to $t+n$ for a given currency pair $k$ via FX swaps using:

$$Net^{k}_{t,t+n} = \frac{(\text{USD in at } t+n) - (\text{USD out at } t+n)}{(\text{USD in at } t+n) + (\text{USD out at } t+n)}. \tag{5}$$

[14]Du et al. (2018) also discuss dollar demand in terms of interest rates—AUD is the highest interest rate currency, and therefore expected to have the least carry trade demand for dollars, and vice-versa for JPY.

Table 3 provides a simple example to illustrate the logic underpinning $Net^k_{t,t+n}$ construction. Suppose the bank is buying and selling JPY swaps, with spot rate $S_t = 115$ and forward exchange rate $F_{t,t+7} = 110$. In the first swap, the bank lends \$100 dollars at the near leg and receives $\$100 \times S_t = ¥11,500$. Separately, and simultaneously, the bank receives \$95 in a second swap and pays $\$95 \times S_t = ¥10,925$. The bank has paid \$5 more than it received, equivalent to lending \$5. At maturity, the two swaps unwind at the forward price. The net variable is the ratio of the net dollars lent to the notional dollars: $2.6\% = 5/195 = 5.23/204$. Notice that the net variable is the same regardless of whether it's based on the near or far leg flows.

When $Net^k_{t,t+n} > 0$, the bank lends out more dollars today than it borrows against currency $k$ with maturity $t+n$. This is because the bank will receive more dollars in at maturity on $t+n$ than it pays out. We divide by the notional dollar flows because some markets are larger than others.[15]

We can calculate two flavors of our net lending measure by aggregating across different levels. $Net^k_{t,t+n}$is our primary measure, and it reflects the entire intermediary sector because it aggregates lending across all banks at the date × maturity × currency level. We also calculate a bank-specific measure (indicated by the $i$ superscript), $Net^{k,i}_{t,t+n}$, that aggregates at the date × maturity × currency × bank level. Since banks report only non-zero values, we set $Net^k_{t,t+n} = 0$ when there is no lending or borrowing data for a given observation.

### 4.2.2 *Net* Summary Statistics

Table 4 gives the average and standard deviation of $Net^k_{t,t+n}$by currency and tenor.[16] With few exceptions, $Net^k_{t,t+n}$is small and near zero, indicating that the intermediary sector broadly matches its dollars in and out. The table highlights three facts: first, the intermediary sector tends to borrow at shorter tenors and lend at longer tenors, with the average flipping

[15]Although net lending is possible through swaps and a combination of spot transaction with forwards, we calculate net using only swaps since there is no upfront exchange of principal for forwards and futures, and we are unable to connect forward transactions with spot transactions forming basis trades. We expect our results to be unaffected if forward and future demand is directionally similar to swap demand. We find little evidence that including forwards and futures meaningfully affects our results, for example, in our analysis of safe asset ratios in Appendix Table A4. Similarly, our net lending measure excludes intermediate interest payments since we are unable to match those payments to the underlying swap.

[16]Appendix Table A1 shows the level of $Net^k_{t,t+n}$in billions of dollars. Average $Net^k_{t,t+n}$is largest in level terms for JPY at \$5.4b, and AUD is the smallest with $Net^k_{t,t+n}$at −\$1.3b.

from negative (reflecting banks are borrowing) to positive (indicating banks are lending) around 3 months. Second, there is considerable variation in net dollar lending across the currencies—averaging across all tenors, intermediaries tend to borrow in AUD (on average $Net^{AUD}_{t,t+n}$ of 3.6 percent) and lend EUR (1.1 percent). Third, net lending is more volatile at shorter maturities than it is at longer maturities. At less than a month, the average time-series standard deviation is 30 percent compared to less than 10 percent for maturities beyond 5 months.

Appendix Table A2 shows $Net^{k}_{t,t+n}$'s correlations with several variables, and shows generally intuitive correlations with several variables. $Net^{k}_{t,t+n}$is increasing in maturity, size, and an analogous measure of net lending provided using the coarser data from the Traders in Financial Futures Report from the Commodity Futures Trading Commission, as used in Hazelkorn et al. (2022). It is lower at quarter- and month-ends. Our measure is also positively related to the market's convenience yield—estimated using a Nelson-Siegel-Svensson yield curve model. Finally, $Net^{k}_{t,t+n}$is weakly pro-cyclical, given that it is larger when the VIX and the Baa-Aaa spread are lower. There is no obvious relationship between $Net^{k}_{t,t+n}$and the SPX return.

As an example, we plot the net variable for 1-week, 6-month, and 1-year EUR in Figure A2. At shorter maturities, net lending is normally negative but often turns positive for brief periods. The middle panel shows the level of net lending in billions of dollars. Average net lending at the 1-week tenor is −\$1.2 billion but grows to \$12 billion at the 1-year tenor. The bottom panel shows the gross notional dollar flows—the denominator of our net variables. There is an increasing trend across all tenors, but the trend is most obvious at the 1-year maturity, where notional approached \$800 billion in 2023. The recurring spikes in the shorter maturities reflect window dressing.

Figure A3 shows a histogram of $Net^{k}_{t,t+n}$by currency across all tenors. The peaks near the zero net lending line indicate that the banking system generally runs a matched book, lending as much as it borrows. The figure makes clear that we expect $Net^{k}_{t,t+n}$to be near zero or tightly bounded around zero, rather than large directional positions, either long or short.[17]

[17]In the online appendix, Figure A4 shows that net is closer to zero for the largest markets.

## 4.3 *Net* vs. Bases

With the $Net^k_{t,t+n}$variable in hand, we test Prediction 3, and how cross-sectional variation in $Net^k_{t,t+n}$captures cross-sectional variation in bases, with banks lending more in markets where obtaining dollar funding is most expensive. We run the regression

$$\text{Basis}^k_{t,t+n} = \alpha + \beta Net^k_{t,t+n} + \gamma\prime X_t + \varepsilon^k_{t,t+n} \tag{6}$$

where a negative basis indicates that swapping foreign currency for dollars is expensive, and $X_t$ is a vector of controls. Our model predicts that $\beta < 0$—bases are the most negative in the markets with the most dollar demand. The relationship arises from the search costs and risks that intermediaries face when supplying dollars in different markets, due to the scarcity of foreign safe assets.[18] The alternative—$\beta > 0$ would mean that banks are lending USD exactly in the markets where there is the least USD shortage.

Table 5 shows the regression results. $\beta < 0$ across specifications with different controls. The first row shows a robust negative relationship between bases and $Net^k_{t,t+n}$ after including tenor and time fixed effects and weighting by the square root of the market's share of the total daily gross notional. We use weighted regression because some markets are magnitudes larger than others. Column (4) is the benchmark estimate which includes the full set of fixed effects and weights by notional share. The coefficient shows that when $Net^k_{t,t+n}$ is 1 percentage point (pp) larger then the basis is 0.4 bps smaller. A one-standard-deviation change in $Net^k_{t,t+n}$weighted by its daily notional share is about 10pp, corresponding to a basis that is 4.3 bps lower using column 4's coefficient. Columns (5) and (6) split the sample into short- and long-term tenors, with a threshold of one year. The relationship is much stronger for longer-tenor lending, with a coefficient roughly 9 times larger than short-dated tenors. A one-standard-deviation-change in $Net^k_{t,t+n}$for these longer tenors corresponds to a basis that is 9 bps lower ($5.9 \times -1.5$).

In the online appendix, we provide additional results on the relationship between bases and lending. We show this two ways: first, Figure A5 scatterplots the average basis against the average $Net^k_{t,t+n}$ for a given currency and tenor, and Figure A6 shows the average basis

[18] $\beta < 0$ is implicitly assumed in several other papers, for example, Greenwood et al. (2023), Liao and Zhang (2021) and Du and Huber (2023). An analogous result is also shown to be the case in equity index futures markets by Hazelkorn et al. (2022).

and $Net^k_{t,t+n}$for each currency at several tenors. Second, we show the regression coefficients in Figure A7. The regression coefficient on $Net^k_{t,t+n}$ is near zero for maturities less than 3-months but is significant and negative for all longer maturities.

The cross-sectional relationship between $Net^k_{t,t+n}$and bases is consistent with the interpretation provided by our model: Intermediaries' (shadow) costs that increase with dollar funding demand from a particular currency drive the cross-sectional variation in bases. The remaining predictions of our model are based on decomposing the importance of different constraints in contributing to these costs.

## 4.4 Frictions Contributing to the Basis

We identify three potential sources that influence the size of the basis: foreign safe asset scarcity, intermediary segmentation, and demand concentration. We construct proxies for each measure, which we use together in a regression to jointly test their ability to explain the cross-section of bases. We describe these variables' construction and features in turn.

### 4.4.1 Foreign Safe Asset Scarcity

Prediction 3 links bases with dollar funding demand because it is difficult for intermediaries to find foreign safe assets. Intermediaries can only imperfectly execute CIP arbitrage that requires \$1 of foreign safe asset for every net dollar lent. The trade is not a risk-free arbitrage if the arbitrageur cannot find a perfectly maturity-matched security, denominated in that foreign currency, that pays the risk-free rate.

We supply evidence that intermediaries do not appear to be executing perfectly hedged CIP arbitrages. We find that on average, intermediaries hold only 5 cents of maturity-matched foreign safe assets per dollar of net lending. Banks put the rest of the cash leg of their CIP trades toward maturity-mismatched safe assets and foreign currency risky assets, showing that intermediaries take on substantial risk to meet customer demand in currency swap markets.

To illustrate this point, we define foreign safe assets as the sum of level 1 HQLAs that are unencumbered assets, unsettled asset purchases, and assets pledged to the central bank. We include a full list of types of level 1 HQLA securities in the online appendix section A.2. Sovereign bonds make up most of these securities. The BIS describes HQLAs as assets that

"can be easily and immediately converted into cash at little or no loss of value."[19] The measure is specific to the market: there are daily observations for each currency and tenor. We define

$$\text{Safe Asset Ratio}^k_{t,t+n} = \frac{\text{Level 1 HQLAs (\$)}^k_{t,t+n}}{\text{Net (\$)}^k_{t,t+n}}.$$

If $\text{Safe Asset Ratio}^k_{t,t+n} = 1$, then banks perfectly match every dollar of net dollar lending for currency $k$ on date $t$ with maturity $t+n$ to a dollar of foreign safe asset in the same currency and the same tenor. The definition of CIP arbitrage requires perfect matching. Safe asset ratios less than one indicate that intermediaries' swap and cash positions do not offset one another and that the intermediary holds an imperfect CIP position involving some risk.

To better understand where intermediaries may place their foreign currency, we calculate a broader definition

$$\text{Broad Asset Ratio}^k_{t,t+n} = \frac{\text{Broad Assets (\$)}^k_{t,t+n}}{\text{Net (\$)}^k_{t,t+n}}$$

where $\text{Broad Asset Ratio}^k_{t,t+n}$ covers assets beyond HQLAs in the three categories listed above. It includes unencumbered assets, unsettled asset purchases, capacity, central bank deposits, reverse repurchases, or securities borrowings. For example, a German government bond would be in both ratios, but a German corporate bond would show up only in the broad asset ratio since it is not a level 1 HQLA. Beginning in 2022, the data includes encumbered assets, and we treat this subsample separately as a point of comparison given the large value of the banking system's encumbered assets. Note that these measures are gross long positions, so they are an upper bound on the banking system's net position in these markets. If a bank were short the security, their net long measure would be lower.[20]

We separately address cases when banks net lend dollars from cases when banks net

[19] See `https://www.bis.org/basel_framework/chapter/LCR/30.htm`.

[20] The broad asset ratio measure may include double counting when the bank pre-positions collateral it receives from a secured financing transaction, like a reverse repo, with a central bank. For example, if a bank lent AUD in a 1-month repo and pre-positioned the AUD-denominated collateral (say with a 2-year maturity) it received with a central bank, then our broad ratios for 1-month and 2-year AUD would reflect this exposure. The data does not allow us to distinguish if the pre-positioned securities are received through a secured financing transaction. As a robustness, we calculate the broad asset ratios excluding pre-positioned securities. The correlation between the broad asset ratios including and excluding capacity is 0.997, so the effect of any double counting is small.

borrow dollars. When $Net^k_{t,t+n} > 0$, banks lend dollars, receive foreign currency, and therefore have demand for foreign safe assets. When $Net^k_{t,t+n} < 0$, banks lend foreign currency, receive dollars, and thus need USD-denominated safe assets. We separate these two cases in the following analysis since we have strong priors that USD safe asset holdings are different for U.S.-based G-SIBs.

Figure 4 summarizes our estimates of safe and broad asset ratios, which illustrate the presence of foreign safe asset scarcity. This analysis focuses particularly on the case where $Net^k_{t,t+n} > 0$, where intermediaries have demand for foreign safe assets. The top panel reports the average median Safe Asset Ratio across currencies, which is 0.05. Intermediaries hold 5 cents of unencumbered, maturity-matched safe assets per dollar of net lending. The panel plots the ratios constructed using two alternative methods. First, plotted in orange, we include forwards in the estimate of FX net lending, since a spot transaction paired with a forward transaction can be economically similar to an FX swap. The data doesn't allow us to pair these two separate transactions together, so including forwards will likely overestimate the amount of net lending. Second, plotted in light blue, we calculate asset ratios net of the banks' short positions in those securities. Both methods yield nearly identical results.

The definition of maturity-matching used in the construction of the safe asset ratio is restrictive because it requires the maturity of safe assets to exactly match the maturity of net swap lending. For example, a 6-day maturity safe asset does not count as hedging 7-day maturity swap lending. We relax the restriction by rounding the tenor of safe assets and FX lending to the nearest benchmark tenor for which we have estimated CIP violations, and we round tenors less than 7 days to the 1-week bucket.[21] For example, we round a 2-day swap to the 1-week bucket, a 10-day swap to the 1-week bucket, and a 9-month swap to the 1-year bucket. The resulting safe asset ratios allow small deviations in intermediaries' cash positions to count towards hedging their swap market positioning.

Figure 4 reports the rounded tenor results for safe asset ratios in the top panel. The average median safe asset ratio increases to 0.13, indicating that small maturity mismatches can account for an additional 8 cents of safe asset positioning, but is still markedly below 1.

[21] Recall, the tenors in the confidential supervisory data are daily increments up to 60 days, weekly increments from 61 days to 90, monthly increments to 180 days, 6-month increments to 1-year, and yearly increments beyond that. The benchmark tenors for which we estimate CIP violations are 1w (7d), 2w (14d), 3w (21d), 1m (28d), 2m (61d), 3m (91d), 4m (121d), 5m (151d), 6m (181d), 1y (366d), 2y (731d), 3y (1096d), and 4y (1461d).

We further relax the definition of safe assets to include encumbered assets—which cannot be used as collateral, entail counterparty risk, and are not counted as level 1 HQLAs—but which may help intermediaries hedge the currency exposure coming from their net lending. Figure 4 illustrates that this more permissive definition of the safe asset ratio has an average median value of 0.13 across currencies when requiring strict maturity matching, and the value jumps to 0.48 with maturity rounding. Intermediaries hold only 48 cents of foreign safe assets per dollar lent, even under a generous definition of safe assets that overlooks counterparty risk and some maturity mismatch. What do banks do with the remaining unaccounted for 52 cents?

We turn to broad asset ratios, which include riskier foreign assets in addition to safe assets in the numerator of the ratio calculation. The bottom panel of Figure 4 plots the same values as the top panel, except for broad asset ratios. Focusing on unencumbered assets, the average median broad asset ratio requiring exact maturity matching for broad assets and net lending is 0.11, indicating that intermediaries hold only maturity-matched assets against approximately 11 cents per dollar of net lending they perform. Using the rounded tenor approach, the broad asset ratio jumps to 0.59, indicating that we can account for 59 cents per dollar of net lending by relaxing the perfect maturity matching requirement. Finally, the broad asset ratios jump to over 1 after we further add encumbered assets. Although intermediaries hedge foreign currency exposures, a large part of their hedging relies on maturity-mismatched, encumbered, non-safe assets.

The analysis of safe and broad asset ratios indicates that intermediaries' currency swap transactions entail significant risk. Intermediaries, unable to take positions in foreign safe assets that CIP arbitrage requires, end up taking positions in risky foreign assets. Interpreted through the lens of our model, imperfect hedging shows why $Net^k_{t,t+n}$variable—capturing intermediaries' net dollar lending—helps explain cross-sectional variation in bases.

Figure 4 plots safe and broad asset ratios medians across markets. There is substantial variation in safe asset ratios across currencies and maturities, as we present in more detail in the appendix. Heterogeneity is relevant for understanding cross-sectional variation in bases as well. Corollary 1 shows that differences in safe asset ratios—stemming from differences in intermediaries' ability to source foreign safe bonds in foreign currency—amplifies cross-sectional variation in bases, even after we control for differences in net dollar lending across currencies. This is because differences in safe asset ratios across currencies indicate different

riskiness associated with CIP trades. Lower safe asset ratios imply riskier CIP trades.

Appendix Table A3 reports details on safe and broad asset ratios across currencies and shows the median safe asset ratio across currencies ranges from \$0.00 (CHF) to \$0.16 (EUR). For our broadest measures—broad assets including both unencumbered and encumbered assets—the ratios range between 0.02 (CHF) and 0.80 (EUR). The bottom of the table also shows the corresponding measures for instances in which the banks are net borrowing dollars—which means they need to hold USD-denominated safe assets. This number is consistently much higher than 1, ranging from 2 to 29, consistent with our prediction that U.S.-based G-SIBs hold substantially more dollar-denominated safe assets. Figure A8 breaks out the median ratio by currency and tenor for unencumbered safe and broad assets, and Figure A9 plots the median ratio when using the rounded tenor method. Across the maturity dimension, the takeaway is that safe asset ratios are slightly higher at longer maturities than shorter maturities, partially because the width of the maturity buckets is larger for the longer maturities.

As robustness, Appendix Table A4 calculates the ratios using two different methods. First, we include forwards in the estimate of FX net lending, since a spot transaction paired with a forward transaction can be economically similar to an FX swap. The data doesn't allow us to pair these two separate transactions together, so including forwards will likely overestimate the amount of net lending. Second, we calculate asset ratios net of the banks' short positions in those securities.

Our analysis of safe asset ratios estimates an upper bound for the safe assets that intermediaries hold corresponding with their swap exposure, because intermediaries may hold safe assets for reasons unrelated to meeting customer demand and hedging CIP trades. We explore this dimension in a regression framework presented in Appendix Table A5, which estimates the extent to which intermediaries may hold foreign currency assets even if their net lending exposure is zero. For example, Column 1 of the table shows that banks hold an average of \$1.8 billion of unencumbered EUR-denominated safe assets across the tenors, after controlling for $Net^{k}_{t,t+n}$.[22]

Our discussion of foreign safe asset scarcity primarily focuses on cross-sectional variation.

[22]Moreover, the currency fixed effects in the regression are lower bounds since we merge safe asset holdings to the net lending data based on the net borrowing currency. Observations where the bank holds safe assets in a currency but is not net borrowing that currency are not included in the panel.

However, we also note that foreign safe asset scarcity may be important for understanding the dramatic post-2008 increase in the size of bases as well. There has been a steep drop in the foreign safe assets available to intermediaries, because of decreases in the supply of safe assets (e.g., see Caballero et al. (2017)), and decreases in their re-use as collateral in transactions. Systematic data on collateral re-use are not readily available before the 2008 financial crisis, but to the extent they are available, the evidence indicates a reduction. Gorton et al. (2020) use data from the 10Qs of six broker-dealers and banks to show that collateral pledged halved between 2007 and 2009, amounting to a decline of more than $2.5 trillion. Along these lines, the existing evidence indicates there has been a large reduction in collateral velocity—the amount of re-use of the same safe asset in multiple transactions. For example, per the numbers reported in Jank et al. (2022), there has been an approximate 30% reduction in collateral re-use of European sovereign bonds from 2008 to 2017. As Inhoffen and van Lelyveld (2023) note, the decrease in collateral velocity connects to higher balance sheet costs after the crisis, attributed to the balance sheet-intensive nature of repos.

### 4.4.2 Supply Segmentation

Prediction 5 suggests that variation in intermediaries' expertise in substituting away from safe assets in a market leads to market segmentation. Given the riskiness of CIP arbitrage, the basis for a given currency is more inelastic when intermediary supply is more segmented, since there is less risk-sharing.

We show that banks do specialize in markets and provide evidence consistent with heterogeneous expertise. Without frictions or market segmentation, banks' net exposures across tenors and currencies should be in equal proportion to the size of their FX books. In this case, banks face the same marginal search costs and risks across markets when lending dollars, with their net lending being highly correlated. But, if markets are segmented, a bank lending more in its own market faces increasing marginal costs and risk, pushing the basis increasingly negative—consistent with Prediction 5.

We empirically estimate dollar supply concentration by calculating an HHI of an individual bank's notional exposure in each market using

$$\text{Supply HHI}^{k}_{t,t+n} = \sum_{i \in \text{bank}} (\text{Market Share}^{k,i}_{t,t+n})^2$$

where

$$\text{Market Share}_{t,t+n}^{k,i} = \frac{\text{Bank } i\text{'s USD In} + \text{Bank } i\text{'s USD Out}}{\text{Industry USD In} + \text{Industry USD Out}}.$$

Our HHI measure is not directly comparable to HHI measures in other settings because our sample is limited to nine banks and therefore the lowest possible value is $9 \times [(1/9) \times 100]^2 \approx 1,111$.

