

# IMPACTS OF POLICY REFORMS ON LABOR MIGRATION FROM RURAL MEXICO TO THE UNITED STATES\*

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## Abstract

Using new survey data from Mexico, a dynamic econometric model is estimated to test the effect of policy changes on the flow of migrant labor from rural Mexico to the United States and test for differential effects of policy changes on male and female migration. We find that both IRCA and NAFTA reduced the share of rural Mexicans working in the United States. Increased U.S. border enforcement had the opposite effect. The impacts of these policy variables are small compared with those of macroeconomic variables. The influence of policy and macroeconomic variables is small compared with that of migration networks, as reflected in past migration by villagers to the United States. The effects of all of these variables on migration propensities differ, quantitatively and in some cases qualitatively, by gender.

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## **IMPACTS OF POLICY REFORMS ON LABOR MIGRATION FROM RURAL MEXICO TO THE UNITED STATES**

Immigrant workers from Mexico are a critical component of the supply of labor to agriculture and many non-agricultural sectors in the United States. They constitute 3.5 percent of U.S. labor force but are heavily concentrated into two types of sectors. Twenty-five percent are in services, and twenty-nine percent are involved in production and transportation occupations (Grieco and Ray, 2004). Most are in nonfarm jobs. However, according to the National Agricultural Worker Survey (NAWS), Mexico-born persons represented an estimated 77 percent of the U.S. farm workforce in 1997-98 (up from 57 percent in 1990; U.S. Department of Labor, 2000 and 1991). The majority of these workers (52 percent) were unauthorized. An overwhelming majority originate from households in rural Mexico (U.S. Commission on Immigration Reform, 1997).

Two major policy changes, The North American Free Trade Agreement (NAFTA) and the 1986 Immigration Reform and Control Act (IRCA), together with intensified enforcement along the southern U.S. border, were aimed wholly or partially at curtailing the flow of unauthorized Mexico-to-U.S. migration, potentially reducing the supply of labor to these U.S. economic sectors. The exact impact of NAFTA, IRCA and increased border enforcement is theoretically ambiguous and must be estimated econometrically.

In this paper, we estimate a dynamic econometric model to test the effect of these policy changes on the flow of migrant labor from rural Mexico to the United States. Recognizing that policy changes may have differential effects on male and female labor migration, we estimate the effects of policy changes by the gender of migrant flows, as well. The models are estimated using retrospective data from the 2003 National Mexico Rural Household Survey.

### ***Conceptual Framework***

Given individual, household and community characteristics, policy changes may alter the larger milieu within which migration decisions take place. In this paper we isolate the impact of three policies, IRCA, NAFTA and increased expenditure on enforcement along the U.S.-Mexican border. Each of these policies has counteracting effects on migration, making the overall impact on migration ambiguous.

IRCA had two main components. First, it made employers who hired illegal aliens subject to fines and/or imprisonment. These penalties were meant to discourage the hiring of unauthorized immigrants and reduce migration by dampening the employment expectations of migrants. Second, IRCA provided amnesty to illegal aliens who have lived in the U.S. continually since 1982, if they applied before 1988. This policy legalized U.S. migration contacts for households throughout rural Mexico. It may have encouraged migration by family members of newly legalized migrants, while sending a signal to rural Mexicans that future amnesty deals might be forthcoming. Therefore, these two components of IRCA potentially have counteracting effects on immigration.

NAFTA was only partially motivated by migration concerns but was expected to have far-reaching impacts on migration flows. In the long run, trade liberalization policies open North American markets to Mexico, encouraging export of goods and

decreasing migration pressures.<sup>1</sup> That is, in the long run, trade and migration may be substitutes. After Mexico joined NAFTA, Mexican agricultural exports to the United States did, indeed, increase. However, in the short run, NAFTA could displace rural workers as production shifts from importables to exportables and labor markets adjust to new market realities. Computable general equilibrium models predicted that the increase in labor demand generated by exports to the United States would be insufficient to absorb workers displaced from agricultural activities that had been protected by government policies prior to NAFTA. This, in turn, would stimulate out-migration from rural Mexico (Levy and van Wijnberger: 1992, and Robinson *et.al.*: 1991). Displacement of workers was expected to be a result of dismantling protectionist agrarian policies, as required by NAFTA. Mexico phased out price supports for eleven agricultural field crops and the processing, storing, and marketing activities of the state-run National Company of Popular Subsistence (CONASUPO, Yunez and Barceinas: 2004). Credit subsidies were reduced sharply. For rural workers displaced by policies related to NAFTA, migration may have been a vehicle to overcome short term financial shocks.

The third policy that we evaluate is the increase in enforcement along the U.S. - Mexico border. Increases in border enforcement were meant to curtail unauthorized immigration. However, they could have the opposite effect, by discouraging unauthorized immigrants from returning to their home countries and thus extending their stays in the United States. Increased border enforcement raises smuggler fees, but family members may be willing to pay the increased cost in order to reunite with relatives who extend their stays in the United States.

