China and the World Bank – How Contrasting Development Approaches affect the Stability of African States

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China's development model challenges the approaches of traditional Western donors like the World Bank, and many observers fear Chinese aid poses a threat to stability in recipient countries. We argue that as the WB, China has an interest in avoiding outright conflict, but differs in the norms propagated to achieve stability. We consider a broad range of georeferenced measures of stability at the subnational level in Africa, ranging from outright conflict to protests and citizens' attitudes. Using fixed effects and IV estimations, we find that aid by both the WB and China does not increase outright conflict, on average. Both even reduce one-sided lethal government violence against civilians and do not increase any type of citizen protests. Still, Chinese aid is associated with more government repression and an increased acceptance of authoritarian norms, while WB projects strengthen democratic values.

Keywords: Development Models, Development Aid, Stability, Conflict, Repression, World Bank, China, Africa, Geolocation
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1 Introduction

As part of what The Economist describes as the "The new scramble for Africa", emerging donors, in particular China (Dreher and Fuchs, 2015; Dreher et al., 2018), are challenging the predominance of traditional donors in affecting African development. The big question is whether this time, as the magazine asks, African countries will be the benefactors of foreign engagement? Africa is of particular interest as a central focus of traditional donors as well as a key target region for China. While there was a huge drop in global poverty rates thanks to rapid growth mostly in Asian countries, many African states lag behind. In particular conflict-prone states plagued by re-igniting battles pose a challenge, which is why those are also labeled as the "new frontier of development".[1]

While the literature on the growth effects of aid converges towards an on average small, positive effect (Clemens et al., 2011; Dreher and Langlotz, 2019; Galiani et al., 2017; Kilby, 2015), researchers are divided about the impact of aid on stability and conflict (e.g., Bluhm et al., 2016; Child, 2018; Crost et al., 2014; Nunn and Qian, 2014). Some perceive Chinese aid as a crucial step forward that brings growth and stability to Africa, while others regard it as a big risk that narrowly focuses on Chinese self-interest, enriches elites (Dreher et al., 2019), fosters conflict, and exports repression, along with surveillance tactics and autocratic norms (Kishi and Raleigh, 2016).[2] We are shedding light on this crucial question by systematically contrasting the Chinese approach to development with that of the World Bank (WB), one of the most important traditional donors, and analyzing their effect on stability in recipient regions.

Moving beyond the partly subjective public rhetoric, we argue that Chinese foreign aid needs to be considered with all its nuances. China's "no strings attached"approach to development differs strongly from the expert-driven, conditional approach of traditional donors like the WB and many other Western DAC donors. Both are interested in growth, but while the World Bank regards democracy, transparency and humanitarian rights as a critical part of that, the Chinese model highlights social and political stability as the key ingredient to prosperity. To comprehensively understand the impact of these approaches, requires a holistic definition of stability. This paper defines stability as a broad continuum ranging from outright conflicts with at least a certain number of battle-related deaths, to lower level-conflict events like citizen protests and government repression, as well as attitudes related to stability.

Although the more economic growth and stability-oriented perspective of China and the rule- and expert-based democratic perspective of the World Bank might be seen as two ends of the

[1] , The Economist (2017), accessed 30.01.2019.

[2] See Washington Post Monkey Cage and Council on Foreign Relations for the direct citation, accessed 31.04.2019. See Freedom House on East African states adapting Chinese internet censorship policies, The New York Times on exporting the surveillance state, and US News about China's web surveillance model, and Quartz about Zimbabwe using Chinese large-scale facial recognition software. See The Economist about China training foreign officials and bureaucrats, and promoting "their political model" as an alternative to democracy. All accessed 31.04.2019.

spectrum of development policies, their impact on stability is complex. Even if its motive would be mere self-interest, China also has an incentive to protect its investments as well as its workers in Africa, and one should not expect China to turn a blind eye on recipient governments starting outright conflicts. Both donors will try to stop recipient governments from engaging in conflicts that they deem avoidable or unnecessary, and given their size have some leeway over recipient governments.[3] At the same time, when regarding stability more broadly than just focusing on the outright conflicts, China is likely to build on its own domestic development experience, which combines growth with an autocratic and stability-oriented rule. Therefore, there are good reasons to believe that China would be more willing to accept recipient governments' use of autocratic policies and non-lethal repression to enhance stability, while the WB emphasizes democracy and humanitarian values more strongly. This paper does not take a normative stance which perspective is ultimately superior from the perspective of a developing country, but carefully carves out the most important conceptual differences between the two donors and their potential effect on state stability.

To then investigate the causal effects of the two donors on stability, we precisely link new detailed geo-referenced datasets on development projects by China (Strange et al., 2017) and the WB (Dreher et al., 2017) with geo-referenced measures of stability at the sub-national level in Africa. Our dataset allows us to match the location of aid projects and conflicts more precisely than earlier studies, and enables us to flexibly eliminate potential biases arising from, for instance, unobserved conflict trends, region-specific time-invariant factors, and country-level time-varying factors. Moreover, our identification strategy adapts an established instrumental variable (IV) approach by Nunn and Qian (2014). Our instrument is the interaction between the pre-determined probability of a region to receive aid with exogenous temporal variation in the WB's IDA liquidity (Dreher et al., 2017) or in Chinese domestic (over-)production of commodities (Bluhm et al., 2018).

Our results provide several important insights. First, a wide range of fixed effects (FE) and IV specifications reject the idea that aid by either donor, on average, fuels conflict at the sub-national level. In the FE specification, a one standard deviation change in WB aid decreases the conflict likelihood by about 1.6 percentage points. The effect remains negative and of similar magnitude, but becomes statistically insignificant when using IV. When studying China, both strategies yield negative and small, insignificant coefficients.

Moving beyond this main IV effect, we show that when considering the actors involved in and responsible for a conflict, both WB and Chinese engagement consistently leads to a reduction in

[3] Based on the different approaches, the distribution of aid also differs. Chinese aid goes more often directly to governments and the home regions of ethnic leaders (Dreher et al., 2019), and its distribution is associated with more corruption (Isaksson and Kotsadam, 2018a) and weaker labor unions (Isaksson and Kotsadam, 2018b). Still, its less bureaucratic approach was also found to lead to more evenly distributed economic activity within countries (Bluhm et al., 2018).

lethal violence by governments against civilians. For both donors, we also find no positive effect on protest events like demonstrations, riots, and strikes. At the same time, there are crucial differences with regard to how stability is secured. Among the two, only Chinese aid is associated with more government repression in recipient regions. Afrobarometer responses suggest that both donors have different effects on measures of security, democratic norms and attitudes, as well as on perceptions of government behavior. While WB aid is linked to a higher perceived security and stronger support for democratic values, Chinese aid tends to result in a stronger emphasis on rule following behavior and a higher acceptance of autocratic regimes.

This paper contributes in several ways to better understand the role of donors in influencing recipient country stability, as well as the channels and mechanisms linking aid to various types of conflict. We combine the strengths of existing approaches on the country level (e.g., Bluhm et al., 2016; Nielsen et al., 2011; Nunn and Qian, 2014), with the advantages of studies focusing on sub-national aid data in specific sectors in selected countries (e.g., Berman et al., 2011; Child, 2018; Crost et al., 2016; Sexton, 2016; Van Weezel, 2015). The aim is to deliver the best possible compromise between using micro-data with causal identification strategies and estimating externally valid results for more than one country. Truly randomly allocated aid projects in individual countries of course possess a higher internal validity. Still, their results could be driven by the specific country- or aid-type context, and it is impossible to replicate them at large scale for a full continent. We apply identification strategies that are well established in the literature, and by considering a large set of all aid-eligible countries our results can be meaningfully interpreted beyond the context of an individual country.

Besides using new data and providing more precise estimates about the causal effect of aid on more comprehensive measures of stability, we want to emphasize three main contributions. First, our paper adds to the scarce evidence on the incentives for different actors and their choices created by development projects. Crost et al. (2014), for instance, focus on how aid changes the incentives for rebel groups.[4] Our finding of a significant reduction in lethal violence enacted by recipient governments against civilians supports the idea of the "cost of shame" (Lebovic and Voeten, 2009), that the fear of loosing aid money notably changes the incentives of recipient.

Second, we shed some light about the hopes and fears associated with emerging donors (Asmus et al., 2017a; Fuchs and Vadlamannati, 2013a). In particular China's increased global engagement, like the Belt and Road initiative and the intense China-Africa Cooperation, was one of the most crucial geopolitical changes in the last two decades, which will continue to create tensions in the future. Existing papers have either focused on outright conflict, or on the impact of, for instance, Chinese aid in Latin America on attitudes towards China (**?**). Still, convincing causal evidence

[4] As a robustness test, we also show results on sectoral differences, which augment previous results on intersectoral differences within specific countries (e.g., Child, 2018; Crost et al., 2016; Berman et al., 2011).

that provides a comprehensive picture of the impact of Chinese aid on stability in a broad sense was missing.

Third, by contrasting Chinese aid with the World Bank as a prototypical example of a traditional, multilateral donor that involves development experts and accounts for democracy and humanitarian values, we provide a useful reference point. Western newspapers and NGOs have widely complained about the active export of Chinese surveillance technology and policies, as well as about the potentially detrimental impact that the apparent success of the Chinese approach to development has on developing countries. Our results paint a more nuanced picture. China's engagement is not associated with an increase in outright conflict, and even with more stability when considering less lethal conflicts by governments against civilians. Still, it comes along with increased government repression and a higher prevalence of autocratic norms. Hence, the approach at least in some dimensions makes a difference. How to assess these differences depends on the normative perspective of the observer.

The paper proceeds as follows. Section 2 summarizes the existing literature and how the two approaches are linked to different measures of stability. Section 3 explains the data and the corresponding sources and provides descriptive statistics. Section 4 presents the specification and empirical strategy. Section 5 shows and discusses the results, and Section 6 concludes.

2 Theoretical considerations and existing literature

This section first defines our concept of stability, and then contrasts the policies by China and the WB, and how they might affect outright conflict as well as individual dimensions of stability. To really understand the impact of the two approaches to development in a comprehensive way requires a holistic definition of stability. More specifically, we think of outright conflict as a lethal fight that caused at least a specific number of battle related deaths. Besides the average effects on conflict, we distinguish between the actors involved in a conflict, either government related or non-state actors like rebels. Each of those can be involved in a two-sided conflict with the respective other group, or start a one-sided conflict against civilians. Besides these lethal conflict events, lower-level conflict like citizens protests against governments, and government policies against potential rebels or other minority groups in the country also characterize stability. Finally, attitudes are both reflecting the results of these other events, but are also themselves signs of stability, for instance attitudes about the quality of democratic processes or rule-following behavior.

Outright conflict: Aid can lower conflict if it raises income and hence the opportunity costs on fighting. In that regard, the aid effectiveness literature converges towards either a null effect (Doucouliagos and Paldam, 2009), or small positive effects (Galiani et al., 2017) of aid on growth.

Berman et al. (2013) hypothesize that projects are more successful in reducing violence if they require the integration of development experts, and Minasyan et al. (2017) demonstrate the importance of donor quality. The WB built up a large expertise over the decades since its foundation, which might increase the effectiveness of its projects in raising income.

At the same time, traditional donors have also been criticized for a lack of "ownership" and underutilizing local knowledge in recipient countries. Scholars writing about emerging donors like India (Fuchs and Vadlamannati, 2013b) and China (Humphrey and Michaelowa, 2018) also highlight less complicated, bureaucratic processes with quicker implementation times. Hence, China's flexibility and emphasis on economic "mutual benefits" may boost growth even more than the WB approach (Dreher et al., 2017). Thus, in this dimensions both donors could reduce the incentives to engage in conflict by fostering growth in their own ways.[5]

Besides growth effects, the distribution of potential gains is important as the literature on resource-related income shocks highlights (e.g. Berman et al., 2017; Dube and Vargas, 2013; Gehring et al., 2018). Whether potential gains from aid are used for short-term consumption, invested in fostering development, or end up in the foreign bank accounts of government officials, affects the impact on conflict. If the projects contribute to rising inequality this could trigger conflicts. WB projects were found to be less politically motivated than other types of aid (e.g., Dreher et al., 2009), and the Bank aims for aid allocation in line with conflict prevention policies accounting for humanitarian aspects and security. In contrast, Dreher et al. (2019) find that Chinese projects in Africa are more likely to benefit the birth regions of the respective leader. Isaksson and Kotsadam (2018a) suggest that Chinese engagement is associated with higher local corruption, which could increase inequality, and with lower trade union membership (Isaksson and Kotsadam, 2018b), which could decrease the labor share of profits. However, in particular Chinese infrastructure projects are found to lead to a more equal distribution of economic activity (Bluhm et al., 2018). Hence, the theoretical predictions are to some degree contradictory, leaving this an empirical question.

Finally, traditional Western donors often impose conditions and require certain processes in aid-receiving countries. The WB often uses conditions regarding governance, equality, anti-discrimination, among others. The Bank is also considered to be a global leader in "conflict-sensitive programming" (Van der Windt and Humphreys, 2016; World Bank, 2011). This involves the identification of conflict escalators using a detailed Conflict Analysis Framework (CAF) (Wam, 2006) to help

[5] The literature also describes "aid as a price" that can be acquired as a result of winning a fight or conflict. This "aid as a price" theory has both a direct goods-related and a political dimension. Regarding goods, Nunn and Qian (2014) show that US food aid leads to more conflict, as it can be looted and sold on black markets. Expensive equipment associated with investments in healthcare and communication infrastructure can also be sold on black markets. To remedy these issues, some traditional donors like the WB seek to "conflict-proof" their aid by avoiding projects that provide lootable/fungible resources over which warring parties might fight and instead provide aid in a more discrete manner, such as social programs (Berman et al., 2013; Crost et al., 2014; Imai et al., 2018). We investigate aid in different sectors separately in a robustness test.

WB staff to understand country-specific sources of conflicts. The WB's Operational Procedures instruct WB staff how to act within a conflict-affected country (World Bank, 2001; Bannon, 2010).

Officially, Chinese aid has less strings attached.[6] Still, even for skeptical observers who assume China is largely interested in securing resources and providing employment for Chinese workers, it is implausible that China would welcome recipient governments engaging in unprovoked, avoidable conflicts. This would endanger existing investments and the health not only of African, but also the large number of Chinese workers in Africa (officially 227,407 by 2016). Moreover, stability is a crucial part of the Chinese development model; in a speech at the 2008 central party congress Hu Jintao mentioned the word stability alone 21 times (freedom did not appear a single time).[7]

To sum up, while politicians, newspapers and some scholars raise concerns about specific aspects of Chinese aid that could give rise to more conflict than the rule-driven approach of the WB, we argue that the net impact on outright conflict is less straightforward. There are reasons to expect WB aid could be more successful in lowering the average risk of conflict, but based on its self-interest and emphasis on stability, China has incentives to "unofficially" set conditions to avoid instability.[8]

Actors and types of conflict: Two-sided conflict between governments and rebels could be affected differently by aid than one-sided conflict against civilians. Generally, neither donor should be interested in outright conflict. They can threaten to withhold future aid payments to prevent recipients from engaging in conflicts they deem harmful and unnecessary. Lebovic and Voeten (2009) label this the "cost of shame". We argue that this threat is less likely to matter for two-sided conflicts as it is much easier for both sides to justify their actions as a necessary reaction to the other side. One-sided conflict actions against civilians, in contrast, are harder to justify. If governments use excessive violence against citizens, public pressure in donor countries can stop in particular international organizations like the WB from aid payments (Tir and Karreth, 2018). At the same time, China as an autocratic one-party state where decision-making is less constrained could be able to more credibly threaten to cut aid payments. Generally, both donors have the incentives and means to exert pressure on recipient governments to avoid unprovoked one-sided conflicts, while their direct influence on rebel groups is limited.[9]

[6] Anthony Germain on CBC, "China in Africa: No strings attached," last accessed 31.01.2019.
[7] Anthony Germain on BJ Review last accessed 31.06.2019.
[8] The Guardian also postulates that "Chinese aid to Africa is going to come with all sorts of strings attached, despite the "no-conditionality" rhetoric (The Guardian: "The west has no right to criticise the China-Africa relationship" last accessed 31.01.2019.)
[9] Donors might also encourage rebels to fight an opposed regime as in the case of covert aid to Angolan UNITA under president Reagan (Lagon, 1992). Our data cover almost exclusively projects implemented in accordance with the government, so this aspect should be of lesser importance.

Lower level conflict- Government policies and protests: Besides outright, lethal, conflict, we are also interested in lower-level types of conflict that are not necessarily directly causing a significant number of casualties. Lower-level conflict has two dimensions. Protest events like strikes, riots or demonstrations can be understood as bottom-up actions by citizens against governments, whereas government policies like repression are top-down measures by governments to avoid conflicts and such protests. The latter affects the costs of the former, and empirically, we are only able to observe the equilibrium outcome of both dimensions.

WB and Chinese aid can affect the reasons as well as costs of protest by fostering state capacity. On the one hand side, infrastructure projects like highways, bridges, railroads, and ports strengthen the capacity of the state by extending the spatial reach of its monopoly. Agents of the state – e.g. police officers, judges, and tax collectors – can use their increased capacity in different ways. If they wield it to impartially enforce the rule of law, levy taxes, and deliver public services, improvements in capacity and legitimacy may result in a "virtuous circle" of better state capability (Levi et al., 2009), conflict reduction (Berman et al., 2011) and less reasons to protest. On the other hand, if state agents exploit their increased capacity to enrich themselves, favor some groups over others, or weaken political opponents (Wig and Tollefsen, 2016), this can trigger protests.[10]

The WB uses an independent "Inspection Panel" to investigate complaints about human rights abuses or local conflict provoked by the WB (Zvogbo and Graham, 2018). It pursues an approach to actively build trust and social cohesion in post-conflict and conflict-affected countries (Bannon, 2010). This approach includes, for example, projects with a focus on community-driven development, and capacity building with regards to accountability and public service delivery. The Kecamatan Development program in Indonesia, for instance, attempted to reduce protests via transparency through a particularly participatory approach (Gibson and Woolcock, 2005; Barron et al., 2011).[11] To the best of our knowledge, China does not have an analogous set of policies, institutions, or operational tools in place to encourage conflict-sensitive development programming.[12] Hence, all else equal, Chinese projects could be related to more protests.

However, the equilibrium impact of both donors is more complex. Citizens deciding whether to engage in protests also weigh reasons for against the costs of these actions. WB policies that foster democratic participation and transparency might actually be linked, all else equal, to a higher likelihood to protest. Better informed citizens in more democratic states where the political costs of protesting and the fears of its consequences are lower, might be more willing to engage themselves

[10] For instance, insurgents might sabotage projects if they would not benefit sufficiently and government success weakens their support in the population (Crost et al., 2014).

[11] This community driven development approach inspired also the National Solidarity Program - a large scale development program, which was evaluated to increase governmental support in conflict-ridden Afghanistan (Beath et al., 2016).

[12] China only established its first specialized aid Agency CIDCA with a centralized evaluation mandate in 2018. Heiner Janus on DIE, "Next Steps for China's New Development Agency," last accessed 22.02.2019.

politically.

At the same time, Chinese aid could increase the costs of organizing protests, as it decreases trade union membership (Isaksson and Kotsadam, 2018b). Moreover, the enhancement of state capacity also affects the ability to handle protests. China emphasizes social stability as part of its growth model domestically, including the use of force to constrain opposition forces or protesters. Such repression can incite anger and unrest, but also enhance stability via a deterrence effect. An article, for instance, describes how "Chinese officials use advances in facial recognition technology and big data to identify potential troublemakers and reduce the risk of large-scale public demonstrations."[13]

The country is also accused of financially supporting repressive governments in Africa and exporting such repression to recipient countries (Kishi and Raleigh, 2016). For example, Uganda could turn to China after Western donors protested against strict "anti-gay" laws in the country.[14] Several reports describe how China exports its approaches regarding surveillance and censorship. One describes how China "propagate its model abroad by conducting large-scale training of foreign officials" of 36 mostly developing countries.[15] Many of those like Angola, Ethiopia, The Gambia, Kenya, Libya, Morocco, Nigeria, Rwanda, South Africa, Sudan, Zambia, Zimbabwe are in Africa. Another article describes how Uganda and Tanzania introduced cybersecurity laws that are similar to Chinese law after attending training sessions.[16] Freedom house emphasizes how Chinese support helps governments in Sub-Saharan Africa to censor the internet and social media.[17]

China's projects might thus provide more reasons to protest, but repressive policies raise the cost of protests. WB policies might provoke less protests due to the implemented safeguards, but stronger democratic standards and less fear of expressing opinions in public make protests more likely. Hence, we expect an increase in repressive government policies related to Chinese aid, but the equilibrium impact of both donors regarding protests remains an empirical question.

Attitudes: The Chinese government regards stability as central for development, and portrays itself as a "rock of stability". [18] However, it does not regard democracy, democratic participation or equal democratic rights as necessary to achieve stability, or sometimes even sees them as an obstacle to that. Dagong, a Chinese rating agency, writes that "centralized political power enabled [East-Asian countries] to concentrate on solving the most urgent issues in the economic reform

[13] See Nikkei.com, accessed 31.04.2019.
[14] Washington Post, "When China gives aid to African governments, they become more violent," last accessed 31.01.2019.
[15] See US News, accessed 31.04.2019.
[16] See Nextgov.com, accessed 31.04.2019.
[17] See Freedom House, accessed 31.04.2019. In addition to training, China reportedly exports surveillance technology like cameras, but also advanced artificial intelligence technology. For instance, China signed an agreement with Zimbabwe, Angola, and Ethiopia to deploy a new facial recognition software to monitor their population.
[18] The Economist last accessed 31.01.2019.

step by step", while "countries copying the western system encountered many political obstacles in maintaining stability".[19] A Chinese scholar describes the common perception that developing countries experiencing "chaotic" democratization "inevitably plunge into a chaotic situation marked by soaring prices, shortage of essential supplies, frequent violent conflicts and a precarious state of life and property." This also entails that "the ability to establish and maintain an effective internal order [...] is the most important of all national capacities", with higher priority than "democratic accountability in a country's political development process".[20]

China is keen on spreading its development model and emphasizing its advantages. Chinese development projects sometimes contain conditions on broadcasting Chinese radio or TV broadcasts in exchange for financial support. For instance, a radio station set up in Kenya reserves a specified amount of hours to promote Chinese culture and values, China supplies text books for schools in Liberia, Ghana and Tanzania, and organizes cultural events in South Africa. Cultural centers aim at spreading Chinese culture and values. Note that this is not good or bad per se; Western donors and the WB are engaging in the same efforts to spread the values and norms they want to propagate. The motivation to do so might be mere self-interest or the honest conviction that the respective development model is the best to raise developing countries out of poverty. Empirically, we are interested whether WB projects are related to more positive perceptions of democracy and governance, and if citizens in regions receiving Chinese aid are more likely to accept autocratic, strong states and strict rules to achieve prosperity.

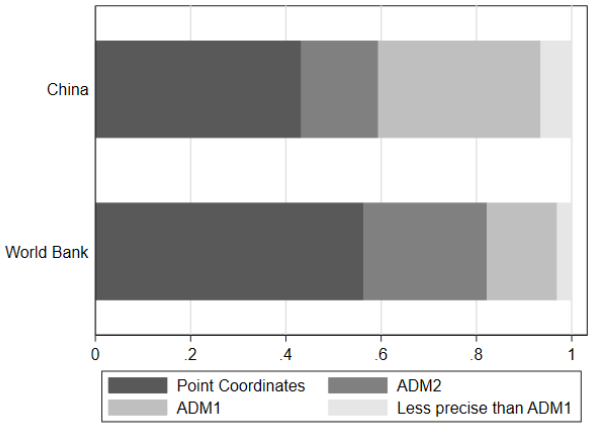
3 Data

3.1 Aid Data: World Bank and China

We focus our analysis on Africa, considering all countries with more than one million inhabitants on the OECD's DAC recipient list in 1995, the initial year of our sample period. Our unit of observation is the country-region-year, with region as the unit of analysis referring to the first level of sub-national administrative division (ADM1: "provinces", "states", or "regions") (data from Hijmans et al., 2010). This level is the most suitable choice, as it allows us to distinguish between considerable sub-national variation, while still capturing over 90% of the overall spending by China and the WB (see Figure 1).[21] Moreover, this administrative level is also highly relevant for aid allocation, as many projects are assigned to specific regions, and the regional governments can influence how, or where, to spend the funds.

