

***Fiscal Divergence and Business Cycle Synchronization:
Irresponsibility is Idiosyncratic***

Zsolt Darvas, Andrew K. Rose and György Szapáry*

Draft Revised as of: May 20, 2005

Preliminary: Comments Welcome!

Abstract

Using a panel of 21 OECD countries and 40 years of annual data, we find that countries with similar government budget positions tend to have business cycles that fluctuate more closely. That is, fiscal convergence (in the form of persistently similar ratios of government surplus/deficit to GDP) is systematically associated with more synchronized business cycles. We also find evidence that reduced fiscal deficits increase business cycle synchronization. The Maastricht “convergence criteria,” used to determine eligibility for EMU, encouraged fiscal convergence and deficit reduction. They may thus have indirectly moved Europe closer to an optimum currency area, by reducing countries’ abilities to create idiosyncratic fiscal shocks. Our empirical results are economically and statistically significant, and robust.

Keywords: European; monetary; union; policy; Maastricht; criteria; optimum; Mundell.

JEL Classification Numbers: F42

Zsolt Darvas, Math. Econ.
Fovam ter 8
Corvinus University
H-1093 Budapest, Hungary
Tel: +36 (1) 482-5279
Fax: +36 (1) 482-5029
zsolt.darvas@uni-corvinus.hu

Andrew K. Rose
Haas School of Business
University of California
Berkeley, CA USA 94720-1900
+1 (510) 642-6609
+1 (510) 642-4700
arose@haas.berkeley.edu

György Szapary
Szabadsag ter 8-9
Magyar Nemzeti Bank
H-1850 Budapest Hungary
+36 (1) 428 2770
+36 (1) 428 2581
szaparygy@mnbb.hu

* Darvas is Assistant Professor, Department of Mathematical Economics and Economic Analysis, Corvinus University. Rose is B.T. Rocca Jr. Professor of International Trade and Economic Analysis and Policy in the Haas School of Business at the University of California, Berkeley, NBER research associate and CEPR Research Fellow. Szapary is deputy governor of the Magyar Nemzeti Bank. The idea of this paper stemmed from a conversation with Miriam Green. Rose thanks the Banco de España for hospitality during the course of this research. For comments we thank: Dave Backus, Marianne Baxter, Peter Benczur, Paul de Grauwe, Antonio Fatás, Jeff Frankel, Sebnem Kalemli-Ozcan, Pat Kehoe, Roberto Perotti, Eric van Wincoop, and Charles Wyplosz. The views expressed below do not represent those of the Magyar Nemzeti Bank or its staff. A current version of this paper, key output, and the STATA 8 data set used in the paper are available at <http://faculty.haas.berkeley.edu/arose>.

1. Introduction

In 1998, European countries qualified for entry into European Monetary Union (EMU) on the basis of five “convergence criteria.” The criteria were enshrined in the 1992 Maastricht Treaty and quantify targets concerning inflation, long-term bond yields, exchange rates, government debt, and the government budget. The Maastricht convergence criteria are of more than of historical relevance, since they will also be applied to future EMU entrants. Further, the 1997 “Stability and Growth Pact” implies that the fiscal criteria are still, in principle, binding.¹

Most economists – particularly non-Europeans – view the Maastricht convergence criteria with skepticism. The reason is simple: they have little to do with standard economic arguments concerning “optimal currency areas,” monetary unions that are desirable and sustainable. The consensus in economics is that from a theoretical viewpoint, monetary unions make sense for countries with synchronized business cycles, integrated markets, flexibility, and mechanisms to share risk. The overlap between the Maastricht convergence criteria and the optimum currency area criteria is small.²

Clearly the direct correspondence between the (Maastricht) criteria actually applied for EMU entry and the appropriate (optimum currency area) criteria is poor. In this paper we ask if there is an *indirect* connection. We focus on the most controversial Maastricht criteria – the total government budget deficit/GDP ratio – and link it empirically to arguably the most important optimum currency area criterion, namely the synchronization of business cycles. Using a panel of data that includes twenty-one countries and forty years of data, we show that countries with divergent fiscal policies (i.e., large average cross-country differences in the ratio of general government net lending/borrowing to GDP) tend to have less synchronized business cycles. We estimate that each percentage point of fiscal divergence between a pair of countries tends to lower the correlation coefficient of their business cycles by between .03

and .12. This effect is both statistically and economically significant. We also show that reduced *levels* of primary fiscal deficits (or increased primary surpluses) tend to increase the level of business cycle synchronization, though the evidence for this effect is somewhat weaker.

A concrete example may clarify things. When the Maastricht Treaty was signed in 1992, the total Italian budget deficit was 10.7% of GDP, and had been hovering at or above 11% of GDP for a decade. This was in sharp contrast to the typical German deficit, which was 2.6% of GDP in 1992.³ The drive to enter EMU – that is, to satisfy the Maastricht criteria – encouraged this gap to shrink by around eight percentage points; by the 1999 start, Italy's budget deficit had fallen to 1.7%, similar to the German deficit of 1.5%. In this paper, we ask: could such fiscal convergence have an effect on the synchronization of business cycles between Germany and Italy? Alternatively, the (cross-country) standard deviation of the government budget position/GDP ratio was 4.1% for the EURO-12 in 1991, and only 2.1% in 1999; did this convergence in fiscal positions affect business cycle synchronization at the start of EMU?⁴ We find that the answer is generally positive; a larger panel of OECD data indicates that fiscal convergence (in either the total or primary budget balance) is systematically associated with more synchronized economic activity. *Whether or not it was intentional, the application of the Maastricht convergence criteria may have moved the EMU entrants closer to being an optimum currency area, since fiscal convergence tends to synchronize business cycles!*

We stress at this point that we know of no theoretical model formally linking fiscal convergence to business cycle synchronization. Still, we do not think it is difficult to understand our results. Fiscal convergence, by our definition, usually occurs because a country that has been fiscally irresponsible – that is, a country that has run persistently high budget deficits – reforms and closes the fiscal gap with other countries. Intuitively, countries

that are fiscally irresponsible – i.e., countries that run persistently high budget deficits – are also countries that create idiosyncratic fiscal shocks. (This seems a natural association to us; irresponsible behavior is often idiosyncratic, for individuals as well as fiscal authorities.) In this case, reducing the budget deficit of a country simultaneously reduces its scope for idiosyncratic fiscal shocks, raising the coherence of its business cycle with the business cycle of others. That is, fiscal convergence raises business cycle synchronization since responsible fiscal behavior tends to be less idiosyncratic fiscal behavior.

We mention in passing that we know of no deliberate intent on the part of the creators of the Maastricht convergence criteria to affect the optimum currency criteria, either directly or indirectly. Our effect seems to have been an unintended side-benefit to the convergence process.

In section 2 we describe our methodology. Our results on the link between fiscal convergence and business cycle synchronization are presented in section 3; we link budget deficits to business cycle volatility more directly in the following section. The paper ends with a brief conclusion.

2. Methodology

What should the effect of persistent fiscal divergence be on business cycle synchronization? To our knowledge, there is no formal treatment of this topic in the extant literature.

Countries are subject to asymmetric shocks (e.g. exchange rate and/or wage shocks). Further, similar shocks (e.g. oil price shocks) can have asymmetric effects across countries because of differing propagation mechanisms. If these asymmetries are persistent, and are partially offset with discretionary fiscal policy or automatic fiscal stabilizers, then fiscal divergence can, in principle, be associated with greater business cycle synchronization. For

example, suppose that Austria and Belgium begin with identical budget positions and perfectly synchronized business cycles. Austria receives a persistent negative shock, and responds with expansionary fiscal policy that neutralizes any effect on its cycle. In this case, Austria's business cycle remains synchronized with the Belgian economy *ceteris paribus*, while the Austrian deficit diverges from the Belgian.

However, fiscal policy can also be a source of shocks. Suppose that Austrian fiscal policy expands in the absence of shocks to either Austria or Belgium, and generates an Austrian expansion. In this case fiscal divergence will be associated with reduced business cycle synchronization.

From a theoretical viewpoint then, the matter is ambiguous. If fiscal policy divergence is a response to asymmetric shocks then it may be associated with enhanced business cycle coherence; if fiscal shocks themselves cause business cycles, then the opposite may be true. Without persistent shocks (or shocks with persistent effects), there may be no relationship at all between fiscal policy divergence and business cycle synchronization. The question is thus ultimately empirical. While the absence of a formal structural framework makes us uncomfortable, we see no alternative but to take the issue to the data.

The literature gives only a few hints about the matter. Several authors argue that a world business cycle exists (e.g. Gerlach, 1988; Lumsdaine and Prasad, 1997; Darvas and Szapary 2004; Canova et al. 2004), consistent with the absence of important asymmetries. Fatás and Mihov (2003a) studied discretionary fiscal policy for 91 countries and conclude (p1419) "governments that use fiscal policy aggressively induce significant macroeconomic instability" i.e., output volatility. Similarly, Fatás and Mihov (2004) study the American states and conclude that budgetary restrictions lead to lower fiscal policy volatility and smoother business cycles; they conclude (p23) that "fiscal policy is a significant source of business cycle volatility among US states, and, as a result, constraints on politicians leads to

less volatile economic fluctuations.” Perhaps the work closest to ours is that of Kose et al. (2003) who study determinants of coherence of a country’s business cycle with a global business cycle. One interpretation of their findings (p 62) is that “fiscal policies exacerbate country-specific fluctuations.”

Still, to our knowledge, no one has explored the link between differences of national fiscal policies and the synchronization of their business cycles. We now turn to that task.