All three policies' possible impacts on migration thus are complex and theoretically ambiguous. The net effects of these policy shocks on the migration of labor from rural Mexico to the United States can only be determined empirically. In order to isolate the effects of policy changes on migration we need to control for the plethora of individual, household and community variables influencing migration decisions over time, as well as macroeconomic shocks that may have affected migration.

The propensity to migrate and obtain employment in the United States is a function of migration networks and sending-area characteristics that influence the benefits and costs of migrating relative to staying at home. We control for sending-area characteristics and community level heterogeneity by using fixed effects. Migration networks or contacts with employed migrants in the United States are represented by lagged stocks of employed villagers in the United States. Networks may be gender-specific. For example, females may base their migration decision on the knowledge that other females in the village have succeeded in crossing the border and obtaining employment in the United States. In order to evaluate gender-specific network effects we include separately the lagged stocks of male and female villagers employed in the United States.

Macroeconomic variables also may influence the benefits and costs of international migration. We control for several macroeconomic variables, including changes in the peso-dollar exchange rate and in per-capita GDPs of both countries. Mexican currency devaluations increase the purchasing power of dollars remitted in Mexico. Changes in U.S. GDP are included as a proxy for the availability of jobs in the United States. They are expected to be positively related to migration. The impact of an

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<sup>1</sup> Presidents Salinas and Bush (Senior) argued this to gain support for NAFTA.

increase in Mexico's GDP is ambiguous. On one hand, it could reflect employment growth that discourages migration. On the other hand, higher income in Mexico could provide households with the liquidity to finance investments, including investments in international migration, i.e., the costs associated with crossing the U.S.-Mexican border and establishing oneself in a foreign labor market. Our econometric analysis controls explicitly for these variables in order to isolate the impact of our three policy variables.

### **Theoretical Model**

At the micro level, international migration is only observed for households and family members that choose to participate in migration, which is a discrete decision. Migrants are individuals for whom the expected benefits of migration,  $R$ , exceed the (unobserved) migration "reservation wage,"  $\omega$ . The migration reservation wage depends on local opportunities on and off the farm. Following Mincer, the local wage is a function of human capital that affects the marginal productivity of labor. Let  $X_W$  denote a vector of human capital characteristics influencing wage income in the local labor market. The productivity of family members' on local farm and off-farm activities is shaped both by these human capital variables and by family assets  $\bar{K}$ . Remittances are a function of migrants' human capital, which affects earnings, as well as migrants' motivations to remit, which may be influenced by both human capital and family assets (Lucas and Stark, Taylor, 1985). Contacts at migrant destinations,  $\overline{MK}$ , are a form of migration capital that can enhance the labor-market prospects of migrants (Munshi, 2003).

Migrant remittances and reservation wages have both deterministic and stochastic components; thus,  $R = R(X_R) + u$  and  $\omega = \omega(X_\omega) + v$ , where  $X_\omega = [X_W, \bar{K}]$ ,  $X_R = [X_W, \bar{K}, \overline{MK}]$ , and  $u$  and  $v$  are stochastic errors. Letting  $\delta_i = 1$  if household member  $i$  migrates and 0 otherwise, the migration participation decision becomes:

$$(1) \quad \delta_i = \begin{cases} 1 & \text{if } \eta_i < R(X_R) - \omega(X_\omega) \\ 0 & \text{otherwise} \end{cases}$$

where  $\eta_i = v_i - u_i$ . Total migration is simply the sum of individuals who migrate; that is,  $M = \sum_i \delta_i$ . Let  $\theta_t$  represent the joint distribution of variables  $X_R$  and  $X_\omega$  in

community  $j$  at time  $t$ . Then

$$(2) \quad M_{jt} = M(\theta_{jt}, Z_{jt})$$

where  $Z_{jt}$  is a vector of community variables influencing the productivity of labor in local activities and remittances. In the econometric model, we control for the influences of  $\theta_{jt}$  and  $Z_{jt}$  by including lagged migration ( $M_{j,t-1}$ ), fixed effects for communities, and a time trend.

### **Data**

The data used to estimate the model are from a nationwide rural household survey carried out jointly by the University of California, Davis, and El Colegio de Mexico in Mexico City. The Mexico National Rural Household Survey (Encuesta Nacional a Hogares Rurales de Mexico, or ENHRUM) provides retrospective data on migration by individuals from a sample of rural households that is both nationally and regionally representative (see <http://precesam.colmex.mx>). Past studies of Mexican labor supply to

the U.S. employment sectors used proxies including border apprehensions (e.g., Torok and Huffman, 1986) or data from surveys of small numbers of villages. Usually surveys have not collected migration flows over extended periods of times and thus are unable to evaluate if policies have a long term impact on the dynamics and trends of migration. We hope that our dataset will fill this lacuna in the literature.