[19] QZ.com last accessed 31.01.2019.
[20] CGTN last accessed 31.01.2019.
[21] Lower level administrative regions (ADM2) would only capture between 60 and 80%. Using smaller grid cells would require only relying on projects with exact data on latitude and longitude, which is only about 50% for the WB and less than 50% for China.

Figure 1: Disbursement/Commitment Amounts by Precision Codes



Our approach to assigning aid projects to regions is the following. Precisely georeferenced projects, as well as projects where we possess information about the ADM2 and ADM1 regions, are assigned to the respective ADM1 region. To cope with the fact that most projects have several project locations, we assume that aid is distributed equally across locations, following Dreher and Lohmann (2015). This means that for a project implemented in 10 locations, with four locations in region A and six in region B, 40% of the project volume would be assigned to A and 60% to B. The data appendix provides more details about the procedure.[22]

WB aid disbursements are from AidData (Strandow et al., 2011), covering the 1995 to 2012 period. We focus on disbursements by IDA, the WB's arm for development aid (qualified as financial support with a significant concessionary component). Disbursements that we can assign to the ADM1 level sum up to USD 29.4 billion, distributed over 1,472 projects in 25,041 locations in Africa. For China, we use the media-based data set on Chinese ODA-like commitments from (Dreher et al., 2019), which were georeferenced by Strange et al. (2017). Since graduating from IDA eligibility in 1999 (Galiani et al., 2017), China's overseas portfolio of grants, loans, and export credits has rapidly expanded as part of its 'Going Out' strategy. In Africa, the total volume of Chinese official financing was roughly comparable to U.S. official financing between 2000 and 2012 (Strange et al., 2017). In total, Chinese aid amounts to USD 13.2 bn, from 333 projects in 1,308

[22] Hence, our aid attribution formula is: $Aid_{pijt} = \frac{Aid_{pit}}{\int Locations_{pi}} * \int Locations_{pj}$, where p is the project, i is the country, j is the region, and t is the period for which we estimate the allocation shares. For robustness, Tables A 53 and 54 display the main results using population weights. For instance, if a project has project locations in two regions of a country, two million inhabitants reside in region A, and three million reside in region B, 40% of project funds are allocated to region A and 60% to region B. Here, the aid attribution formula is $Aid_{pijt} = \frac{Aid_{pit}}{\int Population_{pi}} * Population_{pj}$. Our population data build on the gridded population data provided by the Center for International Earth Science Information Network (CIESIN) Columbia University (2016). As global population censuses have to build on strong assumptions and yearly data has to be imputed, the data is subject to a certain degree of measurement error. The remaining project locations, with less precise location information, are mostly non-geocoded aid accruing directly to the government, which we assign to the capital region in a robustness test when considering potential spill-overs. We show results using the ADM2 regions as a robustness test in the appendix, and incorporate ethnic group homelands by intersecting those with the administrative regions.

locations.

Table 1: Donor Comparison: WB vs. China

	WB Aid	Chinese Aid
Total Disbursements/Commitments (USD):	29.4bn	13.2bn
Active in No. of Countries:	35	41
Number of Projects:	1,472	333
Number of Locations:	25,041	1,308
Mean Number of Locations per Project	17	4
Mean per Project (USD):	19.97m	39.63m
Mean per Location (USD):	1.17m	10.09m

Notes: Aid is measured in constant 2011 USD.

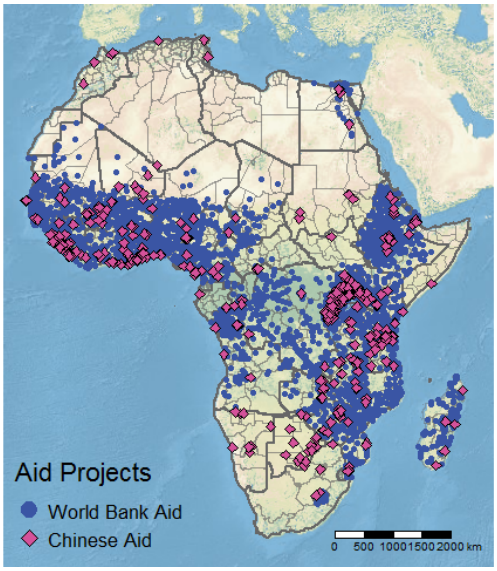
The remaining sample comprises 728 ADM1 regions in 45 countries. Table 1 shows a comparison of the two donors in some important dimensions. Both the WB and China are active in most African countries, with significant overlap in their presence between countries; the WB is active in 35 countries, and China is active in 41 countries (Humphrey and Michaelowa, 2018). While information for aid disbursements by the WB's IDA is available from 1995 to 2012, information on Chinese aid commitments in Africa is constrained to the years 2000 to 2012.[23] One interesting difference is that the WB actually finances a larger number of projects than China; these projects are present in more locations across countries. China finances fewer projects, however they are larger in scope. Accordingly, China spends nearly twice as much per project, and nearly ten times as much per project location.[24]

3.2 Conflict Data

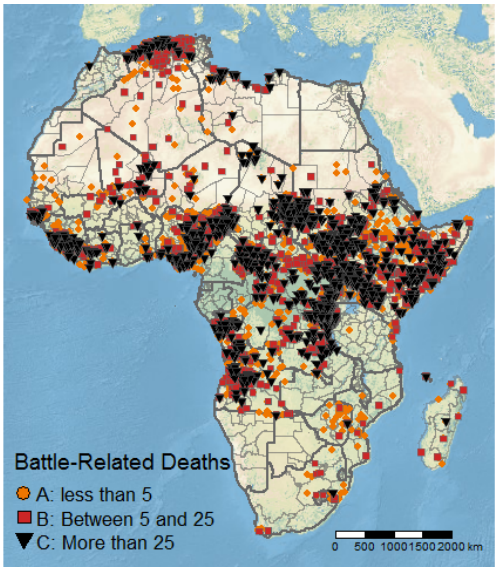
Our main specification defines conflict based on the number of battle-related deaths (BRD) from the Uppsala Conflict Data Program's (UCDP) Georeferenced Event Dataset (GED) (Croicu and Sundberg, 2015). Derived from media and NGO reports, as well as secondary sources, namely field reports and books, GED provides the most reliable and comprehensive data on incidences of

[23] This analysis focuses on Official Development Aid (ODA) flows in contrast to other official finance (OOF). OOF plays a large role in China's finance portfolio, but has a less development oriented focus. The WB also augments its ODA with the International Bank for Reconstruction and Development (IBRD), which provides development finance in the form of loans with interest rates closer to market rates. However, we expect a clearer relationship between aid and conflict than with less concessionary development finance. One reason is that the domestic government's role in distributing concessionary development aid might increase the risk of distributive conflicts. Moreover, as development finance is acquired on a loan basis, the respective government has to pay it back and, hence, has larger incentives to invest it in a sustainable way.

[24] AidData cannot distinguish exactly how much money from the Chinese commitments is disbursed in a particular year. However, for the share of projects for which we know disbursments, they seem to follow on average on year after the commitment Moreover, an examination of further lags in Table A13 suggests that this timing is not driving the subsequently reported results.



(a)



(b)

Figure 2a Chinese (2000-2012) and WB (1995-2012) development aid. Authors' depiction based on AidData (2017) and Dreher et al. (2019).
Figure 2b Conflict 1996-2014. Authors' depiction based on Croicu and Sundberg (2015).
Category 1 (binary) = B+C, Category 2 (binary) = C, Category 3 (continuous) = {A, B, C}
Notes: Depicted borders refer to countries (thick line) and first administrative divisions (thin line).

violence including the involved parties, casualties, and location.[25] Table 2 shows that the range of battle-related deaths per region-year varies between 0 and 33,417. Figure 2b shows a map with all conflict events in our sample period, distinguishing between conflict with less than 5 BRD, between 5 and 25, and more than 25 BRD. Studies at the country level usually use thresholds of 25 or 100 to define a conflict. As our study is at the smaller regional level, we choose 5 per region-year as the threshold for our binary conflict indicator at the region-year. We show robustness tests using 25 BRD as the threshold as well as the log of battle-related deaths as a continuous indicator. To examine non-lethal conflict, we define similar binary variables for smaller scale conflict events like demonstrations, strikes or riots, as well as for non-lethal government repression based on the Social Conflict Analysis Database (SCAD, Salehyan et al., 2012). Table A10 in the data appendix provides a more detailed overview about the different outcome variables.

Table 2 provides summary statistics of the most important variables at the country-region-year level. With regard to the main treatment variables, namely WB and Chinese Aid, it becomes visible that WB aid is, on average, higher per region-year: USD 2.2 million versus USD 1.4 million, respectively. However, the standard deviation of Chinese aid indicates the large differences in scale. The maximum Chinese spending per region-year is USD 900 million; nearly twice as large as the highest value for the WB. This corresponds to many Chinese "mega-projects", like railroads,

[25] Alternatives are the ACLED and PRIO Gridded datasets, which rely on similar primary data as UCDP. One issue with PRIO Gridded data is that neighboring cells in a 50km radius are also coded as conflict-affected, which might lead to erroneous conflict coding of neighboring administrative and ethnic regions (Tollefsen et al., 2012). ACLED is broader in coverage than UCDP data, but is criticized for its partly ambiguous inclusion criteria and vague geo-coding (Eck, 2012).

dams and power plants. Figure 2a, which depicts the geographical distribution of development aid locations, shows that the engagement of both donors exhibits sufficient variation across, as well as within, countries.

Table 2: Descriptive statistics - ADM1 Region

	Mean	SD	Min	Max
World Bank Aid	2,240,340	8,991,909	0	488,643,178
ln(WB Aid)	6	9	-5	20
Chinese Aid	1,391,272	22,843,120	0	900,000,000
ln(Chinese Aid)	-4	4	-5	21
Battle Related Deaths	21	342	0	33,417
Conflict Incidence in Percent	12	32	0	100

Notes: Descriptive statistics for our main variables. ln(Aid) is based on aid +0.01USD.

3.3 Control Variables

Even though we will not decisively rely on control variables due to the bad control problem, we provide specifications using the most important aspects highlighted in the previous literature. Initial regional development is proxied using nighttime light (Henderson et al., 2012). Regional population matters for aid allocation and to scale the potential for conflict for regions of different size (Hegre and Sambanis, 2006), and is calculated based on the *Gridded Population of the World* dataset (Center for International Earth Science Information Network (CIESIN) Columbia University, 2016). From the PRIO Gridded data (Tollefsen et al., 2012), we use several natural resource indicators including oil, gold, gemstones, and narcotics, as well as measures on temperature and precipitation, that can be linked to conflict (Miguel et al., 2004). In order to match the gridded data to the respective region-year, we intersect the PRIO-Grid with the AMD1 shapefile and calculate area-weighted averages for each region. Robustness tests use data from Cederman et al. (2014) and Wucherpfennig et al. (2011) about thedistribution of ethnic groups.

4 Empirical Strategy

It is important to note that aid projects are not randomly allocated. Donors might be more or less likely to select a region based on its conflict potential, which causes concerns about endogenous selection. Over the long term, reverse causality might also cause problems if regions formerly plagued by conflict receive more aid afterward. Considering Figures 2a and 2b again helps to understand our two different approaches to identification. The first approach is to use the possibilities of the sub-national data and conditions step-by-step on more and more observables and unobservables through various fixed effects, time trends, and controls.

First, precise coding helps. Angola, for instance, receives more aid projects in regions that also experience more conflict. In contrast, the regions in Sudan that often receive aid are not the ones that experience conflict. Country-level studies, in contrast, would code both countries as cases where a country received aid and also experienced conflict. Second, the correlation between aid and conflict is affected by unobserved region-specific factors that can make both receiving aid projects and conflict more likely. Region-fixed effects can eliminate time-invariant differences that affect this joint likelihood of receiving aid and experiencing conflict.

Third, country times year (from now on country-year) fixed effects go one step further and eliminate the effect of any spurious event at the country-year level that could affect conflict and, by chance, coincides with changes in aid allocation, which could include a change in political regime.One problem in economics that was recently emphasized is that very restrictive specifications might eliminate too much variation in the data, and thus, in our case, falsely conclude that there is no conflict-fueling effect. For that reason, our first table eliminates variation step-by-step to transparently show how the relationship between aid and conflict changes when eliminating further variation. Of course, this does not fully eliminate concerns about endogenous selection. We will, in the following sections,assess the direction of bias and propose an IV strategy for each donor.

4.1 Linear models with fixed effects, time trends and control variables

Our baseline empirical specifications are

$$C_{i,c,t} = \beta_1 A_{i,c,t-1/t-2} + \delta_i + \tau_t + \Delta_i T + X'^{Ex}_{i,c,t}\beta_2 + \epsilon_{i,c,t}, \tag{1}$$

$$C_{i,c,t} = \beta_1 A_{i,c,t-1/t-2} + \delta_i + \tau_t + \Delta_i T + X'^{Ex}_{i,c,t}\beta_2 + \kappa_{c,t} + \epsilon_{i,c,t}, \tag{2}$$

where $C_{i,c,t}$ is our conflict indicator of interest in region i, in country c and year t. $A_{i,c,t-1/t-2}$ is log of per capita aid. We lag WB aid disbursements by one year, and follow the literature (Dreher et al., 2019, 2017) in taking a two year lag for Chinese aid commitments, given that disbursements seem to occur on average one year after the commitment.

Our specifications includes time, and region fixed effects, δ_i and τ_t. Furthermore, we add regional linear time trends $\delta_i T$ to control for any differing linear conflict trends across regions. Including country-year fixed effects $\kappa_{c,t}$ asks a subtly different question: conditional on the whole country being in conflict or not in a particular year, how did previous aid receipts affect the conditional likelihood of a particular region to be in conflict? For that reason, the following sections always consider one specification without (eq. 1) and one with country-year fixed effects (eq. 2).

We distinguish between three types of control variables. First, exogenous controls such as climatic shocks. Second, we account for the effect of time-invariant controls like elevation or ruggedness of terrain by interacting them with year dummies. These first two sets are contained

in $X^{Ex}_{i,c,t}$, as they are not at risk of being bad controls. Third, we twice lag potentially "bad controls" like nighttime light (as a proxy for economic activity), or population, $X^{End}_{i,c,t-2}$, which can be affected directly by aid projects. Using "pre-determined" values solves the bad control issue only if we assume sequential exogeneity. For that reason, those variables are tested but not part of our preferred specifications. The error term is denoted as $\epsilon_{ir,t}$.

Standard errors are two-way clustered at both the country-year and the regional level (Cameron et al., 2011). This allows for arbitrary correlation within a country and year, which is important as conflicts often have a strong spatial component and tend to spill over. Also allowing for correlation within a region over time is important as conflict also tends to exhibit strong persistence over time. Tables A50 and A51 show that the results are similar for other clustering options.

4.2 Instrumental Variable approach

Our IV strategies exploit the heterogeneous impact of a plausibly exogenous time-series interacted with either a pre-determined or fixed cross-sectional difference (as in , e.g., Nunn and Qian, 2014).[26] The identifying assumption is that, in absence of a change in the time series, there would be common trends in aid allocation, within low and high aid probability recipient regions. As in any Difference-in-Difference (DiD) setup, both regression stages control for the main constituting terms forming the interaction, and only the interaction term is used as the conditionally exogenous instrument in the first stage. For both the WB and China, we use a cumulative, either initial or pre-determined, probability over the whole sample period as the cross-sectional difference. This is computed by dividing the number of years a region i has received aid in the past, by the number of years passed until year t.[27] The identification strategies for WB aid and Chinese aid, hence, differ in the donor-specific probability, and in the time-varying factor T_t, used to induce variation in project allocations over time.

[26] Nunn and Qian exploit temporal variation in US wheat production, which they interact with the aid recipient's probability to receive US food aid. In essence, this strategy is similar to Bartik instruments used, e.g., in the labor economics literature (Autor et al., 2013) or the shift-share instruments common in the migration literature (Altonji and Card, 1991). In contrast to most Bartik and shift-share instruments, where cross-sectional units differ in many dimensions, e.g., different industry shares or immigrant enclave sizes, the units in our approach differ only along one dimension, the probability to receive aid.

[27] If our sample begins in 1995, and a region received aid in three out of five years, the value of the probability in 1999 would be 0.6. If aid receipts stop in 1999, the probability would decline to 0.5 in 2000 as the country would have received aid in three out of six years. The constant probability used in Nunn and Qian (2014) or Bluhm et al. (2018) relies on all observed treatment values per unit, i.e., the term for region i in year t also depends on the values in $t+1, t+2, ...$ These future values can themselves be a function of conflict. Nizalova and Murtazashvili (2016) show that under certain assumptions the interaction of an exogenous variable with an endogenous variable can be interpreted as exogenous when controlling for the endogenous factor (in this case the constant probability). Nonetheless, using initial or pre-determined values minimizes endogeneity concerns.

4.2.1 Application to WB aid

Based on discussions with WB staff, as well as recipient country personnel, the mechanism we exploit and document for identification is the following. We exploit the heterogeneous effect of yearly variation in the availability of additional IDA resources on regions with an initially lower or higher likelihood of receiving aid.[28] If there are more funds available, the Bank has an interest to exhaust the funds and allocate them to recipient countries. Countries and regions which were, or currently are, already involved in projects receive a larger share of the additional funds, partly because the costs of information screening and other preparation are lower.[29] Variation in the funding position, defined as "the extent to which IDA can commit to new financing of loans, grants, and guarantees given its financial position at any point in time" (World Bank, 2015), can be caused by internal adjustments, the timing of payments by the shareholders, and repayments by large borrowers like India. Conflict in any individual sub-national African region, in particular conditional on country or even country-year fixed effects, cannot plausibly affect the position significantly.[30]

From 1995 to 2007 we rely on the reconstructed time series by Dreher et al. (2017); starting in 2008, we use the measure publicly disclosed in the annual financial reports.[31] This time-varying variable is interacted with the pre-determined probability of a region to receive aid, $p_{i,c,t-2}$, in order to capture that higher probability regions should profit more from higher funding positions. For simplicity we do not display the fixed effects and time trends here. The first stage equation then has the following form:

$$Aid_{i,c,t-1} = \alpha_1 p_{i,c,t-2} + \alpha_2 IDA_{t-1} + \alpha_3 p_{i,c,t-2} IDA_{t-1} + \epsilon_{i,c,t-1} \tag{3}$$

One potential problem associated with approaches like ours is that, even if the temporal variation is plausibly exogenous, trends in the time series might overlap with differing trends in the outcome variable, leading to a spurious IV effect. This risk is exacerbated if the time series is relatively short and dominated by long-term trends (Christian and Barrett, 2017). The left-hand side of Figure 2 shows how differences in the long term conflict trend, within both low and high probability regions, could lead to such biased estimates. The right-hand side figure then shows that

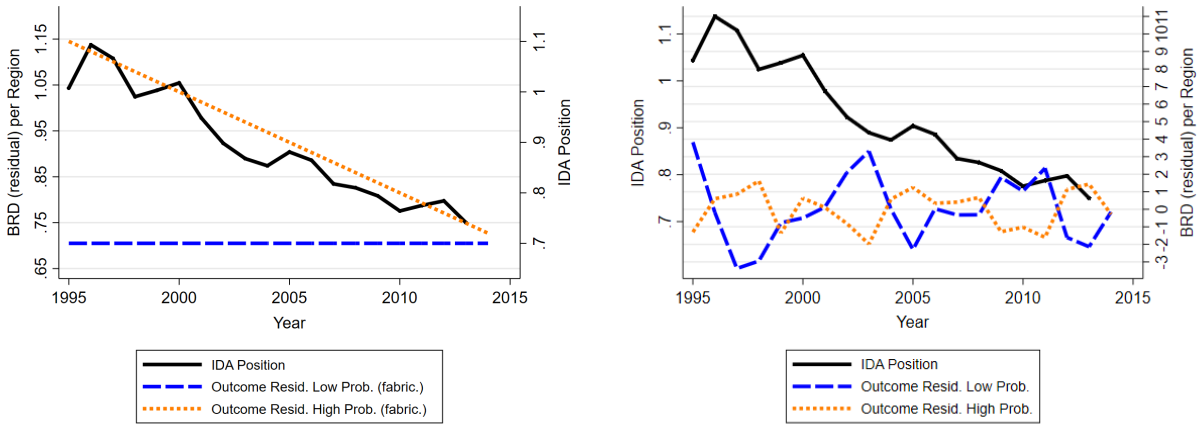
[28] The idea is based on Lang (2016) and Gehring and Lang (2018), who employ such a supply-push identification approach using variation in the IMF's liquidity.

[29] Galiani et al. (2017) use Gross National Income (GNI) as a threshold for IDA eligibility. We prefer the liquidity over graduation for three distinct reasons. First, the continuous liquidity treatment has a broader Local Average Treatment Effect than the IDA threshold where the latter refers to a subset of mostly stable countries on a growth path. Second, Kerner et al. (2017) suggest that countries have a certain leeway to postpone graduation by reporting lower GNI estimates. Third, applying the IDA threshold in our setting, we find that the threshold does not always imply a strict reduction in IDA allocations.

[30] One worry would be that the global level of conflict correlates with free resources, and at the same time more strongly correlated with conflict in high probability regions. Controlling for global conflict levels interacted with the probability in Tables A25 and A26 does not affect the first or second stage results.

[31] Because the WB's fiscal year ends in June, the reported position in the fiscal years t and t-1 can both affect disbursements in t-1. Using only the position in t-1 is a viable alternative and also works well in first stage estimations, which is demonstrated in Table A19. Using both fiscal years t and t-1 to compute the funding position appears more coherent and is applied subsequently.

the actual residual variation net of fixed effects and time trends in conflict that we exploit exhibits no such trend. Moreover, despite a general decline in the funding position, there is also sufficient year-on-year variation. For instance, the position initially increases between 1996 and 1997, before falling sharply in the following years. Table A15 and A16 indicate that this works both through the extensive and intensive margin. High probability regions have a higher likelihood to profit by receiving aid in a particular year, and conditional on receiving aid in a given year, the size of the disbursements also becomes larger.[32]



(a) Problematic trends (fabricated) in outcomes

(b) Actual trends in outcomes

Figure 2: WB- IDA funding position and conflict outcomes for low and high probability regions.

Note: Figure (a) displays the temporal variation we use in our interacted instrument, the IDA Funding Position (solid line), along with fabricated trends in the conflict outcomes for low (long-dashed line) and high probability (short-dashed line) recipient regions. The trends are fabricated to illustrate potentially problematic trend differences that could induce a spurious correlation. Figure (b) displays the IDA Funding Position (solid line), along with the actual trends in the conflict outcomes for low (long-dashed line) and high probability (short-dashed line) recipient regions. The displayed outcomes in (b) are the residuals net of the fixed effects and time trends that we use in Table 3, column (4), the remaining unexplained variation in the outcomes used in our preferred specification.

4.2.2 Application to China

Regarding China, we make use of the fact that the economic structure and political incentives frequently lead to excess domestic production in various sections. To clear markets and protect domestic companies from potential losses, China commits to more aid projects abroad (Dreher et al., 2019), a pattern not entirely unknown from European agricultural overproduction. These additional projects are often large-scale infrastructure projects that directly use overproduced commondities as inputs (Bräutigam, 2011), but Bluhm et al. (2018) show that commodity (over-)production also induces variation in other sectors like education or health. As the Chinese "mega-deals" (Strange et al., 2017) cannot easily be duplicated or scaled within regions, and the country tries expanding its

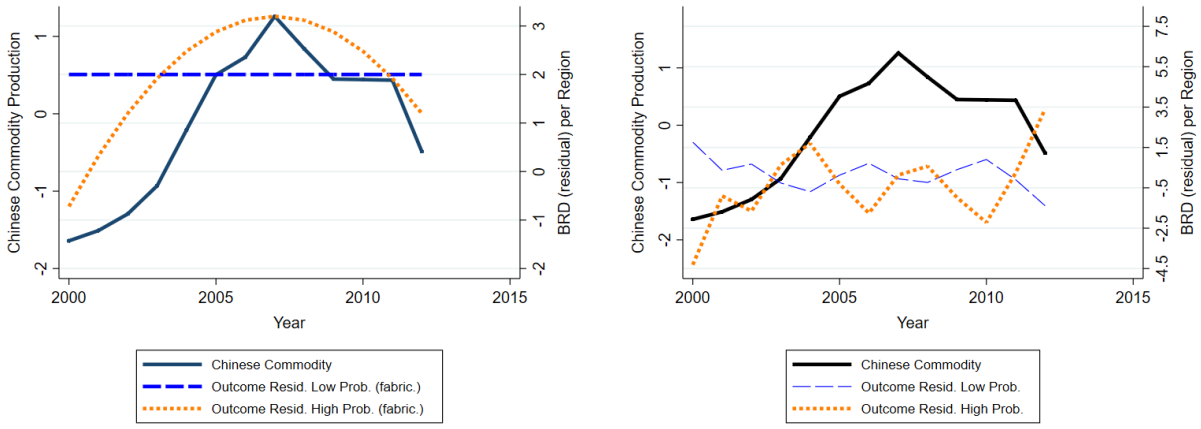
[32] To allow the reader to assess the trends in the treatment variables, Figure A11 depicts the time series for the means of logged WB and Chinese aid per high and low exposure regions.

influence during our sample period, the additional projects are often implemented in low probability regions that had initially no, or very few, projects.