Empirical Framework

We are interested in investigating the empirical linkages between persistent cross-country differences in the fiscal policy and business cycle synchronization. We are also interested in the effects of the average cross-country *level* of aggregate fiscal policy on business cycle synchronization.⁵

Our primary measure of fiscal divergence is the difference between countries in the general government budget surplus (+) or deficit (-), measured as a percentage of national GDP. In 1999, the Austrian deficit was 2.3% of GDP, while the Belgian deficit was .4%. Thus our measure of Austrian-Belgian fiscal divergence in 1999 is 1.9%. Taking the average of this over a decade of annual data yields our measure of fiscal divergence (.98 for average Austrian-Belgian fiscal divergence during 1994-2003). That is, we measure fiscal divergence as:

$$\text{FiscalDiverge}_{ij\tau} \equiv .1 * \sum_{\tau} (|\text{Budg}_{it} - \text{Budg}_{jt}|)$$

where Budg_{it} is the general government budget surplus (+) or deficit (-) at time t expressed as a percentage of nominal GDP for country i , and the averaging is done over a decade of annual

data. A larger value of FiscalDiverge corresponds to higher average divergence between the fiscal positions of the two countries over the relevant period of time.

The total government budget position as a percentage of GDP is of great relevance; the Maastricht convergence criteria focus on this measure. However, we also examine the analogue using the cyclically adjusted primary budget position (also measured as a percentage of GDP). Since the primary balance excludes interest payments (and thus the impact of the government debt level), it better captures discretionary fiscal policy (as well as acting as a robustness check).

We note that our measure of international fiscal divergence indicates little about the pro- or counter-cyclic nature of national fiscal policy. A standard argument used against the Stability and Growth Pact is that countries that are constrained to have the same monetary policy should have good access to counter-cyclic fiscal policy. But the average *level* of the budget deficit is unrelated to its counter- or pro-cyclic stance, especially when the data is smoothed over a decade. Countries that use fiscal policy counter-cyclically sometimes have persistent deficits, but so do countries with pro-cyclic fiscal policy.⁶ In any case, our focus is on the average *difference* between fiscal positions.

Fiscal policy was highly divergent at the signing of the Maastricht Treaty. In 1992, four European countries had total government budget deficits in excess of 6% of GDP (Belgium 8%; Greece 12.2%; Italy 10.7%; and UK 6.5%), while another four had deficits of less than 3% of GDP (Austria 1.9%; and Denmark 2.2%; Germany 2.6%; Luxembourg .3%).⁷ The Maastricht treaty encouraged fiscal convergence since it pointed potential EMU entrants towards lower deficits. For this reason, we find it interesting to determine the consequences, if any, of fiscal convergence. But clearly the treaty encouraged members to converge to *lower* deficits (of no more than 3% of GDP), not to similar deficits irrespective of their level.

Accordingly, we also examine the effect of the average cross-country *level* of the total government budget deficit, measured as a percentage of GDP. We measure this by:

$$\text{AvgFiscal}_{ijt} \equiv .1 * \sum_{\tau} (\text{Budg}_{it} + \text{Budg}_{jt}) / 2$$

Again, we also examine the analogy for the primary budget position.

Our other important variable is business cycle synchronization. We focus on this because it is arguably the most important criteria of the traditional Mundell optimum currency area criteria. Regions with more synchronized business cycles have less need of individual monetary policies, and are thus better candidates for currency union. While it is by no means the only criteria, it seems a natural place to search for an overlap between Maastricht and Mundell.

We are interested in the bilateral correlation between real activity in country *i* and country *j* over decade τ . There is no obvious single measure for this; accordingly, we construct a number of proxies. We begin by using two standard measures of real economic activity: (the natural logarithm of) real GDP and the unemployment rate. We then de-trend the variables so as to focus on business cycle fluctuations (i.e., the combination of shocks and propagation mechanisms), in two different ways: 1) we take simple first-differences of annual variables; 2) we use the well-known Hodrick-Prescott (“HP”) filter (with the standard annual smoothing parameter of 100). After de-trending our variables over the entire available sample, we are able to compute bilateral correlations for real activity. These correlation coefficients are estimated (for a given concept of real economic activity and de-trending technique), *between* two countries *over* a given span of time. Thus, for instance, we estimate the correlation between (HP-de-trended real) Austrian and Belgian GDP, between 1964 and 1973. We also investigate a number of other measures of business cycle synchronization

below to ensure that our results are insensitive to the underlying measure of economic activity, the de-trending technique, etc. Thus we also use industrial production, we de-trend with the Baxter-King “BK” (1999) band-pass filter, and so forth.

The Data Set

Our sample includes twenty-one OECD countries; these are listed in Appendix Table A1. We stick to the OECD *Economic Outlook* data set because it is both high quality and the most relevant for e.g., questions concerning EMU.⁸ Our underlying data set consists of annual observations (though with some gaps); we also use quarterly data (which has more holes) as a robustness check. The data set spans 1964 through 2003, which we then split into four decades (1964-73, 1974-83, 1984-93, and 1994-2003). We are thus left with a panel of data; the maximum possible sample size is 840 observations; 210 bilateral country-pair “dyads” $[(21 \times 20) / 2]$, with four decadal observations per country-pair.⁹ Descriptive statistics for key variables are provided in Appendix Table A2.

Figure 1 provides a set of four simple scatterplots of our four default measures of business cycle synchronization (GDP/Unemployment, differenced/HP-filtered) graphed against budget divergence. Non-parametric data smoothers are also provided in the graphs; these demonstrate a loose negative relationship between the two variables. Figures 2 and 3 are analogues that portray observations from the most recent (1994-2003) decade and EMU members respectively. Figure 4 is the analogue that portrays divergence in the *primary* (instead of the total) fiscal balance. Finally, Figures 5 and 6 are scatterplots of business cycle synchronization against the average cross-country levels of the total and primary budget positions respectively. There is reasonably consistent ocular evidence of a negative relationship between fiscal divergence and business cycle synchronization. However, there is

no sign of a strong link between the latter and the average total fiscal *level*, though the correlations are higher for the average *primary* budget position.

Estimation

Our general empirical strategy follows that of Frankel and Rose (1998) who focused on the endogeneity of business cycle synchronization with respect to trade.

The benchmark regressions we estimate are non-structural and take the simple form:

$$\text{Corr}(v,s)_{i,j,\tau} = \alpha + \beta \text{FiscalDiverge}_{i,j,\tau} + \varepsilon_{i,j,\tau}.$$

$\text{Corr}(v,s)_{i,j,\tau}$ denotes the correlation coefficient between country i and country j over decade τ for activity concept v (corresponding to log real GDP or the unemployment rate), de-trended with method s (corresponding to differencing or HP-filtering). $\text{FiscalDiverge}_{i,j,\tau}$ denotes the average (over decade τ) absolute difference in the government budget position (measured as a percentage of national GDP) between countries i and j . Finally, $\varepsilon_{i,j,\tau}$ represents the myriad influences on bilateral activity correlations above and beyond the influences of fiscal divergence (hopefully unrelated to our regressor), and α and β are the regression coefficients to be estimated.

The object of interest to us is the slope coefficient β . A negative estimate of β indicates that an increase in fiscal divergence is associated with reduced business cycle coherence. That is, fiscal policy convergence is linked to more synchronized business cycles.

A simple OLS regression of bilateral activity income correlations on fiscal divergence might be inappropriate for a couple of reasons. First, there may be non-trivial measurement error in fiscal divergence (especially since measuring the general government budget position itself seems difficult). A potentially more important worry is simultaneity. Suppose that for

some exogenous reason a high-deficit country decides to engage in long-term fiscal consolidation. If this leads to a recession, *ceteris paribus*, we might expect fiscal convergence to coincide with lower business cycle synchronization, at least over a short period of time.¹⁰ Alternatively, suppose that a high-deficit country decides to engage in fiscal consolidation and convergence simultaneously (e.g., during the drive to EMU); in this case, the effect goes the opposite way.

Accordingly, our default estimation is conducted with both OLS and instrumental variables. Our instrumental variables are associated with (cross-country differences in) the size and composition of public sector activity, since the public finance/political economy literature has shown these to be of relevance (e.g., Alesina and Perotti, 1997). Thus we use expenditure variables (such as government investment and non-wage consumption), as well as revenue variables (e.g., direct business and household taxes), all expressed as percentages of GDP. We check that our OLS and IV results are consistent and also show that our results are insensitive to the exact choice of instrumental variables.

3. Empirics

Benchmark Results on Fiscal Convergence and Business Cycle Synchronization

Our main results are presented in Tables 1 and 2. These display estimates of β , the estimated effect of fiscal divergence on business cycle synchronization. Robust standard errors (clustered by country-pair dyads) for the slope coefficients are presented beneath the coefficients in parentheses. One (two) asterisk(s) mark a coefficient that is significantly different from zero at the .05 (.01) confidence level. Table 1 presents OLS results, while our IV estimates are tabulated in Table 2.

The first row of each table present four benchmark estimates, one for each of our four default ways of measuring business cycle synchronization (arranged in columns). All four

coefficients are negative and distinguishable from zero with a high level of statistical confidence, for both OLS and IV. Moreover, the effects are economically important. A simple average of the four coefficients is -.034 for OLS. This implies that a reduction in fiscal divergence of (say) 2.5 percentage points – equal to one standard deviation in fiscal divergence – around its mean tends to raise the correlation of business cycles between a pair of countries, *ceteris paribus*, by around .085. Since the *average* correlation coefficient in the sample is around .3, this effect is neither trivial nor implausible. The IV results are approximately four times larger, and remain highly statistically significant. We try to be conservative in estimating the magnitude of our effect (especially when the model is so simple), but are reassured by the fact that OLS and IV deliver the same sign.

Succinctly, our initial results show that fiscal convergence tends to raise business cycle synchronization.

Sensitivity Analysis

Our benchmark estimates are derived from a simple setup; before taking them seriously, it is critical to establish their robustness. The remainder of tables 1 and 2 is devoted to sensitivity analysis. In particular, we explore the robustness of our finding to: a) differences in the estimation technique; b) differences in the sample; c) the inclusion of other controls; and d) different measures of fiscal policy. None of these alters our basic finding that fiscal convergence is associated with increased business cycle synchronization.