The survey was carried out in January and February 2003 in all of Mexico's five census regions (See Figure 1). INEGI (*Instituto Nacional de Estadística, Geografía e Informática*), Mexico's national census office, designed the sampling frame to provide a statistically reliable characterization of Mexico's population living in rural areas, defined by the Mexican government as communities with fewer than 2,500 inhabitants. For reasons of cost and tractability, individuals in hamlets or disperse populations with fewer than 500 inhabitants were not included in the survey. The survey was designed to be representative both nationally and regionally. The result is a sample of 1,760 households from 80 villages that is representative of more than 80 percent of the population that the Mexican census office considers to be rural.

The ENHRUM survey assembled complete migration histories from 1980 through 2002 for (a) the household head, (b) the spouse of the head, (c) all individuals who lived in the household 3 months or more in 2002, and (d) a random sample of sons and daughters of either the head or his/her spouse who lived outside the household longer than 3 months in 2002. In theory, the ENHRUM survey collected retrospective information only about labor migrants.<sup>2</sup> However, data on place of residence of all family members were gathered for 2002, regardless of employment status. For 2002, virtually all migrants were considered by family members in the village to be labor migrants. If the same is true for earlier years, our counts of labor migrants will also reflect total migration. For each year, the survey provides information on the migrant's sector of employment, agricultural or non-agricultural, and the state in which he/she worked. The survey provides the most reliable longitudinal data on migration from rural Mexican communities to the United States.

The survey asked individuals to recall employment information for each migrant from 1980-2002. Individuals may be unable to remember their (or their migrant sons' and daughters') employment histories for 22 years. However, when employment is coupled with a life event such as international migration, there is a smaller likelihood that data will be misreported. A study by Smith and Thomas (2003) showed that when respondents are asked to recall information linked to salient events, such as marriage or birth of a child, misreporting is insignificant. Also, individuals asked to recall labor or migration histories reported more accurately moves that involved either a long distance or extended stays.

Our sample includes 1,760 households from the 5 rural regions of Mexico.<sup>3</sup> To implement the survey, Mexico was divided into 5 regions, reflecting INEGI's standard regionalization of the country: Center, South-Southeast, Center-West, Northwest, and

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<sup>2</sup> The question asked in the survey was: "¿Qué años trabajó \_\_\_ en los E.U. desde 1980? ¿En qué trabajó y en qué estado? ¿Por un salario o por cuenta propia?" ("In what years did \_\_\_ work in the U.S. since 1980? In what job and which state? For a salary or self-employed?").

<sup>3</sup> One of these 5 regions, the West-Central, is the focus of Mexico Migration Project (MMP) surveys (Population Studies Center, University of Pennsylvania, Philadelphia (producer and distributor), [www.pop.upenn.edu/mexmig/welcome.html](http://www.pop.upenn.edu/mexmig/welcome.html)). The MMP surveyed a random sample of households within communities, but the sample of MMP communities is not random.

Northeast. Table 1 summarizes migration from households in rural Mexico. Sixteen percent of all households in the sample had a family member living in the United States at the start of 2002, the year of the survey, and 26 percent had a family member living in another part of Mexico. Many households had more than one migrant. The number of U.S. migrants per household ranged from 0 to 9, while the number of internal migrants ranged from 0 to 10. The average household in the sample had 0.35 U.S. migrants and 0.71 internal migrants in 2002—or 1.06 migrants in total.

As indicated in the Table, there are sharp differences in migration experience among the five rural regions. West-Central Mexico traditionally has been the largest sender of migrants to the United States, with far and away the highest current participation in international migration and the most international migration experience. In this region, nearly 28% of all households have at least one family member in the United States, and the average household had .62 U.S. migrants. By contrast, 7.5% of households in the south-southwest have U.S. migrants, with an average of .10 U.S. migrants per household.

### ***Econometric Model***

We econometrically estimate the impact of policy reforms for three dependent variables: 1) the share of villagers employed in the U.S.; 2) the share of female villagers employed in the U.S.; and 3) the share of male villagers employed in the U.S.

For the first dependent variable we estimate two fixed-effects panel data models for all employed immigrants. The first model is intended to capture the basic dynamics of rural Mexico-to-U.S. migration. The share of village population observed as labor migrants in the United States at time  $t$ ,  $M_{jt}$ , is regressed on the same share lagged one year ( $M_{jt-1}$ ) and a time trend ( $t$ ), controlling for a vector of village fixed effects,  $\alpha_j$ :<sup>4</sup>

$$(3a) \quad M_{jt} = \alpha_j + \gamma t + \delta M_{jt-1} + u_{jt}$$

Essentially, equation (3a) estimates the dynamic nature of employed migration over time and allows us to evaluate the role of networks and the inertia of employed migration over time. Village fixed effects,  $\alpha_j$ , control for community-specific migration dynamics that shape the migration trend; each village is permitted to have its own migration trajectory. The village fixed-effects model makes it possible to isolate the underlying migration trend ( $t$ ) and influence of networks and inertia (the lagged-migration variable) from policy and macroeconomic variables shaping migration.

