

Inflation dynamics, marginal cost, and the output gap: Evidence from three countries

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Abstract

Recent studies by Galí and Gertler (1999), Galí, Gertler, and López-Salido (2001), and Sbordone (2001, 2002) have argued that the New Keynesian Phillips curve (Calvo pricing model) is empirically valid, provided that real marginal cost rather than detrended output is used as the variable driving inflation. One interpretation of these results is that real marginal cost is not closely related to the output gap, and so models for monetary policy need to include labor market rigidities. An alternative interpretation is that marginal cost and the output gap are closely related, but that the latter needs to be measured in a manner consistent with dynamic general equilibrium models. To date, there has been little econometric investigation of this alternative interpretation. This paper provides estimates of the New Keynesian Phillips curve for the United States, the United Kingdom, and Australia using theory-based estimates of the output gap. Using theory to measure the output gap leads to a considerable improvement in the empirical performance of output-gap-based Phillips curves.

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1 Introduction

Recent contributions by Galí and Gertler (1999), Galí, Gertler, and López-Salido (2001, 2002), and Sbordone (2001, 2002) have provided empirical support for what Roberts (1995) termed the “New Keynesian Phillips curve” (NKPC). These are encouraging findings for the use of dynamic general equilibrium models in monetary policy analysis, as they suggest that the observed dynamics of inflation can be understood with models derived from microeconomic foundations. In particular, as shown in Rotemberg (1987) and Roberts (1995), the forward-looking dynamics that underlie the New Keynesian Phillips curve emerge from optimal firm responses to obstacles to adjusting prices of the type introduced by Rotemberg (1982) and Calvo (1983). Previous work, e.g. Fuhrer (1997), had suggested that the NKPC was highly counterfactual, and that backward-looking models were required to understand empirical inflation dynamics. The recent studies suggest, however, that the NKPC does work empirically if marginal cost is used as the process driving inflation, instead of a measure of output relative to trend.

At the same time, the implications for macroeconomic modeling of recent results on the NKPC are unclear. A crucial issue is how to interpret the empirical success of NKPCs that use marginal cost as the driving process and the failure of those that use detrended output. One possibility, raised by Galí and Gertler (1999) for example, is that the results imply that the relationship between real marginal cost and the output gap is weak.¹ According to New Keynesian models, a simple structural relationship between inflation and the output gap does not hold in general—it holds only if the labor market is perfectly competitive. If the labor market is not competitive, labor frictions become crucial. The minimal number of endogenous variables needed in a realistic monetary policy analysis is then five: inflation, output, nominal interest rates, real marginal cost, and labor input (see Erceg, Henderson, and Levin (EHL), 2000, Table 1).² In particular, one then needs to model the “wage markup” produced by monopoly power in labor supply, which drives a wedge between real marginal cost and the output gap (see GGL, 2001). We will refer to this interpretation of the NKPC findings as the “wage markup” interpretation.

¹ For example, Galí and Gertler (1999, p. 204) state that a “fundamental issue, we believe, is that even if the output gap were observable the conditions under which it corresponds to marginal cost may not be satisfied.”

² Of course, by substitution, one could reduce the number of variables in the analysis. But this may not be a straightforward procedure if there are several types of imperfection in the labor market.

An alternative interpretation, recognized but not endorsed in the above papers, is that the poor performance of detrended output-based NKPCs is not evidence against *output-gap*-based NKPCs; rather, it is evidence of difficulties in measuring the output gap. Under this interpretation, real marginal cost has a closer relationship to the true output gap than do traditional, trend- or filter-derived measures of the latter. Real shocks produce fluctuations in the natural level of output, which is therefore not well approximated as smooth. This “output gap proxy” interpretation of marginal-cost-based NKPCs is endorsed, and evidence in its favor provided, by Galí (2002), Neiss and Nelson (2001), and Woodford (2001a). This interpretation implies that monetary policy analysis can be validly conducted using compact systems consisting of three variables: inflation, output, and nominal interest rates.³

Distinguishing between the “wage markup” and “output gap proxy” interpretations of the NKPC is important not only for choosing the appropriate model for monetary policy, but also the appropriate targets of policy. EHL’s (2000) analysis suggests that price inflation targeting is suboptimal in hybrid sticky price/sticky wage models; rather, central banks should target a mixture of price and nominal wage inflation. If, however, the “output gap proxy” interpretation of NKPCs is valid, then the goods market is the only source of nominal rigidity in the economy that is distorting outcomes for real variables, and price inflation targeting is optimal (Goodfriend and King, 2001; Woodford, 2001b).

In this paper, we provide evidence that the output gap proxy interpretation of the NKPC deserves reconsideration. Using data for three countries—the United States, the United Kingdom, and Australia—we find that output-gap-based NKPCs deliver correctly signed and interpretable estimates, and are competitive in fit with cost-based NKPCs, provided the potential GDP series is derived in a manner consistent with theory. To date, direct estimation of output-gap-based Phillips curves with the gap measured in a theory-consistent manner has been impeded by technical obstacles to defining and measuring potential GDP in the realistic case when potential is partly determined by endogenous state variables (such as the capital stock). The algorithm used in Neiss and Nelson (2001) overcomes these obstacles and enables us to generate a theory-consistent gap series.

We note two other criticisms of sticky price models addressed in this paper. First, EHL (2000, p. 298) argue that models with only nominal price rigidities cannot rationalize a disturbance (or “cost-push shock”) term in empirical NKPCs. We will argue that such a shock term can be rationalized even in the absence of labor market

³ In models with endogenous investment, consumption and the capital stock would be added to this list.

rigidities, once one interprets it as a “price-level shock.” Second, Christiano, Eichenbaum, and Evans (2001) argue that labor market rigidities are required to explain the sluggish response of inflation to a monetary policy shock. We contend that this sluggish response can be largely accounted for, provided that the reaction of the output gap is protracted. This sluggishness, in turn, may arise largely from intrinsic dynamics in output behavior (e.g. from habit formation and capital adjustment costs).

Our paper offers an alternative interpretation of the NKPC literature’s results, but our overall results strongly support the general approach to modeling inflation of GG, GGL, and Sbordone. The empirical work establishes that their approach—modeling inflation using labor cost data and explicit optimizing analysis—can be usefully extended to other countries, extended sample periods, and to CPI inflation. Many of our results provide evidence against claims, such as in Rudd and Whelan (2001), that GG’s results are fragile or consistent with backward-looking inflation dynamics.

Our paper is organized as follows. Section 2 discusses recent empirical work on the NKPC, focussing on the contrast between “wage markup” and “output gap proxy” interpretations of this work. Section 3 gives a model of potential GDP determination used for our empirical measurement of output gaps. Section 4 describes our data and specification. In Section 5 we estimate the NKPC for the United States, the United Kingdom, and Australia, using our model-consistent output gap series. Section 6 concludes.

2 Existing estimates of the New Keynesian Phillips curve: a summary and reinterpretation

In this section we provide a brief discussion of output-gap-based and marginal-cost-based New Keynesian Phillips curves (Section 2.1); and offer an interpretation that is more favorable to the output-gap-based NKPC than the existing literature (Section 2.2).

2.1 Empirical work on the NKPC: costs vs. gaps

The basic idea behind the New Keynesian Phillips curve (NKPC) is that the profit-maximizing response of firms to obstacles to adjusting prices, is to solve dynamic optimization problems. The first order conditions for optimization then imply that expected future market conditions matter for today’s pricing decisions. In aggregate, and combined with the assumption of competitive factor markets, this implies the following NKPC describing the behavior of annualized quarterly inflation (π_t^a):

$$\pi_t^a = \beta E_t \pi_{t+1}^a + \lambda_y (y_t - y_t^*) + u_t, \quad (1)$$

where constant terms are suppressed, y_t is log output, and y_t^* is log potential output. The parameter β corresponds to the discount factor and so should be close to one; the parameter λ_y is a function of firm structure and price adjustment costs, and satisfies $\lambda_y > 0$. See e.g. Roberts (1995), Walsh (1998), or Sbordone (2002), for derivations. The disturbance term u_t has been labelled a “cost-push” shock, although (as discussed below) we prefer the term “price-level shock.” As we discuss in Section 4.4, various rationalizations of the u_t term have been advanced, but for the moment we simply assume that it is exogenous in the sense that it is not proxying for omitted dynamics or excluded endogenous variables.

There have been numerous problems with empirical estimates of equation (1). While Roberts (1995) did find a positive and significant value of λ_y on annual U.S. data, analogous tests on quarterly data have been less favorable. GGL (2001, p. 1251) find that instrumental variables estimates of (1) on quarterly U.S. and euro area data deliver coefficients on $E_t \pi_{t+1}$ near 1 (in keeping with the theory), but *negative* coefficients on the “output gap” term. Indeed, they find this coefficient is *significantly* negative ($t = 3.5$) for the U.S. Using quarterly U.S. data, Fuhrer (1997) finds that when both lags and leads of inflation are included in (1) instead of the single lead of π_t , the coefficient sum on the leads of π_t is near zero, again apparently rejecting the NKPC.⁴ Estrella and Fuhrer (1999) estimate (1) on quarterly U.S. data for 1966–1997; they do find a positive estimate of λ_y , but it is highly insignificant ($t = 0.4$) and they emphasize the NKPC’s poor fit relative to backward-looking specifications. Importantly, as we discuss in Section 2.2, these papers uniformly use (log) GDP relative to a trend or filter as the empirical measure of the output gap, $(y_t - y_t^*)$. For example, Roberts (1995) and GG use quadratically detrended log GDP, while Fuhrer (1997) uses deviations of log GDP from a broken-linear trend to measure the output gap, and finds similar results when potential output is modeled as following linear, quadratic, or spline-based trends. Estrella and Fuhrer (1999) use the Congressional Budget Office (CBO) “output gap” series, which closely resembles broken-trend-based output-gap measures.

While some have interpreted the above results as rejections of the NKPC, recent work has suggested that forward-looking price-setting dynamics may be empirically important, but that equation (1) is too restrictive a representation of this behavior.

⁴ Rudebusch (2002), using survey data for expectations, reports a somewhat higher and more significant weight on expected future inflation (0.29).