Our analysis examines pairs of countries over different periods of time. It is thus natural to add country pair-specific (dyadic) fixed effects. When we do so, β remains negative; its statistical significance falls somewhat, while its economic importance grows substantially with IV, and shrinks with OLS. Further, the fixed effects themselves are jointly insignificant at standard levels (except for two of the OLS equations). It seems that dyadic

fixed effects are not the reason for our finding of a negative β . Similarly, removing the decade (time-specific) fixed effects does not change our conclusion.

Our results seem insensitive to the exact handling of EMU observations. Dropping country-pairs that eventually joined EMU does not destroy our result; neither does adding a separate intercept for EMU dyads.¹¹ Our significantly negative β estimate also survives dropping observations from the first two decades of our sample, and dropping all observations with residuals lying more than two standard deviations from zero.

Frankel and Rose (1998) demonstrated that trade integration had the effect of raising business cycle synchronization. Might including trade in the regression reduce the effect of fiscal divergence? No. We add bilateral trade between countries i and j , normalized by the ratio of their GDPs, using four geographic determinants of the gravity model of bilateral trade as instrumental variables.¹² As expected, trade has a positive and usually significant effect on business cycle synchronization, but its presence makes little difference to the effect of fiscal divergence on business cycles.¹³ Our results are also not substantially affected when we include the four gravity variables directly in our equation.

Our next sensitivity analysis uses four different variants of the fiscal divergence regressor. First, we use the absolute value of the average (over time) gap between the two countries' budget positions, instead of using the average of the absolute value. Since budget balances are persistent, this variant delivers almost identical results to our benchmark. Second, we use (averages of absolute values of) primary budget deficits instead of total budget deficits; this delivers economically large results that remain statistically significant.¹⁴ Next, we use the gap between the two countries' actual government budget deficits and the Maastricht targets of a maximal 3% deficit/GDP ratio.¹⁵ Here we find weaker results; there is a statistically significant result only when we use unemployment. That is, cross-country

deviations from the Maastricht convergence criteria (and thus the Stability Pact's ceiling of 3% deficits) do not seem to have a substantial consistent effect on cycle synchronization.¹⁶

Towards the bottom of Table 1, we also use the standard deviation (computed over the ten years inside each decadal observation) of the gap between the two countries' budget/GDP ratios, in place of our default measure of fiscal divergence. OLS estimates indicate that variation in the budget deficit positions between the countries tends to lower their business cycle synchronization, which support our benchmark results.

It is comforting to us that OLS and IV estimates both sign β negatively. Nevertheless, we do not have vast confidence in our instrumental variables themselves.¹⁷ (Our first stage is tabulated in Appendix Table A3; while three of the instrumental variables are significant, the R^2 of the first stage is only .18.) Accordingly, we use four different sets of instrumental variables, combining measures of government revenue and expenditure series in different ways. We tabulate these results towards the bottom of Table 2. Both the economic and statistical significance of β varies depending on the estimator and measure of business cycle coherence. Still, all the estimates are negative, and the vast majority are significantly so.

We also check whether our finding (that fiscal divergence lowers business cycle synchronization) is immune to the addition of the average *level* of the government budget position. That is, we add AvgFiscal to our default equation and re-estimate. As can be seen from the bottoms of Tables 1 and 2, the effect of fiscal divergence on business cycle synchronization is unaffected when we control for the level of the average (cross-country) fiscal deficit; β remains economically and statistically significant.

Further Robustness Checks

Table 3 provides more sensitivity checks, using a number of different measures of business cycle synchronization. Rather than rely on a single measure of business cycle

coherence in the benchmark results, we used four measures in Tables 1 and 2. Still, there is no reason not to try others. The rows of Table 3 correspond to the estimated effect of fiscal divergence on fifteen further measures of business cycle synchronization. In different columns we provide OLS and IV estimates of β .

The first rows of Table 3 use industrial production (rather than GDP or unemployment) as the underlying measure of economic activity. Next, we follow Alesina, Barro and Tenreyro (2002) in measuring business cycle divergence. Alesina et al first construct the ratio of the two countries' log real GDP; they then regress that ratio on two of its lags and an intercept. The root mean squared error of the residual is their measure of business cycle divergence. Since a smaller number implies greater synchronization, we expect the sign of β to be reversed (compared with that of the correlation coefficient of de-trended business cycles). We construct Alesina-Barro-Tenreyro measures for log real GDP, log real GDP per capita, the unemployment rate, and the log of industrial production.

A third set of checks uses the Baxter-King (1999) band-pass filter to de-trend the underlying data (we use 2-8 years, corresponding to their 6-32 quarters). Finally, we switch to using underlying quarterly data rather than annual data. The finer frequency comes at a cost of a smaller data span.

None of the results in Table 3 alter our conclusions. Five of the six results using the Alesina-Barro-Tenreyro are positive as expected (though only two are statistically significant); the last is negative but statistically negligible. All the other checks work well in the sense that β remains significantly negative.

As an additional robustness check, we broadened the country coverage to include developing countries as well. This extended database covers 115 countries (hence it has maximum of 6555 [=115*114/2] bilateral country-pairs) for four decades. Since the unemployment rate and our instrumental variables are missing for many observations, we are

constrained to use only GDP and OLS,. The results are tabulated in Table A6. As in Tables 1 and 2, we find a negative and mostly significant relationship between fiscal divergence and business cycle synchronization (though when pair-specific effects are included, the coefficients lose significance).

Does the Average Budget Position have an Effect on Business Cycle Synchronization?

Thus far we have found strong evidence that persistent cross-country *differences* in government budget positions have a (negative) effect on the synchronization of their business cycles. An interesting but different question is whether the average (cross-country) *levels* of government budget positions also affect business cycle synchronization. We now investigate that issue.¹⁸

Table 4 contains estimates of the effect of the average (across pair of countries) government budget position on business cycle synchronization. Since we analyze two underlying concepts of economic activity (GDP and unemployment), three de-trending techniques (HP-filtering, differencing, and BK-filtered), two estimators (OLS and IV), and two budget concepts (total and primary), we provide twenty-four ($=2*3*2*2$) different point estimates and their standard errors.

We find little evidence that the total budget deficit has a consistent effect on business cycle synchronization. Seven of the twelve estimates are negative (two of those are statistically significant), while five are positive (non significant). All are small. However, all twelve of the coefficients for the primary budget effects are positive, three-quarters of them significantly so. We interpret the evidence as indicating that lower primary fiscal deficits (or higher primary surpluses) enhance business cycle synchronization. Further, when we use our extended sample of 115 countries, the average total budget balance has a positive and significant effect on synchronization, as can be seen from the last column of Table A6.

Still, we do not wish to over-interpret our findings. The average primary budget position is negatively correlated with our default measure of fiscal divergence (as can be seen from Table A2). When we include both fiscal divergence and the average primary budget position in our regressions, the former remains significantly negative (as can be seen from Tables 1 and 2), while the latter effect loses the horse-race (its effect becomes economically and statistically small, and varies across specifications). We have searched without success for a non-linear or interactive effect, and consider this to be a good topic for future research. That is, there is evidence that primary fiscal consolidation enhances business cycle synchronization, but it is weak. By way of comparison, there is strong evidence that fiscal divergence (of both total and primary balances) reduces the coherence of business cycles.¹⁹

4. Interpretation: Fiscal Irresponsibility tends to be Idiosyncratic

In section 3, we established that fiscal convergence seems to induce greater business cycle synchronization. If one takes the finding as given, the question remains: Why? We think the answer is that fiscal divergence tends to occur when one country runs a substantially and persistently higher budget deficit than other countries, and simultaneously creates fiscal shocks. That is, irresponsible fiscal policy (a persistently high deficit) coincides with idiosyncratic (fiscal) instability. When the budget deficit is closed (fiscal convergence), the fiscal shocks diminish; business cycles tend to become more synchronized. Succinctly, fiscal policy that is irresponsible is also fiscal policy that creates idiosyncratic shocks and thus macroeconomic volatility. This idea is both intuitive and consistent with the literature (e.g., Fatás and Mihov, 2003a, 2004).

Direct Evidence on Budgets and Macroeconomic Volatility via a Unilateral Panel

We now test our intuition in a straightforward way. We are interested in testing for a (negative) link between a country's average budget position and its business cycle volatility. Our evidence thus far has relied on bilateral data, comparing fiscal policy of pairs of countries to the synchronization of their business cycles. It is also possible to check this idea more directly using a unilateral (though still non-structural) approach. Accordingly, we gather a panel of annual data for 115 countries between 1960 and 2003 (with gaps), consisting of data on real GDP and the total government budget position (as a percentage of GDP; surpluses are positive, deficits negative).²⁰ We then de-trend the output data using both the HP and BK filters to create measures of business cycle fluctuations. We compare both the average absolute value of these business cycle deviations, and their volatility – proxied by the standard deviation (estimated for a country over time) – to the average level of the government's fiscal position. A negative relationship between the two indicates that smaller deficits or larger surpluses are associated with reduced business cycle volatility, consistent with our hypothesis.

We exploit our (country x year) panel of data in three different ways. First, we estimate panel regressions of the effect of the government budget position on business cycle deviations from trend at the annual frequency. Second, we split our 44-year data set into four eleven-year periods, so that each country contributes a maximum of four observations. Finally, we average over all 44 years, creating a single cross-section where each country contributes a single observation. For the first two cases, we estimate our models with differing sets of country- and time-specific fixed effects.

Our results are contained in Table 5. The top panel portrays annual results; the middle presents results estimated at the 11-year frequency; and the bottom shows cross-sectional results that average out the entire 44-year sample.

The point estimates from our annual results are all negative; a higher fiscal surplus (or lower deficit) is associated with smaller (in absolute value) business cycle deviations from trend. The results are statistically significant at conventional levels for twelve perturbations. When we shift to a lower frequency, we can examine both the average (over eleven years) of the absolute value of business cycle deviations, and the volatility of business cycles (the standard deviation of de-trended log real GDP). 19 of the 24 point estimates are negative, eight significantly so; none of the positive coefficients is economically or statistically large. Finally, when we examine a single cross-section of our countries, we again find that larger fiscal surpluses/smaller deficits are associated with lower business cycle volatility. At this very low frequency, all six point estimates are negative and half of them are significantly different from zero at standard confidence levels.