Model two includes three policy variables: dummy variables for IRCA (1 for all time periods beginning in 1986, the year of IRCA's implementation) and NAFTA (1 beginning in 1994, 0 before), and a continuous variable measuring the percentage change in border enforcement expenditures ( $\Delta BE_t$ ). We also include macroeconomic variables, including the percentage changes between time  $t$  and  $t-1$  in the peso-dollar exchange rate ( $\Delta ER_t$ ) and the US and Mexico GDPs ( $\Delta USGDP_t, \Delta MGDP_t$ ).

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<sup>4</sup> We use the percentage rather than the sum of villagers who migrated because of our concern that the size of village populations in the synthetic cohorts created using retrospective data is biased downward as one goes back in time, as individuals are removed from the population due to death (and thus are not available to be counted in 2003).

$$(3b) \quad M_{jt} = \alpha_j + \gamma t + \delta M_{jt-1} + \beta_1 IRCA_t + \beta_2 NAFTA_t + \beta_3 \Delta BE_t + \theta_1 \Delta ER_t + \theta_2 \Delta USGDP_t + \theta_3 \Delta MGDP_t + u''_{jt}$$

Our use of dummy variables to evaluate the impact of NAFTA and IRCA warrants some explanation. Other methods are possible, including controls for trade flows or changes in real wages in Mexico and the U.S. But these variables may not be exogenous to migration. For example, changes in real agricultural wages in the United States clearly are related to the supply of migrant labor from rural Mexico. The inclusion of policy dummy variables in the regressions makes it possible to evaluate the long run impact of exogenous policy shocks on the rate and dynamic of labor migration. Inclusion of the lagged-migration variable allows for the impacts of policy shocks to unfold gradually over time. The vector of fixed effects,  $\alpha_j$ ,  $\delta$ ,  $\beta_k$ ,  $k = 1, \dots, 3$  and  $\theta_l$ ,  $l = 1, \dots, 4$  are parameters to be estimated, and  $u_{jt}$ ,  $u'_{jt}$ , and  $u''_{jt}$  are stochastic errors. Under the null hypothesis of no policy impacts on migration the coefficients  $\beta_k = 0 \forall k$ .

We estimate three types of dynamic fixed-effects models for labor migration by gender (g=m,f). The first model is similar to equation (3a). We estimate the share of males (females) in employed migration as a function of a time trend (t) and the lagged share of male (female) migrants in village populations:

$$(4a) \quad M_{jgt} = \alpha_j + \gamma t + \delta_1 M_{jgt-1} + u''_{jt}$$

We re-estimated each equation including the lagged stock of other-gender migrants to evaluate the gender sensitivity of networks:

$$(4b) \quad M_{jsgt} = \alpha_j + \gamma t + \delta_1 M_{jmt-1} + \delta_2 M_{jft-1} + u''_{jt}$$

If both male and female networks shape female migration, then both  $\delta_1$  (lagged-female migration) and  $\delta_2$  (lagged-male migration) will be significant, and conversely for the male migration regression.

In the final estimation we estimate include, in each gender-specific migration equation, all our policy and macroeconomic variables. This model is similar to equation (3b) except that it includes the lagged stock of other gender's participation in employed migration:

$$(4c) \quad M_{jsgt} = \alpha_j + \gamma t + \delta_1 M_{jmt-1} + \delta_2 M_{jft-1} + \theta_1 \Delta IRCA_t + \theta_2 \Delta NAFTA_t + \theta_3 \Delta BE_t + \beta_1 \Delta ER_t + \beta_2 \Delta USGDP_t + \beta_3 \Delta MGDP_t + u''_{jt}$$

### ***Estimation and Results***

The series of models given by 3(a)-3(b) and 4(a)-4(c) was estimated using the standard least square dummy variable (LSDV) method. This method results in a downward bias in the estimate of  $\delta$ ; however, this bias diminishes as the number of observations in the time dimension (T) increases (Judson and Owen, 1999). Judson and Owen show that the bias becomes negligible at T approaches 30. In our dataset, T=23 (1980-2002, inclusive), which indicates that there will be some bias, but it will be small. Judson and Owen also conclude that when T is greater than 20 the bias in the other parameter estimates is negligible. The key hypotheses that we wish to test involve not  $\delta$ , but rather, the other parameters in the model (i.e., the effects of policy variables).

