Labor Market Polarization and The Great Divergence: Theory and Evidence*

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Abstract

Two of the most important features of advanced labor markets in the past quarter century are labor market polarization and the great divergence. The first of these concerns the growth of jobs in high and low wage categories and the disappearance of middle wage jobs. The second is an explicitly spatial theory about the intensificiation of development particularly at the high end in large, already developed cities relative to smaller, less developed cities. This paper addresses how the two phenomena are interrelated. The great divergence is typically contemplated in a two factor setting with skill-biased technical change. Labor market polarization is instead considered in an explicitly three-factor setting, specifically rejecting the simpler framework. We develop a theory in which the driving forces of labor market polarization alone give rise to both phenomena, building on Autor and Dorn (2013), Davis and Dingel (2014) and Davis and Dingel (forthcoming). Key to this is that the productivity advantages in large cities are biased toward high skilled tasks, so that a uniform shock to technology leads to labor market polarization with a biased impact on cities of different sizes, giving rise to the great divergence. We examine the model using detailed data for a sample of 117 French cities and find the patterns in the period 1994-2015 accord well with the theory.

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1 Introduction

Among the most important economic developments in the last three decades is labor market polarization – robust job growth at the bottom and top of the occupational skill distribution and a decline in middle skill jobs Acemoglu and Autor (2011). Researchers have documented labor market polarization in the United States (Autor et al., 2006; Autor and Dorn, 2013) and in many European countries (Goos and Manning, 2007; Goos et al., 2009).

Ongoing dramatic falls in the cost of automation, information and communications technologies loom large in the discussion of driving forces. On one hand, this is tied to the routinization hypothesis, whereby these technologies are particularly adept at substituting for routine tasks previously carried out by middle skill labor Autor et al. (2003). Likewise, advances in information and communications technologies also facilitated the growth in offshoring of tasks from high to low wage countries, notably China.

Researchers have begun to explore how exposure to affected occupations manifests itself in local labor markets (Autor et al., 2013; Goos et al., 2014; Acemoglu and Restrepo, 2017). Harrigan et al. (2016) explore how the effects of offshoring at the firm level are mediated by the export-import orientation of the firm and the presence of specific workers capable of translating offshoring opportunities into action.

The literature tying these impacts to local labor markets has made important contributions, but also has limitations. These have frequently assumed that some (e.g. Autor and Dorn, 2013) or all (Autor et al., 2013) types of skills are immobile across locations. Thus employment of some or all factors in a location, which should be an object of investigation, is instead an assumption. For a phenomenon such as labor market polarization, which develops over decades (see Barany and Siegel, 2018), this is an important limitation that needs to be relaxed.

Recently researchers have developed a literature exploring heterogeneity of workers or firms for cities in spatial equilibrium. These include Behrens et al. (2014), Gaubert (2018), and Davis and Dingel (2014, forthcoming). The latter of these provides the most suitable setting for our work.

Our theoretical work both replicates key prior theoretical results and goes beyond them. While our model departs in important ways from Davis and Dingel (2014) and Davis and Dingel (forthcoming), it likewise features log supermodularity of skills in city sizes and a skill premium that rises with city size. Similar to models of labor market polarization such as Autor and Dorn (2013) and Autor et al. (2013), our model features labor market polarization in the aggregate in response to technology and globalization shocks.

Our theory also goes beyond the prior work in a number of significant ways. First, with the distribution of all skills across space endogenous we are able to provide a condition for the larger city to have lower initial exposure to middle-skill jobs in equilibrium. Second we show that the forces that generate aggregate labor market polarization do so in both large and small cities considered separately. Third, we show that the response of the large and small city in our framework can differ

importantly. Specifically we provide a condition under which in response to the shock the large city has a sharper percentage point decline in middle-skill jobs in spite of its low initial exposure; a greater percentage point rise in high-skill jobs; and a smaller percentage point increase in low-skill jobs.

We take these theoretical predictions to the data for France in the period 1994-2015, employing a rich variety of data sources covering wages, employment, educational attainment, and geographical location, among other variables. Our empirical work provides solid support for our theory. Large cities are more skilled and specialize in higher skill activities as per log supermodularity. They also have a lower exposure to middle-skill jobs than smaller cities. In the face of shocks that generate labor market polarization, these common forces work themselves out differently in large and small cities. There is a sharper percentage point decline in middle-skill jobs in large cities in spite of the lower initial exposure. The polarization is itself polarized in the sense that large cities have a strong bias toward job growth in the high-skill sectors and likewise for small cities in the low skill sectors. In sum, city size is a strong conditioning factor in the local effects of labor market polarization.

2 Theory

In this section, we build a model integrating the core framework of job polarization in Autor and Dorn (2013) with the system of cities model of Davis and Dingel (2014). This allows to obtain in equilibrium key features of data. First of all, it permits to have a log-supermodular distribution of skills across cities. Second, polarization of the job market occurs (both in the aggregate and in each city) when the price of capital or offshored tasks decreases. Furthermore, we obtain that, when the comparative advantage of high-skilled workers to live and work in the large city is sufficiently large, the labor market polarization is stronger and high-skill biased in the large city – with more destruction of middle-skilled jobs and more creation of high-skilled jobs – despite an initially lower share of exposure to middle-skilled jobs.

2.1 The environment

The set of cities is $c \in \{1, 2\}$.[1] In each city, there is a continuum of locations $\tau \in [0, \infty)$.

The economy is populated by households that provide labor, consume and decide where to live and work.

Households. Each household inelastically provides 1 unit of labor but households have different skills that we denote by ω. ω is distributed on $[\underline{\omega}, \overline{\omega}]$ with a pdf $n(.)$ and $\underline{\omega} \geq 1$.

Households consume a single final good and 1 unit of housing. They freely choose where they live (the city c and the internal location $\tau \geq 0$). We denote the rental price of location (c, τ) by

[1] We extend our framework to N cities in Appendix A.2.

$r(c,\tau)$. We use the price of the final good as the numeraire and we normalize the prize of unoccupied locations to 0 so that $r(c,\tau) \geq 0$.

Households can also decide in which sector σ they work.

Finally, we denote by $f(c,\tau,\omega,\sigma)$ the endogenous pdf of the distribution of households.

Production. Production in this economy involves different sectors: final goods are produced out of capital Z and intermediate goods $\{h,m,l\}$ that are also produced within the economy. All goods are traded with zero transportation costs except non-traded housing.

Final goods. The final good is produced by a continuum of identical competitive firms. To produce, they use intermediate goods $\{h,m,l\}$ as well as Z.

The production function of the representative firm is:

$$Q = \left(a(h)q(h)^{\gamma} + \left(a(m)q(m)^{\frac{1}{\theta}} + a(z)Z^{\frac{1}{\theta}} \right)^{\gamma\theta} + a(l)q(l)^{\gamma} \right)^{1/\gamma} \tag{1}$$

where q_j and p_j, $j \in \{h,m,l\}$, are the quantity and the price of intermediate good j, p_Z is the price of capital and/or an offshorable input with the rest being technological parameters.

As we are using the final good of numeraire, the profits of the representative firms can be written as:

$$\Pi = Q - p(h)q(h) - p(m)q(m) - p(l)q(l) - p(z)Z \tag{2}$$

Intermediate goods. The intermediate goods $\{h,m,l\}$ are freely tradable across cities and are produced with a constant return to scale technology only using labor.

There is one sector to produce each of the $\{h,m,l\}$ goods. Consistently, we label sectors by $\sigma \in \{h,m,l\}$ where h stands for high, m for middle and and l for low skilled.

Each individual with skill ω, living in city c and in a location τ has a productivity:

$$H(\omega,\sigma,c)T(\tau) \tag{3}$$

Here τ denotes the distance from an ideal location inside a city. This can be interpreted in a variety of ways, including as commuting distance to a central business district, or as remoteness from the core of a productive cluster with positive spillovers. Thus $T(.)$ is a decreasing function identifying the cost in productivity of being remote from the most productive location in a city. In each city c, we assume that the supply of locations at least as good as τ is $S(\tau)$ with $S(0) = 0$, $S(.)$ strictly increasing and twice continuously differentiable.

$H(\omega,\sigma,c)$ is log supermodular in (ω,σ) and $H(.,\sigma,c)$ is increasing. For simplicity, we assume

H has the following functional form:

$$H(\omega, l, c) = A(l, c)\omega^{\phi} \text{ with } \phi \in (0, 1), \tag{4}$$

$$H(\omega, m, c) = A(m, c)\omega, \tag{5}$$

$$H(\omega, h, c) = A(h, c)e^{\eta\omega} \text{ with } \eta > 1. \tag{6}$$

As a result of log-supermodularity in (ω, σ), $H(\omega, m, c)/H(\omega, l, c)$ and $H(\omega, h, c)/H(\omega, m, c)$ are increasing functions of ω. In the end, individual skills matter for productivity even more for the high-skilled sector, while they matter much less for the low-skilled sector.

In the interest of simplicity we make the assumption of exogenous productivity differences between the two cities. Endogenous productivity differentiation across cities arises naturally in this type of framework, but is not the focus of this paper. Davis and Dingel (2014, forthcoming) provide alternative approaches to endogenous productivity differences across cities even when fundamentals are symmetric.[2]

Assumption 1. *City 1 has an absolute advantage in all sectors:*

$$A(1, j) > A(2, j) \text{ for } j \in \{l, m, h\}$$

and a comparative advantage in higher skilled sectors:

$$\frac{A(1, h)}{A(2, h)} > \frac{A(1, m)}{A(2, m)} > \frac{A(1, l)}{A(2, l)}.$$

Assumption 1 directly implies that $H(\omega, \sigma, 1) \geq H(\omega, \sigma, 2)$ for all (ω, σ).

Finally, we assume that there is perfect competition in all the three sectors so that in each sector the wage per efficiency unit of labor equals the price of the intermediate good $p(\sigma)$.

Capital good/offshoring intermediate good The intermediate good z is produced by transforming final goods using the following technology:

$$Z = \frac{1}{\xi} q, \tag{7}$$

where q is the amount of final goods and ξ is a technology parameter. There, perfect competition implies:

$$p_z = \xi. \tag{8}$$

[2]In appendix A.1, we provide a way to obtain endogenous productivity differences consistent with patterns of labor market polarization.

The intermediate good z has two interpretations. The first is that it is a capital good that substitutes for middle-paying labor as in Autor and Dorn (2013). Note that as in Autor and Dorn (2013), this capital good would fully depreciate with production. With this view, ξ is a parameter that governs the efficiency of producing the capital good. The second interpretation is that Z is imported intermediates and ξ is the terms of trade. In this offshoring interpretation, the driving force is technical progress in, for example, the Chinese intermediate export sector (which it trades for imports of final output).

2.2 Household decisions

Let us first investigate location and sector decisions by agents and how these decisions depend on factor prices.

The utility flow obtained by an agent with skill ω, location decisions (c, τ) and intermediate good sector σ is:

$$H(\omega, \sigma, c)T(\tau)p(\sigma) - r(c, \tau) \tag{9}$$

We are interested in understanding in which city and in which sector a household with intrinsic skill ω decides to work, that is, how the household maximizes (9) with respect to c, τ and σ.

Sectoral decisions. In each city c, we can define two thresholds $\omega(m, c)$ and $\omega(h, c)$:

$$H(\omega(m, c), m, c)p(m) = H(\omega(m, c), l, c)p(l) \tag{10}$$

$$H(\omega(h, c), h, c)p(h) = H(\omega(h, c), m, c)p(m) \tag{11}$$

The threshold $\omega(m, c)$ is such that a household in a given city c is indifferent between working in the low-skill and the middle skill sectors and the threshold $\omega(h, c)$ is such that the same household is indifferent between the middle and the high skill sectors.

The following lemma shows that these two thresholds are sufficient for characterizing sectoral decisions by households:

Lemma 1. *A household living in city c and with skill ω works in sector l when $\omega \leq \omega(m, c)$, in sector m when $\omega \in (\omega(m, c), \omega(h, c))$ and in sector h when $\omega \geq \omega(h, c)$.*

Across cities, these thresholds satisfy:

$$\omega(1, h) < \omega(2, h) \text{ and } \omega(1, m) < \omega(2, m) \tag{12}$$

Proof. See Appendix B.1 □

The differences in the thresholds across cities result from the comparative advantage of the larger cities in more skilled sectors associated with the increasing reliance on individual skills in higher skilled sectors. For the same prices of intermediate goods, the same individual is relatively more productive in higher skilled sectors in the larger city and, thus, has more incentive to work in these sectors.

Let us note that, in principle, it is possible that a sector does not exist in at least one of the two cities, even though the production function guarantees that this sector should exist in at least one city. This happens, for example, when $\omega(1,m) \leq \underline{\omega}$. In this case, there is no low-skill sector in city 1.

This result contrasts with Davis and Dingel (2014) where the sectors and ωs in the less productive city were a strict subset of the sectors and ωs in the larger city. This comes from the assumption that the productivity gains in a given city are different depending on the sector ($A(c,\sigma)$ is a function of σ). In the case where these gains are constant across sectors ($A(c,\sigma) = A(c)$), the thresholds would be the same in the two cities and we would be back to the framework of Davis and Dingel (2014). In addition, note that different thresholds also imply the same ω need not produce in the same σ in both cities even when present.

In the end, Lemma 1 defines a function M such that $M(\omega,c)$ is the optimal sectoral decision for a household with skill ω in city c.

Location decisions. Let us now turn to location decisions. First note that a household with skill ω will decide to work in city 1 and in a given location only if it is not better off working in the other city or in any other location τ, that is:

$$\max_{\sigma,\tau} H(\omega,\sigma,1)T(\tau)p(\sigma) - r(1,\tau) \geq \max_{\sigma',\tau'} H(\omega,\sigma',2)T(\tau')p(\sigma') - r(2,\tau'). \tag{13}$$

When this holds with equality the skill ω is present in the two cities.

Location within cities. Let us start by describing the decision location within each city. Here we closely follow Davis and Dingel (2014).

To start with, the set of locations occupied in city c is a bounded set. We denote by $\overline{\tau}(l,c)$ the maximum value of τ occupied in city c. More desirable locations have higher rental prices:

Lemma 2. *Housing prices $r(c,\tau)$ are decreasing on $[0,\overline{\tau}(l,c)]$ and $r(c,\overline{\tau}(l,c)) = 0$. Finally, for all $\tau \in [0,\overline{\tau}(l,c)]$:*

$$S(\tau) = L\int_0^{\tau}\int_{\sigma}\int_{\omega} f(c,x,\omega,M(\omega,c))d\omega d\sigma dx \tag{14}$$

Proof. See Appendix B.2 □

Furthermore, higher skilled households occupy more desirable locations. We find this by obtaining a mapping between skill ω and location (c, τ):

Lemma 3. *There exists a function K such that: $f(c, \tau, \omega, M(\omega, c)) > 0 \Leftrightarrow K(c, \tau) = \omega$. The function $K(c, .)$ is continuous and strictly decreasing.*

In addition, there exist $\underline{\tau}(1)$ and $\underline{\tau}(2)$ such that $K(2, \underline{\tau}(2)) = K(1, \underline{\tau}(1)) = \underline{\omega}$ and $K(1, 0) = \overline{\omega}(1) = \overline{\omega}$ and $K(2, 0) = \overline{\omega}(2)$.

Proof. See Appendix B.3 □

The proof of this lemma closely follows Davis and Dingel (2014) but extended to the case where households' productivity in a given sector is city specific.

Previously, we noted that the different sectors need not be in both cities. In contrast, both cities have the least skilled person $\underline{\omega}$, so that the larger city's skill set is a strict superset of that in the smaller city.

Using the results of Lemma 1 and 3, we can connect location decisions with the sectoral decisions and show that more skill intensive sectors concentrate in the most attractive locations in the city.

Lemma 4 (Sorting within cities)**.** *In each city c, there exists $\overline{\tau}(h, c) \leq \overline{\tau}(m, c) \leq \overline{\tau}(l, c)$ such that:*

- *If $\omega \geq \omega(h, c)$ then $\tau \leq \overline{\tau}(h, c)$,*
- *If $\omega \in [\omega(m, c), \omega(h, c)]$ then $\tau \in [\overline{\tau}(h, c), \overline{\tau}(m, c)]$,*
- *If $\omega \leq \omega(m, c)$ then $\tau \in [\overline{\tau}(m, c), \overline{\tau}(l, c)]$.*

In particular, $f(\omega, \sigma, c, \tau) = 0$ for all ω, σ, c and $\tau \geq \overline{\tau}(l, c)$.

Proof. See Appendix B.4 □

Locations across cities. To start with, there are locations in city 1 where the productivity of worker is strictly higher than what it could be in city 2. This happens for locations τ where productivity in city 1 strictly exceeds what can be obtained in city 2, even in the best location. More formally:

$$H(\omega(\tau), M(\omega(\tau), 1), 1)T(\tau) > H(\omega(\tau), M(\omega(\tau), 2), 2)T(0) \tag{15}$$

where $\omega(\tau) = K(1, \tau)$ is the value of ω occupying location τ in city 1. This defines a maximum value for the skill in city 2, $\bar{\omega}(2)$ for which inequality (15) holds with equality.

Below the productivity $\bar{\omega}(2)$, for each ω and for each τ, there exists $\tau' < \tau$ such that the productivities in city 1 and in city 2 are the same:

$$H(\omega(\tau), M(\omega(\tau), 1), 1)T(\tau) = H(\omega(\tau), M(\omega(\tau), 2), 2)T(\tau'). \tag{16}$$

To give further intuition, when there are high-skill workers in both cities, (16) can be rewritten, depending on the value of ω as:

$$\begin{aligned}
\forall \omega \in [\omega(2,h), \bar{\omega}(2)], \quad & A(1,h)T(\tau) & &= A(2,h)T(\tau') \\
\forall \omega \in [\omega(1,h), \omega(2,h)], \quad & A(1,h)T(\tau)e^{\eta\omega}p(h) & &= A(2,m)T(\tau')p(m) \\
\forall \omega \in [\omega(2,m), \omega(1,h)], \quad & A(1,m)T(\tau) & &= A(2,m)T(\tau') \\
\forall \omega \in [\omega(1,m), \omega(2,m)], \quad & A(1,m)T(\tau)\omega p(m) & &= A(2,l)T(\tau')\omega^{\phi}p(l) \\
\forall \omega \leq \omega(1,m), \quad & A(1,l)T(\tau) & &= A(2,l)T(\tau')
\end{aligned}$$

In the end, for an $\omega \leq \bar{\omega}(2)$ and a τ, there exists a single τ' in city 2. This defines a function $\Gamma(\omega, \tau) = \tau'$. We then have that, for all ω:

$$H(1, \omega, M(\omega, 1))T(\tau) = H(2, \omega, M(\omega, 2))T(\Gamma(\omega, \tau))$$

In equilibrium, in the location τ, if the agent with skill ω is the marginal buyer, we then have that:

$$r(1, \tau) = r(2, \Gamma(\omega, \tau)).$$

In Davis and Dingel (2014), the function Γ would be constant with respect to ω, but, as the larger city has also a comparative advantage in higher-skill sector, we obtain the following result:

Lemma 5. *For all ω, $\Gamma(\omega, .)$ is continuously increasing in τ and, for any τ, $\Gamma(., \tau)$ is continuous and weakly decreasing in ω.*

Proof. See Appendix B.5 □

In the end, for $\omega \in [\underline{\omega}, \bar{\omega}(2)]$, households are indifferent between a less desirable location in the more productive and larger city 1 or a more desirable location in the less productive and smaller city 2.