We do not consider this evidence to be overwhelming. Since we have essentially no structure in our empirical model, our results are suggestive rather than definitive. Still, we have not found evidence inconsistent with our hypothesis either in the literature or in our own empirical work. The hypothesis that larger fiscal deficits tend to be associated with greater business cycle volatility seems reasonable and awaits further scrutiny.

5. Conclusion

The motivation for this paper is simple. The criteria that make a currency area optimal were established long ago by Mundell and have essentially no intersection with the “Maastricht convergence” criteria used to govern the actual entry of countries into European Monetary Union. In this paper, we ask: does Maastricht indirectly overlap with Mundell?

The answer is positive. We find that fiscal convergence – similarity in the aggregate budget positions across countries – is systematically associated with enhanced business cycle synchronization. Fiscal convergence raises business cycle synchronization by eliminating

idiosyncratic fiscal shocks. We find evidence that reduced primary fiscal deficits (or higher surpluses) also increase the coherence of business cycles across countries. The Maastricht convergence process encouraged both fiscal convergence and reduced deficits for the Euro-12 during the run-up to EMU. Our results indicate that this fiscal convergence would have raised their business cycle coherence, making them better candidates for currency union. Even if not by design, Maastricht mimics Mundell!

There is a different (though consistent) interpretation of our results. Conventional wisdom tells us that national fiscal policy is the sole macroeconomic tool to smooth the business cycle when a country is hit by asymmetric shocks in a currency union. Yet the Maastricht criteria impose convergence of budget deficits at low levels. Consequently, Maastricht *could* reduce business cycle synchronization and increase volatility. In fact though, fiscal convergence seems to *increase* cycle synchronization by reducing volatile fiscal shocks.

If our finding is corroborated, it is of more than academic interest. The Maastricht criteria continue to govern future entry into the euro zone. Further, the Stability and Growth Pact continues, in principle, to constrain fiscal policy for the EU. If either or both of these institutions induce fiscal convergence, they indirectly enhance the desirability and sustainability of EMU. Two cheers!

Table 1: Effect of Fiscal Divergence on Business Cycle Synchronization, OLS

	GDP, HP-Filtered	GDP, Differenced	Unemployment, HP Filtered	Unemployment, Differenced
Benchmark	-.036** (.006)	-.024** (.005)	-.048** (.006)	-.028** (.005)
Pair-Specific Fixed Effects	-.022** (.008)	-.010 (.007)	-.034** (.009)	-.005 (.008)
Without Decade Effects	-.027* (.006)	-.0013** (.005)	-.032** (.006)	-.016** (.006)
Without EMU Pairs	-.039** (.007)	-.026** (.006)	-.050** (.007)	-.029** (.006)
Add EMU-Pairs Intercept	-.036** (.006)	-.024** (.005)	-.048** (.006)	-.028** (.005)
Last Half of Sample	-.055** (.009)	-.040** (.007)	-.073** (.010)	-.045** (.009)
Without 2σ Outliers	-.040** (.006)	-.024** (.004)	-.046** (.006)	-.028** (.006)
Add Trade/GDP Ratio	-.030** (.006)	-.018** (.005)	-.042** (.006)	-.022** (.005)
With Gravity Regressors	-.036** (.006)	-.023** (.005)	-.050** (.006)	-.028** (.005)
Regressor Variant	-.031** (.006)	-.023** (.005)	-.044** (.005)	-.027** (.005)
Primary Deficit Measure	-.009* (.004)	-.008* (.003)	-.010** (.003)	-.005** (.002)
Deviation from Maastricht Criteria	-.013 (.009)	-.012 (.007)	-.041** (.008)	-.023** (.007)
Std Dev (not mean) of Budget	-.084** (.014)	-.049** (.011)	-.077** (.015)	-.034* (.014)
With Average Fiscal Position	-.044** (.006)	-.026** (.005)	-.050** (.006)	-.027** (.006)
With Avg Primary Fiscal Position	-.040** (.008)	-.026** (.01)	-.057** (.008)	-.032** (.008)

Regressand is correlation coefficient (computed over decades) between country i and j de-trended series.

Coefficients recorded are effect of (average of absolute-value of differential of) government budget surplus/deficit, as percentage of GDP.

Robust standard errors (clustered by country-pair dyads) recorded in parentheses. Decade effects and constant included but not recorded.

Coefficients significantly different from zero at .05 (.01) level marked with one (two) asterisk(s)

OLS estimation unless noted.

Data set has maximum of $21*20/2=210$ country pairs for four decades (1964-73, 1974-83, 1984-93, 1994-2003).

Regressor variant is absolute value of average of differential (not average of absolute-value of differential). Std Dev is standard deviation over time of absolute-value of differential of government budget surplus/deficit, % GDP.

Table 2: Effect of Fiscal Divergence on Business Cycle Synchronization, IV

	GDP, HP-Filtered	GDP, Differenced	Unemployment, HP Filtered	Unemployment, Differenced
Benchmark	-0.16** (.04)	-0.11** (.03)	-0.15** (.04)	-0.11** (.03)
Pair-Specific Fixed Effects	-0.23** (.08)	-0.16** (.06)	-0.25** (.08)	-0.13* (.06)
Without Decade Effects	-0.16** (.04)	-0.13** (.04)	-0.16** (.04)	-0.11** (.03)
Without EMU Pairs	-0.12** (.03)	-0.07** (.02)	-0.11** (.03)	-0.07** (.02)
Add EMU-Pairs Intercept	-0.12** (.03)	-0.08** (.02)	-0.11** (.03)	-0.07** (.02)
Last Half of Sample	-0.19** (.04)	-0.11** (.03)	-0.16** (.04)	-0.11** (.03)
Without 2σ Outliers	-0.19** (.03)	-0.13** (.02)	-0.19** (.04)	-0.14** (.03)
Add Trade/GDP Ratio (Gravity IV)	-0.09** (.02)	-0.05** (.01)	-0.06** (.02)	-0.04* (.02)
With Gravity Regressors	-0.08** (.02)	-0.05** (.02)	-0.06** (.02)	-0.03 (.02)
Regressor Variant	-0.14** (.03)	-0.10** (.03)	-0.14** (.03)	-0.09** (.02)
Primary Deficit Measure	-0.15** (.04)	-0.13** (.03)	-0.19** (.05)	-0.10** (.03)
Deviation from Maastricht Criteria	-0.03 (.04)	-0.05 (.03)	-0.09* (.04)	-0.06* (.03)
With Average Fiscal Position	-0.15** (.04)	-0.12** (.03)	-0.16** (.04)	-0.11** (.03)
With Avg Primary Fiscal Position	-0.14** (.05)	-0.09** (.03)	-0.12** (.04)	-0.10** (.03)
IV Variant 1	-0.16** (.05)	-0.12** (.04)	-0.29** (.06)	-0.25** (.06)
IV Variant 2	-0.14** (.04)	-0.08* (.03)	-0.08** (.03)	-0.03 (.03)
IV Variant 3	-0.18** (.06)	-0.10* (.05)	-0.12** (.05)	-0.07* (.03)
IV Variant 4	-0.19** (.06)	-0.15** (.05)	-0.20** (.05)	-0.15** (.04)

Regressand is correlation coefficient (computed over decades) between country i and j de-trended series.

Coefficients recorded are effect of (average of absolute-value of differential of) government budget surplus/deficit, as percentage of GDP. Robust standard errors (clustered by country-pair dyads) recorded in parentheses. Decade effects and constant included but not recorded. Coefficients significantly different from zero at .05 (.01) level marked with one (two) asterisk(s).

Instrumental Variable estimation unless noted. IVs include: a) government non-wage consumption/GDP; b) government investment/GDP; c) direct business taxes/GDP; and d) direct household taxes/GDP. IVs are average of absolute-value of cross-country differentials.

Data set has maximum of $21 \times 20 / 2 = 210$ country pairs for four decades (1964-73, 1974-83, 1984-93, 1994-2003).

Regressor variant is absolute value of average of differential (not average of absolute-value of differential). Std Dev is standard deviation over time of absolute-value of differential of government budget surplus/deficit, % GDP.

IV Variant 1: a) government non-wage consumption/GDP; b) government investment/GDP; c) effective labor taxes as percentage of labor costs; and d) indirect taxes/GDP. Variant 2: a) government social benefits/GDP; b) government wages/GDP; and c) direct business taxes/GDP. Variant 3: a) direct household taxes/GDP; b) indirect taxes/GDP; and c) direct business taxes/GDP. Variant 4: a) government non-wage consumption/GDP; b) government wages/GDP; and c) government investment/GDP.

Gravity regressors are: 1) log distance; 2) log product land area; 3) common land border dummy; 4) common language dummy.

Table 3: Fiscal Divergence and Different Measures of Business Cycle Synchronization

	OLS	IV
Industrial Production, HP-Filtered	-.03** (.005)	-.08** (.02)
Industrial Production, Differenced	-.014** (.005)	-.06** (.02)
GDP, Alesina-Barro-Tenreyro	.0005** (.0001)	.0019** (.0005)
GDP p/c, Alesina-Barro-Tenreyro	.0004** (.0001)	.0018** (.0005)
Unemployment, Alesina-Barro-Tenreyro	.03** (.01)	.027 (.026)
Industrial Production, Alesina-Barro-Tenreyro	.0009** (.0002)	.0046** (.0010)
GDP, Baxter-King	-.028** (.005)	-.15** (.03)
Unemployment, Baxter-King	-.030** (.005)	-.11** (.03)
Industrial Production, Baxter-King	-.023** (.005)	-.07** (.02)
Quarterly GDP, HP-Filtered	-.012* (.005)	-.09** (.02)
Quarterly GDP, Differenced	-.023** (.007)	-.12** (.03)
Quarterly Industrial Production, HP-Filtered	-.016** (.003)	-.05** (.02)
Quarterly Industrial Production, Differenced	-.016** (.004)	-.05** (.02)
Quarterly GDP, Alesina-Barro-Tenreyro	.0003** (.0001)	.0022** (.0005)
Quarterly Industrial Production, Alesina-Barro-Tenreyro	.0008** (.0001)	.0013* (.0006)

IVs include: a) government non-wage consumption/GDP; b) government investment/GDP; c) direct business taxes/GDP; and d) direct household taxes/GDP. IVs are average of absolute-value of cross-country differentials.