Figure 2 presents estimated shares of populations from the surveyed villages in U.S. farm and nonfarm jobs from 1980 to 2002. It shows an upward trend in migration to

the United States for both males and females. However the trend is steeper for males than females.<sup>5</sup> Female migration is steady, lower than male migration, and has a steady increase over the 23-year period. Table 2 presents variable definitions and means for variables used in the econometrics.

The data set for this sample provides information on migration from 80 villages over 23 years (from 1980 to 2002); however, one year (80 observations) was lost as a result of lagged right-hand-side variables. Thus, the total sample size is 1,760 (22 years x 80 villages).

Table 3 reports the econometric results for the village share of labor migrants using ordinary least squares, controlling for fixed effects. Column 1 of Table 3 shows results from the model that only controls for the lagged stock of migration. The time trend is significant and positive. The coefficient on lagged migration is also significant and positive, indicating that networks created through past migration influence current migration.

Macroeconomic variables and policies, such as NAFTA and IRCA, can change the overall trend and influence of migration rates. Results from Model 2, which includes these variables, are reported in column 2 of Table 3<sup>6</sup>. The time trend and coefficient on lagged migration remain positive, large and significant. The dummy variables for NAFTA and IRCA have a significant and negative impact on migration. Therefore, the supply of migrant labor from rural Mexico to the U.S. decreases after the implementation of IRCA in 1982 and once again following NAFTA in 1994. These findings support the hypothesis that NAFTA relieved migration pressures. IRCA also curtailed migration, but to a smaller degree than NAFTA. In contrast to these policies, the increase in border enforcement *increased* the share of villagers working in the United States. This finding supports the hypothesis that migrants continue to enter the U.S. and/or that increased border enforcement discourages return migration.

The macroeconomic variables are all significant and larger in magnitude than the policy variables. A one percentage change in the peso-dollar exchange rate devaluation increases the rate of migration. This is of the expected sign, inasmuch as the devaluation raises the returns to migration (remittances) in pesos. Changes in both country GDPs increase migration. Economic expansion in the U.S. pulls rural Mexicans into the country. GDP growth in Mexico encourages migration. This finding is consistent with the argument that income growth enables rural households to finance the cost of crossing the border and establishing migrants in U.S. labor markets.

Gender-specific regression results appear in Table 4, for female migrants, and Table 5, for male migrants. The column labeled “Model I” in each table shows results for the basic dynamic regression model. When we estimate the migration model by gender, the trend remains significant and positive but flat. However, the lagged migration participation rate is both statistically and quantitatively significant. When the lagged migration rate of males is added to the estimate of female migration (Model II in Table

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<sup>5</sup> The surge in migration to the United States in the 1990s is mirrored in U.S. Census 2000 data. The U.S. Census does not provide information on where migrants originate in Mexico (e.g., from rural or urban areas). However, they show an unexpectedly large increase in Mexico-born persons living in the United States.

<sup>6</sup> Trends in migration may also be influenced by regional dynamics. When separate repressors are estimated for Mexico’s five census regions, there were no differences in the signs of significant variables among regions for all of the models presented in this paper.

4), it has no significant impact. That is, there is no evidence of significant cross-gender network effects on female migration. The effect of lagged female migration on male migration (Table 5) is positive and significant at the 10% level, but it is quantitatively small. A one-percentage-point increase of the share of village females working in the United States, other things being equal, is associated with a .07-percent increase in male participation in international migration. This finding suggests that labor migration networks are gender specific. That is, estimates of a gender's participation in labor migration do not improve appreciably when the other gender's migration network is included in our regressions.

Macroeconomic and policy variables (Model III in Tables 4 and 5) significantly increase the predicative power of the migration models for both genders.<sup>7</sup> Qualitatively, policy changes have similar effects on male and female migration, with the exception of GDP growth. However, quantitatively the results differ according to gender. The decrease in the male migrant share after NAFTA is three times greater than the drop in the female share. This suggests that female migration was more resilient to NAFTA-related policy changes. In contrast, IRCA appears to have been a greater deterrent to female than male migration. Increased border enforcement has a larger positive influence on male migration than on female migration (.022 and .010, respectively). This indicates either that border enforcement increases male stays more than female stays, or perhaps more plausibly, that border controls are more of a deterrent to border crossings by females than males.

The effects of changes in Mexico's GDP are significant only for female migration. By contrast, changes in US GDP are significant only for male migration. These findings may suggest that female migration is more sensitive to liquidity constraints that can be loosened by income growth in Mexico, while female migrant labor demand is robust to U.S. GDP growth. Male migration, on the other hand, appears to be sensitive to U.S. economic growth.