The firm optimization problem underlying the NKPC leads to the output-gap-based Phillips curve (1) only under additional conditions, notably the assumption of flexible wages. These assumptions imply that the output gap and log real marginal cost are perfectly correlated. In the more general case, equation (1) may not hold, but the firm’s optimality condition, relating its pricing decisions to the stream of current and expected future marginal costs, would continue to hold. Under conditions discussed in Sbordone (2002), this first order condition implies in aggregate the following marginal-cost-based NKPC:

$$\pi_t^a = \beta E_t \pi_{t+1}^a + \lambda mc_t + u_t \quad (2)$$

where mc_t is the log of real marginal cost, and $\lambda > 0$. Estimates of this equation have been far more satisfactory than equation (1). For example, Galí and Gertler (GG) (1999, p. 207) obtain on 1960–1997 quarterly U.S. data an estimate of β of 0.942, and a coefficient of 0.092 ($t = 1.9$) on mc_t ; and GGL (2001, p. 1250) find on 1970–1998 quarterly euro area data estimates of $\beta = 0.91$ and $\lambda = 0.352$ ($t = 2.1$). These papers use instrumental variables estimates; alternative estimation techniques in Sbordone (2001, 2002) also support the marginal-cost-based NKPC. In light of these results, GGL (2001) argue for models with forward-looking price setting combined with labor rigidities, which imply that equation (2) holds but the stricter, output-gap-based NKPC (1) does not.⁵

On the surface, then, the empirical evidence seems consistent with Sbordone’s (2001, p. 6) characterization that “inflation dynamics is well explained when real marginal cost is approximated by unit labor costs, but is not well modeled when marginal cost is approximated by output gap.” Yet we will argue instead that the output-gap-based Phillips curve (1) is a reasonable approximation in modeling inflation, and is not greatly inferior to equation (2) in its empirical performance. We now provide a reinterpretation of the existing evidence that is the basis for our argument.

2.2 The NKPC with the output gap reconsidered

In questioning the implications of the results reported, we focus on the fact that dynamic stochastic general equilibrium (DSGE) models define “potential output” or “the natural level of output” differently from that typically used in empirical work, and in a way that is explicitly related to the underlying real shocks in the economy. Specifically, the natural level of output in a DSGE model corresponds to the output level that would prevail if there were no nominal rigidities in the economy, i.e. if

⁵ Recent work on wage rigidities in optimizing models includes Jeanne (1998), EHL (2000), Ravenna (2000), Christiano, Eichenbaum and Evans (2001), Sbordone (2001), and Smets and Wouters (2001).

prices and wages were fully flexible.⁶ This definition implies two properties that distinguish it from standard gap measures, including those used by GGL, Roberts, and others in their tests of the output-gap-based Phillips curve.

First, potential output is affected by real shocks over the business cycle, and so does not follow a smooth trend as implied for example by detrended or filtered output-based measures of the output gap. Detrended potential output is not constant over the cycle in DSGE models. As Amato and Laubach (2002) put it, “In non-optimizing models, the output gap is constructed, both conceptually and empirically, as deviations of output from a smooth trend, whereas, in optimizing models, the notion of potential output is different... [and] in general could be very *volatile*.” While the NKPC specification itself does arise from an optimizing model, empirical work on the NKPC has not always recognized the nature of potential output in optimizing models. Rather, as Sbordone (2001, p. 6) notes, “empirical estimates of the NKPC curve usually approximate potential output Y^p_t by some deterministic function of time...”

Second, the issue does not boil down to disputes on how to detrend output (e.g. HP filtering vs. linear detrending).⁷ Rather, the issue is what phenomenon the output gap is supposed to capture. In DSGE models, the output gap captures that portion of the movement in output that can be attributed solely to the existence of nominal rigidities in the economy. Interpreted in this way, the output gap is *not* a measure of the business cycle. For example, output may respond cyclically to real demand and supply shocks, but if prices are flexible, the output gap is zero, even though there is business-cycle-frequency variation in output. Similarly, there are serious limitations to judging the plausibility of an output gap series by whether it becomes negative during recessions. A recession (negative growth in actual output) is consistent with the output gap being positive, negative, or zero, depending on whether the source of the economic downturn is nominal or real, and on where output is in relation to potential at the beginning of the recession.

⁶ For example, Goodfriend and King (1997, p. 261) define the output gap as the percentage difference between output and “the flexible price level of output, i.e. that obtained in a noncompetitive RBC model.” Clarida, Galí, and Gertler (2000, p. 170) define the “natural rate” of output as “the level of output that would obtain under perfectly flexible prices.” McCallum and Nelson (1999, p. 23) define “capacity output” as the level of output that would prevail “if there were no nominal frictions.” Woodford (2001b, p. 14) defines “the *natural* rate of output” as “the equilibrium level of output under price flexibility.” An early discussion is that by Obstfeld and Rogoff (1984), who define excess demand as the percentage difference between output and “flex-price equilibrium output... the equilibrium that would obtain if all prices were fully flexible” (pp. 160, 163).

⁷ Thus Fuhrer’s (1997) check that his rejections of the NKPC are robust to alternative techniques for detrending output would not resolve the problem of measuring potential discussed here.

We found in simulations of quantitative DSGE models in Neiss and Nelson (2001) that the difference in output gap concepts is a far from trivial distinction. Rather, using our procedure (see Section 3.3 below) for obtaining the output gap in a DSGE model with habit and capital formation, and simulating the model under a policy reaction function estimated from the data, we found the output gap and detrended output had a negative correlation (-0.68), and that this result was robust to changing the specification of price adjustment, policy rule, preferences, and production. Cyclical variation in potential output was sizable. Our model was one where the output-gap-based NKPC (1) held near-exactly; yet attempts to estimate it would produce unfavorable results, if they were to follow the standard practice of measuring the gap by detrended output. On the other hand, equation (2) also holds in that model, and attempts to estimate it would be successful if real marginal cost was measured reliably. Thus an “output gap proxy” interpretation of the existing work on cost-based NKPCs is that their relative success reflects not the failure of the output-gap-based NKPC to hold in the data, but instead the fact that marginal cost is a better index of output gap fluctuations than is detrended output.

GG (1999, p. 204) express some sympathy with output-gap measurement problems as an explanation for the poor performance of detrended-output NKPCs and better performance of unit labor cost NKPCs, but ultimately conclude that it is “problematic” whether “correcting for [output gap] measurement error alone” could account for the differences. Similar reservations are expressed in GGL (2001), who suggest that unit labor costs and the output gap may not be closely related in practice. If valid, this would be grounds for introducing model features that break the relation between the gap and unit labor costs (and so, between the gap and inflation). But we contend that, because existing estimates of output-gap-based NKPCs do not use a potential GDP series consistent with theory, the invalidity of the original NKPC (1) has not been firmly established.

As discussed in (e.g.) EHL (2000) and GGL (2001, Section 5), the output-gap-based Phillips curve emerges by substitution only in the special case of price stickiness, plus a competitive and flexible-wage labor market. The marginal-cost formulation of the Phillips curve is the more general relationship, and this is stressed by GGL (2001) and Gagnon and Khan (2001) as an advantage of estimating Phillips curves of this type.

To see the issue at stake, we follow GGL (2001, p. 1261) in defining the (log) wage markup $\log(\mu_t^w)$ as the percentage markup of the real wage over the ratio of marginal

utilities of leisure and consumption that would normally determine what a worker would accept as payment for its labor in a competitive labor market:

$$w_t - p_t = \log(u_n/u_c) + \log(\mu_t^w) \quad (3)$$

where $w_t - p_t$ is the log real wage, and u_n and u_c denote the household's marginal (dis)utility from labor supply and consumption respectively in period t . With log real marginal cost given by $mc_t = w_t - p_t - (y_t - n_t)$ (i.e., log real unit labor cost),⁸ the forcing process in the cost-based NKPC (2) is increasing in the wage markup:

$$\pi_t^a = \beta E_t \pi_{t+1}^a + \lambda [\log(\mu_t^w) + \log(u_n/u_c) - (y_t - n_t)], \quad (4)$$

so inflation is increasing in the degree of monopoly power exerted by workers. If this term varies widely from period to period, then the output gap will not do well in describing inflation dynamics. On the other hand, if wage-markup variations are not a major source of fluctuations in inflation, or if the wage markup is itself closely related to excess-demand conditions, then the output gap should be sufficient to model inflation dynamics. The rest of this paper establishes whether this “output gap proxy” interpretation of cost-based NKPCs has merit. To begin this investigation, we introduce a model of potential output.

3 A model of potential output

In this section we describe a dynamic stochastic general equilibrium model, based on Neiss and Nelson (2001), that we use to obtain potential output series for the estimation of output-gap-based Phillips curves in our empirical work in Section 5.

3.1 Model equilibrium conditions

We concentrate on the household side of the model, since that is from where most of the conditions that define potential output emerge. The representative household has a utility function of the form $E_0 \sum_{t=0}^{\infty} \beta^t u(C_t, C_{t-1}, 1-N_t)$, where C_t is its period t consumption of a Dixit-Stiglitz aggregate and N_t is the fraction of time worked in period t . Utility from money holding can be neglected in computing potential GDP. The period utility function we use is of the form $u(\bullet) = \lambda_t^c (\sigma/\sigma-1) (C_t/C_{t-1})^{(\sigma-1)/\sigma} + b(1-N_t)$, $\sigma \in (0,1)$, $b > 0$. Parameter $h \in [0,1)$ indexes the degree of habit formation.

⁸ The entire discussion in this paper will be in the context of models where unit labor cost and marginal cost coincide. For an examination of the empirical importance of relaxing this assumption, see Gagnon and Khan (2001) and Sbordone (2002).

Period utility has a multiplicative disturbance, λ_t^c , to consumption preferences—an “IS” or “real demand” shock, whose effects we discuss below.

For our purposes, the key terms in the period t household budget constraint are:

$$C_t + (B_{t+1}/P_t) + X_t = w_t^r N_t + (1 + R_{t-1})B_t - \phi X_{t+1}^2, \quad (5)$$

where w_t^r is the real wage in period t , and B_t is the quantity of short-term securities (redeemed for $1+R_{t-1}$ units of period t output) carried over from $t-1$ (so R_t is the short-term nominal interest rate in period t). The variable X_t , “quasi-investment,” is related to the household’s capital stock K_t by $X_t = K_{t+1} - (1-\delta)K_t$, $\delta \in [0,1)$. If there were no capital adjustment costs, X_t would correspond to investment expenditure. The size of capital adjustment costs is determined by the parameter $\phi \geq 0$.

After substituting in the first-order condition for consumption and the Fisher relation $(1+R_t) = E_t(1+r_t)(1+[P_{t+1}/P_t])$, three of the key first-order conditions used in our computation of potential output are those for consumption, labor supply and bond holding:

$$\psi_t = \lambda_t (C_t/C_{t-1})^{h(\sigma-1/\sigma)} (1/C_t) - \beta h E_t \lambda_{t+1} (C_{t+1}/C_t)^{h(\sigma-1/\sigma)} (1/C_t) \quad (6)$$

$$b = w_t^r \psi_t \quad (7)$$

$$\psi_t = \beta (1+r_t) E_t \psi_{t+1}, \quad (8)$$

where ψ_t is the Lagrange multiplier on the period t wealth constraint. An optimality condition for capital accumulation will also be relevant in determining potential.