Coefficients recorded are effect of (average of absolute-value of differential of) government budget surplus/deficit, as percentage of GDP.

Coefficients significantly different from zero at .05 (.01) level marked with one (two) asterisk(s)

Data set has maximum of $21 \times 20 / 2 = 210$ country pairs for four decades (1964-73, 1974-83, 1984-93, 1994-2003).

Decade effects and constant included but not recorded.

Alesina-Barro-Tenreyro measure is root mean squared error of residual from AR(2) of log ratios (lower => greater co-movement).

Robust standard errors (clustered by country-pair dyads) recorded in parentheses.

Table 4: Average Budget Positions and Business Cycle Synchronization

	GDP, HP	GDP, Diff.	GDP, BK	Unem, HP	Unem, Diff	Unem, BK
Total Budget (% GDP) IV	-.04 (.02)	-.00 (.02)	-.04 (.02)	.00 (.02)	.00 (.02)	-.01 (.02)
Total Budget (% GDP) OLS	-.02* (.01)	-.00 (.01)	-.02** (.01)	.01 (.01)	.01 (.01)	.01 (.01)
Primary Budget (% GDP) IV	.11** (.03)	.09** (.03)	.12** (.03)	.10** (.04)	.03 (.03)	.07** (.03)
Primary Budget (% GDP) OLS	.03** (.01)	.02* (.01)	.05** (.01)	.02 (.01)	.01 (.01)	.03** (.01)

IVs include: a) government non-wage consumption/GDP; b) government investment/GDP; c) direct business taxes/GDP; and d) direct household taxes/GDP. IVs are average of absolute-value of cross-country differentials.

Coefficients recorded are effect of cross-country average level of total/primary government budget surplus/deficit, as percentage of GDP.

Coefficients significantly different from zero at .05 (.01) level marked with one (two) asterisk(s)

Data set has maximum of $21 \times 20 / 2 = 210$ country pairs for four decades (1964-73, 1974-83, 1984-93, 1994-2003).

Decade effects and constant included but not recorded.

Robust standard errors (clustered by country-pair dyads) recorded in parentheses.

Table 5: Government Budgets and Business Cycle Volatility**Annual Panel Results**

	Hodrick-Prescott	Baxter-King	Differenced
Common intercept	-0.057** (.014)	-0.050** (.011)	-0.080** (.016)
Year Effects	-0.038** (.014)	-0.040** (.011)	-0.072** (.017)
Country Effects	-0.058** (.015)	-0.042** (.012)	-0.066** (.019)
Year and Country Effects	-0.038** (.015)	-0.032** (.012)	-0.060** (.019)
Observations	3371	2944	3308

Regressands are the absolute value of detrended log real GDP, either a) Hodrick-Prescott filtered, b) Baxter-King band-pass filtered or c) differenced (country specific mean growth removed from differences before taking absolute values). Regressor is government budget, % GDP.

Long-Run Panel Results (for data averaged over 11-year periods)

	Standard deviation			Mean absolute value		
	Hodrick-Prescott	Baxter-King	Differenced	Hodrick-Prescott	Baxter-King	Differenced
Common intercept	-0.062* (.035)	-0.067** (.033)	-.083 (.057)	-0.070** (.033)	-.051* (.027)	-.115** (.040)
Period Effects	-.039 (.036)	-.052 (.033)	-.068 (.059)	-.046 (.036)	-.040 (.027)	-.111** (.044)
Country Effects	-.033 (.048)	-.029 (.046)	.010 (.072)	-.076** (.038)	-.032 (.035)	-.073* (.043)
Period, Country Effects	.012 (.047)	.000 (.046)	.039 (.071)	-.032 (.040)	-.010 (.035)	-.072 (.047)
Observations	365	349	364	368	354	365

Regressands are either a) standard deviation or b) mean absolute value of log real GDP, either a) Hodrick-Prescott filtered, b) Baxter-King band-pass filtered or c) differenced (country specific mean growth removed from differences before taking absolute values) over four 11-year long periods. Regressor is mean of government budget, % GDP.

Cross-Sectional Results (for data averaged over entire sample)

	Standard deviation			Mean absolute value		
	Hodrick-Prescott	Baxter-King	Differenced	Hodrick-Prescott	Baxter-King	Differenced
Intercept	-.064 (.070)	-.117** (.047)	-.139* (.073)	-.025 (.050)	-.058* (.030)	-.077 (.049)
Observations	115	115	115	115	115	115

Regressands are either a) standard deviation or b) mean absolute value of log real GDP, either a) Hodrick-Prescott filtered, b) Baxter-King band-pass filtered or c) differenced (country specific mean growth removed from differences before taking absolute values) over entire period, 1960-2003. Regressor is the mean of government budget, % GDP.

Notes for all blocks.

Coefficients from OLS regressions, multiplied by 100. Robust standard errors (clustered by country) in parentheses (also multiplied by 100).

Coefficient significantly different from zero at .01 (.05) marked by two (one) asterisks.

Based on annual data for 115 countries, 1960-2003 (with gaps).