Sectoral decisions and factor prices. As this can be observed from equations (10) and (11), the two thresholds are functions of intermediate good prices $p(l)$, $p(m)$ and $p(h)$. The following lemma clarifies how the thresholds moves as a function of $p(m)$.

Lemma 6. *A decline in $p(m)/p(h)$ implies a relatively larger decline for $\omega(1,h)$ than for $\omega(2,h)$. An increase in $p(l)/p(m)$ implies a relatively larger increase in $\omega(2,m)$ than for $\omega(1,m)$.*

Proof. See Appendix B.6 □

When the relative price of middle-paying work declines, the incentive for a middle-skill worker to become a high-skill worker increases for a larger set of people in the large than in the small city.

The reason is that the difference in productivity in the middle-skill sector (that increases linearly with skill ω) and the high-skill sector (which increases exponentially) is lower in the large city for more workers given the lower threshold $\omega(1,h)$ in the larger city.[3]

Similarly, the incentive for a middle-skill worker to become a low skill worker also increases in both cities. Yet, as the relative productivity for middle-paying jobs is lower in the smaller city, this incentive for middle-skill workers to become low skill increases by more in the smaller city: this leads $\omega(2,m)$ to increase by more than $\omega(1,m)$ for the same variation of intermediate good prices.

2.3 Implications

We are now able to describe the main implications of our model. In particular, we show four empirical implications along the lines on the evidence that we report.

Allocation of skills and city exposure to the middle-skill sector Let us first describe the allocation of skills and the exposures to different skills across cities.

Initial exposure to middle-skill jobs Our first implication is about the exposure of a given city to middle-paying jobs, that is the share of total jobs in the middle-paying sector:

Proposition 7. *When $\frac{A(1,h)}{A(2,h)}$ is sufficiently large relative to $\frac{A(1,m)}{A(2,m)}$, the share of middle skill sector jobs is smaller in the larger city.*

Proof. See Appendix B.7. □

The lower exposure of the large city to middle-skilled jobs reflects this city's comparative advantage in high-skilled jobs. Indeed, the relative importance of the three different sectors in the larger city depends on the relative productivity gains of concentrating in that city households working in each of these sectors. When the productivity gains of concentrating workers in the high-paying sectors in the large city are sufficiently large, the share of this sector becomes relatively large and crowds out the presence of the other sectors, thus leading the larger city to be more exposed to high-skill jobs and less exposed to middle-skill jobs.

Log-supermodularity Our second implications concern the distribution of skills in the two cities:

Proposition 8. *Let us assume that the supply of locations in each city has a sufficiently decreasing elasticity. Then, the distribution of skills $f(c,\omega)$ is strictly log-supermodular.*

[3]More specifically, in the large city, the indifference condition between being the middle-skill and the high-skill sectors leads to a lower threshold $\omega(1,h)$ than the one in the smaller city. As a result, a similar variation in the gains of becoming a high-skill leads to a larger variation for $\omega(1,h)$, where productivity is relatively flatter with respect to ω than for $\omega(2,h)$ where productivity is steeper with respect to ω.

Proof. See Appendix B.8 □

Let us remind that a distribution is strictly log-supermodular when, for $c > c'$ and $\omega > \omega'$, $f(c,\omega)f(c',\omega') > f(c',\omega)f(c,\omega')$, which means that there are relatively more high skill workers in the larger city.

Given our previous result in Proposition 7 where we obtained conditions under which the share of middle-paying jobs is smaller in the larger city, we can also characterize the elasticity of the middle-paying jobs with respect to the size of the city:

Corollary 9. *Under the conditions of Proposition 7, the elasticity for middle-skill workers with respect to the size of the city is lower than* 1*.*

One implication of this result associated with the fact that larger cities have a lower initial share of middle-skill workers as shown in Proposition 7 is that occupations do not need to be log-supermodular as this is the case for skills. More specifically, the total number of jobs in the middle-skill occupations may be lower in the larger city compared with the smaller one.

Labor market polarization Let us now investigate how a decrease of the price of the intermediate good z affects the distribution of jobs in our economy, as in Autor and Dorn (2013).

Labor market polarization in the aggregate Let us first observe how this decline of the price of capital affects labor markets overall and in each city:

Proposition 10. *A decline in p_z reduces the share of middle-skill jobs in the aggregate and in each city.*

Proof. See Appendix B.9. □

As in Autor and Dorn (2013), a decline in the price of capital goods/offshoring intermediary goods leads firms to substitute middle-paying jobs by capital. This leads workers to reallocate, either to the high-paying or to the low-paying sectors, depending on workers' skills and, overall, the labor market becomes more polarized.

Importantly, this polarization does not only occur in the aggregate, as already is shown by Autor and Dorn (2013), but also in each city: in both cities, the share of middle-paying jobs declines.

Labor market polarization across cities Yet, labor market polarization features some striking differences depending on the size of the city: how pronounced is the decline of middle-skill jobs and the rise in low- and high-paying jobs depends on the relative productivity gains of the different sectors in the different cities:

Proposition 11. *When $\frac{A(1,h)}{A(2,h)}$ is sufficiently large relative to $\frac{A(1,m)}{A(2,m)}$, then in the large city:*

(i) *The increase in high-skill jobs is larger in percentage points.*

(ii) *The decline in middle-skill jobs is larger in percentage points.*

(iii) *The increase in low-skill jobs is smaller in percentage points.*

Proof. See Appendix B.10. □

The decrease in the price of capital/offshored goods corresponds to a negative demand shock for middle-skilled jobs. The effects of this shock are the strongest where the opportunity cost of being a middle-skill is the largest. For each level of skill ω, this opportunity cost is the one associated with not selecting a high- or a low-skill job.

When the comparative advantage of being a high-skill the large city is sufficiently large – that is when $A(1,h)/A(2,h)$ is sufficiently large compared with $A(1,m)/A(2,m)$ but also compared with $A(1,l)/A(2,l)$ given Assumption 1 – the opportunity cost to be a middle-skill is the highest for agents in the large city that are close to selecting a high-skill job. Thus, in that case, a decrease in the price of capital/offshored goods leads to a stronger decline in middle-skill jobs in the largest city and, for the same reason, in a larger increase in the share of high-skill jobs. In contrast, the smaller city has a comparative advantage in low-skill jobs, which leads to an increase in the share of these jobs in that city. Note that this result would not appear in Davis and Dingel (2014), given that there is no specific comparative advantage for high-skill jobs to be in the larger city.

In the end, the decline in the price of capital/offshored goods leads to a reinforcement of the largest city into high-skill jobs: it is indeed in that city with the lowest exposure to middle-skill jobs (and the more specialized in high-skill jobs as noted by Proposition 7) that further specializes in high-skill jobs and reduces the number of middle-skill jobs.

To obtain these results, the mobility of workers across cities and sectors with play a key role: agents need to find optimal to live in a city and work in a given sector. In the absence of mobility of workers across sectors, the specialization of the large city in high skill activities would mechanically lead the smaller city to host more middle-skill jobs – as a result, no labor market polarization would occur in that smaller city, which we will find to be at odds with what we observe in the data. The mobility of workers also ensures that the stronger destruction of middle-skill jobs in the large city does not lead to a stronger creation of low-skill jobs in that city. This contrasts with Autor and Dorn (2013), where for low levels of skill, there is mobility only between the low- and the middle-skill sectors and the corresponding workers cannot move from one city to another: there, the stronger destruction of middle-skill jobs leads to more creation of low skill jobs in the large city.

3 Data description

We focus on a few key questions: the characteristics and the evolution of labor market polarization in the aggregate and by metropolitan area and the log supermodularity of skills. These require

exhaustive data on job characteristics (e.g. their routine or offshorable nature), hours worked and wages by occupations. We also need a measure of skills such as educational attainment. The data should be geographically detailed at the metropolitan area level and comparable through time.

3.1 DADS-Postes data

Our main data source is DADS-Postes for the years 1994-2015, which is a part of the publicly available DADS ("Déclaration Annuelle des Données Sociales") data set[4]. This data is provided by INSEE, the French national statistical institute, and is based on mandatory annual reports by all French companies[5]. It includes data about all legally held job positions (postes), detailed at the plant level. The initial year 1994 is chosen as this is the first year of data that has comprehensive coverage of hours worked while 2015 was used as the last vintage available. For each worker, for a particular job position, the main reported data are the hours worked, remuneration (total compensation before taxes), occupation type, age and gender[6]. Establishment location information is available at the *commune* level, the lowest administrative unit. There were 36,169 communes in metropolitan (mainland) France as of January 1, 2015.

We use data only for private companies, excluding privatized firms or those that changed status from public to private incorporation (which impacts for example the public or private law under which labor contracts are offered) in the period 1994-2015[7]. We use data for mainland France (without Corsica or the overseas departments). We limit the sample to workers 25-64 years of age. To minimize erroneous entries (for example employees recorded with few working hours with abnormally high income) one needs to filter the job positions. We retain all positions where there were at least 120 hours worked in a year[8].

3.2 Occupations and their classification

In DADS-Postes the information on occupations is available at a 2-digit level according to the French occupation classification called PCS ("Nomenclature des professions et categories socio-profesionnelles"). It has been developed by French statistical authorities to classify occupations

[4]The French Labor Survey is available since 1982 but for early years has approximately 60,000 observations per year and has only data at the department level. This allows to document some general facts about labor market polarization starting from 1982, but the DADS-data is exhaustive and gives inter alia more geographical details.

[5]This includes public and private firms. Data on self-employed are not reported.

[6]We cannot observe education data or job tenure, and it is not possible to aggregate incomes by individual workers at each year level.

[7]Some firms in the finance, insurance and real estate sectors reported pre-2001 their employees from branches at few establishments for example at the department level which may introduce minor errors given the scale of the problem when we use metropolitan area-level data. Exclusion of these sectors from our analysis does not change our results considerably and does not impact our conclusions. We include Table 16 without these sectors as a replication of Table 7 in a robustness test.

[8]We do not observe, however, a material difference in our results if no filtering is applied or filtering based on end-of-year presence with at least 30 days in the firm. INSEE provides filtering in the DADS data set, but it is not consistent between 1994 and 2015.

according to their "socio-professional" status and does not have a clean and direct correspondence at this level to other internationally used classifications such as e.g. the International Standard Classification of Occupations (ISCO). The broad 1-digit codes represent CEOs or small-business owners (CS category "2"), "cadres" (high-skilled professionals, code "3"), medium-skilled professions (codes starting with "4"), low-skilled employees (codes with "5" as the first digit) and blue-collar workers (codes starting with "6"). The 2-digit categories provide more detail, allowing us to use 18 different CS 2-digit categories (see Table 19 for representative occupations within each category) [9]. We exclude artisans, agriculture-related and public-sector occupations.

The list of 2-digit CS categories we use is provided in Table 1 along with a short description, their in-sample employment share, average wages in considered cities in 1994 and 2015, as well as their routine (RTI) occupation and offshorability (OFF-GMS) ranking. The exact index values for each category are given in the Appendix Table 20.

Table 1 – Basic statistics by 2 digit CS categories.

CS 2-digit	description	employment share in %		average city wage (in 2015 euros)		Routine ranking	Offshorable ranking
		1994	2015	1994	2015		
	high-paying occupations						
23	CEOs	1.0	0.9	42.82	59.20	16	17
37	**managers and professionals**	**6.2**	**10.2**	32.52	38.56	15	16
38	**engineers**	**5.1**	**9.0**	30.36	33.68	17	10
35	creative professionals	0.5	0.5	22.82	31.80	14	11
	medium-paying occupations						
48	**supervisors and foremen**	**4.1**	**2.7**	18.03	21.86	3	3
46	**mid-level professionals**	**12.3**	**7.6**	17.54	21.20	13	6
47	**technicians**	**5.7**	**6.3**	17.15	20.60	11	7
43	mid-level health professionals	0.8	1.5	15.05	18.05	10	13
62	**skilled industrial workers**	**14.1**	**9.3**	13.52	17.99	4	2
54	**office workers**	**11.8**	**11.2**	13.17	16.98	1	4
65	**transport and logistics personnel**	**2.9**	**3.0**	11.96	16.00	5	5
63	**skilled manual workers**	**8.0**	**8.3**	11.90	15.50	7	8
64	**drivers**	**5.0**	**5.5**	11.50	14.46	18	18
67	**unskilled industrial workers**	**10.9**	**5.7**	11.02	14.72	2	1
	low-paying occupations						
53	security workers	0.7	1.4	10.60	14.60	9	12
55	**sales-related occupations**	**5.4**	**8.3**	10.44	13.74	6	15
56	personal service workers	2.2	4.8	9.97	12.63	12	14
68	**unskilled manual workers**	**3.3**	**3.8**	9.11	13.27	8	9

Notes: In-sample values. Employment share for metropolitan mainland France (excluding Corsica). "Routine" ranking based on the RTI measure of Autor et al. (2003) while "Offshorable" on the OFF-GMS measure from Goos et al. (2014), both mapped into PCS 2-digit employment categories. Occupations with employment shares above 2.5% in 1994 in bold.

[9] Firms should report their data using much finer 4-digit codes, but many fail to do so, and as a result this data is not usable. After the 2003 revision the difference at the 2-digit level is a new category, 31 ("liberal" professionals such as lawyers etc.) that was previously included in 37. In all our data we merge the two together.

To classify high, medium and low-paying occupations and obtain the exposures to automation and offshoring we proceed as follows. We first merged into the 1994 French Labor Survey the exposure classifications of Goos et al. (2014) and mapped their 2-digit ISCO-based ones into the 2-digit CS used in French data[10]. We classify as high paying occupations those of "cadres" and CEOs (CS codes 23, 35, 37, 38). Apart from a different legal (e.g. special retirement treatment) and social status, there is a clear gap in terms of wages between these and the remainder of the occupations ("non-cadres"). To determine the low-paying occupations in the 2-digit CS classification, we retained those for which the share of low-paying occupations classified as such in Goos et al. (2014) in terms total hours worked in 1994 was over 50%. The resulting 4 occupations are indeed ranked as the least paid in 1994 in our data, and they are typically much less routine and offshorable than the least-paid middle-paying job category "unskilled industrial workers".

The low-paying occupations are then security workers, retail workers, personal service workers and unskilled manual laborers (CS 53, 55, 56, 68)[11]. As a measure of exposure to automation we retain the Routine Task Intensity index (RTI) used by Autor and Dorn (2013), where this is used to identify occupations for which computers may be able to substitute. For the measure of offshorability, we use the standardized index developed by Goos et al. (2014), where this indicates occupations that can be readily substituted by imports. We observe the CS category 54 (clerks) as being most routine and 67 (unskilled industrial workers) as most offshorable.

We note that the set of the four most routine and the four most offshorable occupations are the same (CS 48, 54, 62 and 67), comprising 40.9% of hours worked in 1994 in our private-sector employment sample and spanning the entire wage distribution of middle-paying jobs. We will refer to this group as RTI4 jobs. The focus on these 4 most routine and offshorable occupations corresponds well with our theory. In our setup medium skill intermediate goods (equation 7) can be produced either by capital goods or imported intermediates. The fall in the prices of both types of goods will have the same effects in our model, and will be captured similarly in our data.

3.3 Cities considered and final sample

For most of our empirical exercises, we are going to be concerned with across-city comparisons. We will therefore use principally data on jobs that are performed in cities (metropolitan areas) above 50,000 inhabitants as of 2015 unless otherwise noted. We aggregate the commune-level data to the metropolitan area level ("unité urbaine") with city boundaries defined by INSEE as of 2010 unless otherwise indicated. There are 117 such cities in 2015 with the largest 55 above 100,000 inhabitants shown in Figure 1 with population data by category in Table 8. The characteristics of

[10]The ISCO and CS categories are both available only in the French Labor Survey and not directly in the DADS data. We used hours worked in 1994 in the Survey as weights.

[11]Our results are not materially affected when we use a narrower set of occupations – only CS 55, 56, 68. The reason for such a robustness check is that the CS 53 category is quite close in terms of wages in 1994 to the category CS 67 that we classify as a middle-paying job.

the final sample are given in Table 9.

The cities above 50,000 inhabitants have 54% of total population of metropolitan France. At the same time, the jobs therein are responsible in 2015 for 73% (73% in 1994) of wages paid and 68% (69% in 1994) of hours worked in the mainland in the non-farm private sector. There were 396,540 (out of 596,430 for which we have data) and 633,845 (out of 998,455) firms active in these metropolitan areas in 1994 and 2015 respectively. After all above-mentioned exclusions we remain with a sample that accounts for 65% of total wages paid and 58% of hours worked in metropolitan France in both 1994 and 2015 with data from 364,398 and 596,441 firms in 1994 and 2015 respectively[12].

We group the cities into six major categories for our analysis. Paris, given its size (10.7m inhabitants in the metropolitan area and 37.5% of jobs in our final sample) is a category by itself. Then, we use 2 categories of cities above 0.5m cities: 0.5-0.75m and 0.75m and above (except Paris). Such a choice is warranted because there is a considerable size difference between the seventh largest metropolitan areas – Bordeaux (904 thousand inhabitants) and the eighth – Nantes (634 thousand people). Moreover, cities with metropolitan areas of "0.75m and above" have also "urban areas" as defined by INSEE of over 1m inhabitants[13]. For other divisions we follow the ones of INSEE: 0.2-0.5m (size categories "71" and "72") , 0.1-0.2m (sizes "61" and "62") and 0.05-0.1m ("51" and "52"). We took the city size of 50,000 as a cutoff - below which individual city labor markets are small.

3.4 Other data

We also use other data sources to provide additional statistics. When we examine log-supermodularity, we use the detailed part of the Census of 1999 (5% of population), also provided by INSEE. It provides data on education, nationality of respondents, and allows us to identify their location at the commune level. Population data was taken from the INSEE for the relevant years.

4 Core empirics

Our primary aim is to understand the consequences of technological and offshoring shocks on labor market polarization across cities. However these are best understood if we take as a preliminary understanding key cross-sectional patterns in the data.

[12]Establishment-level output data is unavailable in the French data.

[13]This is the "unité urbaine" – all communes in the metropolitan area plus all communes where at least 40% of residents have employment in the same metropolitan area.