We follow Bluhm et al. (2018) and use principal component analysis to construct a time series on Chinese domestic commodity over-production, $T_{i,c,t}$. Again, the time-varying variable is then interacted with the pre-determined probability of a region to receive aid, $p_{i,c,t-3}$, to capture that lower probability regions should profit more from Chinese commodity overproduction. The first stage equation has the following form:

$$Aid_{i,c,t-2} = \alpha_1 p_{i,c,t-3} + \alpha_2 Commodity_{t-3} + \alpha_3 p_{i,c,t-3} Commodity_{t-3} + X^{Ex}_{i,c,t}\alpha_4 + \epsilon_{i,c,t-2} \quad (4)$$

Again, the left-hand side of Figure 3 illustrates a fabricated, problematic, relationship with differing long-term conflict trends in low and high probability regions. The commodity time series variable is inverse U-shaped; if conflict trends in either low or high probability regions would, for other reason than aid, also follow such a pattern, the IV results would be driven by these spurious trends. The right-hand side graph assures us that there is no overlap with such trend in the variation we exploit. Table A15 supports our theoretical prior that the first stage relationship is mainly driven by the extensive margin, e.g., the likelihood of having at least one active project in a specific region-year. Regions without pre-existing projects are more likely to receive a project as the Chinese development budget expands.



(a) Problematic trends (fabricated) in outcomes

(b) Actual trends in outcomes

Figure 3: China: Chinese commodity production and conflict outcomes for low and high probability regions.

Notes: Figure (a) displays the temporal variation we use in our interacted instrument, the Chinese Commodity Production (solid line), along with fabricated trends in the conflict outcomes for low (long-dashed line) and high probability (short-dashed line) recipient regions. The trends are fabricated to illustrate potentially problematic trend differences that could induce a spurious correlation. Figure (b) displays the Chinese commodity (over-)production (solid line), along with the actual trends in the conflict outcomes for low (long-dashed line) and high probability (short-dashed line) recipient regions. The displayed outcomes in (b) are the residuals net of the fixed effects and time trends that we use in Table 3, column (4), the remaining unexplained variation in the outcomes used in our preferred specification.

5 Results

5.1 OLS, fixed effects and time trends

To allow readers to evaluate a potential trade-off between eliminating bias and over-controlling, we begin by showing simple correlations, and then add fixed effects, time trends, and different categories of control variables step-by-step. Beginning with WB aid in Table 3, we find that the raw correlation with conflict incidence is negative. Adding country and year fixed effects shifts the coefficient upward (column 2); adding country-specific linear and quadratic trends to capture country-specific conflict dynamics moves the coefficient slightly downward to -0.05 (column 3). When adding region fixed effects, which capture region-specific, time-invariant attributes, that can explain heterogeneity within countries, the point estimate nearly quadruples in size to -0.21 and becomes statistically significant at the 1%-level (column 4).

Adding exogenous controls, and time-invariant region characteristics, interacted with year dummies to capture their potentially time-varying influence (column 5), as well as adding region-specific linear time trends, changes the coefficient only slightly (column 6). Column 8 goes one step further by controlling for country-year fixed effects. The remaining variation then is only due to differences in aid across regions within country-years, conditional on the country as a whole being in conflict or not. Despite the strict specification, the robust negative relationship between WB aid and conflict

does not disappear and remains significant at the 5%-level. The coefficient of -0.1772 suggests that a one standard deviation change in log WB aid decreases the conflict likelihood by $9\times0.1772 \approx 1.59$ percentage points. To put this into perspective, the average of conflict incidence with our threshold of five battle-related deaths (BRD) is 12 percent; accordingly, this is small, however it is a non-trivial change. The coefficient becomes insignificant when controlling for lagged values of factors that are potentially endogenous controls (columns 7 and 9), but remains negative. Although these are only conditional correlations, the fact that 8 out of 9 coefficients are negative suggests that there is no conflict-fueling effect of WB aid, on average.

Turning to China, our theoretical prior was that certain arguments suggest a positive relationship with conflict to be more likely when involved with Chinese aid. Nonetheless, the raw correlation with conflict is also negative. The coefficient drops drastically in size when adding country and time fixed effects, as well as country-specific time trends (columns 2 and 3), but loses significance. Overall, the coefficients are much smaller and closer to zero than those for the WB. Remarkably, however, there is not a single positive coefficient, also suggesting no signs of a conflict-inducing effect of Chinese aid. Our preferred specifications in columns 6 and 8 indicate that increasing log Chinese aid by one standard deviation decreases the conflict likelihood by $4\times0.0654 \approx 0.26$ percentage points.

Table 3 reveals how many degrees of freedom researchers possess in selecting their preferred specification in such a setting. What we find reassuring is that throughout all these different specifications there is no sign of a conflict-inducing effect for either WB or Chinese projects. Relating to the ideas about assessing coefficient changes when moving towards more restrictive specifications in Altonji et al. (2005), we also see that the effect of adding additional FE, trends, and covariates neither suggests a strong systematic upward, nor a downward bias. The confidence interval comprises negative, zero, and some positive effects. Still, considering the rich set of specifications we examined, it seems highly unlikely that other unobserved factors would push the average effect towards a statistically, and economically, significant conflict-fueling effect. Even if there would be large changes in Chinese aid, their average effect on conflict would be rather small, when compared to the average likelihood of conflict of 12 percent. For further tests, we continue examining specifications based on the set of control variables in columns 6 and 8, and we address the potentially remaining selection bias with our IV estimations.

Table 3: OLS results - Aid and conflict likelihood

Panel A: WB Aid	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1918*	0.0010	-0.0496	-0.2129***	-0.2057***	-0.1608**	-0.0419	-0.1772**	-0.1420
	(0.0989)	(0.0776)	(0.0683)	(0.0659)	(0.0701)	(0.0782)	(0.0849)	(0.0847)	(0.1048)
N	13104	13104	13104	13104	13050	13050	11017	13050	11017
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	-0.1753**	-0.0233	-0.0026	-0.1090*	-0.0663	-0.0654	-0.0641	-0.0347	-0.0369
	(0.0865)	(0.0705)	(0.0642)	(0.0572)	(0.0783)	(0.0827)	(0.0877)	(0.1015)	(0.0916)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Conflicts are considered for the WB from 1996 to 2013 and for Chinese aid from 2002 to 2014 due to the lag structure. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

5.2 Instrumental Variables

Table 4 shows the IV results for our preferred specifications. The first stages work well for both donors. The interaction term between the prior probability to receive aid and the IDA position, respectively Chinese commodity production, is highly significant with and without country-year fixed effects. On average, the first stage works better for the WB (F= 99/86) than for China (F=36/31); all F-statistics, however, are well above the critical value of 10. In addition to being relevant, the signs of the coefficients align with our priors. Regions with a higher pre-determined probability profit more from a higher WB liquidity, and regions with an initially lower probability profit more from an expansion of the Chinese aid budget.

The second stage results largely confirm the OLS results. Both specifications yield negative coefficients for the WB and China. The coefficients for the WB are somehow smaller (larger) in the specification without (with) country-year FE, and become statistically insignificant. The coefficients for China become much more negative, however they remain insignificant. There is again no evidence for a conflict-fueling effect of aid projects for either of the two donors. Taken at face value, increasing log WB aid by a one standard deviation decreases the conflict likelihood by about $9 \times 0.2252 \approx 2.03$ percentage points. Raising log Chinese aid by one standard deviation would decrease conflict by $4 \times 0.1886 \approx 0.75$ percentage points.

Examining those results with more scrutiny raises the question to what degree they represent a local average treatment effect (LATE) that might be different from the average effect. By definition, the IV estimate is identified using a particular kind of variation in the variable of interest that is caused by the excluded instrument, the interaction term. Nonetheless, comparing the IV point estimates with OLS shows only minor differences in size, and no difference with regard to the direction of the effects.

We can check whether the direction of the changes, when moving from OLS to IV estimations, is plausible by running OLS specifications using leads and lags of our variable of interest. More specifically, Table A13 shows three lags, the contemporaneous value, and a lead term of the treatment variable. For the WB, there are no clear indications of a pre-trend that would signal selection bias. For China, however, the lead terms are positive in both cases. This indicates that China selects into regions that are more likely to have experienced a conflict in the past years; this is potentially due to China being less worried about conflict, or because of attempts to fill up the space left by other donors who are more hesitant to enter that type of region.[33] This suggests an upward bias in the OLS coefficients, which is in line with the IV coefficients for China being more negative. For the WB, without apparent pre-trends, IV and OLS results are very similar. Despite

[33] In this regard, Strange et al. (2017) demonstrate that after withdrawal of Western aid Chinese commitments fill gaps and, hence, can reduce conflict risk.

signaling a null or slightly negative effect on average, the rather large standard errors suggest that this average effect hides considerable heterogeneity.

Table 4: IV results - Aid and conflict likelihood at the ADM1 level

Panel A: World Bank Aid	(1)	(2)
IV Second stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1014	-0.2252
	(0.3752)	(0.4192)
N	11600	11600
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	99.639	86.724
IV First stage: World Bank		
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	70.9363***	80.8832***
	(7.1065)	(8.6854)
N	12325	12325
Panel B: Chinese Aid	**(1)**	**(2)**
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	-0.2582	-0.1886
	(0.4282)	(0.5257)
N	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.457	31.087
IV First stage: World Bank & Chinese Commodity		
$\text{Position}_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	57.3235***	63.8053***
	(12.0425)	(24.1932)
N	7975	7975
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include exogenous (time-varying) controls. Year and region fixed effects as well as time trends are included in all regressions. Time Trends include linear and squared country-specific time trends and a linear regional trend. The constituent term of the probability is depicted in Table A17. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

5.3 Channels - Actors

We can distinguish between state violence, rebel (labeled non-state-) violence, and one-sided actions by either of those two groups against civilians not directly involved in the conflict. The UCDP Codebook describes one-sided violence as "[...] the use of armed force by the government of a state or by a formally organized group against civilians [...]" (Eck and Hultman, 2007).

Table 5 shows the results using these distinctions with and without country-year FE. The coefficients for state-based violence against rebels (column 1 and 2), conflict between different rebel groups (column 3 and 4), and rebel violence against civilians (column 7 and 8) are partly of an

economically significant size, but even though they vary in the direction of the effect are all far from being statistically significant. When considering state conflict against civilians (column 5 and 6), we find that in a region that receives more WB aid, there is significantly less conflict. A one standard deviation change in log WB aid decreases the likelihood of violence against civilians with at least 5 BRD by $9 \times 0.29 \approx 2.61$ percentage points. This is plausible as the WB is known to punish human right violations by governments. For instance, the WB suspended important aid payments in Indonesia so as to push the government towards finding peaceful bargaining solutions on the island of Timor (Tir and Karreth, 2018).

Although Tir and Karreth (2018) focus their arguments on international organizations like the WB, which impose strong conditionality, we find a conflict-reducing effect for Chinese aid as well. Changing log Chinese aid by one standard deviation decreases the conflict likelihood substantially by $4 \times 0.5673 \approx 2.27$ percentage points. Even without officially imposing conditions about human rights violations, governments in Africa seem to abstain from lethal actions against civilians when China supports a project in a particular region. Besides business interests, the presence of Chinese workers might be another reason to convince recipient governments to abstain from engaging in actions that cause civilian casualties and endanger stability.[34]

5.4 Channels - Protest and government repression

Raleigh et al. (2010) emphasize "dire consequences" of Chinese aid, and state that "political violence rates involving state forces also increase." Should we then conclude that these fears are unwarranted? Not necessarily. Our analysis has focused on violent conflict that involves battle-related deaths, but Kishi and Raleigh highlight that states "use this aid to finance their hold on power by repressing political competitors." It seems plausible that China is interested in avoiding outright battles; however using government repression to ensure stability is also in line with its domestic approach and ideology. In that regard, some observers claim that Chinese aid purposefully supports building up recipient countries' surveillance capacities to repress elements of civil society.[35]

[34] While Tir and Karreth (2018) argue that the prospect of gaining access to aid can also constrain rebels, the coefficients for one-sided rebel violence against civilians are negative for both donors, but remain statistically insignificant.

[35] Washington Post, last accessed 02.02.2019.

Table 5: Aid and conflict types by actors

Panel A: World Bank Aid - IV	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
IV: World Bank - Actors	State vs. N-State		N-State vs. N-State		State vs. Civilans		N-State vs. Civilians	
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.4177	-0.4319	0.1252	0.1488	-0.3579*	-0.2939*	-0.0961	-0.1417
	(0.3174)	(0.2630)	(0.2096)	(0.2447)	(0.1885)	(0.1739)	(0.2072)	(0.2704)
N	12325	12325	12325	12325	12325	12325	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Kleibergen-Paap weak identification F-statistic	99.639	86.724	99.639	86.724	99.639	86.724	99.639	86.724
Panel B: Chinese Aid - IV								
IV: Chinese Commodity - Actors	State vs. N-State		N-State vs. N-State		State vs. Civilans		N-State vs. Civilians	
$\ln(Chinese\,Aid_{\text{t-2}})$	0.2749	0.2200	0.2462	0.4178	-0.5336**	-0.5673**	-0.3273	-0.3553
	(0.2104)	(0.2280)	(0.1924)	(0.2637)	(0.2300)	(0.2877)	(0.2520)	(0.3066)
N	7975	7975	7975	7975	7975	7975	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.578	31.190	36.578	31.190	36.578	31.190	36.578	31.190
Country-Year FE	No	Yes	No	Yes	No	Yes	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD$\geq$5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Exogenous (time-varying) controls are included in all regressions. Time Trends included, consist of linear and squared country-specific time trends as well as linear regional time trends. "State vs. N-State" refers to state-based violence against non-government actors, "N-State vs. N-State" refers to non-government violence against the other organized non-state groups, and "State vs. Civilians" refers to one-sided violence versus civilians by the government and "N-State vs. Civilians" refers to one-sided violence versus civilians by non-state actors. The categories are mutually exclusive. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Table A37 depicts corresponding OLS results.

To evaluate this hypothesis, we rely on the Social Conflict Analysis Database (SCAD) (Salehyan et al., 2012). The particular strength of this database is that it covers types of social and political disorder, that are usually overlooked in other conflict datasets, with georeferenced data available from 1990-2016. We are, in particular, interested in two types of variables. First, we code binary variables that take on a value of one if there was at least one riot, strike, or demonstration in a district in order to measure potential civil unrest or protests against projects related to China. Second, we code whether there was at least one event recorded as repression by the government, focusing on non-lethal repression to distinguish these regressions from our previous results.

Panel A of table 6 shows the results for our two main specifications, but now, replacing the outcome variable with an indicator, measuring whether at least one demonstration, riot, or strike took place.[36] For the WB, both specifications yield a negative coefficient but remain statistically insignificant. Regarding China, we observe negative coefficients, which are of modest size (100% more aid increase the likelihood of riots by 0.07%) and remain statistically insignificant. Accordingly, despite reports indicating increasing protests against the presence of Chinese buisness (Wegenast et al., 2017), we find no clear relationship between Chinese aid and citizen protests over our sample period.[37]

Recipient governments might achieve this absence of protests and outright conflict by intensifying non-lethal repression. Panel B of table 6 tests whether aid is related to more reports of non-lethal government repression.[38] The results indicate neither a positive nor significantly negative relationship for the WB. The results for China contrast our previous findings and establish that repression intensifies in regions where China is present. A 100% increase in Chinese aid increases the likelihood of experiencing repression by about 0.77%, which is significant considering an average of 2.26%.

[36] Table A38 depicts corresponding OLS results. Tables A31, A32 and A33 show OLS regressions separately for demonstrations, riots and strikes; Table A34 separate IV estimates. None of them turns out significant once region FE are included.

[37] See, for instance, The Telegraph, last accessed 02.02.2019.

[38] Table A36 reports results for a count variable of non-lethal pro-government violence events, which are robust to this change in the outcome variable. Table A35 verifies that this is driven by events recorded in SCAD that are distinct from the UCDP events, by coding only those region-years as a one that did not experience lethal government violence against civilians according to UCDP.

Table 6: Protests and non-lethal government repression [SCAD]

Panel A: Riots, demonstrations and strikes	(1)	(2)
IV Second stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.3854	-0.2032
	(0.3092)	(0.3362)
N	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	99.639	86.724
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	-0.1599	-0.0742
	(0.3964)	(0.4453)
N	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.457	31.087
Panel B: Non-lethal Government Repression		
IV Second stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.1543	0.0885
	(0.1042)	(0.1177)
N	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	99.639	86.724
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	0.6103**	0.7696**
	(0.2873)	(0.3439)
N	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.457	31.087
Country-Year FE	No	Yes

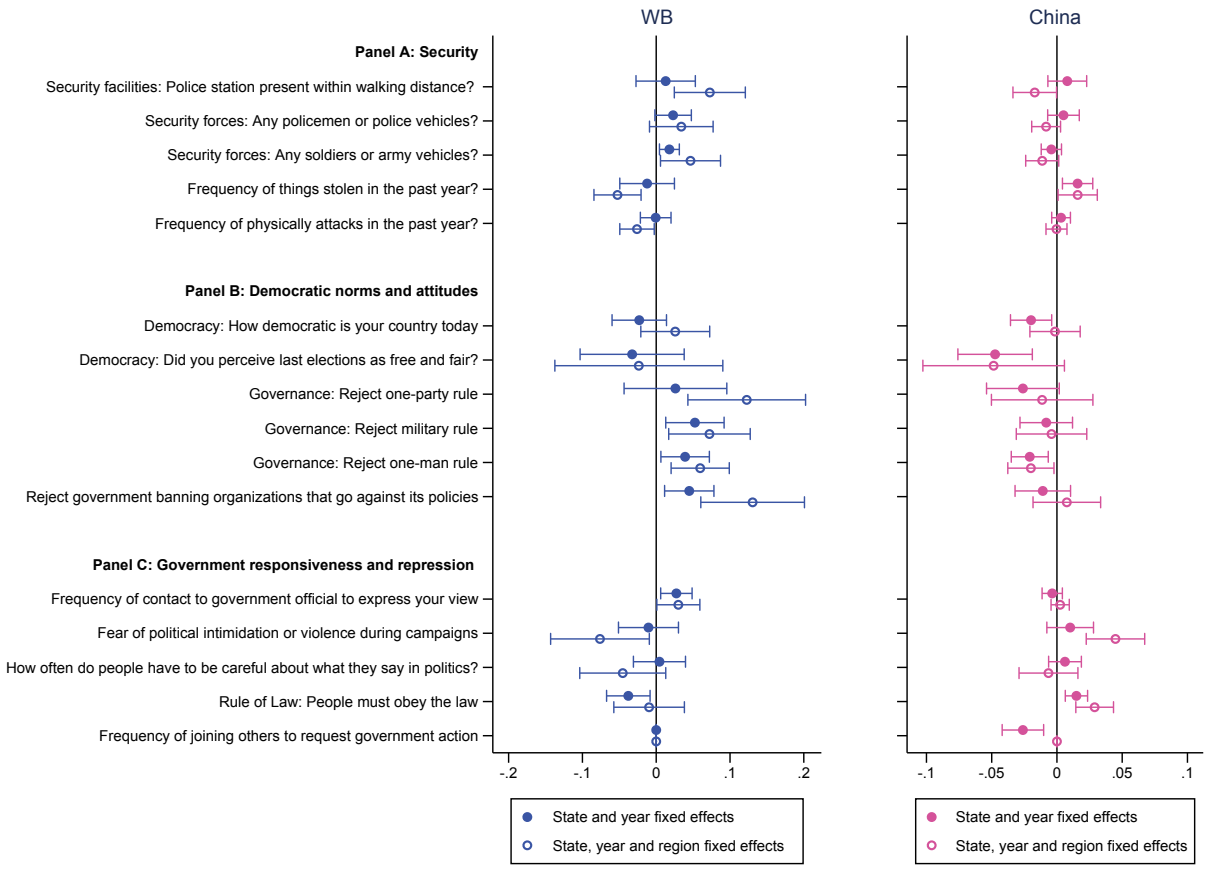
Notes: The table displays regression coefficients for a binary dependent variable of the occurrence of riots, demonstrations and strikes in Panel A and non-lethal repression (pro-government violence) in Panel B. The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level.Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends.
* $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

5.5 Channels - Attitudes

Examing the associated mechanisms for all effects is beyond the scope of this paper. Still, we can present some correlational evidence using georeferenced Afrobarometer data to investigate the plausibility of some of our results. To do so, we match data, from all Afrobarometer waves, to the regions and years in our sample, and compute the region-year level average of each question we use. Details are provided in Appendix Table A11. Note that the survey covers varying subsets of all African countries in selective years, so that the resulting dataset comprises an unbalanced panel with gaps. The temporal variation is not sufficient for a strong first stage using the IV, and we can

only use less restrictive sets of fixed effects than in our main specifications. Figure 4, thus, plots the coefficients from individual OLS regressions of selected relevant questions on WB and Chinese aid: model 1 uses country and time FE, model 2 region and time FE.

Figure 4: OLS regressions on mechanisms using Afrobarometer for WB and China



Notes: The figure shows coefficient plots along with 90% confidence intervals of individual OLS regressions of log WB and log Chinese aid on the respective questions from Afrobarometer. All outcome question responses were standardized with mean zero. Respondents were matched to the ADM1 regions using the provided geocoordinates. Table A46 provides the full regression results. Afrobarometer was conducted in the years 1999-2015 for a varying number of 12 to 36 countries, resulting in an unbalanced panel with uneven gaps between years.

The results are grouped in three categories. Panel A refers to questions signaling the presence of state security forces as a measure for state capacity within the area, and the ability to maintain a monopoly of violence. Moreover, we use two questions asking whether respondents or their families were the victims of robbery or physical attacks in the past year. The results suggest that the WB engagement is associated with an increase in security forces and a reduction in crimes. There is no such increase for China. However, one needs to keep in mind that these are conditional correlations, and China might select into regions more likely to experience conflict and a deterioration in state capacity.

Panel B examines democratic norms and attitudes. Even though the results are not necessarily causal, some differences stand out that reflect the differential approaches the two donors take. There are some indications that the perception of democracy, and the fairness of elections, deteriorate

in regions with Chinese aid projects. The WB seems to have a consistent impact on democratic norms. Respondents are more likely to reject one-party rule, military rule, and one-man rule, which is not the case for China. With the coefficients being consistently significant in both models regarding one-man rule, respondents are less likely to reject these authoritarian governance forms. This could indicate that China helps some authoritarian regimes to stay in power. Note that in a more detailed examination Isaksson and Kotsadam (2018a) also find a deterioration in norms, and an increase in local corruption, associated with Chinese projects.

Panel C examines questions indicating the way the government interacts with its citizens and its use of repression. In regions with more WB aid, people report being more apt to frequently contact their government officials and express their views. There is no such effect for China. In regions with WB aid, the fear of political intimidation or violence is lower, while it is higher in regions with Chinese aid activities. At the same time, there is no clear difference in whether people think they have to be careful what they say privately about politics. Finally, two results stand out. In regions with more Chinese aid, respondents state much more clearly that people must always obey the law. Moreover, there is a negative correlation between Chinese aid and the willingness to join others to request government action. These correlations correspond to the different norms and conditions of the WB and China that we described above.