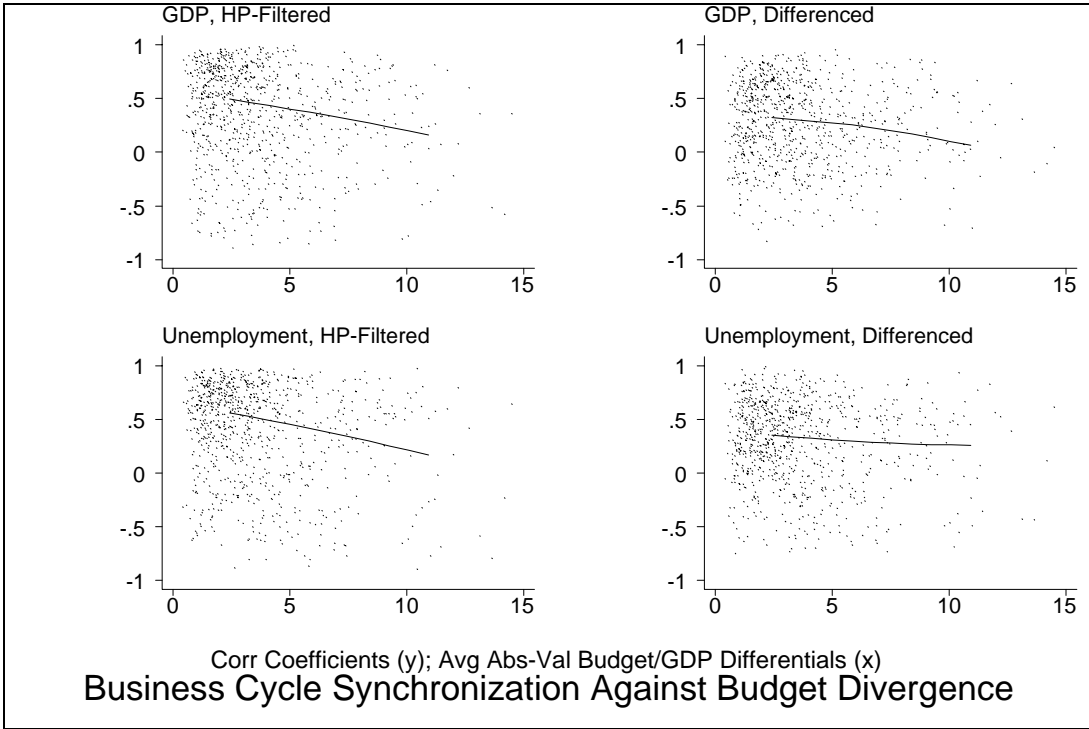


Figure 1: Simple Scatterplots of Key Variables, 1964-2003

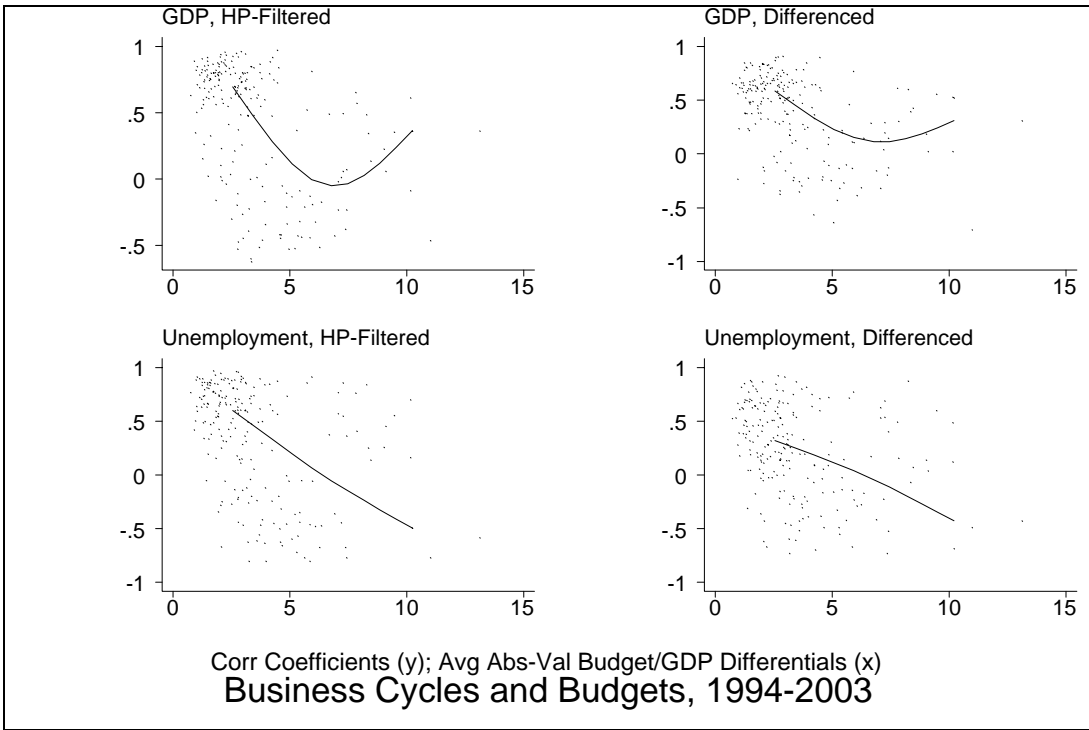


Figure 2: Scatterplots for Most Recent Decade

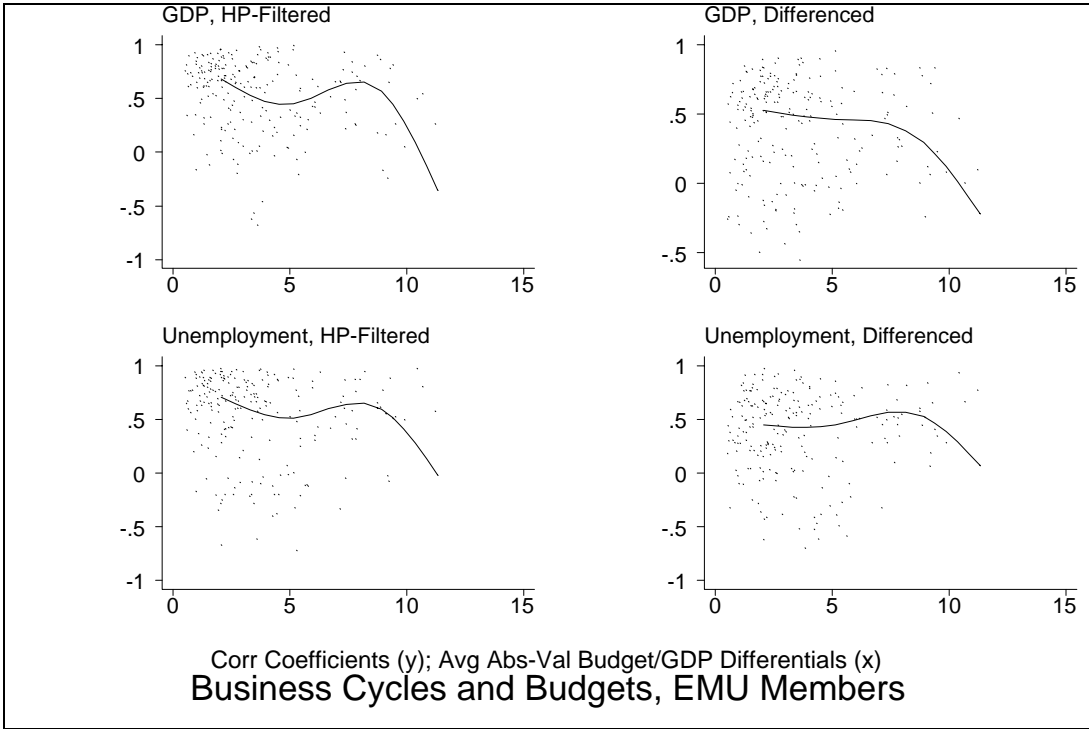


Figure 3: Scatterplots for the Ins, 1964-2003

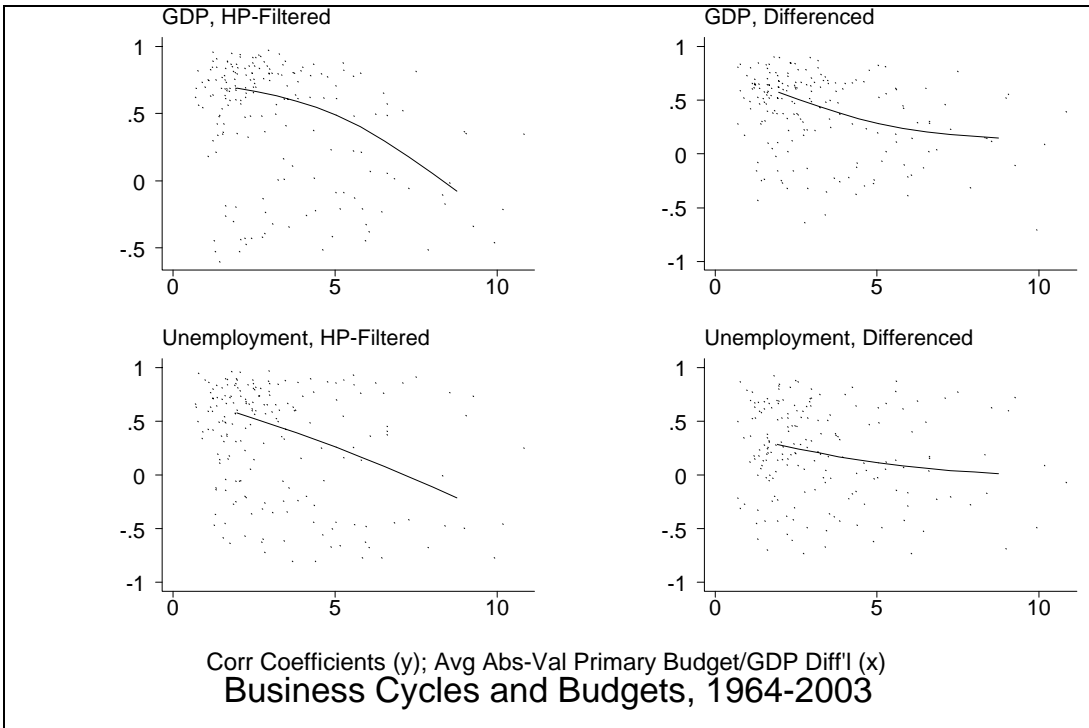


Figure 4: Primary Fiscal Divergence

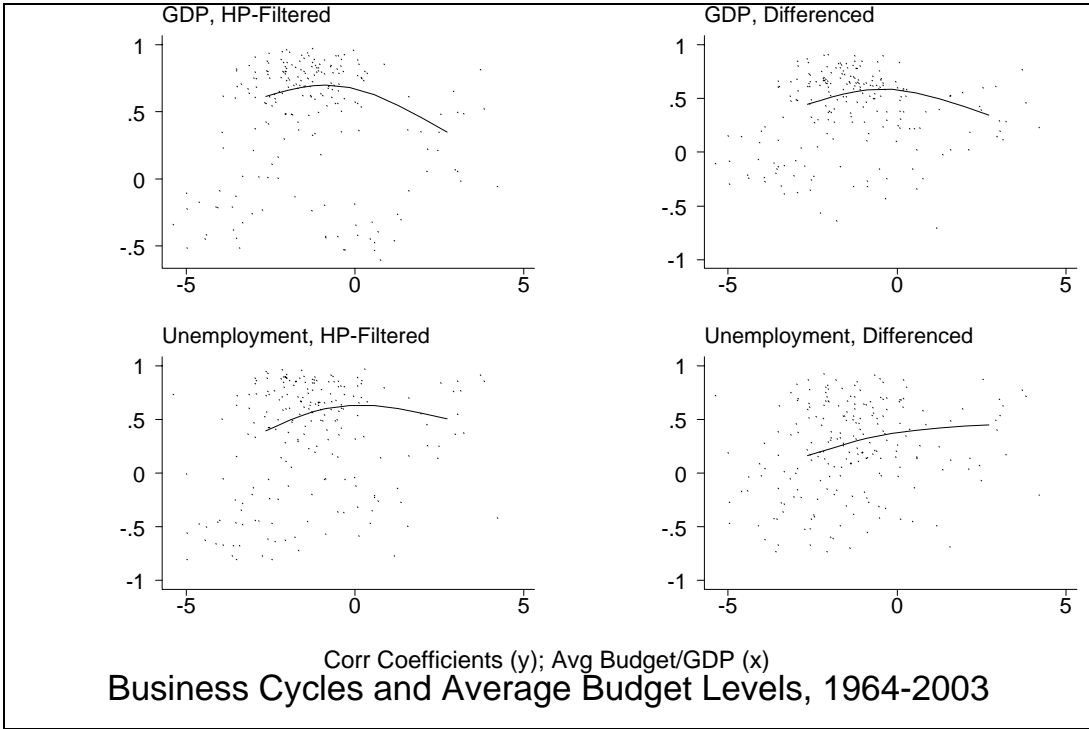


Figure 5: Average Total Government Budget Level

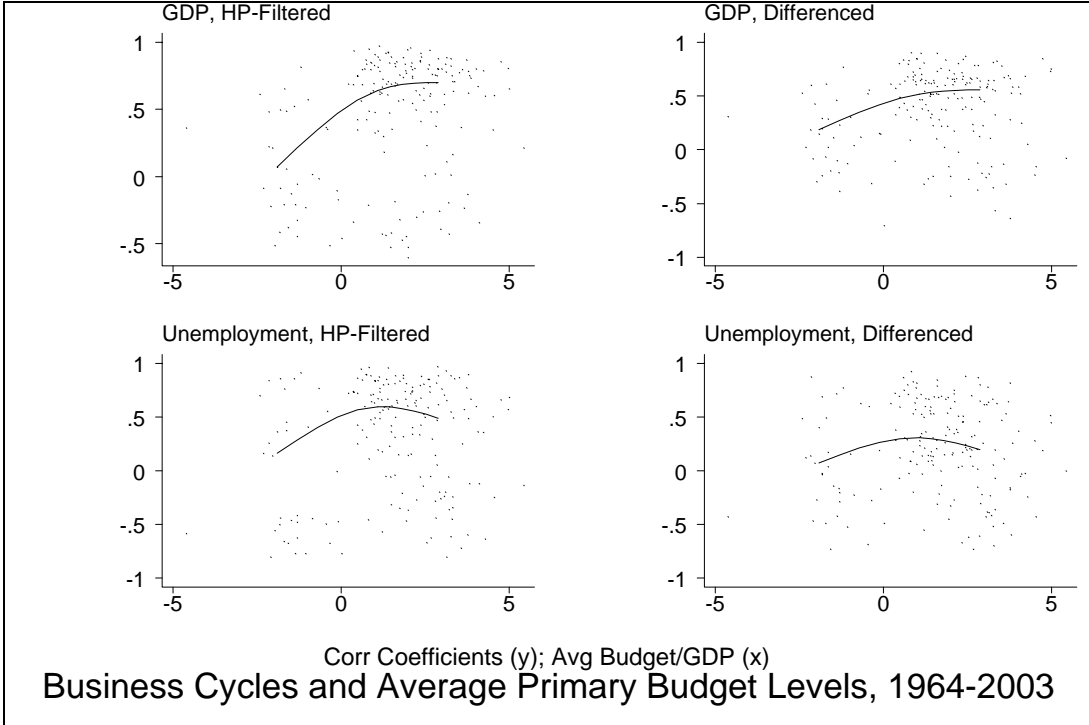


Figure 6: Average Primary Government Budget Level

Table A1a: Countries in Default OECD Sample

Australia	Austria	Belgium	Canada	Denmark	Finland
France	Germany	Greece	Ireland	Italy	Japan
Netherlands	Norway	New Zealand	Portugal	Spain	Sweden
Switzerland	UK	USA			

Table A1b: Additional Countries in Wide Sample

Argentina	Bahrain	Bangladesh	Barbados	Belarus	Belize
Bhutan	Bolivia	Botswana	Brazil	Bulgaria	Burk. Faso
Burundi	Cameroon	Chile	China	Colombia	Congo
Costa Rica	Croatia	Cyprus	Czech Rep.	Dominican R.	Egypt
El Salvador	Estonia	Fiji	Ghana	Guatemala	Guyana
Haiti	Honduras	Hungary	Iceland	India	Indonesia
Iran	Israel	Jamaica	Jordan	Kazakhstan	Kenya
Korea	Kyrgyz Re.	Latvia	Lesotho	Lithuania	Madagascar
Malawi	Malaysia	Malta	Mauritius	Mexico	Mongolia
Morocco	Myanmar	Nepal	Nicaragua	Nigeria	Oman
Pakistan	Panama	Pap. N. Guinea	Paraguay	Peru	Philippines
Poland	Romania	Russia	Rwanda	Saudi Arabia	Senegal
Seychelles	Sierra Leone	Singapore	Slovak Republic	Slovenia	South Africa
Sri Lanka	St. Lucia	St. Vin. & Gren.	Swaziland	Syria	Thailand
Tunisia	Turkey	Uganda	Ukraine	Uruguay	Venezuela
Vietnam	Yemen	Zambia	Zimbabwe		

Table A2: Descriptive Statistics

	Obs.	Avg.	Std. Dev.	Min.	Max.	Corr.
Correlation Coefficient, GDP, HP-Filtered	840	.36	.44	-.88	.99	-.23
Correlation Coefficient, GDP, Differenced	840	.27	.37	-.83	.96	-.13
Correlation Coefficient, Unemployment, HP-Filtered	840	.39	.45	-.89	.98	-.31
Correlation Coefficient, Unemployment, Differenced	840	.29	.39	-.74	.99	-.22
Government Budget/GDP Divergence	840	3.65	2.52	.41	14.5	n/a
Average Government Budget/GDP Level	840	-2.77	2.47	-11.9	4.2	-.14
Primary Government Budget/GDP Divergence	619	3.44	5.89	.14	101.8	.48
Average Primary Government Budget/GDP Level	617	-.03	2.04	-6.63	5.43	-.41
Gov't Budget/GDP Divergence, Maastricht Deviation	840	1.91	2.23	0	9.82	.70
Government non-wage consumption/GDP Divergence	800	2.46	1.76	.15	9.89	-.16
Government investment/GDP Divergence	722	1.00	.66	.06	4.01	.08
Direct business taxes/GDP Divergence	638	1.27	.86	.10	5.05	.26
Direct household taxes/GDP Divergence	602	5.25	4.36	.17	21.95	.01
Trade/GDP Ratio	840	.49	.77	.01	7.21	-.07
Inflation Divergence	840	3.48	3.18	.36	18.2	.10
Long Interest Rate Divergence	742	2.55	2.44	.08	16.3	.21
Government Debt/GDP Divergence	592	28.0	20.2	.58	106.8	.38
Standard Deviation of Exchange Rate	840	.12	.09	.003	.58	.02
Maximum Change of Exchange Rate	840	.28	.13	.02	.67	.13

Corr. is simple correlation coefficient between variable and government budget/GDP.