Figure 5 illustrates cross-sectional variation in HHI, plotting the average measure by market against that market's size. Lower HHI indicates more competition and less concentration, implying less segmentation. Larger markets are clearly less segmented. Short-term contracts for CAD and AUD are more segmented, and CHF is the most segmented. The least segmented markets are JPY and EUR at longer tenors, especially at 1-year.[23]

Why FX markets are segmented is an important question. Our model ascribes segmentation to heterogeneous expertise. We provide support for this mechanism using cross-sectional HHI variation in Section 5 after we discuss our regression results capturing the relationship between various frictions and bases. There, we also discuss another implication of segmentation: that shocks to the specialists in a market should particularly impact *that* market.

### 4.4.3 Demand Concentration

Demand concentration also affects the basis. Prediction 4 shows that the basis is increasing in the demand concentration. High demand concentration is equivalent to high counterparty risk. A bank can manage its counterparty risk by lending to a wide set of counterparties in each market. Lending becomes riskier when a bank deals with fewer counterparties, increasing its counterparty risk. Intermediaries should require compensation for this risk.

We empirically proxy for demand concentration with a separate HHI measuring counterparty concentration. The data does not provide firm-specific counterparty names but instead provides two dozen counterparty types—e.g., broker-dealer, non-financial corporate, non-regulated fund.[24] Since we do not view individual counterparties, we calculate demand

[23] We estimate similar supply HHI's looking just across tenors and currencies in the online appendix, see Figures A11 and A11.

[24] The full list of FX counterparties: Bank, Broker Dealer, Central Bank, Debt Issuing Special Purpose Entity, Financial Market Utility, Government Sponsored Entity, Investment Company or Advisor, Multilateral Development Bank, Non-Financial Corporate, Non-regulated Fund, Other, Other Supervised Non-Bank

concentration assuming that firms within a category have highly correlated risks. Such an assumption is plausible: lending exclusively to levered funds is riskier than lending to a mix of levered funds and non-financial companies given the two lines of business are exposed to different risks.

We calculate the demand HHI similarly to the supply HHI:

$$\text{Demand HHI}^{k}_{t,t+n} = \sum_{j \in (\text{c.p. type})} (\text{Counterparty Market Share}^{k,j}_{t,t+n})^2$$

where

$$\text{Counterparty Market Share}^{k,j}_{t,t+n} = \frac{\text{Counterparty } j\text{'s USD In} + \text{Counterparty } j\text{'s USD Out}}{\text{Industry USD In} + \text{Industry USD Out}}.$$

Figure 6 plots demand concentration by market. There is no clear pattern between the market size and its demand concentration, although currencies tend to bunch together—longer maturity CAD are among the most demand-concentrated. The range of demand HHIs is much higher, partly reflecting banks' large share as counterparties to one another.

Counterparty data are available only beginning in May 2022. We estimate the full sample demand concentration using two different items available during the full sample: loan counterparties and settlement types. The data for unsecured and secured loans includes counterparty types, and we calculate loan demand concentration HHI analogously. Banks also report FX settlement types: bilateral, continuous linked settlement (CLS), or other. CLS transactions are settled payment-versus-payment to reduce Herstatt risk and cover 30 percent of all FX transactions.[25] Bilateral trades are often, but not exclusively, transactions that banks do with their clients, like hedge funds. Hedge funds trade through their prime brokers using *give up trades* which are typically bilateral. The bilateral share is a rough measure for the relative importance of investors like hedge funds in each market. For each market, we calculate that day's share of bilateral transactions compared to the total transactions in that market. We estimate Demand $\text{HHI}^{k}_{t,t+n}$ for the full sample with

$$\text{Demand HHI}^{k}_{t,t+n} = \alpha + \beta_1 \text{Loan Demand HHI}^{k}_{t,t+n} + \beta_2 \text{Bilateral Share}^{k}_{t,t+n} + \varepsilon^{k}_{t,t+n}.$$

Financial Entity, Other Supranational, Pension Fund, Public Sector Entity, Retail, Small Business, Sovereign.

[25] See `https://www.newyorkfed.org/medialibrary/microsites/fxc/files/2020/FX_settlement_risk_CLS.pdf`.

We show the estimated coefficients in the online appendix, Table A6. We use the regression coefficients to estimate $\widehat{\text{Demand HHI}}^{k}_{t,t+n}$ for the full sample. Both measures of demand concentration—directly measured using the shorter sample with counterparty information, or the longer sample that uses settlement type and loan counterparties—likely have considerably more error than the supply concentration measure.

## 4.5 Decomposing Deviations from CIP

Now we compare how each of the frictions contributes to cross-sectional variation in bases. We run the following regression:

$$
\begin{aligned}
|\text{Basis}^{k}_{t,t+n}| = \alpha & \\
&+ \beta_1 \left(\text{Demand HHI}^{k}_{t,t+n} \times \mathbb{I}(Net^{k}_{t,t+n} \geq 0)\right) \\
&+ \beta_2 \left(\text{Demand HHI}^{k}_{t,t+n} \times \mathbb{I}(Net^{k}_{t,t+n} < 0)\right) \\
&+ \beta_3 \left(\text{Supply HHI}^{k}_{t,t+n} \times \mathbb{I}(Net^{k}_{t,t+n} \geq 0)\right) \\
&+ \beta_4 \left(\text{Supply HHI}^{k}_{t,t+n} \times \mathbb{I}(Net^{k}_{t,t+n} < 0)\right) \\
&+ \beta_5 \left(\text{Safe Asset Ratio}^{k}_{t,t+n} \times \mathbb{I}(Net^{k}_{t,t+n} \geq 0)\right) \\
&+ \beta_6 \left(\text{Safe Asset Ratio}^{k}_{t,t+n} \times \mathbb{I}(Net^{k}_{t,t+n} < 0)\right).
\end{aligned}
$$

The dependent variable is the absolute value of the basis for a given market on a given day. A larger basis dislocation—either expensive or cheap dollar funding compared to benchmark rates—increases the absolute value of the basis. We use the absolute value of the basis because we expected increases in the magnitude of any of our frictions to push the basis away from zero. Since we expect there may be differences in the frictions when the bank is net lending dollars or net borrowing dollars, we include dummies to capture this asymmetry. We expect foreign safe asset scarcity to behave differently than USD safe asset scarcity since the banks are based in the U.S.

We include controls for the risk of a safe asset sovereign issuer, which we proxy using the country's CDS spread, as well as the value-weighted average CDS spread for the banks lending in that market. We merge government CDS spreads based on which safe asset an arbitrageur would hold, the foreign government CDS when $Net^{k}_{t,t+n} \geq 0$ and the U.S spread

when $Net^k_{t,t+n} < 0$. [26] We include tenor and date fixed effects, and we weight the regression by the square root of the market's share of the total daily gross notional. To make the coefficients directly comparable, we transform the independent variables to modified $z$-scores using each variable's full sample median and standard deviation. We use the sample median rather than mean to control for the high skewness of the data.

Table 6 reports our main regression results using the sample from January 2016 to March 2023. Due to data availability, the regressions use estimated demand concentration, rather than the directly measured demand concentration HHI, and we compute safe asset ratios including unencumbered assets with matched maturities. The first two columns in the table exclude estimated demand concentration, while the last two columns include it.

The first two rows show that supply concentration is an important friction and is symmetric across markets with long or short dollar demand. A one standard deviation increase in the supply HHI increases the absolute value of the basis by 10 to 14 bps. Such a result is consistent with our expectation that markets relying on a more concentrated set of banks will have larger dislocation. That is, the elasticity of bases is affected by the concentration of intermediaries that meet dollar funding demand in each market.

The next set of rows illustrates the importance of foreign safe asset scarcity and banks' search costs for foreign safe bonds. The coefficients on $Net^k_{t,t+n}$are positive and significant, with values ranging from around 4 to 7.5 bps, indicating that a one standard deviation increase in $Net^k_{t,t+n}$corresponds with a basis 4 to 7.5 bps higher basis, consistent with the results presented in Table 5. Interpreted through the lens of our model, these coefficients capture the effect of foreign safe asset scarcity if safe asset ratios are the same across markets and if differences across markets stem from the total amount of dollar demand.

However, there are also meaningful differences in safe asset ratios across markets. The next four rows account for the effect of variation in foreign safe asset scarcity across different markets, stemming, for example, from differences in search costs for foreign safe assets across markets. Focusing on markets where there is net dollar demand, we observe significant

[26] CDS spreads are from Markit. For both banks and sovereigns, we use CDS spreads for 5y tenor of senior unsecured tier for the primary curves and coupons, as identified by Markit. The euro CDS spread is a simple average of Italy and German CDS spreads (both quoted in USD). We use the MM14 contract for DB since Markit denotes both MM and MM14 as primary curves. We create a market-specific bank CDS spread by value-weighting the individual banks' CDS spreads based on their gross position in the market as a share of the total gross positions across all banks in that market.

negative coefficients of 8 to 11 bps in columns (1) and (2) on the safe asset ratio, indicating that a one standard deviation decrease in the safe asset ratio moves the basis by an additional 8 to 11 bps. When using the broad asset ratios in the regressions rather than safe asset ratios (in columns (2) and (4)), these coefficients increase in magnitude to 11 and 14.

The next two rows examine the effect of demand concentration on the basis. We observe that in markets where banks are net lending dollars, a one standard deviation increase in demand concentration coincides with an approximately 7 bps higher basis. This effect is limited to dollar net lending, and the coefficients are insignificant when looking at markets where banks are borrowing in dollars.

The regression evidence indicates the importance of safe asset scarcity and supply concentration in explaining cross-sectional variation in CIP deviations. Additionally, though demand concentration is measured with more error than the other variables, the regression suggests that demand concentration may play a role in amplifying basis dislocations.

We also present additional variations of our regressions in Appendix Table A7, where we present analysis including encumbered assets and using directly measured demand concentration from the shorter sample, and in Appendix Table A8 where we present regression results separately considering each independent variable. The results are similar to the results reported in the main regression, and suggest an even stronger relationship between the independent variables measures and bases.

The regression shows that both foreign safe asset scarcity and intermediary segmentation are important frictions that help describe the cross-section of bases. We interpret these frictions as arising from heterogeneous expertise across intermediaries. We next turn to further exploring intermediary segmentation.

# 5 Intermediary Segmentation

The earlier cross-sectional regressions are tests of equilibrium relationships that hold in the data. The analysis reveals that intermediary segmentation is a key driver of bases, making them more inelastic to demand.

We now dive deeper to understand intermediary segmentation better. First, we supply evidence that segmentation arises from heterogeneous expertise across intermediaries in substituting away from safe assets. Second, we use an event study analysis around the

Silicon Valley bank run to explore an additional prediction of segmentation (Prediction 6): specializing intermediaries' constraints should help explain cross-sectional variation in bases.

## 5.1 The Sources of Intermediary Segmentation

We present several facts consistent with intermediary segmentation being driven by banks having heterogeneous expertise in operating in different markets, with banks specializing in the markets where they are better able to manage risk. In particular, the evidence is consistent with banks having market-specific expertise stemming from other business areas, and more ready access to counterparties in the markets in which they specialize.

First, we show that markets with higher supply segmentation rely on banks with larger FX books. Table 7 column 1 regresses a market's supply HHI on (log of 1 plus) the FX swap notional of the banks active in that market, weighted by their market share. The independent variable in the regression estimates the FX book size of the average bank in that market. There is a strong relationship, showing that more segmented markets rely on larger intermediaries.[27]

Segmentation is persistent, with segmented markets remaining so over time. The first four columns of Table 7 report regression results from a market's supply HHI on its 1-month, 1-year, and 5-year lags. The evidence indicates a strong relationship. A 1-point increase in the supply HHI is associated with a 0.51-point increase in the supply HHI one year later. The relationship is stronger at shorter lags, but the coefficient is still large and significantly different from zero even with a 5-year lag.

Individual banks' market shares are also persistent. We regress bank $i$'s market share on lags of its market share. We calculate a bank's market share as its share of the total gross notional of swaps in that market. The last four columns of Table 7 show the regression results using a panel at the bank-tenor-currency-date level. These persistence coefficients are similarly large and different from zero, ranging from 0.87 at a 1-month lag to 0.74 at a 5-year lag.

These facts suggest that segmentation is real, with banks specializing in the same markets over time, consistent with having and developing market-specific expertise. We present further evidence supporting this story by showing that banks specialize not just in specific

[27] Figure A13 in the online appendix provides a scatterplot of these two variables.

markets—currency and tenor—but also in counterparty segments. One bank, for example, may have a large rolodex of Canadian insurance counterparties, while another may cater more to Asian sovereign funds. We show this by calculating a bank's market share of a given counterparty and currency, collapsing across all maturities. We calculate the market share as the notional FX swap exposures with that counterparty-currency pair as a share of the bank's total notional FX swaps on that day. We denote this Bank FX Share$^{i,k}_{t,ctpty}$ where $i$ denotes the bank and *ctpty* denotes the counterparty. For example, if bank $z$ had \$100 of gross notional swaps outstanding and had \$10 of gross notional with insurance companies denominated in AUD, we would set Bank FX Share$^{z,AUD}_{t,insurance} = 10\%$. We exclude bank counterparties since they are the largest counterparty by an order of magnitude in most markets, and recall that FX counterparty data is available beginning in May 2022.

We calculate two related measures: first, we calculate the average market share of all banks except bank $i$, which we call Other Bank FX Share$^{i,k}_{t,ctpty}$. Second, we estimate a bank's Bank Loan Share$^{i,k}_{t,ctpty}$ to study how segmentation works across a bank's lines of business. We define bank loans as the items reported by the bank in their inflows-secured and inflows-unsecured tables in the FR2052a data.[28]

Table 8 shows how banks specialize in specific counterparty-by-currency markets. The first column shows that the bank's specialization in counterparty-currency markets is persistent over time as the lag of Bank FX Share$^{i,k}_{t,ctpty}$ strongly predicts its current values. However, one concern is that banks all uniformly service the same counterparty-currency markets. Column two adds a variable, Other Bank FX Share$^{i,k}_{t,ctpty}$, and rejects such a possibility since the coefficient on the other bank variable is not different from 0. Column three shows that banks specialize in currencies and counterparties across their lines of business. It regresses the bank FX share on Bank Loan Share$^{i,k}_{t,ctpty}$ and finds a strong positive relationship. Adding controls for other banks' loan shares (column 4) does not change the result. Column 5 adds Other Bank FX Share$^{i,k}_{t,ctpty}$ along with the loan share variables, showing that the bank's loan book is still informative above and beyond what other banks' FX activities are in predicting a bank's FX activities. The last column, column 6, includes the bank's lagged FX share, which dominates the results. Note, however, that the bank's loan share variable is the only

[28] Values are reported on a gross basis and not netted. We exclude collateral swaps. These tables include offshore and onshore placements, operational balances, outstanding draws on unsecured or secured revolving facilities, other unsecured loans, short-term investments, reverse repos, securities borrowing, dollar rolls, margin loans, other secured loans, synthetic customer longs, and synthetic firm sourcing.

other variable with a positive coefficient, even if not statistically significant.

The evidence supports the notion that banks specialize in particular markets because they have expertise and easier access to counterparties in those markets, stemming, for example, from other parts of their business—that is, it is likely easier for them to manage risky assets in the markets they operate, and more generally, operating costs are lower for them.

## 5.2 Event Study: Silicon Valley Bank Run

In March 2023, Silicon Valley Bank suffered a bank run following disclosures about losses on its hold-to-maturity portfolio. Data from the publicly available H.8 shows that some depositors moved from smaller to larger banks, including the G-SIBs in our sample. Between March 8 and March 15, the H.8 data shows that large banks gained $120bn of deposits while small banks lost $108bn.[29] The average large bank had deposit inflows of $120/25 =$4.8 billion since H.8 data spans the largest 25 domestically charted commercial banks. Publicly available call report data shows that the average bank in our sample lost $11 billion in deposits between 2022 Q4 and 2023 Q1, ranging from −$44 billion to $32 billion, with a standard deviation of $21 billion. [30]

The unexpected flow of deposits to large banks is an exogenous increase of those banks' risk-bearing capacity, at least in the short term. We use this plausibly exogenous variation in banks' abilities to intermediate FX markets to test Prediction 6 of our model: that markets should reflect the risk-bearing capacity of *their* specializing banks. The particular source of variation exploited by the shock is the cross-sectional variation in deposit inflows corresponding with the shock.

We calculate the change in the dollar-denominated deposits between March 9, one day before the event, and two weeks after the event for each bank. We calculate the value-weighted change in deposits for each market (tenor and currency pairs) in our sample, which we define as $\Delta Deposits^k_{t+n}$. We calculate value-weights using a bank's gross notional in a given market relative to the total gross notional for the same market. We fix our value weights from March

[29] See `https://www.federalreserve.gov/releases/h8/20230324/`.

[30] We match the banks in our FX data with their main affiliated banks using the following RSSDs: JPM (852218), BAC (480228), WFC (451965), C (476810), GS (2182786), BK (541101), MS (1456501 and 2489805), SST (35301), and DB (214807). We calculate the change in deposits using RCON2200 (domestic deposits) and RCFN2200 (foreign deposits).

9 to exclude the effects of banks rebalancing their activity after the SVB shock. Hence, the variable has no $t$ subscript: it is fixed through the sample and varies only over tenor and currency. The method provides a market-specific deposit inflow. For example, if two banks each had half of the 1-year JPY gross notional on March 9, and bank 1 had deposit flows of $x$ and bank 2 had deposit flows of $y$, then the value-weighted deposit flow we assign to 1-year JPY would be $(x + y)/2$. We transform $\Delta Deposits^k_{t+n}$ into a modified $z$-score using the median rather than the mean. To give a sense of magnitudes, the average deposit inflow to a given market was \$20.3 billion, with a standard deviation of \$6.5 billion.

The shock is an inflow of dollars, so we expect markets where banks lend dollars will have an effect. Therefore, we limit the sample to the markets where $Net^k_{t,t+n}$was positive on average over the two weeks leading into the event. In Table 9, we run a regression:

$$|\text{Basis}^k_{t,t+n}| = \alpha + \gamma_1 \mathbb{I}(\text{Post}) + \gamma_2 \Delta Deposits^k_{t+n} + \gamma_3 \mathbb{I}(\text{Post}) \times \Delta Deposits^k_{t+n} + \varepsilon^k_{t,n}. \quad (7)$$

$\mathbb{I}$(Post) is equal to 1 for days after March 9 and 0 otherwise. Our window is the two weeks before and after the event, which gives us the longest time frame before the last week of March, when quarter-end window-dressing could confound the estimates.

The first three columns of Panel A show the main results. The columns vary the fixed effects and whether the regression is weighted by the markets' sizes. The coefficient on the post dummy is positive and significant, indicating that basis dislocations worsened following the event, consistent with a risk-off sentiment. The bases were larger on average in markets with larger deposit inflows. This is likely because deposit flows went toward banks perceived as the safest, and those same banks lend dollars most in markets with larger CIP dislocations.

The key result is the interaction row. After the stress, markets that relied more on banks with the largest inflows had smaller dislocations, as indicated by the significant and negative coefficient. The result is robust across specifications. The specification with all the controls, column 3, shows that a one standard deviation increase in deposit inflows decreases the absolute value of the basis by 2 or 3 bps after the event. The effect is economically large: the average basis absolute value is 30 bps, so the coefficients imply a 7 percent decline.

Why would the basis relatively improve in the treated markets? The last three columns of Panel A show that treated markets saw increased net dollar lending. The deposit inflows likely allowed banks to lend some marginal dollars. The panel runs the same event study

regression except changes the dependent variable to $Net^k_{t,t+n}$. The main result is the third row: markets with larger deposit inflows were the same in which banks increased $Net^k_{t,t+n}$. The result holds when the markets are weighted by their size, so small markets behaved differently. Using column 6, a one standard deviation increase in deposit flows after the event increased $Net^k_{t,t+n}$by 0.6 percentage points.

Panel B runs a placebo event study using data from one month before. The regression is exactly the same except sets the post dummy equal to 1 for days after February 9. There was no salient market volatility in mid-February 2023, so we do not expect the interaction term to have a significant effect. Moreover, the deposit flows captured by the variable had not yet occurred. The third row of Panel B shows that the placebo event had no apparent effect on the basis, and if anything, the basis increased during the period. The last three columns of the table show no obvious pattern in the interaction term's effect on net lending.

The evidence from the SVB bank run supports the implication of intermediary segmentation that markets reflect their specializing banks' risk-bearing capacity. Moreover, it supports the broader claim that segmentation is important for intermediaries' impact on asset prices.

# 6 Conclusion

CIP bases vary across currencies, a puzzle for prevailing theories of financial intermediation. We show that banks lend the most dollars in precisely the markets with the largest CIP violations. Why? We show that intermediaries execute CIP arbitrage imperfectly due to the scarcity of foreign safe assets available for the cash legs of their CIP trades. Foreign safe asset scarcity makes CIP trading risky. Moreover, intermediaries have heterogeneous expertise in conducting CIP trades in different markets, leading markets to reflect their specializing banks' constraints, as evidenced by the fallout following the Silicon Valley Bank run. The combination of foreign safe asset scarcity and risky CIP arbitrage indicates that segmentation makes bases more inelastic to demand, limiting intermediary risk sharing. Our results highlight the importance of market segmentation and search frictions in even the most liquid markets in the world.