On the firm side, a generic firm $j \in [0,1]$ faces a Dixit-Stiglitz demand function for its output Y_{jt} . It has a Cobb-Douglas production function $Y_{jt} = A_t N_{jt}^\alpha K_{jt}^{1-\alpha}$, where A_t is a technology shock, and $\alpha \in (0,1)$. In a symmetric equilibrium, the profit-maximizing choice of N_{jt} will lead to the condition

$$\alpha(Y_t/N_t) = \mu_t^G w_t^r, \quad (9)$$

as well as an analogous condition for the optimal rental of capital. Here, $\mu_t^G > 1$ is the gross markup of price over marginal cost, present because of monopolistic competition in the goods market.

3.2 Potential output with no capital or habit formation

As a preliminary step to obtaining an expression for potential output, it is useful to look at potential GDP behavior in a stripped-down version of the above model. This stripped-down version does not include capital or habit formation. Combining labor supply and demand conditions (7) and (9), and loglinearizing the other conditions, the result is the compact system:

$$y_t = E_t y_{t+1} - \sigma r_t + \sigma(1-\rho_\lambda) \lambda_t \quad (10)$$

$$(1-(1/\sigma))y_t - n_t - \mu_t + \lambda_t = 0 \quad (11)$$

$$y_t = a_t + \alpha n_t. \quad (12)$$

Here y_t is log output, n_t log labor input, μ_t the log markup, and λ_t the log of λ_t^c . Logged variables should be regarded as deviations from their steady-state values (and r_t as in units of quarterly fractional deviations from its steady-state value). The “IS shock” λ_t has an AR(1) parameter denoted ρ_λ .

An important property of this small system is that λ_t appears as a shock term not only in the IS equation (10), but also in the labor market equilibrium condition (11). This type of IS shock—a multiplicative stochastic term in the households’ preferences over consumption, which in turn leads to a term in λ_t in the expression for log marginal utility of consumption—is used in the preference specifications of McCallum and Nelson (1999), Amato and Laubach (2001), Ireland (2001), Neiss and Nelson (2001), Woodford (2001b), and others. GGL (2002) carry out an identification scheme to obtain an estimate of the “inefficiency gap” in US data, which they use as the basis for drawing implications about U.S. output-gap behavior. GGL’s identification scheme rests on an assumption of negligible variation in any preference shock term that appears in the labor condition (11). They can then interpret any residual left over from the linear combination of y_t , n_t , and μ_t as a “wage markup” term reflecting the existence of labor market rigidity. But, because it is part of the marginal utility of consumption expression, the preference shock λ_t appears in our labor equilibrium condition even though the household’s preferences over leisure are nonstochastic. GGL’s identification scheme is thus quite restrictive in the sense that it rules out as a special case the preference specification of many existing quantitative sticky-price DSGE models. A corollary of this is that these standard models do not need to rely on labor-market imperfections to justify a stochastic relationship between y_t , n_t , and μ_t .

The flexible-price version of the above system corresponds to the case where y_t , n_t , and r_t are equal to their natural values y_t^* , n_t^* , and r_t^* . In the class of model used here, it is the case that when output is equal to potential, the markup is constant; so $\mu_t = 0$ for all t .⁹ The flexible-price values of real variables can then be obtained from the implied versions of equations (10)–(12), with n_t^* removed by substitution:¹⁰

$$y_t^* = E_t y_{t+1}^* - \sigma r_t^* + \sigma(1-\rho_\lambda) \lambda_t \quad (13)$$

$$(1-(1/\sigma)-(1/\alpha))y_t^* + (1/\alpha)a_t + \lambda_t = 0. \quad (14)$$

The minimum-state-variable solutions for both (the log of) potential output y_t^* and the natural real interest rate can be expressed in terms of productivity and real demand shocks:

$$y_t^* = \pi_1 a_t + \pi_2 \lambda_t \quad (15)$$

$$r_t^* = \pi_3 a_t + \pi_4 \lambda_t, \quad (16)$$

where the π_i coefficients can be obtained by substituting these expressions into (13)–(14) and applying the method of undetermined coefficients. The implied values for the π_i are:

$$\pi_1 = 1/(1-\alpha+[\alpha/\sigma]) > 0, \quad (17)$$

$$\pi_2 = \sigma/(1-\sigma + [\sigma/\alpha]) > 0, \quad (18)$$

$$\pi_3 = -\pi_1([1-\rho_a]/\sigma) < 0, \quad (19)$$

$$\pi_4 = (1-\rho_\lambda)(\pi_2-\sigma)/\sigma > 0. \quad (20)$$

Here, ρ_a denotes the AR(1) parameter driving the productivity shock. These expressions illustrate that potential output and the natural real interest rate in the above model respond to temporary shocks to both productivity and real demand. Productivity shocks raise potential output but reduce the natural interest rate, while real demand shocks raise both potential output and the natural rate.

This simple case highlights several important properties of the model-derived concept of the output gap that carry over to more general environments. First, potential output is a function of both supply and real demand shocks. Describing fluctuations in

⁹ This is because under Dixit-Stiglitz monopolistic competition, the optimal markup of prices over costs is constant under price flexibility. Then real marginal cost, which is the inverse of this markup, is also constant.

¹⁰ Examination of eqs. (11) and (14) verifies that the output gap/real marginal cost relationship stressed by GG (1999) holds in this model. Substituting (12) into (11) and expressing relative to (14), we have: $(1-(1/\sigma)-(1/\alpha))(y_t - y_t^*) - \mu_t = 0$. Since $-\mu_t = mc_t$, the output gap and real marginal cost have a perfect loglinear relationship.

potential output as arising solely from “aggregate supply shocks” should, in general, be regarded only as shorthand.¹¹ Second, since the potential output series defined in equation (15) pertains to the deviation of potential output from trend, it immediately follows that, as noted in Section 2, detrended potential GDP is not constant over the cycle. This renders invalid not only trend-based approaches to measuring the output gap, but also standard “production-function approaches” to measuring the output gap (such as the CBO output gap series in Estrella and Fuhrer, 1999). Production-function approaches fail to address fully the aspects of potential output behavior stressed here. They typically recognize changes in the full-employment labor force due to demographic and labor market developments.¹² But potential GDP can vary cyclically in a dynamic general equilibrium model even if the economy’s demographic characteristics are unchanged. For example, government spending shocks, terms of trade shocks, tax changes, and (as here) taste shocks all affect households’ decisions regarding the timing and magnitude of consumption, and so produce variations in labor supply and potential output. Production-function approaches typically neglect the effect of real shocks on the dynamics of input supplies and so on short-run potential output dynamics.

3.3 Potential output with endogenous state variables

With habit formation ($h > 0$), the previous period’s consumption level becomes a state variable. And with capital formation present, the capital stock also enters the state vector. We denote the flexible price values of lagged log consumption and log capital by c_{t-1}^* and k_t^* . The resulting expression for y_t^* is:

$$y_t^* = \pi_{11}a_t + \pi_{12}\lambda_t + \pi_{13}c_{t-1}^* + \pi_{14}k_t^*, \quad (21)$$

where the π_{ij} are solution coefficients. The solution for potential output can no longer be written as an exact, finite distributed lag of the two real shocks. But since the solutions for c_t^* and k_{t+1}^* are themselves linear functions of the same state vector, repeated substitution allows (21) to be expressed as:

$$y_t^* = f(a_t, a_{t-1}, a_{t-2}, \dots, \lambda_t, \lambda_{t-1}, \lambda_{t-2}, \dots), \quad (22)$$

¹¹ For example, while Rotemberg and Woodford (1997, p. 316), describe their potential output series as consisting of “exogenous supply disturbances,” this labelling is a shorthand to denote all the real shocks, both demand and supply, that affect potential GDP in their model. “Stochastic variability in demands and supplies” was part of Friedman’s (1968) concept of the natural rate of unemployment.

¹² They also may allow for changes in the economy’s potential growth rate, due to the feedback from investment in physical capital to productivity growth.

and model simulations can establish how many lags of a_t and λ_t are required to obtain an accurate finite-lag approximation of the infinite-distributed lag solution given by (22). That is the procedure we use to obtain expressions for potential GDP and the natural interest rate in Neiss and Nelson (2001), and that we use in this paper to obtain empirical, model-based estimates of the output gap.

Note also that in (22) the effects of real shocks on potential GDP are spread over time, as potential output evolves in response to the real shock. The response of output under sticky prices—which we refer to as “actual” output—to real and monetary shocks will also tend to be spread over time, depending on the richness of the model. Since, according to the NKPC, inflation depends on the stream of current and expected output gaps, persistence in the gap stream, due to the dynamics of potential GDP, may help in explaining the observed persistence in inflation.¹³

Galí (2002), Woodford (2001a), and Neiss and Nelson (2001) provide evidence that the output gap and detrended output are not closely related. Galí (2002) and Woodford (2001a) plot real marginal cost against traditional detrended-output-based measures of the output gap, as evidence that, in the words of Galí (2002, p. 2), “the output gap... for the postwar U.S. shows little resemblance with traditional output gap measures.” In Neiss and Nelson (2001) we also argue that the output gap and detrended output are negatively correlated, doing so not on the basis of the marginal cost profile but on evidence from simulations of a sticky-price dynamic general equilibrium model, using equation (22) to generate the potential output and output-gap series under conditions of habit and capital formation. In that model, marginal cost and the output gap have a near-exact relation, but simulations established that detrended output and the output gap were negatively related. Galí’s and Woodford’s evidence is subject to the criticism that it is equally consistent with the “wage markup” interpretation of estimated NKPCs, and hence does not establish that detrended output and the output gap are closely related. Our previous work is subject to the criticism that it is based on simulated data and does not attempt to generate model-consistent output gap series using data from actual economies. We undertake that exercise in this paper, and so provide direct empirical evidence for the “output gap proxy” interpretation of Phillips curves.

¹³ By contrast, GGL (2001, pp. 1242–1243) conclude, “Overall, the output-gap based formulation of the new Phillips curve cannot account for the persistence of inflation either for the U.S. or for the euro area.”

3.4 Completing the model of potential

To complete the model of potential GDP implied by our model, we report the first-order conditions for the households' choices of the capital stock for period $t+1$:

$$\psi_t(1+2\phi X_t) - \beta E_t \psi_{t+1} [(1-\delta)(1+2\phi X_{t+1}) + z_{t+1}] = 0, \quad (23)$$

where z_t is the factor price for capital. An expression for potential output in terms of underlying exogenous shocks, as in equation (22), emerges from solving a log-linear system whose key equations include approximations of equations (6)–(9) and (23).

4 Data and specification

We will estimate equation (1), the output-gap-based NKPC, for three countries. Here we describe our measurement and interpretation of the three variables in equation (1): annualized inflation π_t^a , the output gap ($y_t - y_t^*$), and the shock u_t . We also describe the labor share data that we use to estimate the cost-based NKPC (2).