Table A3: First Stage

Government non-wage consumption/GDP	-.23** (.06)
Government investment/GDP	.44** (.14)
Direct Business Taxes/GDP	.44** (.11)
Direct Household Taxes/GDP	-.02 (.02)

Regressand is (average of absolute-value of differential of) government budget surplus/deficit, as percentage of GDP. Coefficients estimated via OLS. Standard errors recorded in parentheses.

All regressors are average of absolute-value of cross-country differentials.

Data set has maximum of $21 \times 20 / 2 = 210$ country pairs for four decades (1964-73, 1974-83, 1984-93, 1994-2003).

Decade effects and constant included but not recorded.

Coefficients significantly different from zero at .05 (.01) level marked with one (two) asterisk(s)

Table A4: Business Cycle Synchronization in Different Fiscal Regimes**A: GDP HP-Filtered, Total Deficit**

	Surplus/Deficit < 1%	Deficit in (1,6)%	Deficit > 6%
Surplus/Deficit < 1%	.30 (85)		
Deficit in (1,6)%	.35 (278)	.42 (293)	
Deficit > 6%	.07 (32)	.35 (136)	.46 (16)

B: GDP Differenced Total Deficit

	Surplus/Deficit < 1%	Deficit in (1,6)%	Deficit > 6%
Surplus/Deficit < 1%	.30 (85)		
Deficit in (1,6)%	.26 (278)	.29 (293)	
Deficit > 6%	.13 (32)	.27 (136)	.38 (16)

C: Unemployment HP-Filtered Total Deficit

	Surplus/Deficit < 1%	Deficit in (1,6)%	Deficit > 6%
Surplus/Deficit < 1%	.39 (85)		
Deficit in (1,6)%	.35 (278)	.47 (293)	
Deficit > 6%	-.11 (32)	.38 (136)	.49 (16)

D: Unemployment Differenced Total Deficit

	Surplus/Deficit < 1%	Deficit in (1,6)%	Deficit > 6%
Surplus/Deficit < 1%	.38 (85)		
Deficit in (1,6)%	.25 (278)	.32 (293)	
Deficit > 6%	.02 (32)	.30 (136)	.38 (16)

E: GDP HP-Filtered, Primary Surplus/Deficit

	Primary Surplus > 1.5%	Balance in (-1.5,1.5)%	Primary Deficit > 1.5%
Primary Surplus > 1.5%	0.45 (62)		
Balance in (-1.5,1.5)%	0.45 (145)	0.44 (132)	
Primary Deficit > 1.5%	0.18 (56)	0.39 (165)	0.35 (57)

F: GDP Differenced, Primary Surplus/Deficit

	Primary Surplus > 1.5%	Balance in (-1.5,1.5)%	Primary Deficit > 1.5%
Primary Surplus > 1.5%	0.34 (62)		
Balance in (-1.5,1.5)%	0.39 (145)	0.34 (132)	
Primary Deficit > 1.5%	0.19 (56)	0.28 (165)	0.25 (57)

G: Unemployment HP-Filtered, Primary Surplus/Deficit

	Primary Surplus > 1.5%	Balance in (-1.5,1.5)%	Primary Deficit > 1.5%
Primary Surplus > 1.5%	0.27 (62)		
Balance in (-1.5,1.5)%	0.39 (145)	0.53 (132)	
Primary Deficit > 1.5%	0.15 (56)	0.41 (165)	0.36 (57)

H: Unemployment Differenced, Primary Surplus/Deficit

	Primary Surplus > 1.5%	Balance in (-1.5,1.5)%	Primary Deficit > 1.5%
Primary Surplus > 1.5%	0.15 (62)		
Balance in (-1.5,1.5)%	0.32 (145)	0.40 (132)	
Primary Deficit > 1.5%	0.17 (56)	0.30 (165)	0.26 (57)

Deficits expressed as percentages of national GDP. Number of observations recorded in parentheses.

Data tabulated are average correlations of business cycles. Thus for the (85) cases where both countries are in total surplus or have deficits < 1% GDP, the average correlation of de-trended GDP is .30.