Importantly, all of these results on mechanisms need to be interpreted cautiously, and do not necessarily signal causality. Still, they underline that the different approaches taken by the two donors matter. Against this background, it is important to reconsider that aid by both donors is, if anything, leading to less conflict. The results on mechanisms suggest that, for the WB, this is going along with an improvement in the democratic norms and security provision by the government. For China, one interpretation is that the country is exporting stability which results in a reduction in the likelihood of certain types of conflict. Still, this increase in stability seems to come at the cost of increased government repression in addition to a weakening of democratic processes.

5.6 Sensitivity

Modifiable area unit problem - different aggregation levels: First, we aggregate at the country level. This allows us to see the aggregate impact of potential spill-overs to other regions, and enables us to compare our main results to studies at the country level. We show results both with and without controlling for the share of aid projects that could not be assigned to a particular ADM1 region. These are, to a large extent, projects where money flows directly to the central government. The coefficients are also negative for both donors in both specifications. Thus, our

results at the local level do not seem to be driven by choosing a particular spatial unit.[39]

In Table A42 (A43), we move towards OLS (IV) regressions at a lower level of aggregation, the ADM2 level. Note that we are capturing a smaller share of all projects at this level due to the precision level in the georeferencing. The OLS results for the WB and China are both similar to the ones at the ADM1 level, with all coefficients being insignificant and the majority being negative, especially, when conditioning on more restrictive fixed effects. The IV point estimates differ somehow, but never become statistically significant.

Choice of conflict indicator: As we discuss in the data section, there is no "correct" coding of the dependent variable, just more and less plausible choices. Table A27 (A28) presents alternative regression results with a higher conflict threshold of at least 25 BRD per region-year using the OLS (IV) specifications. Table A29 (Table A30) considers the log of battle-related deaths (+0.01) as a continuous measure of conflict intensity instead of looking at a binary indicator of conflict incidence using OLS (IV). We find largely negative OLS coefficients for the WB and slightly positive ones for China. However, with IV, both coefficients turn negative, in line with previous results.

Instrumental variable: We conduct the majority of robustness tests with regard to our IV strategy. As outlined, we take the concern serious that our instrumental variable might intersect with a spurious trend as suggested by Christian and Barrett (2017). In this regard, when taking non-stationarity (Table A20) of the time series into account by taking first differences of conflict, aid and our liquidity indicator. The second stage results in Table A21 remain clearly indistinguishable from zero. While it is arguably unlikely that conflict in a specific sub-national region determines our global liquidity indicators, we also assess the robustness of the instrumental variables by controlling for global conflict levels in Tables A25 and A26. The strength of the instruments remains virtually unaffected and the point estimates remain negative and statistically insignificant.

The second component of the IV, the probability term, may be computed in different ways. We test various plausible options. Using the cumulative probability is advantageous, as it only uses pre-determined values; yet, it could create problems if the probability in the first year(s) is not sufficiently informative. Table A22 drops the first year of the respective panel (starting at 1998 for the WB's IDA, and 2003 for Chinese Commodity Production), so that the first probability is already based on at least two observations. Table A23 uses a constant probability from the third year of the respective sample onwards, i.e., 1998 for the WB's IDA, and 2003 for Chinese Commodity Production. Table A24 drops the 10 highest leverage region-year observations. Figures A13 and A14 display the IV estimates dropping country-by-country, so as to avoid the possibility

[39] Point estimates for the less precisely coded aid can be found in Table A45. Although the coefficient for non-geocoded WB aid at the country level turns positive it remains small and insignificant, further supporting that there is also a null effect at the country level. OLS and IV point estimates for geo-coded aid aggregatred at the country level can be found in Table A44. The coefficients remains small and insignificant, as well.

of the relationship being driven by one particular state. Both first and second stage results are robust to all these choices and specifications.

Moreover, Table A18 reports reduced-form estimates. Table A14 uses a lead of aid as a placebo treatment in the first stage, which always shows up as statistically insignificantly. Table A17 reports the first stage, including the coefficient for the probability.

Both donors in same specification A second trade-off is whether to show both donors over the same period and in the same equation. The advantage is increased comparability and accounting for aid from one donor as a potential omitted variable in the other donors equation. Table A55 (Table A56) shows that the OLS (IV) results also suggest no conflict-fueling effects when including both donors jointly. When estimating the IV specifications jointly for both donors and the restricted time period, the K-P F-statistics for the WB are much smaller (Table A56), giving rise to concerns about weak IV problems. Still, this table also shows that the two instruments actually capture distinct variation: the interaction instrument for the WB is still significant in explaining variation in WB aid, and the IV for China still significant in explaining variation in Chinese aid. Even with the weak IV, the table indicates no conflict-fueling effects for both donors. Moreover, Humphrey and Michaelowa (2018) find no systematic relationship between the selection of locations by the two donors, also indicating that this choice does not bias our results.

Non-linear estimators: In line with Berman et al. (2017), we also run a Poisson Pseudo Maximum Likelihood estimation in Table A47, which is suitable for binary outcomes with a large fraction of zeros. Moreover, we implement a conditional logit estimation in Table 48. The results are generally in line with the main findings in terms of coefficient signs. However, note that the models converge only when restricting us to the use of year fixed effects.

Temporal dependence: As conflict might be highly persistent over time, we include a lagged dependent variable in Table A49. The results are very similar, with mostly negative and partly significant coefficients for the WB and China.

6 Conclusion

China increases its engagement in Africa, which raises both hopes and rejections among political as well as academic observers. The big question is whether African countries will benefit or suffer from this foreign engagement? To answer this question, we examine aid from the World Bank and China, two donors that represent strongly contrasting approaches to development. The former is a multilateral donor that emphasizes human right conditions, expert knowledge, and specifically engages in conflict-sensitive programming. The latter is the major emerging donor, emphasizing mutual economic benefits without the weight of numerous explicit conditions for recipient governments and without officially accounting for potential conflict risks (Asmus et al., 2017b; Hernandez, 2017). Our paper contributes to the literature on aid and conflict by providing, as we hope, the most comprehensive analysis of the relationship between development aid and stability in a multi-country analysis at the sub-national level this far.

Our results using aid projects and conflict in the same region show no signs of a conflict-fueling effect. The WB tends to have a conflict-reducing effect in some fixed effects specifications, but when using instrumental variable strategies, estimates for both donors are negative and insignificant, on average. Looking at heterogeneity with regard to actors and types of conflict, we find that the threat of losing out on future aid payments leads to a reduction in lethal violence by governments against civilians related to Chinese and WB projects.

In contrast to a substantial amount of media reports, we also find no net effect of Chinese aid projects on civilian unrest and protests in Africa. At the same time, we do, however, observe that regions in which China is engaged experience an increased likelihood of government repression against targeted individuals or groups. Thus we cannot say with certainty whether the non-significant result on protest reflects the higher costs of protesting due to repression or that there is no reason to protest. WB aid has neither a significant net effect on protests or on government repression.

Nonetheless, when considering attitudes from Afrobarometer surveys, our results suggests that WB aid has positive effects on perceived safety, democratic norms, and democratic values. Chinese aid is associated with attitudes related to stability like a higher adherence to the rule of law, but also with a higher acceptance of autocratic approaches. The results suggest a rationale where China is eager to export stability and avoid violent conflict that endangers its workers and investment, while also being more supportive of repression and autocratic rule than traditional Western-influenced donors like the WB.

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Appendix

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A Data Appendix

A.1 Sources

Table 7 lists descriptions and sources of our independent, dependent and control variables.

Table 7: Data Sources

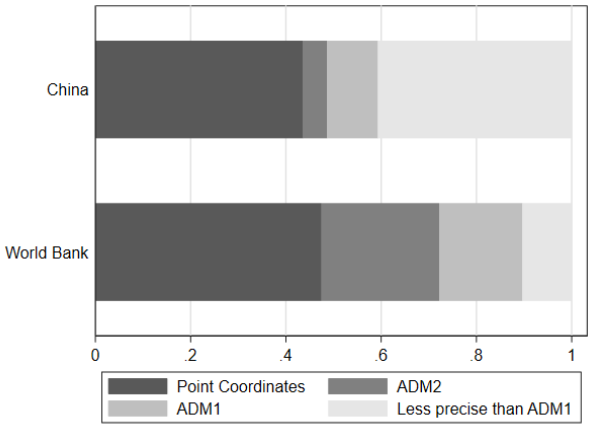
Variable Name	*Variable Description*	Time Period	*Variable Source*
WB Aid	log of WB Aid disbursements in a given region-year	1995-2012	Strandow et al. (2011)
Chinese Aid	log of Chinese Aid commitments in a given region-year	2000-2012	Dreher et al. (2017)
Strikes, Riots, Demonstrations	Binary indicator (100;0) if any violent event of this type in a given region-year took place	1995-2012	Salehyan et al. (2012)
Intensity 1/2	Binary indicator (100;0) 1 if >=5/>=25 persons were killed in a given region-year	1995-2014	(Croicu and Sundberg, 2015)
Population	Continuous indicator of regional population	1995-2014	(CIESIN 2016)
Drought (end of rainseason)	SPI value of drought severity of the region's entire rainy season	1995-2014	Tollefsen et al. (2012) and Guttman (1999)
Drought (start of rainseason)	SPI value of drought severity during the first month of the region's rainy season	1995-2014	Tollefsen et al. (2012) and Guttman (1999)
Temperature	Mean temperature (in degrees Celsius) per region-year	1995-2014	Tollefsen et al. (2012) and Fan and Van den Dool (2008)
Precipitation	Total amount of precipitation (in millimeter) per region-year	1995-2014	Tollefsen et al. (2012) and Schneider et al. (2015)
Chinese Commodity (Over-) Production	Chinese Commodity production (factor, standardized)	1999-2013	Data Stream, World Steel Association (2019), ...
IDA Funding Position	"Bank‘s net investment portfolio & its non-negotiable, non-interest-bearing demand obligations (on account of members‘ subscriptions and contributions)" divided by "sum of the Bank‘s undisbursed commitments of development credits and grants."	1995-2012	Dreher et al. (2017)
Elevation	Standard deviation of regional elevation as an indicator of ruggedness of terrain	Constant	USGS Global 30 Arc-Second Elevation (GTOPO30)
Ocean, Rivers, Lakes	Binary indicator of presence of rivers, lakes or ocean in a given ADM1 region	Constant	Natural Earth, available at Natural Earth.com
Landarea	Area of a given region	Constant	Hijmans et al. (2010)
Travel Time (Mean)	Gives the mean regional estimate of the travel time to the nearest major city	Constant	Tollefsen et al. (2012) and Uchida and Nelson (2009)
Borders	Binary indicator if a given ADM1 region borders another country	Constant	Own estimations based on Hijmans et al. (2010)

A.2 Independent Variables (Development Aid)

WB's IDA & IBRD disbursements

For our analysis, we draw on the "WB IBRD-IDA, Level 1, Version 1.4.1" provided by the AidData consortium, which covers approved loans under the IBRD-IDA lending line between 1995 and 2014.[40] These data correspond to project aid disbursed from 5,684 projects in 61,243 locations. The data build on information provided by the WB, including the disbursement dates, project sectors and disbursed values. These values were deflated to 2011 values. In an effort to allow for more fine-grained analysis of aid projects, AidData's coders filtered the location names from aid project documentation and assigned these to specific locations. While for some projects exact locations including latitude and longitude were assigned, other projects, which had a more policy or regulation oriented purpose, could only be assigned to an administrative level (e.g., the first level of sub-national regions (provinces) or the second level (districts). In order to include as many disbursements as possible, but to be also able to grasp the advantages of georeferenced data, we focus our analysis on these administrative levels. For our administrative boundaries, we build on the GADM dataset constructed by Hijmans et al. (2010). One difficulty with these data is that for some countries, including more populous nations like Armenia, more fine-grained administrative distinctions are missing. As the size of administrative regions is not fixed by size across countries, we assume in this cases that our ADM1 regions would be ADM2 regions.

Figure 5 displays the development finance locations coded by donor, distinguishing all projects (precision 1-8), projects coded at least at the first administrative level (precision 1-4), projects coded at least at the second administrative level (precision 1-3) and projects coded more precise (precision 1-2).



(a)

Figure 5: No. of Project Locations by Precision Codes

[40] As the number of documented projects declines steeply after 2012, we focus on the 1995-2012 period.

One challenge arises in projects with a multitude of locations, where it is not possible to derive a distinct value of disbursements. In this regard, we suggest two solutions.

First, we allocate disbursements by the number of locations. In line with previous research by Dreher and Lohmann (2015) we assume that aid is distributed equally across locations and allocate aid proportionally to the locations per region. For instance, for a project with 10 locations, where 4 locations are in region A and 6 locations are in region B, 40% of project disbursements would be accounted in region A and 60% in region B.

Second, we calculate population weighted disbursements. Here, we assume that aid is allocated based on the regional population shares. For instance, if a project would have project locations in two regions of a country, where two million inhabitants would reside in region A and three million would reside in region B, 40% of project disbursements would be accounted in region A and 60% in region B. Here, the aid attribution formula would write as follows: $Aid_{pijt} = \frac{Aid_{pit}}{\int Population_{pi}} * Population_{pj}$, where p is the project, i is the country, j is the region and t is the period for which we estimate the allocation shares.

Finally, our dataset comprises development finance from IBRD and IDA. However, only IDA disbursements can be classified as Official Development Assistance. For this purpose, disbursements were disentangled into IDA (development aid) and IBRD (development finance) disbursements.

Allocation scheme (more detailed)

Location weighting

The WB geocoded data release comes in the format of projects and several corresponding locations. For instance, a typical project report would mention the transaction amounts, the project purpose as well as different project locations. The latter can be classified in different degrees of precision (e.g., precision codes smaller than 4 correspond to locations that refer to an ADM2 region or even more precise, while precision code 4 corresponds to locations at the ADM1 level). When allocating the development aid across locations on the ADM1 and ADM2 level, we make the following assumptions based on a three-step procedure.[41] First, we subtract the share of development aid, which corresponds to locations, which are coded less precise than ADM1 (e.g., large geographic regions or aid at the country level). For example, if three out of 10 locations in a project are coded less precise than ADM1, the further analysis focuses on the remaining 70% of development aid. Second, we then allocate all aid with precision codes 1-3 to the corresponding ADM2 regions. This is done by taking the location share (either by equal or population weights) of the transaction amount per location. A certain ADM2 regions might have several locations per project or even several projects,

[41] Throughout the paper we allocate the aid either assuming equal weights per location or weighting each location by population.

we collapse our data by ADM2 region. Third, we then allocate all aid with precision code 4 to the corresponding ADM1 regions. This is done by taking the location share (either by equal or population weights) of the transaction amount per location. A certain ADM1 regions might have several locations per project or even several projects, we collapse our data by ADM1 region. In order to allow for inference on the ADM2 level, we make the assumption that transactions coded with precision 4 are attributable equally to all corresponding ADM2 regions. In practice, this is done by merging the ADM1 regions with all corresponding ADM2 regions and then splitting the aid with location or population weights. Finally, data with precision codes 1-3 and precision code 4 can be simply added upon the ADM2 level yielding our treatment variable of interest. For inference on the ADM1 level, totals of ADM2 level development assistance are created on the geounit-year level.

Table 9: Aid Allocation Formula Example

Example of Weighted Aid Allocation										
ID	Year	Aid Value	Loc. ID	ADM1 ID	ADM2 ID	Prec. Code	ADM1 Weight	Prec.4 Aid to ADM2	Prec. 1–3	Total Aid
	1995	100	2	1	1	1	1/7		14.29	14.29
1	1995	100	3	1	2	2	1/7		14.29	14.29
1	1995	100	4	2	1	4	1/7	14.29		14.29
1	1995	100	5	3	1	3	1/7		14.29	14.29
1	1995	100	6	3	2	1	1/7		14.29	14.29
1	1995	100	6	3	3	4	(1/7)*(1/3)	4.76		4.76
1	1995	100	6	3	1	4	(1/7)*(1/3)	4.76		4.76
1	1995	100	7	3	2	4	(1/7)*(1/3)	4.76		4.76
1	1995	100	8	4	1	4	1/7	14.29		14.29
Totals:								42.86	57.14	100.00

Population weighting

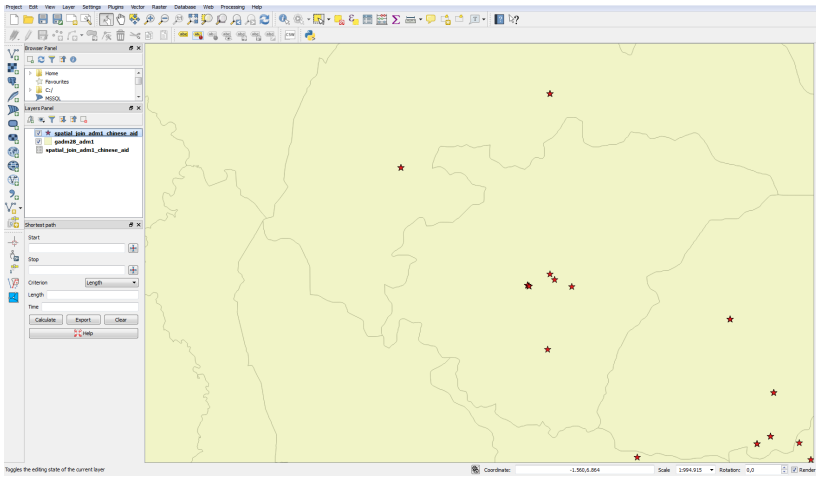
Analogous to the location weighted aid, we also distribute aid with population weights. Our population data are from the Center for International Earth Science Information Network (CIESIN) Columbia University (2016). However, some projects only consist of locations without population estimates (e.g., deserts). In this case, we assume a population of one citizen per location in order to be able to distribute those aid disbursements. We then consequently attribute population of ADM1 regions to project locations, which are coded at the ADM1 level (precision 4), and ADM2 populations to project locations, which are coded at least as precise as the ADM2 level (precision 1-3).

Similar to the location-weighing, we construct the total population of each project-year $pop_{project}$. For the projects coded with precision 4, we then attribute disbursements via the regional share in population pop_{ADM1}. This is then divided by $pop_{project}$ and multiplied with the project disburse-

ments $TransactionValue_{proj}$ in each year: $ADM1Precision_4 = \frac{pop_{ADM1}}{pop_{proj}} * TransactionValue_{proj}$. As there might be several active projects per ADM1 region, we aggregate the disbursements on the ADM1 level. In order to break those numbers down to the ADM2 level, we merge all corresponding ADM2 regions to the ADM1 regions. We then divide the population in each ADM2 region by the population in each ADM1 region and multiply this share with the yearly disbursements per region, $ADM2Precision_4 = \frac{pop_{ADM2}}{pop_{ADM1}} * ADM1Precision_4$. For the precision codes 1-3 (at least coded as precise as the ADM2 level), we then attribute disbursements via the regional share in population divided by $pop_{project}$. This is then multiplied with the project disbursements in each year: $ADM2Precision_{123} = \frac{pop_{ADM2}}{pop_{proj}} * TransactionValue_{proj}$. As there might be several active projects per ADM2 region, we aggregate the disbursements on the ADM2 level. Finally, we merge the precision code 1-3 and 4 data on the ADM2 level to obtain our variables of interest. Those can then be aggregated on the ADM1 level.

Chinese Aid (ODA-like and OOF flows)

In order to create our data on the ADM2 and ADM1 level, we make use of the feature that aid can be defined on the ADM2 level and then aggregated to the ADM1 level. One challenge with the data is, however, that we lack information on the ADM2 regions for some countries (as there are no ADM2 regions in small countries). Therefore, we create two spatial joins of ADM1 and ADM2 regions from the GADM dataset with Chinese aid point features. This yields matches of the specific project locations with the administrative regions as depicted in Figure 6.



Notes: Graphical depiction based on Quantum GIS.

Figure 6: Chinese Aid ADM1 Spatial Join

In order to create our data, we first load our ADM2 data into Stata and drop the ADM0 and ADM1 identifiers in order to be later able to rely on the identifiers from the ADM1-Aid spatial join. The next step involves merging the ADM2-Aid spatial join with the ADM1-Aid spatial join by the target-fid, which uniquely identifies the points from the Dataset "aiddata_china_1_1_1.xlsx" by

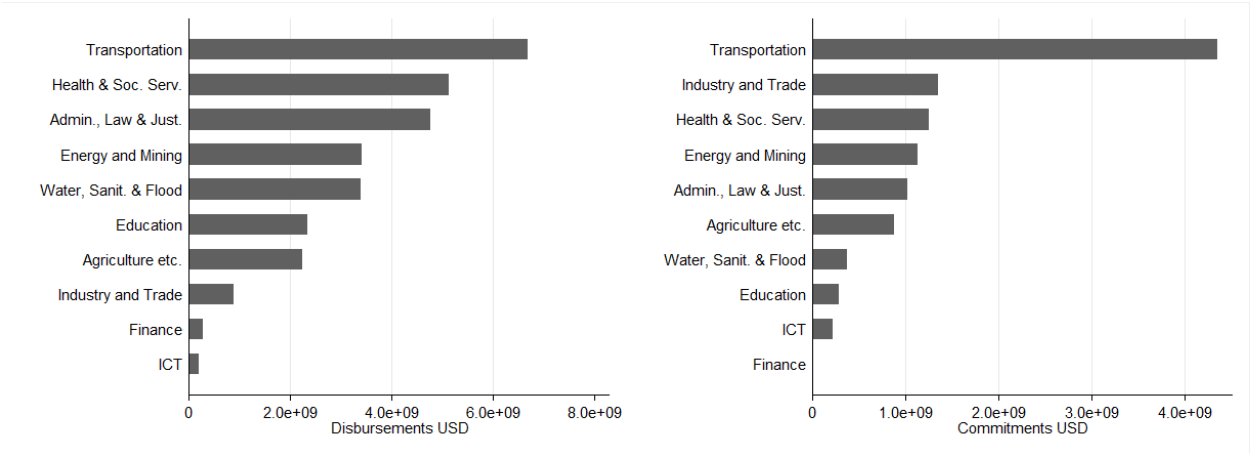
(Dreher et al., 2019) and Strange et al. (2017). Based on this data, we create unique identifiers for all ADM1 and ADM2 regions, whereby we treat ADM1 regions as ADM2 regions in cases that ADM2 regions are missing (e.g., in Cape Verde). This assumption can be made as sizes of administrative regions are rather arbitrary and several ADM2 regions are larger than other countries' ADM1 regions. After getting the regional identifiers right, we can merge (a) the spatial joins of ADM regions & Chinese aid locations with (b) data on flows of Chinese aid. In a first step, we clean these data from entries that only relate to pledges of Chinese aid (information is from the variable status254). Although the data on Chinese finance to Africa also contain information on official investment, the focus of this paper is on development aid. Thus, we focus on flows, which correspond to "ODA-like" funds as those would correspond closest to development aid (following individual correspondence with the authors of Strange et al. (2017)). The data are then merged with population data from the gridded population of the world data in order to be able to allocate financial flows with population weights in case one project had commitment locations in different administrative regions. Yet, one further challenge has to be resolved before allocating the commitments to regions, as the Chinese aid commitments are coded like WB disbursements with different precision (e.g., some are coded only for geographic features, which involve several administrative regions or are funds which go to central ministries or the government). For our commitment allocation, we only consider those projects, which are at least coded at the ADM1 level. This means that we proportionally exclude commitments, which provide information on the central level and on sub-regional level as indicated before. We furthermore distinguish between projects, which are coded only at the ADM1 level and ones that provide information on the ADM2 level (or more precise). The former are proportionally split over the underlying ADM2 regions. Although the latter can be precisely traced back to the ADM2 region, it might happen that projects have commitments in several ADM2 regions. In this case, we also split the commitments proportionally by locations or population as indicated earlier.

To exploit sectoral variation in development finance both for the WB and China, we make use of the information provided by Strange et al. (2017) on Chinese aid's sectoral allocation using the OECD's Creditor Reporting System (CRS) codes. To achieve comparability with the broad sectors indicated for the WB, we assign sectors as follows: "Agriculture, Fishing and Forestry" (CRS-310:"Agriculture, Forestry and Fishing"), "Public Administration, Law and Justice" (CRS-150), "Information and communication" (CRS-220: "Communications"), "Education" (CRS-110: "Education"), "Finance" (CRS-240: "Banking and Financial Services"), "Health and other social services" (CRS-120: "Health," CRS-160: "Other Social infrastructure and services"), "Energy and mining" (CRS-230: "Energy Generation and Supply"), "Transportation" (CRS-210: "Transport and Storage"), "Water, sanitation and flood protection" (CRS-140: "Water Supply and Sanitation"),

"Industry and Trade" (CRS-330: "Trade and Tourism," CRS-320: "Industry, Mining, Construction").