4.1 Inflation series

While we will report results for GDP deflator inflation for the United Kingdom and the United States, the consumer price index (CPI) is our baseline series for calculation of inflation rates in the three countries in our sample. CPI inflation, unlike GDP deflator inflation, corresponds to the series actually targeted by central banks and so is of greatest concern for monetary policy. It is also of interest, as in GG (1999, p. 206), to see how the NKPC performs in describing “a standard broad measure of inflation.” CPI inflation has been used for this purpose by Roberts (1995, 2001) and Estrella and Fuhrer (1999).

Another important reason for our choice is that the use of GDP deflator rather than CPI inflation in estimating the NKPC may understate the importance of the output gap for aggregate inflation outcomes. Consider a wage-push driven increase in unit labor costs in an open economy. Even if unit labor cost increases are passed through to domestically produced goods prices, a monetary policy that restrains the output gap could lead to offsetting pressure on import prices, provided, as seems likely, import prices are set partly in response to domestic demand conditions. The lack of monetary accommodation translates labor cost changes into relative price changes

rather than aggregate price level changes. A broad aggregate such as the CPI is needed to internalize these substitution effects.

An alternative view is that closed-economy NKPCs should not be used to analyze CPI inflation. On this view, import prices behave dissimilarly from other elements of the CPI and are largely associated with exchange rate and world price changes, leading to a strong influence on the CPI from open-economy factors. But this view does not have much support in the data (see e.g. Stock and Watson, 2001), even for highly open economies like the U.K., and so we proceed with an analysis that does not have explicit open-economy elements.

4.2 Labor share data

For the U.S., following GG, GGL and Sbordone, we use BLS nonfarm business sector data to obtain the labor share. For the U.K., unit labor cost data are an updated version of the labor share data in Batini, Jackson, and Nickell (2000).

For Australia, nonfarm employment and nominal wage data are updated versions of the data used by Gruen, Pagan, and Thompson (1999).¹⁴ The product of these series was expressed as a ratio to nominal GDP and logged to produce a *ulc* series ($ulc = (w - p^{gdp} - (y - n))$, where p^{gdp} is the log GDP deflator).

Data plots and correlations: Figures 1–3 plot labor share against quarterly annualized CPI inflation for the three countries, while contemporaneous correlation matrices for 1961 Q1–2000 Q4 are given in Table 1. The correlations refer to CPI inflation (π_t^a), GDP deflator inflation (π_t^a), log real unit labor costs (labor share) (ulc_t), HP-filtered log output (y_t^{hp}), and quadratically detrended log output (y_t^{quad}). The latter two series are, of course, frequently labelled the “output gap.” Inflation rates are defined as the annualized quarterly percentage change in the relevant price index. Only the CPI inflation rate is used for Australia because the GDP deflator inflation series, especially before 1980, is dominated by swings in export prices.

For the U.K., while both inflation measures are positively correlated with unit labor costs, the relation is noticeably stronger for CPI inflation. This may appear surprising from the perspective of models that stress domestic-goods inflation as being driven by marginal cost, and import price inflation as determined by a different set of factors.

¹⁴ We are grateful to Ben Dolman and David Gruen for provision of these series.

But we have argued that there are grounds for regarding import price inflation, and so total CPI inflation, as driven heavily by domestic-economy factors.

Unit labor cost is negatively correlated with measures of detrended output in five out of six cases. For example, the correlation between HP filtered log GDP and real unit labor cost is -0.36 for the United Kingdom, -0.24 for Australia, and -0.22 for the United States. We have argued that this is indirect evidence that detrended output is a poor proxy for the output gap—*not* that marginal cost and the output gap are poorly related in reality. We provide evidence on this in Section 5.

Regardless of detrending method, detrended output it is quite weakly correlated with contemporaneous inflation in all three countries; the correlation is always below 0.3. This partly reflects the fact that output tends to lead inflation; allowing for this delivers $\text{Corr}(\pi_t^a \text{cpi}, y_{t-6}^{\text{quad}}) = 0.50$ for the U.S., $\text{Corr}(\pi_t^a \text{cpi}, y_{t-8}^{\text{quad}}) = 0.49$ for the U.K., and $\text{Corr}(\pi_t^a \text{cpi}, y_{t-6}^{\text{quad}}) = 0.41$ for Australia. But it is a harbinger of the poor performance of NKPCs estimated with detrended output for all three countries. GG (1999, p. 201) report that the detrended-output NKPC delivers an incorrect sign on U.S. data, and GGL (2001) report the same for the euro area. They attribute these weak results to the fact that the output-gap-based NKPC specifies inflation as a function of current and expected future gaps, while empirically, detrended output leads inflation. It follows that NKPCs with detrended output as the forcing process will not work. For our own dataset, we reaffirm this finding in Table 2. It shows that negative coefficients in detrended-output-based NKPC estimates hold regardless of inflation definition for the U.S., the U.K., and Australia.

4.3 Potential output and output gap series

We generate a theory-consistent potential GDP series for each country by calibrating key preference and production parameters, generating the implied IS and technology shocks from the data, and using the method of Section 3.3 to express potential GDP as a distributed lag of these shocks. We give a brief description here; the Appendix provides details.

The key production function parameters are set as follows: $\alpha = 0.7$, $\delta = 0.025$. These parameter choices are sufficient to generate a Solow residual from each country's output and factor-input data. The technology shock is defined as the HP-filtered values of these Solow residuals.

For preferences, we restrict ourselves to the case where the IS shock is white noise, so that habit formation is the only source of persistence in consumption behavior. We fix the habit formation parameter h for each country to $h = 0.8$, Fuhrer's (2000) estimate. We then calibrate σ , the intertemporal elasticity of substitution of consumption, to a value consistent with λ_t in (6) being white noise for each country. This gives $\sigma = 0.7$ for the U.S., $\sigma = 0.9$ for the U.K., and $\sigma = 0.6$ for Australia.¹⁵

We obtain potential GDP series by solving the model of Section 3 under price flexibility and obtaining expression (22). We solve the model under two settings: (i) the basic RBC-style model without capital adjustment costs ($\phi = 0$ in equation (5)); and (ii) the model with capital adjustment costs (specifically, $\phi = 0.35$). Each setting implies different profiles for potential GDP. For example, the relative importance of IS shocks increases when there are capital adjustment costs, because it becomes less easy to switch the composition of a given amount of potential GDP between consumption and investment (see our 2001 paper).

Each version of the model needs to be solved separately for each country because the countries differ in their value of σ as well as ρ_a , the AR(1) parameter for the technology shocks. The “output gap” is then defined as HP-filtered GDP minus our potential GDP series.¹⁶

Figures 4–6 plot the resulting output gap series, under the assumption of capital adjustment costs, for each country against a more standard measure of the output gap—quadratically detrended output. We note three properties of the gap series. First, we find that the theory-based output gap and detrended output are not very closely related. Second, the output-gap series are of much smaller amplitude than detrended output, a result consistent with our simulation evidence in Neiss and Nelson (2001) and with the hypothesis that much observed GDP variation reflects variation in potential GDP. Third, the theory-based output gap tends to lag detrended output. Mechanically, this arises from the fact that potential output is a function of lagged shocks, so the output gap tends to be more inertial than detrended output. The figures therefore answer a reservation about the gap-based NKPC expressed by GG (1999, pp. 198, 201–202, 204). They noted that the gap-based NKPC implies inflation leads the gap, and questioned whether this is consistent with detrended output leading

¹⁵ We also set $\beta = 0.99$ in equation (6).

¹⁶ No detrending is necessary for our potential GDP series; as our estimates of the underlying real shocks that appear in equation (22) are generated from detrended or growth-rate data, our series for y_t^* corresponds conceptually to the detrended level of natural GDP.

inflation. We suggest that the two observations are not irreconcilable, because the gap itself lags detrended output.

4.4 Rationalization of the disturbance term u_t

In our estimates of equations (1) and (2), we will include the disturbance term u_t , i.e., a structural shock to the Phillips curve. Such a shock is needed for realistic empirical work with these equations. EHL (2000, p. 298) argue that this term is itself evidence against sticky-price models, on the grounds that models without wage rigidities cannot provide a microeconomic foundation for the presence of this disturbance.

We argue, however, that in the case where u_t is white noise, this disturbance does have an interpretation in sticky-price models, namely as a *price-level shock*. By this we mean a shock that permanently raises the price level, but (provided that monetary policy is non-accommodative)¹⁷ only temporarily increases inflation by raising π_t relative to $E_t\pi_{t+1}$ ^a. As Meltzer (1977, p. 183) puts it, “a one-time change in tastes, the degree of monopoly, or other real variables changes the price level... [W]e require a theory that distinguishes between once-and-for-all price changes and maintained rates of price change.” Similarly, Clarida, Galí, and Gertler (2000, p. 147) distinguish “transitory periods of sharp increases in the general price level” from “persistent high inflation.”

Changes in the real steady state of the economy, arising from (e.g.) changes in the degree of competition, produce movements in the steady-state level of natural output. For a given measure of actual output y_t , these changes in the steady state value of potential y_t^* should produce an output gap. A correct measure of the output gap that included variations in potential GDP due to high-frequency changes in the economy’s steady state, would capture the effects Meltzer mentions. In practice, however, the coefficients in our potential GDP expression (22) are based on the assumption of a constant steady state. In general, the log-level of (detrended) log potential output in period t is composed of a steady-state component, plus technology and IS shocks that drive fluctuations around the steady state:

$$y_t^* = m_t(y_t^*) + b(L)a_t + c(L)\lambda_t, \quad (24)$$

where $m_t(y_t^*)$ is the steady state log-level of potential, and $b(L)$ and $c(L)$ are lag polynomials. If the steady state of the economy was constant, then $m_t(y_t^*)$ would be a

¹⁷ In the sense that it keeps current and expected deviations of output from potential unchanged.

constant and the a_t and λ_t shocks then fully index the variation in potential GDP. Suppose, however, that the degree of monopoly power in the economy underwent a discrete change during the sample period under consideration.¹⁸ Then treating $m_t(y_t^*)$ as a constant would only be an approximation, and while there would be no *mean* error in measuring y_t^* , there would be some variation in y_t^* not captured by the demand and supply shocks alone. Equation (1) can then be written:

$$\pi_t^a = \beta E_t \pi_{t+1}^a + \lambda_y (y_t - y_t^{s*}) - \lambda_y m_{y,t}, \quad (25)$$

where y_t^{s*} is the portion of potential output variation due to the shocks, and $m_{y,t} = m_t(y_t^*) - E[m_t(y_t^*)]$ is the deviation of steady state log potential in period t from a constant. Clearly, with $u_t \equiv -\lambda_y m_{y,t}$, price-level shocks that move steady-state potential provide a rationale for a disturbance term in the output-gap-based NKPC (1). Furthermore, if these changes in the steady state of the economy were discrete in nature, they would provide a rationale for u_t being approximately white noise.