4.1 Log-supermodularity of skills.

One implication – Proposition 8 – of the model is, as in Davis and Dingel (2014), that the distribution of skills $f(c, \omega)$ is log-supermodular in city size. To obtain a measure of skills we turn to the 1999 Census data that has both good data on diplomas and the commune of residence. We measure skills by the highest obtained diploma by individuals. The results are shown in Table 2 with a more detailed classification given in the Appendix Table 10.[14] We observe that, as expected, for low skilled categories we obtain coefficient estimates that are statistically significantly below 1 at conventional significance levels. The opposite is true for those with higher education that are more prevalent in larger cities. In particular, the category of workers with a graduate diploma has a population elasticity of 1.184. We note that the elasticities for middle-skilled workers are not significantly different from 1. These observations carry over when we consider only French-born individuals: the presence of low-skilled immigrants does not change these patterns.

Table 2 – Log-supermodularity, population elasticities by diploma (4 categories) in the 1999 Census data.

Dependent variable: ln $f(c, \omega)$	All workers	French born	Population share	French born share
Below high school x ln pop	0.93**	0.91***	0.24	0.84
	(0.027)	(0.023)		
High school professional diploma (CAP, BEP) X ln pop	0.91***	0.91***	0.31	0.96
	(0.023)	(0.022)		
End of high school diploma (Bac) X ln pop	1.00	0.99	0.15	0.95
	(0.024)	(0.023)		
Higher education X ln pop	1.10***	1.09***	0.3	0.96
	(0.030)	(0.030)		

Notes: 112 cities > 50,000 inhabitants defined by INSEE as of 1999. The variable "ln pop" is the natural logarithm of metropolitan area population from 1999. Exclusions in terms of CS and age as in main sample. Robust standard errors in parantheses. ***, **, and * denote statistical significance at the 1 %, 5 %, and 10 % levels for the test of hypotheses whether a given coefficient is equal to one.

We reconfirm these results using our classification of high, middle and low paying jobs and the broad 1-digit CS categories in Tables 11-12. It is not a coincidence that the population elasticity coefficients for high-paying jobs and "cadres" (respectively 1.14 and 1.16) are similar: "cadres" perform the bulk of high-paying occupations. The coefficients on middle- and low- paying jobs that are below 1 (in a statistically significant manner) show that larger French cities have not only fewer low paying jobs, but also fewer middle-paying jobs. This conforms with Corollary 9.

4.2 Employment shares of different occupation categories.

The model has implications on the cross-section employment shares across cities. We use the DADS-Postes data set for 1994 and 2015 and calculate hours worked in different occupations across

[14] In this table, CAP or *Certificat d'aptitude professionnelle* is obtained at the age of 16, the BEP or *Brevet d'études professionnelles* is also obtained at the age of 16 but is a prerequisite for obtaining the more advanced bac professionnel at the age of 18.

Table 3 – Share of 4 highest-paying occupations per metropolitan area size.

Agglo.size	Paris	> .75M	.5-.75M	.2-.5M	.1-.2M	.05-.1M
1994	0.23	0.14	0.12	0.10	0.09	0.08
2015	0.37	0.25	0.21	0.16	0.14	0.12
change in ppct	0.13	0.11	0.09	0.06	0.05	0.04
growth in %	0.57	0.77	0.71	0.63	0.61	0.49

Note: Percentage point changes and growth rates not directly calculable from the upper two lines of the table due to rounding.

metropolitan areas (Tables 3-5).

Table 4 – Share of 10 middle-paying occupations per metropolitan area size.

Agglo.size	Paris	> .75M	.5-.75M	.2-.5M	.1-.2M	.05-.1M
1994	0.65	0.74	0.75	0.77	0.79	0.79
2015	0.45	0.57	0.60	0.64	0.66	0.67
change in ppct	-0.20	-0.17	-0.15	-0.13	-0.14	-0.12
growth in %	-0.31	-0.23	-0.20	-0.17	-0.17	-0.15

Note: Percentage point changes and growth rates not directly calculable from the upper two lines of the table due to rounding.

We observe that the share of high-paying occupations in total employment increases monotonically with city size in both years, as required by log-supermodularity shown by Proposition 8 and documented in the previous section. The differences are sizable, especially when comparing the extremes – the Parisian metropolitan area and cities between 50 and 100 thousand of population. In 1994, the share of high-paying occupations in Paris was 23.3 %, as opposed to 10.1 % for cities between 0.2 and 0.5m, and only 8.0 % in cities between 0.05 and 0.1m of population. This diverged to respectively 36.7 %, 16.4 %, and 11.9 % in 2015.

Moreover, the share of middle-paying jobs monotonically declines with city size in line with Proposition 7, both in 1994 and 2015 (see Table 4). The share of lowest-paid occupations is highest in smallest cities in either of the years, although the across-city variation is modest. The decline of low-paid occupation shares with city size is, however, very clear when one measures the share of hours worked for 3-lowest paying jobs (sales-related occupations, personal service workers and unskilled manual workers; see Table 13).

Overall, these patterns along with evidence on log-supermodularity support our theoretical model. They also contradict for France the extreme-skill complementarity hypothesis stated by Eeckhout et al. (2014).

Table 5 – Share of 4 lowest-paying occupations per metropolitan area size.

Agglo.size	Paris	> .75M	.5-.75M	.2-.5M	.1-.2M	.05-.1M
1994	0.12	0.13	0.12	0.13	0.12	0.13
2015	0.18	0.18	0.18	0.20	0.20	0.21
change in ppct	0.07	0.06	0.06	0.07	0.08	0.08
growth in %	0.59	0.48	0.48	0.55	0.68	0.64

Note: Percentage point changes and growth rates not directly calculable from the upper two lines of the table due to rounding.

As automation (Autor and Dorn (2013)) and offshoring (Goos et al. (2014)) are believed to be driving labor market polarization, it is important to observe the exposure of different cities to routine and offshorable tasks. In Table 6 we show that the employment shares across cities in 1994 and 2015 of most routine and offshorable jobs (the same 4 categories that belong to middle-paying group) are declining in city size – in line with the patterns for middle-paying occupations overall exhibited in Table 4 and our Proposition 7.

Table 6 – Share of the 4 most routine and offshorable occupations (CS 48, 54, 62 and 67) per metropolitan area size.

Agglo.size	Paris	> .75M	.5-.75M	.2-.5M	.1-.2M	.05-.1M
1994	0.29	0.36	0.39	0.41	0.45	0.45
2015	0.19	0.25	0.27	0.29	0.31	0.32
change in ppct	-0.11	-0.12	-0.12	-0.12	-0.13	-0.12
growth in %	-0.36	-0.32	-0.31	-0.29	-0.29	-0.27

Note: Percentage point changes and growth rates not directly calculable from the upper two lines of the table due to rounding.

4.3 Labor market polarization in France.

Turning to changes in employment patterns in time, we first state that labor market polarization – a fall in the employment share of medium-paying occupations and the rise in the share of high- and low-paying ones occurred in France in the period 1994-2015. The high-, medium, and low-paying occupations, respectively account for one-eighth, three-quarters, and one-eighth of employment in 1994 and roughly one-fifth, three-fifths, and one-fifth in 2015.

We can examine these changes in overall employment shares exhibited in Table 1. The share of middle-paying jobs declined from 75.5% to 61.1% between 1994-2015. The bulk of job losses in this category occurred in the RTI4 jobs – the 4 most automation and offshoring exposed occupations

(supervisors and foremen; office workers; skilled and unskilled industrial workers), and their share in hours worked fell from 40.9% to 28.9%. The only other occupation experiencing a large employment share drop was CS 46 (mid-level associate professionals) the share of which fell from 12.3% to 7.6% in the labor force. It is ranked 6th highest in our classification of offshorable categories.

At the same time, the overall shares of high- and low-paying jobs increased respectively from 12.8% to 20.6% and from 11.6% to 18.3%.

The patterns detailed at the 2-digit CS-level are exhibited in Figure 3 and confirm the U-shaped relationship in France for the years 1994-2015 studied by (Autor et al., 2006; Autor and Dorn, 2013) for the U.S. and documented by (Goos and Manning, 2007; Goos et al., 2009, 2014) for Europe[15]. They are also consistent with observations made by Harrigan et al. (2016) for France for the time period 1994-2007.

Our theory – Proposition 10 – that labor market polarization is induced by increased automation and/or a lower offshoring cost – is compatible with the above patterns.

4.4 Labor market polarization in large and small cities.

One can observe labor market polarization in France in the aggregate data. Our model (Proposition 11) implies that we should observe polarization at the city level as well. However, labor market polarization should have different consequences for large and small cities. First, middle-paying jobs should decline the most in large cities. Second, high-paying jobs should increase by more in larger metropolitan areas, while low-paying jobs increase more in smaller cities. The strength of these relationships need not depend on initial exposure to routine or offshorable occupations, and, indeed, such exposure should be lowest in large cities as indicated by Proposition 7 and exhibited in Section 4.2. All such patterns are upheld in our data.

Labor market polarization in cities at the 2-digit CS level. We start by discussing labor market polarization in cities of different sizes in France (Paris and Lyon – the two largest cities; cities between 0.2-0.5m and 0.05-0.1m inhabitants) comparing employment change patterns shown in Figure 4.

First of all, the figure confirms that labor market polarization occurred in different cities[16]. However, it is immediately clear that the patterns are markedly different across metropolitan areas. Figures 5 and 6, where we superimpose the employment share changes in Paris and Lyon vs. those in cities with population between 0.05-0.1m, further attest to this. The changes in the Parisian labor market in comparison to smallest cities between 0.05-0.1m inhabitants are larger (e.g. the average of absolute percentage point changes in the individual categories – 2.4 vs. 1.9 respectively

[15]We exclude here the category of CEOs - CS category 23. It is an outlier with highest pay that has a rather constant population elasticity in sample – its employment share varies between 0.8% and 1.1% in 1994. The change in the share is less than 0.3 percentage points in absolute terms across cities in the given years.

[16]Similar patterns can be observed for all size categories considered in Tables 3-6.

– are higher), despite a lower exposure to jobs with the highest routine or offshorability indexes (Table 6).

Turning to individual categories, there is a much higher percentage point (pp) increase in the employment share of categories CS 37 (managers and professionals) and 38 (engineers) in Paris and Lyon in comparison to other cities. Moreover, different categories of middle-paying occupations are declining in large and small cities. While in smallest cities the highest employment declines are observed for industrial workers (whether skilled or unskilled - categories 62 and 67) that happen to be the two most offshorable categories in our data as well, it is category CS 46 (mid-level associate professionals, ranked as sixth most offshorable occupation) whose share shrinks most in Paris. The category that is most routine in sample – office workers (CS 54) – declines also most strongly in Paris (by 3.1 pp) while in small cities employment in this category grows (by 0.5 pp). The CS category 47 (technicians) proportion of hours fell in Paris by 1.4 pp but increased by 1.3 pp in the smallest cities. Among the low-paying occupations, there is a relatively strong increase in the CS category 55 (sales related occupations) employment share by 4.5 pp in small cities. The patterns in Lyon and cities between 0.2-0.5m inhabitants show intermediate developments between the two polar cases, suggesting common forces at play that may be related to city size.

Some of these differences in labor market developments for individual categories across cities may be due to several factors that should be mentioned but cannot be fully addressed empirically given the limitations of data at our disposal.

First, the employment share of industrial workers (categories 62 and 67) in large cities is lower in 1994 at the outset of our sample than in small cities. Therefore, the additional adjustment in terms of percentage points we observe in these categories over the period 1994-2015 may be less pronounced as well. Data from the Census or the French Labor market survey (available since 1982) allow us to document that labor market polarization was ongoing for several years beforehand and it was strongest (in terms of middle-paying jobs' employment share losses) in respectively larger cities or departments with large metropolitan areas. The evolution of data on individual occupations points that initially the process was especially related to deindustrialization, as the categories that decreased first and strongest among middle-paying jobs were the ones related with manufacturing (categories 48, 62 and 67). One of the reasons manufacturing was moving out of large cities in the earlier period could be the development of road and train networks in France. Further details are provided in Appendix D.

One explanation for the decline in the share of back-office or support jobs like office workers (CS 54) or technicians (CS 47) in Paris with their coincident expansion in small cities within the 1994-2015 period can be internal offshoring from large to small cities permitted by the Internet and communication technologies. Such tendencies are consistent with our model (all goods, including intermediates, are traded) though we do not model nor cannot fully verify empirically supply chain developments that are internal or external to firms.

Labor market polarization in cities by 3 occupation groups. We discuss now the changes in shares in hours worked in different categories of occupations by metropolitan area sizes.

In Table 4 we see that the percentage point changes in employment shares are highest in the largest cities. In Paris, over the period 1994-2015 middle-paying jobs share declined by 20.1 percentage points. In contrast, this decline was lower in smaller cities - only 12.1 percentage points in metropolitan areas between 50 and 100 thousand inhabitants, and this is despite the lower initial share of middle-paying jobs in larger cities. We note that the middle-paying job destruction rates increase with city size. Comparing these patterns with the changes experienced in the 4 most routine and offshorable jobs (Table 6) – we see that the fall of shares in this category of middle-paying jobs is similar across metropolitan areas – between 10.5 and 13.1 percentage points. However, as discussed in Appendix D, the decline in the most routine/offshorable jobs might have occured more strongly in large cities before 1994. Given that the exposure of the largest metropolitan areas is lowest in 1994, the destruction rates of these occupations still monotonically rises with city population[17]. The overall decline in middle-paying occupations in the largest metropolitan areas exceeds that of the destruction of the 4 most routine and offshorable jobs while it is smaller in cities between 50-100 thousand inhabitants (e.g. 12.1 vs. 12.2 percentage points respectively).

The greater decrease in middle-paying jobs in larger cities can be accounted for by our model, as exhibited by Proposition 11. There is a direct effect of the substitution of labor by capital or imported intermediates in the middle-skill task that may push – given the differences in the production functions across sectors – many more workers from the middle-skill task into the high-skill task in the large city in comparison to the smaller one[18].

Table 3 shows that the percentage point increase in high-paying jobs is monotonic in metropolitan area size. In Paris and metropolitan areas above 1m inhabitants the increase in such occupations is above 10 percentage points over the period 1994-2015. The smallest cities – those between 50 and 100 thousand inhabitants – have the lowest gain of 3.9 pp. Finally, as depicted in Table 5 the percentage point increase in low-paying jobs is highest for smallest cities. Indeed, our theory predicts that the change in the indifference thresholds $\omega(m, c)$ between working in the l and m-skill sectors will be higher in small cities by Lemma 6. We note that there is a much a sharper tradeoff of high tasks for medium skills (especially in the large cities than in smaller cities) while the changes in the low-paying jobs, although displaying behavior according with the model, are less important in terms of size.

[17]In Table 14 we show the patterns for the 6 most offshorable jobs encompassing 4 most routine occuaptions, transport and logistics personnel and mid-level professionals. The shares of such jobs in total employment are monotonically decreasing with city size whether in 1994 or 2015. Percentage point fall in the employment shares of these occupations between 1994 and 2015, however, is higher in larger cities.

[18]This tendency could be reinforced by several mechanisms that we do not model. For example, it could be due to productivity gains for high-skill workers either because of an increase in metropolitan area productivity gains $A(h, c)$ or a complementarity between high skill labor and capital (as in the Autor and Dorn (2013) model).

Individual city-level evidence. The changes in employment shares coming from overall summation of hours by metropolitan areas discussed above may not hold at the individual city level if they are driven by a few sizeable outliers. Therefore we present in Table 7 the comparison of mean percentage point changes and growth rates between the 11 largest (above 0.5m of inhabitants) and 62 smallest (between 50-100 thousand people) cities in our sample[19]. Calculations for cities between 100 thousand and 0.5 million inhabitants show intermediate patterns between the two groups.

There is not a statistically significant difference in the percentage point change in the decline of the 4 most routine and offshorable jobs (but see the discussion of patterns in Figure 7 below), though we observe a larger destruction of such jobs in larger cities with the six most routine categories included (RTI6) or the six most offshorable ones (OFF6). Average destruction of any such jobs is always stronger in larger cities. The added occupations when one moves from the narrower RTI4 category to the RTI6 and OFF6 deserve thus closer attention. Category CS 65 (transport and logistics workers) is the fifth most routine and offshorable occupation in our sample, but its employment share in the overall sample increases. However, it's employment share falls in cities >0.5m by 0.2 pp while it increases in the smallest cities by 0.3 pp. The CS 55 (sales-related occupations) ranks as the sixth most automatizable job - it contains inter alia such jobs as cashiers or telemarketers. It is a low-paying occupation, and even though some positions might have been supplanted by machines, its share in employment increases in all cities – and in accordance with our theory more strongly in smaller cities (4.5 pp) than in large ones (2 pp). The CS 46 (mid-level professionals) occupation is the sixth most-offshorable – though heterogenous – and it contains several back-office jobs that may now be performed from afar (accountants, sales administration, banking and insurance staff, interpreters, graphic designers etc.). It was the one that declined considerably more in larger cities than in smaller ones (7.1 pp vs. 3.7 pp respectively)[20]. The differences in the behavior of these CS 65, 55 and 46 categories between large and small cities may be further driven by ongoing internal offshoring of tasks within France.

As in the aggregate data by metropolitan areas, we observe a higher destruction overall of middle-paying jobs (by 18.1 percentage points on average) in cities of over 0.5m in comparison with smallest cities (11.5 percentage points). Low-paying occupations increase more in smallest cities while the highest paying in large cities as well. As lower panels of Table 7 indicate, there is a statistically significant difference in the growth rates of the discussed categories as well except in the low-paying occupations. The patterns exhibited by the aggregate data by metropolitan area sizes are confirmed: large and small cities task composition diverges considerably in the studied period.

[19]In Table 15 we show the means of employment shares in the presented categories across these cities while in Table 16 the same changes and growth patterns as in in Table 7 in a sample without FIRE industries.

[20]Another hypothesis for the decline in this category might have to do with the internal organization of the firm. Some of the professions in this category – such as interpreters, photographers, graphic designers, journalists embedded in companies – might have been internally offshored to freelancers. If that were the case, our DADS data would not capture those as we do not have data on self-employed.

Table 7 – Comparison of means of changes in different occupations, cities >0.5m vs. 0.05-0.1m.

Item	RTI4	RTI6	OFF6	middle-paid	low-paid	high-paid
mean change ppct, cities >0.5m	-0.107	-0.091	-0.178	-0.181	0.065	0.116
mean change ppct, cities 0.05-0.1m	-0.109	-0.063	-0.145	-0.115	0.079	0.036
difference in pct points	0.002	-0.027***	-0.033***	-0.065***	-0.015***	0.080***
mean growth, cities >0.5m	-0.33	-0.228	-0.351	-0.265	0.539	0.633
mean growth, cities 0.05-0.1m	-0.249	-0.114	-0.255	-0.149	0.61	0.453
difference in growth	-0.081***	-0.114***	-0.096***	-0.116***	-0.071	0.18***

Notes: 2015 population weighted, robust standard errors. N=73 (11 cities > 0.5m). ***, **, and * denote statistical significance at the 1 %, 5 %, and 10 % levels for the tests of equality of means between the groups of small and large cities. Differences remain statistically significant at least at the 5 % level without weighting or weighted by city population as of 1990. Individual mean changes or growth rates are statistically singificant different from zero at the 1 % level.