Table A5: Different Criteria and Business Cycle Synchronization

OLS	GDP, HP-Filtered	GDP, Differenced	Unemployment, HP Filtered	Unemployment, Differenced
Inflation	-.01 (.01)	-.01 (.01)	-.02** (.01)	-.02** (.01)
Long Interest Rate	-.02* (.01)	-.01 (.01)	-.03** (.01)	-.02** (.01)
Government Debt/GDP	-.001 (.001)	-.001 (.001)	-.004** (.001)	-.003** (.001)
Standard Deviation of Exchange Rate	-.43 (.23)	-.36 (.21)	-.92** (.23)	-.77** (.21)
Maximum Change of Exchange Rate	-.42* (.18)	-.40** (.14)	-.61** (.16)	-.53** (.14)

IV	GDP, HP-Filtered	GDP, Differenced	Unemployment, HP Filtered	Unemployment, Differenced
Inflation	-.04** (.06)	-.09* (.04)	-.13 (.07)	-.05 (.04)
Long Interest Rate	-.13** (.04)	-.13** (.04)	-.18** (.05)	-.09* (.04)
Government Debt/GDP	-.006 (.004)	-.007* (.004)	-.008 (.005)	-.005 (.004)
Standard Deviation of Exchange Rate	-4.99** (1.56)	-4.83** (1.26)	-6.51** (1.95)	-3.42* (1.32)
Maximum Change of Exchange Rate	-1.93** (.57)	-1.79** (.43)	-2.49** (.67)	-1.53** (.52)

Regressand is correlation coefficient (computed for individual decades of annual data) between country i and j de-trended series.

Coefficients recorded are effect of (average of absolute-value of differential of) variable tabulated in left column. Coefficients significantly different from zero at .05 (.01) level marked with one (two) asterisk(s)

IVs include: a) government non-wage consumption/GDP; b) government investment/GDP; c) direct business taxes/GDP; and d) direct household taxes/GDP. IVs are average of absolute-value of cross-country differentials.

Data set has maximum of $21 \times 20 / 2 = 210$ country pairs for four decades (1964-73, 1974-83, 1984-93, 1994-2003).

Decade effects and constant included but not recorded.

Robust standard errors (clustered by country-pair dyads) recorded in parentheses.

Table A6: Fiscal Divergence and Business Cycle Synchronization; OLS on a Wide Panel

	Benchmark Effect of Fiscal Divergence	Pair-Specific Fixed Effects	With Average Fiscal Position	Only Average Fiscal Position
HP-Filtered	-0.005** (.001)	-.001 (.003)	-0.004** (.001)	.007** (.002)
First-Differenced	-0.002** (.001)	.001 (.002)	-0.001 (.001)	.005** (.001)

Regressand is correlation coefficient (computed over decades) between country i and j de-trended series.

Coefficients recorded are effect of (average of absolute-value of differential of) government budget surplus/deficit, as percentage of GDP.

OLS estimation. Robust standard errors (clustered by country-pair dyads) in parentheses.

Decade effects and constant included but not recorded. 14,961 observations

Coefficients significantly different from zero at .05 (.01) level marked with one (two) asterisk(s)

Data Sources

OECD Economic Outlook (Annual series): Consumer Price Index; Direct Taxes, Households; Direct Taxes, Business; Fixed Investment, Government, Value; Government Consumption, Excluding Wages; Government Consumption, Wages; Gross Domestic Product (Market prices), Value; Gross Domestic Product (Market prices), Volume; Gross Government Debt, % GDP; Indirect Taxes; Interest Rate, Long-Term; Interest Rate, Short-Term; Primary Government Balance, Cyclically Adjusted, % Potential GDP; Social Benefits Paid by Government; Unemployment Rate.

OECD Quarterly National Accounts: Gross Domestic Product, Volume.

OECD Tax Database (Annuals series): Income tax plus employee and employer contribution less cash benefits (as % of labor costs), one-earner family with two children.

IMF International Financial Statistics (Annual series): General Government Deficit (-) or Surplus; Gross Domestic Product, Volume and Value (for developing countries included in the wide sample); Industrial Production (Volume). (Quarterly series): Industrial Production. Quarterly data on: Industrial Production (Volume). (Monthly series): Exchange rate (National Currency per US Dollar, line RF)

IMF Direction of Trade Statistics (Annual series for 1980-2003): Exports, f.o.b.; Imports, c.i.f.

Frankel-Rose (1998) (Annual series for 1960-1979): Exports, f.o.b.; Imports, c.i.f. starting in 1980.

EC AMECO database (Annual series): Net lending (+) or net borrowing (-): general government, Percentage of GDP at market prices.

Reinhart-Rogoff (2004) (Monthly series): Parallel or Black Market Exchange Rate

References

- Alesina, Alberto F. and Robert J. Barro (2002) "Currency Unions" *Quarterly Journal of Economics* CXVII-2 May, 409-436.
- Alesina, Alberto F., Robert J. Barro, and Silvana Tenreyro (2002) "Optimal Currency Areas" in *NBER Macroeconomics Annual* (M. Gertler and K. Rogoff, eds; MIT Press, Cambridge).
- Alesina, Alberto F. and Roberto Perotti (1997) "Fiscal Adjustments in OECD Countries: Composition and Macroeconomic Effects" *IMF Staff Papers* 44-2, 210-248.
- Baxter, Marianne and Robert G. King (1999) "Measuring Business Cycles: Approximate Band-Pass Filters for Economic Time Series" *The Review of Economics and Statistics* 81(4), 575-593.
- Canova, Fabio, Matteo Ciccarelli and Eva Ortega (2004) "Similarities and Convergence in G-7 Cycles", CEPR DP No. 4534.
- Clark, Todd E. and Eric van Wincoop (2001) "Borders and Business Cycles" *Journal of International Economics* 55-1, 59-85.
- Darvas, Zsolt and György Szapáry (2004) "Business Cycle Synchronization in the Enlarged EU: Comovements in the New and Old Members", MNB Working Paper 2004/1.
- Fatás, Antonio and Ilian Mihov (2003a) "The Case for Restricting Fiscal Policy Discretion" *Quarterly Journal of Economics* 118-4, 1419-1447.
- Fatás, Antonio and Ilian Mihov (2003b) "Fiscal Policy and EMU: Challenges of the Early Years" in *EMU and Economic Policy in Europe* (Buti and Sapir, eds; Edward Elgar).
- Fatás, Antonio and Ilian Mihov (2004) "The Macroeconomic Effects of Fiscal Rules in the US States" CEPR DP No. 4372.
- Frankel, Jeffrey A. and Andrew K. Rose (1998) "The Endogeneity of the Optimum Currency Area Criteria" *Economic Journal* 108 (449), 1009-1025
- Gavin, Michael and Roberto Perotti (1997) "Fiscal Policy in Latin America" *NBER Macroeconomics Annual* 12, 11-61.
- Gerlach, H.M.S. (1988) "World business cycles under fixed and flexible exchange rates" *Money Market and Banking* 20 (4), 621-632.
- Kose, A.A., E.S. Prasad and M. Terrones (2003) "How does globalization affect the synchronization of business cycles?" *American Economic Review* 93-2, 57-62.
- Lumsdaine, R.L. and E.S. Prasad (1997) Identifying the common component on international economic fluctuations. NBER Working Paper No. 5984.
- Mundell, Robert (1961) "A Theory of Optimum Currency Areas" *American Economic Review* LI, 657-665.
- Reinhart, Carmen M. and Kenneth S. Rogoff (2004) "The Modern History of Exchange Rate Arrangements: A Reinterpretation", *Quarterly Journal of Economics* CXIX-1, 1-48.

Endnotes

¹ In EU terminology, EMU technically refers to Economic and Monetary Union, which is different from the euro area. All EU countries are members of the Economic and Monetary Union, but only twelve members are currently members of the euro area. In the academic literature, EMU generally refers to the European monetary union. In this paper we follow conventional practice and use EMU to refer to the euro area.

² We ignore the design of monetary institutions and policies for the time being. These are relevant to both the Maastricht Treaty and Optimum Currency Area considerations, but are not intrinsically either national or international. In any case, there is considerable overlap between the two sets of criteria in this respect.

³ Table 7A of “Cyclical Adjustment of Budget Balances” produced by ECFIN, EC, Spring 2005, available at http://europa.eu.int/comm/economy_finance/indicators/general_government_data/2005/cabb_spring2005en.pdf

⁴ Again, we use Table 7A of “Cyclical Adjustment of Budget Balances.” For further analysis, see Fatás and Mihov (2003b).

⁵ We also briefly examine effects of other Maastricht criteria, such as those for inflation, exchange rates, etc.

⁶ See, e.g., Gavin and Perotti (1997).

⁷ Ditto, Table 7A of “Cyclical Adjustment of Budget Balances.”

⁸ We see no reason why our analysis could not be extended to other countries, though missing observations are likely to be a serious issue.

⁹ In practice there are often gaps in our data set.

¹⁰ We try to minimize such issues by estimating our business cycle synchronizations using decades, but the issue remains.

¹¹ We do not attempt to estimate our effect using only observations for country-pairs in EMU, since there are few observations (55 dependent dyadic observations, for only half of the last decade).

¹² The four instrumental variables are: 1) the natural logarithm of the great circle bilateral distance between the two countries; 2) the log of the product of the countries’ land areas; 3) a common land border dummy; and 4) a common language dummy.

¹³ This is unsurprising since trade is almost uncorrelated with fiscal divergence.

¹⁴ We use the OECD’s measure “Primary Government Balance, Cyclically Adjusted, % Potential GDP.”

¹⁵ We formalize this as follows. If both countries meet the 3% target, the gap between them is zero. If one meets the criterion and one has a deficit of say 4% of GDP, the gap is 1% (of GDP). If neither meets the criteria, one country’s deficit is 5% and the other’s is 6%, the difference between them is also 1% (of GDP).

¹⁶ This may be unsurprising, since there is little reason to think that convergence to 3% should have a different effect on business cycle synchronization than convergence to another deficit level.

¹⁷ For instance, we cannot exclude the possibility of simultaneity from any available fiscal aggregate.

¹⁸ We have already shown in Tables 1 and 2 that controlling for the average level of the government budget position (i.e., including AvgFiscal in our regressions) has little effect on the economic or statistical significance of β .

¹⁹ We have also briefly investigated the effects of other Maastricht criteria on business cycle synchronization; estimates appear in Table A5. There is some evidence that exchange rate volatility, and divergence in inflation, long interest rates, and government debt levels all tend to lower business cycle synchronization. However, none of the effects is particularly strong or consistent. We view this as an area worthy of future research.

²⁰ We do not know of a source that systematically provides primary fiscal positions for countries outside the OECD.