5 Empirical results

This section presents our NKPC estimates for the three countries. We first report marginal-cost-based NKPCs analogous to those in GG, and then estimate NKPCs using a theory-based output-gap series. We begin with the United States.

5.1 United States

We first report estimates for the U.S. of the following version of equation (2):

$$\pi_t^a = \beta E_t \pi_{t+1}^a + \lambda ulc_t + d_0 + d_1 DNIXON_t + u_t. \quad (26)$$

Here $DNIXON_t$ is a dummy variable for the Nixon price controls; this follows work on backward-looking Phillips curves (Gordon, 1982) and is consistent with our wish to control for price-level shocks. To capture the fact that the controls suppressed a

¹⁸ Giannoni (2000) and Woodford (2001b) argue that time-varying distortions such as a changing degree of monopoly power can create movements in potential GDP. In general, because they affect the structure of the economy, movements in these distortions affect both the steady-state level of potential (our $m_t(y_t^*)$ variable) and the dynamic response of potential to shocks (the $b(L)$ and $c(L)$ coefficients in equation (24)). Our price-level-shock definition is restricted to refer only to the first source of change. An increase in the steady-state markup, or an increase in the factor of proportionality in the utility function (Meltzer's "one-time change in tastes," which does not fall into the category of a changing distortion) would be suitable candidates as price-level shocks, as they affect the steady-state potential level but not the shock response coefficients in the equilibrium expression for potential. Note that, in contrast to Giannoni, our price-level-shock interpretation of the u_t process implies that the $(y - y^*)$ term in equation (25) is still the log-deviation of output from its natural level, rather than from its efficient level.

growing amount of inflation (because monetary policy was relaxed with the controls' introduction), we make the dummy equal to the length of time in which the controls have been in effect. The dummy is constructed on the assumption that all the suppressed inflation was recorded in actual inflation once the controls were lifted.¹⁹

GG (1999) and GGL (2001) obtained β estimates in the 0.9–0.95 range, somewhat lower than the value of about 0.99 suggested by theory. Our own estimates of β were occasionally above 1.0 for some countries and specifications, so we report results conditional on two imposed β values: $\beta = 0.99$ and $\beta = 0.942$ (GG's (1999) estimate).

Estimates of equation (26), first using GDP deflator inflation, are given in Table 3. Using an instrument list like GG's,²⁰ we obtain (for $\beta = 0.942$) a t -value of 1.4 on *ulc*, and a coefficient larger and more significant than reported in Rudd and Whelan's (2001) replication, but not as significant as that ($t = 1.9$) reported by GG for 1961 Q1–1997 Q4. We found, however, that modest changes in the instrument definitions and an instrument-choice strategy close to that of GGL (2001) gave somewhat more significant results. Specifically, GGL set the lag length for instruments other than inflation somewhat shorter than for lagged inflation, the rationale being that the extra lags of inflation contribute heavily to the explanatory power of the “first-stage” regressions that underlie the instrumental variables estimates. In light of this, we formulate an “instrument set 2” that shortens the lag length of non-inflation instruments from 4 to 2. We also replace lags of the long bond/funds rate spread as an instrument with separate lags of the Treasury bill rate and the long bond rate; replace quarterly wage and commodity inflation by the corresponding four-quarter inflation rates; and drop detrended output as an instrument.

With this different instrument set we obtain (setting $\beta = 0.942$) an estimate of $\lambda = 0.098$ with $t = 1.9$, similar to GG's $\lambda = 0.092$ ($t = 1.9$). With CPI inflation we obtain a larger coefficient estimate and the same level of significance. When $\beta = 0.99$, which corresponds more closely to the β value suggested by theory, the size and significance of the λ estimates diminish in all cases. The Nixon-controls dummy coefficient is the right sign throughout, but only contributes minor explanatory power to CPI inflation. For the GDP deflator it is highly insignificant and we drop it from the specification.

¹⁹ The Nixon dummy equals 1 in 1971 Q3, 4 in 1971 Q4, ..., 31 in 1974 Q1, and -31 in 1974 Q2.

²⁰ The differences are our inclusion of the Nixon dummy variable and our use of the percentage change rather than the log change in computing growth rates.

Tables 4 and 5 give NKPC estimates using our theory-based output gap as the forcing process.²¹ In contrast to Table 2 where the output gap was measured by detrended output, the theory-consistent output gaps enter the estimated NKPCs positively. Moreover, when CPI inflation is used and potential output is defined as subject to capital adjustment costs, the coefficients are significant (whether $\beta = 0.942$ or 0.99 is used). Indeed, in these cases our gap measure performs better in explaining inflation than did unit labor costs. The correlations at the bottom of Table 5 indicate that our gap measures are positively correlated with unit labor cost, consistent with the prediction one would make if prices were sticky and labor markets could be approximated as competitive.

One unsatisfactory aspect of the results, however, is that the coefficients on the output gap seem to be on the high side. In Rotemberg and Woodford (1997), for example, λ_y is only 0.11.²² The counterpart of their λ_y estimate is high output gap volatility: large gap movements lead to only modest movements in inflation. In our U.S. results, output-gap variability is mild, but the inflation response to gap movements is quite strong.

Variable capital utilization corrupts Solow residuals as measures of technology shocks (Summers, 1986). As a result, we may be attributing too much observed GDP variability to potential GDP and too little to the gap. To check this possibility, we regenerate technology shocks, adjusted for variable utilization using the Federal Reserve's capacity utilization series. We then construct a model of potential GDP based on the model in Section 3.4, augmented by variable capacity utilization as in Burnside and Eichenbaum (1996).²³ The resulting model of potential GDP features less variable technology shocks but much larger responses of potential to those shocks than the baseline model. A gap series was generated from the new potential GDP estimates.

²¹ As the model underlying our potential GDP series has capital and habit formation, the NKPC (1) holds only approximately (Amato and Laubach, 2001, derive the NKPC in a model with habit formation but no capital). Model simulations suggest, however, that the approximation is reasonable.

²² The estimates in Tables 4 and 5 nevertheless strongly support sticky prices over flexible prices, which correspond to the limiting case of λ_y infinite. Sbordone (2002) finds that even modest amounts of price stickiness can achieve a major improvement in fit over the flexible-price model.

²³ To allow for time varying capital utilization (denoted U_t), equation (23) is amended to:

$$\psi_t(1+2\phi X_t) - \beta E_t \psi_{t+1} [(1-\delta U_{t+1}^{\phi_1})(1+2\phi X_{t+1}) + U_{t+1} z_{t+1}] = 0. \quad (23a)$$

and the system now includes an additional equation:

$$z_t = \delta \phi_1 U_t^{\phi_1} (1+2\phi X_t). \quad (23b)$$

Results using this alternative gap series are given in Table 6. They are favorable to the output gap proxy hypothesis. The coefficient on the gap continues to be positive and is now at a value closer to Rotemberg and Woodford’s estimate. Moreover, it is highly significant. The combination favored by the data is one with variable capacity utilization no capital adjustment costs, as in King and Rebelo (1999).

5.2 United Kingdom

We turn next to the U.K. As for the U.S., we first report NKPC estimates with real marginal cost as the driving process. Compared to previous estimates of marginal-cost-based NKPCs on U.K. data (e.g. Amato and Gerlach, 2000; Batini, Jackson, and Nickell, 2000; Balakrishnan and López-Salido, 2001) we (i) estimate equations on broader definitions of inflation, (ii) allow explicitly for price-level shocks, and (iii) estimate on a longer sample (40 years of quarterly data).

While the u_t disturbance term in equation (2) accounts for price-level shocks of a normal magnitude, we include dummy variables for some specific large price level increases. We include a dummy for the Heath government’s price controls of 1972–1974. Like the Nixon controls, these were imposed alongside increasingly expansionary monetary policy, so we define the U.K. price control dummy as equal to the number of months that the controls have been in effect. We also include dummy variables for increases in indirect tax in 1968, 1973 and 1979, and the 1974 cut. In principle, indirect-tax (from 1973, VAT) changes are relevant primarily for CPI inflation and not GDP deflator inflation. But 1979 Q3 (the date of the 1979 VAT increase) was also the time of the second oil shock as well as sharp increases in the prices of government sector output. Similarly, the 1974 VAT cut was accompanied by the extension of policies to hold down private and government sector prices directly, which could have lowered GDP deflator inflation. So we include dummies for the 1974 and 1979 price level movements in our GDP deflator specification too. For CPI inflation, we also include a dummy to capture the jump in the measured Retail Price Index in 1990 due to the increase in property tax (the introduction of the poll tax) in 1990 Q2.

The tax changes of 1968, 1973, 1979, and 1990 were announced a quarter in advance so their effect on CPI inflation could be rationally anticipated. For the marginal-cost-based NKPC, this implies an empirical specification of the form:

$$\pi_t^a = \beta E_t \pi_{t+1}^a + d_0 + \lambda ulc_t + d_1(D682_t(1-\beta F)) + d_2(D732_t(1-\beta F)) + d_3(DVAT74_t) + d_4(DHEATH_t) + d_5(D793_t(1-\beta F)) + d_6(DPOLL_t(1-\beta F)) + u_t, \quad (27)$$

where F is the lead operator. We estimate (27) for the two β settings used for the U.S.

Table 7 indicates that the *ulc* series explains both CPI inflation and GDP deflator inflation fairly well. For CPI inflation, for example, *ulc* has a t -ratio of 1.7 when $\beta = 0.99$ and 2.2 for $\beta = 0.942$. We find that our output gap series gives a similar but slightly poorer fit in explaining CPI inflation, with the best results being when potential is defined without capital adjustment costs. In that case, the coefficient on the gap has a t -ratio of 1.3 when $\beta = 0.99$. Results with GDP deflator inflation are similar. What we take from the U.K. results is that costs are only marginally better than the gap in explaining CPI inflation, and that this suggests the relevance of the “output gap proxy” interpretation of cost-based NKPC estimates. This interpretation is backed up by the correlations given as a memo to Table 7. They show that the theory-based output gap is (1) positively correlated with inflation; (2) negatively correlated with detrended output; and (3) strongly related to labor costs. As shown in our 2001 paper, all three observations are consistent with inflation being generated by a model with sticky prices but no labor rigidities.