Initial exposure to routine and offshorable jobs. We disregarded so far the role of initial exposure to routine and offshorable jobs. Further scrutiny of employment changes at the city level shown in Figures 7-11 with regressions reported in Tables 17 and 18 reveals instructive patterns. We compare the largest 11 cities (above 0.5m people) and 62 smallest cities between 0.05-0.1m inhabitants to highlight the observable differences.

First of all, it is clear that large cities above 0.5m people have lower initial exposures to routine and offshorable occupations (RTI4)[21]. Although there is considerable variation, Figure 7 confirms the observation of Autor and Dorn (2013) that the initial exposure to the most routine (and, in our context, offshorable) jobs is strongly negatively correlated with their decline as the technological shocks occur. The observations for large cities lie in the lower envelope of observations: conditional on initial exposure, the changes in the employment shares in these cities are larger than in cities with a population between 0.05-0.1m inhabitants. Regression analysis in Tables 17-18 further confirms this point: the interaction of a dummy for large cities with exposure is robustly negatively different from zero. Routine and offhsorable jobs in large cities are destroyed at a higher rate than in small cities in reaction to the same automation or trade shocks.

The picture turns different, however, when we scrutinize the relationships between exposure and the change in middle-paying jobs (that include the RTI4 category) in Figure 8. The population-weighted regression of changes in employment on initial exposure to 4 most routine or offshorable occupations reveals a strongly positive relationship (though the non-population weighted relationship is close to zero). There is clearly a larger destruction of middle-paying jobs in the largest cities witnessed by the sign of the interaction of a dummy for large cities with RTI4 exposure. For many small but highly-exposed cities, the drop in RTI4 jobs is larger than the decline in middle-paying

[21]The large cities with the highest initial exposure are the Douai-Lens and Lille metropolitan areas, both located in the old industrial region in the North of France.

jobs while the opposite is true for largest cities. There may be hence some substitution of the 4 most routine or offshorable jobs by other middle-paying jobs in smaller cities while the destruction of middle-paying jobs is exacerbated in large cities.

Figure 9 shows that the overall relationship between routine and offshorable job exposure is not correlated with the changes the low paying jobs. Not only – as shown in Table 7 – increases in the low-paid job shares in large cities are lower on average than in small cities, but for the largest cities (assuming the same intercept) there is a negative correlation of exposure with the creation of these jobs. Considering each city category separately, however, these relationships is not statistically significant.

A striking pattern can be seen in Figure 10 where the two groups of cities form disparate sets of observations. The increase in high-paying occupations is distinctively higher for most large cities above 0.5m inhabitants (the outliers are Douai-Lens and Toulon). Given that the largest cities are also least exposed, there is a negative population-weighted overall correlation between initial exposure to 4 most routine or offshorable jobs and the change in the employment of high-paying occupations. Larger cities create more high-skilled jobs than small cities for any level of exposure. Further regressions (not shown) reveal that the relationship (population-weighted) between initial most routine/offshorable occupation exposure and high-paid jobs creation within city categories is negative (though not statistically significant) for large cities while positive for small cities.

This overall lack of relationship between changes in low-paid occupations, and exposure to RTI4 jobs, a negative correlation between high-paying occupations and initial RTI4 share plus a larger destruction of middle-paying jobs in larger cities (that are less exposed) as a result of labor market polarization is expected from our theory – Proposition 11. It is clear that what matters is the city size and the interaction of the technology changes on middle-skill tasks.

Similar patterns to that for 4 most routine or offshorable jobs can be observed for the 6 most offshorable occupations. For example, in Figure 11 higher initial exposure to offshorable jobs leads to their greater decrease in the studied period. This time, however, large cities are relatively more exposed to offshorable jobs in comparison to RTI4 occupations only as – in particular – they have on average a higher share of the CS 46 category, mid-level professionals (cf. also Table 14). Again, conditional on exposure, offshorable jobs employment shrinks by more in large cities in the studied period.

5 Conclusions

In a country subject to the forces of labor market polarization, how will this play out differently in large and small cities? As a framework for understanding this, we build on elements of Autor and Dorn (2013), Davis and Dingel (2014) and Davis and Dingel (forthcoming) with amendments that strengthen the relative advantage of large cities for the more skilled. Our theoretical results

predict that the large city will have smaller initial exposure to middle skill jobs, but that in the process of polarization these will decline even more sharply in the large than the small city. Under these same conditions, our theory predicts a sharper rise of high skill jobs in the large city and of low skill jobs in the small city. We take this to the data for France from 1994-2015. The results strongly confirm the key implications of the theory. In large cities there is much stronger growth of high skill jobs and decline in middle skill jobs than in small cities. Small cities do have stronger growth in low skill jobs, although the differences are modest.

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A Extensions of the model

A.1 Endogenous productivity

In the model, we treat productivity terms A as exogenous. In this subsection, we extend our results to the case where A is endogenous to the composition of labor in the city.

More specifically, let us consider the following form for productivity terms:

$$A(c,j) = G\left(\int_{\omega \geq \omega(c,h)} g(\omega) f(\omega, c) d\omega, j\right)$$

where $G(.,j)$ is an increasing function for all $j \in \{l,m,h\}$ and $G(x,j)$ is log-supermodular in $\{x,j\}$. Both assumptions ensure that, in equilibrium, Assumption 1 is satisfied.

Under these conditions, we obtain that:

Proposition 12. *A decline in p_z leads to*

(i) *an increase in the absolute advantage of city 1, i.e. $A(1,j)/A(1,j)$ increases for $j \in \{l,m,h\}$,*

(ii) *an increase in the comparative advantage of city 1 in higher-skill activities, that is $A(1,h)/A(2,h) - A(1,m)/A(2,m)$ and $A(1,m)/A(2,m) - A(1,l)/A(2,l)$ are increasing.*

By increasing the share of the population in the high-skill sector, the polarization shock increases productivity. Yet, as shown in Proposition 11, the increase in the share of the population in the high-skill sector is more important in the larger than in the smaller city. Thus, this increases productivity by more in the larger city, reinforcing the absolute advantage of this city. Given that we assume that productivity reacts by more for higher-skill occcupations, this also leads to a reinforcement of the comparative advantage in the larger city for higher skill occupations.

Remark. A potential pitfall with endogenous productivity is that it can lead to multiple equilibria. For example, in our model, symmetric cities can also be an equilibrium outcome if productivity is endogenous. To extend our results to endogenous productivity would then require to maintain the assumption that we stay close to the selected equilibrium. We also refer the interested reader to Davis and Dingel (2014) for a discussion of the possibility of multiple equilibria in a related setting.

A.2 N cities

The benchmark model only considers two cities. We extend here the model to N cities.

Let us then index cities by $c \in \{1,2,...,N\}$. We order cities so that for any $i,j \in \{1,2,...,N\}$ so that if $i > j$, we assume that city i has an absolute advantage over city j in all occupations $A(i,\sigma) > A(j,\sigma)$ for $\sigma \in \{l,m,h\}$ and it has a comparative advantage in higher skill occupations: $A(i,h)/A(j,h) > A(i,m)/A(j,m) > A(i,l)/A(j,l)$.

As this can be observed, if $i > j > k$, the absolute advantage of i over j and the absolute advantage of j over k leads to an absolute advantage of i over k. Similarly, we obtain such a transitivity for the comparative advantage. To illustrate, the comparative advantage of city i in skill h with respect to city j and the same comparative advantage for city j with respect to city k leads to $A(i,h)/A(i,m) > A(j,h)/A(j,m) > A(k,h)/A(k,m)$, which implies $A(i,h)/A(k,h) > A(i,m)/A(k,m)$, that is that city i has a comparative advantage in skill h compared with city k.

Sectoral decisions As a result of these assumptions and extending Lemma 1, the thresholds $\omega(m,c)$ and $\omega(h,c)$ are decreasing with city size. Furthermore, we can extend Lemma 6 and obtain that a change in p_z leads to a stronger decline in $\omega(h,c)$ in large cities and $\omega(m,c)$ increase by more in smaller cities.

Location decisions The description of location decision within a city as described in Lemmas 2, 3 and 4 given that these results apply for any city c. We then only need to describe how agents decide to choose locations between the different N cities.

To start with, for any $i \leq N-1$, there are locations in city $c \in \{1,...,i\}$ where the productivity of worker is strictly higher than what it could be in any city $c \geq i+1$. This happens for locations τ where productivity in city $c \leq i$ strictly exceeds what can be obtained in city $c > i$, even in the best location. More formally:

$$H(\omega(\tau(c)), M(\omega(\tau(c)), c), c)T(\tau(c)) > H(\omega(\tau(c)), M(\omega(\tau(c)), i+1), i+1)T(0) \tag{17}$$

where $\omega(\tau(c)) = K(c, \tau(c))$ is the value of ω occupying location $\tau(c)$ in city c. This defines a maximum value for the skill in city $i+1$, $\bar{\omega}(i+1)$ for which inequality (15) holds with equality. As a result, any agent with a skill higher than $\bar{\omega}(i+1)$ will decide to live only in cities $c \leq i$.

Below this threshold $\bar{\omega}(i+1)$, for each ω and for each τ, there exists $\tau' < \tau$ such that the productivities in city 1 and in city 2 are the same:

$$H(\omega(\tau), M(\omega(\tau, c), c), c)T(\tau) = H(\omega(\tau), M(\omega(\tau), i+1), i+1)T(\tau'). \tag{18}$$

which implies that this agent is indifferent in living between, at least, any city $c \leq i+1$. In the end, households are indifferent between a less desirable location in the more productive and larger city $c \leq i$ or a more desirable location in the less productive and smaller city $i+1$. Similarly, between two locations c and c' such that $c \leq c' \leq i$, households hesitate between more desirable locations in city c' and less desirable ones in city c.

Results As for the 2-city case, labor market polarization will happen in the aggregate and across cities. This results from sectoral decisions. As for the 2-city case, the distribution of skills is going

to be log-supermodular.

Similarly, if, for any city $c \in \{1, ..., N-1\}$ the comparative advantage of c over city $c-1$ in high skill occupations is sufficiently large, i.e. $A(i,h)/A(i-1,h)$ is sufficiently large compared with $A(i,m)/A(i-1,m)$ for all $i \leq N-1$, we also obtain the results of Proposition 7 about initial exposures and of Proposition 11.

B Proofs of Propositions

B.1 Proof of Lemma 1

First, $\omega(m,c)^{1-\phi} = A(l,c)/A(m,c)p(l)/p(m)$. This latter ratio is larger in city 2 than in city 1 given Assumption 1 and thus $\omega(1,m) < \omega(2,m)$ as $\phi < 1$.

Second, $\omega(h,c)$ solves the following equation:

$$\frac{A(h,c)}{A(m,c)}\frac{p(h)}{p(m)} = \omega(h,c)e^{-\eta\omega(h,c)} \tag{19}$$

On $[\underline{\omega}, \overline{\omega}]$, the rhs term is a decreasing function of ω. Given Assumption 1, the ratio $A(h,c)/A(m,c)$ is larger in city 1 than in city 2, thus implying $\omega(1,h) < \omega(2,h)$.

B.2 Proof of Lemma 2

The proof of Lemma 1 in Davis and Dingel (2014) still holds: otherwise, there exists $\tau' < \tau''$ such that $r(c,\tau') \leq r(c,\tau'')$. Thus, $U(c,\tau',\sigma,\omega) > U(c,\tau'',\sigma,\omega)$ for all σ and all ω. This contradicts the fact that τ'' has to maximize utility for some individual with some skill ω and sectoral decision σ.

$$A(h,1)e^{\gamma\omega(1,h)}T(\tau_1)p(h) - r(1,\tau_1) = A(h,2)e^{\gamma\omega(1,h)}T(\tau_2)p(h) - r(1,\tau_2) \tag{20}$$

$$\geq A(m,2)\omega T(\tau_1)p(m) - r(1,\tau_1) \tag{21}$$

B.3 Proof of Lemma 3

Here, we follow Davis and Dingel (2014), lemma 2 and Lemma 1 in Costinot and Vogel (2010). Let us first define

$$f(\omega,c,\tau) = \int_\sigma f(\omega,c,\tau,\sigma)d\sigma$$

$$\Omega(\tau,c) = \{\omega \in \Omega, f(\omega,c,\tau) > 0\}$$

$$\mathcal{T}(\omega,c) = \{\tau \in [0,\bar{\tau}(c)], f(c,\omega,\tau) > 0\}$$

Using these objects, we obtain:

(i) $\Omega(\tau, c) \neq \emptyset$ for $0 \leq \tau \leq \bar{\tau}(c)$ and $\tau(\omega, c) \neq \emptyset$ for at least one city as $f(\omega) > 0$.

(ii) $\Omega(\tau, c)$ is a non-empty interval for $0 \leq \tau \leq \bar{\tau}(c)$. If not, there exist $\omega < \omega' < \omega''$ such that $\omega, \omega'' \in \Omega(\tau)$ but not ω'. This means that there exists τ' such that $\omega' \in \Omega(\tau')$. Without loss of generality, suppose that $\tau' > \tau$. Utility maximization for both ω and ω' implies:

$$T(\tau')G(\omega', c) - r(c, \tau') \geq T(\tau)G(\omega', c) - r(c, \tau) \tag{22}$$

$$T(\tau)G(\omega, c) - r(c, \tau) \geq T(\tau')G(\omega, c) - r(c, \tau') \tag{23}$$

This jointly implies that $(T(\tau') - T(\tau))(G(\omega', c) - G(\omega, c)) \geq 0$, but this cannot be with $\tau' > \tau$ and $\omega' > \omega$. The same reasoning can be applied when $\tau' < \tau$. We can also conclude that for any $\tau < \tau'$, if $\omega \in \Omega(\tau)$ and $\omega' \in \Omega(\tau')$, then $\omega \geq \omega'$.

(iii) $\Omega(\tau, c)$ is a singleton for all but a countable subset of $[0, \bar{\tau}(c)]$. For any $\tau \in [0, \bar{\tau}(c)]$, $\Omega(\tau, c)$ is measurable as it a non-empty interval. Let $\mathcal{T}_0(c)$ denote the subset of locations τ such that $\mu(\Omega(\tau, c)) > 0$, μ being the Lebesgue measure over $\mathcal{R}$. Let us show that $\mathcal{T}_0(c)$ is a countable sets – any other $\Omega(\tau, c)$ where $\tau \notin \mathcal{T}_0(c)$ is a singleton as it is a interval with measure 0. For any $\tau \in \mathcal{T}_0(c)$, let us define $\underline{\omega}(\tau) \equiv \inf \Omega(\tau, c)$ and $\overline{\omega}(\tau) \equiv \sup \Omega(\tau, c)$. As $\mu(\Omega(\tau, c)) > 0$, $\underline{\omega}(\tau) < \overline{\omega}(\tau)$. Thus there exists a integer j such that $j(\overline{\omega}(\tau) - \underline{\omega}(\tau)) > (\bar{\omega}(c) - \underline{\omega})$. Given that $\mu(\Omega(\tau, c) \cap \Omega(\tau', c)) = 0$ for $\tau \neq \tau'$, for any j, we can then have at most j elements $\{\tau_1, ...\tau_j\} \equiv \mathcal{T}_j^0$ verifying $j(\overline{\omega}(\tau_i) - \underline{\omega}(\tau_i)) > (\bar{\omega}(c) - \underline{\omega})$. Thus $\mathcal{T}_j^0$ is countable. Given that $\mathcal{T}^0 = \bigcup_{j=1}^{\infty} \mathcal{T}_j^0$ and that the countable union of countable sets is also countable, we conclude that $\mathcal{T}^0$ is countable.

(iv) $\mathcal{T}(\omega, c)$ is a singleton for all but a countable subset of Ω. As in Davis and Dingel (2017), we use the arguments as in steps 2 and 3.

(v) $\Omega(\tau, c)$ is a singleton for any $\tau \in [0, \bar{\tau}(c)]$. Suppose not: there exists $\tau \in [0, \bar{\tau}(c)]$ so that $\Omega(\tau, c)$ is not a singleton. Given step (ii), it is then an interval with strictly positive measure. Step (iv) implies that $\mathcal{T}(\omega, c) = \{\tau\}$ for almost all $\omega \in \Omega(\tau, c)$ Hence we obtain:

$$f(c, \omega, \tau) = f(\omega)\delta^{\text{Dirac}}\left(1 - 1_{\Omega(c,\tau)}\right) \text{ for almost all } \omega \in \Omega(c, \tau). \tag{24}$$

This contradicts assumptions on $S(\tau)$ as this implies that $S'(\tau) = \infty$. TBC.

In the end, in city c, for any $\tau \in [0, \bar{\tau}(c)]$, there exists a unique ω such that $\omega \in \Omega(c, \tau)$. This does defines a function K_c such that $K_c(\tau) = \omega$. This function is weakly decreasing as shown by step (ii). Furthermore, as $\Omega(\tau) \neq \emptyset$ for all $\tau \in [0, \bar{\tau}(c)]$, K_c is continuous and satisfies $K_c(0) = \overline{\omega}(c)$ and $K_c(\bar{\tau}(c)) = \underline{\omega}$.

B.4 Proof of Lemma 4

By using the function $K_c(\tau)$ that is continuous and weakly decreasing from Lemma 3, there exist unique $\bar{\tau}(h,c)$ such that $K_c(\bar{\tau}(h,c)) = \omega(h,c)$ and $\bar{\tau}(m,c)$ such that $K_c(\bar{\tau}(m,c)) = \omega(m,c)$.

B.5 Proof of Lemma 5

$\Gamma(\omega,.)$ inherits the properties of the function T. For $\Gamma(.,\tau)$, the function is continuous and either constant or decreasing in each segment defined by the thresholds $\omega(h,c)$ and $\omega(m,c)$. Given the definition of the thresholds, the function is continuous everywhere and, thus, given it is either constant or decreasing in each segment, it is globally weakly decreasing.