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# 7 Figures

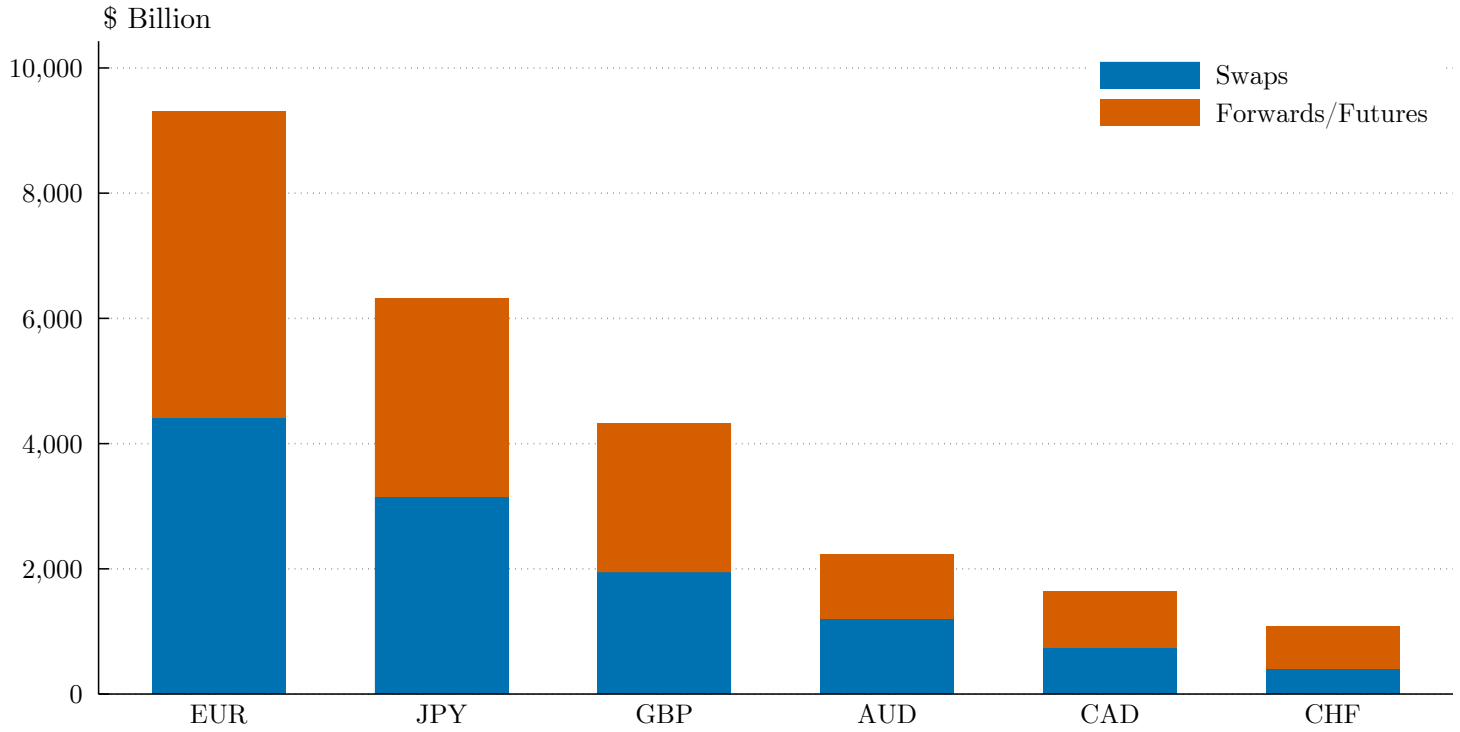



**Figure 1: Foreign Exchange Notional Exposure**. Figure provides the average notional FX exposures in our sample by currency across swaps and forwards/futures before limiting to the tenors with OIS rates. Sample includes only transactions in the six currencies versus the dollar.

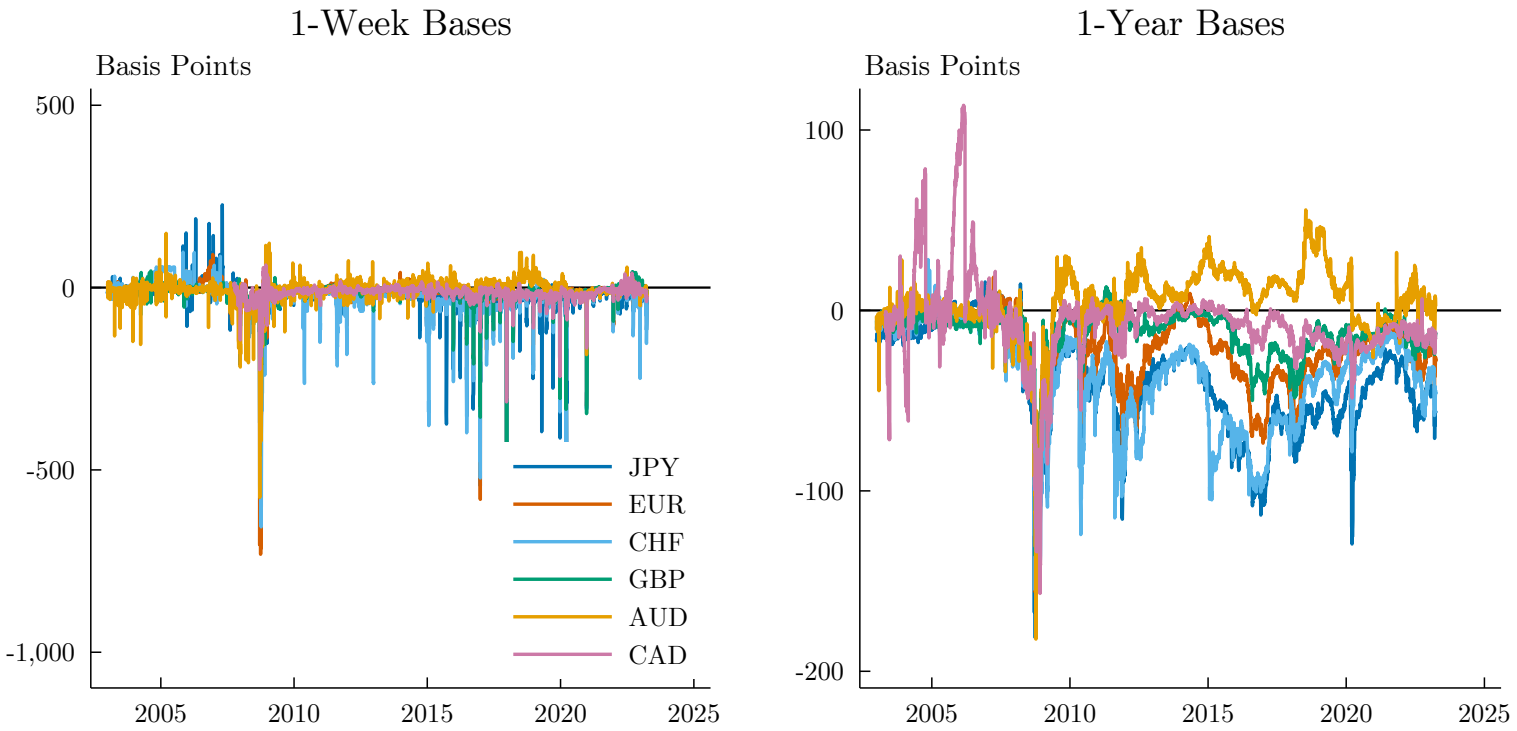



**Figure 2: Covered-Interest Parity Violations**. Figure plots the CIP bases across two tenors—1-week and 1-year—across several currencies.

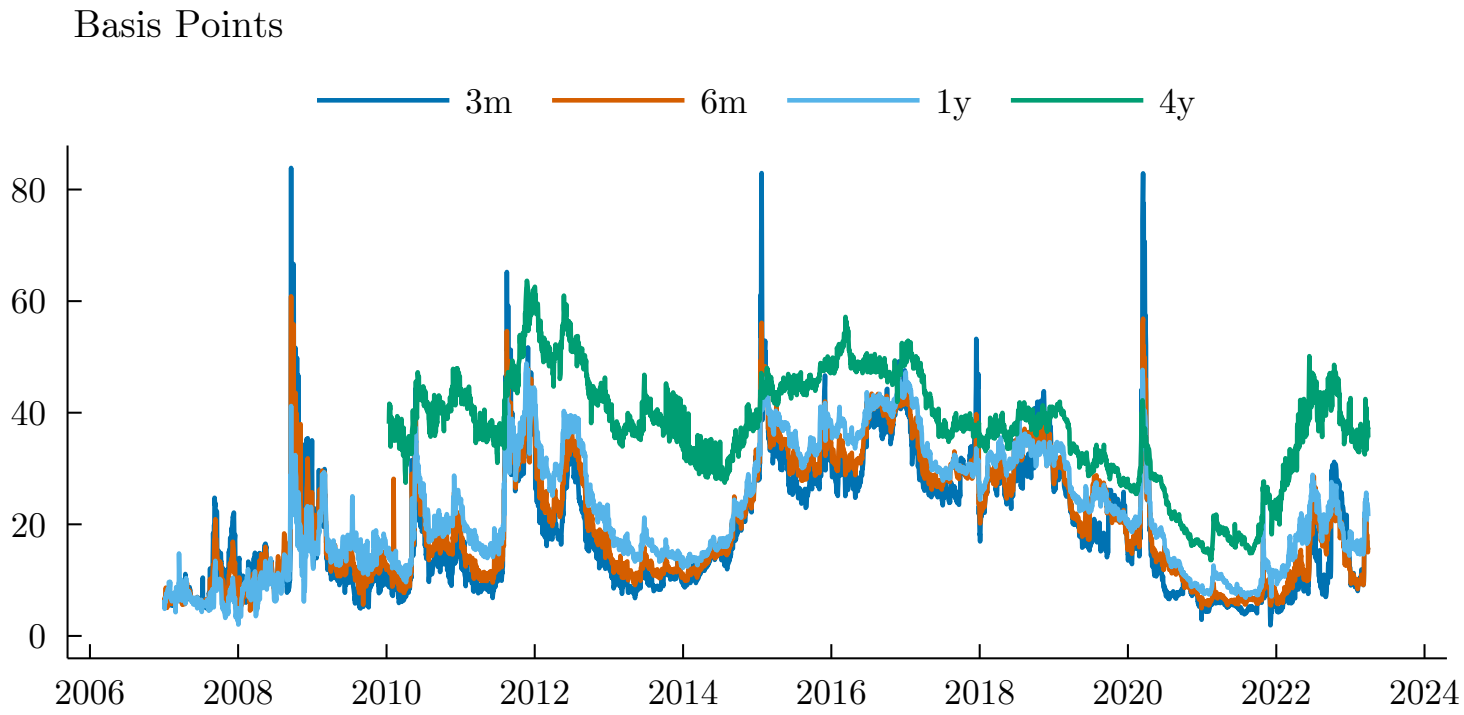



**Figure 3: Cross-Sectional Standard Deviation of Covered-Interest Parity Violations**. Figure plots the cross-sectional standard deviation of CIP bases for the given tenor. Sample includes only periods when we have observations for all of the six currencies against the dollar.




**Figure 4: Summary of Safe Asset Ratios** Figure plots the average median safe and broad asset ratio across all currencies when $Net^k_{t,t+n}$is positive excluding the USD. Matched tenor shows the asset ratio when the tenor of the underlying asset and the swap have the same maturity. Rounded tenor buckets swaps and the assets into the nearest benchmark tenor. Figure also shows the analogous net measurement when including forwards and futures (affecting the denominator of the ratios) and when netting out the firms' short positions in the asset (affecting the numerator).

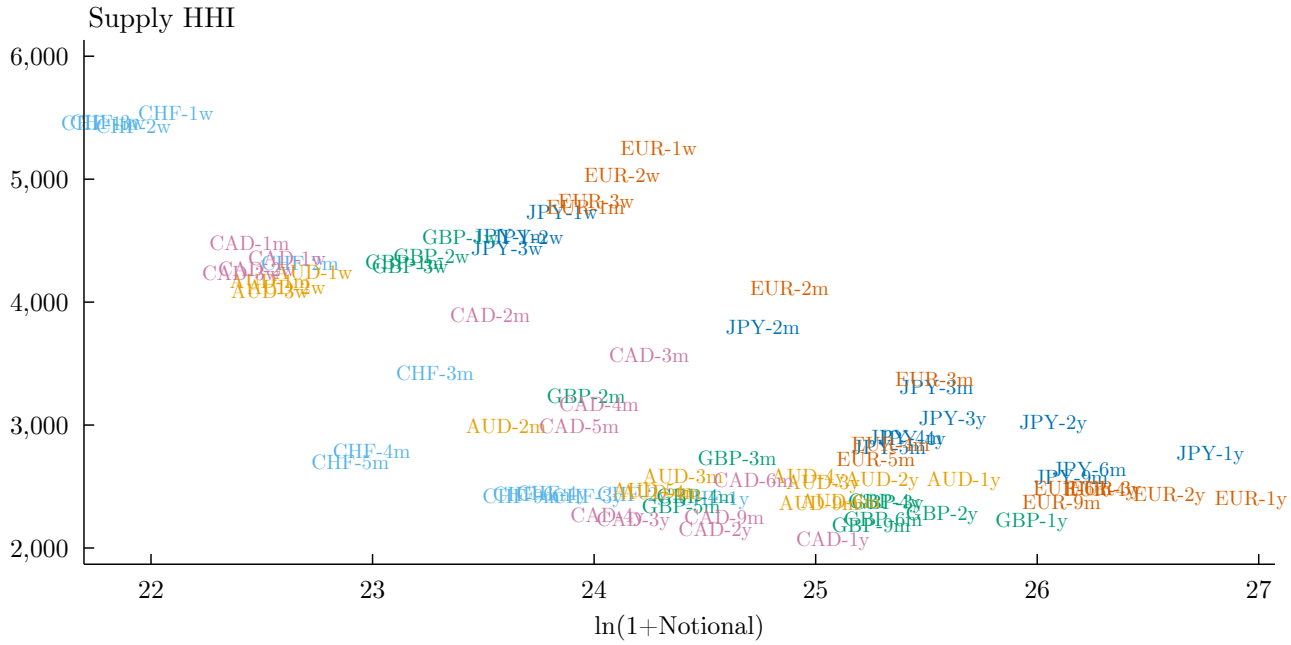



**Figure 5: Supply Segmentation.** Figure plots the average of daily Supply $\text{HHI}^k_{t,t+n}$ against the (log of 1 plus) the average notional of that market.

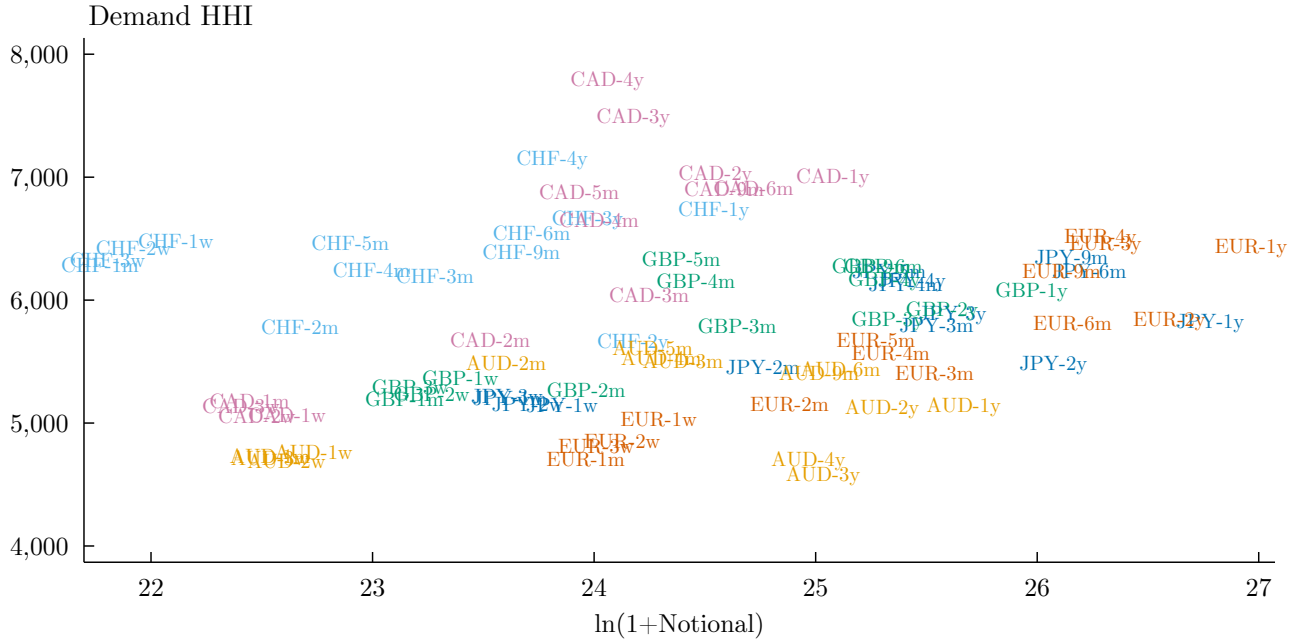



**Figure 6: Demand Segmentation.** Figure plots the average of daily Demand $\text{HHI}^k_{t,t+n}$ against the (log of 1 plus) the average notional of that market.

# 8 Tables

| *Summary Statistics, bps* | All | AUD | CAD | CHF | EUR | GBP | JPY |
|---|---|---|---|---|---|---|---|
| 2008–2023 | | | | | | | |
| Mean | −24.1 | 6.2 | −12.8 | −46.6 | −26.8 | −14.7 | −49.9 |
| Std. Dev. | 28.4 | 21.9 | 16.9 | 24.3 | 20.3 | 14.8 | 23.3 |
| 2016–2023 | | | | | | | |
| Mean | −26.3 | 9.6 | −14.3 | −45.6 | −30.9 | −19.4 | −57.5 |
| Std. Dev. | 27.0 | 14.8 | 6.6 | 22.9 | 15.1 | 9.9 | 20.5 |

**Table 1: 1-Year Covered-Interest Parity Violations**. Table shows mean and standard deviation of the CIP deviations at a one-year tenor. See section 3.2 for the calculation details.

| *by currency* | Swaps | | Forwards & Futures | |
|---|---|---|---|---|
| *$ billions* | Mean | Std. Dev. | Mean | Std. Dev. |
| AUD | 689 | 92 | 331 | 137 |
| CAD | 411 | 92 | 290 | 99 |
| CHF | 215 | 28 | 207 | 55 |
| EUR | 2,344 | 345 | 1,610 | 270 |
| GBP | 914 | 172 | 729 | 166 |
| JPY | 1,827 | 180 | 1,146 | 181 |

| *by tenor* | Swaps | | Forwards & Futures | |
|---|---|---|---|---|
| *$ billions* | Mean | Std. Dev. | Mean | Std. Dev. |
| 1w | 91 | 78 | 308 | 295 |
| 2w | 78 | 72 | 251 | 260 |
| 3w | 71 | 67 | 222 | 240 |
| 1m | 69 | 68 | 208 | 226 |
| 2m | 188 | 120 | 456 | 334 |
| 3m | 384 | 147 | 721 | 338 |
| 4m | 324 | 78 | 420 | 122 |
| 5m | 302 | 67 | 338 | 94 |
| 6m | 730 | 104 | 564 | 127 |
| 9m | 679 | 101 | 352 | 71 |
| 1y | 1,396 | 201 | 277 | 43 |
| 2y | 868 | 133 | 109 | 9 |
| 3y | 632 | 101 | 53 | 6 |
| 4y | 589 | 92 | 34 | 7 |

**Table 2: Average daily gross FX notional across all banks**. Table shows average daily notional across all banks and tenors for the matched covered-interest parity tenors by currency.

| *Bank's Cash Flows* | | $t$ | $t+7$ |
|---|---|---|---|
| **Swap #1: lend USD vs. JPY** | | | |
| | USD pay | −$100.00 | |
| | JPY receive | ¥11,500 | |
| | | | |
| | USD receive | | $104.55 |
| | JPY pay | | −¥11,500 |
| **Swap #2: lend JPY vs. USD** | | | |
| | USD receive | $95.00 | |
| | JPY pay | −¥10,925 | |
| | | | |
| | USD pay | | −$99.32 |
| | JPY receive | | ¥10,925 |
| **Total** | | | |
| $(a)$ | Net Dollars Lent | $5.00 | $5.23 |
| $(b)$ | Notional Dollars | $195.00 | $203.87 |
| $a/b$ | $Net_{t,t+7}^{JPY}$ | 2.6% | 2.6% |

**Table 3: Net Calculation Example**. Table shows the bank's cash flows across two swaps, one receiving dollars and the other paying dollars with spot exchange rate $S_t = 115$ and the forward exchange rate is $F_{t,t+7} = 110$.