5.3 Australia

We next present estimated NKPCs for Australia. Throughout, we include three dummy variables for jumps in the CPI series (and so rises in measured inflation). These are in 1975 Q3, 1975 Q4, 1976 Q4, and 2000 Q3. The 1975 Q4 and 2000 Q3 increases correspond to pre-announced shifts by the federal government to greater indirect taxation; while 1975 Q3 and 1976 Q4 saw swings in the measured cost of living due to changes in the financing of universal health insurance.²⁴

Our estimates of equation (2), the cost-based Phillips curve, on Australian data are reported in Table 8. The *ulc* series has t -value 1.3 when $\beta = 0.99$ is imposed and 1.8 with $\beta = 0.942$. We also report estimates of equation (1) with the theory-based output gap. In contrast to the Table 2 estimates with detrended output, the gap coefficient is correctly signed. The best results are when potential GDP is defined assuming no capital adjustment costs; the gap then has t -value 1.8 with $\beta = 0.99$ imposed and 1.9 with $\beta = 0.942$. We interpret these results as quite supportive of the output gap proxy hypothesis: the gap does about as well as, or better than, *ulc* in explaining inflation.

²⁴ The 1976 *OECD Economic Survey of Australia* noted (p. 19) that “[m]uch of the uneven development of the consumer price index is attributable to the effects of the introduction of the Medibank scheme in the September quarter of 1975, which greatly reduced the user-cost of health services; and to the increases in public charges and indirect taxes in the following quarter.” The sizable impact of the 1976 health insurance change on the CPI is evident in Figure 1 of de Brouwer and Ericsson (1998).

One reason why the gap series is not more significant in these regressions is clear from Figure 6: the gap is not big enough to explain the breakout of double-digit inflation in the 1970s. Most likely, our use of Solow residuals as measures of technology shocks means too much of the early 1970s boom is attributed to productivity shocks and too little to expansionary monetary policy. Our conjecture is that a cleaner measure of technology shocks, which abstracted from the component of Solow residuals that responds endogenously to monetary policy shocks, would improve our ability to model inflation using the output gap, for all the countries in our sample.²⁵ As it is, the fact that our gap-based NKPCs for Australia are able to deliver an explanation of inflation behavior close to that of cost-based NKPCs is notable.

6 Conclusions

The results in this paper suggest the following conclusions.

First, we were able to find strong support for the approach advocated by GG (1999), GGL (2001) and Sbordone (2001, 2002), of modeling inflation using unit labor cost data. We found their structural modeling of price setting remained valid for a wider range of countries, broader measures of inflation, and extended sample periods. Their approach of basing empirical modeling on explicit optimizing behavior, and of using theory to measure the process driving inflation, appears to be a very useful way of analyzing inflation.

Second, we offer an interpretation of these results that suggests labor market rigidities (such as real or nominal wage stickiness) are not crucial for modeling of inflation or monetary policy analysis. Once the output gap is defined in a manner consistent with theory, gap-based Phillips curves fit the data about as well as cost-based Phillips curves. Our evidence supports Galí's (2002) and Woodford's (2001a) "output gap proxy" interpretation of the empirical success of cost-based Phillips curves, according to which the main source of nominal rigidity is in prices rather than wages.

Third, our approach suggests several problems with traditional methods of measuring the output gap, such as by detrending. Estimates of the NKPC using detrended output to measure the output gap do not provide valid tests of the gap-based NKPC. The negative coefficient on the "output gap" in such estimates does not imply that the true

²⁵ We could not perform a simple correction of the Solow residual series for variable capacity utilization as we did for the U.S, because the capacity utilization series for Australia (described in Otto, 1999) is not in comparable units to the U.S. series.

coefficient is negative; rather, it reflects the fact that detrended output is a poor approximation for the output gap. Working with theory-based gap estimates, we find that NKPC estimates deliver consistently positive coefficients on the output gap.

Our estimates of the output gap indicate that it is closely related to real marginal cost, again consistent with models where price but not wage stickiness is central. If unit labor costs and the output gap are closely related, detailed modeling of labor market rigidities is not a high priority in analyzing inflation. On the other hand, we find that modeling the dynamic effects of real shocks—not only productivity shocks but also preference shocks—on potential GDP is crucial for understanding inflation behavior. Because they explicitly relate potential output dynamics to underlying shocks, it is important to use optimizing models in monetary policy analysis. Taken together, our results suggest that monetary policy analysis with sticky-price optimizing models (such as in Clarida, Galí, and Gertler 2000, Sect. IV) is valid.

We note some caveats about our analysis. Firstly, the real shocks that drive potential output are not observable, and the proxies for the shocks that we have used can be improved upon. Secondly, the evidence we have provided against sticky-wage models is only indirect. Beyond looking at the relationship between the output gap and marginal cost, we have not attempted to examine directly the implications of sticky-price models for the labor market. In particular, we have not attempted to estimate wage-adjustment equations. System estimation of optimization-based wage and price-adjustment equations would help distinguish between models that rely primarily on price stickiness from those that emphasize sticky wages.

Data Appendix

For all countries, the following were used:

- Inflation formula: $\pi_t^a = (P_t/P_{t-1})^4 - 1$, where P_t is quarterly average of price index.
- IS shock generation: λ_t series backed out from loglinearized versions of equation (6) and (8), with VAR projections proxying for the expectations of π_{t+1} , Δc_{t+1} , and Δc_{t+2} .

United States

P_t : For CPI, seasonally adjusted quarterly average. Sources for CPI and GDP deflator: Federal Reserve Bank of St. Louis.

y_t^{hp} (labelled y_{us}^{hp} henceforth): HP-filtered log real GDP. HP-filtering over 1947 Q1–2001 Q2.

c_t : log of quarterly nondurables consumption. Source: Federal Reserve Bank of St. Louis.

R : quarterly average Treasury bill rate. Source: Federal Reserve Bank of St. Louis.

ulc : log labor share, based on nonfarm business sector data. Source for labor data: BLS web page.

a_t : HP filtered version of Solow residual $y - 0.7n$. Implied ρ_a : 0.77. With capacity utilization adjustment, $a = y - 0.7n - 0.3 \log(CAPU)$, and implied ρ_a : 0.59.

Projections of Δc and π^a used for IS shock generation are forecasts from 1955 Q1–2000 Q4 VAR(8) in R , Δc , π_{cpi}^a , y_{us}^{hp} , plus Nixon price control dummy.

United Kingdom

P : For CPI, seasonally adjusted quarterly average of RPI/RPIX. For GDP deflator, ONS series code YBHA divided by ABMI.

y^{hp} : HP-filtered log real GDP. Series code ABMI, HP-filtered and logged over 1955 Q1–2001 Q2.

c_t : log of quarterly nondurables consumption.

R : quarterly average Treasury bill rate.

ulc : updated version of Batini-Jackson-Nickell (2001) series with our (2001) N series used pre-1978.

a : HP filtered version of Solow residual $y - 0.7n - 0.3k$, where k is obtained by logging cumulated private investment. Implied ρ_a : 0.73.

Projections of Δc and π^a used for IS shock generation are forecasts from 1959 Q1–2000 Q4 VAR(8) in R , Δc , π_{cpi}^a , y^{hp} , y_{us}^{hp} , plus tax and price control dummies.

Australia

P : *IFS* CPI series except that observations from 1984 Q1 are rescaled to take into account the break in the series induced by the reintroduction of universal health insurance. The use of CPI data that are adjusted for the 1984 break, but are otherwise equal to the headline series, is common in empirical work; see e.g. Gruen and Stevens (2000, Figure 5).

y : log of seasonally adjusted real GDP. To obtain this, the official seasonally adjusted real GDP series (RGDP) (expenditure definition) was downloaded from Reserve Bank of Australia web page (www.rba.gov.au). Then a regression of $\Delta \log(\text{RGDP})$ on seasonal dummies revealed significant seasonality for the sample up to 1979 Q4. The residuals from this regression, with mean of the dependent variable restored, were taken as the new seasonally adjusted Δy series for 1959 Q4–1979 Q4 and were integrated and spliced into the log of the official seasonally adjusted series to obtain a time series for y for 1959 Q3–2001 Q2.

Unit labor cost data: Average weekly earnings and employment data provided by Ben Dolman and David Gruen. Product expressed as ratio to nominal GDP (source for nominal GDP: RBA web page).

y^{hp} : y HP-filtered over 1959 Q3–2001 Q2.

a : Technology shock series. HP-filtered Solow residual, where the Solow residual is defined as $y - 0.7n$, with n series the employment data used in labor cost construction. Implied ρ_a : 0.59.

c : log of private consumption series. Source: RBA web page.

R : Three-month Treasury note rate, quarterly average. Corresponds to the *IFS* series, but RBA data (*Reserve Bank of Australia Bulletin* and RBA web page) were used as source in obtaining quarterly averages for pre-1969 Q3 and from 1992 Q3 on.

Projections of Δc and π^a used for IS shock generation are forecasts from 1962 Q2–2000 Q4 VAR(2) in R , Δc , π^a , y^{hp} , y_{us}^{hp} , plus tax dummies.

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Table 1: Data correlations, 1961 Q1–2000 Q4

Correlation matrix: United States

	π_t^a	π_t^a	ulc_t	y_t^{hp}	y_t^{quad}
π_t^a	1.00	0.90	0.44	0.27	0.27
π_t^a		1.00	0.52	0.10	0.21
ulc_t			1.00	-0.22	-0.15
y_t^{hp}				1.00	0.73
y_t^{quad}					1.00

Correlation matrix: United Kingdom

	π_t^a	π_t^a	ulc_t	y_t^{hp}	y_t^{quad}
π_t^a	1.00	0.84	0.44	-0.01	0.11
π_t^a		1.00	0.34	0.06	0.13
ulc_t			1.00	-0.36	-0.12
y_t^{hp}				1.00	0.66
y_t^{quad}					1.00

Correlation matrix: Australia

	π_t^a	ulc_t	y_t^{hp}	y_t^{quad}
π_t^a	1.00	0.52	0.14	0.23
ulc_t		1.00	-0.24	0.11
y_t^{hp}			1.00	0.55
y_t^{quad}				1.00

Table 2: Estimates of NKPC with HP-filtered log GDP as forcing process

Specification: $\pi_t^a = \beta E_t \pi_{t+1}^a + \lambda_y y_t^{hp} + \text{constant} + \text{price shock dummies} + u_t$

	λ_y
United States, CPI inflation, 1961 Q1–2000 Q4	-0.008 ($t = 0.1$)
United States, GDP deflator inflation, 1961 Q1–2000 Q4	-0.090 ($t = 1.3$)
United Kingdom, RPIX inflation, 1961 Q4–2000 Q4	-0.392 ($t = 1.8$)
United Kingdom, GDP deflator inflation, 1961 Q4–2000 Q4	-0.389 ($t = 1.4$)
Australia, CPI inflation, 1962 Q3–2000 Q4	-0.265 ($t = 1.1$)

Note: $\beta = 0.99$ imposed in estimation throughout. Instruments: U.K. regression, as in Table 7. Australia, as in Table 8. U.S., as in Table 3, with instrument list 1, with y_t^{hp} lags 1–4 instead of y_t^{quad} .