Sectoral distribution of aid disbursements

We use additional information on the financier for each disbursement for each project. Based on this information, we can construct sectoral distributions of aid flows. While both donors are investing heavily in transportation across Africa, further priorities differ. The WB supports Health and Social Services strongly, whereas China commits a large share of its funds to Industry & Trade.



(a) (b)

Figure 7: Sectoral Distribution of Aid: (a) WB's IDA; (b) China

A.3 Dependent Variables (Conflict data)

Table 10 provides an overview about the different conflict outcomes considered in this paper. The construction of the data and sources are described in more detail in the subsequent paragraphs.

Table 10: Descriptive statistics - ADM1 Region

	Mean	SD	Min	Max
Conflict Incidence	11.65	32.08	0	100
State Based Conflict	7.01	25.54	0	100
Non-State Based Conflict	3.74	18.97	0	100
State Violence vs Civilians	1.83	13.39	0	100
Non-State Violence vs Civilians	3.41	18.14	0	100
Riots, Strikes, Demonstrations	13.59	34.27	0	100
Riots	8.08	27.26	0	100
Strikes	7.53	26.40	0	100
Demonstrations	2.92	16.83	0	100
Non-lethal Pro-GVMT Violence	1.16	10.71	0	100

Notes: Descriptive statistics for our main outcome variables. The sample period is 1995-2014 in order to account for the different lag structures. Click here to go back to section 3.2.

UCDP Data

AidData and UCDP use the same coding framework, so we can use similar coding rules and restrict us to events coded at least at the ADM1 level (precision codes 1-4). For the more precise data (precision codes 1 and 2), we again use a point to polygon analysis on the ADM level. As one conflict event is always coded in one discernible location (Croicu and Sundberg, 2015), we do not need to make additional distributional assumptions by location number or population size for conflict data, because we do not face issues of multiple project locations, which we had in the aid data. Yet, for conflict observations on the ADM1 level (precision code 4), we do not distribute battle-related deaths by population weights across ADM2 regions.

One further useful feature of the UCDP data is that it is possible to discern three different types of violence. Those are namely the government against organized groups (type 1), organized non-governmental groups versus the government (or against another non-governmental group) (type 2), and one-sided violence by the government against civilians (type 3 governmental) and by non-governmental groups against civilians (type 3 non-governmental).[42] UCDP data can be considered as comprehensive for our 1995 to 2012 sample, despite for Syria for which no battle-related deaths information are provided. Hence, all missing values are treated as zeros except for the Syrian case,

[42] For a more detailed decription of the different types of violence, please consult Croicu and Sundberg (2015).

which is not part of our analysis.

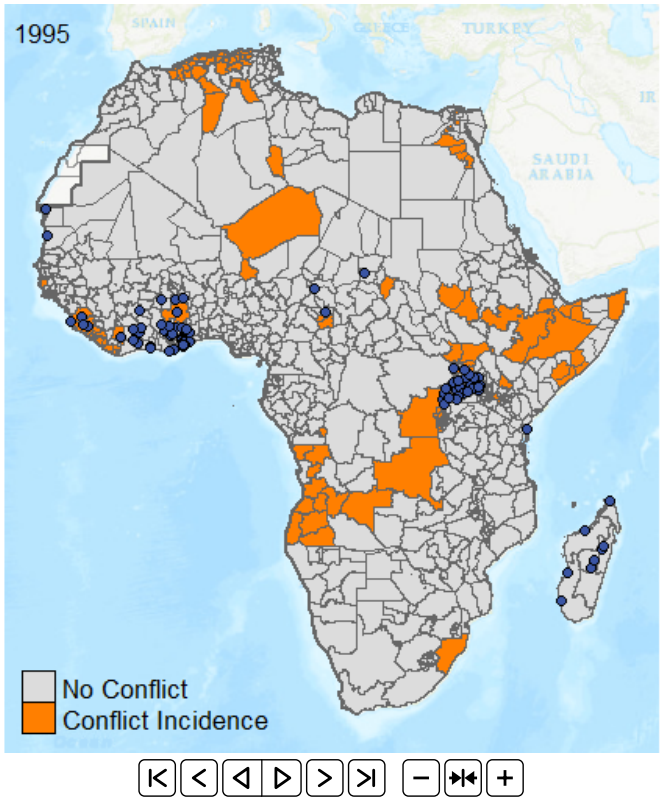


Figure 8: WB Aid and Conflict - By Year

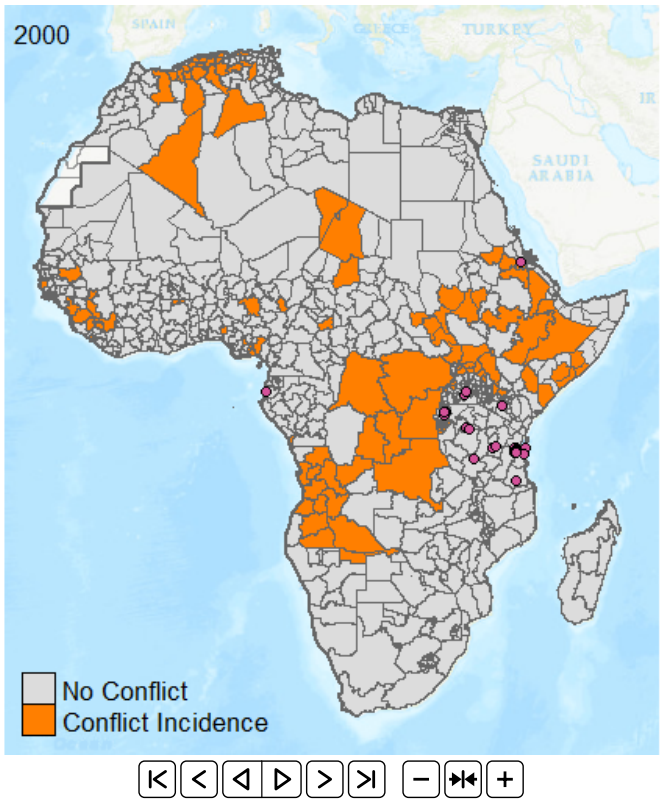


Figure 9: Chinese Aid and Conflict - By Year

SCAD data

UCDP data focus on organized violence with lethal outcomes. However, along with the different theories, it could be hypothesized that discontent and aid appropriation do not necessarily need to be linked to full-fledged conflict. What is more, recent empirical work by Bluhm et al. (2016) underscores the role of aid in conflict dynamics. Thus, we also consider social conflict as a further outcome, in terms of demonstrations and repressions, based on the Social Conflict Analysis Database (Salehyan et al., 2012). SCAD involves demonstrations, riots, strikes, coups, pro-, anti- and extra-government violence, which can, but do not necessarily have to involve casualties. In this way, SCAD complements the UCDP data.[43] SCAD mainly builds on data compiled by the Lexis-Nexis services from searches of Agence France Presse and Associated Press. Based on the available information, data are georeferenced by web searches of the locations mentioned in the event reports. Analogous to UCDP data, precision codes are provided, which are used to allocate events in a similar manner.

[43] Prior to 2014 armed conflict was not included in SCAD data and is now also distinguished from "social disturbances."

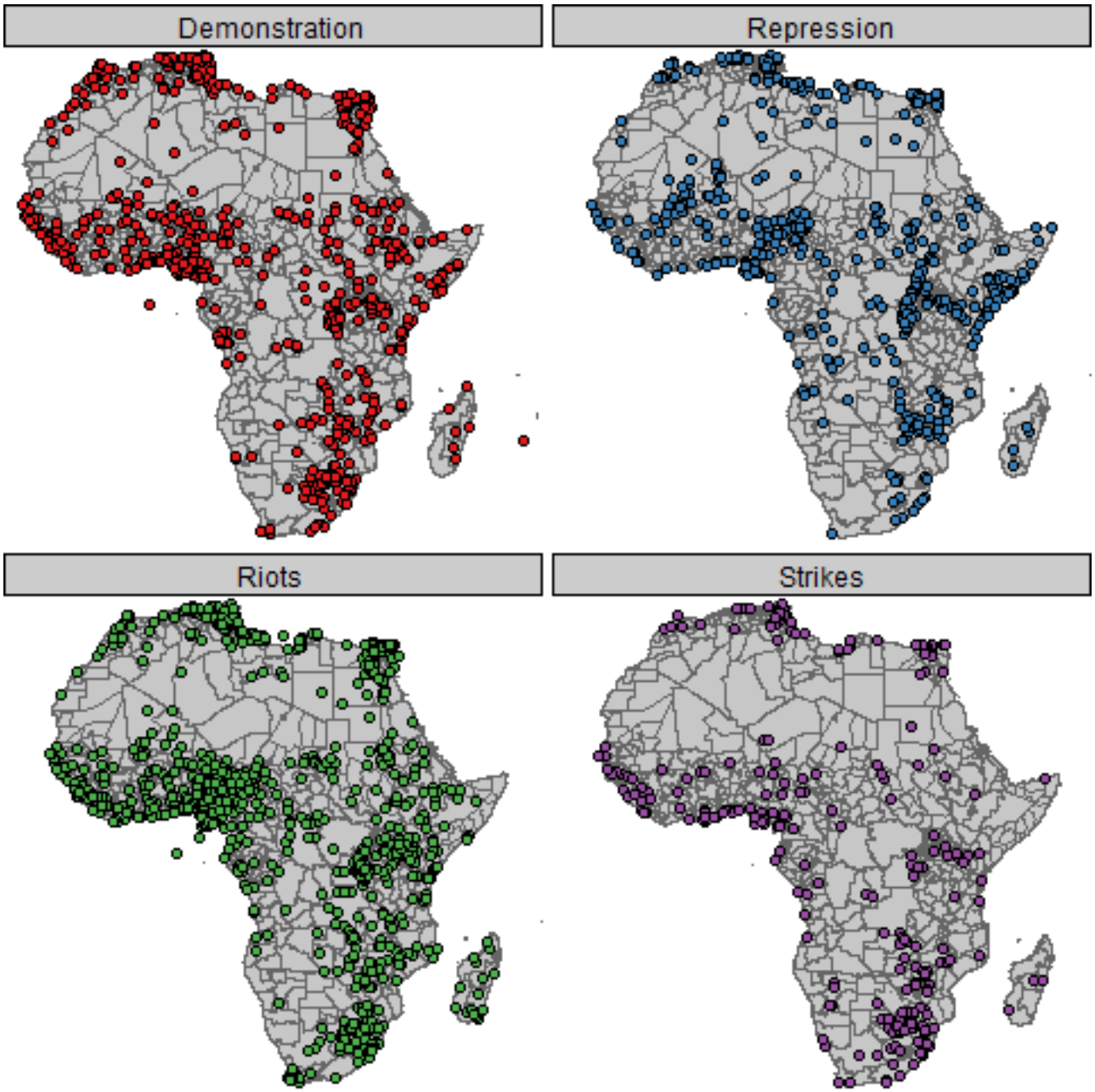


Figure 10: SCAD Data for precision codes 1-4

Matching EPR to GREG

To measure ethnic homelands, we use the GREG dataset (Weidmann et al., 2010), which is a georeferenced version of the initial locations of ethnic homelands based on the Soviet Atlas Narodov Mira. The information about the power status comes from the time-variant Ethnic Power Relations (EPR) dataset (Vogt et al., 2015). Wherever possible, we match the group power status from EPR in a particular year to one of the time-invariant GREG group homelands. The original dataset assigns eight different power statuses to groups. The differences are sometimes marginal and hard to interpret, which is why to minimize measurement error we only use the more precise information on whether a group was part of the governing coalition or not. We then intersect the ethnic group polygons with the administrative regions to classify regions as one of the three categories.

A.4 Afrobarometer

Table 11: Afrobarometer - Labels, questions and sources

Variable Name	Variable Description	Availability	Code
Panel A: Security			
Security facilities: Police station present within walking distance?	Are the following facilities present in the primary sampling unit/enumeration area, or within easy walking distance: Police station?	2008-2009, 2011-2014	ea-fac-c
Security forces: Any policemen or police vehicles?	Are the following facilities present in the primary sampling unit/enumeration area, or within easywalking distance: Police station?	2008-2009, 2011-2014	ea-sec-a
Security forces: Any soldiers or army vehicles?	In the PSU/EA, did you (or any of your colleagues) see: Any soldiers or army vehicles?	2008-2009, 2011-2014	ea-sec-b
Frequency of things stolen in the past year?	During the past year, have you or anyone in your family: Had something stolen from your house?	2002-2006, 2008-2009, 2011-2014	q11a-x
Frequency of physical attacks in the past year?	During the past year, have you or anyone in your family: Been physically attacked?	2002-2006, 2008-2009, 2011-2014	q11b-x
Panel B: Democratic norms and attitudes			
Democracy: How democratic is your country today?	In your opinion how much of a democracy is your country today?	1999-2006, 2008-2009, 2011-2014	q40
Democracy: Did you perceive last elections as free and fair?	On the whole, how would you rate the freeness and fairness of the last national election, held in your country?	1999-2001, 2005-2006, 2008-2009, 2011-2014	q22-x
Governance: Reject one-party rule	There are many ways to govern a country. Would you disapprove or approve of the following alternatives: Only one political party is allowed to stand for election and hold office?	1999-2006, 2008-2009, 2011-2014	q28a
Governance: Reject military rule	There are many ways to govern a country. Would you disapprove or approve of the following alternatives: The army comes in to govern the country?	1999-2006, 2008-2009, 2011-2014	q28b
Governance: Reject one-man rule	There are many ways to govern a country. Would you disapprove or approve of the following alternatives: Elections and Parliament are abolished so that the president can decide everything?	1999-2006, 2008-2009, 2011-2014	q28c
Reject government banning organizations that go against its policies	Which of the following statements is closest to your view? Choose Statement 1 or Statement 2. Statement 1: Government should be able to ban any organization that goes against its policies. Statement 2: We should be able to join any organization, whether or not the government approves of it.	2005-2006, 2008-2009, 2011-2014	q16-x
Panel C: Government responsiveness and repression			
Frequency of contact to government official to express your view	During the past year, how often have you contacted any of the following persons about some important problem or to give them your views: An official of a government agency?	1999-2006, 2008-2009, 2011-2014	q24c-x
Fear of political intimidation or violence during campaigns	During election campaigns in this country, how much do you personally fear becoming a victim of political intimidation or violence?	2008-2009, 2011-2014	q49-x
How often do people have to be careful about what they say in politics?	In your opinion, how often, in this country: do people have to be careful of what they say about politics?	2002-2006, 2008-2009, 2011-2014	q51a-x
Rule of Law: People must obey the law	For each of the following statements, please tell me whether you disagree or agree: The police always have the right to make people obey the law.	2002-2006, 2008-2009, 2011-2014	q42b
Frequency of joining others to request government action	Here is a list of actions that people sometimes take as citizens when they are dissatisfied with government performance. For each of these, please tell me whether you, personally, have done any of these things during the past year. If not, would you do this if you had the chance: Joined others in your community to request action from government.	2014	q27a

A.5 World Bank Aid in the Financial Sector

A deeper classification exercise on the World Bank's financial sector aid reveals that sectoral reforms may play a crucial part in mitigating conflict. To classify IDA projects that are sufficiently targeted at the financial sector, we select projects where at least 10% of disbursementsare directed at the recipient's financial sector. Moreover, we restrict the classification to projects that are precisely coded, i.e., projects where money flow to ADM1 regions is traceable. Finally, we obtain the reports for each project and develop a classification of IDA aid in the financial sector based on the project goals and descriptions.

Table 12 shows that 50% of all aid projects that are significantly targeted at the financial sector are aimed at sectoral reforms. Projects in this category supports existing government efforts for sectoral reforms and development, but includes mainly new projects that are launched outside the initiative of the recipient government.

Class.	Classification Name	Share of Projects	Description
I.	Support services to enterprises	15%	Financial and non-financial support to (selected) enterprises or enterprise sectors
II.	Support services to NGOs	2.5%	Financial and non-financial support to NGOs or welfare organisations
III.	Support services to individuals or groups	15%	Financial and non-financial support to individuals, socio-economic or geographical groups
IV.	Capacity building	10%	Capacity building in socio-economic or geographical groups or supporting other capacity building projects
V.	Sectoral reforms	50%	New projects or support of existing government efforts that primarily target sectoral adjustment and reforms
VI.	Environmental Protection	2.5%	Projects aimed at protecting or improving the environment or wildlife
VII.	Emergency support	2.5%	Projects providing emergency support
VIII.	Research support	2.5%	Research or evaluation focused projects

Specific project examples		
Class.	Project Number	Project goals
I.	P083082 Micro, Small and Medium Enterprise Project, Nigeria	Increase performance and employment levels of micro, small and medium enterprises in selected non-oil industry sub-sectors + 3 targeted states of the country through i.) Improving access to financial services, ii.) Developing the market for business development services, iii.) Development of business climate etc. **http://documents.worldbank.org/curated/en/333691474574170700/pdf/000020051-20140625225024.pdf**
III.	P052186 Microfinance Project, Madagascar	Improve income and living standards of low-income Malagasy by i.) Establishing appropriate legal, regulatory and supervisory framework for microfinance, ii.) Expanding micro-financial skills and iii.) Developing strong and sustainable local institutions. **http://documents.worldbank.org/curated/en/933341474899762755/pdf/000020051-20140625070634.pdf**
V.	P035620 Financial Institutions Development Project, Tanzania	i.) Restructuring and privatizing the National Bank of Commerce and restructuring the smaller People's Bank of Zanzibar for competition and efficiency in the banking sector, ii.) Continuation of strengthening of Bank Supervision Directorate, iii.) Improving payments system, iv.) Creating a private credit information bureau, v.) Developing the insurance industry and capital markets. **http://documents.worldbank.org/curated/en/899741468311395554/pdf/multi-page.pdf**

Table 12: World Bank Aid in the Financial Sector

B Analytical Appendix

B.1 Instrumental Variable

B.1.1 Motivation of Instrumental Variable

In order to reduce the risk of the instrument being subject to spurious trends and correlations, we need to understand the underlying mechanisms. This section is dedicated to providing a more detailed description. In a first step, Table 13 shows OLS correlations of our conflict measure with two leads and lags of aid. The second lead of Chinese aid is correlated with conflict, suggesting China selects into post-conflict settings. This correlation might also correspond to a geographically more selective allocation of Chinese funds as described in Figure 2a. We also test more formally if the instrument is suitable to tackle the selection bias, by regressing conflict on an instrumented lead term and find no significant relationship in Table 14. The instrumental variable approach is, thus, warranted to reduce selection bias.

Figure 11 depicts the funding positions for both donors along with corresponding aid flows for high and low probability regions. Evidently, aid flows in high probability regions respond more strongly to changes in the funding positions. In line with stronger first stage Kleibergen-Paap F-statistics, the relationship is more nuanced for the WB. Table 15 suggest that the instrumental variables for both donors affect the extensive margin (e.g., the probability to have at least one active aid project in a given region-year). Table 16, in turn, indicates that for the WB the intensive margin matters as well (e.g., given at least one active aid project, how much funds does a region receive?).

Table 18 depicts the reduced form estimates. In line with the main results, both interacted instruments are not significantly correlated with lethal conflict outcomes at the regional level.[44] For transparency, Table 17 displays the first stage including the constituent probability term, which, however, is not an instrument itself as we control for it in the second stage (see Section 4).

[44] While the probability constituent term enters significantly, it is not part of the instrument and we control for it in the second stage.

Table 13: ADM1 - Leads and further Lags

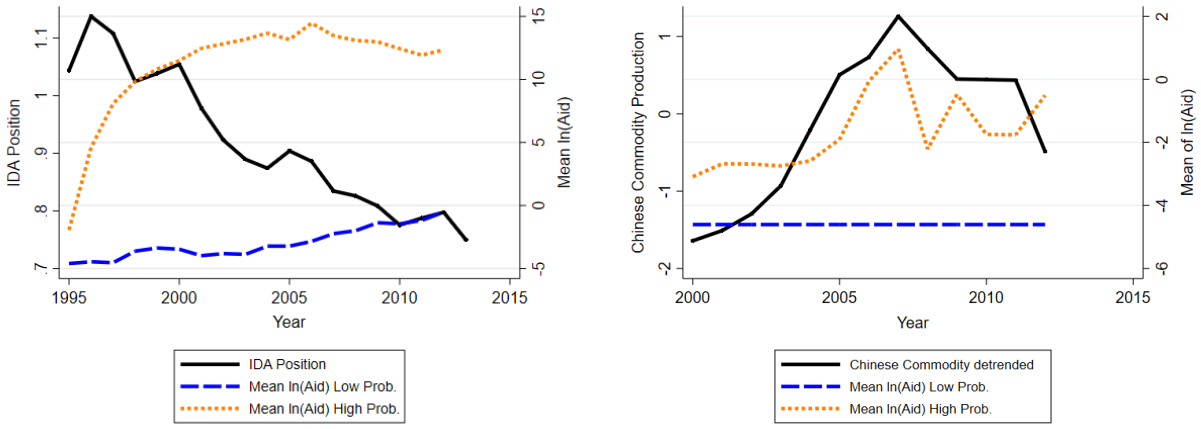
	(1)	(2)
Panel A: WB Aid		
Two Leads and Lags: World Bank		
$ln(World\,Bank\,Aid_{t+1})$	-0.0059	0.1559
	(0.1298)	(0.1199)
$ln(World\,Bank\,Aid_{t})$	-0.1089	-0.2128*
	(0.1152)	(0.1157)
$ln(World\,Bank\,Aid_{t-1})$	0.0214	-0.0933
	(0.0973)	(0.0956)
$ln(World\,Bank\,Aid_{t-2})$	0.0516	0.1424
	(0.0939)	(0.1212)
$ln(World\,Bank\,Aid_{t-3})$	-0.0811	-0.0535
	(0.0877)	(0.1076)
N	10150	10150
Panel B: Chinese Aid		
Lead and Lag: China		
$ln(Chinese\,Aid_{t+1})$	0.1681	0.2083*
	(0.1244)	(0.1258)
$ln(Chinese\,Aid_{t})$	-0.0127	0.0231
	(0.1268)	(0.1358)
$ln(Chinese\,Aid_{t-1})$	-0.0086	-0.0481
	(0.1514)	(0.1600)
$ln(Chinese\,Aid_{t-2})$	0.0121	-0.0506
	(0.1165)	(0.1313)
$ln(Chinese\,Aid_{t-3})$	0.0572	-0.0308
	(0.0986)	(0.1102)
N	6525	6525
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country- × Year	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Conflicts are considered for the WB from 1996 to 2013 and for Chinese aid from 2002 to 2014 due to the lag structure. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$
Click here to go back to section 5.2.

Table 14: ADM1 - Absence of Pre-Trends with IV. Regression with Instrumented Lead of Aid

	(1)	(2)
Panel A: WB Aid		
Placebo (Lead): World Bank		
$ln(World\,Bank\,Aid_{t+1})$	0.2299	0.2332
	(0.3586)	(0.3704)
N	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	99.481	86.444
Panel B: Chinese Aid		
Placebo (Lead): China		
$\ln(Chinese\,Aid_{t+1})$	0.0396	-0.3753
	(0.2888)	(0.3351)
N	8700	8700
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	34.263	29.941
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country × Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.



(a) (b)

Figure 11: (a) WB IDA funding position and mean of ln(WB Aid) and (b) Chinese commodity production and mean of ln(Chinese Aid).

Note: Figure 11 (a) displays the IDA funding position (solid line), the mean of logged WB aid disbursements per low probability recipient regions (long-dashed line) and the mean of logged WB aid disbursements per high probability recipient regions (short-dashed line). Figure 11 (b) displays the Chinese Commodity Production (solid line), the mean of logged Chinese aid per low probability recipient regions (long-dashed line) and the mean of logged Chinese aid per high probability recipient regions (short-dashed line). Click here to go back to section 4.2.1.