Due to mortality, some (mostly older) individuals disappear from our synthetic cohorts of migrants and villagers as we go back in time—that is, they are not alive to be counted at the time of the survey. If old villagers are less likely to migrate, this will result in an upward bias in the estimated share of villagers in the United States (and thus a downward bias in the estimated migration trend), and this bias will be larger the farther back in time one goes. We explored the ramifications of this bias by estimating the model separately for younger age cohorts of villagers (i.e., those who were 16-35 years old in 1980 and thus less likely to be affected by mortality over the study period). There were no significant changes in our econometric findings.

## Conclusions

The impacts of NAFTA, IRCA and increased U.S. border enforcement are ambiguous a-priori. Each policy change potentially has both positive and negative influences on migration. In the past, data limitations have made it difficult to test for impacts of policy shocks on Mexico-to-U.S. migration dynamics. The Mexico National Rural Household Survey provides retrospective migration histories from a nationally random sample of rural Mexicans. This makes it possible to isolate migration trends and

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<sup>7</sup> An F test of restricted versus unrestricted regressions, for both male and female migration shares rejected the joint hypothesis that the macroeconomic and policy variables were jointly zero.

control for place-of-origin characteristics while measuring the impacts of policy shocks on the share of rural Mexicans working in the United States.

Several general findings emerge from our analysis. First, international labor migration from rural Mexico has followed an upward trend from 1980-2002 but is driven overwhelmingly by past migration, reflecting the central role of migration networks. Second, policy variables significantly influence migration, but not as much as macroeconomic variables. NAFTA and IRCA had some impact on curtailing migration; however, increased border enforcement appears to have the opposite effect. No policies are able to counteract the effects of a changing macroeconomic environment. Third, the influences of both policy and macroeconomic variables are small compared with network effects embodied in past migration.

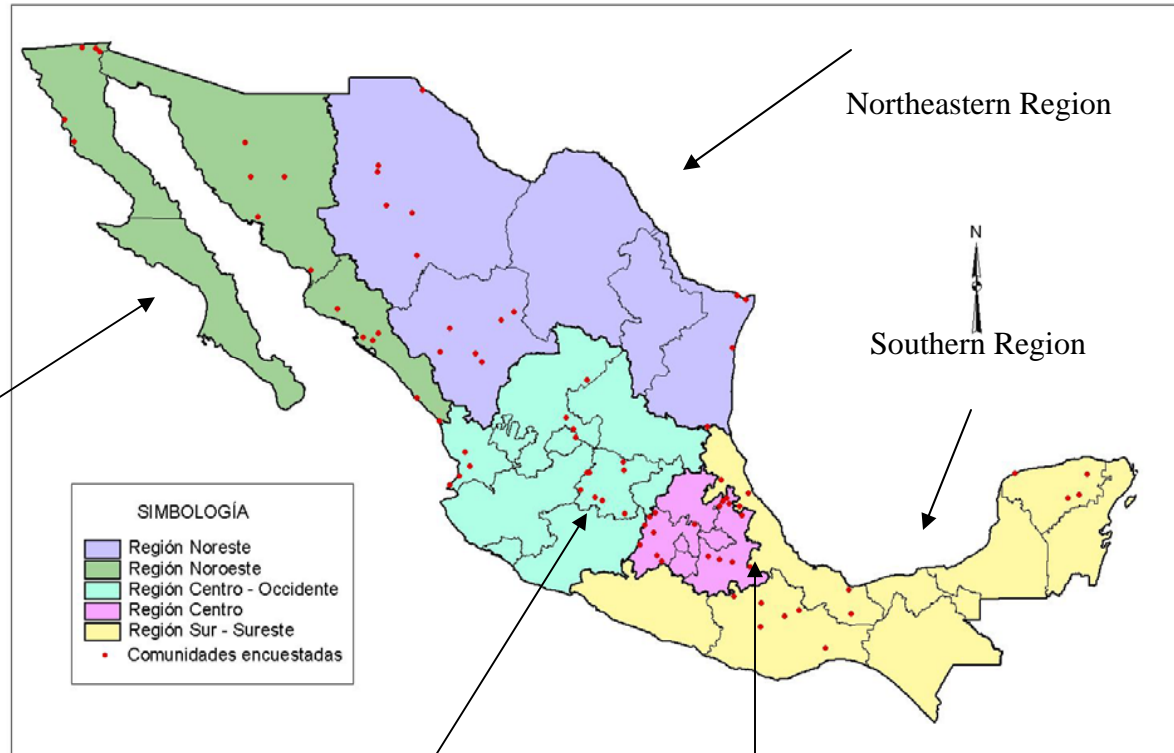
A unique contribution of this analysis is the insight it offers into the dynamics underlying female and male migration. Policy shocks and macroeconomic variables have differential effects on female and male migration, quantitatively (in the case of NAFTA and IRCA) and in some cases qualitatively (in the case of Mexico and U.S. GDP growth). The role of Mexico GDP growth in loosening liquidity constraints on migration appears to be more important for females than males, while the impact of U.S. income growth is greater for males.