Table 3: United States: Estimates of Marginal Cost NKPC
Sample Period 1961 Q1–2000 Q4

Specification: $\pi_t^a = b_0 + \beta E_t \pi_{t+1}^a + \lambda ulc_t + \text{dummy} + u_t$								
	GDP deflator inflation				CPI inflation			
	Instrument set 1		Instrument set 2		Instrument set 1		Instrument set 2	
	$\beta =$ 0.99	$\beta =$ 0.942	$\beta =$ 0.99	$\beta =$ 0.942	$\beta =$ 0.99	$\beta =$ 0.942	$\beta =$ 0.99	$\beta =$ 0.942
<i>Coefficient</i>								
β	0.99 (—)	0.942 (—)	0.99 (—)	0.942 (—)	0.99 (—)	0.942 (—)	0.99 (—)	0.942 (—)
λ	0.033 ($t = 0.6$)	0.070 ($t = 1.4$)	0.062 ($t = 1.2$)	0.098 ($t = 1.9$)	0.063 ($t = 0.8$)	0.101 ($t = 1.3$)	0.109 ($t = 1.4$)	0.147 ($t = 1.9$)
Nixon price controls dummy	−0.0002 ($t = 0.1$)	−0.0002 ($t = 0.2$)	—	—	−0.0003 ($t = 1.1$)	−0.0003 ($t = 1.0$)	−0.0003 ($t = 1.3$)	−0.0003 ($t = 1.2$)

Constant term is included in all regressions.
Instrument set 1: lags 1–4 of: π^a , ulc , quadratically detrended log output, quarterly wage inflation, quarterly commodity price inflation, bond/funds rate spread, plus Nixon dummy.
Instrument set 2: lags 1–4 of π^a , lags 1–2 of ulc , nominal Treasury bill rate, 10-year bond rate, annual wage inflation, annual commodity price inflation, plus Nixon dummy.

Table 4: United States: Estimates of NKPC with theory-based output gap Sample Period 1961 Q1–2000 Q4					
	Specification: $\pi_t^a = b_0 + \beta E_t \pi_{t+1}^a + \lambda_y (y_t - y_t^*) + \text{dummy} + u_t$				
	GDP deflator inflation, $\beta = 0.99$ imposed				
	Instrument set 1		Instrument set 2		
	y_t^* defn: No K adj costs ($\varphi = 0$)	y_t^* defn: K adj costs ($\varphi = 0.35$)	y_t^* defn: No K adj costs ($\varphi = 0$)	y_t^* defn: K adj costs ($\varphi = 0.35$)	
Coefficient					
β	0.99 (—)	0.99 (—)	0.99 (—)	0.99 (—)	
λ_y	0.090 ($t = 1.1$)	0.123 ($t = 0.5$)	0.099 ($t = 1.2$)	0.222 ($t = 0.9$)	
	CPI inflation, $\beta = 0.99$ imposed				
	Instrument set 1		Instrument set 2		
	y_t^* defn: No K adj costs ($\varphi = 0$)	y_t^* defn: K adj costs ($\varphi = 0.35$)	y_t^* defn: No K adj costs ($\varphi = 0$)	y_t^* defn: K adj costs ($\varphi = 0.35$)	
Coefficient					
β	0.99 (—)	0.99 (—)	0.99 (—)	0.99 (—)	
λ_y	0.197 ($t = 1.6$)	0.808 ($t = 2.3$)	0.186 ($t = 1.5$)	0.801 ($t = 2.3$)	
Nixon price controls dummy	–0.0001 ($t = 0.5$)	–0.0002 ($t = 0.6$)	–0.0001 ($t = 0.5$)	–0.0002 ($t = 0.6$)	
Memo item: Correlations 1961 Q1–2000 Q4					
Correlation with:					
	$\pi_{t\ cpi}^a$	$\pi_{t\ def}^a$	ulc_t	y_t^{hp}	y_t^{quad}
Output gap (y_t^* : $\varphi = 0.0$)	0.03	0.12	0.39	–0.77	–0.53
Output gap (y_t^* : $\varphi = 0.35$)	0.30	0.30	0.37	–0.22	–0.08
Constant term is included in all regressions.					
Instrument set 1: lags 1–4 of: π^a , ulc , $(y - y^*)$, quarterly wage inflation, quarterly commodity price inflation, bond/funds rate spread, plus Nixon dummy.					
Instrument set 2: lags 1–4 of π^a , lags 1–2 of ulc , $(y - y^*)$, nominal Treasury bill rate, 10-year bond rate, annual wage inflation, annual commodity price inflation, plus Nixon dummy.					

Table 5: United States: Estimates of NKPC with theory-based output gap Sample Period 1961 Q1–2000 Q4					
Specification: $\pi_t^a = b_0 + \beta E_t \pi_{t+1}^a + \lambda_y (y_t - y_t^*) + \text{dummy} + u_t$					
GDP deflator inflation, $\beta = 0.942$ imposed					
Instrument set 1			Instrument set 2		
	y_t^* defn: No K adj costs ($\varphi = 0$)	y_t^* defn: K adj costs ($\varphi = 0.35$)	y_t^* defn: No K adj costs ($\varphi = 0$)	y_t^* defn: K adj costs ($\varphi = 0.35$)	
<i>Coefficient</i>					
β	0.942 (—)	0.942 (—)	0.942 (—)	0.942 (—)	
λ_y	0.118 ($t = 1.5$)	0.247 ($t = 1.0$)	0.109 ($t = 1.4$)	0.341 ($t = 1.4$)	
CPI inflation, $\beta = 0.942$ imposed					
Instrument set 1			Instrument set 2		
	y_t^* defn: No K adj costs ($\varphi = 0$)	y_t^* defn: K adj costs ($\varphi = 0.35$)	y_t^* defn: No K adj costs ($\varphi = 0$)	y_t^* defn: K adj costs ($\varphi = 0.35$)	
<i>Coefficient</i>					
β	0.942 (—)	0.942 (—)	0.942 (—)	0.942 (—)	
λ_y	0.213 ($t = 1.7$)	1.031 ($t = 2.8$)	0.202 ($t = 1.6$)	1.033 ($t = 2.8$)	
Nixon price controls dummy	–0.0001 ($t = 0.3$)	–0.0001 ($t = 0.5$)	–0.0001 ($t = 0.3$)	–0.0001 ($t = 0.3$)	
<i>Memo item: Correlations 1961 Q1–2000 Q4</i>					
Correlation with:					
	π_{cpi}^a	π_{def}^a	ulc_t	y_t^{hp}	y_t^{quad}
Output gap (y_t^* : $\varphi = 0.0$)	0.03	0.12	0.39	–0.77	–0.53
Output gap (y_t^* : $\varphi = 0.35$)	0.34	0.32	0.35	–0.10	0.01
Constant term is included in all regressions. Instrument set 1: lags 1–4 of: π^a , ulc , $(y - y^*)$, quarterly wage inflation, quarterly commodity price inflation, bond/funds rate spread, plus Nixon dummy. Instrument set 2: lags 1–4 of π^a , lags 1–2 of ulc , $(y - y^*)$, nominal Treasury bill rate, 10-year bond rate, annual wage inflation, annual commodity price inflation, plus Nixon dummy.					

Table 6: United States: Estimates of NKPC with theory-based output gap Variable Capacity Utilization Sample Period 1961 Q1–2000 Q4				
	GDP deflator inflation		CPI inflation	
	$\beta = 0.942$	$\beta = 0.99$	$\beta = 0.942$	$\beta = 0.99$
y_t^* defn: No K adj costs ($\varphi = 0$)	$\lambda_y = 0.030$ ($t = 1.5$)	$\lambda_y = 0.022$ ($t = 1.0$)	$\lambda_y = 0.143$ ($t = 4.0$)	$\lambda_y = 0.126$ ($t = 3.6$)
y_t^* defn: K adj costs ($\varphi = 0.35$)	$\lambda_y = 0.007$ ($t = 0.1$)	$\lambda_y = -0.016$ ($t = 0.3$)	$\lambda_y = 0.219$ ($t = 2.4$)	$\lambda_y = 0.179$ ($t = 1.9$)
Instrument set 2 used. Nixon dummy included in CPI specification.				

Table 7: New Keynesian Phillips curve estimates: United Kingdom						
Cost-based Phillips curve: $\pi_t^a = b_0 + \beta E_t \pi_{t+1}^a + \lambda ulc_t + \text{tax dummies} + u_t$ Sample Period 1961 Q4–2000 Q4						
	$\beta = 0.99$		$\beta = 0.942$			
CPI inflation	$\lambda = 0.240$ ($t = 1.7$)		$\lambda = 0.311$ ($t = 2.2$)			
GDP deflator inflation	$\lambda = 0.349$ ($t = 1.8$)		$\lambda = 0.413$ ($t = 2.2$)			
Output gap-based Phillips curve: $\pi_t^a = b_0 + \beta E_t \pi_{t+1}^a + \lambda_y (y_t - y_t^*) + \text{tax dummies} + u_t$ Sample Period 1961 Q4–2000 Q4						
	CPI inflation		GDP deflator inflation			
	$\beta = 0.99$	$\beta = 0.942$	$\beta = 0.99$	$\beta = 0.942$		
y_t^* definition: No capital adj costs ($\phi = 0$)	$\lambda_y = 0.179$ ($t = 1.3$)	$\lambda_y = 0.215$ ($t = 1.6$)	$\lambda_y = 0.284$ ($t = 1.6$)	$\lambda_y = 0.309$ ($t = 1.8$)		
y_t^* definition: Capital adj costs ($\phi = 0.35$)	$\lambda_y = 0.211$ ($t = 0.8$)	$\lambda_y = 0.318$ ($t = 1.2$)	$\lambda_y = 0.413$ ($t = 1.2$)	$\lambda_y = 0.519$ ($t = 1.5$)		
<i>Memo item 1. Correlations</i>						
	Correlation with:					
	$\pi_{t\ cpi}^a$	$\pi_{t\ def}^a$	ulc_t	y_t^{hp}	y_t^{quad}	
<i>Output gap</i> (y_t^* : $\phi = 0$)	0.40	0.36	0.56	-0.24	0.02	
<i>Output gap</i> (y_t^* : $\phi = 0.35$)	0.26	0.21	0.60	-0.60	-0.31	
<i>Memo item 2. Coefficients on dummy variables</i>						
The estimated NKPCs include dummy variables for price-level shocks. Estimated coefficients on these dummies are similar across different imposed values of β and different forcing variables in the NKPC. In the case $\beta = 0.99$ and ulc as the forcing process, the estimated dummy variable coefficients are:						
	Heath price controls	1968 Q2 indirect tax increase	1973 Q2 VAT introduction	1974 Q3 VAT cut and subsidy changes	1979 Q3 VAT increase	1990 Q2 poll tax introduction
CPI inflation	-0.004 ($t = 3.6$)	0.035 ($t = 1.6$)	0.048 ($t = 2.2$)	-0.064 ($t = 1.8$)	0.238 ($t = 10.8$)	0.075 ($t = 3.5$)
GDP deflator inflation	-0.004 ($t = 2.3$)	—	—	-0.103 ($t = 2.2$)	0.211 ($t = 7.6$)	—
Constant term is included in all regressions. Instruments are: constant, price shock dummies, lags 1–4 of inflation, and lags 1–3 of bill rate, bond rate, labor share, HP filtered output. For output-gap-based NKPC, additional instruments are lags 1–3 of $(y - y^*)$.						