| Mean (Percent) | | | | | | | |
|---|---|---|---|---|---|---|---|
| | AUD | CAD | CHF | EUR | GBP | JPY | Mean |
| 1w | −5.2 | −4.3 | −4.2 | −3.0 | −6.6 | −6.6 | −5.0 |
| 2w | −5.5 | −4.3 | −0.5 | −1.6 | −7.3 | −5.1 | −4.0 |
| 3w | −6.2 | −3.3 | −1.0 | −1.6 | −8.1 | −4.1 | −4.0 |
| 1m | −5.8 | −3.9 | −2.0 | −1.6 | −8.2 | −3.2 | −4.1 |
| 2m | −5.6 | −0.2 | −7.0 | 2.4 | −2.1 | −3.8 | −2.7 |
| 3m | −3.7 | 3.0 | −1.6 | 2.0 | −0.2 | 0.8 | 0.1 |
| 4m | −3.5 | 1.8 | −0.1 | 2.5 | 2.5 | 2.5 | 0.9 |
| 5m | −3.6 | 0.9 | 0.0 | 3.2 | 2.4 | 2.1 | 0.8 |
| 6m | −3.9 | −0.6 | −1.5 | 4.3 | 2.3 | 4.5 | 0.8 |
| 9m | −2.8 | −5.0 | −2.4 | 4.8 | 3.5 | 3.5 | 0.3 |
| 1y | −1.9 | −3.1 | −3.7 | 2.3 | 3.4 | 5.7 | 0.4 |
| 2y | 0.3 | −3.4 | −6.1 | 1.4 | 1.1 | 5.4 | −0.2 |
| 3y | −1.9 | −1.4 | −7.3 | 1.2 | 0.3 | 5.0 | −0.7 |
| 4y | −1.2 | −0.3 | −3.1 | −1.2 | 1.2 | 7.9 | 0.5 |
| Mean | −3.6 | −1.7 | −2.9 | 1.1 | −1.1 | 1.1 | |

| Standard Deviation (Percent) | | | | | | | |
|---|---|---|---|---|---|---|---|
| | AUD | CAD | CHF | EUR | GBP | JPY | Mean |
| 1w | 30.2 | 30.2 | 41.4 | 21.3 | 28.9 | 23.5 | 29.3 |
| 2w | 32.3 | 33.1 | 44.3 | 22.9 | 31.0 | 24.7 | 31.4 |
| 3w | 34.1 | 34.6 | 45.8 | 24.0 | 31.6 | 25.5 | 32.6 |
| 1m | 34.2 | 35.7 | 45.9 | 23.6 | 31.9 | 26.0 | 32.9 |
| 2m | 18.8 | 20.4 | 27.6 | 12.9 | 18.3 | 12.8 | 18.5 |
| 3m | 11.2 | 13.7 | 20.1 | 8.5 | 12.8 | 9.7 | 12.7 |
| 4m | 10.6 | 14.1 | 19.8 | 8.2 | 14.3 | 9.5 | 12.8 |
| 5m | 11.2 | 15.2 | 19.6 | 8.8 | 14.7 | 9.9 | 13.2 |
| 6m | 7.4 | 9.8 | 12.5 | 6.1 | 8.0 | 9.0 | 8.8 |
| 9m | 9.0 | 10.5 | 12.7 | 6.3 | 8.4 | 8.5 | 9.2 |
| 1y | 5.6 | 7.2 | 6.1 | 3.4 | 5.1 | 5.8 | 5.5 |
| 2y | 6.9 | 7.5 | 5.5 | 3.2 | 5.8 | 7.2 | 6.0 |
| 3y | 4.6 | 8.6 | 3.3 | 3.2 | 4.1 | 6.0 | 4.9 |
| 4y | 4.6 | 8.2 | 4.4 | 3.3 | 4.9 | 6.7 | 5.4 |
| Mean | 15.8 | 17.8 | 22.1 | 11.1 | 15.7 | 13.2 | |

**Table 4: *Net* Summary Statistics**. Top panel plots the average daily $Net^k_{t,t+n}$ for a given currency $k$ and maturity $t+n$. Bottom panel plots the time-series standard deviation of $Net^k_{t,t+n}$.

| | All Tenors | | | | Short-Term | Long-Term |
|---|---|---|---|---|---|---|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| $Net^k_{t,t+n}$ | $-0.0544^{*}$ | $-0.601^{***}$ | $-0.432^{***}$ | $-0.431^{***}$ | $-0.171^{**}$ | $-1.548^{***}$ |
| | $(-1.82)$ | $(-3.97)$ | $(-3.36)$ | $(-3.40)$ | $(-2.55)$ | $(-3.67)$ |
| $N$ | 151,655 | 149,337 | 149,337 | 149,337 | 106,005 | 43,332 |
| $R^2$ | 0.00 | 0.04 | 0.03 | 0.02 | 0.01 | 0.11 |
| Tenor FE | No | No | No | Yes | Yes | Yes |
| Time FE | No | No | Yes | Yes | Yes | Yes |
| Weighted | No | Yes | Yes | Yes | Yes | Yes |

**Table 5: Net and Bases**. Table presents the regression of the basis on $Net^k_{t,t+n}$: $\text{Basis}^k_{t,t+n} = \alpha + \beta Net^k_{t,t+n} + \gamma\prime X_t + \varepsilon^k_{t,t+n}$. Currencies include AUD, CAD, CHF, EUR, GBP, and JPY and tenors include: 1w, 2w, 3w, 1m, 2m, 3m, 4m, 5m, 6m, 9m, 1y, 2y, 3y, and 4y. Constant omitted. Columns with weights are weighted by the square root of the market's daily gross notional share. Short-term column limits swaps to less than 1-year maturities, and long-term is greater than or equal to 1-year maturities. $Net^k_{t,t+n}$is in percent and basis is in basis points. Within $R^2$ reported. $t$-statistics shown using robust standard errors clustered by market and date where $^{*}$ $p < 0.10$, $^{**}$ $p < 0.05$, $^{***}$ $p < 0.01$.

| | (1) | (2) | (3) | (4) |
|---|---|---|---|---|
| *Supply Segmentation* | | | | |
| Supply $\text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | $10.81^{***}$ | $10.81^{***}$ | $10.12^{***}$ | $10.12^{***}$ |
| | (6.00) | (6.00) | (6.05) | (6.05) |
| Supply $\text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | $13.75^{***}$ | $13.76^{***}$ | $11.65^{***}$ | $11.65^{***}$ |
| | (6.63) | (6.63) | (6.23) | (6.23) |
| *Safe Asset Scarcity* | | | | |
| $\text{Net}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | $5.22^{**}$ | $5.22^{**}$ | $4.35^{**}$ | $4.34^{**}$ |
| | (2.11) | (2.11) | (2.00) | (2.00) |
| $\text{Net}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | $7.14^{***}$ | $7.14^{***}$ | $7.50^{***}$ | $7.50^{***}$ |
| | (3.13) | (3.12) | (3.23) | (3.23) |
| Safe Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | $-7.77^{*}$ | | $-9.34^{*}$ | |
| | $(-1.94)$ | | $(-1.97)$ | |
| Safe Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | $-0.50$ | | $-0.36$ | |
| | $(-1.39)$ | | $(-1.15)$ | |
| Broad Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | | $-10.82^{*}$ | | $-13.52^{**}$ |
| | | $(-1.94)$ | | $(-2.24)$ |
| Broad Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | | $-0.25$ | | $-0.17$ |
| | | $(-1.06)$ | | $(-0.87)$ |
| *Demand Concentration* | | | | |
| $\widehat{\text{Demand HHI}}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | | | $7.42^{***}$ | $7.42^{***}$ |
| | | | (3.37) | (3.38) |
| $\widehat{\text{Demand HHI}}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | | | 1.27 | 1.26 |
| | | | (0.82) | (0.82) |
| *Controls* | | | | |
| Bank $\text{CDS}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | $-97.97^{***}$ | $-97.97^{***}$ | $-97.35^{***}$ | $-97.34^{***}$ |
| | $(-6.50)$ | $(-6.50)$ | $(-6.86)$ | $(-6.86)$ |
| Bank $\text{CDS}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | $-99.21^{***}$ | $-99.22^{***}$ | $-98.13^{***}$ | $-98.13^{***}$ |
| | $(-6.47)$ | $(-6.47)$ | $(-6.85)$ | $(-6.85)$ |
| Govt $\text{CDS}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | $-2.14^{***}$ | $-2.14^{***}$ | $-1.71^{**}$ | $-1.71^{**}$ |
| | $(-2.85)$ | $(-2.85)$ | $(-2.43)$ | $(-2.43)$ |
| Govt $\text{CDS}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | $-2.55$ | $-2.53$ | $-4.32^{*}$ | $-4.31^{*}$ |
| | $(-1.22)$ | $(-1.21)$ | $(-1.75)$ | $(-1.75)$ |
| $N$ | 149,337 | 149,337 | 149,293 | 149,293 |
| $R^2$ | 0.21 | 0.21 | 0.23 | 0.23 |
| Tenor FE | Yes | Yes | Yes | Yes |
| Time FE | Yes | Yes | Yes | Yes |
| Weighted | Yes | Yes | Yes | Yes |

**Table 6: Regression of the absolute value of the basis on marginal cost measures**. Table presents the regression described in section 4.5 using demand HHI measures estimated from Table A6. The dependent variable is the absolute value of the basis. To make the coefficients directly comparable, we transform the independent variables to modified $z$-scores using each variable's full sample median and standard deviation. Regression includes tenor and date fixed effects and we weight the regression by the square root of the market's share of the total daily gross notional. Within $R^2$ reported. $t$-statistics shown using robust standard errors clustered by market and date where $^{*}$ $p < 0.10$, $^{**}$ $p < 0.05$, $^{***}$ $p < 0.01$.

| | HHI Supply$^k_{t,t+n}$ | | | | Market Share$^{k,i}_{t,t+n}$ | | |
|---|---|---|---|---|---|---|---|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
| ln(1 + Avg. Book Size$^k_{t,t+n}$) | 122.6*** | | | | | | |
| | (35.63) | | | | | | |
| HHI Supply$^k_{t-1m,t-1m+n}$ | | 0.510*** | | | | | |
| | | (20.18) | | | | | |
| HHI Supply$^k_{t-1y,t-1y+n}$ | | | 0.510*** | | | | |
| | | | (22.36) | | | | |
| HHI Supply$^k_{t-5y,t-3y+n}$ | | | | 0.377*** | | | |
| | | | | (15.78) | | | |
| Market Share$^{k,i}_{t-1m,t-1m+n}$ | | | | | 0.866*** | | |
| | | | | | (24.09) | | |
| Market Share$^{k,i}_{t-1y,t-1y+n}$ | | | | | | 0.838*** | |
| | | | | | | (18.23) | |
| Market Share$^{k,i}_{t-5y,t-5y+n}$ | | | | | | | 0.736*** |
| | | | | | | | (9.39) |
| $N$ | 151,655 | 149,595 | 130,487 | 46,655 | 1,408,437 | 1,230,588 | 441,360 |
| $R^2$ | 0.08 | 0.26 | 0.28 | 0.19 | 0.75 | 0.72 | 0.57 |

**Table 7: Segmentation Persistence**. Table presents the regression of segmentation measures on its lags as well as (1 plus the log of the) average notional book size of banks in that market. The notional book size of the banks active in that market, weighted by their market share in that market. Lags for 1 month are 21 business days, 1 year is 250 business days, and 5 years is 1,250 business days. Regression includes date fixed effects. Constant omitted. Within $R^2$ reported. $t$-statistics shown using robust standard errors clustered by market and date for the first four columns and clustered by bank and date for the last three columns, where * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

| | Bank FX Share$^{i,k}_{t,ctpty}$ | | | | | |
|---|---|---|---|---|---|---|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Bank FX Share$^{i,k}_{t-6m,ctpty}$ | 0.951*** | 0.938*** | | | | 0.927*** |
| | (43.50) | (32.10) | | | | (24.60) |
| Other Bank FX Share$^{i,k}_{t,ctpty}$ | | 0.0539 | | | 0.571*** | 0.0211 |
| | | (1.50) | | | (4.41) | (0.66) |
| Bank Loan Share$^{i,k}_{t,ctpty}$ | | | 0.234*** | 0.174** | 0.155** | 0.108 |
| | | | (3.95) | (2.36) | (2.33) | (1.19) |
| Other Bank Loan Share$^{i,k}_{t,ctpty}$ | | | | 0.265*** | 0.0332 | −0.00849 |
| | | | | (3.39) | (0.74) | (−0.28) |
| $N$ | 93,636 | 93,636 | 208,386 | 208,386 | 208,386 | 93,636 |
| $R^2$ | 0.59 | 0.59 | 0.05 | 0.06 | 0.14 | 0.60 |

**Table 8: Segmentation across Counterparties**. Table presents the regression of Bank FX Share$^{i,k}_{t,ctpty}$ on several variables. Bank FX Share$^{i,k}_{t,ctpty}$ is the bank's notional FX swap exposures with that counterparty-currency pair as a share of the bank's total notional FX swaps on that day. Other Bank FX Share$^{i,k}_{t,ctpty}$ is the average market share of all banks except bank $i$. Bank Loan Share$^{i,k}_{t,ctpty}$ and Other Bank Loan Share$^{i,k}_{t,ctpty}$ are calculated analogously using secured and unsecured loans rather than FX positions. Constant omitted. Within $R^2$ reported. $t$-statistics shown using robust standard errors clustered by date and bank where $^{*}$ $p < 0.10$, $^{**}$ $p < 0.05$, $^{***}$ $p < 0.01$.

| Panel A: Event Study. | | | | | | |
|---|---|---|---|---|---|---|
| | Dependent Var.: $\lvert\text{Basis}^k_{t,t+n}\rvert$ | | | Dependent Var.: $Net^k_{t,t+n}$ | | |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| $\mathbb{I}$(Post) | $14.59^{***}$ | $14.96^{***}$ | $14.52^{***}$ | $-1.78^{***}$ | $-1.89^{***}$ | $-1.90^{***}$ |
| | (7.82) | (7.94) | (7.47) | (−3.93) | (−3.30) | (−3.13) |
| $\Delta\text{Deposits}^k_{t+n}$ | $16.38^{***}$ | $23.70^{***}$ | $21.11^{***}$ | −0.26 | 1.65 | 1.63 |
| | (6.10) | (6.96) | (5.21) | (−0.29) | (1.51) | (1.34) |
| $\mathbb{I}(\text{Post}) \times \Delta\text{Deposits}^k_{t+n}$ | $-2.11^{***}$ | $-2.37^{**}$ | $-1.95^{*}$ | $0.53^{**}$ | $0.60^{*}$ | 0.60 |
| | (−3.21) | (−2.23) | (−1.98) | (2.52) | (1.73) | (1.61) |
| Supply $\text{HHI}^k_{t,t+n}$ | | | 11.60 | | | 0.10 |
| | | | (1.54) | | | (0.05) |
| Constant | $10.07^{***}$ | 3.56 | $7.68^{*}$ | $5.66^{***}$ | $4.01^{***}$ | $4.04^{**}$ |
| | (4.72) | (1.06) | (1.79) | (4.63) | (3.10) | (2.57) |
| $N$ | 959 | 959 | 959 | 959 | 959 | 959 |
| $R^2$ | 0.51 | 0.54 | 0.55 | 0.01 | 0.04 | 0.04 |
| Tenor FE | No | Yes | Yes | No | Yes | Yes |
| Weighted | Yes | Yes | Yes | Yes | Yes | Yes |
| Panel B: Placebo Event Study. | | | | | | |
| | Dependent Var.: $\lvert\text{Basis}^k_{t,t+n}\rvert$ | | | Dependent Var.: $Net^k_{t,t+n}$ | | |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| $\mathbb{I}$(Post) | $-1.46^{***}$ | $-1.64^{***}$ | $-1.76^{***}$ | $1.44^{**}$ | $1.72^{***}$ | $1.70^{***}$ |
| | (−5.66) | (−41.16) | (−9.70) | (2.73) | (2.93) | (2.91) |
| $\Delta\text{Deposits}^k_{t+n}$ | $15.61^{***}$ | $20.26^{***}$ | $17.91^{***}$ | 0.01 | 1.95 | 1.54 |
| | (7.19) | (7.81) | (5.57) | (0.01) | (1.54) | (1.01) |
| $\mathbb{I}(\text{Post}) \times \Delta\text{Deposits}^k_{t+n}$ | 0.47 | $0.62^{**}$ | $0.70^{***}$ | 0.19 | 0.00 | 0.02 |
| | (1.31) | (2.81) | (3.35) | (0.71) | (0.00) | (0.05) |
| Supply $\text{HHI}^k_{t,t+n}$ | | | 11.22 | | | 1.94 |
| | | | (1.62) | | | (0.73) |
| Constant | $12.49^{***}$ | $8.36^{***}$ | $12.37^{***}$ | 2.49 | 0.70 | 1.40 |
| | (6.50) | (3.15) | (3.40) | (1.59) | (0.47) | (0.70) |
| $N$ | 910 | 910 | 910 | 910 | 910 | 910 |
| $R^2$ | 0.60 | 0.57 | 0.59 | 0.01 | 0.04 | 0.05 |
| Tenor FE | No | Yes | Yes | No | Yes | Yes |
| Weighted | Yes | Yes | Yes | Yes | Yes | Yes |

**Table 9: March 2023 Event Study**. Table shows the results of the regression $\lvert\text{Basis}^k_{t,t+n}\rvert = \alpha + \gamma_1\mathbb{I}(\text{Post}) + \gamma_2\Delta\text{Deposits}^k_{t+n} + \gamma_3\mathbb{I}(\text{Post}) \times \Delta\text{Deposits}^k_{t+n} + \varepsilon^k_{t,n}$. We transform $\Delta\text{Deposits}^k_{t+n}$ to a modified $z$-score using their median and standard deviation; supply HHI is transformed to a modified $z$-score using its full sample median and standard deviation. $\mathbb{I}$(Post) is equal to 1 for days after March 9, and 0 otherwise. The window is the 2 weeks before and after the event. Panel B is a placebo test which shifts the treatment date to February 9. Regression includes tenor fixed effects and we weight the regression by the square root of the market's share of the total daily gross notional. Within $R^2$ reported. $t$-statistics shown using robust standard errors clustered by market and date where $^{*}$ $p < 0.10$, $^{**}$ $p < 0.05$, $^{***}$ $p < 0.01$.

# A Online Appendix

## A.1 Model

### Derivation of Expression for the Basis (No Heterogeneity)

For risk, and safe asset scarcity, the relevant part of the first-order condition can be expressed as

$$\frac{\partial}{\partial Z_{i,k}} \sum_{k'} \frac{\gamma}{2} (\alpha_{i,k'} Z_{i,k'})^2 \sigma^2_{i,k'} + \frac{\lambda_{s,k'}}{2} \left((1-\alpha_{i,k'}) Z_{i,k'}\right)^2 = \alpha^2_{i,k} \gamma \sigma^2_{i,k} Z_{i,k} + \lambda_{s,k} (1-\alpha_{i,k})^2 Z_{i,k}.$$

We combine the risk and safe asset scarcity terms by noting that

$$\gamma \sigma^2_{i,k} \alpha^2_{i,k} Z_{i,k} + \lambda_{s,k} (1-\alpha_{i,k})^2 Z_{i,k} = \frac{\lambda_{s,k} \gamma \sigma^2_{i,k}}{\lambda_{s,k} + \gamma_i \sigma^2_{i,k}} Z_{i,k}.$$

When intermediaries are identical, this reduces to:

$$-\frac{\lambda_{s,k} \gamma \sigma^2_k}{\lambda_{s,k} + \gamma \sigma^2_k} \frac{X_k}{N_i} \tag{8}$$

For balance sheet costs, the relevant part of the first-order condition can be expressed as

$$\frac{\partial}{\partial Z_{i,k}} \left( \sum_k{}' |Z_{i,k'}| \right)^2 = \text{Sign}(Z_{i,k}) \sum_{k'} |Z_{i,k'}| = -\text{Sign}(X_k) \sum_{k'} |Z_{i,k'}|.$$

Averaging across intermediaries yields $-\text{Sign}(X_k) \sum_{k'} |X_{k'}|$. For counterparty costs, the relevant part of the first-order condition is

$$\begin{aligned}
\frac{\partial}{\partial Z_{i,k}} \frac{1}{2} \lambda_{CP} \sum_c \left( \sum_{k'} |Z_{i,k',c}| \right)^2 =& \frac{\partial}{\partial Z_{i,k}} \frac{1}{2} \lambda_{CP} \sum_c \left( \sum_{k'} \frac{X_{c,k'}}{X_{k'}} |Z_{i,k'}| \right)^2 \\
=& -\lambda_{CP} \text{Sign}(X_k) \sum_c \frac{X_{c,k}}{X_k} \sum_{k'} \frac{X_{c,k'}}{X_{k'}} |Z_{i,k'}| \\
=& -\lambda_{CP} \text{Sign}(X_k) \sum_c \left( \frac{X^2_{c,k}}{X^2_k} |Z_{i,k}| + \frac{X_{c,k}}{X_k} \sum_{k' \neq k} \frac{X_{c,k'}}{X_{k'}} |Z_{i,k'}| \right).
\end{aligned}$$

Averaging across intermediaries, we observe that:

$$\frac{1}{N_i}\lambda_{CP} \times \text{Sign}(X_k) \sum_i \sum_c \left( \frac{X_{c,k}^2}{X_k^2}|Z_{i,k}| + \frac{X_{c,k}}{X_k} \sum_{k' \neq k} \frac{X_{c,k'}}{X_{k'}}|Z_{i,k'}| \right)$$

$$= \frac{1}{N_i}\lambda_{CP} \times \text{Sign}(X_k) \left( \underbrace{|X_k| \sum_c \frac{X_{c,k}^2}{X_k^2}}_{\text{Demand Concentration in } k} + \sum_c \frac{X_{c,k}}{X_k} \sum_{k' \neq k} |X_{c,k'}| \right).$$

**Derivation of Segmentation Predictions**

For the segmentation predictions, we set $\lambda_{CP} = \lambda_{BS} = 0$. We first derive equilibrium quantities in the model focusing only on intermediaries that participate in the basis trade. Based on these quantities, we then turn to financial intermediaries' participation decisions based on the fixed participation costs.