Although own-gender migration networks are significant and large, cross-gender network effects are small or nonexistent. Past research has suggested that female migrants follow males, for example, for purposes of family reunification. However, we find that past labor migration by male villagers has no significant effect on female labor migration. That is, *controlling for community effects and long run migration dynamics*, labor migration networks are gender specific. Future economic research is warranted on gender asymmetries in networks and their influence on migration propensities.

ENCUESTA NACIONAL A HOGARES RURALES DE MÉXICO (ENHRUM), 2003  
REGIONES Y DISTRIBUCIÓN GEOGRÁFICA DE COMUNIDADES ENCUESTADAS

**Figure 1:**  
ENHRUM 2003 Regions  
and Distribution of  
Communities Surveyed

Northwestern  
Region



Northeastern Region

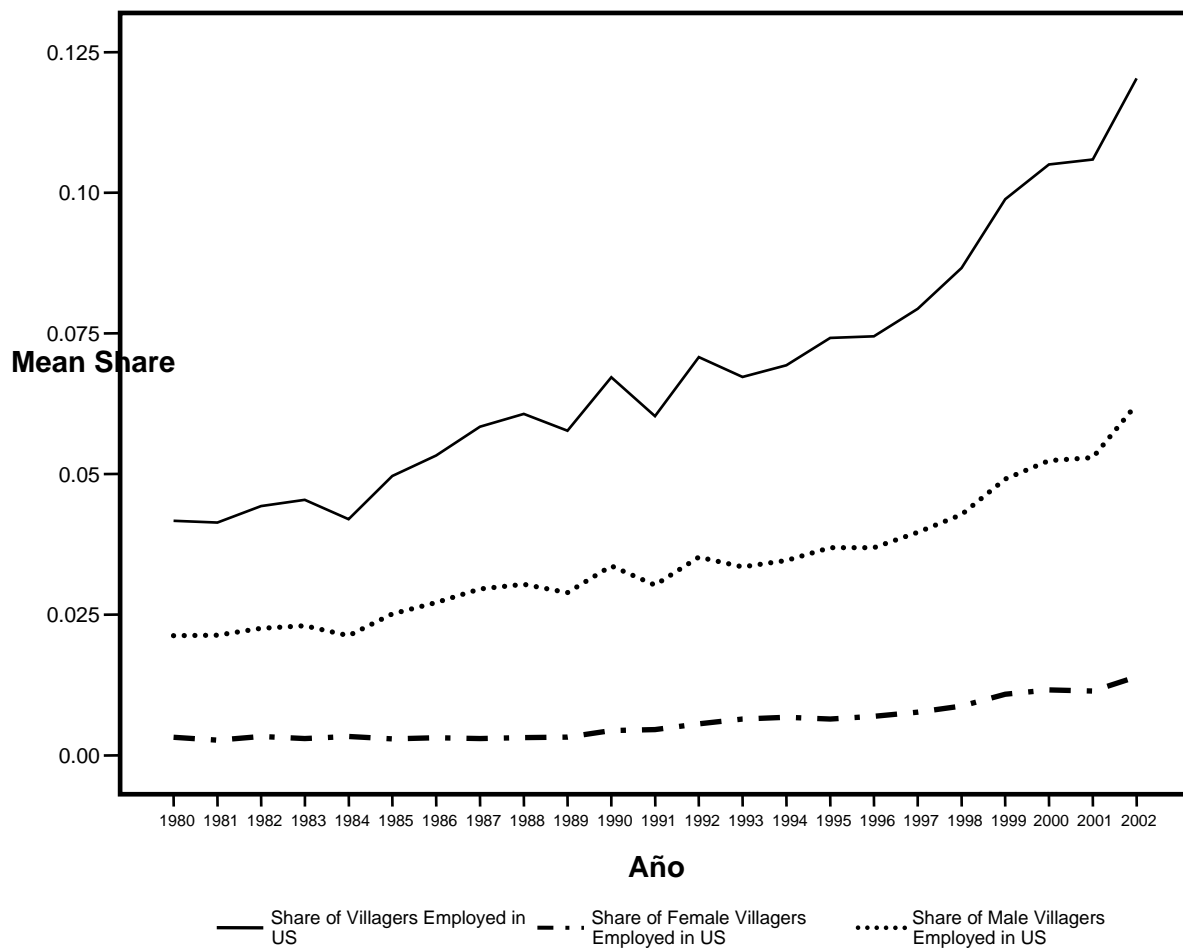
Southern Region

West Central Region

Central Region

Fuente: Elaborado en el Laboratorio de Análisis Espacial, Coordinación de Servicios de Cómputo, El Colegio de México con base en datos de la ENHRUM, 2003.