Table 15: ADM1 IV (First Stage - Extensive Margin (Likelihood of at least one active project))

	(1)	(2)
Panel A: WB Aid		
IV FS Extensive Margin: World Bank		
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	4.0782***	4.8249***
	(0.4140)	(0.5238)
$Cum.\,Prob_{\text{t-2}}$	-4.3155***	-5.0339***
	(0.4512)	(0.5508)
N	12325	12325
Panel B: Chinese Aid		
IV FS Extensive Margin: Chinese Steel		
$Chinese\,Commodity_{\text{t-3}} \times Cum.\,Prob_{\text{t-3}}$	-0.7267***	-0.6591***
	(0.1205)	(0.1163)
$Cum.\,Prob_{\text{t-3}}$	-2.2910***	-2.0452***
	(0.2510)	(0.2313)
N	7975	7975
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The table displays regression coefficients the first stage of the IV regression, when instead of the aid amount a binary indicator of aid receipts is used. The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. The constituent term of the probability is depicted in the appendix. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 4.2.1.

Table 16: ADM1 IV (First Stage - Intensive Margin)

	(1)	(2)
Panel A: WB Aid		
IV FS Intensive Margin: World Bank		
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	4.4155	8.5243**
	(3.3360)	(3.7977)
$Cum.\,Prob_{\text{t-2}}$	-2.3430	-6.3455
	(3.8699)	(4.3885)
N	7091	7081
Country-Year FE	No	Yes
Regional Time Trend	Yes	Yes
Country Time Trend:	Yes	Yes
$CountryTimeTrend^2$:	Yes	Yes
Panel B: Chinese Aid: IV FS Intensive Margin: Chinese Steel		
$Chinese\,Commodity_{\text{t-3}} \times Cum.\,Prob_{\text{t-3}}$	-0.6974	0.0592
	(1.4881)	(2.3579)
$Cum.\,Prob_{\text{t-3}}$	-3.3488	-5.8048
	(5.7090)	(3.5031)
N	232	232
Country-Time Trends	No	Yes

Notes: The table displays regression coefficients the first stage of the IV regression, when constraining the sample only on recipient regions. The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. All regressions include exogenous controls, region fixed effects and year fixed effects. Country-Year fixed effects and more rigid time trends are not included for Chinese Aid due to the more limited variation. The constituent term of the probability is depicted in the appendix. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 4.2.1.

Table 17: ADM1 IV (First Stage with probability constituent term)

Panel A: WB Aid	(1)	(2)
IV First stage: World Bank		
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	70.9363***	80.8832***
	(7.1065)	(8.6854)
$Cum.\,Prob_{\text{t-2}}$	-72.7723***	-82.0994***
	(7.7291)	(9.2698)
N	12325	12325
Panel B: Chinese Aid		
IV First stage: China		
$Chinese\,Commodity_{\text{t-3}} \times \text{Cum. Prob}_{\text{t-3}}$	-19.9994***	-19.9994***
	(4.5891)	(4.5891)
$Cum.\,Prob_{\text{t-3}}$	-54.0892***	-54.0892***
	(10.9932)	(10.9932)
N	1311	1311
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The table displays regression coefficients the first stage of the IV regression, displaying additionally the constituent term of the probability, which was also used in Table 4. The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 18: ADM1 Reduced Form

	(1)	(2)
Panel A: WB Aid		
Reduced Form: World Bank		
$Cum.\,Prob_{\text{t-2}}$	10.8281	19.2994
	(27.3795)	(33.4583)
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	-7.1921	-18.2132
	(26.5498)	(33.5818)
N	12325	12325
Panel B: Chinese Aid		
Reduced Form: Chinese Commodity		
$Cum.\,Prob_{\text{t-3}}$	-8.1658	-14.3840
	(9.7645)	(10.2370)
$Chinese\,Commodity_{,\text{factor1}} \times \text{Cum. Prob}_{\text{t-3}}$	6.6166	5.1407
	(6.6143)	(7.1646)
N	7250	7250
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country × Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

B.1.2 Robustness of Instrumental Variable

As a main specification we use the rolling average of the WB's IDA position (e.g., averaging across t and $t-1$) because the Bank's fiscal year ends already in June. For robustness, Table 19 depicts instrumental variable results using only the variation in $t-1$. Results are largely unchanged.

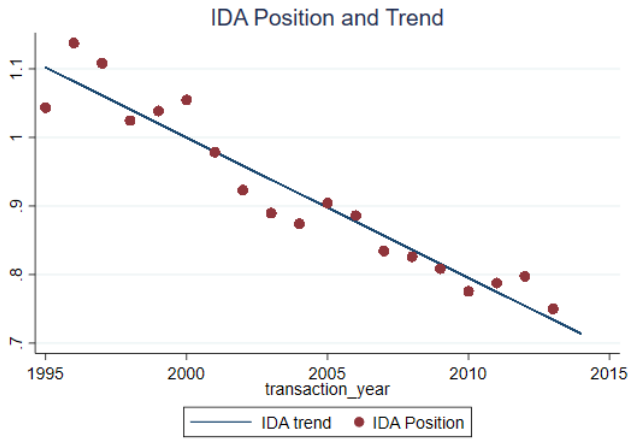
Moreover, there are several degrees of freedom regarding the definition of the interacted probability term. We indicate the robustness of an insignificant conflict-aid link when using an interacted instrument based on an initial probability from the first three sampling years (1995 to 1997 for the WB's IDA; 2000 to 2002 for Chinese Steel) in Table 23 or if excluding probability observations based only on the first sampling year in Table 22.

Finally, first stage results might be susceptible to a small share of very influential observations. Table 24 indicates that results are qualitatively unchanged if we exclude the ten high leverage region-years from the sample. Figures 13 and 14 display the first stage relationship leaving out single countries, suggesting that there are no individual states driving the relationship.

Table 19: ADM1 IV (IDA-Position$_{t-1}$)

	(1)	(2)
Panel A: WB Aid		
IV Second Stage: World Bank (t-1)		
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1294	-0.0251
	(0.3976)	(0.3868)
IV FS: World Bank (t-1)		
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	51.3655***	65.1984***
	(5.6627)	(6.9103)
$Cum.\,Prob_{\text{t-2}}$	-52.8484***	-67.1407***
	(6.2620)	(7.5204)
N	12325	12325
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. Instead of a running sum of IDA funding position in "t" and "t-1" only the variation in "t-1" is used. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 4.2.1.



Notes: Yearly values of $IDA - Position_t$ based on Dreher et al. (2017).

Figure 12: IDA Position

The Hadri test assesses the null hypothesis "All Panels are (trend) stationary". Table 20 indicates that there are at least some panels being non-stationary. For this reason, we assess the robustness of main results in Table 21, where we take first differences of outcome, treatment and instrumental variables.

Table 20: Test for (Trend) Stationarity - Hadri type

	(1)	(2)
ln(World Bank Aid $_t$)	125.8488***	89.4980***
	(0.0000)	(0.0000)
ln(Chinese Aid $_t$)	2.9868***	1.1684
	(0.0014)	(0.1213)
IDA Position detrend $_{t-1}$ × Cum. Prob $_{t-2}$	145.3093***	121.1980***
	(0.0000)	(0.0000)
Steel Prod detrend $_{t-3}$ × Cum. Prob $_{t-3}$	48.2231***	91.4337***
	(0.0000)	(0.0000)
Conflict	56.8260***	23.5170***
	(0.0000)	(0.0000)
Linear Trend	No	Yes

Notes: Conflict refers to category 1 binary conflict indicator (100 if BRD≥5, 0 if BRD<5). Hadri type test coefficient estimates for the five variables indicated in rows. p-values in columns refer to the null hypothesis "All panels are (trend) stationary." * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 21: ADM1 IV (First Difference WB & Chinese aid no detrend)

	(1)	(2)
Panel A: WB Aid		
IV Second Stage: World Bank		
$ln(World\,Bank\,Aid_{\text{(t-1)-(t-2)}})$	0.7606	0.2460
	(1.2448)	(1.2600)
N	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	64.794	83.036
IV First stage: World Bank		
IDA Position $_{\text{(t-1)-(t-2)}} \times$ Cum. Prob $_{\text{t-2}}$	18.7864***	29.3949***
	(2.3314)	(3.2222)
Cum. Prob $_{\text{t-2}}$	-31.9905***	-29.5644***
	(0.4492)	(0.4978)
Panel B: Chinese Aid		
IV First stage: China		
$ln(Chinese\,Aid_{\text{(t-2)-(t-3)}})$	-0.3690	-0.5025
	(0.4856)	(0.6253)
N	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	20.972	13.774
First Stage: Chinese Commodity		
$Chinese\,Commodity_{,\text{factor1}} \times$ Cum. Prob $_{\text{t-3}}$	-13.5621***	-10.5846***
	(2.9568)	(2.8471)
$Cum.\,Prob_{\text{t-3}}$	-81.0911***	-75.7980***
	(4.6647)	(4.8310)
Exogeneous Controls	Yes	Yes
Exogeneous Controls $\times$ Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. The constituent term of the probability is depicted in the appendix. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 4.2.2 or section 5.6.

Table 22: ADM1 IV (Without first year)

	(1)	(2)
Panel A: WB Aid		
IV Second stage:World Bank $ln(World\,Bank\,Aid_{\text{t-1}})$	-0.2904	-0.2681
	(0.4172)	(0.3975)
N	11600	11600
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	80.438	78.004
IV First stage: World Bank $IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	68.5810***	88.1297***
	(7.6467)	(9.9784)
Panel B: Chinese Aid		
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	-0.5634	-0.5104
	(0.5775)	(0.7229)
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	22.620	16.927
IV First stage: China		
$Chinese\,Commodity_{\text{t-3}} \times \text{Cum. Prob}_{\text{t-3}}$	-16.3668***	-16.3668***
	(2.4057)	(2.4057)
N	2420	2420
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The constituent term of the probability is depicted in the appendix. Click here to go back to section 5.6.

Table 23: ADM1 IV (Initial Probability)

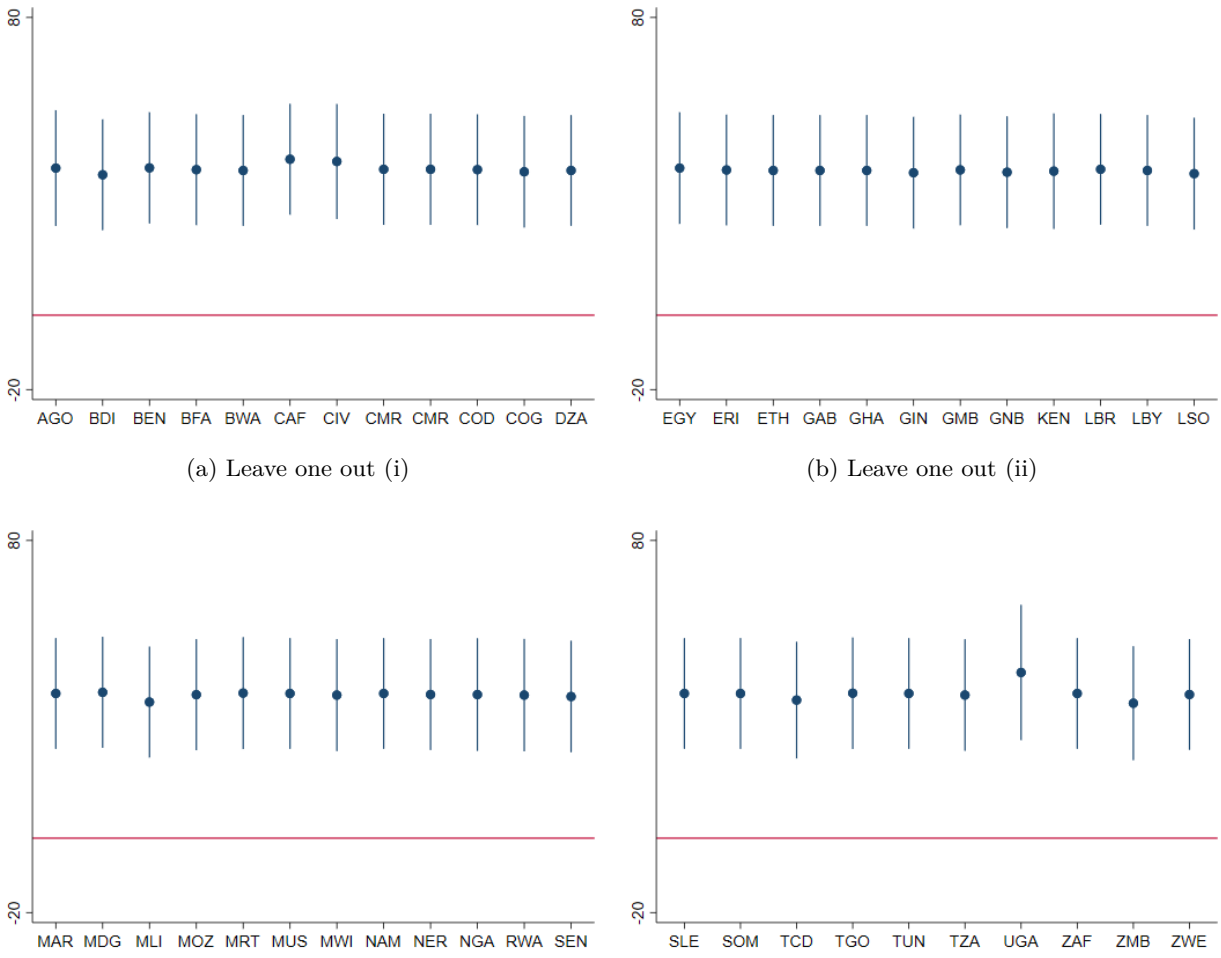
	(1)	(2)
Panel A: WB Aid		
IV Second Stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.2253	-0.3389
	(0.7469)	(0.6206)
N	11600	11600
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	27.090	26.027
IV First stage: World Bank		
$IDA\,Position_{\text{t-1}} \times \text{Con. Prob}_{98}$	43.4391***	61.1537***
	(8.3414)	(11.9820)
Panel B: Chinese Aid		
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	-1.4689	-1.2846
	(1.3446)	(1.4723)
Kleibergen-Paap underidentification test p-value	0.001	0.002
Kleibergen-Paap weak identification F-statistic	13.035	9.925
IV First stage: China		
$Chinese\,Commodity_{\text{t-3}} \times \text{Con. Prob}_{03}$	-5.8046***	-5.7207***
	(1.6049)	(1.8125)
N	7250	7250
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. The probability is based on the third year in the corresponding sample (1998 for the WB's IDA; 2003 for Chinese Commodities) and held thereafter constant. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 24: ADM1 IV (Without high leverage region)

	(1)	(2)
Panel A: WB Aid		
IV Second stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.0990	-0.2268
	(0.3761)	(0.4197)
N	12317	12291
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	99.363	86.752
IV First stage: World Bank		
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	70.8414***	80.8936***
	(7.1068)	(8.6851)
Panel B: Chinese Aid		
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	-0.2592	-0.1934
	(0.4281)	(0.5251)
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.571	31.181
IV First stage: China		
$Chinese\,Commodity_{\text{t-3}} \times \text{Cum. Prob}_{\text{t-3}}$	-14.0197***	-12.6973***
	(2.3183)	(2.2739)
N	7974	7974
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.



(a) Leave one out (i)

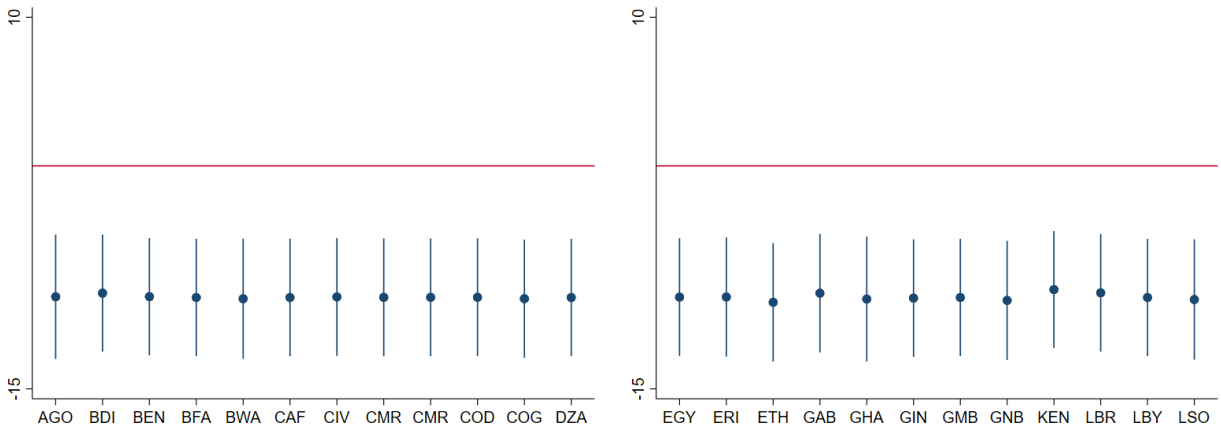
(b) Leave one out (ii)

(c) Leave one out (iii)

(d) Leave one out (iv)

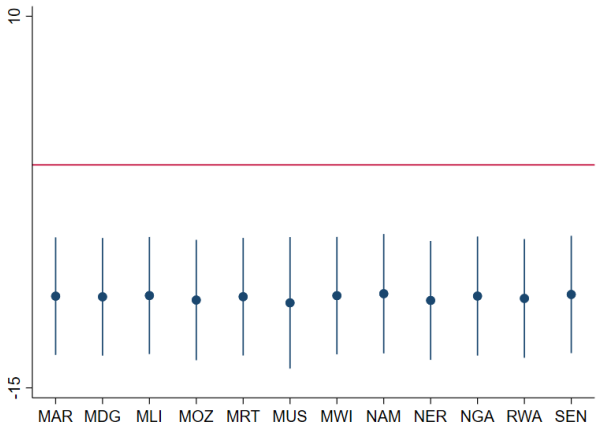
Figure 13: Robustness of first stage for World Bank Aid - Leaving one country out

Note: Results depict coefficients of the instrumental variable $probability_{i,c,t-2} \times IDAPosition_{t-1}$ for different regressions leaving one country out from the estimation. Labels in the graph refer to ISO codes of recipients. Click here to go back to section 5.6.

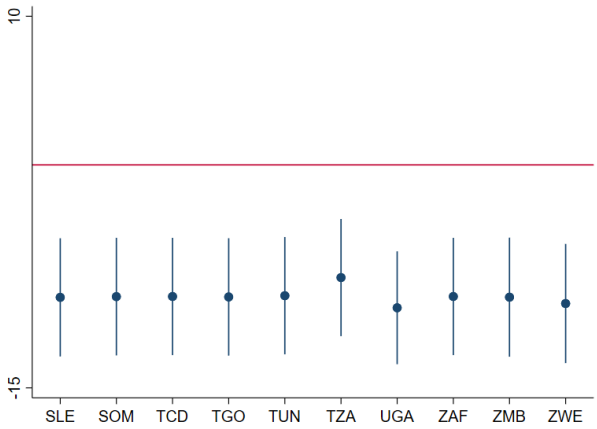


(a) Leave one out (i)

(b) Leave one out (ii)



(c) Leave one out (iii)



(d) Leave one out (iv)

Figure 14: Robustness of first stage for Chinese Aid - Leaving one country out

Note: Results depict coefficients of the instrumental variable $probability_{i,c,t-3} \times ln(Chinese\,Commodity_{t-3})$ for different regressions leaving one country out from the estimation. Labels in the graph refer to ISO codes of recipients. Click here to go back to section 5.6.

Table 25: ADM1 IV (WB - Global Time Series)

	(1)	(2)
Panel A: WB Second Stage		
IV Second Stage: World Bank		
	Glob. Conflict	Glob. Conflict
$\ln(\text{World Bank Aid}_{t-1})$	-0.1484	-0.0605
	(0.3637)	(0.3788)
N	12325	12325
Kleibergen-Paap under-ID p-val.	0.000	0.000
Kleibergen-Paap weak ID F-stat	99.771	99.659
Panel B: WB First Stage		
IV First stage: World Bank		
$\text{IDA Position}_{t-1} \times \text{Cum. Prob}_{t-2}$	76.5677***	95.7504***
	(7.6581)	(9.6052)
$\ln(\text{BRD}_{t-1}) \times \text{Cum. Prob}_{t-2}$	-0.9489	-2.9929***
	(0.5927)	(1.0353)
N	12325	12325
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. All regressions consider interactions of the cumulative probability of aid receipts (as a proxy for treatment intensity) and global battle related deaths, which could potentially cause spurious trends. All regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. Click here to go back to section 5.6. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

Table 26: ADM1 IV (China - Global Time Series)

	(1)	(2)
Panel A: China Second Stage		
IV Second Stage: China		
	Glob. Conflict	Glob. Conflict
$\ln(\text{Chinese Commodity}_{\text{t-2}})$	-0.2070	-0.1037
	(0.4766)	(0.5948)
N	7975	7975
Kleibergen-Paap under-ID p-val.	0.000	0.000
Kleibergen-Paap weak ID F-stat	31.432	26.756
Panel B: China First Stage		
IV First stage: China		
$\text{Chinese Commodity}_{\text{t-3}} \times \text{Cum. Prob}_{\text{t-3}}$	-13.0574***	-11.8020***
	(2.3254)	(2.2780)
$\ln(\text{BRD}_{\text{t-3}}) \times \text{Cum. Prob}_{\text{t-3}}$	2.6516	2.1932
	(2.3520)	(2.3782)
N	7975	7975
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. All regressions consider interactions of the cumulative probability of aid receipts (as a proxy for treatment intensity) and global battle related deaths, which could potentially cause spurious trends. All regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. Click here to go back to section 5.6. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

B.2 Alternative Outcome Variables

Robustness of results on lethal violence (UCDP measures)

As thresholds of five battle-related deaths or one incidence per region-year are arbitrary, we depict for robustness also other intensity thresholds. First, aid could matter for rather more intense conflicts in line with the evidence on conflict dynamics made by Bluhm et al. (2018). Tables 27 (OLS) and 28 (IV) indicate for a higher threshold of 25 battle-related deaths mainly insignificant coefficients, which also remain negative for the few significant OLS results. While the IV specifications indicate medium-sized negative for the WB and small positive coefficients for China, both stay insignificant. Second, this also holds in Tables 29 (OLS) and 30 (IV) when using a continuous measure of logarithmized battle-related deaths.

Robustness of results on non-lethal violence (SCAD)

The measurement of conflict is non-trivial and in this respect we display in the main part beyond lethal violence measures of social conflict based on Salehyan et al. (2012). Both anecdotal evidence and research studies alike suggest increased social conflict linked to Chinese investment activities. We take these concerns serious by disentangling the results from Table 6 from the main part. We consider the effects on demonstrations, riots and strikes separately with OLS in Tables 31 ,32 and 33 as well as using IV in Table 34. Results do not correspond to a statistically significantly positive effect of aid on neither riots, demonstrations and strikes. An explanation could be that these accounts mostly cover commercial investment activities, which are not conflict sensitively programmed (Wegenast et al., 2017; Christensen, 2017).

Additionally, we consider robustness of the main results relating to repression fueling effects of Chinese aid. First, to separate clearly between regions with lethal pro-government and non-lethal pro-government activities, we constrain the sample on regions, which *did not* encounter any one-sided violence by the government registered in the UCDP dataset. Results in Table 35 support a robust link between Chinese aid and repression. Second, when using instead of a dichotomous repression measure from SCAD a continuous indicator, a consistently positive effect of Chinese aid on repression is suggested by the IV estimates of Table 36.