**Equilibrium for Participating Intermediaries**. Assuming they participate ($Z_{i,k} \neq 0$), intermediary $i$'s first order condition can be written as

$$\text{Basis}_k = -\frac{\lambda_{s,k}\gamma\sigma_{i,k}^2}{\lambda_{s,k} + \gamma\sigma_{i,k}^2} Z_{i,k}. \tag{9}$$

This holds for each participating intermediary $i$, meaning that for any two intermediaries $i$ and $j$,

$$\frac{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}{\lambda_{s,k} + \gamma_i\sigma_{i,k}^2} Z_{i,k} = \frac{\lambda_{s,k}\gamma_j\sigma_{j,k}^2}{\lambda_{s,k} + \gamma_j\sigma_{j,k}^2} Z_{j,k}, \tag{10}$$

or equivalently,

$$Z_{j,k} = \frac{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}{\lambda_{s,k} + \gamma_i\sigma_{i,k}^2} \frac{\lambda_{s,k} + \gamma_j\sigma_{j,k}^2}{\lambda_{s,k}\gamma_j\sigma_{j,k}^2} Z_{i,k}. \tag{11}$$

The market clearing condition is $\sum_j Z_{j,k} = -X_k$, which can be written as

$$Z_{i,k} \frac{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}{\lambda_{s,k} + \gamma_i\sigma_{i,k}^2} \sum_j \frac{\lambda_{s,k} + \gamma_j\sigma_{j,k}^2}{\lambda_{s,k}\gamma_j\sigma_{j,k}^2} = -X_k, \tag{12}$$

which provides a solution for $Z_{i,k}$:

$$Z_{i,k} = -\frac{X_k}{\sum_j \frac{\lambda_{s,k} + \gamma_j\sigma_{j,k}^2}{\lambda_{s,k}\gamma_j\sigma_{j,k}^2}} \times \frac{\lambda_{s,k} + \gamma_i\sigma_{i,k}^2}{\lambda_{s,k}\gamma\sigma_{i,k}^2}. \tag{13}$$

We can substitute this into the first-order condition to derive an expression for the basis.

$$\text{Basis}_k = \frac{X_k}{\sum_j \frac{\lambda_{s,k}+\gamma_j\sigma_{j,k}^2}{\lambda_{s,k}\gamma_j\sigma_{j,k}^2}}. \tag{14}$$

We can further write $\sum_j \frac{\lambda_{s,k}+\gamma_j\sigma_{j,k}^2}{\lambda_{s,k}\gamma_j\sigma_{j,k}^2}$ as

$$\begin{aligned}\sum_i \frac{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2}{\lambda_{s,k}\gamma_i\sigma_{i,k}^2} =& \frac{\sum_i(\lambda_{s,k}+\gamma_i\sigma_{i,k}^2)\prod_{j\neq i}\lambda_{s,k}\gamma_i\sigma_{j,k}^2}{\prod_i \lambda_{s,k}\gamma_i\sigma_{i,k}^2} \\ =& \lambda_{s,k}\sum_i \frac{\prod_{j\neq i}\lambda_{s,k}\gamma_j\sigma_{j,k}^2}{\prod_j \lambda_{s,k}\gamma_j\sigma_{j,k}^2} + \sum_i \frac{\gamma_i\sigma_{i,k}^2\prod_{j\neq i}\lambda_{s,k}\gamma_j\sigma_{j,k}^2}{\prod_j \lambda_{s,k}\gamma_j\sigma_{j,k}^2} \\ =& \frac{N_{p,k}}{\lambda_{s,k}} + \sum_i \frac{1}{\gamma_i\sigma_{i,k}^2},\end{aligned}$$

where $N_{p,k} \leq N_i$ is the number of intermediaries participating in the basis trade. Hence, we can write the basis as

$$\text{Basis}_k = \frac{X_k}{\frac{N_p}{\lambda_{s,k}} + \sum_i \frac{1}{\gamma_i\sigma_{i,k}^2}}. \tag{15}$$

**Existence of Equilibrium**. We focus on Nash equilibria. For an equilibrium to exist, there has to be at least one intermediary for whom basis profits justify the fixed costs. We derive the conditions for an intermediary to participate as the only intermediary, and then discuss the more general behavior of intermediaries in Nash equilibria.

For intermediary $i$ to be a sole participant, we need that

$$\lambda_{PC,k} < -\text{Basis}_k Z_{i,k} - \frac{\gamma_i}{2}(\alpha_{i,k}Z_{i,k})^2\sigma_{i,k}^2 - \frac{\lambda_{s,k}}{2}(1-\alpha_{i,k})^2 Z_{i,k}^2.$$

By assumption, we have that

$$\text{Basis}_k = \frac{X_k}{\frac{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2}{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}}, Z_{i,k} = -X_k,$$

so, the condition becomes

$$\begin{aligned}\lambda_{PC,k} <& \underbrace{\frac{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2}X_k^2}_{\text{Basis}_k X_k} - \underbrace{\frac{1}{2}\frac{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2}X_k^2}_{\text{Risk and Safe Asset Scarcity}} \\ =& \frac{1}{2}\frac{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2}X_k^2.\end{aligned}$$

This condition reduces to:

$$\gamma_i \sigma_{i,k}^2 > \frac{1}{\frac{X_k^2}{2\lambda_{PC,k}} - \frac{1}{\lambda_{s,k}}}. \tag{16}$$

The condition provides a lower bound for $\sigma_{i,k}$. Equilibrium exists as long as there is at least one intermediary for whom the alternative to currency $k$ safe assets is sufficiently risky that intermediary $i$ does not have the capacity to drive all basis profits to be lower than the fixed cost of participation in currency $k$.

**Intermediaries' Participation Decisions**. In order to participate in the basis trade in currency $k$, intermediaries must find it worthwhile to pay the fixed cost of participation in equilibrium. Using the equilibrium quantities previously derived, this means that, assuming intermediary $i$ participates,

$$\begin{aligned}
\lambda_{PC,k} <& -\text{Basis}_k Z_{i,k} - \frac{\gamma_i}{2}(\alpha_{i,k} Z_{i,k})^2 \sigma_{i,k}^2 - \frac{\lambda_{s,k}}{2}(1-\alpha_{i,k})^2 Z_{i,k}^2 \\
=& \underbrace{\frac{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2} Z_{i,k}^2}_{\text{Basis}_k Z_{i,k}} - \underbrace{\frac{1}{2}\frac{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2} Z_{i,k}^2}_{\text{Risk and Safe Asset Scarcity}} \\
=& \frac{1}{2}\frac{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2} Z_{i,k}^2.
\end{aligned}$$

For ease of notation, we denote $A \equiv \frac{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2}{\lambda_{s,k}\gamma_i\sigma_{i,k}^2} = \frac{1}{\lambda_{s,k}} + \frac{1}{\gamma_i\sigma_{i,k}^2}$, and $B \equiv \sum_{j\neq i} \frac{\lambda_{s,k}+\gamma_j\sigma_{j,k}^2}{\lambda_{s,k}\gamma_j\sigma_{j,k}^2}$, where $j$ are the participating intermediaries (other than intermediary $i$). We want to solve for $A$ to characterize $i$'s participation threshold (where a higher $A$ indicates that substitution from safe assets is less risky for intermediary $i$). We can re-write our participation condition as:

$$\begin{aligned}
\lambda_{PC,k} <& \frac{1}{2}\frac{\lambda_{s,k}\gamma_i\sigma_{i,k}^2}{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2} Z_{i,k}^2 \\
=& \frac{1}{2A}\frac{A^2 X_k^2}{(A+B)^2}.
\end{aligned}$$

This yields a quadratic inequality for $A$ to justify intermediary $i$'s participation:

$$\lambda_{PC,k} A^2 + \left(2B\lambda_{PC,k} - \frac{1}{2}X_k^2\right) A + \lambda_{PC,k} B^2 < 0.$$

This yields a lower and upper bound for $A$:

$$\frac{\frac{1}{2}X_k^2 - 2B\lambda_{PC,k} - \sqrt{\left(\frac{1}{2}X_k^2 - 2B\lambda_{PC,k}\right)^2 - 4\lambda_{PC,k}^2 B^2}}{2\lambda_{PC,k}} \leq A$$
$$\leq \frac{\frac{1}{2}X_k^2 - 2B\lambda_{PC,k} + \sqrt{\left(\frac{1}{2}X_k^2 - 2B\lambda_{PC,k}\right)^2 - 4\lambda_{PC,k}^2 B^2}}{2\lambda_{PC,k}}.$$

Denoting the lower and upper bounds as $\underline{A}$ and $\overline{A}$, we note that the bounds imply that intermediary $i$ enters if and only if

$$\underline{A} \leq \frac{1}{\lambda_{s,k}} + \frac{1}{\gamma_i \sigma_{i,k}^2} \leq \overline{A},$$

which implies that

$$\frac{1}{(\overline{A} - 1/\lambda_{s,k})} \leq \gamma_i \sigma_{i,k}^2 \leq \frac{1}{(\underline{A} - 1/\lambda_{s,k})}. \tag{17}$$

The upper bound for $\gamma_i \sigma_{i,k}^2$ means that intermediary $i$ must not face too much risk in substituting into risky assets (relative to participating intermediaries); otherwise, the fixed participation cost is too high for them. The lower bound for $\gamma_i \sigma_{i,k}^2$ means that intermediary $i$ cannot have a risky alternative that is *too* close to risk-free; otherwise, they could drive the profits of the basis trade to zero, making the trade sufficiently unprofitable to justify their participation.

Importantly, with sufficient heterogeneity in $\sigma_{i,k}^2$ across intermediaries, we have limited participation in market $k$.

**Basis with Balance Sheet Costs, Counterparty Costs, and Heterogeneity**

In currency $k$, we denote the number of participating intermediaries as $N_{p,k}$. We take participating intermediary $i$'s first order condition with respect to $Z_{i,k}$.

For risk, and safe asset scarcity, the relevant part of the first-order condition can be expressed as

$$\frac{\partial}{\partial Z_{i,k}} \sum_k \frac{\gamma}{2} (\alpha_{i,k} Z_{i,k})^2 \sigma_{i,k}^2 + \frac{\lambda_{s,k}}{2} \left((1 - \alpha_{i,k}) Z_{i,k}\right)^2 = \alpha_{i,k}^2 \gamma \sigma_{i,k}^2 Z_{i,k} + \lambda_{s,k} (1 - \alpha_{i,k})^2 Z_{i,k}.$$

We combine the risk and safe asset scarcity terms by noting that

$$\gamma \sigma_{i,k}^2 \alpha_{i,k}^2 Z_{i,k} + \lambda_{s,k} (1 - \alpha_{i,k})^2 Z_{i,k} = \frac{\lambda_{s,k} \gamma \sigma_{i,k}^2}{\lambda_{s,k} + \gamma_i \sigma_{i,k}^2} Z_{i,k}.$$

For balance sheet costs, the relevant part of the first-order condition can be expressed as

$$\begin{aligned}\frac{\partial}{\partial Z_{i,k}}\frac{1}{2}\lambda_{BS}\left(\sum_{k'}|Z_{i,k'}|\right)^2 =&\lambda_{BS}Z_{i,k}+\lambda_{BS}\text{Sign}(Z_{i,k})\sum_{k'\neq k}|Z_{i,k'}|\\ =&\lambda_{BS}Z_{i,k}-\lambda_{BS}\text{Sign}(X_k)\sum_{k'\neq k}|Z_{i,k'}|.\end{aligned}$$

For counterparty costs, the relevant part of the first-order condition is

$$\begin{aligned}\frac{\partial}{\partial Z_{i,k}}\frac{1}{2}\lambda_{CP}\sum_c\left(\sum_{k'}|Z_{i,k',c}|\right)^2 =&\frac{\partial}{\partial Z_{i,k}}\frac{1}{2}\lambda_{CP}\sum_c\left(\sum_{k'}\frac{X_{c,k'}}{X_{k'}}|Z_{i,k'}|\right)^2\\ =&-\lambda_{CP}\text{Sign}(X_k)\sum_c\frac{X_{c,k}}{X_k}\sum_{k'}\frac{X_{c,k'}}{X_{k'}}|Z_{i,k'}|\\ =&\lambda_{CP}Z_{i,k}\sum_c\frac{X_{c,k}^2}{X_k^2}-\lambda_{CP}\text{Sign}(X_k)\sum_c\sum_{k'\neq k}\frac{X_{c,k}}{X_k}\frac{X_{c,k'}}{X_{k'}}|Z_{i,k'}|.\end{aligned}$$

So, the first order condition can be expressed as

$$\begin{aligned}\text{Basis}_k=&-\left(\frac{\lambda_{s,k}\gamma\sigma_{i,k}^2}{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2}+\lambda_{BS}+\lambda_{CP}\sum_c\frac{X_{c,k}^2}{X_k^2}\right)Z_{i,k}\\ &+\lambda_{BS}\text{Sign}(X_k)\sum_{k'\neq k}|Z_{i,k'}|+\lambda_{CP}\text{Sign}(X_k)\sum_c\sum_{k'\neq k}\frac{X_{c,k}}{X_k}\frac{X_{c,k'}}{X_{k'}}|Z_{i,k'}|.\end{aligned}$$

which we can re-write in terms of $Z_{i,k}$:

$$\begin{aligned}Z_{i,k}=&\left(-\text{Basis}_k+\lambda_{BS}\text{Sign}(X_k)\sum_{k'\neq k}|Z_{i,k'}|+\lambda_{CP}\text{Sign}(X_k)\sum_c\sum_{k'\neq k}\frac{X_{c,k}}{X_k}\frac{X_{c,k'}}{X_{k'}}|Z_{i,k'}|\right)\\ &\times\left(\frac{\lambda_{s,k}\gamma\sigma_{i,k}^2}{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2}+\lambda_{BS}+\lambda_{CP}\sum_c\frac{X_{c,k}^2}{X_k^2}\right)^{-1}.\end{aligned}$$

Summing over participating intermediaries,

$$X_k=\sum_i^{N_{p,k}}\frac{\text{Basis}_k-\lambda_{BS}\text{Sign}(X_k)\sum_{k'\neq k}|Z_{i,k'}|-\lambda_{CP}\text{Sign}(X_k)\sum_c\sum_{k'\neq k}\frac{X_{c,k}}{X_k}\frac{X_{c,k'}}{X_{k'}}|Z_{i,k'}|}{\frac{\lambda_{s,k}\gamma\sigma_{i,k}^2}{\lambda_{s,k}+\gamma_i\sigma_{i,k}^2}+\lambda_{BS}+\lambda_{CP}\sum_c\frac{X_{c,k}^2}{X_k^2}}.$$

This can then be re-written as:

$$\begin{aligned}\text{Basis}_k =& X_k \left( \sum_i^{N_{p,k}} \frac{1}{\frac{\lambda_{s,k}\gamma\sigma^2_{i,k}}{\lambda_{s,k}+\gamma_i\sigma^2_{i,k}} + \lambda_{BS} + \lambda_{CP}\sum_c \frac{X^2_{c,k}}{X^2_k}} \right)^{-1} \\ &+ \underbrace{\lambda_{BS}\text{Sign}(X_k)\left( \sum_i^{N_{p,k}} \sum_{k' \neq k} |Z_{i,k'}| \right)}_{\text{Balance Sheet usage of participating intermediaries}} \left( \sum_i^{N_{p,k}} \frac{1}{\frac{\lambda_{s,k}\gamma\sigma^2_{i,k}}{\lambda_{s,k}+\gamma_i\sigma^2_{i,k}} + \lambda_{BS} + \lambda_{CP}\sum_c \frac{X^2_{c,k}}{X^2_k}} \right)^{-1} \\ &+ \underbrace{\lambda_{CP}\text{Sign}(X_k)\left( \sum_c \sum_{k' \neq k} \frac{X_{c,k}}{X_k}\frac{X_{c,k'}}{X_{k'}} |Z_{i,k'}| \right)}_{\text{Counterparty Exposure of Participating Intermediaries}} \left( \sum_i^{N_{p,k}} \frac{1}{\frac{\lambda_{s,k}\gamma\sigma^2_{i,k}}{\lambda_{s,k}+\gamma_i\sigma^2_{i,k}} + \lambda_{BS} + \lambda_{CP}\sum_c \frac{X^2_{c,k}}{X^2_k}} \right)^{-1}.\end{aligned}$$

The expression for the basis yields the same predictions as the main text. In addition, we note that the expression highlights the logic behind Prediction 6 in a more general sense: the balance sheet usage and counterparty exposure of the *participating* intermediaries are the features that are relevant for understanding the basis in currency $k$, not the balance sheet usage and counterparty exposure of the overall financial sector.

## A.2 Data Details

### FR 2052a Complex Institution Liquidity Monitoring Report

We limit the sample to firms that report daily data through the sample, focusing on the largest consolidated entity firm rather than individual material entities. We exclude transactions with internal counterparties. We drop foreign exchange options except when it appears the foreign exchange flag is a data entry error. Foreign exchange options represent a small share of the notional positions for banks in our sample, so including foreign exchange options that could provide dollar lending in our sample does not meaningfully change the results. We drop transactions in which one leg's currency is denominated as "other" and keep transactions that include USD. We exclude transactions where the first leg has not yet settled, since we are interested in actual lending rather than future obligations to lend. We include transactions that have likely already settled, as is the case for a handful of transactions when the forward start date and maturity date are reported to be the same although this convention typically indicates a forward transaction that will have an open maturity. We drop roughly half a dozen dates with outliers, and we drop dates where there are fewer than seven filers' data available. Since the set of banks reporting daily has changed over time, we focus on subset of

the largest banks that have consistently been daily reporters over the whole sample.[31] The data collection instructions were modestly updated in April 2022 to provide additional details on certain segments and we clean the data so the data before and after 2022 are directly comparable. During the brief period that banks reported two sets of data, one to satisfy the pre-April 2022 instruction data and the other to satisfy post-April 2022 instructions, we use the previous instruction data. We also adjust the maturities of the contracts to be consistent through the sample as the maturity buckets for some increased by 1 day with the updated instructions (e.g., 1 to 2 year contracts, previously reported at 366 days, started being reported at 365).

**Level 1 HQLA Assets**

The following security types are considered level 1 HQLAs so long as they meet the asset-specific tests in section 20 of Regulation WW:

- Cash
- Debt issued by the U.S. Treasury
- U.S. Government Agency-issued debt (excluding the U.S. Treasury) with a US Government guarantee
- Vanilla debt (including pass-through MBS) guaranteed by a U.S. Government Agency, where the U.S. Government Agency has a full U.S. Government guarantee
- Structured debt (excluding pass-through MBS) guaranteed by a U.S. Government Agency, where the U.S. Government Agency has a full U.S. Government guarantee
- Other debt with a U.S. Government guarantee
- Debt issued by non-U.S. Sovereigns (excluding central banks) with a 0% RW
- Debt issued by multilateral development banks or other supranationals with a 0% RW
- Debt with a non-U.S. sovereign (excluding central banks) or multilateral development bank or other supranational guarantee, where guaranteeing entity has a 0% RW

[31] Over time, a handful of firms move from daily to monthly filing, or vice versa. Including these firms does not materially affect our results since the firms changing their filing frequency account for a comparatively small share of dollar lending.

- Debt issued or guaranteed by a non-U.S. Sovereign (excluding central banks) that does not have a 0% RW, but supports outflows that are in the same jurisdiction of the sovereign and are denominated in the home currency of the sovereign
- Securities issued or guaranteed by a central bank with a 0% RW
- Securities issued or guaranteed by a non-U.S. central bank that does not have a 0% RW, but supports outflows that are in the same jurisdiction of the central bank and are denominated in the home currency of the central bank

**OIS and FX Rates**

We adjust the OIS rates for two currencies: CHF and EUR. CHF OIS rates were based on TOIS fixings until December 29, 2017 when it switched to SARON fixings. As a result, we split the CHF OIS rates to use the TOIS swaps before that date and the SARON swaps after that date. Bloomberg also does not have a full timeseries for the 3w, 4m, and 5m OIS rate; when it is missing, we linearly interpolate the rate by estimating the curve each day. For the 3w CHF tenor, we estimate it based CHF OIS tenors with fewer than 100 days maturity; for the 4m and 5m, we use CHF OIS tenors with maturities between 28 and 181 days, exclusive. EURO OIS rates were based on EONIA until January 2, 2022 when the benchmark changed to ESTR. ESTR was introduced in October 2019 but Bloomberg provides backfilled rates for all tenors in our sample except for 3 weeks. We use the ESTR OIS rates when they are available, otherwise we use EONIA OIS rates.

We clean five data points for JPY forward points in April 2019 that appear incorrect because the days to maturity for consecutive contracts are the same. For example, the 2w and 3w forwards list the same days to maturity on April 11, 2019. In these five cases, we manually change the days to maturity to 7 $\times$ the contract's maturity in weeks.