B.6 Proof of Lemma 6

Let us now compute how a change in price of intermediate goods modifies the thresholds. By rewriting the indifference condition as:

$$\frac{H(\omega(h,c),h,c)}{H(\omega(h,c),m,c)} = \frac{p(m)}{p(h)} \tag{25}$$

we obtain, by differentiating both the right and the left hand terms:

$$\frac{d\left(\frac{H(\omega(h,c),h,c)}{H(\omega(h,c),m,c)}\right)}{\frac{H(\omega(h,c),h,c)}{H(\omega(h,c),m,c)}} = \frac{d\left(\frac{p(m)}{p(h)}\right)}{\frac{p(m)}{p(h)}} \tag{26}$$

Let us compute the different terms separately:

$$\frac{d\left(\frac{H(\omega(h,c),h)}{H(\omega(h,c),m)}\right)}{\frac{H(\omega(h,c),h)}{H(\omega(h,c),m)}} = \left(\frac{H_\omega(\omega(h,c),h)}{H(\omega(h,c),h)} - \frac{H_\omega(\omega(h,c),m)}{H(\omega(h,c),m)}\right) d\omega(h,c) \tag{27}$$

As a result, the effect of a relative decline in prices is such that:

$$d\omega(h,c) = \frac{1}{\Gamma(\omega(h,c),c)} \frac{d\left(\frac{p(m)}{p(h)}\right)}{\frac{p(m)}{p(h)}} \tag{28}$$

Given that H is log-supermodular,

$$\Gamma(\omega(h,c),c) = \frac{H_\omega(\omega(h,c),h,c)}{H(\omega(h,c),h,c)} - \frac{H_\omega(\omega(h,c),m,c)}{H(\omega(h,c),m,c)} > 0. \tag{29}$$

As a result, a decline in $p(m)/p(h)$ then leads to a decline in $\omega(h,c)$. Similarly, we obtain:

$$d\omega(m,c) = \frac{1}{\Gamma(\omega(m,c),c)} \frac{d\left(\frac{p(l)}{p(m)}\right)}{\frac{p(l)}{p(m)}} \tag{30}$$

As a result, an increase in $p(l)/p(m)$ then leads to an increase in $\omega(m,c)$.

We now want to know where the decline in $\omega(h,c)$ and the increase in $\omega(m,c)$ are the stronger.

For the first point, this amounts to comparing $\Gamma(\omega(1,h),1)$ and $\Gamma(\omega(2,h),2)$, that is to determine the sign of:

$$\frac{H_\omega(\omega(1,h),h,1)}{H(\omega(1,h),h,1)} - \frac{H_\omega(\omega(1,h),m,1)}{H(\omega(1,h),m,1)} - \frac{H_\omega(\omega(2,h),h,2)}{H(\omega(2,h),h,2)} + \frac{H_\omega(\omega(2,h),m,2)}{H(\omega(2,h),m,2)} \tag{31}$$

For the second point, this amounts to comparing $\Gamma(\omega(1,m),1)$ and $\Gamma(\omega(2,m),2)$, that is to determine the sign of:

$$\frac{H_\omega(\omega(1,m),m,1)}{H(\omega(1,m),m,1)} - \frac{H_\omega(\omega(1,m),l,1)}{H(\omega(1,m),l,1)} - \frac{H_\omega(\omega(2,m),m,2)}{H(\omega(2,m),m,2)} + \frac{H_\omega(\omega(2,m),l,2)}{H(\omega(2,m),l,2)} \tag{32}$$

Let us make some further assumption on the H function. Using our assumption on H and given that $\phi < 1$, this simplifies the two expressions into:

$$\frac{1}{\omega(1,m)} - \frac{1}{\omega(2,m)} \geq 0 \tag{33}$$

which is positive as $\omega(1,m) \leq \omega(2,m)$ and:

$$\frac{H_\omega(\omega(1,h),h,1)}{H(\omega(1,h),h,1)} - \frac{1}{\omega(1,h)} - \frac{H_\omega(\omega(2,h),h,2)}{H(\omega(2,h),h,2)} + \frac{1}{\omega(2,h)} \tag{34}$$

Let us investigate the sign of this expression. Note that it is negative as long as:

$$\frac{H_\omega(\omega(1,h),h,1)}{H(\omega(1,h),h,1)} - \frac{H_\omega(\omega(2,h),h,2)}{H(\omega(2,h),h,2)} \leq \frac{1}{\omega(1,h)} - \frac{1}{\omega(2,h)} \tag{35}$$

which is satisfied.

B.7 Proof of Proposition 7

In city c, the population of individuals in the middle-skill task, i.e. with skill ω is between $\omega(h,c)$ and $\omega(m,c)$, is:

$$L\int_{\omega(m,c)}^{\omega(h,c)} f(x,c)dx = S\left(T^{-1}\left(h(\omega(m,c),c)\right)\right) - S\left(T^{-1}\left(h(\omega(h,c),c)\right)\right) \tag{36}$$

where $K(T^{-1}(h(\omega,c)),c) = \omega$.

The share of middle skill agents in city c is:

$$s(m,c) = \frac{\int_{\omega(m,c)}^{\omega(h,c)} f(x,c)dx}{\int_{\underline{\omega}}^{\overline{\omega}(c)} f(x,c)dx} = \frac{S\left(T^{-1}\left(h(\omega(m,c),c)\right)\right) - S\left(T^{-1}\left(h(\omega(h,c),c)\right)\right)}{S\left(T^{-1}\left(h(\underline{\omega},c)\right)\right) - S\left(T^{-1}\left(h(\overline{\omega}(c),c)\right)\right)} \tag{37}$$

Using the continuity of the different function and given that $\omega(h,c)$ is decreasing in $A(c,h)/A(c,m)$ and $s(m,c) = 0$ when $A(c,h)/A(c,m) \to \infty$, we thus obtain that when $A(1,h)/A(1,m)$ is sufficiently large compared with $A(2,h)/A(2,m)$, shares satisfy $s(m,1) \leq s(m,2)$.

B.8 Proof of Proposition 8

To start with, let us derive the pdf of the distribution, $f(\omega,c)$. The population of individuals with skills between ω and $\omega + d\omega$ is:

$$L\int_{\omega}^{\omega+d\omega} f(x,c)dx = S\left(T^{-1}\left(h(\omega,c)\right)\right) - S\left(T^{-1}\left(h(\omega+d\omega,c)\right)\right) \tag{38}$$

Taking the derivative with respect to $d\omega$ and taking $d\omega \to 0$ yield:

$$f(\omega,c) = -\frac{\partial}{\partial\omega} S\left(T^{-1}\left(h(\omega,c)\right)\right) = h'(\omega,c)V(h(\omega,c)) \tag{39}$$

with $V(.) = -\frac{\partial}{\partial\omega} S\left(T^{-1}(.)\right)$.

Let us first note that $f(\omega,c)$ is log-supermodular if and only if, for all $\omega > \omega'$ and $c > c'$, we have:

$$f(\omega,c)f(\omega',c') > f(\omega',c)f(\omega,c').$$

When $f(\omega,c')$ and $f(\omega',c')$ are different than 0, this condition amounts to verifying than $f(\omega,c)/f(\omega,c')$ is strictly increasing or, equivalently that:

$$f'(\omega,c)f(\omega,c') > f'(\omega,c')f(\omega,c). \tag{40}$$

Using the fact that $f(\omega, c) = h'(\omega, c)V(h(\omega, c))$, we can compute:

$$f'(\omega, c) = h''(\omega, c)V(h(\omega, c)) + (h'(\omega, c))^2 V'(h(\omega, c))$$

By denoting $\eta(V, h(\omega, c)) = h(\omega, c)V'(h(\omega, c))/V(h(\omega, c))$, we obtain that:

$$\begin{aligned} f'(\omega, c) =& h''(\omega, c)V(h(\omega, c)) + (h'(\omega, c))^2 \eta(V, h(\omega, c))V(h(\omega, c))/h(\omega, c) \\ =& f(\omega, c)\left(\frac{h''(\omega, c)}{h'(\omega, c)} + \eta(V, h(\omega, c))\frac{h'(\omega, c)}{h(\omega, c)}\right) \end{aligned}$$

Replacing $f'(\omega, c)$ and $f'(\omega, c')$ by their values in (40), we then obtain the following condition:

$$\frac{h''(\omega, c)}{h'(\omega, c)} + \eta(V, h(\omega, c))\frac{h'(\omega, c)}{h(\omega, c)} > \frac{h''(\omega, c')}{h'(\omega, c')} + \eta(V, h(\omega, c'))\frac{h'(\omega, c')}{h(\omega, c')} \tag{41}$$

A straightforward implication of this necessary and sufficient condition is the following.

Lemma 13. **(i)** *If, for ω and ω' and for c and c', the occupation decisions are the same across cities, that is $M(\omega, c) = M(\omega, c')$ and $M(\omega', c) = M(\omega', c')$, then*

$$f(\omega, c)f(\omega', c') > f(\omega', c)f(\omega, c')$$

if and only if $\eta(V, x)$ is decreasing in x.

(ii) *If productivities are constant across occupations, $A(c, h) = A(c, m) = A(c, l)$ as in Davis and Dingel (2014), a necessary and sufficient condition for $f(\omega, c)$ to be log-supermodular is that $\eta(V, x)$ is decreasing in x.*

Proof. Suppose that $M(\omega, c) = M(\omega, c')$ and $M(\omega', c) = M(\omega', c')$, then, in equilibrium:

$$A(c, M(\omega, c))H(\omega, M(\omega, c))h(\omega, c) = A(c', M(\omega, c'))H(\omega, M(\omega, c'))h(\omega, c')$$

and thus $h(\omega, c) = h(\omega, c')$. By continuity, $M(\omega, c) = M(\omega, c')$ on a (right- or left-) neighborhood of ω and thus $h(., c) = h(., c')$ on this neighborhood, thus ensuring that locally $h''(., c) = h''(., c')$ and $h'(., c) = h'(., c')$ and in particular that $h''(\omega, c) = h''(\omega, c')$ and $h'(\omega, c) = h'(\omega, c')$. In the end, (41) simplifies into:

$$\eta(V, h(\omega, c)) > \eta(V, h(\omega, c')) \tag{42}$$

which is satisfied as long as V features decreasing elasticity.

□

The conclusion of the second point is that, with V featuring decreasing elasticity, we obtain that

f is log supermodular on subsets where the occupation decisions are the same, that is $[\omega(2,h),\overline{\omega}]$, $[\omega(2,m),\omega(1,h)]$ and $[\underline{\omega},\omega(1,m)]$.

Let us now turn to the segments $[\omega(1,h),\omega(2,h)]$ and $[\omega(1,m),\omega(2,m)]$, where households have different occupation choices depending on cities. Let us first show that it is sufficient to show that f is log-supermodular on each of these two segments to obtain log-supermodularity on $[\underline{\omega},\overline{\omega}]$.

Lemma 14. *Suppose that $f(x,c)$ is log-supermodular in $\{x,c\}$ on $[\underline{x},\overline{x}]$ and $[\overline{x},\overline{\overline{x}}]$, then $f(x,c)$ is log-supermodular in $\{x,c\}$ on $[\underline{x},\overline{\overline{x}}]$*

Proof. Let us consider any x and x' in $[\underline{x},\overline{\overline{x}}]$ such that $x > x'$. Let us also consider two cities c and c' such that $c > c'$. If x and x' are both in the same segment, either $[\underline{x},\overline{x}]$ or $[\overline{x},\overline{\overline{x}}]$, we already have log-supermodularity. So, let us consider the case where $x \geq \overline{x} \geq x'$.

Using log-supermodularity on $[\overline{x},\overline{\overline{x}}]$, we have:

$$\frac{f(x,c)}{f(x,c')} > \frac{f(\overline{x},c)}{f(\overline{x},c')}$$

Using log-supermodularity on $[\underline{x},\overline{x}]$, we have:

$$\frac{f(\overline{x},c)}{f(\overline{x},c')} > \frac{f(x',c)}{f(x',c')}$$

Combining these two equations, we obtain:

$$\frac{f(x,c)}{f(x,c')} > \frac{f(x',c)}{f(x',c')}.$$

In the end, f is then log-supermodular on $[\underline{x},\overline{\overline{x}}]$.

□

We now need to establish log-supermodularity on $[\omega(1,h),\omega(2,h)]$ and $[\omega(1,m),\omega(2,m)]$.

Let us start with some properties on the $h(\omega,c)$ function. The indifference condition between location implies that:

$$\phi(\omega) = H(\omega, M(\omega,c),c)h(\omega,c) = H(\omega,M(\omega,c'),c')h(\omega,c')$$

Given that $M(\omega,c) \geq M(\omega,c')$ due to the comparative advantage of the large city and that $H(\omega,M(\omega,c),c) \geq H(\omega,M(\omega,c'),c')$, we have that $h(\omega,c) \leq h(\omega,c')$. Furthermore, given that $H(\omega,M(\omega,c),c)/H(\omega,M(\omega,c'),c')$ is an increasing function of ω, we obtain that $h(\omega,c')/h(\omega,c)$ is an increasing function of ω and $h'(\omega,c) \leq h'(\omega,c')$. Finally, $H(\omega,M(\omega,c),c)$ being log-supermodular, we obtain that

$$h(\omega,1)h(\omega',2) \leq h(\omega',1)h(\omega,2)$$

and that

$$\frac{h'(\omega,1)}{h(\omega,1)} \leq \frac{h'(\omega,2)}{h(\omega,2)}$$

A first conclusion is then that when $\eta(V) \leq 0$ and decreasing, we obtain that:

$$\eta(V,h(\omega,1))\frac{h'(\omega,1)}{h(\omega,1)} > \eta(V,h(\omega,2))\frac{h'(\omega,2)}{h(\omega,2)}. \tag{43}$$

Second, note that (41) is invariant to equilibrium prices.

In the end, when $\eta(V,h(\omega,1))$ is sufficiently decreasing, condition (41) is satisfied.

B.9 Proof of Proposition 10

The change in percentage points of the share of middle skill is:

$$ds(m,c) = \frac{\left(S\left(T^{-1}\left(h(\omega(m,c),c)\right)\right)\right)' d\omega(m,c) - \left(S\left(T^{-1}\left(h(\omega(h,c),c)\right)\right)\right)' d\omega(h,c)}{S\left(T^{-1}\left(h(\underline{\omega},c)\right)\right) - S\left(T^{-1}\left(h(\overline{\omega}(c),c)\right)\right)} \tag{44}$$

Given that $d\omega(m,c) > 0$ and $d\omega(h,c) < 0$ when $dp(k) < 0$, we obtain that $s(m,c)$ declines in both cities c. As these shares decline in both cities, it also declines overall.

B.10 Proof of Proposition 11

We want to compute the variation in the share of middle-skill workers in a given city,

$$ds(m,c) = \frac{V\left(h(\omega(h,c),c)\right) d\omega(h,c) - V\left(h(\omega(m,c),c)\right) d\omega(m,c)}{S\left(T^{-1}\left(h(\underline{\omega},c)\right)\right) - S\left(T^{-1}\left(h(\overline{\omega}(c),c)\right)\right)} \tag{45}$$

We obtain:

$$ds(m,c) = \frac{\left(\frac{V(h(\omega(h,c),c))}{\Gamma(\omega(h,c),c,h)} + \frac{V(h(\omega(m,c),c))}{\Gamma(\omega(m,c),c,m)}\right)}{S\left(T^{-1}\left(h(\underline{\omega},c)\right)\right) - S\left(T^{-1}\left(h(\overline{\omega}(c),c)\right)\right)}\frac{dp}{p} \tag{46}$$

Using the expressions for $\Gamma(\omega(m,c),c,m)$ and $\Gamma(\omega(h,c),c,h)$, the coefficient can be rewritten as:

$$\frac{\left(V\left(h(\omega(h,c),c)\right)\frac{\omega(h,c)}{\gamma\omega(h,c)-1} + V\left(h(\omega(m,c),c)\right)\omega(m,c)(1-\phi)\right)}{S\left(T^{-1}\left(h(\underline{\omega},c)\right)\right) - S\left(T^{-1}\left(h(\overline{\omega}(c),c)\right)\right)} \tag{47}$$

We have easily that:

$$V\left(h(\omega(1,h),1)\right)\frac{\omega(1,h)}{\gamma\omega(1,h)-1} > V\left(h(\omega(2,h),2)\right)\frac{\omega(2,h)}{\gamma\omega(2,h)-1}. \tag{48}$$

On the other hand, the sign of the following term is ambiguous:

$$V\left(h(\omega(1,m),1)\right)\omega(1,m) - V\left(h(\omega(2,m),2)\right)\omega(2,m) \tag{49}$$

When $A(1,m)$ is close to $A(2,m)$, we then find that a sufficient condition for the decline to be larger in the smaller city is:

$$\frac{\left(V\left(h(\omega(1,h),1)\right)\frac{\omega(1,h)}{\gamma\omega(1,h)-1}\right)}{S\left(T^{-1}\left(h(\underline{\omega},1)\right)\right) - S\left(T^{-1}\left(h(\overline{\omega}(1),1)\right)\right)} > \frac{\left(V\left(h(\omega(2,h),2)\right)\frac{\omega(2,h)}{\gamma\omega(2,h)-1}\right)}{S\left(T^{-1}\left(h(\underline{\omega},2)\right)\right) - S\left(T^{-1}\left(h(\overline{\omega}(2),2)\right)\right)} \tag{50}$$

this condition is satisfied when $A(1,h)$ is sufficient large. Indeed, it can be observed that the left hand term diverge to ∞ when $A(1,h) \to \infty$.

C Figures and tables

Table 8 – City categories: number of cities, population and the share of hours worked in 2015.

city size	number	total population	share of hours worked
>2,000,000	1	10,706,072	.375
750,000-2,000,000	6	7,060,599	.206
500,000-750,000	4	2,219,618	.055
200,000-500,000	22	6,691,222	.169
100,000-200,000	22	3,245,887	.083
50,000-100,000	62	4,414,317	.112
Total	117	34,337,715	

Figure 1 – Map of France with largest metropolitan areas in 2015.