Table 8: New Keynesian Phillips curve estimates: Australia				
Cost-based Phillips curve: $\pi_t^a = b_0 + \beta E_t \pi_{t+1}^a + \lambda ulc_t + \text{tax dummies} + u_t$ Sample Period 1963 Q4–2000 Q4				
	$\beta = 0.99$	$\beta = 0.942$		
CPI inflation	$\lambda = 0.113 (t = 1.3)$	$\lambda = 0.145 (t = 1.8)$		
Output gap-based Phillips curve: $\pi_t^a = b_0 + \beta E_t \pi_{t+1}^a + \lambda_y (y_t - y_t^*) + \text{tax dummies} + u_t$				
	$\beta = 0.99$	$\beta = 0.942$		
y_t^* definition: No capital adj costs ($\varphi = 0$)	$\lambda_y = 0.270$ ($t = 1.8$)	$\lambda_y = 0.290$ ($t = 1.9$)		
y_t^* definition: Capital adj costs ($\varphi = 0.35$)	$\lambda_y = 0.300$ ($t = 1.0$)	$\lambda_y = 0.359$ ($t = 1.3$)		
<i>Memo item 1. Correlations</i>				
Correlation with:				
	$\pi_{t\ cpi}^a$	ulc_t	y_t^{hp}	y_t^{quad}
<i>Output gap</i> ($y_t^*: \varphi = 0$)	0.15	0.40	-0.55	-0.19
<i>Output gap</i> ($y_t^*: \varphi = 0.35$)	0.25	0.32	-0.11	0.12
<i>Memo item 2. Coefficients on dummy variables</i>				
The estimated NKPCs include dummy variables for price-level shocks. Estimated coefficients on these dummies are similar across different imposed values of β and different forcing variables in the NKPC. In the case $\beta = 0.99$ and ulc as the forcing process, the estimated dummy variable coefficients are:				
	1975 Q3	1975 Q4	1976 Q4	2000Q3
	Healthcare tax cut	Indirect tax increase	Healthcare tax increase	Indirect tax increase
CPI inflation	-0.095 ($t = 3.9$)	0.129 ($t = 5.8$)	0.163 ($t = 8.2$)	0.122 ($t = 6.4$)
Constant term is included in all regressions. Instruments are: constant, tax dummies, lags 1–5 of π^a , and lags 1–2 of bill rate, ulc , y^{hp} , y_{us}^{hp} . For gap-based NKPC, lags 1–2 of $(y - y^*)$ replace lags of ulc .				

Figure 1 : U.S. quarterly annualized CPI inflation (π_{cpi}^a) and log real unit labor cost (ulc)

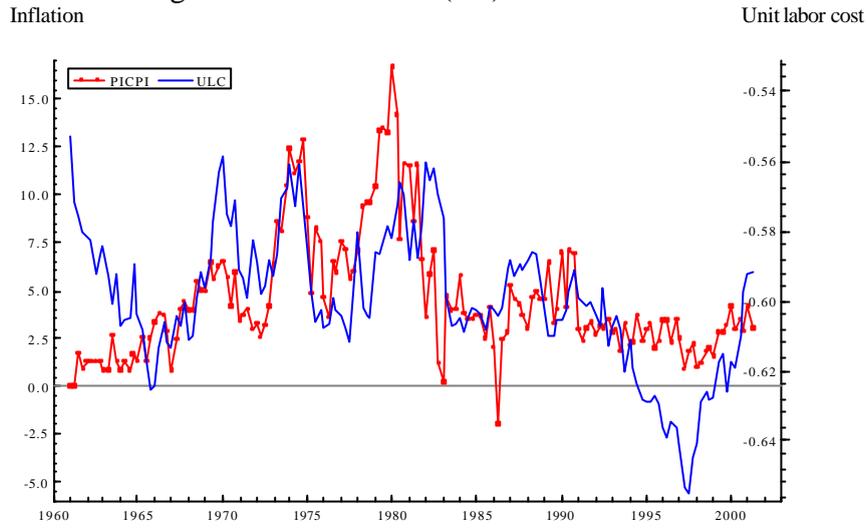


Figure 2: U.K. quarterly annualized CPI inflation (π_{cpi}^a) and log real unit labor cost (ulc)

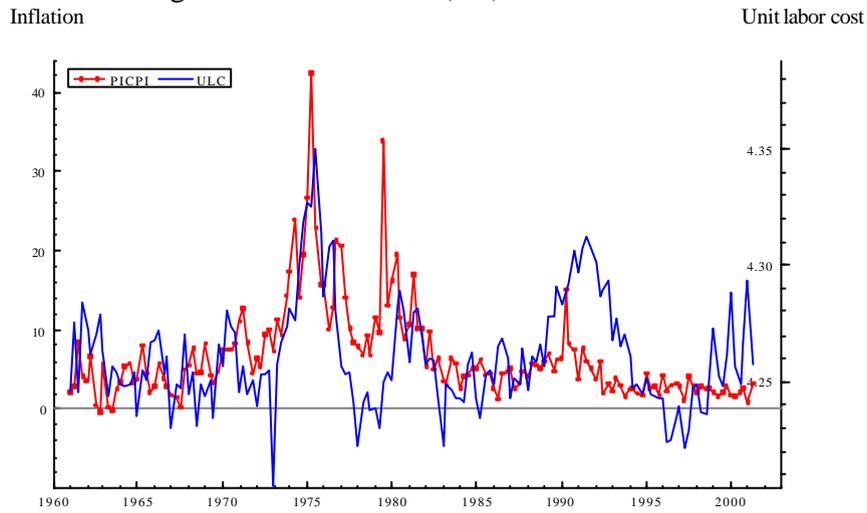


Figure 3: Australia quarterly annualized CPI inflation (π_{cpi}^a) and log real unit labor cost (ulc)

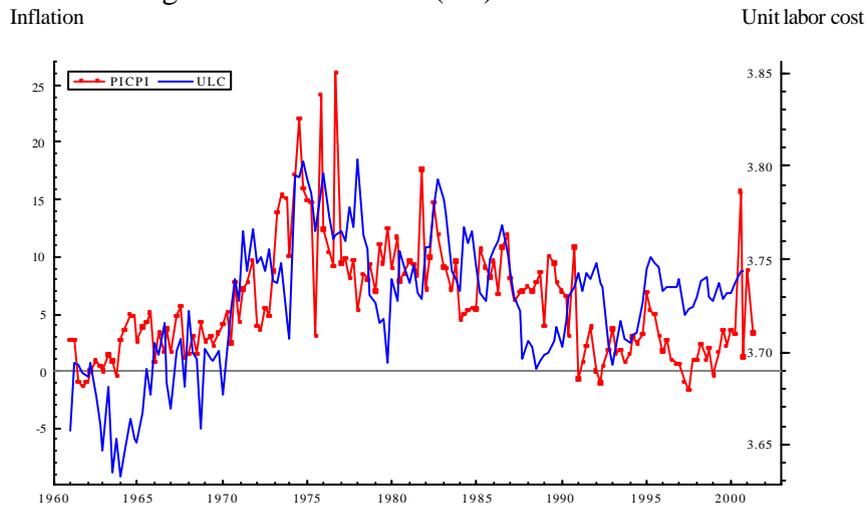


Figure 4: U.S. quadratically detrended log GDP ($YDETR$) vs. theory-based output gap ($YGAP$)

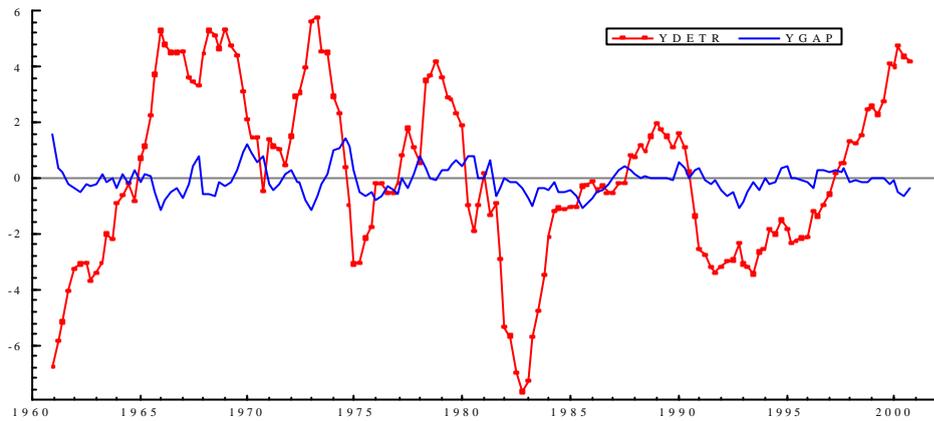


Figure 5: U.K. quadratically detrended log GDP ($YDETR$) vs. theory-based output gap ($YGAP$)

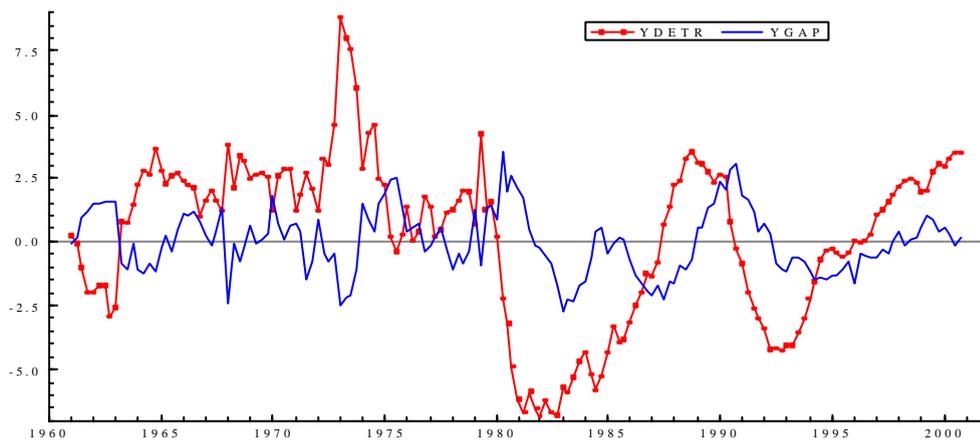


Figure 6: Australia quadratically detrended log GDP ($YDETR$) vs. theory-based output gap ($YGAP$)