## A.3 Figures

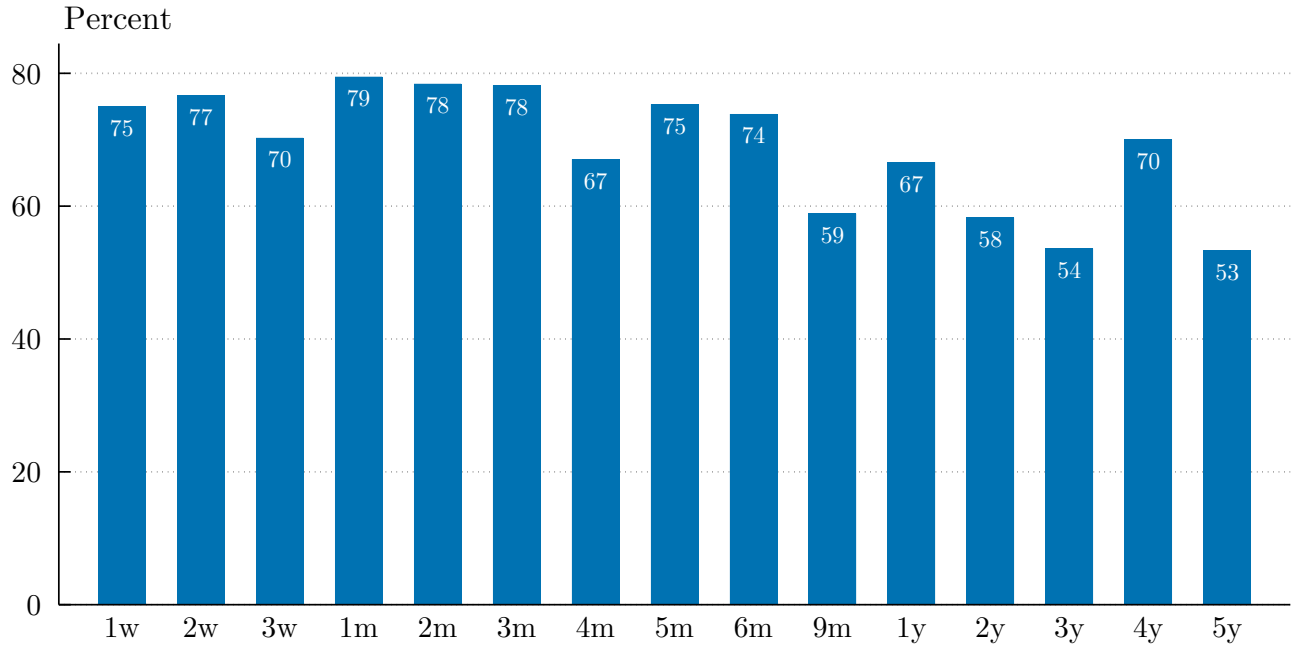



**Figure A1: Proportion of CIP Deviation Variance Explained by First Principal Component**. Figure plots the proportion of the variance explained by the first principal component after estimating 6 principal components across the signed CIP deviations by tenor.

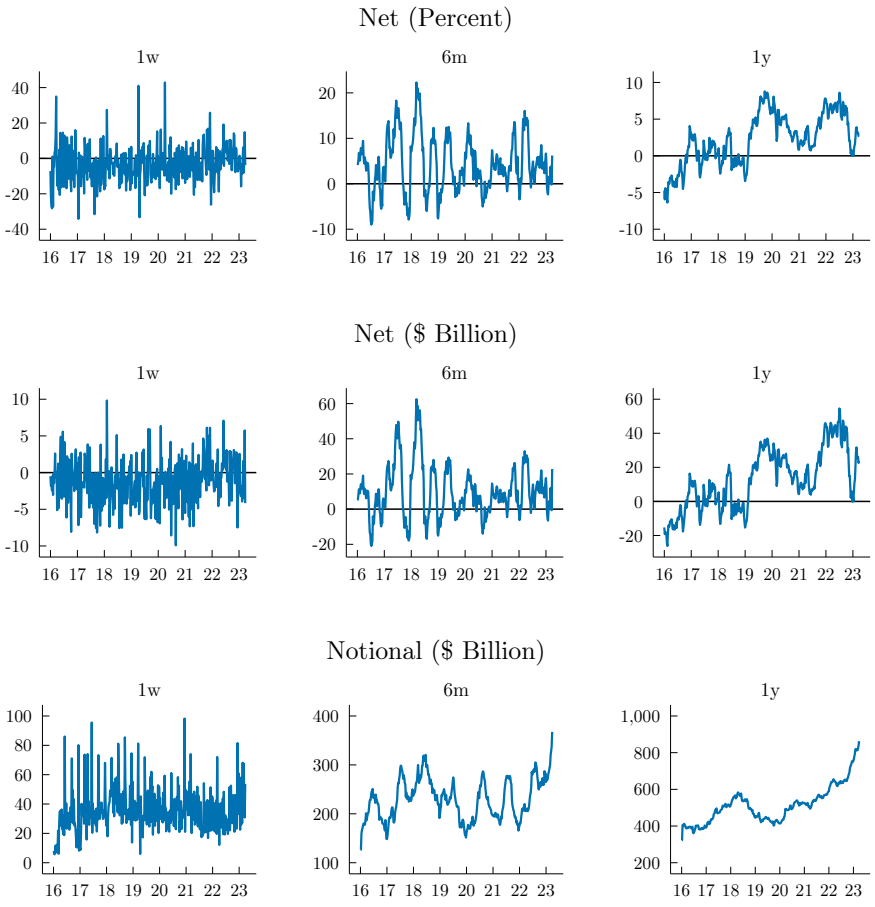



**Figure A2: *Net* for EUR across several tenors**. Top panel plots our main measure, $Net^k_{t,t+n}$, for EUR contracts across various tenors. Middle panel plots the level of net dollar lending. Bottom panel plots the notional dollars across borrowing and lending transactions for EUR across the highlighted tenors. All figures plot weekly averages based on daily observations.

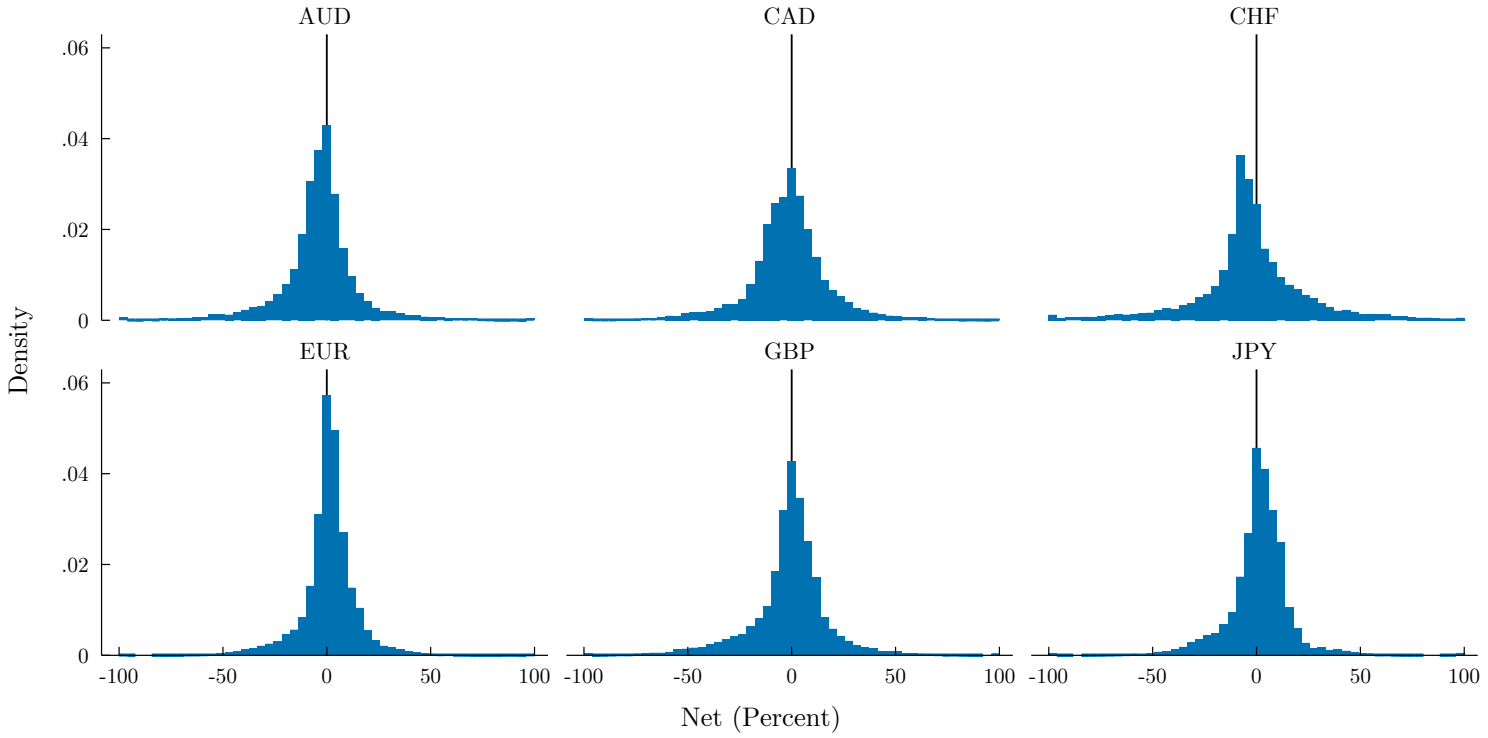



**Figure A3: Most Markets are Nearly Matched Books.** Figure plots the histogram of $Net^k_{t,t+n}$across all tenors within a given currency.

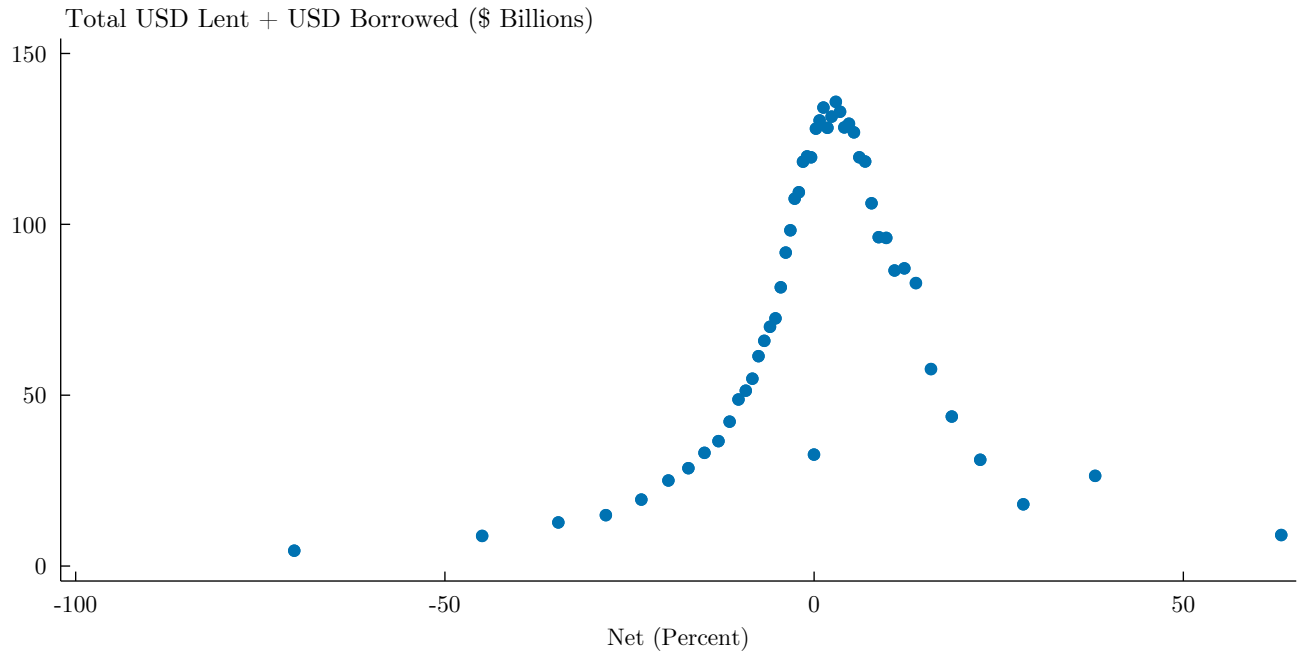



**Figure A4: Size vs. *Net*.** Figure presents the binscatter of size of the market—defined as the sum of dollars lent and borrowed in billions—at the daily-currency-tenor level.

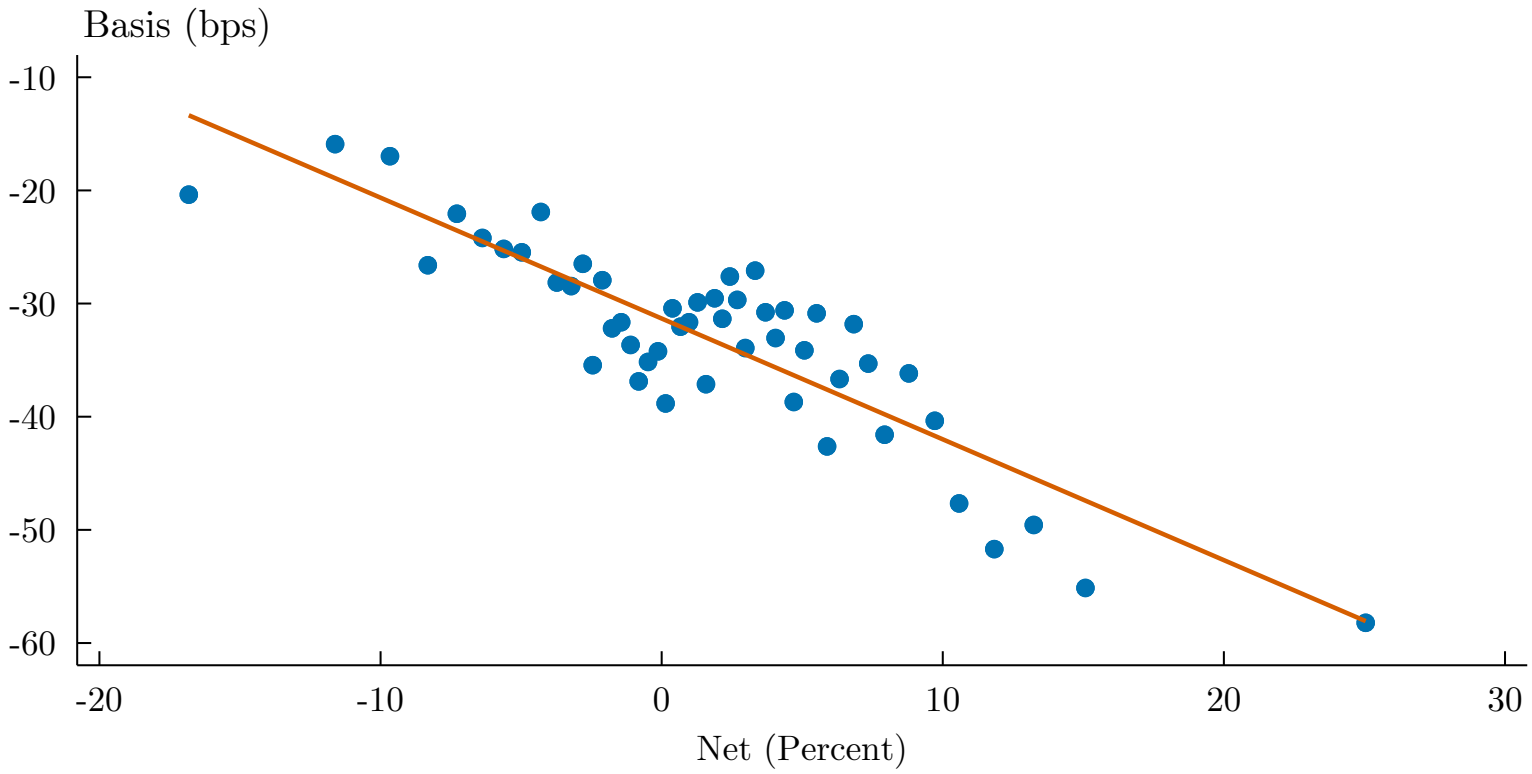



**Figure A5: Basis vs. *Net*.** Figure presents the binscatter of the basis on $Net^k_{t,t+n}$after averaging on a monthly frequency. $Net^k_{t,t+n}$is at the month by currency by tenor level.

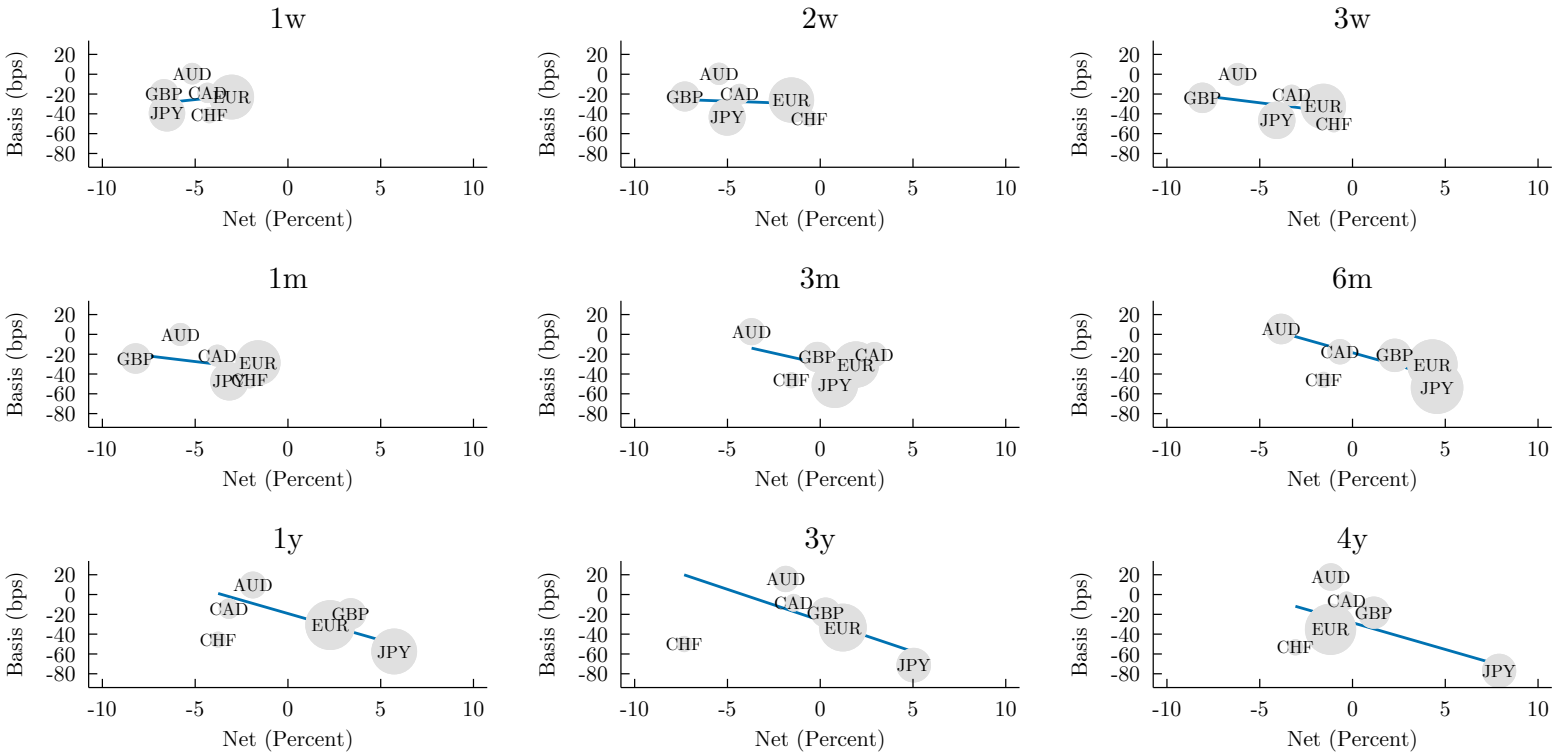



**Figure A6: By Tenor: Basis vs. *Net*.** Figure presents the scatter of the average basis on the average $Net^k_{t,t+n}$.