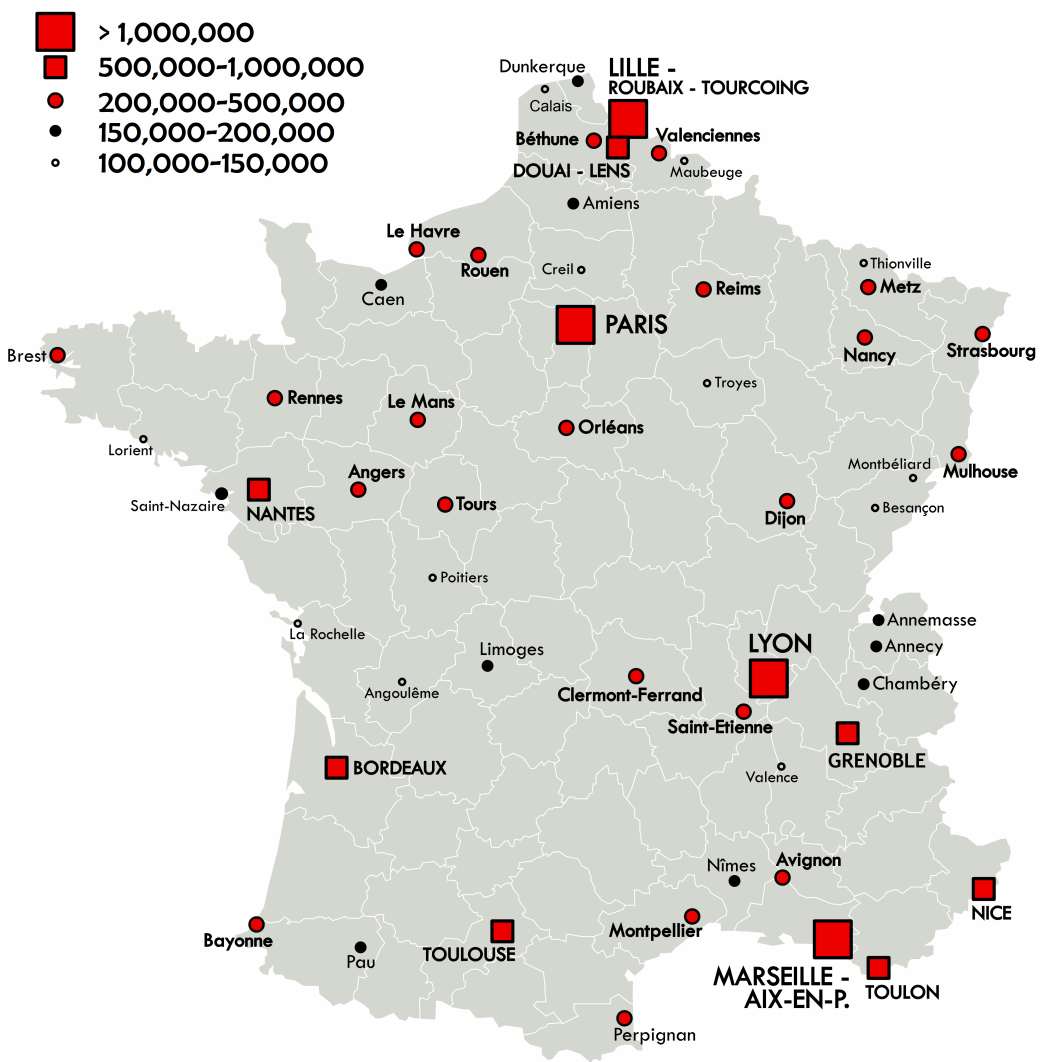
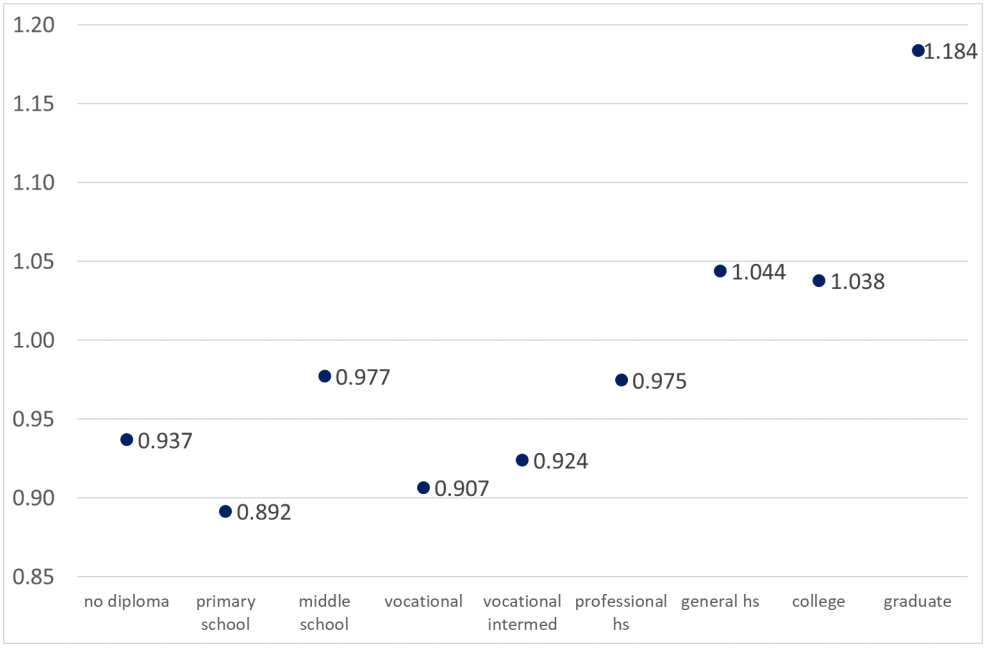


Table 9 – Summary statistics at the city level.

Item		year	mean	stdev	min	max
population		2015	293,485	1,007,302	50,571	10,706,072
number of firms with jobs in the city		1994	3,522	141,922	529	13,283
		2015	5,728	222,251	881	20,903
employment share	high paying jobs	1994	0.090	0.027	0.052	0.233
		2015	0.139	0.049	0.080	0.367
	middle-paying jobs	1994	0.775	0.050	0.601	0.872
		2015	0.651	0.059	0.449	0.799
	low paying jobs	1994	0.135	0.041	0.063	0.312
		2015	0.210	0.051	0.084	0.440
	RTI4 jobs	1994	0.419	0.078	0.240	0.625
		2015	0.307	0.055	0.175	0.489
	RTI6 jobs	1994	0.523	0.067	0.355	0.705
		2015	0.451	0.059	0.266	0.581
	OFF6 jobs	1994	0.564	0.065	0.379	0.728
		2015	0.415	0.055	0.258	0.576
employment share percentage change 1994-2015	high paying jobs		0.048	0.029	- 0.005	0.153
	middle-paying jobs		- 0.124	0.030	- 0.204	- 0.003
	low paying jobs		0.076	0.024	- 0.022	0.137
	RTI4 jobs		- 0.112	0.041	- 0.255	0.022
	RTI6 jobs		- 0.072	0.036	- 0.172	0.044
	OFF6 jobs		- 0.150	0.036	- 0.255	0.026
employment share growth 1994-2015	high paying jobs		0.532	0.251	- 0.049	1.508
	middle-paying jobs		- 0.161	0.042	- 0.309	- 0.004
	low paying jobs		0.601	0.231	- 0.097	1.420
	RTI4 jobs		- 0.263	0.074	- 0.411	0.081
	RTI6 jobs		- 0.136	0.065	- 0.298	0.095
	OFF6 jobs		- 0.265	0.057	- 0.387	0.059

Figure 2 – Log-supermodularity, population elasticities by diploma (9 categories) in the 1999 Census data.



Notes: 112 cities > 50,000 inhabitants defined by INSEE as of 1999, population figures from 1999. Exclusions in terms of CS and age as in main sample.

Table 10 – Log-supermodularity, population elasticities by diploma (9 categories) in the 1999 Census data.

Dependent variable: ln $f(c,\omega)$	All workers	French born	Population share	French born share
No diploma X ln pop	0.94*	0.91***	.12	.75
	(0.032)	(0.0254)		
End of primary school X ln pop	0.89***	0.88***	.08	.89
	(0.033)	(0.029)		
End of middle school (collège) X ln pop	0.98	0.97	.07	.94
	(0.024)	(0.023)		
Vocational school diploma (CAP) X ln pop	0.91***	0.90***	.20	.96
	(0.025)	(0.024)		
Vocational high school intermediate diploma (BEP) X ln pop	0.92***	0.92***	.10	.96
	(0.022)	(0.022)		
High school vocational diploma (bac technologique or professionnel) X ln pop	0.98	0.97	.09	.97
	(0.026)	(0.026)		
General high school diploma (Bac) X ln pop	1.04	1.04	.06	.93
	(0.029)	(0.028)		
Undergraduate studies X ln pop	1.04	1.03	.15	.97
	(0.026)	(0.026)		
Graduate studies X ln pop	1.18***	1.18***	.14	.94
	(0.037)	(0.037)		

Notes: 112 cities > 50,000 inhabitants defined by INSEE as of 1999. The variable "ln pop" is the natural logarithm of metropolitan area population from 1999. Exclusions in terms of CS and age as in main sample. Robust standard errors in parantheses. ***, **, and * denote statistical significance at the 1 %, 5 %, and 10 % levels for the test of hypotheses whether a given coefficient is equal to one.

Table 11 – Population elasticities by high, middle and low paying categories in the 1999 Census data.

Dependent variable: ln $f(c,\omega)$	All workers	French born	Population share	French born share
High paying X ln pop	1.14***	1.14***	.17	.85
	(0.037)	(0.036)		
Middle-paying X ln pop	0.95*	0.95**	.65	.94
	(0.025)	(0.024)		
Low-paying X ln pop	0.94***	0.92***	.18	.96
	(0.018)	(0.015)		

Notes: 112 cities above 50,000 inhabitants as defined by INSEE as of 1999. The variable "ln pop" is the natural logarithm of metropolitan area population from 1999. Exclusions in terms of CS and age as in main sample. Robust standard errors in parantheses. ***, **, and * denote statistical significance at the 1 %, 5 %, and 10 % levels for the test of hypotheses whether a given coefficient is equal to one.

Table 12 – Population elasticities by 1-digit CS categories in the 1999 Census data.

Dependent variable: ln $f(c,\omega)$	All workers	French born	Population share	French born share
Cadres (CS 3) X ln pop	1.16***	1.15***	.17	.96
	(0.038)	(0.037)		
Intermediate professionals (CS 4) X ln pop	1.02	1.02	.30	.97
	(0.026)	(0.026)		
Low-skill employees (CS 5) X ln pop	0.97	0.96**	.27	.92
	(0.021)	(0.019)		
Blue-collar workers (CS 6) X ln pop	0.88***	0.86***	.26	.86
	(0.027)	(0.026)		

Notes: 112 cities above 50,000 inhabitants as defined by INSEE as of 1999. The variable "ln pop" is the natural logarithm of metropolitan area population from 1999. Exclusions in terms of CS and age as in main sample. CS 23 category – CEOs – not included in the category "cadres". Robust standard errors in parantheses. ***, **, and * denote statistical significance at the 1 %, 5 %, and 10 % levels for the test of hypotheses whether a given coefficient is equal to one.

Table 13 – Share of 3 lowest-paying occupations per metropolitan area size.

Agglo.size	Paris	> .75M	.5-.75M	.2-.5M	.1-.2M	.05-.1M
1994	0.11	0.12	0.12	0.12	0.11	0.12
2015	0.16	0.16	0.17	0.18	0.19	0.20
change in ppct	0.05	0.05	0.05	0.06	0.08	0.08
growth in %	0.51	0.42	0.43	0.54	0.69	0.64

Note: Percentage point changes and growth rates not directly calculable from the upper two lines of the table due to rounding.

Table 14 – Share of 6 most-offshorable occupations per metropolitan area size.

Agglo.size	Paris	> .75M	.5-.75M	.2-.5M	.1-.2M	.05-.1M
1994	0.49	0.54	0.54	0.56	0.58	0.59
2015	0.30	0.36	0.38	0.41	0.42	0.43
change in ppct	-0.19	-0.17	-0.16	-0.16	-0.16	-0.16
growth in %	-0.39	-0.33	-0.29	-0.28	-0.28	-0.27

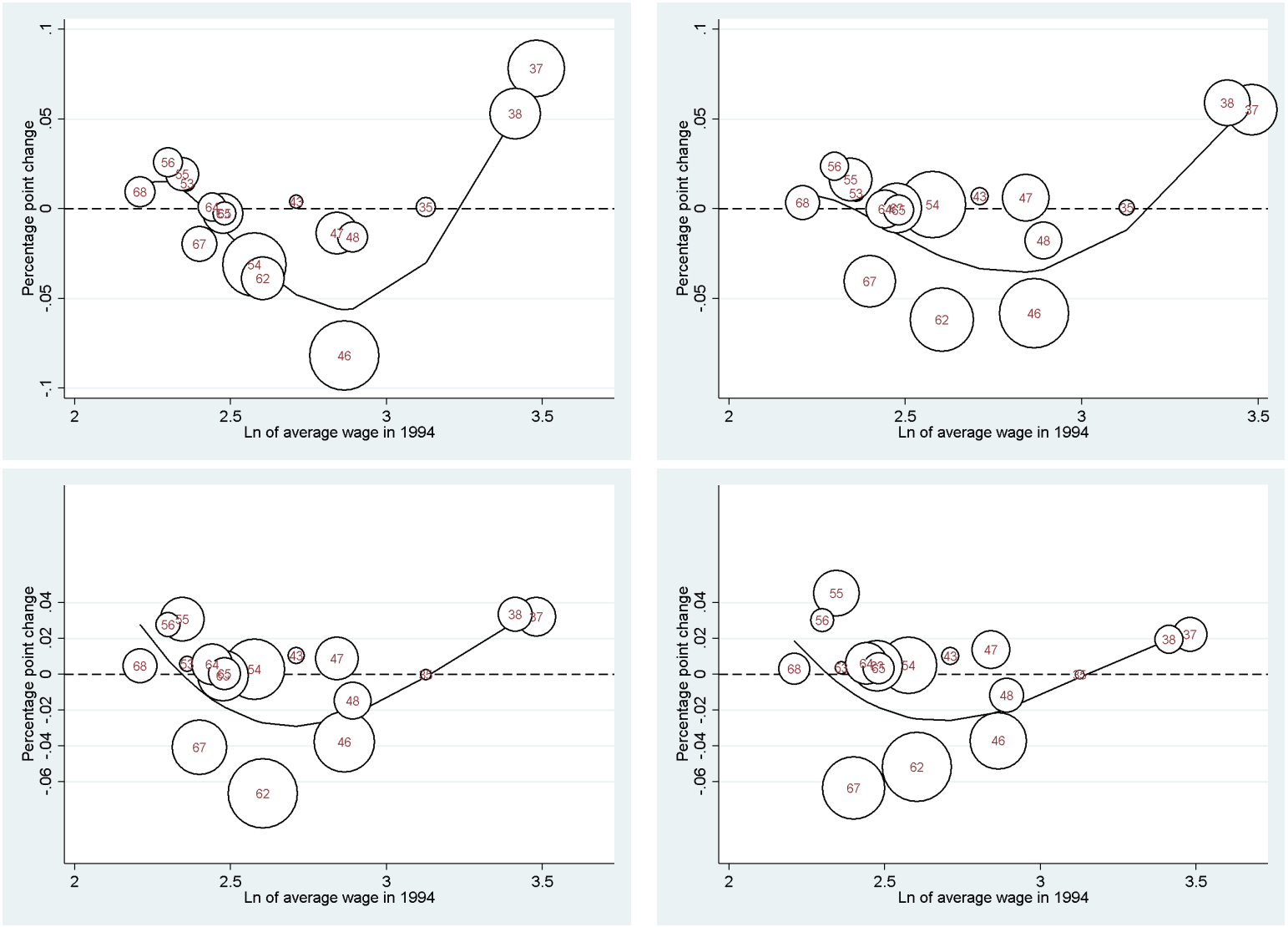
Note: Percentage point changes and growth rates not directly calculable from the upper two lines of the table due to rounding.

Figure 3 – Labor market polarization in France 1994-2015.



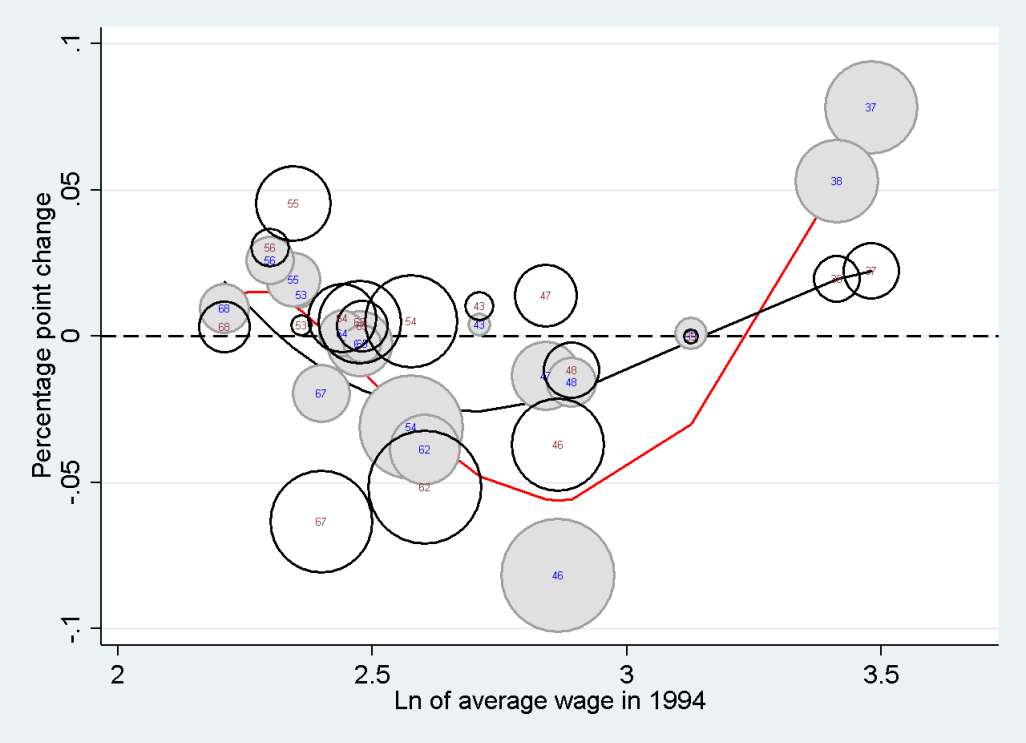
Notes: The figure shows the percentage point change in employment of the considered 2-digit CS categories plotted against their average wage in cities > .05m in 1994. Numbers pertain to 2-digit CS categories represented. Circle sizes correspond to the employment shares. The line shows a cubic relationship between the average wage and the percentage point change. The CS category "23" - CEOs excluded.

Figure 4 – Labor market polarization across cities 1994-2015.



Left-upper panel: Paris. Right-upper panel: Lyon. Left-lower panel: cities between 0.2-0.5m. Right-lower panel: cities between .05-.1m inhabitants.
Notes: Figures show the percentage point change in employment of the considered 2-digit CS categories plotted against their average wage in cities > 0.05m in 1994. Numbers pertain to 2-digit CS categories represented. Circle sizes correspond to the employment shares. The line shows a cubic relationship between the average wage and the percentage point change. The CS category "23" - CEOs excluded.

Figure 5 – Comparing labor market polarization 1994-2015 in Paris and cities between .05-.1m inhabitants.



Notes: The figure shows the percentage point change in employment of the considered 2-digit CS categories plotted against their average wage in cities > .05m in 1994. Numbers pertain to 2-digit CS categories represented. Grey circles stand for Parisian while white for small city shares. Circle sizes correspond to the employment shares (same scale for the two compared groups). The two lines shows a cubic relationship between the average wage and the percentage point changes in employment for Paris (red) and cities between .05-.1m inhabitants (black) respectively. The CS category "23" - CEOs excluded.

Table 15 – Comparison of means of employment shares in different occupations, cities >0.5m vs. 0.05-0.1m.