Table 27: ADM1 OLS results (Intensity 2)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Panel A: WB Aid									
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1061	-0.0440	-0.0703	-0.1810***	-0.1522**	-0.1528**	-0.0544	-0.1386*	-0.1453
	(0.0659)	(0.0551)	(0.0536)	(0.0528)	(0.0669)	(0.0668)	(0.0747)	(0.0764)	(0.0927)
N	13104	13104	13104	13104	13050	13050	11017	13050	11017
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	-0.0917	-0.0209	0.0184	-0.0285	-0.0140	0.0059	-0.0001	-0.0022	-0.0099
	(0.0614)	(0.0504)	(0.0378)	(0.0446)	(0.0530)	(0.0496)	(0.0543)	(0.0568)	(0.0645)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥25, 0 if BRD<25). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 28: IV (Intensity 2)

	(1)	(2)
Panel A: WB Aid		
IV Second Stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1437	-0.4581
	(0.3075)	(0.3301)
IV First stage: World Bank		
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	70.9363***	80.8832***
	(7.1065)	(8.6854)
N	12325	12325
Panel B: Chinese Aid		
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	0.1289	0.1652
	(0.2757)	(0.3140)
N	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.578	31.190
IV First stage: China		
$Chinese\,Commodity_{\text{t-3}} \times \text{Cum. Prob}_{\text{t-3}}$	-14.0193***	-12.6964***
	(2.3180)	(2.2734)
N	7975	7975
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥25, 0 if BRD<25). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. The constituent term of the probability is depicted in the appendix.
* $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 29: OLS results (Battle-related Deaths)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Panel A: WB Aid									
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.0164*	-0.0014	-0.0025	-0.0174***	-0.0165**	-0.0142*	-0.0019	-0.0142*	-0.0100
	(0.0092)	(0.0071)	(0.0065)	(0.0060)	(0.0068)	(0.0074)	(0.0083)	(0.0081)	(0.0093)
N	13104	13104	13104	13104	13050	13050	11017	13050	11017
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	-0.0119	0.0034	0.0068	-0.0055	-0.0008	0.0004	0.0007	0.0034	0.0029
	(0.0087)	(0.0065)	(0.0054)	(0.0048)	(0.0072)	(0.0066)	(0.0068)	(0.0064)	(0.0071)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: The table displays regression coefficients with the log of battle-related deaths + 0.01 as dependent variable (category 3). Standard errors in parentheses, two-way clustered at the country-year and regional level. The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 30: IV (Battle-Related Deaths)

	(1)	(2)
Panel A: WB Aid		
IV Second stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.0179	-0.0340
	(0.0340)	(0.0358)
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	99.639	86.724
IV First stage: World Bank		
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	70.9363***	80.8832***
	(7.1065)	(8.6854)
N	12325	12325
Panel B: Chinese Aid		
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	-0.0312	-0.0180
	(0.0337)	(0.0420)
N	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.457	31.087
IV First stage: China		
$Chinese\,Commodity_{\text{t-3}} \times \text{Cum. Prob}_{\text{t-3}}$	-19.9994***	-19.9994***
	(4.5891)	(4.5891)
N	1311	1311
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The table displays regression coefficients for the log of battle-related deaths +0.01 as dependent variable (category 3). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. The constituent term of the probability is depicted in the appendix.
* $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 31: OLS results (Demonstrations)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Panel A: WB Aid									
$ln(World\,Bank\,Aid_{t\text{-}1})$	0.0578	0.1247*	0.3399***	0.0514	0.0414	0.0491	-0.0224	0.0390	0.0364
	(0.0684)	(0.0708)	(0.0705)	(0.0472)	(0.0699)	(0.0763)	(0.0816)	(0.0745)	(0.0824)
N	13104	13104	13104	13104	13050	13050	11017	13050	11017
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{t\text{-}2})$	0.7830***	0.8995***	0.9203***	-0.1090	-0.0865	-0.0781	-0.0704	-0.1094	-0.0888
	(0.1899)	(0.1649)	(0.1700)	(0.0766)	(0.0919)	(0.0985)	(0.1011)	(0.1233)	(0.1236)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: The table displays regression coefficients with a binary indicator for demonstrations as dependent variable. The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.4.

Table 32: OLS results (Riots)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Panel A: WB Aid									
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.0920	0.0037	0.2350***	0.0129	-0.0060	-0.0060	-0.0831	-0.0853	-0.1080
	(0.0620)	(0.0856)	(0.0617)	(0.0533)	(0.0559)	(0.0617)	(0.0682)	(0.0804)	(0.1049)
N	13104	13104	13104	13104	13050	13050	11017	13050	11017
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	0.4258***	0.5248***	0.5289***	0.0006	0.0399	0.0316	0.0521	0.0424	0.0613
	(0.1482)	(0.1261)	(0.1292)	(0.0814)	(0.0956)	(0.0986)	(0.0991)	(0.1200)	(0.1313)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: The table displays regression coefficients with a binary indicator for riots as dependent variable. The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.4.

Table 33: OLS results (Strikes)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Panel A: WB Aid									
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.0020	0.0302	0.1288***	-0.0197	-0.0252	-0.0377	-0.0549	-0.0717	-0.0758
	(0.0310)	(0.0391)	(0.0377)	(0.0309)	(0.0445)	(0.0578)	(0.0656)	(0.0582)	(0.0695)
N	13104	13104	13104	13104	13050	13050	11017	13050	11017
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	0.1611*	0.1832**	0.1931**	-0.1785**	-0.2042**	-0.1845*	-0.1800*	-0.1620	-0.1605
	(0.0847)	(0.0810)	(0.0846)	(0.0712)	(0.0887)	(0.1043)	(0.1036)	(0.1073)	(0.1122)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: The table displays regression coefficients with a binary indicator for strikes as dependent variable. Standard errors in parentheses, two-way clustered at the country-year and regional level. The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.4.

Table 34: IV (Riots, Demonstrations & Strikes [SCAD])

Panel A: WB Aid	(1)	(2)	(3)	(4)	(5)	(6)
IV Second stage: World Bank						
	Demonstr.	Demonstr.	Riots	Riots	Strikes	Strikes
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.2232	-0.1458	0.0106	-0.1950	0.0289	-0.0184
	(0.2514)	(0.2808)	(0.2543)	(0.2294)	(0.1793)	(0.1463)
N	12325	12325	12325	12325	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000	0.000	0.000	0.000	0.000
Kleibergen-Paap weak identification F-statistic	99.639	86.724	99.639	86.724	99.639	86.724
Panel B: Chinese Aid						
IV Second Stage: China						
	Demonstr.	Demonstr.	Riots	Riots	Strikes	Strikes
$\ln(Chinese\,Aid_{\text{t-2}})$	0.0498	0.0686	-0.0629	0.0424	-0.1489	-0.0776
	(0.4018)	(0.4707)	(0.3622)	(0.4312)	(0.4183)	(0.5076)
N	7975	7975	7975	7975	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000	0.000	0.000	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.578	31.190	36.578	31.190	36.578	31.190
Country-Year FE	No	Yes	No	Yes	No	Yes

Notes: The table displays regression coefficients for any violence of these three types as dependent variable. The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. OLS results are depicted in the appendix. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ 5.4

Table 35: ADM1 IV (Repression (non-lethal) - Regions with UCDP violence against civilians coded as zero)

	(1)	(2)
Panel A: WB Aid		
IV: World Bank - Actors		
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.1543	0.0885
	(0.1042)	(0.1177)
N	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	99.639	86.724
Panel B: Chinese Aid		
IV: Chinese Commodity - Actors		
$\ln(Chinese\,Aid_{\text{t-2}})$	0.6103**	0.7696**
	(0.2873)	(0.3439)
N	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.578	31.190
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The table displays regression coefficients for a binary pro-governmental violence indicator as dependent variable. The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. The constituent term of the probability is depicted in the appendix. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.4.

Table 36: Non-lethal Repression [SCAD] - Continuous measure

Panel A: WB Aid	(1)	(2)
IV Second stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.0011	0.0012
	(0.0014)	(0.0013)
N	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	99.639	86.724
Panel B: Chinese Aid		
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	0.0072**	0.0092**
	(0.0032)	(0.0045)
N	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.457	31.087
Country-Year FE	No	Yes

Notes: The table displays regression coefficients for a continuous measure of non-lethal pro-government violence as dependent variable. The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.4.

Comparison with OLS estimates

Finally, in order to see if results substantially change when using OLS, we consider the results corresponding to the IV estimates on actors (Table 5) and the aggregated outcome for riots, demonstrations and strikes (Table 6). Table 37 suggests mostly neutral effects, while significantly negatively coefficients of WB aid occur for state-based violence. Regarding riots, demonstrations and strikes, Table 38 shows that the different actors' results become insignificant once we condition on regional level fixed effects.

Table 37: Actors

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Panel A: WB Aid - OLS								
OLS: WB - Actors	State vs N-State		N-State vs N-State		State vs Civilans		N-State vs Civilians	
$ln(World\,Bank\,Aid_{t\text{-}1})$	-0.1229*	-0.1365*	-0.0348	-0.0784	-0.0596	-0.0372	-0.1040**	-0.0979*
	(0.0650)	(0.0707)	(0.0492)	(0.0679)	(0.0452)	(0.0430)	(0.0521)	(0.0578)
N	13050	13050	13050	13050	13050	13050	13050	13050
Panel B: Chinese Aid - OLS								
OLS: China - Actors	State vs N-State		N-State vs N-State		State vs Civilans		N-State vs Civilians	
$ln(Chinese\,Aid_{t\text{-}2})$	-0.0009	0.0122	-0.0162	0.0016	-0.0702	-0.0625	-0.0338	-0.0334
	(0.0548)	(0.0663)	(0.0554)	(0.0769)	(0.0483)	(0.0542)	(0.0349)	(0.0439)
N	8700	8700	8700	8700	8700	8700	8700	8700
Country-Year FE	No	Yes	No	Yes	No	Yes	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Conflicts are considered for the WB from 1996 to 2013 and for Chinese aid from 2002 to 2014 due to the lag structure. Exogenous (time-varying) controls are included in all regressions. Time Trends included, consist of linear and squared country-specific time trends as well as linear regional time trends. "State vs N-State" refers to state-based violence against non-government actors, "N-State vs N-State" refers to non-government violence against the other organized non-state groups, and "State vs Civilians" refers to one-sided violence versus civilians by the government and "N-State vs. Civilians" refers to one-sided violence versus civilians by non-government (NG) actors. The categories are mutually exclusive. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to Table 5.

Table 38: OLS results (Riots, Demonstrations & Strikes [SCAD])

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Panel A: WB Aid									
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.1194	0.1291	0.4360***	0.0106	-0.0140	-0.0035	-0.1421	-0.0092	-0.0447
	(0.0912)	(0.1028)	(0.0885)	(0.0641)	(0.0751)	(0.0848)	(0.1063)	(0.0954)	(0.1133)
N	13104	13104	13104	13104	13050	13050	11017	13050	11017
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	0.8761***	1.0301***	1.0445***	-0.1026	-0.0468	-0.0182	-0.0009	0.0141	0.0387
	(0.2247)	(0.1888)	(0.1939)	(0.0880)	(0.1027)	(0.1050)	(0.1013)	(0.1268)	(0.1301)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: The table displays regression coefficients with a binary indicator for any violence of these three types as dependent variable. The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.4.

B.3 Channels - Aid Sectors

As outlined before, aid in different sectors could be more or less likely to fuel or calm down a conflict. We examine aid projects in eight subcategories for our two preferred specifications, with and without country-year FE. Note that, in almost all cases, the country-year FE only affect the coefficients' sizes, not their signs. For the WB, the IV strategy works well using sector-specific probabilities. For China, there are severe weak IV problems due to limited observations in certain sectors. We show results for China using OLS, building on the fact that OLS and IV results turned out to be very similar before.

Interesting differences across sectors emerge, suggesting that aid in different sectors indeed has different effects on subsequent conflict. Table 39 shows that there are positive coefficients of WB (Chinese) aid in a few categories, but it never becomes statistically significant. The insignificant negative average effects in previous tables seem to be driven by significantly negative, conflict-reducing effects for the sectors "finance" (WB only) and "transportation" (WB and China), both in the less and more restrictive specification with country-year FE. In the latter specification, a 100% increase in WB finance aid leads to a 1.59 percentage point reduction in the likelihood of conflict – relative to the baseline likelihood of 12 percent. Our investigation of a sample of the 1,361 projects in this sector shows that finance projects typically support both existing, and new projects to induce structural or sectoral reforms; these projects also provide technical assistance and consulting, concerning topics like regulation and financial or business services.[45] The actual monetary disbursements are rather small; hence, the main impact must stem from the knowledge transfer and technical support to modernize and develop capital markets, banks and insurances. This includes privatization programs, the development and restructuring of banks, as well as technical assistance to enhance transparency and regulation.

Regarding the transportation sector, a 100% increase in WB (Chinese) aid leads to a 6 (3.4) percentage points reduction in the likelihood of conflict . This sector comprises many projects, often large-scale infrastructure projects, as well as large disbursements in dollar terms. The negative effect suggests that existing constraints on movement or high transportation costs were significant obstacles for exchange, consumption, public good provision and eventually economic growth (see also Berman and Couttenier, 2015; Storeygard, 2016). This seems to dominate both potentially negative effects on corruption (Isaksson and Kotsadam, 2018a), and disputes over land usage. It is in line with Bluhm et al. (2018), who show that Chinese infrastructure projects reduce economic inequality and, hence, potential reasons for conflict.[46]

[45] Out of 40 projects, 26 were in one of those categories. Appendix section A.5 documents how we retrieve detailed information on World Bank aid in the finance sector.

[46] While the improvements in transportation infrastructure are likely to be linked to a higher accessibility for the media and correlated also with mobile phone coverage, this would induce an upward bias to our estimates (Weidmann, 2016; Von Borzyskowski and Wahman, 2019).

Overall, these heterogeneities across aid categories within the sector-specific results are a first explanation for the relatively broad confidence interval when studying the average effect of WB and Chinese aid. It is reassuring that we find no significant conflict-fueling effect on any aid sector for either one of the two donors. The overall average negative relationship does not seem to mask strong conflict-fueling effects in certain sectors.[47]

[47] Table A52 presents the regressions for the WB using OLS and for China using IV. The OLS results differ in some cases with regard to the sign of the effect, but again there is no significant positive effect for any sector. Note that generally, one caveat of these regressions is that due to high collinearity and insufficient power we cannot run the regressions with all individual sectoral aid variables jointly included, in particular with IV estimators.

Table 39: Aid sectors and conflict

World Bank Aid Sectors - IV	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
Panel A: No Country-Year FE	AX	BX	CX	EX	FX	JX	LX	TX	WX	YX
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.2179	-0.2102	0.3423	0.5525	-1.6744**	0.2773	-0.1658	-0.7843**	0.5021	-0.4463
	(0.3572)	(0.4195)	(0.3016)	(0.4572)	(0.7877)	(0.4321)	(0.2858)	(0.3323)	(0.5593)	(0.3647)
Kleibergen-Paap underidentification test p-value	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Kleibergen-Paap weak identification F-statistic	58.309	80.342	39.353	50.568	16.781	73.307	33.666	64.555	40.026	31.887
Panel B: Country-Year FE										
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.4793	-0.4087	0.2652	0.2253	-1.5963*	0.2952	-0.1206	-0.6667*	-0.2726	-0.3717
	(0.3152)	(0.4445)	(0.2709)	(0.4771)	(0.9361)	(0.4020)	(0.2764)	(0.3570)	(0.6850)	(0.3299)
N	12325	12325	12325	12325	12325	12325	12325	12325	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Kleibergen-Paap weak identification F-statistic	59.949	61.188	56.632	31.111	12.238	73.686	36.219	28.587	23.180	33.957
Chinese Aid Sectors - OLS										
Panel C: No Country-Year FE	AX	BX	CX	EX	FX	JX	LX	TX	WX	YX
$ln(Chinese\,Aid_{\text{t-2}})$	-0.3165	-0.2123	0.1770	-0.0830	N.A.	-0.0168	0.3516	-0.2780*	-0.2974	0.8388
	(0.2007)	(0.1391)	(0.1325)	(0.1637)	(N.A.)	(0.1448)	(0.2661)	(0.1611)	(0.1935)	(0.8093)
Panel D: Country-Year FE										
$ln(Chinese\,Aid_{\text{t-2}})$	-0.1946	-0.1881	0.1281	-0.0484	N.A.	0.0287	0.3241	-0.3378*	0.0377	0.7787
	(0.2239)	(0.1434)	(0.1329)	(0.1703)	(N.A.)	(0.1561)	(0.2848)	(0.2018)	(0.2138)	(0.7893)
N	8700	8700	8700	8700	8700	8700	8700	8700	8700	8700

Notes: The dependent variable is a binary conflict indicator (100 if BRD$\geq$5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Conflicts are considered for the WB from 1996 to 2013 and for Chinese aid from 2002 to 2014 due to the lag structure. Regressions account for (time-varying) exogenous controls and time trends. Time Trends include linear and squared country-specific time trends as well as a linear regional trend. AX - "Agriculture, fishing, and forestry" BX - "Public Administration, Law, and Justice" CX - "Information and communications" EX - "Education" FX - "Finance" JX - "Health and other social services" LX - "Energy and mining" TX - "Transportation" WX - "Water, sanitation and flood protection" YX - "Industry and Trade" Standard errors in parentheses, two-way clustered at the country-year and regional level: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

B.4 Channels - Ethnic groups and governing coalition

Conflicts are not only driven by economic considerations, but often strongly influenced by existing cleavages between groups. Ethnic identities are among the most salient traits and ethnicities constitute a very important reference group in most African countries. To measure ethnic homelands, we use the GREG dataset (Weidmann et al., 2010), which is a georeferenced version of the initial locations of ethnic homelands based on the Soviet Atlas Narodov Mira. These locations were determined before our sample, and, even though immigration becomes more important over time, prior studies suggest that a large share of Africans still live in their ethnic home region (Nunn and Wantchekon, 2011). This makes those group polygons a noisy, but still informative measure.

A first important question is whether the effect of aid projects differs between more and less ethnically fractionalized regions. Theoretically, one might expect more potential for dissatisfaction about an unequal allocation of projects or the distribution of the associated benefits in ethnically fractionalized regions. We compute standard fractionalization measures in line with the literature (Alesina and Ferrara, 2005; Fearon and Laitin, 2003), and split the sample between countries in regions with fractionalization above or below the median. Appendix Table 41 shows no large differences. When including country-year FE, the negative relationship between aid and conflict becomes even a bit stronger, but the difference is small. Even in the more fractionalized regions, it does not turn positive. [48]

More important than considering ethnic cleavages in general is to define which ethnic groups are allies and form a joint coalition and which groups are outside that coalition. To classify administrative regions, our unit of analysis, we distinguish whether all groups (Coalition), at least one group (Mixed), or no group (N-Coalition) in a region is part of the governing coalition in a particular year. The information about the power status comes from the time-variant Ethnic Power Relations (EPR) dataset (Vogt et al., 2015). Wherever possible, we match the group power status from EPR in a particular year to one of the time-invarying GREG group homelands. The original dataset assigns 8 different power statuses to groups. The difference are sometimes marginal and hard to interpret, which is why we only use the more precise information on whether a group was part of the governing coalition or not. We then intersect the ethnic group polygons with the administrative regions to classify regions as one of the three categories.

This distinction aims at testing the plausibility of the existing results, and at uncovering heterogeneous effects that might be hidden in the averages. For instance, it might be that there is no conflict-inducing effect on average. However, assuming that aid project benefit governing groups more often, existing tensions and conflict might be fueled especially in mixed districts where other

[48] Note that for individual aid sectors, the IV does not perform sufficiently well for China when splitting the samples. Therefore, we show the OLS specifications for all the sample splits for China. We intend to conduct a more in-depth analysis of aid inequality and ethnic groups in an accompanying paper.

groups observe these distributional differences. In contrast, rapacity theory would predict that governing coalition regions with large aid inflows become more attractive for rebels to capture.

We find several interesting differences in Table 40. The results for the WB always change signs depending on the inclusion of country-year fixed effects. Nonetheless, there is again never a significant conflict-inducing effect. For China, both coefficients for mixed regions are positive. However, all coefficients are statistically insignificant. Even when considering governing coalition structures, on average Chinese aid does not increase conflicts with at least 5 BRDs.[49] Moreover, we control in all regressions for fractionalization, which we define in this case as $1-\sum s^2$, where s is the ethnic groups area share in the administrative region. In order to account for the important role that ethnic fractionalization takes in the politico-economic literature (e.g., Alesina et al., 2003), we consider also a sample split at the median of ethnic fractionalization in Table 41. In the subsample the instrumental variable retains strength. Although coefficients change signs, when considering the more fractionalized regions, results support robustness of the neutral effects.

[49] This finding is robust to defining the coalition only as the more powerful senior, dominant or monopoly groups and excluding junior partners. Results are available upon request from the authors.

Table 40: ADM1 results (Power status - Member of Coalition Group)

Panel A: WB - IV	(1)	(2)	(3)	(4)	(5)	(6)
Conflict in region belonging to ...	N-Coalition	N-Coalition	Coalition	Coalition	Mixed	Mixed
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.7052	0.2016	0.0686	-0.6372	0.1552	-0.3712
	(0.9362)	(1.3680)	(0.4500)	(0.4716)	(0.5181)	(0.5339)
N	2144	2075	3750	3651	4569	4537
Kleibergen-Paap underidentification test p-value	0.000	0.003	0.000	0.000	0.000	0.000
Kleibergen-Paap weak identification F-statistic	35.086	18.726	41.902	26.417	63.396	66.952
Panel B: China- IV:						
Conflict in region belonging to ...	N-Coalition	N-Coalition	Coalition	Coalition	Mixed	Mixed
$\ln(Chinese\,Aid_{\text{t-2}})$	-0.6513	0.7011	-0.7345	-1.2272	0.6919	1.1403
	(1.0808)	(3.4968)	(0.5935)	(0.7612)	(0.6681)	(0.9162)
N	1335	1285	2487	2420	2944	2924
Kleibergen-Paap underidentification test p-value	0.033	0.055	0.000	0.000	0.000	0.000
Kleibergen-Paap weak identification F-statistic	15.575	3.709	59.921	46.322	22.702	19.653
Country $\times$ Year FE	No	Yes	No	Yes	No	Yes
Control for Fractionalization	Yes	Yes	Yes	Yes	Yes	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD$\geq$5, 0 if BRD$<$5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Conflicts are considered for the WB from 1996 to 2013 and for Chinese aid from 2002 to 2014 due to the lag structure. Both regressions include (time-varying) exogenous controls, year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends as well as linear regional time trends. Columns (1) & (2) refer to all regions without members of the governing coalition, whereas columns (3) & (4) to mixed regions with some groups in and out of the coalition, and columns (5) & (6) to regions that contain groups exclusively from the coalition. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

Table 41: Sample-split: Median Fractionalization

Panel A: WB Aid - IV:				
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.2585	-0.6189	0.1471	-0.0455
	(0.4163)	(0.4904)	(0.5688)	(0.7054)
N	5474	5474	4998	4998
Kleibergen-Paap underidentification test p-value	0.000	0.000	0.000	0.000
Kleibergen-Paap weak identification F-statistic	71.721	49.454	75.067	65.391
Panel B: Chinese Aid - IV:				
$\ln(Chinese\,Aid_{\text{t-2}})$	-0.4831	-0.5251	0.0510	0.7714
	(0.5695)	(0.7265)	(0.6113)	(0.7163)
N	3542	3542	3234	3234
Kleibergen-Paap underidentification test p-value	0.000	0.000	0.000	0.000
Kleibergen-Paap weak identification F-statistic	51.569	38.166	23.501	20.763
Country $\times$ Year FE	No	Yes	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD$\geq$5, 0 if BRD$<$5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. The sample is split in regions, which are below the country level median/mean of ethnic fractionalization (0) [columns (1) & (2)] or above the median/mean (1) [columns (3) & (4)]. Ethnic fractionalization is based on $1 - \sum s^2$, where s is the ethnic groups area share in the administrative region. Conflicts are considered for the WB from 1996 to 2013 and for Chinese aid from 2002 to 2014 due to the lag structure. Both regressions include (time-varying) exogenous cont+rols, year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends as well as linear regional time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

B.5 Spatial Dimension (Spill Overs and Aggregation Levels)

Aggregation levels

Despite the many advantages of geospatial analysis (e.g., precision, geographical control variables), robustness is subject to the modifiable are unit problem (MAUP). More specifically, other conflict mechanisms can be at play when considering different levels of aggregation. Testing robustness on different spatial levels, hence, reduces the risk of ecological fallacy (Maystadt et al., 2014), which is specifically relevant in the aid-conflict nexus where different political entities might appropriate funds to engage in violent or peace-building activity. For this reason, we consider conflict and aid in the subordinate ADM2 regions both with OLS (IV) in Table 42 (43). Results are generally consistent with the main finding of a neutral effect of aid on conflict. Although the IV estimates for China turn positive, they do not attain statistical significance at any conventional level.