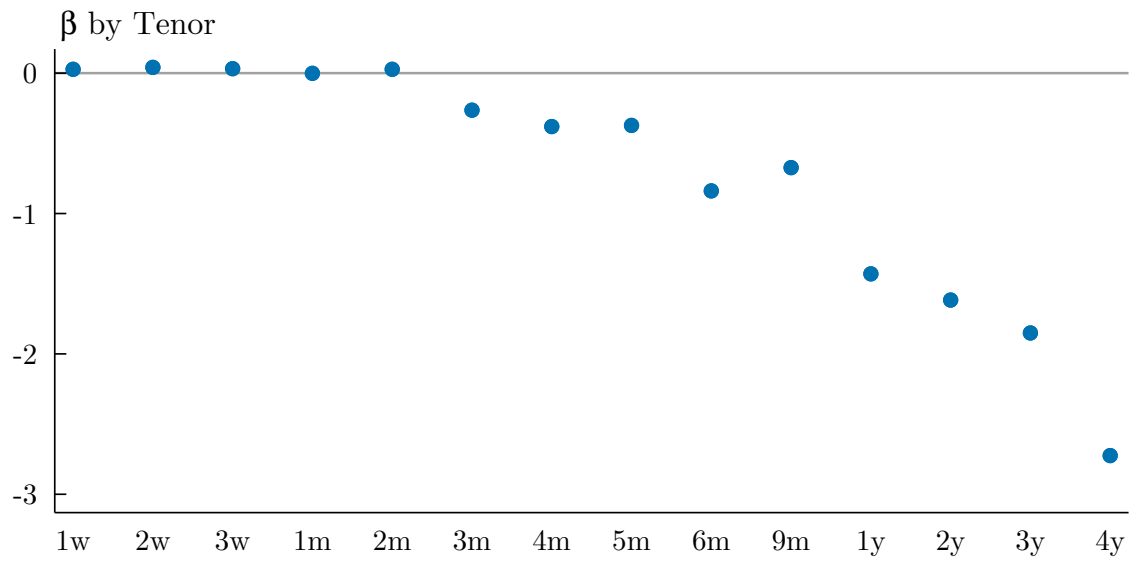



**Figure A7: Regression Coefficient by Tenor: Basis vs. *Net*.** Figure the $\beta$ estimated by running the following regression separately for each tenor: $\text{Basis}^k_{t,t+n} = \alpha + \beta Net^k_{t,t+n} + \varepsilon^k_{t,t+n}$.

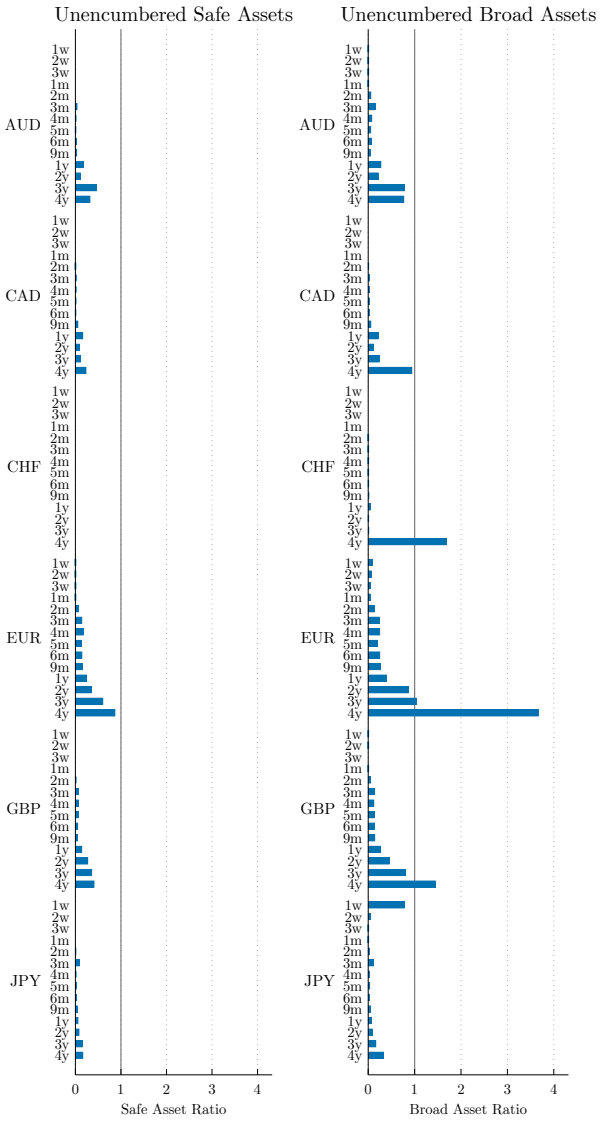



**Figure A8: Safe Asset Scarcity by Market.** Figure plots median safe asset and broad asset ratios when $Net^k_{t,t+n}$is positive and matching the tenor of the net dollar lending and the foreign safe asset.

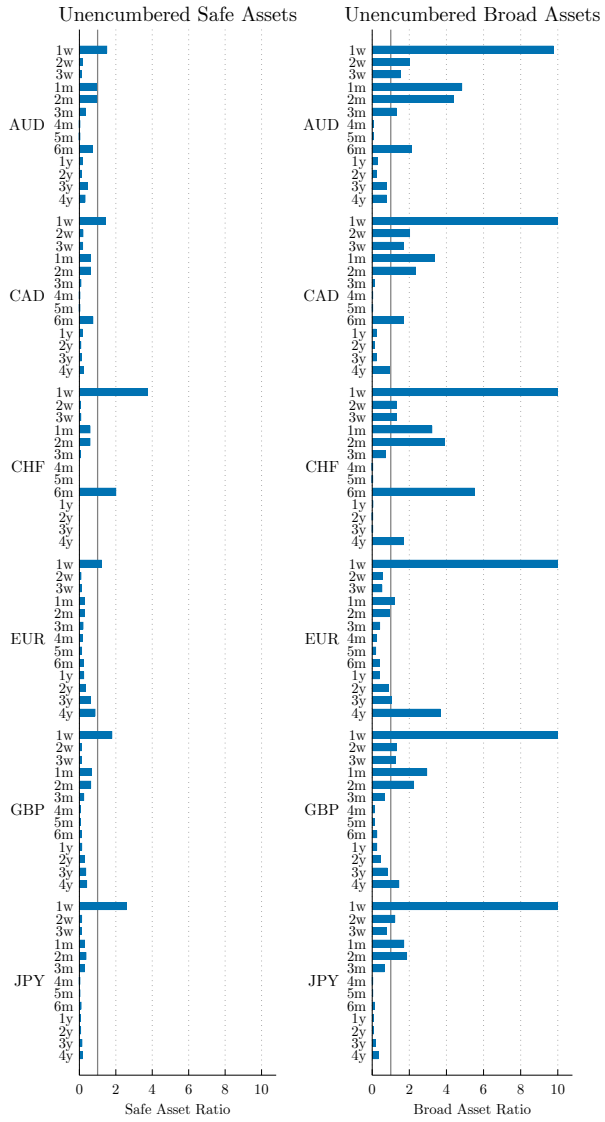



**Figure A9: Safe Asset Scarcity by Market with Rounded Tenors**. Figure plots median safe asset and broad asset ratios when $Net^k_{t,t+n}$is positive and when bucketing net dollar lending and foreign safe assets into the nearest benchmark tenor. Unlike Figure A8 this measure allows for some maturity mismatch in the trade. Values are truncated at 10.

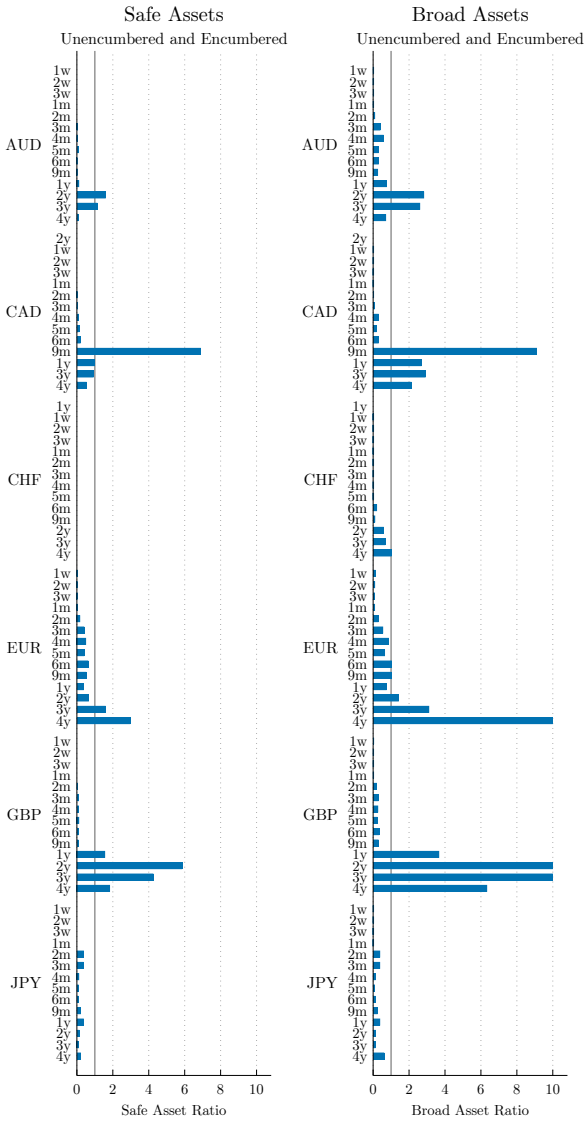



**Figure A10: Safe Asset Scarcity by Market, including encumbered assets.** Figure plots median safe asset and broad asset ratios when $Net^k_{t,t+n}$is positive when matching the tenor of the net dollar lending and the foreign safe asset. Values are truncated at 10.

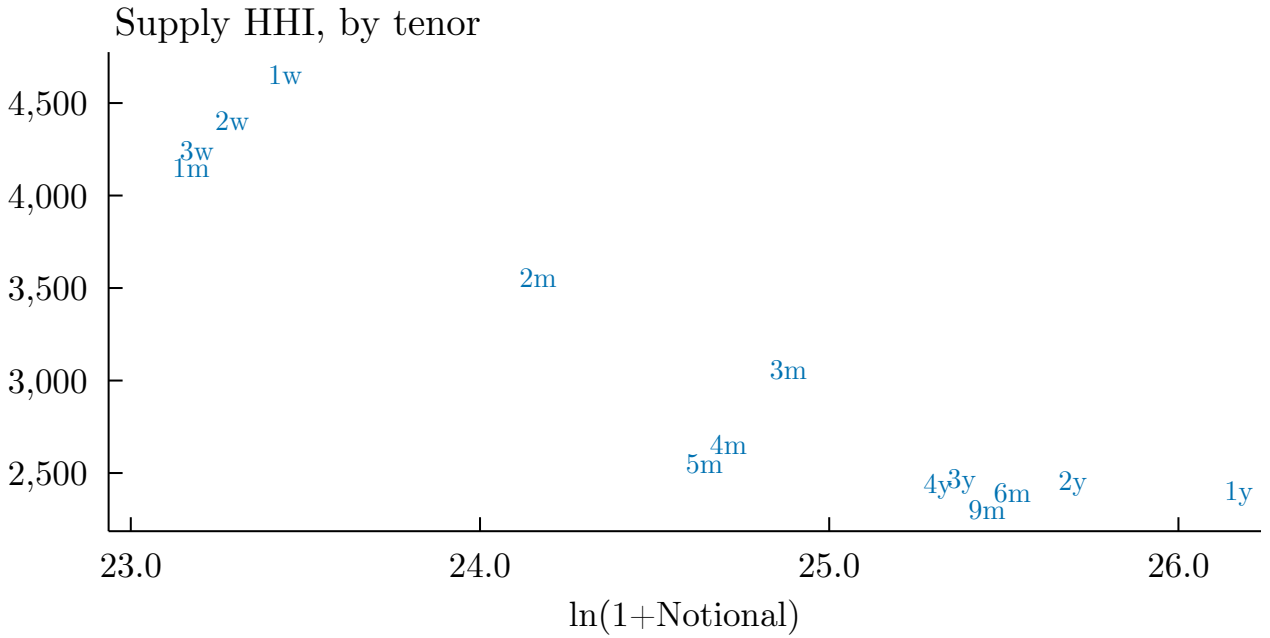



**Figure A11: Supply Segmentation.** Figure plots the average of daily Tenor Supply $\text{HHI}_{t,t+n}$ against (log of 1 plus) the average notional of that market.

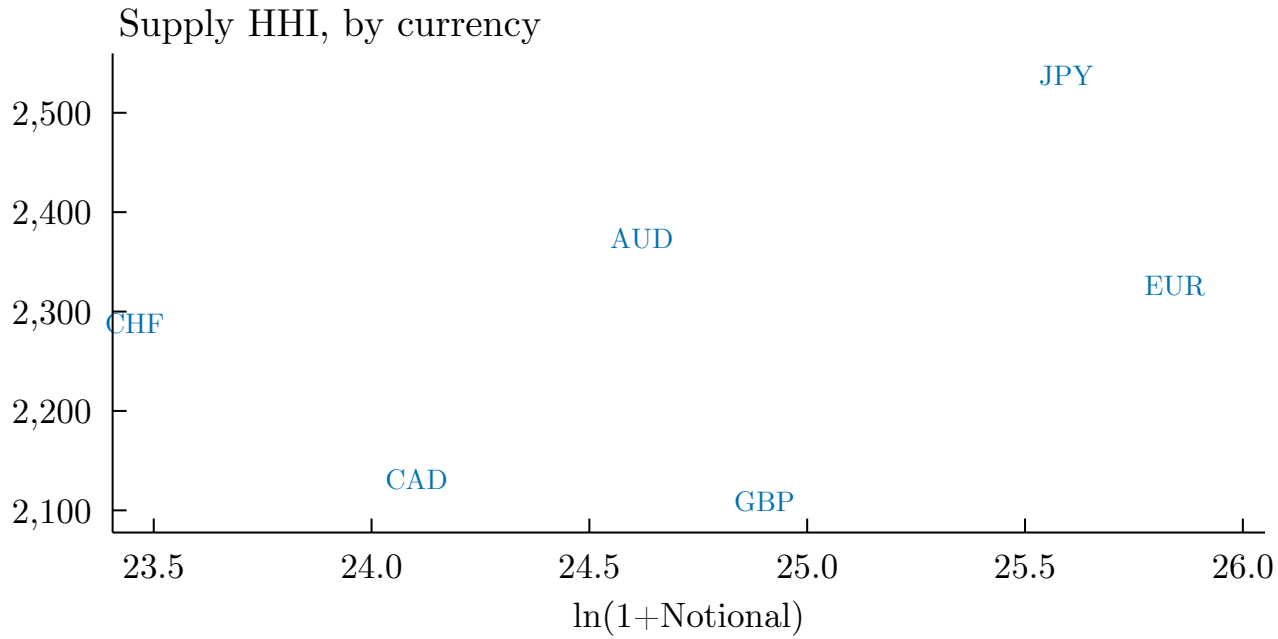



**Figure A12: Supply Segmentation by Currency.** Figure plots the average of daily Currency Supply $\text{HHI}_t^k$ against (log of 1 plus) the average notional of that market.

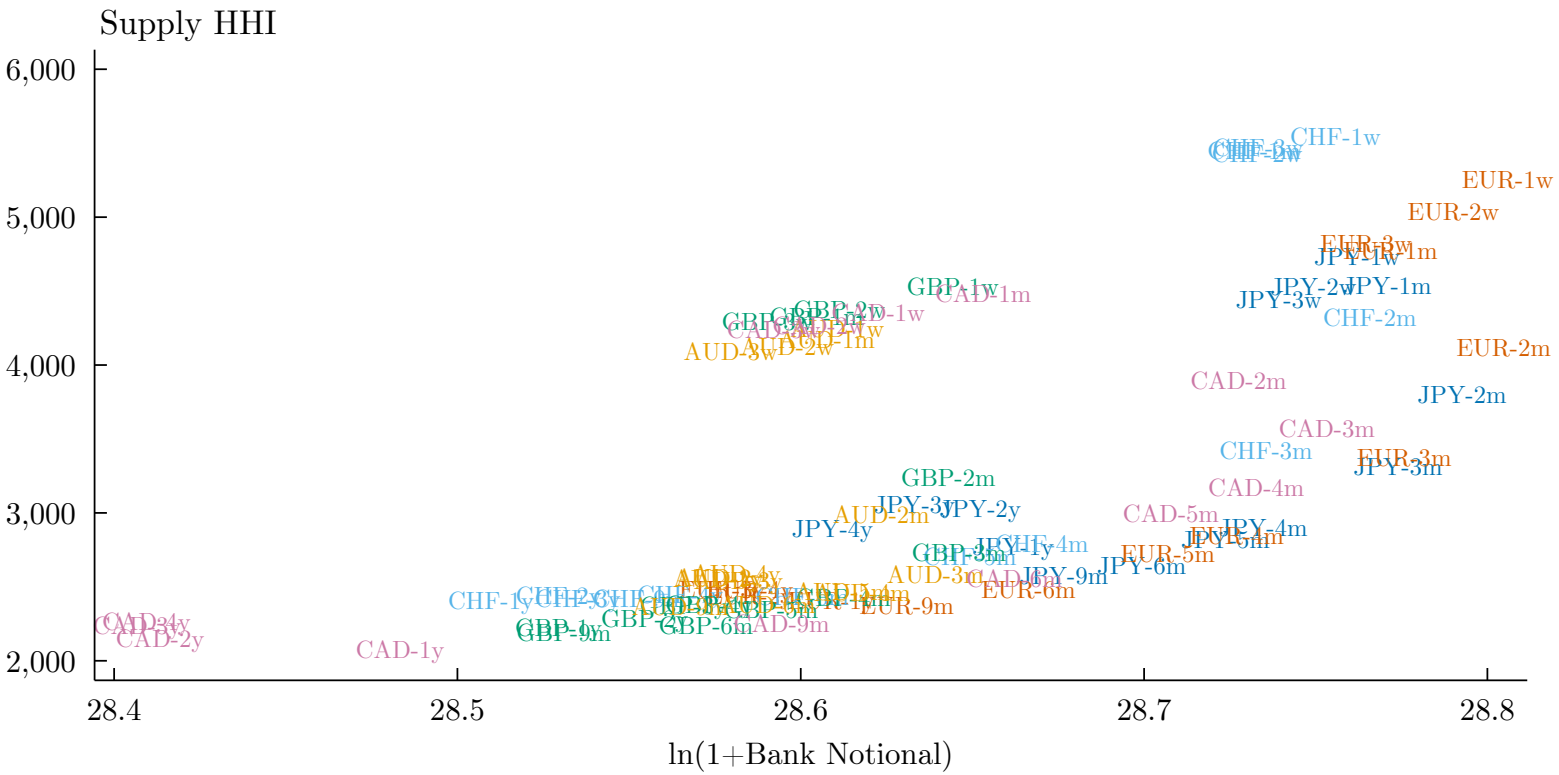



**Figure A13: Supply Segmentation vs. Bank Size** Figure plots the average of daily Supply $\text{HHI}^k_{t,t+n}$ against (log of 1 plus the) the notional book size of the banks active in that market weighted by their market share.

| Mean ($ Billions) | | | | | | | |
|---|---|---|---|---|---|---|---|
| | AUD | CAD | CHF | EUR | GBP | JPY | Mean |
| 1w | −0.4 | −0.2 | 0.0 | −1.2 | −1.1 | −1.7 | −0.8 |
| 2w | −0.4 | −0.2 | 0.2 | −0.8 | −1.2 | −1.1 | −0.6 |
| 3w | −0.4 | −0.2 | 0.1 | −0.8 | −1.2 | −0.9 | −0.5 |
| 1m | −0.4 | −0.2 | 0.1 | −0.7 | −1.2 | −0.8 | −0.5 |
| 2m | −1.0 | 0.0 | −0.4 | 1.1 | −0.9 | −1.9 | −0.5 |
| 3m | −1.6 | 1.2 | 0.0 | 2.4 | 0.0 | 2.3 | 0.7 |
| 4m | −1.5 | 0.7 | 0.0 | 2.6 | 1.3 | 3.6 | 1.1 |
| 5m | −1.5 | 0.4 | −0.1 | 3.1 | 1.2 | 2.7 | 1.0 |
| 6m | −3.7 | −0.9 | −0.3 | 10.0 | 2.3 | 12.9 | 3.4 |
| 9m | −2.8 | −2.9 | −0.5 | 10.2 | 3.1 | 9.7 | 2.8 |
| 1y | −3.0 | −2.0 | −1.8 | 12.9 | 6.3 | 24.3 | 6.1 |
| 2y | 0.7 | −1.2 | −2.0 | 6.1 | 0.5 | 11.0 | 2.5 |
| 3y | −1.2 | −0.4 | −1.9 | 2.9 | −0.1 | 6.6 | 1.0 |
| 4y | −0.8 | 0.0 | −0.7 | −3.6 | 0.6 | 8.3 | 0.6 |
| Mean | −1.3 | −0.4 | −0.5 | 3.2 | 0.7 | 5.4 | |

| Standard Deviation ($ Billions) | | | | | | | |
|---|---|---|---|---|---|---|---|
| | AUD | CAD | CHF | EUR | GBP | JPY | Mean |
| 1w | 2.0 | 1.8 | 1.7 | 6.9 | 4.0 | 4.4 | 3.5 |
| 2w | 1.9 | 1.7 | 1.5 | 6.7 | 4.0 | 4.1 | 3.3 |
| 3w | 1.8 | 1.7 | 1.4 | 6.1 | 3.8 | 4.0 | 3.2 |
| 1m | 1.9 | 1.8 | 1.4 | 5.8 | 3.9 | 4.1 | 3.1 |
| 2m | 3.2 | 3.1 | 1.9 | 7.6 | 5.0 | 6.8 | 4.6 |
| 3m | 4.2 | 4.4 | 2.5 | 10.0 | 6.4 | 14.1 | 6.9 |
| 4m | 3.9 | 3.8 | 2.0 | 8.3 | 6.4 | 14.5 | 6.4 |
| 5m | 4.0 | 3.5 | 1.8 | 8.3 | 6.1 | 12.7 | 6.1 |
| 6m | 6.1 | 5.0 | 2.4 | 15.1 | 8.6 | 31.1 | 11.4 |
| 9m | 6.9 | 4.9 | 2.3 | 13.9 | 8.1 | 28.1 | 10.7 |
| 1y | 7.9 | 5.2 | 3.1 | 17.7 | 9.9 | 25.5 | 11.5 |
| 2y | 6.6 | 3.0 | 1.8 | 11.5 | 7.4 | 15.5 | 7.6 |
| 3y | 3.1 | 2.6 | 1.0 | 8.3 | 4.0 | 8.5 | 4.6 |
| 4y | 3.4 | 2.3 | 1.0 | 8.9 | 4.7 | 7.3 | 4.6 |
| Mean | 4.1 | 3.2 | 1.8 | 9.6 | 5.9 | 12.9 | |