Item	RTI4	RTI6	OFF6	middle-paid	low-paid	high-paid
1994						
mean, cities >0.5m	0.326	0.401	0.509	0.690	0.123	0.187
mean, cities 0.05-0.1m	0.423	0.534	0.566	0.778	0.141	0.081
difference in pct points	-0.097***	-0.133***	-0.057***	-0.088***	-0.019**	0.107***
2015						
mean, cities >0.5m	0.219	0.310	0.331	0.509	0.187	0.303
mean, cities 0.05-0.1m	0.314	0.471	0.421	0.663	0.220	0.117
difference in pct points	-0.095***	-0.161***	-0.090***	-0.153***	-0.033***	0.187***

Notes: 2015 population weighted, robust standard errors. N=73 (11 cities > 0.5m). ***, **, and * denote statistical significance at the 1 %, 5 %, and 10 % levels for the tests of equality of means between the groups of small and large cities.

Table 16 – Comparison of means of changes in different occupations, cities >0.5m vs. 0.05-0.1m. Sample without finance, insurance and real estate sectors.

Item	RTI4	RTI6	OFF6	middle-paid	low-paid	high-paid
mean change ppct, >0.5m	-.109	-.09	-.173	-.177	.072	.105
mean change ppct, 0.05-0.1m	-.117	-.069	-.147	-.115	.084	.031
difference in pct points	.008	-.021***	-.026***	-.062***	-.012**	.074***
mean growth, >0.5m	-.335	-.224	-.354	-.259	.562	.60
mean growth, 0.05-0.1m	-.268	-.124	-.263	-.148	.628	.415
difference in growth	-.067***	-.100***	-.091***	-0.111***	-.066	.185***

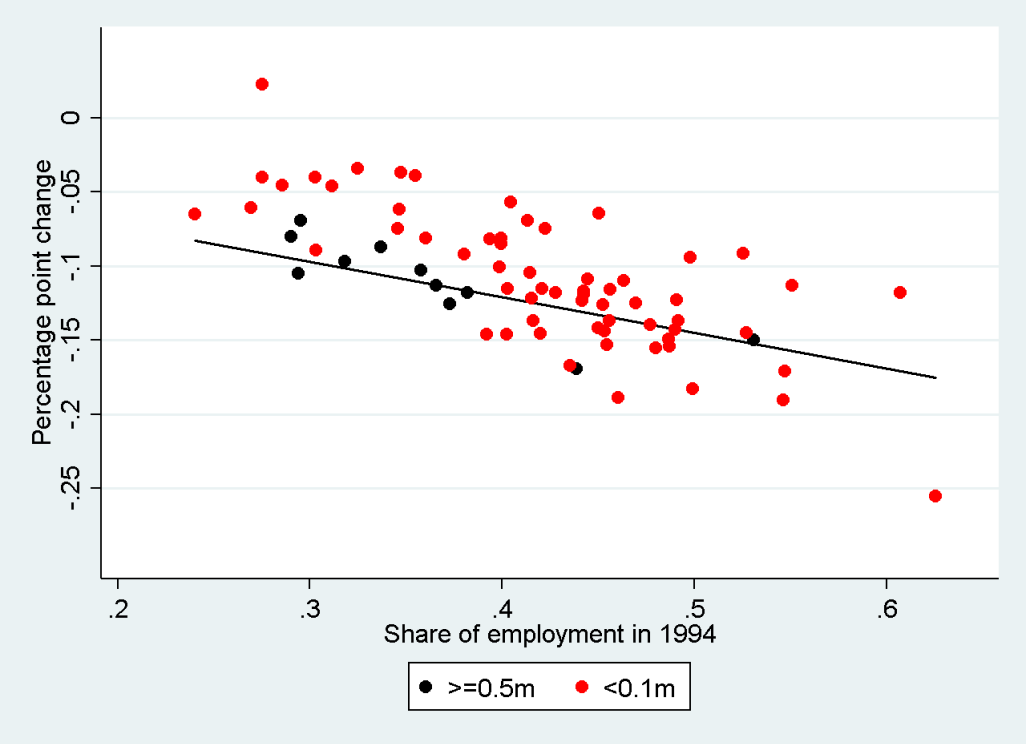
Notes: population weighted, allowing for unequal variances. N=73 (11 cities > 0.5m). ***, **, and * denote statistical significance at the 1 %, 5 %, and 10 % levels for the tests of equality of means between the groups of small and large cities.

Figure 6 – Comparing labor market polarization 1994-2015 in Lyon and cities between .05-1m inhabitants.



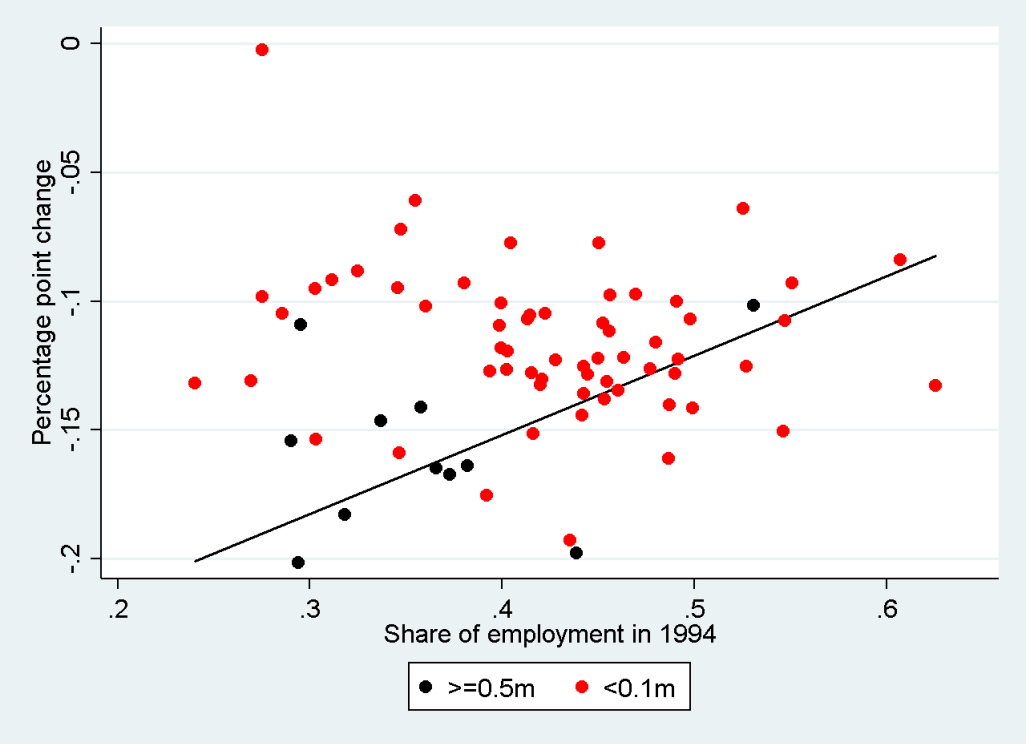
Notes: The figure shows the percentage point change in employment of the considered 2-digit CS categories plotted against their average wage in cities > 0.05m in 1994. Numbers pertain to 2-digit CS categories represented. Grey circles stand for Lyon while white for small city shares. Circle sizes correspond to the employment shares (same scale for the two compared groups). The two lines shows a cubic relationship between the average wage and the percentage point changes in employment for Lyon (red) and cities between 0.05-0.1m inhabitants (black) respectively. The CS category "23" - CEOs excluded.

Figure 7 – Exposure to RTI4 jobs and change in RTI4 jobs in cities >0.5m and cities between 0.05-0.1m inhabitants, 1994-2015.



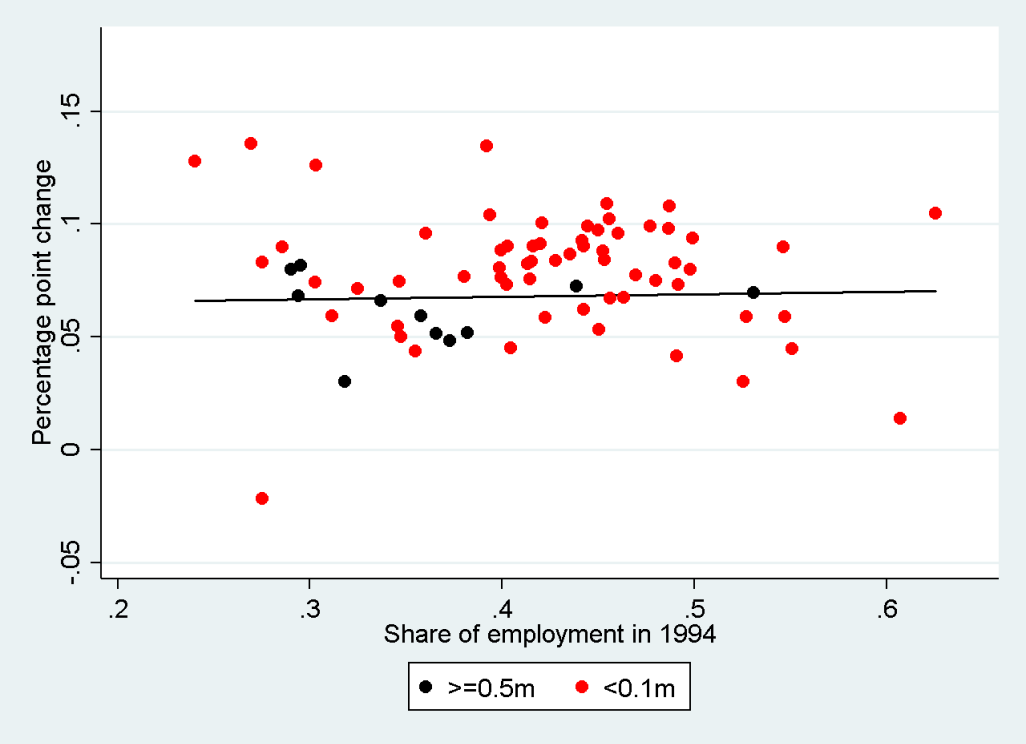
Notes: The figure shows the percentage point change in employment of RTI4 jobs (CS 48, 54, 62 and 67) between 1994-2015 plotted against the share of RTI4 jobs in employment in 1994. Each black (red) dot represents a city above .5m (between 0.05-0.1m). The line shows a linear, population-weighted fit of the relation between employment changes and the initial RTI4 exposure.

Figure 8 – Exposure to RTI4 jobs and change in middle-paying jobs in cities >0.5m and cities between 0.05-0.1m inhabitants, 1994-2015.



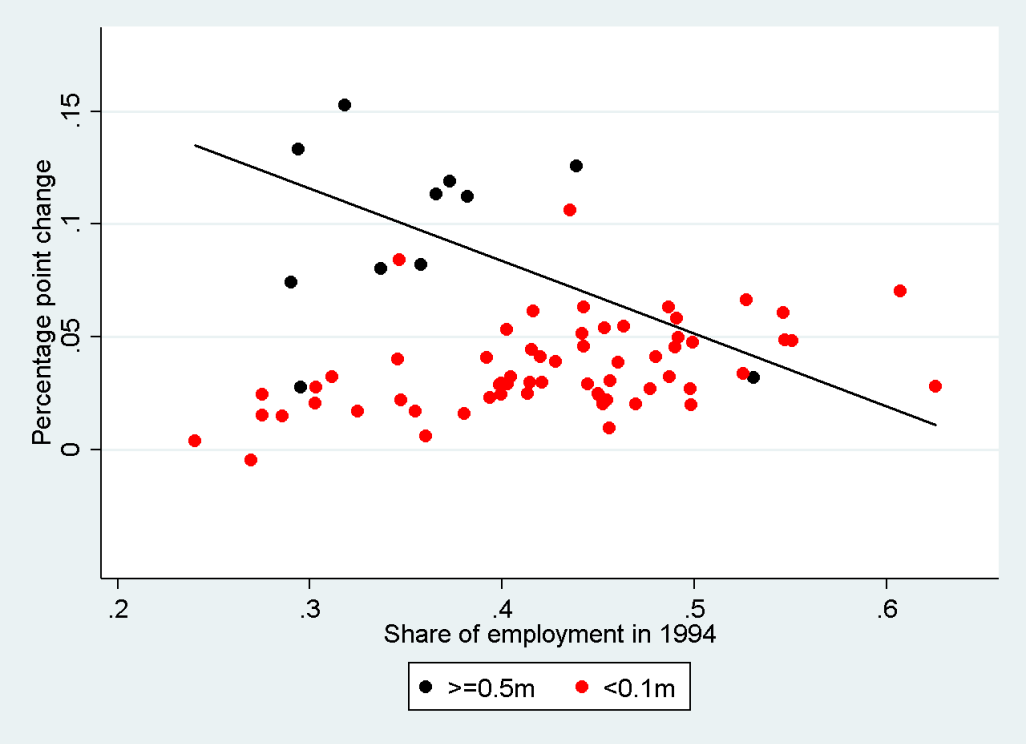
Notes: The figure shows the percentage point change in employment of middle-paying jobs between 1994-2015 plotted against the share of RTI4 jobs (CS 48, 54, 62 and 67) in employment in 1994. Each black (red) dot represents a city above 0.5m (between 0.05-0.1m). The line shows a linear, population-weighted fit of the relation between employment changes and the initial RTI4 exposure.

Figure 9 – Exposure to RTI4 jobs and change in low-paying jobs in cities >0.5m and cities between 0.05-0.1m inhabitants, 1994-2015.



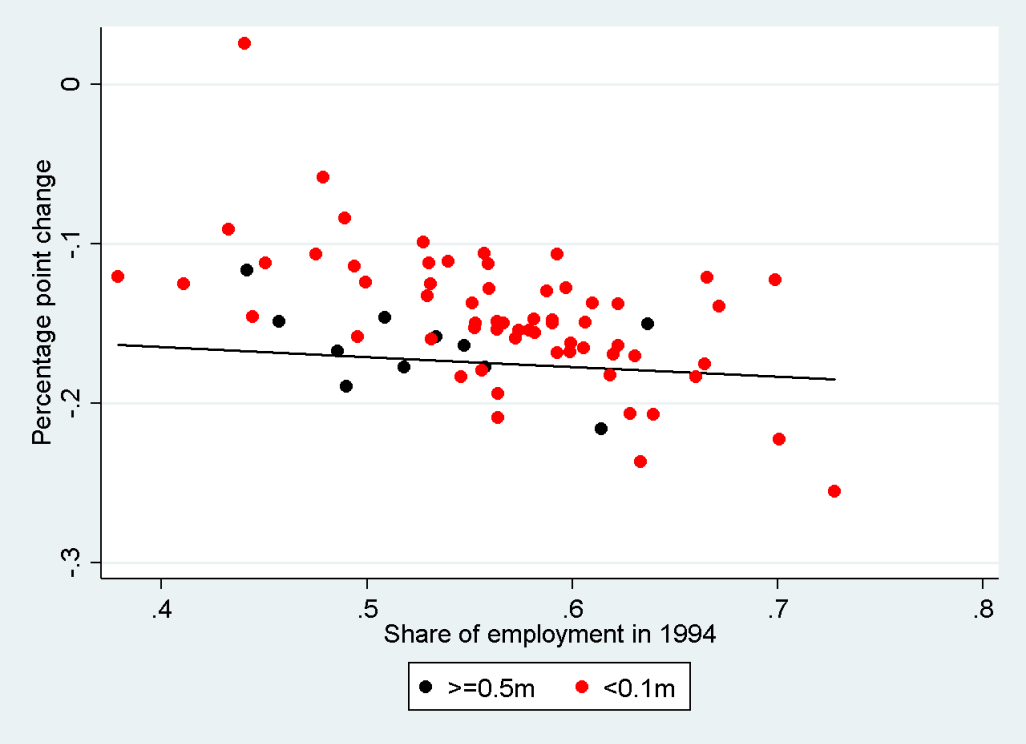
Notes: The figure shows the percentage point change in employment of low-paying jobs (CS 53, 55, 56 and 68) between 1994-2015 plotted against the share of RTI4 jobs (CS 48, 54, 62 and 67) in employment in 1994. Each black (red) dot represents a city above .5m (between 0.05-0.1m). The line shows a linear, population-weighted fit of the relation between employment changes and the initial RTI4 exposure.

Figure 10 – Exposure to RTI4 jobs and change in high-paying jobs in cities >0.5m and cities between 0.05-0.1m inhabitants, 1994-2015.



Notes: The figure shows the percentage point change in employment of high-paying jobs (CS 23, 35, 37 and 38) between 1994-2015 plotted against the share of RTI4 jobs (CS 48, 54, 62 and 67) in employment in 1994. Each black (red) dot represents a city above .5m (between 0.05-0.1m). The line shows a linear, population-weighted fit of the relation between employment changes and the initial RTI4 exposure.

Figure 11 – Exposure to 6 most offshorable jobs and their employment share change in cities >0.5m and cities between 0.05-0.1m inhabitants, 1994-2015.



Notes: The figure shows the percentage point change in employment of jobs with highest OFF-GMS index (CS 46, 48, 54, 62, 65, 67) between 1994-2015 plotted against the share of such jobs in employment in 1994. Each black (red) dot represents a city above .5m (between 0.05-0.1m). The line shows a linear, population-weighted fit of the relation between employment changes and the initial exposure.