Additionally, we turn to an analysis on the country level as conflict might not manifest on the regional level, but spill over to other localities. Also on the country-level Table 44 does provide neither for the WB nor for China any evidence of a significant link between aid and conflict. While both OLS coefficients are negative, the WB IV coefficient turns positive, though insignificant. In order to address concerns that our analysis misses non-geocoded aid flows of the two donors, we make use of the feature that we can include those flows on a country-level. Consistently, results in Table 45 indicate significantly negative to neutral effects.[50] Only one of the coefficients for non-geocoded WB aid turns positive, though remaining statistically insignificant.[51]

[50] Ideally, we would have liked to consider results in Table 45 also via an instrumental variable approach, which was not possible due to weak IV concerns in the first stage.

[51] As those non-geocoded flows are mostly allocated to line ministries or the central government, we consider this question more specifically in the subsequent paragraph.

Table 42: ADM2 level OLS results (Intensity 1)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Panel A: WB Aid									
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.0288	0.0188	0.0068	-0.0740***	-0.0674***	-0.0580**	-0.0354	-0.0627**	-0.0535*
	(0.0209)	(0.0196)	(0.0219)	(0.0245)	(0.0234)	(0.0251)	(0.0294)	(0.0262)	(0.0316)
N	105354	105354	105354	105354	105214	105214	91333	105214	91333
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	0.0105	0.0104	0.0579*	-0.0392	-0.0499	-0.0410	-0.0455	-0.0501	-0.0500
	(0.0407)	(0.0402)	(0.0331)	(0.0318)	(0.0392)	(0.0327)	(0.0347)	(0.0449)	(0.0446)
N	76089	76089	76089	76089	70132	70132	64482	70132	64482
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes second order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 43: ADM2-level IV (Intensity 1)

	(1)	(2)
Panel A: WB Aid		
IV Second stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.2599	0.1522
	(0.1644)	(0.1171)
N	99367	99367
Panel B: Chinese Aid		
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	0.0517	0.0496
	(0.2007)	(0.2748)
N	64285	64285
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes second order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. The constituent term of the probability is depicted in the appendix. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 44: Aggregate - Cross-country Analysis - OLS/IV

Cross-Country Analysis $\ln(WB\,Aid_{\text{t-1}})$	-0.0025	0.0075		
	(0.0024)	(0.0106)		
$\ln(Chinese\,Aid_{\text{t-2}})$			0.0005	0.0118
			(0.0017)	(0.0132)
Kleibergen-Paap underidentification test p-value		0.000		0.002
Kleibergen-Paap weak identification F-statistic		52.295		9.056
OLS:	Yes	No	Yes	No
IV:	No	Yes	No	Yes

Notes: Dependent variable is a binary conflict indicator (100 if BRD≥25, 0 if BRD<25). Estimates refer to the country level, where aid and battle-related deaths were aggregated at the country level. Columns (1) and (2) depict OLS/IV coefficients for WB geocoded aid aggregated at the country level. Columns (3) and (4) depict OLS/IV coefficients for Chinese geocoded aid aggregated at the country level. This includes aid, which is coded at least at the ADM1 level (refer to Figure 1). The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. Conflicts are considered for the WB from 1996 to 2013 and for Chinese aid from 2002 to 2014 due to the lag structure. All regressions include year and country fixed effects. Standard errors in parentheses are clustered at the level of the country. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 45: Aggregate - Cross-country Analysis - OLS

	Geocoded	Non-Geocoded
$ln(WB\,Aid_{\text{t-1}})$	-0.3419	0.2110
	(0.4410)	(0.4843)
$ln(Chinese\,Aid_{\text{t-2}})$	-0.2081*	-0.1678
	(0.1158)	(0.1966)
R^2	0.318	
N	792	
Non-geocoded aid as control:	No	Yes

Notes: Dependent variable: Category 2 binary conflict indicator (100 if BRD≥25, 0 if BRD<25). Estimates refer to the country level, where aid and battle-related deaths were aggregated at the country level. Columns (1) and (2) refer to one regression. Column (1) depicts coefficient for geocoded aid aggregated at the country level. Column (2) depicts coefficients for non-geocoded aid, which is aid coded less precise than the ADM1 level (refer to Figure 1). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. The regression includes country and year fixed effects as well as a linear county-trend. Standard errors in parentheses are clustered at the level of the country. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

B.6 Mechanisms - Afrobarometer

Table 46: Mechanisms - Afrobarometer

	WB	WB	China	China
Panel A: Security				
Security facilities: Police station present within walking distance?	0.001	0.008*	0.002	-0.004***
	(0.003)	(0.003)	(0.002)	(0.003)
Security forces: Any policemen or police vehicles?	0.002	0.004	0.001	-0.002
	(0.002)	(0.003)	(0.002)	(0.002)
Security forces: Any soldiers or army vehicles?	0.002*	0.005***	-0.001	-0.003
	(0.001)	(0.003)	(0.001)	(0.002)
Frequency of things stolen in past year?	-0.001	-0.006**	0.004*	0.004***
	(0.002)	(0.002)	(0.002)	(0.002)
Frequency of phsysical attacks in the past year?	-0.000	-0.003***	0.001	-0.000
	(0.001)	(0.002)	(0.001)	(0.001)
Panel B: Democratic norms and attitudes				
Democracy: How democratic is your country today?	-0.002	0.003	-0.005*	-0.000
	(0.002)	(0.003)	(0.002)	(0.003)
Democracy: Did you perceive last elections as free and fair?	-0.003	-0.003	-0.012**	-0.012
	(0.005)	(0.007)	(0.004)	(0.008)
Governance: Reject one-party rule	0.003	0.013*	-0.006	-0.003
	(0.005)	(0.005)	(0.004)	(0.006)
Governance: Reject military rule	0.006*	0.008*	-0.002	-0.001
	(0.003)	(0.004)	(0.003)	(0.004)
Governance: Reject one-man rule	0.004*	0.006*	-0.005*	-0.005***
	(0.002)	(0.003)	(0.002)	(0.003)
Reject government banning organizations that go against its policies	0.005*	0.014**	-0.003	0.002
	(0.002)	(0.005)	(0.003)	(0.004)
Panel C: Government responsiveness and repression				
Frequency of contact to government official to express your view	0.003*	0.003***	-0.001	0.001
	(0.001)	(0.002)	(0.001)	(0.001)
Fear of political intimidation or violence during campaigns	-0.001	-0.008***	0.003	0.011**
	(0.003)	(0.004)	(0.003)	(0.003)
How often do people have to be careful about what they say in politics?	0.000	-0.005	0.002	-0.002
	(0.002)	(0.004)	(0.002)	(0.003)
Rule of Law: People must obey the law	-0.004*	-0.001	0.004**	0.007**
	(0.002)	(0.003)	(0.001)	(0.002)
Frequency of joining others to request government action			-0.006**	
			(0.002)	
Country FE	Yes	Yes	Yes	Yes
Region FE	No	Yes	No	Yes
Year FE	Yes	Yes	Yes	Yes

Notes: Significance levels: * 0.10 ** 0.05 *** 0.01. Click here to go back to section 5.5.

B.7 Estimations - Miscellaneous

Estimation approach

Data sets with many zero outcome observations can ask for different estimation approaches (Silva and Tenreyro, 2006). Therefore, we also consider a Poisson Pseudo Maximum Likelihood (PPML) estimator in Table 47. In line with the main findings results are mostly non-significant and have a

negative sign if turning statistically significant.[52] Due to the persistent nature of conflicts, the use of lagged dependent variables is a recurring topic in the conflict literature (e.g., Bazzi and Blattman, 2014). Table 49, thus, presents the results including a lagged dependent variable, extending the main model by a lagged conflict indicator:

$$C_{i,c,t} = \beta_1 A_{i,c,t-1/t-2} + \beta_2 C_{i,c,t-1} + \lambda_c + \tau_t + \delta_i + \lambda_c T + \lambda_c T^2 + X^{Ex}_{i,c,t}\beta_2 + \delta_i T + X^{En}_{i,c,t-2}\beta_3 + \kappa_{c,t} + \epsilon_{i,c,t}, \quad (5)$$

None of the coefficients in Table 49 is positive, stressing the robustness of our main findings.

Although less often considered, the choice of standard error clustering can affect results substantially. Tables 50 and 51, thus, depart from our use of two-way clustering on the country-year and regional level, but only cluster on the region. Despite this adaptation, results ensure us that the insignificant findings are not driven by our choice of standard error clustering.

Table 47: PPML

	(1)	(2)	(3)
Panel A: WB Aid			
$ln(World\,Bank\,Aid_{t\text{-}1})$	-0.0005	0.0178	-0.0171
	(0.0063)	(0.0149)	(0.0173)
N	6246	1476	7344
Panel B: Chinese Aid			
$ln(Chinese\,Aid_{t\text{-}2})$	-0.0128*	0.0023	-0.0328*
	(0.0076)	(0.0131)	(0.0189)
N	3783	962	4589

Notes: Dependent variables- In column (1) a binary conflict indicator (100 if BRD≥5, 0 if BRD<5), in column (2) a binary indicator if any event of non-lethal pro-government violence took place, in column (3) a continuous measure of logged battle-related deaths. Standard errors in parentheses, clustered at the regional level. The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. All regressions include year fixed effects. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

[52] A clear caveat is that we can only use year fixed effects with PPML in our setting due to convergence issues. Thus, as results do not differ substantially, we rely in the main part on OLS and instrumental variable estimators.

Table 48: Conditional Logit

	(1)	(2)	(3)
Panel A: WB Aid			
	Intensity 1 (Dummy)	Non-lethal repression (Dummy)	No. of BRD
main			
ln(World Bank Aid $_{t-1}$)	0.0011	0.0201	-0.0035
	(0.0098)	(0.0182)	(0.0086)
R^2			
N	6192	1476	7218
Panel B: Chinese Aid			
	Intensity 1 (Dummy)	Non-lethal repression (Dummy)	No. of BRD
main			
ln(Chinese Aid $_{t-2}$)	-0.0201*	0.0026	-0.0099
	(0.0113)	(0.0178)	(0.0103)
R^2			
N	3731	962	4472

Notes: Dependent variables- In column (1) a binary conflict indicator (100 if BRD≥5, 0 if BRD<5), in column (2) a binary indicator if any event of non-lethal pro-government violence took place, in column (3) a continuous measure of logged battle-related deaths. Standard errors in parentheses, clustered at the regional level. The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. All regressions include year fixed effects. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 49: OLS results: Lagged dependent variable

Panel A: WB Aid	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.0844	-0.0069	-0.0173	-0.1659***	-0.1575**	-0.1406**	-0.0350	-0.1647**	-0.1355
	(0.0520)	(0.0551)	(0.0458)	(0.0585)	(0.0618)	(0.0707)	(0.0812)	(0.0808)	(0.1025)
N	13104	13104	13104	13104	13050	13050	11017	13050	11017
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	-0.0965*	-0.0300	-0.0082	-0.0983*	-0.0634	-0.0661	-0.0645	-0.0345	-0.0365
	(0.0563)	(0.0589)	(0.0588)	(0.0589)	(0.0771)	(0.0871)	(0.0921)	(0.1029)	(0.0913)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: Dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). This regression controls for the first lag of the binary indicator. Standard errors in parentheses, two-way clustered at the country-year and regional level. The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. Applying the lag structure of our regression equation, this means that conflicts are considered for the WB from 1996 to 2013 and for China from 2002 to 2014. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.6.

Table 50: ADM1 OLS results (Clustering at regional level)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Panel A: WB Aid									
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1918***	0.0010	-0.0496	-0.2129***	-0.2057***	-0.1608**	-0.0419	-0.1772**	-0.1420
	(0.0709)	(0.0643)	(0.0666)	(0.0611)	(0.0624)	(0.0672)	(0.0775)	(0.0799)	(0.0906)
N	13104	13104	13104	13104	13050	13050	11017	13050	11017
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	-0.1753**	-0.0233	-0.0026	-0.1090**	-0.0663	-0.0654	-0.0641	-0.0347	-0.0369
	(0.0761)	(0.0664)	(0.0676)	(0.0540)	(0.0605)	(0.0680)	(0.0687)	(0.0743)	(0.0757)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: The table displays regression coefficients with low Intensity Conflict (>5 battle-related deaths) as dependent variable. Standard errors in parentheses, clustered at the regional level. The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 4.1.

Table 51: ADM1 IV (Clustering at Regional Level)

	(1)	(2)
Panel A: WB Aid		
IV Second stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1014	-0.2252
	(0.3276)	(0.3899)
N	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	237.269	132.466
Panel B: Chinese Aid		
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	-0.2582	-0.1886
	(0.4169)	(0.5232)
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	55.714	41.016
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: Dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). Standard errors in parentheses, clustered at the regional level. The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. The constituent term of the probability is depicted in the appendix.
* $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 4.1.

Definition of aid (Sectors and weighting scheme)

Table 52 reports the OLS/IV estimates corresponding to sectoral aid in Table 39. Although significance is affected the negative signs in the transport and finance sectors are retained.

In order to attribute aid across different localities of a given project, we have to make assumptions. In the main part of this paper we assume an equal distribution across localities. An alternative and plausible assumption would be a weighting scheme according to population size. Tables 53 and 54 implement the alternative measure, indicating that results are not driven by this assumption.

Table 52: ADM1 - Aid Subtypes

WB Aid Subtypes - OLS	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
Panel A: No Country-Year FE	AX	BX	CX	EX	FX	JX	LX	TX	WX	YX
$ln(World\,Bank\,Aid_{\text{t-1}})$	0.0293	-0.1873**	0.1229	0.0215	-0.0958	-0.1575**	0.0236	-0.1479**	-0.0339	-0.1125
	(0.0753)	(0.0918)	(0.1575)	(0.0793)	(0.0919)	(0.0798)	(0.0941)	(0.0729)	(0.0898)	(0.0951)
Panel B: Country-Year FE										
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.0617	-0.2672***	0.0048	-0.0209	-0.0912	-0.1667*	-0.0317	-0.1137	0.0013	-0.2080*
	(0.0950)	(0.1031)	(0.1790)	(0.1062)	(0.1474)	(0.0977)	(0.1043)	(0.1021)	(0.1131)	(0.1139)
N	13050	13050	13050	13050	13050	13050	13050	13050	13050	13050
Chinese Aid Subtypes - IV										
Panel C: No Country-Year FE	AX	BX	CX	EX	FX	JX	LX	TX	WX	YX
$\ln(Chinese\,Aid_{\text{t-2}})$	-1.4578**	-2.5352	-0.5066	0.7578	N.A.	-0.1554	1.0265**	-0.4216	0.1977	-7.0545
	(0.7314)	(1.6254)	(0.3691)	(0.6847)	(N.A.)	(0.3838)	(0.4776)	(0.3463)	(0.5326)	(102.9015)
Kleibergen-Paap underid. test p-value	0.176	0.015	0.472	0.120	N.A.	0.062	0.214	0.028	0.554	0.101
Kleibergen-Paap weak id. F-statistic	1.712	11.768	0.484	3.225	N.A.	4.727	1.718	6.006	0.318	7.075
Panel D: Country-Year FE										
$\ln(Chinese\,Aid_{\text{t-2}})$	-1.0048	-2.2549	-0.3010	0.8156	N.A.	0.1259	0.9374*	-0.4798	0.5000	1.3334
	(0.8080)	(1.8415)	(0.3837)	(0.7330)	(N.A.)	(0.4240)	(0.5226)	(0.4277)	(0.6383)	(2.8884)
N	8700	8700	8700	8700	N.A.	8700	8700	8700	8700	8700
Kleibergen-Paap underidentification test p-value	0.232	0.011	0.674	0.095	N.A.	0.064	0.530	0.043	0.626	0.173
Kleibergen-Paap weak identification F-statistic	1.107	10.909	0.156	2.936	N.A.	4.322	0.369	4.282	0.212	2.467

Notes: Dependent variable: Category 1 binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. Conflicts are considered for the WB from 1996 to 2013 and for Chinese aid from 2002 to 2014 due to the lag structure. Regressions account for (time-varying) exogenous controls and time trends. Time Trends include linear and squared country-specific time trends as well as a linear regional trend. AX - "Agriculture, fishing, and forestry" BX - "Public Administration, Law, and Justice" CX - "Information and communications" EX - "Education" FX - "Finance" JX - "Health and other social services" LX - "Energy and mining" TX - "Transportation" WX - "Water, sanitation and flood protection" YX - "Industry and Trade" Standard errors in parentheses, two-way clustered at the country-year and regional level: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section B.3.

Table 53: OLS results: Population Weighted Aid Allocation

Panel A: WB Aid	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1898*	0.0062	-0.0440	-0.2217***	-0.2153***	-0.1664**	-0.0457	-0.1867**	-0.1502
	(0.1005)	(0.0788)	(0.0692)	(0.0667)	(0.0712)	(0.0797)	(0.0856)	(0.0872)	(0.1066)
N	13104	13104	13104	13104	13050	13050	11017	13050	11017
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	-0.1776**	-0.0246	-0.0037	-0.1137**	-0.0718	-0.0696	-0.0679	-0.0390	-0.0408
	(0.0865)	(0.0704)	(0.0648)	(0.0576)	(0.0789)	(0.0833)	(0.0881)	(0.1021)	(0.0919)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: Dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). Standard errors in parentheses, two-way clustered at the country-year and regional level. The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. Conflicts are considered for the WB from 1996 to 2013 and for Chinese aid from 2002 to 2014 due to the lag structure. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 3.1.

Table 54: ADM1 IV: Population Weighted Aid Allocation

Panel A: WB Aid	(1)	(2)
IV Second Stage: World Bank		
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1026	-0.2286
	(0.3798)	(0.4256)
N	12325	12325
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	100.636	88.243
Panel B: Chinese Aid	**(1)**	**(2)**
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	-0.2613	-0.1903
	(0.4333)	(0.5306)
N	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.766	31.399
Country-Year FE	No	Yes

Notes: Dependent variable is a binary conflict indicator (100 if BRD$\geq$5, 0 if BRD$<$5). Standard errors in parentheses, two-way clustered at the country-year and regional level. The sample includes African countries for the sampling period of 1995-2012 for the WB and 2000-2012 for Chinese Aid. Both regressions include exogenous (time-varying) controls. Year and region fixed effects as well as time trends are included in all regressions. Time Trends include linear and squared country-specific time trends and a linear regional trend. The constituent term of the probability is depicted in the appendix. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 3.1.

Both donors

Comparing both donors jointly comes at the disadvantage of losing five years of observations for the WB and - linked to this - a reduction of IV strength. Although the coefficients remain largely negative or insignificant in Tables 55 (OLS) and 56 (IV), the effects for the WB becomes less negative. Tables 55 (OLS) and 56 (IV) indicate that this is mostly driven by the different sampling years, rather than attributable to strong interactions between the two donors. It is important to see in Table 56 that the respective first stages for both donors become weaker when trying to estimate them simultaneously, but the exogenous instruments remains significant for the respective donor (column 2). This further supports that the interaction terms capture a specific variation linked to the allocation process of the two donors, instead of general trends or conflict patterns in the receiving regions. Still, the K-P F-statistics of 3.5 in our preferred specification with country-year FE underlines why we chose to estimate both first stages separately.

Table A57 and Table A58 show that the results also hold when restricting the WB results to the same years data for Chinese aid is available, once for OLS and once for IV.

Table 55: OLS results - Both Donors

WB & Chinese Aid	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1460	0.0571	0.0808	-0.0603	-0.0973	0.0661	0.0674	-0.0793	-0.0948
	(0.1194)	(0.0951)	(0.0913)	(0.0864)	(0.0926)	(0.0904)	(0.0889)	(0.0979)	(0.0958)
$ln(Chinese\,Aid_{\text{t-2}})$	-0.1278	-0.0291	0.0070	-0.1060*	-0.0660	-0.0656	-0.0644	-0.0345	-0.0367
	(0.0854)	(0.0700)	(0.0590)	(0.0595)	(0.0787)	(0.0824)	(0.0880)	(0.1018)	(0.0912)
N	8736	8736	8736	8736	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.1.

Table 56: ADM1 IV - Both Donors (Intensity 1)

	(1)	(2)
IV Second Stage: World Bank		
$\ln(World\,Bank\,Aid_{\text{t-1}})$	-0.7029	-2.3839
	(1.0780)	(1.6965)
$\ln(Chinese\,Aid_{\text{t-1}})$	-0.2482	-0.1655
	(0.4319)	(0.5415)
Kleibergen-Paap underidentification test p-value	0.000	0.005
Kleibergen-Paap weak identification F-statistic	11.573	3.489
IV First stage: World Bank		
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	57.3235***	63.8053***
	(12.0425)	(24.1932)
$Chinese\,Commodity_{\text{t-3}} \times \text{Cum. Prob}_{\text{t-3}}$	-0.2181	-0.1051
	(0.6571)	(0.6166)
N	7975	7975
IV First stage: China		
$Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	-17.9057*	-10.1067
	(9.3878)	(13.2890)
$Chinese\,Commodity_{\text{t-3}} \times \text{Cum. Prob}_{\text{t-3}}$	-13.9921***	-12.7060***
	(2.3178)	(2.2742)
N	7975	7975
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. The constituent term of the probability is depicted in the appendix. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$ Click here to go back to section 5.1.

Table 57: OLS results: (WB Aid - Same Years as Chinese Aid)

Panel A: WB Aid	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.1505	0.0559	0.0811	-0.0606	-0.0976	0.0657	0.0672	-0.0795	-0.0949
	(0.1197)	(0.0949)	(0.0910)	(0.0864)	(0.0922)	(0.0906)	(0.0886)	(0.0981)	(0.0957)
N	8736	8736	8736	8736	8700	8700	8261	8700	8261
Panel B: Chinese Aid									
$ln(Chinese\,Aid_{\text{t-2}})$	-0.1753**	-0.0233	-0.0026	-0.1090*	-0.0663	-0.0654	-0.0641	-0.0347	-0.0369
	(0.0865)	(0.0705)	(0.0642)	(0.0572)	(0.0783)	(0.0827)	(0.0877)	(0.1015)	(0.0916)
N	9464	9464	9464	9464	8700	8700	8261	8700	8261
Country FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time Trends	No	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Region FE	No	No	No	Yes	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Exogeneous Controls × Year FE	No	No	No	No	Yes	Yes	Yes	Yes	Yes
Linear Regional Trends	No	No	No	No	No	Yes	Yes	Yes	Yes
Lagged Endogeneous Controls	No	No	No	No	No	No	Yes	No	Yes
Country × Year FE	No	No	No	No	No	No	No	Yes	Yes

Notes: DThe dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Conflicts are considered for the WB from 2002 to 2013 due to the lag structure. Time Trends include linear and squared country-specific time trends. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$

Table 58: ADM1 IV (WB Aid - Same Years as Chinese Aid)

	(1)	(2)
Panel A: WB Aid		
IV Second stage: IDA Position		
$ln(World\,Bank\,Aid_{\text{t-1}})$	-0.6227	-2.3417
	(1.0568)	(1.6897)
Kleibergen-Paap underidentification test p-value	0.000	0.005
Kleibergen-Paap weak identification F-statistic	22.619	6.960
IV First stage: World Bank		
$IDA\,Position_{\text{t-1}} \times \text{Cum. Prob}_{\text{t-2}}$	57.2759***	63.9080***
	(12.0429)	(24.2241)
N	7975	7975
Panel B: Chinese Aid		
IV Second Stage: China		
$\ln(Chinese\,Aid_{\text{t-2}})$	-0.2582	-0.1886
	(0.4282)	(0.5257)
N	7975	7975
Kleibergen-Paap underidentification test p-value	0.000	0.000
Kleibergen-Paap weak identification F-statistic	36.457	31.087
IV First stage: China		
$Chinese\,Commodity_{\text{t-3}} \times \text{Cum. Prob}_{\text{t-3}}$	-19.9994***	-19.9994***
	(4.5891)	(4.5891)
N	1311	1311
Exogeneous Controls	Yes	Yes
Exogeneous Controls × Year FE	Yes	Yes
Linear Regional Trends	Yes	Yes
Country-Year FE	No	Yes

Notes: The dependent variable is a binary conflict indicator (100 if BRD≥5, 0 if BRD<5). The sample includes first order subnational regions in African countries for the 1995-2012 (WB) and the 2000-2012 period (China). Standard errors in parentheses, two-way clustered at the country-year and regional level. Both regressions include year and region fixed effects as well as time trends. Time Trends include linear and squared country-specific time trends. The constituent term of the probability is depicted in the appendix. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$