**Figure 2: Average Share of Villagers Employed in United States: 1980-2002**



**Table 1. Migration Summary Statistics for Rural Mexico, by Region**

<b>Region</b>	<b>Variable</b>	<b>Percentages</b>	<b>Sample Mean</b>	<b>Standard Deviation</b>
<b>South-South East</b>	Households with US migrants (%)	7.53%	-	0.26
	US Migrants per Household		0.10	0.42
	Household Sample Size		372	
<b>Center</b>	Households with US migrants (%)	14.52%	-	0.35
	US Migrants per Household		0.27	0.89
	Household Sample Size		365	
<b>Center-West</b>	Households with US migrants (%)	27.75%	-	0.45
	US Migrants per Household		0.62	1.29
	Household Sample Size		346	
<b>Northwest</b>	Households with US migrants (%)	12.09%	-	0.33
	US Migrants per Household		0.23	0.79
	Household Sample Size		339	
<b>Northeast</b>	Households with US migrants (%)	19.72%	-	0.40
	US Migrants per Household		0.54	1.43
	Household Sample Size		360	
<b>Total</b>	Households with US migrants (%)	16.22%	-	0.37
	US Migrants per Household		0.35	1.04
	Household Sample Size		1782	

Source: ENHRUM, 2003

Table 2. Variable Definitions and Means

<b>Variable</b>	<b>Description</b>	<b>Mean</b>
T	Time Trend	11
USMIG	Share of Villagers Employed in U.S.	.0403
Female Migration	Share of Female Villagers Employed in U.S.	.0059
Male Migration	Share of Male Villagers Employed in U.S.	.0344
% Change ER	% change in Peso-Dollar exchange rate from previous year	.0138
% Change Border Control	% change in INS border enforcement expenditures in millions of 2000 US\$	.138
NAFTA	Dummy variable = 1 beginning in 1994	0.39
IRCA	Dummy variable = 1 beginning in 1986	0.74
% Change MGDGP	% change Mexico per capita GDP	.025
% Change US GDP	% change US per capita GDP	.0304

**Table 3.**  
**OLS Coefficients for three Dynamic Models—Participation in Migration**  
**(standard errors in Parenthesis)**

Variables	All Regions	
	Model I	Model II
Constant	-.006 (.003)**	-.013 (.004)**
T	.001 (.000)**	.001 (.000)**
USMIGLAG	.846 (.015)**	.845 (.015)**
NAFTA		-.008 (.002)**
IRCA		-.003 (.001)**
% Change Border Control		.011 (.003)**
% Change ER		.015 (.004)**
% Change MGDP		.034 (.016)**
% Change US GDP		.051 (.022)**
R <sup>2</sup>	.947	.948

Dependent Variable: Weighted total of international workers in village

All models were estimated with village fixed effects.

N=1759. \*\*Significance at .05 level \* Significance at .1 level

**Table 4.**  
**OLS Coefficients for three Dynamic Models—Female Participation in Migration**  
**(standard errors are in parentheses)**

Variables	All Regions		
	Model I	Model II	Model III
Constant	-.001 (.001)	-.001 (.001)	-.003 (.001)**
T	.000 (.000)**	.000 (.000)**	.000 (.000)**
Female Migration Lag	.881 (.014)**	.875 (.015)**	.874 (.015)**
Male Migration Lag		.008 (.006)	.007 (.008)
NAFTA			-.002 (.001)**
IRCA			-.001 (.000)*
% Change Border Control			.002 (.001)**
% Change ER			.003 (.001)**
% Change MGDGP			.011 (.005)**
% Change US GDP			.009 (.007)
R <sup>2</sup>	.882	.882	.883

Dependent Variable: Weighted total of female international workers in village

All models were estimated with village fixed effects.

N=1759. \*\*Significance at .05 level \* Significance at .1 level

**Table 5**  
**OLS Coefficients for three Dynamic Models—Male Participation in Migration**  
**(standard errors are in parentheses)**

Variables	All Regions		
	Model I	Model II	Model III
Constant	-.005 (.003)*	-.005 (.003)*	-.010 (.003)**
T	.000 (.000)**	.000 (.000)**	.001 (.000)**
Male Migration Lag	.822 (.016)**	.814 (.017)**	.815 (.017)**
Female Migration lag		.075 (.042)*	.068 (.042)
NAFTA			-.006 (.002)**
IRCA			-.002 (.001)*
% Change Border Control			.010 (.002)**
% Change ER			.012 (.004)**
% Change MGDG			.023 (.015)
% Change US GDP			.042 (.019)**
R <sup>2</sup>	.947	.948	.948

Dependent Variable: Weighted total of male international workers in village

All models were estimated with village fixed effects.

N=1759. \*\*Significance at .05 level \* Significance at .1 level

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