**Table A1: *Net* Level Summary Statistics**. Top panel plots the average level of daily net dollar lending aggregated across all intermediaries in the sample, equal to the numerator of $Net^k_{t,t+n}$for a given currency $k$ and maturity $t+n$. Bottom panel plots the time-series standard deviation of $Net^k_{t,t+n}$.

| | Days to Maturity | $\ln(1+\text{Gross Notional})$ | $\mathbb{I}$(Quarter End) | $\mathbb{I}$(Year End) | $HMV_t^k$ |
|---|---|---|---|---|---|
| $Net_{t,t+n}^k$ | $0.046^{***}$ | $0.046^{***}$ | $-0.006^{**}$ | $-0.008^{***}$ | $0.007^{**}$ |
| | (0.00) | (0.00) | (0.01) | (0.00) | (0.02) |
| $N$ | 151,655 | 151,655 | 151,655 | 151,655 | 129,059 |

| | $CY_{t,t+n}^k$ | $R_t^{SPX}$ | $VIX_t$ | $Baa_t - Aaa_t$ |
|---|---|---|---|---|
| $Net_{t,t+n}^k$ | $0.020^{***}$ | $-0.002$ | $-0.017^{***}$ | $-0.012^{***}$ |
| | (0.00) | (0.38) | (0.00) | (0.00) |
| $N$ | 151,655 | 143,507 | 143,507 | 143,507 |

**Table A2: Correlations**. Table presents the correlation of $Net_{t,t+n}^k$at the day, by currency, by tenor level on days to maturity, $\ln(1+\text{Gross Notional})$, dummies equal to 1 for quarer or year ends (defined as the last week of the quarter or years) and 0 otherwise and $HMV_t^k$ which is a measure of net lending by dealers from CFTC data analogous to Hazelkorn et al. (2022)'s measure. Bottom panel compares with the convenience yield—estimated using the relevant currency's OIS rates and the Nielsen-Siegel-Svensson yield curve model—the return on the SPX, and the levels of the VIX and Baa-Aaa spread. Correlations given with $^{*}$ $p < 0.10$, $^{**}$ $p < 0.05$, $^{***}$ $p < 0.01$.

| | | Matched | | | | Rounded | | | |
|---|---|---|---|---|---|---|---|---|---|
| | | Safe Asset | | Broad Assets | | Safe Asset | | Broad Assets | |
| *Borrowing Currency* | | Unencumbered | Unencumbered & Encumbered | Unencumbered | Unencumbered & Encumbered | Unencumbered | Unencumbered & Encumbered | Unencumbered | Unencumbered & Encumbered |
| AUD | Median | 0.03 | 0.04 | 0.10 | 0.46 | 0.25 | 0.60 | 0.84 | 2.46 |
| | Mean | 0.73 | 0.67 | 1.44 | 2.47 | 6.46 | 14.76 | 31.98 | 119.42 |
| | Std. Dev. | 18.40 | 3.69 | 32.79 | 11.73 | 87.41 | 165.50 | 676.37 | 1,924.34 |
| CAD | Median | 0.01 | 0.06 | 0.03 | 0.13 | 0.17 | 0.45 | 0.64 | 1.93 |
| | Mean | 0.61 | 2.06 | 1.01 | 3.20 | 5.55 | 3.12 | 24.05 | 17.74 |
| | Std. Dev. | 18.78 | 60.50 | 26.70 | 79.98 | 73.17 | 30.69 | 438.50 | 163.84 |
| CHF | Median | 0.00 | 0.00 | 0.00 | 0.01 | 0.00 | 0.00 | 0.53 | 1.21 |
| | Mean | 0.02 | 0.00 | 0.73 | 0.68 | 9.73 | 45.01 | 34.78 | 118.54 |
| | Std. Dev. | 0.53 | 0.02 | 34.88 | 6.73 | 179.17 | 866.07 | 608.15 | 2,054.87 |
| EUR | Median | 0.16 | 0.44 | 0.29 | 0.80 | 0.28 | 0.70 | 0.64 | 1.57 |
| | Mean | 1.02 | 2.16 | 2.12 | 4.49 | 2.54 | 3.87 | 16.91 | 33.77 |
| | Std. Dev. | 11.21 | 19.84 | 21.00 | 35.41 | 36.93 | 25.71 | 705.19 | 835.99 |
| GBP | Median | 0.07 | 0.10 | 0.18 | 0.32 | 0.21 | 0.91 | 0.66 | 3.05 |
| | Mean | 0.93 | 2.85 | 2.11 | 7.58 | 3.10 | 5.55 | 13.08 | 37.85 |
| | Std. Dev. | 31.64 | 23.69 | 55.94 | 59.49 | 49.81 | 31.45 | 211.45 | 535.56 |
| JPY | Median | 0.05 | 0.14 | 0.08 | 0.18 | 0.12 | 0.25 | 0.23 | 0.41 |
| | Mean | 0.44 | 0.61 | 1.19 | 0.89 | 3.32 | 3.53 | 13.34 | 24.13 |
| | Std. Dev. | 5.46 | 5.18 | 19.53 | 7.63 | 136.16 | 57.34 | 328.15 | 666.67 |
| USD | Median | 2.01 | 8.56 | 5.38 | 24.79 | 2.19 | 7.80 | 7.25 | 28.72 |
| | Mean | 50.12 | 94.32 | 107.77 | 234.33 | 55.67 | 97.10 | 335.44 | 342.05 |
| | Std. Dev. | 1,394.56 | 1,142.77 | 2,982.31 | 3,123.05 | 4,609.31 | 1,187.54 | 60,064.85 | 5,107.52 |
| Average of All excl. USD | Median | 0.05 | 0.13 | 0.11 | 0.32 | 0.17 | 0.48 | 0.59 | 1.77 |
| | Mean | 0.63 | 1.39 | 1.43 | 3.22 | 5.12 | 12.64 | 22.36 | 58.57 |
| | Std. Dev. | 14.34 | 18.82 | 31.81 | 33.49 | 93.77 | 196.13 | 494.63 | 1,030.21 |

**Table A3: Safe Asset Ratios**. Table presents the average ratio of safe assets and broad assets, which reflect the value of assets (either unencumbered or both unencumbered and encumbered) relative to the level of net lending in the given market. Unencumbered and encumbered asset column data begins in May 2022.

| | | Matched, incl. Swaps and Forwards | | | | Matched, incl. Firm Shorts | | | |
|---|---|---|---|---|---|---|---|---|---|
| | | Safe Asset | | Broad Assets | | Safe Asset | | Broad Assets | |
| *Borrowing Currency* | | Unencumbered | Unencumbered & Encumbered | Unencumbered | Unencumbered & Encumbered | Unencumbered | Unencumbered & Encumbered | Unencumbered | Unencumbered & Encumbered |
| AUD | Median | 0.01 | 0.02 | 0.05 | 0.13 | 0.03 | 0.04 | 0.09 | 0.36 |
| | Mean | 0.52 | 1.57 | 1.25 | 4.26 | 0.72 | 0.67 | 1.18 | 2.12 |
| | Std. Dev. | 10.90 | 27.26 | 23.64 | 64.98 | 18.38 | 3.69 | 23.08 | 10.74 |
| CAD | Median | 0.01 | 0.04 | 0.01 | 0.10 | 0.01 | 0.06 | 0.03 | 0.13 |
| | Mean | 0.19 | 0.32 | 0.39 | 0.86 | 0.60 | 2.05 | 0.89 | 3.15 |
| | Std. Dev. | 1.91 | 1.59 | 3.91 | 3.94 | 18.74 | 60.50 | 25.96 | 79.98 |
| CHF | Median | 0.00 | 0.00 | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 | 0.01 |
| | Mean | 0.47 | 0.00 | 0.91 | 1.82 | 0.02 | 0.00 | 0.18 | 0.18 |
| | Std. Dev. | 40.98 | 0.00 | 45.30 | 43.28 | 0.51 | 0.02 | 5.31 | 0.95 |
| EUR | Median | 0.12 | 0.30 | 0.20 | 0.55 | 0.15 | 0.43 | 0.24 | 0.78 |
| | Mean | 1.11 | 0.70 | 2.68 | 1.47 | 0.95 | 2.08 | 1.88 | 4.20 |
| | Std. Dev. | 10.57 | 1.93 | 44.88 | 3.91 | 10.49 | 19.20 | 20.02 | 34.28 |
| GBP | Median | 0.06 | 0.11 | 0.15 | 0.29 | 0.07 | 0.10 | 0.15 | 0.31 |
| | Mean | 2.49 | 9.21 | 5.53 | 22.44 | 0.90 | 2.75 | 1.95 | 7.19 |
| | Std. Dev. | 109.26 | 226.64 | 207.99 | 512.93 | 31.62 | 23.34 | 55.74 | 57.98 |
| JPY | Median | 0.05 | 0.15 | 0.08 | 0.18 | 0.05 | 0.14 | 0.07 | 0.16 |
| | Mean | 0.80 | 0.68 | 1.16 | 0.79 | 0.40 | 0.59 | 0.81 | 0.82 |
| | Std. Dev. | 42.26 | 13.44 | 43.91 | 14.63 | 4.99 | 5.17 | 11.49 | 7.61 |
| USD | Median | 1.39 | 5.45 | 3.72 | 16.48 | 1.62 | 8.15 | 4.25 | 24.02 |
| | Mean | 51.31 | 91.14 | 108.04 | 225.04 | 44.16 | 91.63 | 77.27 | 210.52 |
| | Std. Dev. | 1,398.32 | 963.80 | 2,900.46 | 2,396.41 | 1,328.84 | 1,100.99 | 2,612.97 | 2,407.17 |
| Average of All excl. USD | Median | 0.04 | 0.10 | 0.08 | 0.21 | 0.05 | 0.13 | 0.10 | 0.29 |
| | Mean | 0.93 | 2.08 | 1.99 | 5.27 | 0.60 | 1.36 | 1.15 | 2.94 |
| | Std. Dev. | 35.98 | 45.14 | 61.60 | 107.28 | 14.12 | 18.65 | 23.60 | 31.92 |

**Table A4: Safe Asset Ratios**. Table presents the average ratio of safe assets and broad assets, which reflect the value of assets (either unencumbered or both unencumbered and encumbered) relative to the level of net lending in the given market. Unencumbered and encumbered asset column data begins in May 2022. First two columns include net forward positions in addition to net swap lending (which changes the denominator of asset ratios), and last two columns net out firm shorts from banks' net long position in the assets (which changes the numerator).

| Panel A: Safe Assets. | | | | |
|---|---|---|---|---|
| | Unencumbered Safe Assets | | Unencumbered & Encumbered Safe Assets | |
| | (1) | (2) | (3) | (4) |
| $Net^k_{t,t+n}$(\$ Level) | 0.020** | −1.149** | 0.166*** | −2.325 |
| | (2.18) | (−2.01) | (4.74) | (−1.60) |
| AUD | 265.068*** | | 180.328* | |
| | (3.33) | | (1.91) | |
| CAD | 175.109*** | | 96.310 | |
| | (3.40) | | (0.56) | |
| EUR | 1803.268*** | | 3899.425*** | |
| | (4.11) | | (3.94) | |
| GBP | 592.817*** | | 901.963 | |
| | (3.59) | | (1.36) | |
| JPY | 735.252*** | | 530.823 | |
| | (3.88) | | (1.32) | |
| Constant | −27.346* | 22557.509*** | −214.863*** | 61711.154*** |
| | (−1.71) | (5.12) | (−4.91) | (5.17) |
| $N$ | 74,997 | 76,658 | 10,145 | 8,923 |
| $R^2$ | 0.29 | 0.01 | 0.49 | 0.01 |
| Time FE | Yes | Yes | Yes | Yes |
| Sample | Non-USD | USD | Non-USD | USD |
| Panel B: Broad Assets. | | | | |
| | Unencumbered Broad Assets | | Unencumbered & Encumbered Broad Assets | |
| | (1) | (2) | (3) | (4) |
| $Net^k_{t,t+n}$(\$ Level) | 0.187*** | −4.749* | 0.303*** | −5.599* |
| | (3.87) | (−1.76) | (3.72) | (−1.78) |
| AUD | 822.035** | | 986.092** | |
| | (2.20) | | (2.59) | |
| CAD | 239.727 | | 369.490 | |
| | (0.79) | | (0.73) | |
| EUR | 4492.373*** | | 8279.851*** | |
| | (3.16) | | (3.31) | |
| GBP | 2040.245* | | 2593.438 | |
| | (1.83) | | (1.65) | |
| JPY | 18.434 | | 66.421 | |
| | (0.03) | | (0.08) | |
| Constant | −117.260 | 126810.843*** | −264.553* | 145684.497*** |
| | (−1.10) | (5.37) | (−1.97) | (5.40) |
| $N$ | 10,099 | 8,885 | 10,099 | 8,885 |
| $R^2$ | 0.38 | 0.01 | 0.41 | 0.01 |
| Time FE | Yes | Yes | Yes | Yes |
| Sample | USD | Non-USD | Non-USD | USD |

**Table A5: Net vs. Safe Asset Ratios**. Table presents the regression of the level of $Net^k_{t,t+n}$on the level of the assets—either HQLAs or broad assets—with the same maturity and currency in millions of dollars: $\textit{Asset (Level)}^k_{t,t+n} = \alpha + \beta Net^k_{t,t+n}(\text{Level}) + \gamma^k + \varepsilon^k_{t,t+n}$ where $\gamma^k$ is a currency fixed effect. Columns 1 and 3 limit the sample to observations with net dollar lending (e.g., $Net^k_{t,t+n} > 0$). Columns 2 and 4 limit the sample to observations with net dollar borrowing (e.g., $Net^k_{t,t+n} < 0$). Base level is CHF. Within $R^2$ reported. $t$-statistics shown using robust standard errors clustered by market and date where * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

| | (1) |
|---|---|
| Loan Demand $\text{HHI}^k_{t,t+n}$ | 0.0704*** |
| | (10.76) |
| Bilateral $\text{Share}^k_{t,t+n}$ | 2165.9*** |
| | (22.47) |
| Constant | 3640.6*** |
| | (38.99) |
| $N$ | 18,752 |
| $R^2$ | 0.06 |

**Table A6: Demand HHI estimate for full sample**. Table presents the regression of Demand $\text{HHI}^k_{t,t+n} = \alpha + \beta_1 \text{Loan Demand HHI}^k_{t,t+n} + \beta_2 \text{Bilateral Share}^k_{t,t+n} + \varepsilon^k_{t,t+n}$, where Loan Demand $\text{HHI}^k_{t,t+n}$ is the unsecured and secured loan counterparty HHI calculated over the full sample, 2016 to 2023, and bilateral share is the share of bilateral FX transactions compared to the total FX transactions in that market.

| | Unencumbered Assets | | Unencumbered & Encumbered Assets | |
|---|---|---|---|---|
| | (1) | (2) | (3) | (4) |
| *Supply Segmentation* | | | | |
| Supply $\text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | $21.73^{***}$ | $21.67^{***}$ | $21.71^{***}$ | $21.67^{***}$ |
| | (4.93) | (4.94) | (4.93) | (4.94) |
| Supply $\text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | $20.29^{***}$ | $20.27^{***}$ | $20.29^{***}$ | $20.27^{***}$ |
| | (5.63) | (5.63) | (5.63) | (5.63) |
| *Safe Asset Scarcity* | | | | |
| $\text{Net}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | 1.65 | 1.48 | 1.58 | 1.45 |
| | (0.35) | (0.31) | (0.33) | (0.31) |
| $\text{Net}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | 3.06 | 3.05 | 3.06 | 3.05 |
| | (0.96) | (0.96) | (0.96) | (0.96) |
| Safe Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | $-27.02^{*}$ | | $-18.42^{**}$ | |
| | $(-1.76)$ | | $(-2.49)$ | |
| Safe Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | $-0.37$ | | $-0.20$ | |
| | $(-1.40)$ | | $(-1.36)$ | |
| Broad Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | | $-57.00^{***}$ | | $-42.22^{***}$ |
| | | $(-2.64)$ | | $(-3.38)$ |
| Broad Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | | $-0.19$ | | $-0.17$ |
| | | $(-1.29)$ | | $(-1.30)$ |
| *Demand Concentration* | | | | |
| Demand $\text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | 4.54 | 4.56 | 4.55 | 4.57 |
| | (1.39) | (1.40) | (1.39) | (1.40) |
| Demand $\text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | 2.72 | 2.72 | 2.72 | 2.72 |
| | (0.85) | (0.85) | (0.85) | (0.85) |
| *Controls* | | | | |
| Bank $\text{CDS}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | $-84.28^{***}$ | $-83.97^{***}$ | $-84.22^{***}$ | $-83.99^{***}$ |
| | $(-2.85)$ | $(-2.84)$ | $(-2.85)$ | $(-2.84)$ |
| Bank $\text{CDS}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | $-88.45^{***}$ | $-88.21^{***}$ | $-88.42^{***}$ | $-88.23^{***}$ |
| | $(-3.01)$ | $(-3.00)$ | $(-3.01)$ | $(-3.00)$ |
| Govt $\text{CDS}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | $-1.41$ | $-1.42$ | $-1.40$ | $-1.41$ |
| | $(-1.50)$ | $(-1.52)$ | $(-1.50)$ | $(-1.51)$ |
| Govt $\text{CDS}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | $-3.94$ | $-3.98$ | $-3.96$ | $-3.99$ |
| | $(-1.30)$ | $(-1.31)$ | $(-1.30)$ | $(-1.31)$ |
| $N$ | 18,755 | 18,755 | 18,755 | 18,755 |
| $R^2$ | 0.28 | 0.28 | 0.28 | 0.28 |
| Tenor FE | Yes | Yes | Yes | Yes |
| Time FE | Yes | Yes | Yes | Yes |
| Weighted | Yes | Yes | Yes | Yes |

**Table A7: Regression of the absolute value of the basis on measures of frictions**. Table presents the regression described in section 4.5 using the sample where our demand concentration measure is available. The dependent variable is the absolute value of the basis. To make the coefficients directly comparable, we transform the independent variables to modified $z$-scores each variable's full sample median and standard deviation. Regression includes tenor and date fixed effects and we weight the regression by the square root of the market's share of the total daily gross notional. Within $R^2$ reported. $t$-statistics shown using robust standard errors clustered by market and date where $^{*}$ $p < 0.10$, $^{**}$ $p < 0.05$, $^{***}$ $p < 0.01$.

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|---|---|---|---|---|---|
| *Supply Segmentation* | | | | | | |
| Supply $\text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | 18.27*** | | | | | |
| | (4.55) | | | | | |
| Supply $\text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | 17.18*** | | | | | |
| | (4.87) | | | | | |
| *Safe Asset Scarcity* | | | | | | |
| $\text{Net}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | | 8.13** | | | | |
| | | (2.28) | | | | |
| $\text{Net}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | | 4.29** | | | | |
| | | (2.07) | | | | |
| Safe Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | | | −17.35** | | | |
| | | | (−2.04) | | | |
| Safe Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | | | −0.60 | | | |
| | | | (−1.55) | | | |
| Broad Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | | | | −23.29* | | |
| | | | | (−1.91) | | |
| Broad Asset $\text{Ratio}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | | | | −0.36 | | |
| | | | | (−1.23) | | |
| *Demand Concentration* | | | | | | |
| Demand $\text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | | | | | −0.25 | |
| | | | | | (−0.07) | |
| Demand $\text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | | | | | −0.57 | |
| | | | | | (−0.16) | |
| $\widehat{\text{Demand}}\ \text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} \geq 0)$ | | | | | | 12.19*** |
| | | | | | | (3.07) |
| $\widehat{\text{Demand}}\ \text{HHI}^k_{t,t+n} \times \mathbb{I}(Net^k_{t,t+n} < 0)$ | | | | | | 3.40 |
| | | | | | | (1.58) |
| $N$ | 149,337 | 149,337 | 149,337 | 149,337 | 18,755 | 149,293 |
| $R^2$ | 0.08 | 0.02 | 0.00 | 0.00 | 0.00 | 0.04 |
| Tenor FE | Yes | Yes | Yes | Yes | Yes | Yes |
| Time FE | Yes | Yes | Yes | Yes | Yes | Yes |
| Weighted | Yes | Yes | Yes | Yes | Yes | Yes |

**Table A8: Regression of the absolute value of the basis on marginal cost measures**. Table presents the regression described in section 4.5. The dependent variable is the absolute value of the basis. To make the coefficients directly comparable, we transform the independent variables to modified $z$-scores each variable's full sample median and standard deviation. Regression includes tenor and date fixed effects and we weight the regression by the square root of the market's share of the total daily gross notional. Within $R^2$ reported. $t$-statistics shown using robust standard errors clustered by market and date where $^*$ $p < 0.10$, $^{**}$ $p < 0.05$, $^{***}$ $p < 0.01$.