Table 17 – Employment share changes 1994-2015 and exposure to 4 most routine/offshorable occupations (RTI4) in 1994

Employment share change of RTI4 jobs								
employment share of RTI4 in 1994	-0.240***	-0.401***	-0.273***	-0.425***	-0.261***	-0.409***	-0.255***	-0.369***
	(0.060)	(0.057)	(0.063)	(0.058)	(0.063)	(0.058)	(0.060)	(0.049)
large X employment share of RTI4 in 1994			-0.067***	-0.071***	-0.069***	-0.073***	-0.064***	-0.064***
			(0.023)	(0.024)	(0.023)	(0.024)	(0.022)	(0.021)
constant	-0.025	0.057**	0.004	0.070***	0.001	0.066***	-0.003	0.047**
	(0.023)	(0.023)	(0.027)	(0.024)	(0.027)	(0.024)	(0.026)	(0.020)
R^2	0.388	0.535	0.493	0.577	0.478	0.538	0.483	0.514
Employment share change of middle-paid jobs								
employment share of RTI4 in 1994	0.307***	0.009	0.254**	-0.027	0.266**	-0.031	0.039	-0.002
	(0.098)	(0.055)	(0.106)	(0.059)	(0.107)	(0.069)	(0.106)	(0.046)
large X employment share of RTI4 in 1994			-0.106***	-0.108***	-0.106***	-0.110***	-0.099**	-0.089**
			(0.035)	(0.039)	(0.035)	(0.041)	(0.038)	(0.036)
constant	-0.275***	-0.126***	-0.223***	-0.105***	-0.232***	-0.103***	-0.135***	-0.117***
	(0.040)	(0.025)	(0.046)	(0.027)	(0.047)	(0.031)	(0.046)	(0.020)
R^2	0.319	0.001	0.453	0.181	0.460	0.186	0.335	0.165
Employment share change of low-paid jobs								
employment share of RTI4 in 1994	0.011	-0.016	-0.012	-0.033	-0.012	-0.029	-0.015	-0.052
	(0.023)	(0.054)	(0.033)	(0.057)	(0.033)	(0.062)	(0.033)	(0.046)
large X employment share of RTI4 in 1994			-0.046***	-0.051***	-0.043***	-0.047**	-0.048***	-0.056***
			(0.016)	(0.019)	(0.016)	(0.019)	(0.016)	(0.018)
constant	0.064***	0.083***	0.084***	0.093***	0.083***	0.091***	0.086***	0.102***
	(0.008)	(0.023)	(0.014)	(0.025)	(0.014)	(0.027)	(0.014)	(0.020)
R^2	0.002	0.002	0.145	0.066	0.130	0.059	0.177	0.113
Employment share change of high-paid jobs								
employment share of RTI4 in 1994	-0.318***	0.007	-0.242**	0.060	-0.254**	0.060	-0.240**	0.059
	(0.094)	(0.048)	(0.107)	(0.048)	(0.108)	(0.055)	(0.113)	(0.045)
large X employment share of RTI4 in 1994			0.153***	0.160***	0.149***	0.158***	0.146***	0.142***
			(0.036)	(0.047)	(0.035)	(0.048)	(0.035)	(0.045)
constant	0.211***	0.043*	0.144***	0.012	0.149***	0.012	0.144***	0.012
	(0.038)	(0.022)	(0.047)	(0.021)	(0.047)	(0.024)	(0.050)	(0.020)
R^2	0.283	0.000	0.510	0.428	0.512	0.425	0.515	0.402
Observations	73	73	73	73	71	71	71	71
population weighted?	y	n	y	n	y	n	y	n
no outliers in RTI4 share	n	n	n	n	y	y	n	n
no outliers with employment share change	n	n	n	n	n	n	y	y

Notes: RTI4 jobs are the 4 most routine or offshorable occupations with the highest RTI or OFF-GMS indexes (CS 48, 54, 62 and 67). Robust standard errors. 11 cities > 0.5m and 62 cities between 50,000-100,000 inhabitants. ***, **, and * denote statistical significance at the 1 %, 5 %, and 10 % levels.

Table 18 – Employment of most offshorable jobs' share changes 1994-2015 and exposure to 6 most offshorable occupations (OFF6) in 1994

Employment share change of OFF6 jobs								
employment share of OFF6 in 1994	-0.061	-0.312***	-0.140	-0.340***	-0.124	-0.335***	-0.107	-0.259***
	(0.120)	(0.076)	(0.107)	(0.079)	(0.112)	(0.086)	(0.107)	(0.060)
large X employment share of OFF6 in 1994			-0.075***	-0.060***	-0.077***	-0.063***	-0.071***	-0.053***
			(0.019)	(0.019)	(0.019)	(0.020)	(0.018)	(0.017)
constant	-0.140**	0.027	-0.068	0.047	-0.076	0.046	-0.087	0.000
	(0.069)	(0.044)	(0.062)	(0.046)	(0.065)	(0.050)	(0.061)	(0.033)
R^2	0.0127	0.2905	0.2678	0.3662	0.2692	0.3381	0.264	0.3049
Observations	73	73	73	73	71	71	71	71
population weighted?	y	n	y	n	y	n	y	n
no outliers in RTI4 share	n	n	n	n	y	y	n	n
no outliers with employment share change	n	n	n	n	n	n	y	y

Notes: The category of 6 most offshorable occupations (OFF6) are those with highest OFF-GMS index (CS 46, 48, 54, 62, 65, 67). Robust standard errors. 11 cities > 0.5m and 62 cities between 50,000-100,000 inhabitants. ***, **, and * denote statistical significance at the 1 %, 5 %, and 10 % levels.

Table 19 – 2-digit PCS categories and representative 4-digit PCS categories (private sector)

23	**CEOs of firms above 9 employees**
35	**Creative professionals**
	Journalists and writers
	Media, publishing houses and performing arts managers
	Artists
37	**Top managers, professionals and liberal professions (PCS 31)***
	Managers of large businesses
	Finance, accounting, sales, and advertising managers
	Other administrative managers
	Doctors and pharmacists
	Legal and technical liberal professions (lawyers, architects)
38	**Technical managers and engineers**
	Technical managers for large companies
	Engineers and R&D managers
	Eletrical, mechanical, materials and chemical engineers
	Purchasing, planning, quality control, and production managers
	Information technology R&D engineers and managers
	Information technology support engineers and managers
	Telecommunications engineers and specialists
43	**Mid-level health professionals and social workers**
	Nurses
	Masseurs and therapists
	Medical technicians
	Specialized educators
	Leisure and cultural activity organizers
46	**Mid-level professionals**
	Mid-level professionals, various industries
	Supervisors in financial, legal, and other services
	Store, hotel, and food service managers
	Sales and PR representatives
47	**Technicians**
	Designers of electrical, electronic, and mechanical equipment
	R&D technicians, general and IT
	Installation and maintenance of non-IT equipment
	Installation and maintenance of IT equipment
	Telecommunications and computer network technicians
	Computer operation, installation and maintenance technicians
48	**Foremen, Supervisors**
	Foremen: construction and other
	Supervisors: various manufacturing sectors
	Supervisors: maintenance and installation of machinery
	Warehouse and shipping managers
	Food service supervisors Other
53	**Security workers**
	Guards, bodyguards
54	**Office workers**
	Receptionists, secretaries
	Administrative/clerical workers, various sectors
	Computer operators
	Bus/train conductors, etc
55	**Retail workers**
	Retail employees, various establishments
	Cashiers
	Service station attendants
56	**Personal service workers**
	Restaurant servers, food prep workers
	Hotel employees: front desk, cleaning, other
	Barbers, hair stylists, and beauty shop employees
	Child care providers, home health aids
	Residential building janitors, caretakers
62	**Skilled industrial workers**
	Skilled construction workers
	Skilled metalworkers, pipe?tters, welders
	Skilled heavy and electrical machinery operators
	Skilled operators of electrical and electronic equipment
	Skilled workers in various industries
63	**Skilled manual laborers**
	Gardeners
	Master electricians, bricklayers, carpenters, etc
	Skilled electrical and electronice service technicians
	Skilled autobody and autorepair workers
	Master cooks, bakers, butchers
	Skilled artisans (jewelers, potters, etc)
64	**Drivers**
	Truck, taxi, and delivery drivers
65	**Skilled transport workers**
	Heavy crane and vehicle operators
	Warehouse truck and forklift drivers
	Other skilled warehouse workers
67	**Low skill industrial workers**
	Low skill construction workers
	low skill electrical, metalworking, and mechanical workers
	low skill shipping, moving, and warehouse workers
	low skill transport industry workers
	Low skill production workers in various industries
68	**Low skill manual laborers**
	Low skill mechanics, locksmiths, etc
	Apprentice bakers, butchers
	Building cleaners, street cleaners, sanitation workers
	Various low skill manual laborers

Notes: Translation of categories other than PCS 23, 35, 43 and 53 taken from Table 2 of Harrigan et al. (2016). (*): The PCS 31 – liberal professions category was created after the 2003 revision. Since we work with data from earlier years, we merge 37 and 31 together.

Table 20 – Basic statistics by 2 digit CS categories: Full table.

CS 2-digit	description	employment share in %		average city wage (in 2015 euros)		Routine (index values)	Offshorable (index values)
		1994	2014	1994	2015		
		high-paying occupations					
23	CEOs	1.0	0.9	42.82	59.20	-0.75	-0.59
37	**managers and professionals**	**6.2**	**10.2**	32.52	38.56	-0.75	-0.59
38	**engineers**	**5.1**	**9.0**	30.36	33.68	-0.82	-0.39
35	creative professionals	0.5	0.5	22.82	31.80	-0.72	-0.49
		medium-paying occupations					
48	**supervisors and foremen**	**4.1**	**2.7**	18.03	21.86	0.42	1.23
46	**mid-level professionals**	**12.3**	**7.6**	17.54	21.20	-0.48	-0.16
47	**technicians**	**5.7**	**6.3**	17.15	20.60	-0.40	-0.29
43	mid-level health professionals	0.8	1.5	15.05	18.05	-0.35	-0.57
62	**skilled industrial workers**	**14.1**	**9.3**	13.52	17.99	0.38	1.24
54	**office workers**	**11.8**	**11.2**	13.17	16.98	2.03	0.87
65	**transport and logistics personnel**	**2.9**	**3.0**	11.96	16.00	0.33	0.27
63	**skilled manual workers**	**8.0**	**8.3**	11.90	15.50	0.17	-0.33
64	**drivers**	**5.0**	**5.5**	11.50	14.46	-1.50	-0.63
67	**unskilled industrial workers**	**10.9**	**5.7**	11.02	14.72	0.45	2.09
		low-paying occupations					
53	security workers	0.7	1.4	10.60	14.60	-0.28	-0.51
55	**sales-related occupations**	**5.4**	**8.3**	10.44	13.74	0.30	-0.57
56	personal service workers	2.2	4.8	9.97	12.63	-0.43	-0.57
68	**unskilled manual workers**	**3.3**	**3.8**	9.11	13.27	0.06	-0.36

Notes: In-sample values. Employment share for metropolitan (mainland) France. Average city wages in constant 2015 euros. Categories in bold are those with employment shares above 2.5% in 1994 in sample.

D Pre-1994 Labor market developments

In this Appendix we document the relevant changes in the French labor market before the coverage of the detailed DADS data (starting in 1994).

Labor market polarization in France might have begun earlier than 1994: although the modern ICT were not widely used pre-1994 there was a significant advance in automation in manufacturing through CAD/CAM systems, early adoption of basic computer text editors or spreadsheets. There was also an increase in offshoring possibilities for French companies with such developments as the Spanish or Portugese accession to the EEC in 1986 or the opening up of Eastern European countries in 1989. The strength of the automation or offshoring shocks is unclear, however, and the most offshorable occupations (CS 48, 62 and 67) related to manufacturing might have been affected the earliest. It is therefore instructive to detail some of the pre-1994 developments.

There are two data sources that allow to track occupations at the 2-digit PCS level back to 1982 when the PCS classification was introduced: the French Labor Market Survey (yearly data) and the Census (1982, 1990 and 1999). The publicly available Labor Market Survey gives data at the

department but not at the commune level, hence it is impossible to precisely characterize city-level labor markets. The Census, on the other hand, gives the commune location of respondents but does not give data about hours worked or wages. We use the Census as we are interested in the shares of employment in cities, but in contrast to data presented in main text the patterns will refer to shares of people employed and not actual hours worked. We use the publicly available individual data for the 1982, 1990 and 1999 censuses (covering 1/4th for 1982 and 1990, and 1/20th for 1999 of the entire population respectively). The publicly available census data for subsequent years is presented in such a way that misses important details (either in terms of communes or details of occupations) disallowing the same exercise.

The 1982-1999 counterparts to Tables 3-6 using Census data are as follows:

Table 21 – Share of 4 highest-paying occupations per metropolitan area size, Census data 1982-1999.

Agglo.size	Paris	> 0.75M	0.5-0.75M	0.2-0.5M	0.1-0.2M	0.05-0.1M
1982	0.18	0.12	0.10	0.10	0.09	0.09
1990	0.24	0.16	0.13	0.12	0.11	0.11
1999	0.26	0.17	0.14	0.13	0.12	0.11
change in ppct 1982-1990	0.06	0.04	0.03	0.03	0.02	0.02
change in ppct 1990-1999	0.02	0.01	0.01	0.00	0.00	0.00
change in ppct 1982-1999	0.08	0.05	0.04	0.03	0.03	0.02

Table 22 – Share of 10 middle-paying occupations per metropolitan area size, Census data 1982-1999.

Agglo.size	Paris	> 0.75M	0.5-0.75M	0.2-0.5M	0.1-0.2M	0.05-0.1M
1982	0.66	0.71	0.74	0.73	0.74	0.73
1990	0.61	0.68	0.69	0.70	0.71	0.71
1999	0.56	0.64	0.65	0.66	0.67	0.68
change in ppct 1982-1990	-0.05	-0.03	-0.04	-0.04	-0.04	-0.02
change in ppct 1990-1999	-0.05	-0.04	-0.04	-0.03	-0.04	-0.03
change in ppct 1982-1999	-0.10	-0.07	-0.08	-0.07	-0.07	-0.06

The conclusions from this exercise are as follows. First of all, exposure to the most routine and offshorable jobs is much higher in 1982 for large cities above 0.5m inhabitants than in 1994 in the DADS data, and the discrepancies in terms of shares of high- middle- and RTI4 jobs across city sizes are lower. Employment shares of the 4 most routine and offshorable occupations and, more generally, middle-paying jobs indeed decline faster in larger cities whether in 1982-1990 or in the entire 1982-1999 period. The labor market polarization across cities manifests itself as our theory predicts: high paying jobs' shares increase most in largest cities, as found in the exhaustive

Table 23 – Share of 4 lowest-paying occupations per metropolitan area size, Census data 1982-1999.

Agglo.size	Paris	> 0.75M	0.5-0.75M	0.2-0.5M	0.1-0.2M	0.05-0.1M
1982	0.15	0.17	0.16	0.17	0.17	0.18
1990	0.15	0.16	0.18	0.18	0.18	0.18
1999	0.18	0.19	0.21	0.21	0.21	0.22
change in ppct 1982-1990	0.00	0.00	0.01	0.01	0.01	0.01
change in ppct 1990-1999	0.02	0.02	0.03	0.03	0.03	0.03
change in ppct 1982-1999	0.02	0.02	0.04	0.04	0.05	0.04

Table 24 – Share of the 4 most routine and offshorable occupations (CS 48, 54, 62 and 67) per metropolitan area size, Census data 1982-1999.

Agglo.size	Paris	> 0.75M	0.5-0.75M	0.2-0.5M	0.1-0.2M	0.05-0.1M
1982	0.36	0.39	0.43	0.42	0.43	0.43
1990	0.28	0.33	0.34	0.35	0.37	0.38
1999	0.22	0.27	0.29	0.30	0.31	0.33
change in ppct 1982-1990	-0.07	-0.05	-0.08	-0.06	-0.06	-0.05
change in ppct 1990-1999	-0.06	-0.06	-0.06	-0.06	-0.06	-0.06
change in ppct 1982-1999	-0.13	-0.12	-0.14	-0.12	-0.13	-0.11

DADS data for 1994-2015. Low-paying jobs do not increase at all in largest cities in 1982-1990 and increase less in terms of percentage points over 1990-1999 and the entire 1982-1999 period.

Similar patterns obtain for 1982-1994 using the Labor Market Survey data while classifying departments by largest city.

For individual occupational categories, the routine/offshorable job categories whose employment declines most in the studied years 1994-2015 in the DADS data in large cities are in particular PCS 46 and 54 (mid-level professionals and office workers respectively), whereas it is 62 and 67 (skilled and unskilled industrial workers respectively) for small cities (cf. the patterns in Figure 5)[22]. A part of the answer of such a differential evolution may lay in the fact that large cities had different shares of these jobs at the beginning of the 1990s (see Table 25) than small cities, and such a discrepancy existed already in 1982. In particular, the share of mid-level professionals and office workers in employment was higher than that of industrial workers in 1982 in the largest cities above .75m inhabitants while the opposite is true for smaller cities. This feature of data may be explained by different deindustrialization across time and geographies as shown in Table 26 where the share of industry employment at Census years is given for the period 1968-2015. Already over the period 1968-1982 large cities experienced faster deindustrialization than small cities. Reports from research bodies as the INSEE or DATAR (Délégation interministérielle à l'aménagement du territoire et à

[22]The PCS 46 category contains heterogenous professions that were differentially impacted by automation/offshoring. For example, occupations such as drafters, secretaries, photographers, sales in insurance, real estate, finance or advertising included in this category have RTI scores above 2; some of them are also very offshorable.

l'attractivité régionale) indicate the following. Internal offshoring of manufacturing tasks within France might have played a part due to the reduction in internal transport costs (both because of highway and railway construction), environmental regulations to keep polluting industries out of high density areas or a deliberate government policy to decentralize economic activity across France (e.g. moving public engineering schools outside Paris). For Ile-de-France, deindustrialization was largely due to the reorganization of the automobile (that moved out of large cities) and defense industries (idem, with aerospace moving to Toulouse in particular). To an unknown extent firm reorganization and shifting tasks outside the boundaries of firms (e.g. legal services, general and administrative or cleaning premises) that cannot be precisely measured was responsible for the fall in manufacturing value added overall. This, together with moving tasks within multi-establishment firms might have caused some of the tasks to be offshored within France from large to smaller cities.

Table 25 – Employment share of selected middle-paying occupations per metropolitan area size, Census data 1982-1999.

Agglo. Size		Paris	> 0.75M	0.5-0.75M	0.2-0.5M	0.1-0.2M	0.05-0.1M
mid-level professionals (CS 46) and office workers (CS 54)	1982	0.27	0.23	0.19	0.20	0.19	0.19
	1990	0.26	0.23	0.21	0.22	0.21	0.21
	1999	0.26	0.25	0.21	0.22	0.21	0.21
skilled (CS 62) and unskilled industrial workers (CS 67)	1982	0.14	0.19	0.25	0.24	0.25	0.25
	1990	0.10	0.16	0.18	0.19	0.20	0.22
	1999	0.07	0.11	0.14	0.15	0.16	0.18

Notes: Data from the publicly available individual data for the 1982, 1990 and 1999 censuses (covering 1/4th for 1982 and 1990, and 1/20th for 1999 of the entire population respectively). Aggregation to the city-level (based on unites urbaines as of 2015). Includes individuals between 25-64 of age.

Table 26 – Employment share in industry per metropolitan area size, Census data 1968-2015.

share of industry	Paris	>0.75m	0.5-0.75m	0.2-0.5m	0.1-0.2m	0.05-0.1m
1968	0.32	0.30	0.39	0.32	0.34	0.29
1975	0.29	0.28	0.34	0.30	0.34	0.30
1982	0.24	0.23	0.29	0.26	0.30	0.27
1990	0.19	0.20	0.22	0.22	0.24	0.24
1999	0.13	0.15	0.17	0.17	0.19	0.19
2015	0.07	0.11	0.11	0.11	0.13	0.13
ppct chng 68-82	-0.08	-0.07	-0.10	-0.06	-0.04	-0.02
ppct chng 68-90	-0.13	-0.11	-0.16	-0.10	-0.10	-0.05

Notes: Raw, exhaustive data from Censuses were prepared by the INSEE for each commune. Aggregation to the city-level (based on unites urbaines as of 2015) done by authors. Includes individuals between 25-54 of age.